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Lecture 3 : Instrumental Variables

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Reading for lecture 3

- ▶ In the textbook (Wooldridge 2013): Chapter 15
- ▶ In Mostly Harmless Econometrics: Chapter 4

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Revision of Last' Lectures main points

Our objective: Obtain causal effect of policy change

- ▶ In many economic models the **exogeneity assumption is violated**: that is at least one 'explanatory' variable is endogenous (explanatory variables are correlated with the error), or jointly determined with the 'dependent' variable.
 $y = \mathbf{x}\beta + u, E(u|\mathbf{x}) \neq 0.$

- ▶ There are four potential sources of endogeneity that we talked about last week:

(i) omitted variables

(ii) selection bias

(iii) measurement error

(iv) reverse causality or simultaneity

- ▶ In each of these four cases, the OLS is not capable of delivering unbiased/consistent parameter estimates.

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Review of IV

- ▶ RCT is one way to deal with endogeneity.
- ▶ Another way to deal with endogeneity is using **instrumental variables**.
- ▶ Typically, the point of IV is to allow causal inference in a non-experimental setting.

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This example AGAIN: Ability Bias in the Returns to Education

- ▶ Labour economists have been studying returns to education for a very long time.

- ▶ The correctly specified model that is analyzed is:

$$Y_i = a + \rho S_i + \gamma A_i + n_i(*)$$

- ▶ Y_i = log of earnings.
- ▶ S_i = schooling measured in years.
- ▶ A_i = individual ability.
- ▶ Typically the econometrician cannot observe A_i .
- ▶ Suppose you therefore estimate the short regression:

$$Y_i = a + \rho S_i + h_i$$

- ▶ where

$$h_i = \gamma A_i + n_i$$

This example AGAIN: Ability Bias in the Returns to Education

- ▶ The OLS estimator for ρ in this simple case is:

$$\hat{\rho} = \frac{\text{Cov}(Y, S)}{\text{Var}(S)} \quad (1)$$

- ▶ Plugging in the true model for Y

$$\hat{\rho} = \rho + \gamma \frac{\text{Cov}(A, S)}{\text{Var}(S)} \quad (2)$$

- ▶ This is the classic ability bias if $\gamma > 0$ and $\text{Cov}(A, S) > 0$, thus the coefficient on schooling in the short regression would be **upward biased**.

How IV Can be Used to Obtain Unbiased and Consistent Estimates?

- ▶ How can we estimate the true θ if ability is unobserved?
 - ▶ Use an instrument Z .
 - ▶ 2 important conditions for a valid IV:
 - ▶ 1) $Cov(S_i, Z_i) \neq 0$ (1st stage exists).
(hint: S_i is the endogenous regressor)
 - ▶ 2) $Cov(Z_i, h_i) = 0$ (exclusion restriction: Z is uncorrelated with any other determinants of the dependent variable).
- ▶ While we can test whether the 1st condition is satisfied, the second condition cannot be formally tested. As a researcher you have to try to convince your audience that it is satisfied.

First and Second Stages

- First Stage regression:

$$S_i = \alpha_1 + \rho_1 Z_i + \kappa_i$$

where Z_i is the IV and S_i is the endogenous variable

- Second Stage regression:

$$Y_i = \alpha_2 + \rho S_i + \eta_i$$

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- Reduced Form equation:

$$Y_i = \alpha_3 + \rho_3 Z_i + \lambda_i$$

How IV Can be Used to Obtain Consistent Estimates?

- ▶ With one endogenous variable and one instrument the IV estimator is:

$$\hat{\rho}_{iv} = \frac{\text{Cov}(Y_i, Z_i)}{\text{Cov}(S_i, Z_i)} \quad (3)$$

- ▶ Hint: substitute the formula for Y from the correctly specified model (*)
- ▶ Substitute true model for Y :

$$\hat{\rho}_{iv} = \frac{\text{Cov}(a + \rho S_i + \gamma A_i + n_i, Z_i)}{\text{Cov}(S_i, Z_i)}$$

$$= \rho \frac{\text{Cov}(S_i, Z_i)}{\text{Cov}(S_i, Z_i)} + \gamma \frac{\text{Cov}(A_i, Z_i)}{\text{Cov}(S_i, Z_i)} + \frac{\text{Cov}(n_i, Z_i)}{\text{Cov}(S_i, Z_i)}$$

- ▶ Taking plims:

$$\text{plim} \hat{\rho}_{iv} = \rho$$

- ▶ The exclusion restriction says that $\text{Cov}(Z_i, h_i) = 0$, but $h_i = \gamma A_i + n_i$ so $\text{Cov}(Z_i, A_i)$ should be 0 and the $\text{Cov}(Z_i, n_i)$ should be equal to 0. $\text{Cov}(S_i, Z_i) \neq 0$ (due to the 1st stage).
- ▶ **The IV estimator is consistent if the IV assumptions are satisfied.**

Wald Estimator

- ▶ With one endogenous variable and one instrument, the IV estimator is:

$$\rho^{iv} = \frac{\text{cov}(Y_i, Z_i)}{\text{cov}(S_i, Z_i)} \Rightarrow$$

Wald Estimator:

$$\rho^{iv} = \frac{\text{cov}(Y_i, Z_i) / \text{Var}(Z_i)}{\text{cov}(S_i, Z_i) / \text{Var}(Z_i)} (= \frac{\rho_3}{\rho_1})$$

- ▶ The coefficient of interest is the ratio of the population regression of Y_i on Z_i (reduced form) to the population regression of S_i on Z_i (first stage).
- ▶ If there is a first stage, the denominator is different than zero.

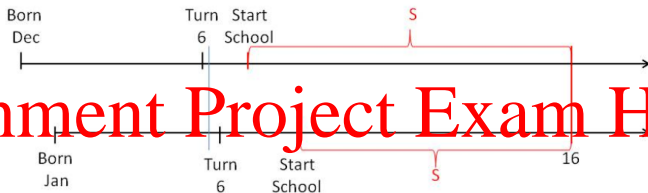
Where can you find a good instrument?

- ▶ Good instruments come from a combination of institutional knowledge and ideas about the process that determines the variable of interest.
- ▶ One source of variation is institutional constraints.
 - ▶ For schooling this can be compulsory schooling laws. Angrist and Krueger (1991) exploit the variation induced by compulsory schooling in a paper that uses an IV coming from a "natural experiment" to eliminate OVB. They would like to examine the effect of education on earnings. However, the naive regression has omitted variable bias.

Instrument for Education using Compulsory Schooling Laws

- ▶ In practice it is often difficult to find convincing instruments
 - ▶ Many potential IVs do not satisfy the exclusion restriction.
- ▶ In the returns to education literature Angrist and Krueger (QJE, 1991) have a very influential study in which they looked at the effect of compulsory schooling requirement on schooling and earnings.
- ▶ This is done by exploring the fact that **school entry age** and **compulsory schooling laws** lead people born in different times of the year to have different average levels of education.
- ▶ In the US you could drop out of school once you turned 16.
- ▶ Typically schools require children to have turned 6 by January 1 of the year in which they enter school.
- ▶ Children have different ages when they start school and thus different lengths of schooling at the time they turn 16 when they can potentially drop out.

Instrument for Education using Compulsory Schooling Laws



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In particular, those born late in the year are young for their grades.
- ▶ In states with a December 31 birthday cutoff, children born in the 4th quarter enter school shortly before they turn 6, while those born in the first quarter enter school at around 6.5.
- ▶ So some stay more in schooling than others based on when they have to start school. Let's assume they drop out when they turn 16.

Is there a first stage?

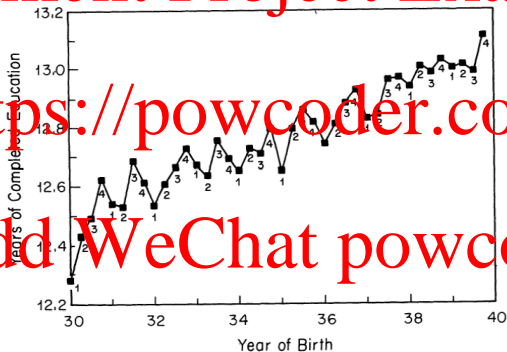
- First Stage Equation:

$$\text{YearsofEducation} = \alpha_1 + \rho_1 \text{QuarterofBirth} + \kappa_i$$

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Men born earlier in the year have lower schooling. This indicates that there is a first stage.

Is there a first stage?

Outcome variable	First cohort	Mean	Quarter-of-birth effect ^a			F-test ^b [P value]
			I	II	III	
Total years of education	1930–1939	12.79	0.124 (0.017)	0.086 (0.017)	-0.019 (0.016)	24.9 [0.0001]
	1940–1949	13.56	-0.085 (0.012)	-0.035 (0.012)	-0.017 (0.011)	18.6 [0.0001]
High school graduate	1930–1939	0.77	-0.019 (0.002)	-0.020 (0.002)	-0.004 (0.002)	46.4 [0.0001]
	1940–1949	0.86	0.015 (0.001)	-0.12 (0.007)	-0.002 (0.001)	54.4 [0.0001]
Years of educ. for high school graduates	1930–1939	13.99	-0.004 (0.014)	0.051 (0.014)	-0.012 (0.014)	6.9 [0.0006]
	1940–1949	14.28	0.005 (0.011)	0.043 (0.011)	-0.003 (0.010)	7.8 [0.0017]
College graduate	1930–1939	0.24	-0.005 (0.002)	0.003 (0.002)	0.002 (0.002)	5.0 [0.0021]
	1940–1949	0.30	-0.003 (0.002)	0.004 (0.002)	0.000 (0.002)	5.0 [0.0018]

Completed years of schooling is lower for men born in the first quarter of the year than the fourth quarter of the year. This indicates that there is a first stage.

Do differences in schooling due to different quarter of birth translate into different earnings?

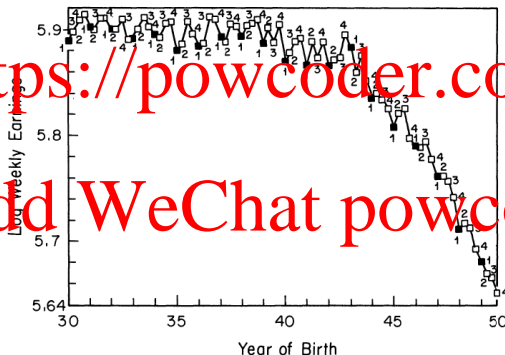
- ▶ Reduced Form Equation:

$$\text{Earnings} = \alpha_1 + \rho_3 \text{QuarterofBirth} + \kappa_i$$

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Younger cohorts tend to have lower earnings.

Two Stage Least Square (2SLS)

- ▶ In practice one estimates IV as a Two-Stage-Least Squares estimator (2SLS).

- ▶ It is called 2SLS because you could estimate it as follows:

1) Obtain the 1st stage fitted values:

- ▶ First Stage regression:

$$\hat{S}_i = \hat{\alpha}_1 + \hat{\lambda}X_i + \hat{\rho}_1Z_i + \kappa_i$$

where $\hat{\rho}_1$ and $\hat{\lambda}$ are OLS estimates of the 1st stage regression.

2) Plug the 1st stage fitted values into the "second-stage equation".

$$Y_i = \alpha_2 + \beta X_i + \gamma \hat{S}_i + error$$

Two Stage Least Square (2SLS)

- ▶ Despite its name the estimation is usually not done in two steps (because then the standard errors would be wrong).
- ▶ STATA or other regression softwares are usually doing the job for you (and get the standard errors right).
- ▶ The intuition of 2SLS, however, is very useful: 2SLS only keeps the variation in S that is generated by quasi-experimental variation (and thus hopefully exogenous).
- ▶ Angrist and Krueger use more than one instrument to instrument for schooling: they include a dummy for each quarter of birth.

$$S_i = X_{0i}\pi_{10} + \pi_{11}Z_{1i} + \pi_{12}Z_{2i} + \pi_{13}Z_{3i} + \psi_{1i}$$

IV Estimates Birth Cohorts 20-29, 1980 Census

Independent variable	(1) OLS	(2) TSLS
Years of education	0.0011 (0.0003)	0.0391 (0.0161)
Race (1 = black)	—	—
SMSA (1 = center city)	—	—
Married (1 = married)	—	—
9 Year-of-birth dummies	Yes	Yes
8 Region-of-residence dummies	No	No
Age	—	—
Age-squared	—	—
χ^2 [dof]	—	25.4 [29]

Additional years of education increase earnings.

Including More Covariates

Independent variable	(1) OLS	(2) TSLS	(3) OLS	(4) TSLS
Years of education	0.0673 (0.0003)	0.0928 (0.0093)	0.0673 (0.0003)	0.0907 (0.0107)
Race (1 = black)	—	—	—	—
SMSA (1 = center city)	—	—	—	—
Married (1 = married)	—	—	—	—
9 Year-of-birth dummies	Yes	Yes	Yes	Yes
8 Region-of-residence dummies	No	No	No	No
50 State-of-birth dummies	Yes	Yes	Yes	Yes
Age	—	—	-0.0757 (0.0617)	-0.0880 (0.0624)
Age-squared	—	—	0.0008 (0.0007)	0.0009 (0.0007)

Including more covariates should not affect much the coefficient of interest.

Including more Instruments

- ▶ They also included specifications where they use 30 (quarter of birth \times year) dummies and 150 (quarter of birth \times state) dummies as IVs (intuition: the effect of quarter of birth may vary by birth year or state).
- ▶ This reduces standard errors.
- ▶ But also comes at the cost of potentially having a weak instrument problem.

Weak Instruments

- ▶ As pointed out by Bound, Jaeger, and Baker (1993-1995), the "Cure can be worse than the disease" when the instruments are only weakly correlated with the endogenous variables.
- ▶ Staiger and Stock (1997) formalized the definition of "weak instruments" and for the case of one IV and one endogenous variable if the **F-statistic** on the instruments in the **first stage** is **greater than 10**, one need worry no further about weak instruments. This is a rule-of-thumb. (Stock, Wright, and Yogo (2002))

Weak Instruments

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- ▶ In the early 1960s a number of papers highlighted that IV estimates can be severely biased. In particular, if the IV is weak (first stage relationship is weak) and if you use many instruments to instrument for one endogenous variable.
- ▶ Adding more weak instruments will increase the bias of 2SLS. **By adding further instruments without predictive power the first stage F-statistic goes towards 0 and the bias will increase.**

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Weak Instruments in Angrist and Krueger

- ▶ Angrist and Krueger present findings while using different sets of instruments:
- ▶ 1) quarter of birth dummies, 3 instruments.
- ▶ 2) quarter of birth + (quarter of birth) \times (year of birth) dummies, 30 instruments.
- ▶ 3) quarter of birth + (quarter of birth) \times (year of birth) + (quarter of birth) \times (state of birth), 180 instruments

Adding Instruments in Angrist and Krueger

	(1) OLS	(2) IV	(3) OLS	(4) IV
Coefficient	.063 (.000)	.142 (.033)	.063 (.000)	.081 (.016)
F (excluded instruments)		13.486		4.747
Partial R^2 (excluded instruments, $\times 100$)		.012		.043
F (overidentification)		.932		.775
<i>Age Control Variables</i>				
Age, Age ²	x	x		
9 Year of birth dummies			x	x
<i>Excluded Instruments</i>				
Quarter of birth		x		x
Quarter of birth \times year of birth				x
Number of excluded instruments		3		30

Adding more weak instruments reduced the 1st stage F-statistic and moves the coefficient towards the OLS coefficient.

Another Example: Angrist (1990) Veteran Draft Lottery

- ▶ In the following, we will often refer to an example from Angrist's paper on the effects of military service on earnings.
- ▶ Consider the following equation:

$$Y_i = \beta + \gamma S_i + h$$

where Y_i are **earnings** and S_i is a **military/veteran dummy**.

- ▶ If one wants to interpret γ as being the direct effect of having served the armed forces, OLS will be biased.
- ▶ Why? Many omitted factors might be correlated with one's willingness to serve the military, and these factors are likely to be correlated with earnings.
- ▶ If we want to estimate the causal effect of S_i on Y_i , we need to find an instrument.
- ▶ This instrument should be correlated with military/veteran status, but it should not affect earning in any other way (only through the effect on veteran status).
- ▶ Why is this important? Debate about whether veterans are adequately compensated for their service.

Angrist (1990) Veteran Draft Lottery

- ▶ Angrist (1990) uses the **Vietnam draft lottery** as in **IV** for **military service**.
- ▶ In the 1960s and early 1970s, young American men were drafted for military service to serve in Vietnam. However, there were **concerns about the fairness of the conscription policy** (certain types of men used to serve the army). This leads to the introduction of a **draft lottery** in 1970.
- ▶ From 1970 to 1972 **random** sequence numbers were assigned to each birth date in cohorts of 19-year-olds.
- ▶ Men with lottery numbers below a cutoff were drafted while men with numbers above the cutoff could not be drafted.
- ▶ *Nevertheless, the draft did not perfectly determinate military service:*
 - ▶ Many draft-eligible men were exempted for health and other reasons.
 - ▶ Exempted men volunteered for service.

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► First stage results:

- Having a low lottery number (being eligible for the draft) increases veteran status by about 16 percentage points (the mean of veteran status is about 27 percent).
- Second stage results: Serving in the army lowers earnings by between 2,050 and 2,741 dollars per year, even long after the service in Vietnam was ended.

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- ▶ Up to this point we only considered models where the causal effect was the same for all individuals (homogenous treatment effects):

- ▶ Be aware that the treatment effects might be heterogeneous for each population group.

- ▶ This will inform us about two types of validity characterizing research designs:

1 Internal validity: Does the design successfully uncover causal effects for the population studied?

2 External validity: Do the study's results inform us about different populations?

- ▶ Different subpopulations:
- ▶ 1) **Compliers**: treated because of being selected.

All the rest are **non-compliers**

- ▶ 2) **Always-takers**: They always take the treatment independently of the draft.
- ▶ 3) **Never-takers**: They never take the treatment independently of the draft.
- ▶ 4) **Defiers**: Non-treated when selected and treated when non-selected.
- ▶ Under some assumptions, IV estimates focus on the average effect of military service on earnings for the sub-population who enrolled in military service, because of the draft but would not have served otherwise.

Card (1995) Use geographical variation in college proximity to estimate the returns to schooling

- ▶ Estimating the returns to schooling using the following regression:

$$\ln wage = \beta_0 + \beta_1 educ + u$$

- ▶ educ is endogenous and $E(educ, u)$ is not 0.
- ▶ We need to find an Instrument Z_i , that is highly correlated with educ
- ▶ Thus, the naive OLS estimate is biased.

Card (1995) paper

Table 2: Estimated Regression Models for Log Hourly Earnings

	(1)	(2)	(3)	(4)	(5)
1. Education	0.074 (0.004)	0.075 (0.003)	0.073 (0.004)	0.074 (0.004)	0.073 (0.004)
2. Experience	0.084 (0.007)	0.085 (0.007)	0.085 (0.007)	0.085 (0.007)	0.085 (0.007)
3. Experience-squared /100	-0.224 (0.032)	-0.229 (0.032)	-0.230 (0.032)	-0.226 (0.032)	-0.229 (0.032)
4. Black Indicator	-0.190 (0.017)	-0.199 (0.018)	-0.194 (0.019)	-0.194 (0.019)	-0.189 (0.019)
5. Live in South	-0.125 (0.015)	-0.148 (0.026)	-0.146 (0.026)	-0.145 (0.026)	-0.146 (0.026)
6. Live in SMSA	0.16 (0.019)	0.136 (0.020)	0.16 (0.020)	0.137 (0.020)	0.138 (0.020)
7. Region in 1966 (8 indicators)	no	yes	yes	yes	yes
8. Live in SMSA in 1966	no	yes	yes	yes	yes
9. Parental Education ^a (mean effects)	no	no	yes	yes	yes
10. Interaction Parental Education Class ^{a,b}	no	no	no	yes	yes
11. Family Structure ^c (2 indicators)	no	no	no	no	yes
12. R-squared	0.291	0.300	0.301	0.303	0.304
13. P-value for family background effects	--	--	0.235	0.462	0.165

Notes: Standard errors in parentheses. Sample size is 3010. The dependent variable in all cases is the log of hourly wages in 1976. The mean and standard deviation of the dependent variable are 6.262 and 0.444.

^a Variables representing years of education of mother and father, plus indicators for missing mother's or father's education.

^b Indicators for 8 classes of mother's and father's education.

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Card (1995) paper

- ▶ Card (1990) used the **proximity to a four-year college**

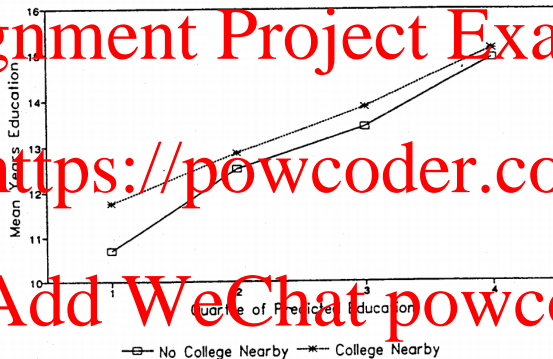
($nearc4_i$) in high school as an IV

- ▶ One who lives in an area without a college faces a higher cost of college education, since the option of living at home is precluded (i.e., they will have to move). One would expect this to reduce investment in higher education, at least for low income families.

- ▶ So $nearc4_i$ is a binary indicator, equal to 1 if a man is near a four-year college in high school.
- ▶ Would expect $x_i = educ_i$ and $z_i = nearc4_i$ to be positively related. They have to be for the first stage to exist.

Always plot the First Stage!

Mean Years of Education
By Quartile of Predicted Education



For every quartile the mean level of education is higher for those who grew up near a college.

Run regressions to see if there is a First Stage!

```
. use card
```

```
. sum educ nearc4
```

Variable	Obs	Mean	Std. Dev.	Min	Max
educ	3010	13.26346	2.676913	1	18
nearc4	3010	.6820598	.4657535	0	1

```
. reg educ nearc4, robust
```

Linear regression

Number of obs =	3010
F(1, 3008) =	60.37
Prob > F =	0.0000
R-squared =	0.0208
Root MSE =	2.6494

	Coef.	Robust Std. Err.	t	Prob > t	[95% Conf. Interval]
educ	.829019	.1066941	7.77	0.000	.6198182 1.03822
nearc4	.829019	.1066941	7.77	0.000	.6198182 1.03822
_cons	12.69801	.0902199	140.75	0.000	12.52112 12.87491

```
. * educ and nearc4 are strongly enough related: being near a 4-year college  
. * increases educ by almost a year. t statistic is pretty large.
```

```
. ivreg lwage (educ = nearc4), robust
```

```
Instrumental variables (2SLS) regression
```

```
Number of obs = 3010  
F(1, 2808) = 1.5  
Prob > F = 0.0000  
R-squared = .  
Root MSE = .55686
```

lwage	Coef	Robust Std. Err.	t	> t	[95% Conf. Interval]
educ	1.1880626	.1261476	7.11	.000	1.31873 .2393217
_cons	3.767472	.346742	10.87	0.000	3.087596 4.447347

```
Instrumented: educ  
Instruments: nearc4
```

```
. * Note that the list of exogenous variables in the lwage equation is empty.  
. * Estimated return to education seems too large. CI is wide, but lower  
. * bound is still 13.7%.
```

. * For comparison, OLS:

. * reg lwage educ robust

Linear regression

Number of obs = 3010
F(1, 3008) = 321.16
Prob > F = 0.0000
R-squared = 0.0987
Root MSE = .42139

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	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]
educ	.0520942	.0029069	17.92	0.000	.0463946 .0577939
_cons	5.570882	.0390935	142.50	0.000	5.49423 5.647535

. * 13.8% versus 5.2%

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▶ Card (1995) paper

- ▶ Card in this paper reveals that men who grew up in local labor markets with a nearby college have significantly higher education and earnings than other men.
- ▶ These gains are concentrated among men with poorly educated parents, who would otherwise stop schooling at low levels.
- ▶ Effects may vary by family background.

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► Why are OLS and IV some cases so different in the Card case?

- A common explanation is that *educ* is **measured with error**, so there is attenuation bias (the estimate is biased towards 0) when using OLS.
- Another explanation is that the **return to schooling is not constant** and IV is picking up the effect for a certain subgroup.
- Another explanation is the possibility that **the instrument is somewhat endogenous**. (i.e. $\text{Corr}(\text{distance to college}, u) \neq 0$)

- Families that place a strong emphasis on education might choose to live near a college.
- The presence of a college might be associated with higher school quality at nearby schools (they could control for school quality.)
- College proximity may be correlated with unobserved geographic wage premiums.

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Two-stage least squares in Stata

- ▶ (in Stata) `ivregress 2sls depvar exogvars (endogvars = insts)`

where `exogvars` is the list of exogenous regressors; `endogvars` is the list of endogenous variables (there can be more than one); and `insts` is the list of instruments.

- ▶ `ivregress 2sls depvar exogvars (endogvars = insts), first`
- ▶ The 'first' option yields estimates of the 1st-stage equation.

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Why not always use IV?

- ▶ It is hard to find variables that meet the definitions of valid instruments: conceptually, most variables that have an effect on endogenous variables x may also have a direct effect on the dependent variable y .
- ▶ The standard errors on IV estimates are likely to be larger than OLS estimates, and much larger if the instrumental variables are only weakly correlated with the endogenous regressors.

Practical Tips for IV Papers

- ▶ Report the first stage
 - ▶ Does it make sense?
 - ▶ Do the coefficients have the right magnitude and size?
- ▶ Report the F-statistic.
 - ▶ Stock, Wright and Yogo (2002) suggest that F-statistics above 10 indicates that you do not have a weak instrument problem.
- ▶ Look at the Reduced Form:
 - ▶ The reduced form is estimated with OLS.
 - ▶ If you can't see the causal relationship of interest in the reduced form it is probably not there.