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ST227 Survival Models-Part II

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Continuous Time Markov Chains

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1 **Stochastic Processes**

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1.1 Introduction

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- A **stochastic process** is a model for a time-dependent random phenomenon.

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- Thus, just as a single random variable describes a static random phenomenon, a **stochastic process** is a **collection of random variables**, $Y(t) = Y_t$, one for each time t in some set J .
- The **process** is denoted $\{Y_t : t \in J\}$. The set of values that the random variables Y_t can take is called the **state space** of the process, S .

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- The **first choice that one faces** when selecting a stochastic process to model a real life situation is that of the nature (**discrete or continuous**) of the time set J and of the state space S .

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Example 1: Discrete state space with discrete time changes

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- A motor insurance company reviews the status of its customers yearly. Three levels of discount are possible (0, 25%, 40%) depending on the accident record of the driver.
- In this case the appropriate state space is $S = \{0, 25, 40\}$ and the time set is $J = \{0, 1, 2, \dots\}$ where each interval represents a year.

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Example 2: Discrete state space with continuous time changes

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- A life insurance company classifies its policyholders as healthy, ill or dead.

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- Hence the state space $S = \{h, i, d\}$.

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- As for the time set, it is natural to take $T = [0, \infty)$ as illness or death can occur at any time. This problem is studied in some detail in what follows (Continuous Time Markov Chains).

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Example 3: Continuous state space

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- Claims of unpredictable amounts reach an automobile insurance company at unpredictable times; the company needs to forecast the **cumulative claim** over $[0, t]$ in order to assess the risk that it might not be able to meet its liabilities.

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- It is standard practice to use $[0, \infty)$ both for S and J in this problem.

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- It is important to be able to **conceptualise the nature of the state space of any process** which is to be analysed and to establish whether it is most usefully modelled using a discrete, a continuous, or a mixed time domain.

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- Usually the **choice of state space will be clear from the nature of the process being studied** (as, for example, with the healthy-ill-dead model), but whether a continuous or discrete time set is used will often depend on the **specific aspects of the process** which are of interest, and upon **practical issues** like the time points for which data are available.

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1.1.1 The Markov property

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- In probability theory and statistics, the term **Markov property** refers to the **memoryless property of a stochastic process**. It is named after the Russian mathematician Andrey Markov.

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- A stochastic process $\{Y(t)\}_{t \geq 0}$ is a Markov process if the conditional probability distribution of future states depends only on the previous state.

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- For example, if $s \geq 0$ and we consider the states i and j , then we can say that $Y(t)$ has the **Markov Property** if $P(Y(t+s) = j | Y(t) = i)$ does not depend on any information before t .
- Hence, the **future development of $Y(t)$ can be predicted from its present state alone, without any reference to its history.**

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2 **Continuous Time Markov Chains**

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2.1 Introduction

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- Up till now our approach has been to specify a future lifetime in terms of a random variable T_x . In this chapter we look at things rather differently, and use a **Continuous Time Markov Chain (or a Markov model)** of transfers between states.
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- In the simplest case, a life can be alive or dead, and this gives a two-state model of mortality which is known as the **dead-or-alive model**. We often represent this in a diagram like:

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- The probability that a life alive at any age x should be dead at some future age is determined by an age dependent force of mortality μ_{x+t} , $t \geq 0$, or **transition intensity**.

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- The **main advantages of the Markov model approach** to modelling mortality over the random variable approach are that:
 - it can be **generalised to multiple decrements or multiple state models**, e.g. the three state model $\{Well, Ill, Dead\}$, and

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- it **deals easily with censoring**, a common feature of mortality data, which we will talk about later on in the chapters ahead.

- We make **two fundamental assumptions** in the 2-state model above.

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1. **(AS1) Markov assumption.** The probability that a life now aged x will be in either state $\{Alive, Dead\}$ at any future time $x + t$ depends only on the age x and the state currently occupied.

2. **(AS2)** The probability ${}_tq_{x+t}$ is given by

$${}_tq_{x+t} = \mu_{x+t}dt + o(dt), t \geq 0. \quad (1)$$

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- Informally, **assumption 1** says that the **future depends only on the present** and not on the past, while **assumption 2** says that the **probability of death in a short interval of time, dt , is approximately proportional to the force of mortality at that time.**

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- **Reminder:** The function $f(h)$ is said to be $o(h)$ or "little o of h " if

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$$\lim_{h \rightarrow 0} \frac{f(h)}{h} = 0.$$

(2)

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In other words, $f(h)$ is $o(h)$ if $f(h) \rightarrow 0$ faster than h .

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2.2 **Computation of ${}_t p_x$ in the dead-or-alive model**

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Our model is defined in terms of the transition intensities μ_x . How can we compute probabilities like ${}_t p_x$, ${}_t q_x$, etc?

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Lemma 1 *Assumptions 1 and 2 imply that*

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$${}_t p_x = \exp \left(- \int_0^t \mu_{x+s} ds \right). \quad (3)$$

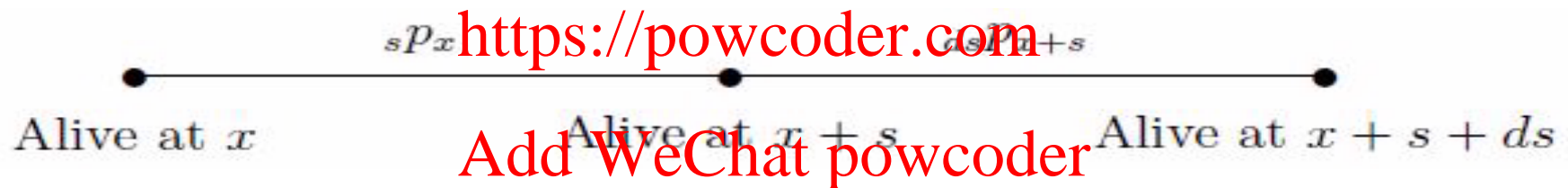
. This agrees with the well known result obtained with the future lifetime approach in Part 1.

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Proof: Let $s < t$. Notice first that ${}_s p_x$ is well defined by the Markov property, i.e. how the life got to age x is irrelevant. We consider the small interval of time *immediately after* $x + s$, and ask: how can a life aged x become a life aged $x + s + ds$? The diagram may help:

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Hence, the probability that (x) survives $s + ds$ years is equal to the probability that they survive s years times the probability that, when at age $x + s$, they survive ds years.

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$$\begin{aligned} {}_{s+ds}p_x &= {}_s p_x \times {}_{ds}p_{x+s} \\ &= {}_s p_x \times (1 - {}_{ds}q_{x+s}) \\ &= {}_s p_x \times (1 - \mu_{x+s} ds + 0(ds)) \quad \text{by Assumption 2.} \end{aligned} \quad (4)$$

Now bring the term ${}_s p_x$ from the right to the left hand side of (4) we get

$${}_{s+ds}p_x - {}_s p_x = -{}_s p_x \mu_{x+s} ds + 0(ds). \quad (5)$$

Then divide both sides of (5) by ds and let $ds \rightarrow 0$. We find

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$$\frac{s + ds p_x - s p_x}{ds} = -s p_x \mu_{x+s} + \frac{0(ds)}{ds}$$
$$\Rightarrow \frac{\partial}{\partial s} s p_x = -s p_x \mu_{x+s} \quad (6)$$

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$$\Rightarrow t p_x = \exp\left(-\int_0^t \mu_{x+s} ds\right)$$

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on integrating (6) from 0 to t , this result will be proved in the class.

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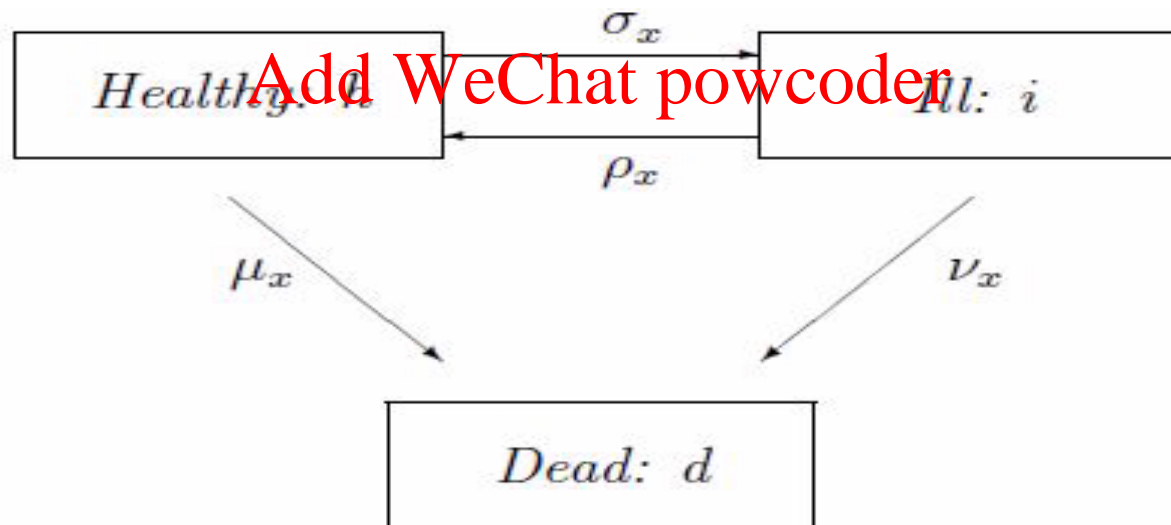
2.3 Multi-state Markov models

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- The 2-state model of mortality can be extended to any number of states. Many insurance products, e.g. Permanent Health Insurance (PHI), can be modelled by a **multi-state model**. The set $S = \{Healthy, Ill, Dead\} = \{h, i, d\}$ is the **state space**. Here is a 3 state model for PHI:

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- Let g and h be any two states. We extend the assumptions for the 2-state model to cope with a multi-state model. We do this in terms of the **transition probability** ${}_t p_x^{gh}$ (analogous to ${}_t p_x$) and the **force of transition (aka transition intensity)** μ_x^{gh} (analogous to μ_x).

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- We define the **transition probability**

$${}_t p_x^{gh} = P_r(\text{In state } j \text{ at time } x+t \mid \text{In state } i \text{ at time } x) \quad (7)$$

for any two states i and j .

- Also, for $z > 0$ define the **force of transition**

$$\mu_x^{gh} = \lim_{z \rightarrow 0^+} \frac{{}_z p_x^{gh}}{z}. \quad (8)$$

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2.4 Fundamental Assumptions for Multi-State models

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1. **(AS1) Markov assumption.** The probability that a life now aged x will be in a particular state at any future time $x + t$ depends only on the age x and the state currently occupied.

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2. **(AS2)** For any two distinct states g and h the transition probability ${}_dt p_{x+t}^{gh}$ is given by

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$${}_dt p_{x+t}^{gh} = \mu_{x+t}^{gh} dt + o(dt), \quad t \geq 0. \quad (9)$$

3. **(AS3)** The probability that a life makes two or more transitions in time dt is $o(dt)$. **Assumption 3** says, in effect, that only one transfer can take place at one time.

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- What does ${}_t p_x^{gg}$ mean?

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- This is the probability that we are in state g at time $x + t$, given that we are in state g at time x .

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- This **does not imply** that we have been in state g for the whole of the time interval from x to $x + t$, for this we define the **occupation probability**.

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- **Occupation Probability:**

$${}_t \overline{p}_x^{gg} = P_r(\text{In state } g \text{ from } x \text{ to } x + t \mid \text{In state } g \text{ at time } x) \quad (10)$$

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- Note that ${}_t p_x^{gg} \leq \overline{{}_t p_x^{gg}}$.

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- Because $\overline{{}_t p_x^{gg}}$ is the occupation probability, i.e. the individual never leaves state g between ages x and $x + t$.

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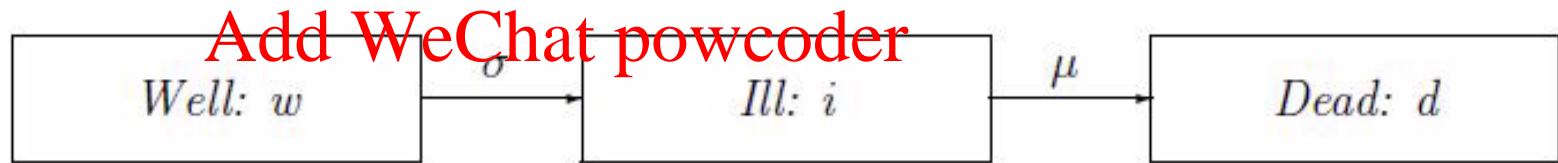
- The important distinction is that ${}_t p_x^{gg}$ includes the possibility that the individual leaves state g between ages x and $x + t$, provided they back in state g at age $x + t$. This result will be shown in the next class.

- However $\overline{{}_t p_x^{gg}}$ will be equal to ${}_t p_x^{gg}$ in one common situation, namely when return to state g is impossible.

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- For example, in this **model of terminal illness**



we have $tP_x^{\overline{ww}} = tP_x^{\overline{gg}}$ since return to the well state is impossible.

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- In a similar fashion, we also have $tP_x^{\overline{ii}} = tP_x^{\overline{ii}}$.

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2.5 Assignment Project Exam Help Kolmogorov forward equations

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- What can we say about the **relationship between the transition intensities μ_{x+t}^{gh} , $g \neq h$ and the transition probabilities ${}_t p_x^{gh}$?**

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- We will look at **two examples** in detail **before giving the general result.**

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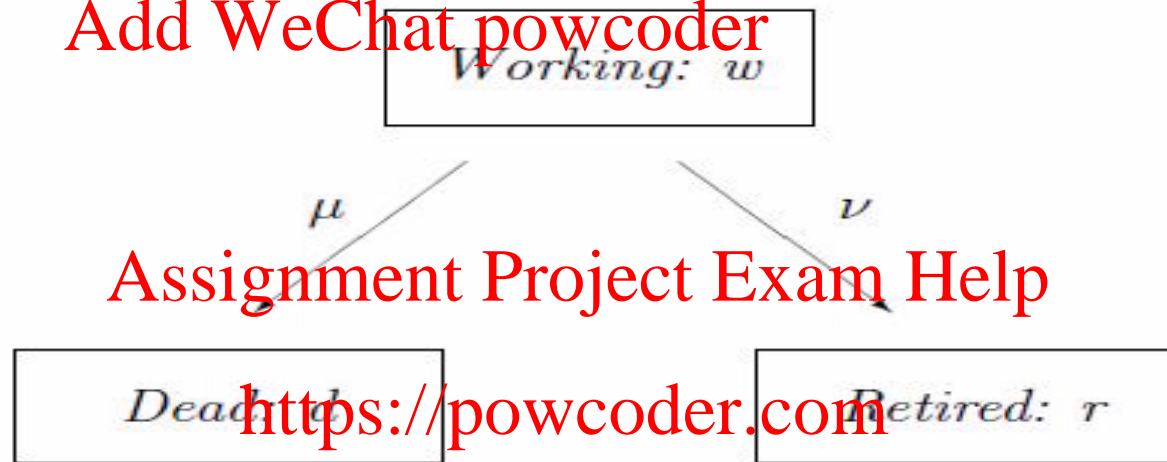
Example 1 Consider the 3-state model for working, retiring and dying.

- In this simple example we will assume **two constant transition intensities** μ (from working to dying) and σ (from working to retiring).

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- There are **three transition probabilities** which correspond to the events:
(a) *Working to Dead*, (b) *Working to Retired* and (c) *Working to Working*.

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- (a) *Working to Dead* or tp_x^{wd} . We use a standard method in all of these kind of problems

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- **Step 1:** We suppose we are in the destination state (here Dead) at time $x + t + dt$.

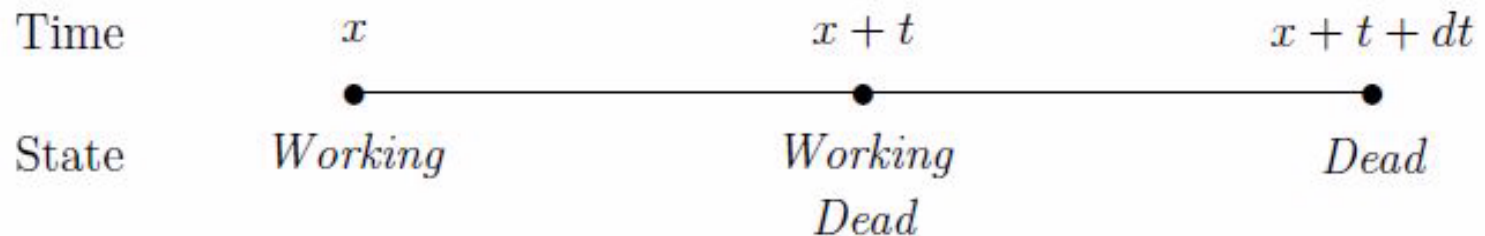
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- **Step 2:** We list the states we could be in at time $x + t$, i.e. just before time $x + t + dt$.

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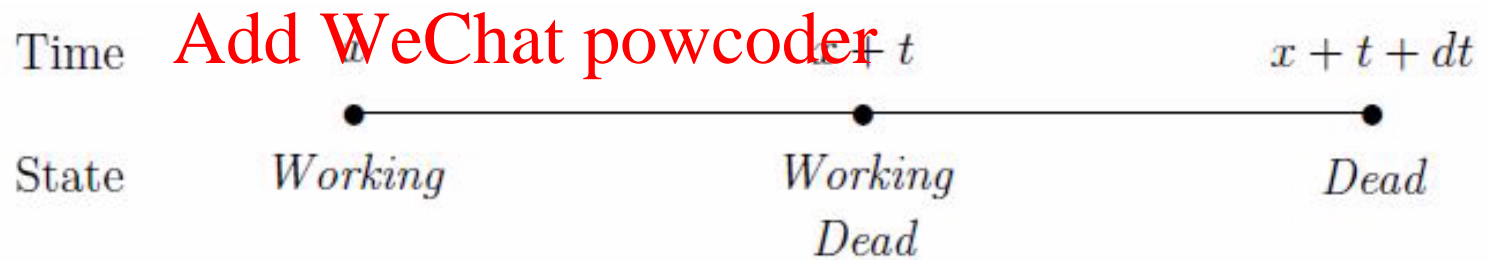
- The diagram might help:

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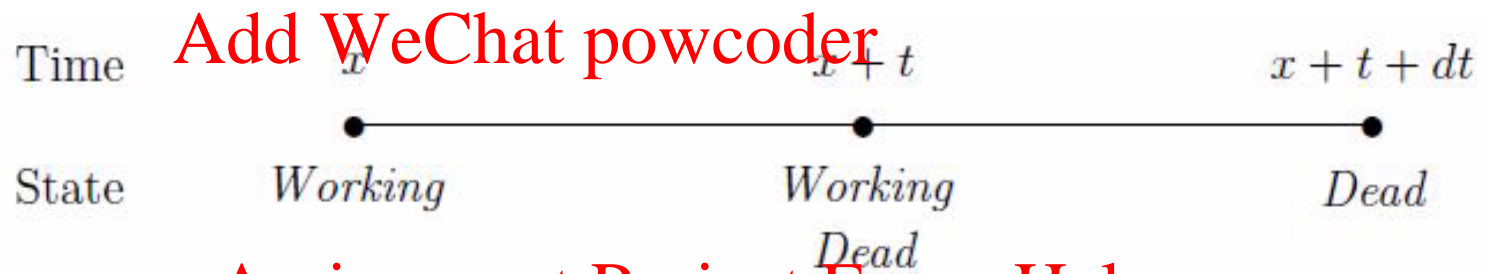


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- The **left end** represents the starting position at time x .
- The **right end** represents the final position at time $x + t + dt$.
- The **middle position** lists the states that can be occupied at $x + t$ **immediately before** the final position at $x + t + dt$.

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- Thus, we have that <https://powcoder.com>

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$${}_{t+dt}p_x^{wd} = {}_t p_x^{wd} + {}_t p_x^{ww} \times {}_{dt} p_{x+t}^{wd} \quad (11)$$

$$= {}_t p_x^{wd} + {}_t p_x^{ww} \times (\mu dt + o(dt)). \quad (12)$$

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Rearranging we get

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$$\frac{{}_{t+dt}p_x^{wd} - {}_t p_x^{wd}}{dt} = \mu \cdot {}_t p_x^{ww} + \frac{o(dt)}{dt} \quad (13)$$

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so taking the limit $dt \rightarrow 0$ we get

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$$\frac{\partial}{\partial t} {}_t p_x^{wd} = \mu \cdot {}_t p_x^{ww}. \quad (14)$$

This is the Kolmogorov forward equation for ${}_t p_x^{wd}$.

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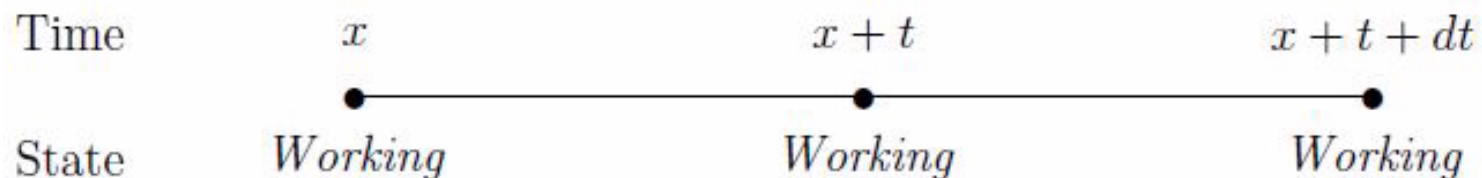
- (b) *Working to Retired* or tp_x^{wr} . We can apply exactly the same argument to tp_x^{wr} but it is better to use the **symmetry of the diagram**. This gives the Kolmogorov forward equation for tp_x^{wr} as

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$$\frac{\partial}{\partial t} tp_x^{wr} = \nu \cdot tp_x^{ww} \quad (15)$$

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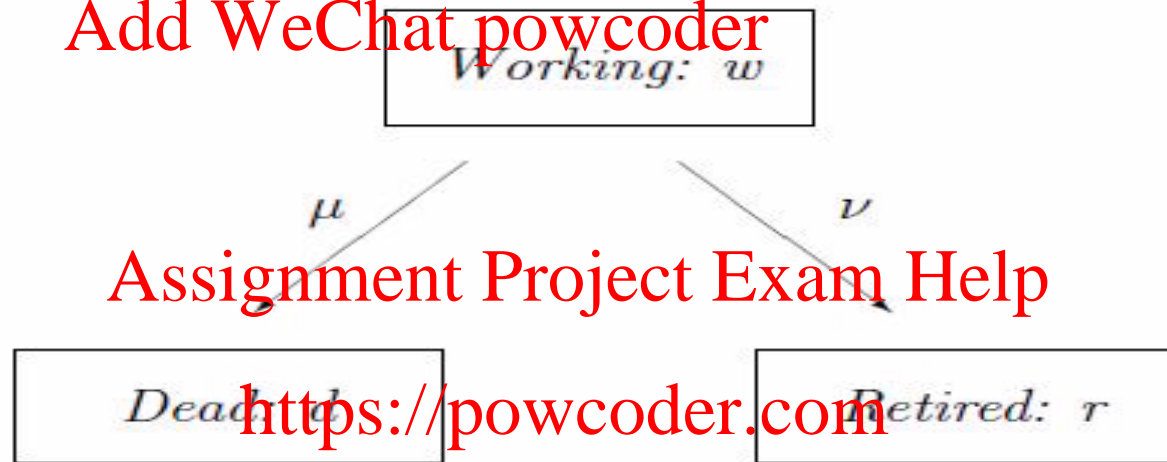
- (c) *Working to Working*. First, notice that return to the *Working* state is impossible so $tp_x^{ww} = tp_x^{\overline{ww}}$. The diagram of possible routes is very simple:



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The probability of transfer out of state *Working* in time dt is

$${}_dt p_{x+t}^{wd} + {}_dt p_{x+t}^{wr} = \mu dt + 0(dt) + \nu \cdot dt + 0(dt) \quad (16)$$

$$= \mu \cdot dt + \nu \cdot dt + 0(dt). \quad (17)$$

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Hence, the probability we **remain in** state *Working* for time dt is

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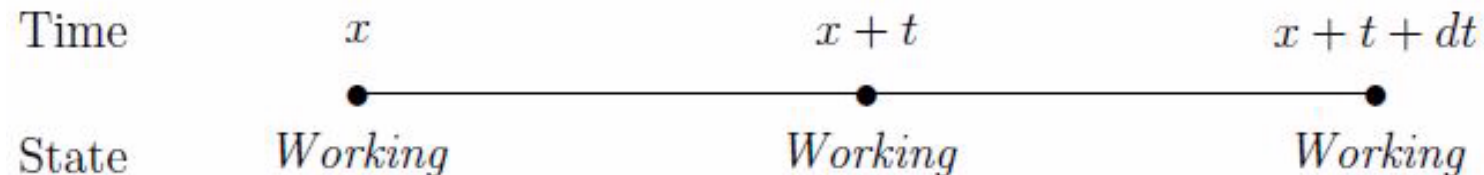
$$1 - \mu \cdot dt - \nu \cdot dt + 0(dt). \quad (18)$$

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Thus, since

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we get

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$$t_{t+dt} p_x^{\overline{ww}} = t p_x^{\overline{ww}} (1 - \mu \cdot dt - \nu \cdot dt + o(dt)). \quad (19)$$

Rearranging and letting $dt \rightarrow 0$, we find

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$$\frac{\partial}{\partial t} t p_x^{\overline{ww}} = -(\mu + \nu) t p_x^{\overline{ww}} \quad (20)$$

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and putting $t p_x^{\overline{ww}} = t p_x^{ww}$ gives the Kolmogorov equation for $t p_x^{ww}$:

$$\frac{\partial}{\partial t} t p_x^{ww} = -(\mu + \nu) t p_x^{ww}. \quad (21)$$

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- We now have a system of three differential equations (14), (15) and (21) for the three unknown transition probabilities, ${}_t p_x^{ww}$, ${}_t p_x^{wr}$ and ${}_t p_x^{wd}$.

- Note that

$${}_t p_x^{ww} + {}_t p_x^{wd} + {}_t p_x^{wr} = 1 \quad (22)$$

since a life in state w at time x must be in some state at time $x + t$.

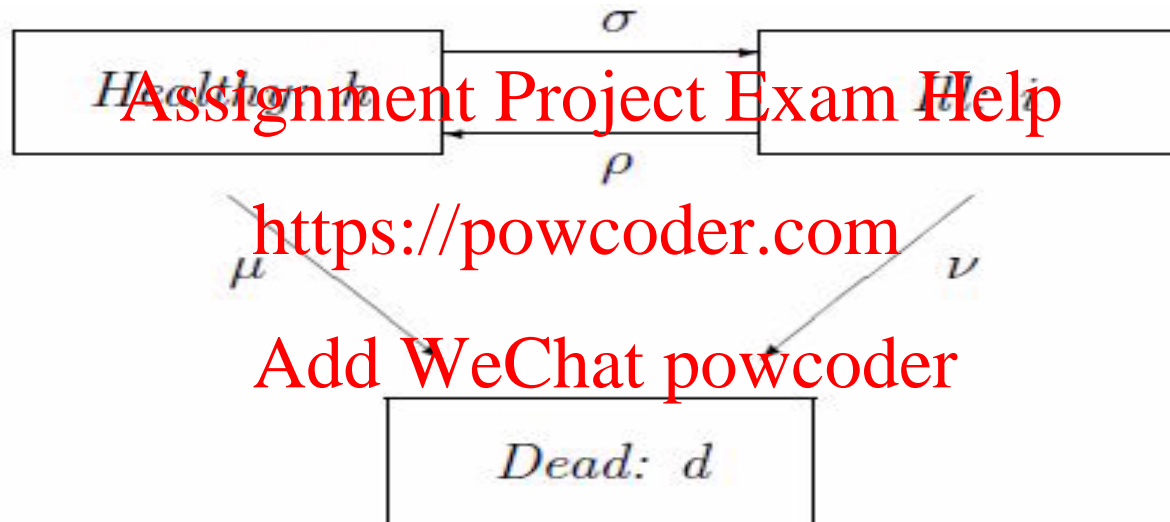
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Example 2 For our second example we return to the 3-state model for PHI.

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- We have assumed that the transition intensities μ , σ , ρ and ν are constant.

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- There are six transition probabilities ${}_t p_x^{hh}$, ${}_t p_x^{hi}$, ${}_t p_x^{hd}$, ${}_t p_x^{ii}$, ${}_t p_x^{ih}$ and ${}_t p_x^{id}$.

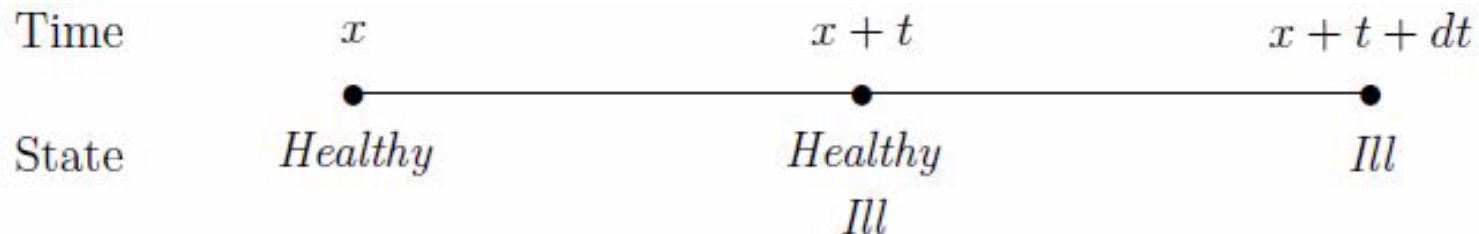
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- We look in detail at the derivation of the Kolmogorov equations for three of these, ${}_t p_x^{hh}$, ${}_t p_x^{hi}$ and ${}_t p_x^{hd}$. (The remaining three equations can then be written down by using symmetry arguments.)

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- (a) *Healthy to Ill* or ${}_t p_x^{hi}$. The *Ill* state at time $x + t + dt$ can be reached from either the *Healthy* or the *Ill* state at time $x + t$. Our diagram is

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Hence

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$${}_{t+dt}p_x^{hi} = {}_t p_x^{hh} \times (\sigma dt + 0(dt)) + {}_t p_x^{hi} \times (1 - \rho dt - \nu dt + 0(dt)) \quad (23)$$

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Rearranging we get

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$$\frac{{}_{t+dt}p_x^{hi} - {}_t p_x^{hi}}{dt} = \sigma \cdot {}_t p_x^{hh} - (\rho + \nu) {}_t p_x^{hi} + \frac{0(dt)}{dt} \quad (24)$$

and taking the limit $dt \rightarrow 0$ gives the Kolmogorov forward equation for ${}_t p_x^{hi}$ (next slide).

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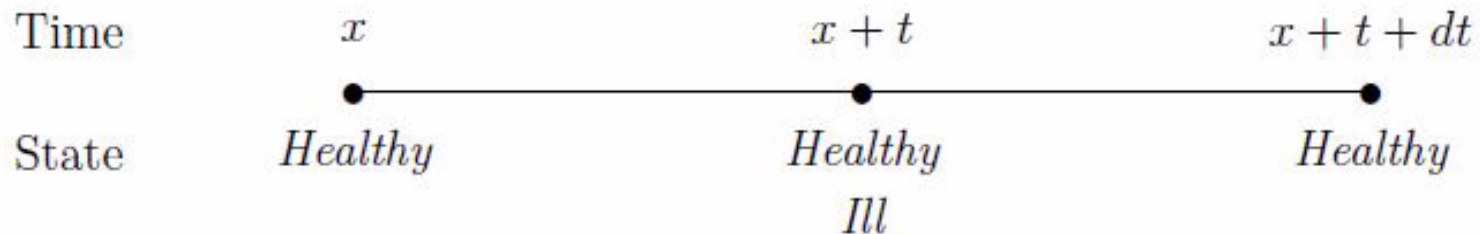
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$$\frac{\partial}{\partial t} {}_t p_x^{hi} = \sigma {}_t p_x^{hi} - (\rho + \nu) {}_t p_x^{hi}. \quad (25)$$

- (b) *Healthy* to *Healthy* or ${}_t p_x^{hh}$. The *Healthy* state at time $x + t + dt$ can be reached from either the *Healthy* or the *Illstate* at time $x + t$. Our diagram is

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Hence

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$${}_{t+dt}p_x^{hh} = {}_tp_x^{hh} \times (1 - \sigma dt - \mu dt + o(dt)) + {}_tp_x^{hi} \times (\rho dt + o(dt)) \quad (26)$$

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Rearranging we get

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$$\frac{{}_{t+dt}p_x^{hh} - {}_tp_x^{hh}}{dt} = \rho \cdot {}_tp_x^{hi} - (\sigma + \mu){}_tp_x^{hh} + \frac{o(dt)}{dt} \quad (27)$$

and taking the limit $dt \rightarrow 0$ gives the Kolmogorov forward equation for ${}_tp_x^{hh}$ (next slide).

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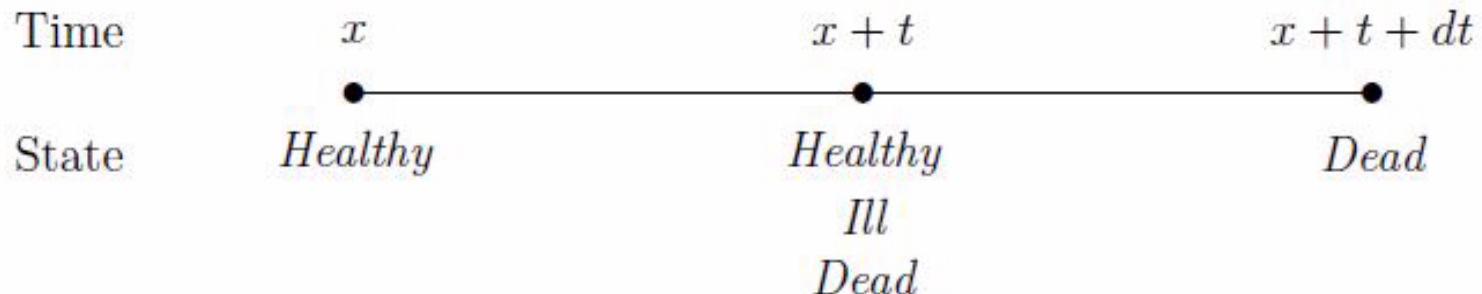
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$$\frac{\partial}{\partial t} {}_t p_x^{hh} = \rho {}_t p_x^{hi} - (\sigma + \mu) {}_t p_x^{hh} \quad (28)$$

- (c) *Healthy to Dead* or ${}_t p_x^{hd}$ The *Dead* state at time $x + t + dt$ can be reached from either the *Healthy*, the *Ill* or the *Dead* state at time $x + t$. Our diagram is

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Hence

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$$_{t+dt}p_x^{hd} = _tp_x^{hh} \times (\mu dt + o(dt)) + _tp_x^{hi} \times (\nu dt + o(dt)) + _tp_x^{hd} \times 1 \quad (29)$$

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The Kolmogorov forward equation for p_x^{hd} follows by rearranging and taking the limit $dt \rightarrow 0$. We find

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$$\frac{\partial}{\partial t} _tp_x^{hd} = \mu \cdot _tp_x^{hh} + \nu \cdot _tp_x^{hi}.$$

(30)

Comment: Note that

$$_tp_x^{hd} = 1 - _tp_x^{hh} - _tp_x^{hi}. \quad (31)$$

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2.6 The general Kolmogorov equations

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- What can we say about the relationship between the transition intensities μ_{x+t}^{gh} , $g \neq h$ and the transition probabilities ${}_t p_x^{gh}$?

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- The **general Kolmogorov equations** generalise the previous two examples. We show that

$$\frac{\partial}{\partial t} {}_t p_x^{gh} = \sum_{j \neq h} \left({}_t p_x^{gj} \mu_{x+t}^{jh} - {}_t p_x^{gh} \mu_{x+t}^{gh} \right), \quad g \neq h \quad (32)$$

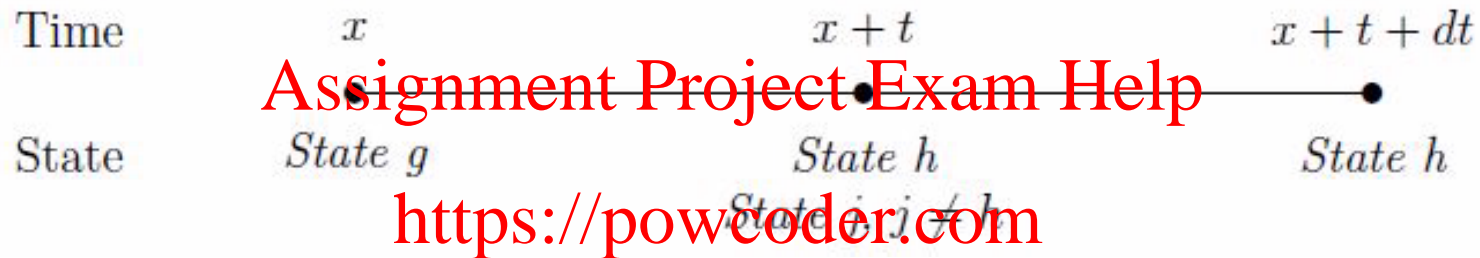
- We are interested in transfers from state g at time x to state h at time $x + t + dt$. So at time $x + t$ we are already in state h , or we have still to reach state h from some other state j .

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- Our diagram is

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- Hence we have that (next slide):

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$$\begin{aligned}
 {}_{t+dt}p_x^{gh} &= {}_t p_x^{gh} \times {}_{dt}p_{x+t} + \sum_{j \neq h} {}_t p_x^{gj} \times {}_{dt}p_{x+t}^{jh} \\
 &= {}_t p_x^{gh} \times \left(1 - \sum_{j \neq h} \mu_{x+t}^{hj} dt + 0(dt) \right) + \sum_{j \neq h} {}_t p_x^{gj} \left(\mu_{x+t}^{hj} dt + 0(dt) \right)
 \end{aligned} \tag{33}$$

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- Rearranging we get

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$$\frac{{}_{t+dt}p_x^{gh} - {}_t p_x^{gh}}{dt} = \sum_{j \neq h} ({}_t p_x^{gj} \mu_{x+t}^{jh} - {}_t p_x^{gh} \mu_{x+t}^{hj}) + \frac{0(dt)}{dt} \tag{34}$$

and the result follows on letting $dt \rightarrow 0$.

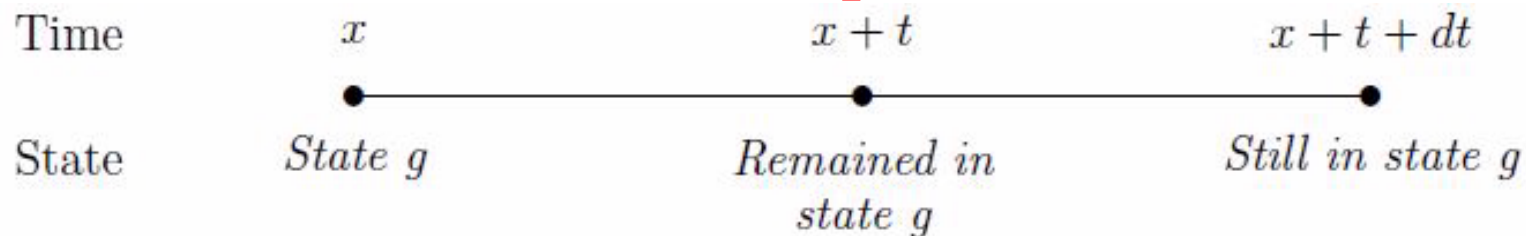
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- We can also apply the same argument to finding the Kolmogorov equation for ${}_t p_x^{\overline{gg}}$, the probability that the state g is occupied continuously from time x to time $x + t$.
- As in the previous examples, the resulting differential equation can be solved to give a closed form expression for ${}_t p_x^{\overline{gg}}$. The diagram

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tells us that (next slide):

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$${}_{t+dt}p_x^{\overline{gg}} = {}_tp_x^{\overline{gg}} \times \left(1 - \sum_{j \neq g} \mu_{x+t}^{gj} dt + o(dt) \right) \quad (35)$$

- Rearranging we get

$$\begin{aligned} \frac{{}_{t+dt}p_x^{\overline{gg}} - {}_tp_x^{\overline{gg}}}{dt} &= -{}_tp_x^{\overline{gg}} \sum_{j \neq g} \mu_{x+t}^{gj} + \frac{o(dt)}{dt} \\ \Rightarrow \frac{\partial}{\partial t} {}_tp_x^{\overline{gg}} &= -{}_tp_x^{\overline{gg}} \sum_{j \neq g} \mu_{x+t}^{gj} \end{aligned} \quad (36)$$

on letting $dt \rightarrow 0$.

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- Integrating (36) we find

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$${}_t p_x^{gg} = \exp \left(- \int_0^t \sum_{j \neq g} \mu_{x+s}^{gj} ds \right) \quad (37)$$

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Comment: This formula generalises the well-known formula:

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$${}_t p_x = \exp \left(- \int_0^t \mu_{x+s} ds \right). \quad (38)$$