# STAT 513/413: Lecture 17 Markov chain Monte Carlo: practical use

(including crash course in Bayesian statistics)

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## Why bother at all?

The possibility of generating from  $\ell(x)$  without knowing the exact value of  $\int \ell(x) \, \mathrm{d}x$ 

Who needs that??

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## Bayesian statistics ("inverse probability")

We observe x, as an outcome of a random variable X that has probability distribution depending on  $\vartheta$ 

Unlike x, the value of  $\vartheta$  is unknown - but we would like to say something about it on the basis of x

To this end, we need to give  $\vartheta$  a probability distribution,  $g(\vartheta)$ 

- which does not have an interpretation in any repeated or potentially repeated property But jeth example of our uncertainty about  $\vartheta$ ; in particular,  $g(\vartheta)$  is interpreted as the "initial", prior distribution: quantities our powerdents observing any x

The distribution of x can be then viewed as  $f(x|\vartheta)$ , conditional distribution of x given  $\vartheta$  we chat powcoder

(all of those may be densities or probability mass functions)

We look then at  $g(\vartheta|x)$ , the conditional distribution of  $\vartheta$  given x -which expresses our uncertainty after observin x

We are thus "inverting the probability":

start with  $f(x|\vartheta)$  (and  $g(\vartheta)$ ), end up with  $g(\vartheta|x)$ 

How do we do that?

## The Bayes theorem

The probabilistic definition of conditional probability/density says:

$$f(x|\vartheta) = \frac{f(x,\vartheta)}{p(\vartheta)} \quad \text{where } f(x,\vartheta) \text{ is the } \textit{joint } \textit{distribution of } X \text{ and } \vartheta$$

(Note:  $f(x, \vartheta)$  is quantity different from  $f(x|\vartheta)$ )

We have then also:  $g(\vartheta|x) = \frac{f(\vartheta,x)}{e \xi t} = \frac{f(x,\vartheta)}{E \xi t}$ Assignment Project Exam Help

because the joint distribution of  $\vartheta$  and x is the same as the joint distribution of x and x are a straightful distribution of  $\varphi$  and  $\varphi$  are distribution of  $\varphi$  and  $\varphi$  are distribution of  $\varphi$  are distribution of  $\varphi$  and  $\varphi$  are distribution of  $\varphi$  and  $\varphi$  are distribution of  $\varphi$  are distribution of  $\varphi$  are distribution of  $\varphi$  are distribution of  $\varphi$  and  $\varphi$  are distribution of  $\varphi$  and  $\varphi$  are distribution of  $\varphi$  and  $\varphi$  are distribution of  $\varphi$  are distribution of  $\varphi$  and  $\varphi$  are distribution of  $\varphi$  are distribution of  $\varphi$  and  $\varphi$  are distribution of  $\varphi$  and  $\varphi$  are distribution of  $\varphi$  are distribution of  $\varphi$  and  $\varphi$  are distribution of  $\varphi$  are distribution of  $\varphi$  and  $\varphi$  are distribution of

How do we obtain f(x), the (marginal) distribution of x? Add WeChat powcoder

$$f(x) = \int f(x, \vartheta) d\vartheta$$
 (replace integral by sum if necessary)

Thus:

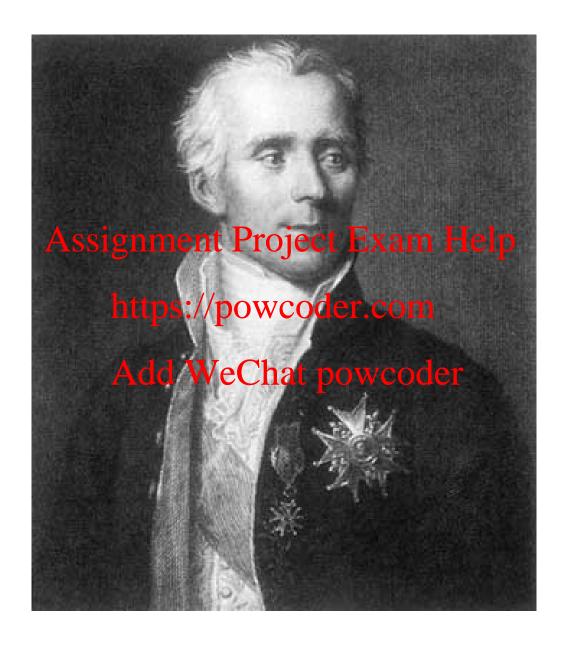
$$g(\vartheta|x) = \frac{f(x,\vartheta)}{f(x)} = \frac{f(x,\vartheta)}{\int f(x,\vartheta) d\vartheta} = \frac{f(x|\vartheta)g(\vartheta)}{\int f(x|\vartheta)g(\vartheta) d\vartheta}$$

Numerator is easy - but denominator may be difficult

# **Reverend Thomas Bayes**



# Pierre-Simon Laplace



### **Example**

We observe the outcomes of  $\mathfrak n$  0-1 independent random variables  $X_{\mathfrak i}$ , which have the same probability  $\vartheta=P[X_{\mathfrak i}=1]$ ; we are interested only in the number of 1's among those

Our X is thus the sum 
$$X = \sum_{i} X_{i}$$

which has binomial distribution Project Exam Help  $\vartheta^x(1-\vartheta)^{n-x}$ 

We assume  $\vartheta$  to have the best dependent of the size of the size

$$g(\vartheta) = \frac{\vartheta^{\alpha-1}(1-\vartheta)^{\beta-}\text{Add} \ \text{WeChat powcoder}}{B(\alpha,\beta)} \text{ for } \vartheta \in [0,1] \text{ (otherwise } g(\vartheta) = 0)$$

Why? Good question... What is  $B(\alpha, \beta)$ ?

Well, something that makes  $g(\vartheta)$  integrating to 1

- otherwise, we do not have to worry about it...

... and be a bit more relaxed

## A blasé way of doing Bayesian derivations

A wondrous symbol:  $\infty$  reads "is proportional to" and means "up to a constant, dependent on context, equal to"

$$g(\vartheta) \propto \vartheta^{\alpha-1}(1-\vartheta)^{\beta-1}$$

$$f(x|\vartheta) \propto \vartheta^x (1-\vartheta)^{n-x}$$

$$f(x,\vartheta) \propto f(x|\vartheta)g(\vartheta) \propto \vartheta^{x+\alpha-1}(1-\vartheta)^{n-x+\beta-1}$$

$$\begin{array}{cccc} f(x,\vartheta) & \propto & f(x|\vartheta)g(\vartheta) \propto \vartheta^{x+\alpha-1}(1-\vartheta)^{n-x+\beta-1} \\ g(\vartheta|x) & \propto & f(x,\vartheta) \end{array}$$
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really: as  $g(\vartheta|x) = \frac{https:/powcoder.com}{f(x)}$  and f(x) is with respect to  $\vartheta$  only a constant (albaid by ending powooder

$$g(\vartheta|x) \propto \vartheta^{x+\alpha-1}(1-\vartheta)^{n-x+\beta-1}$$

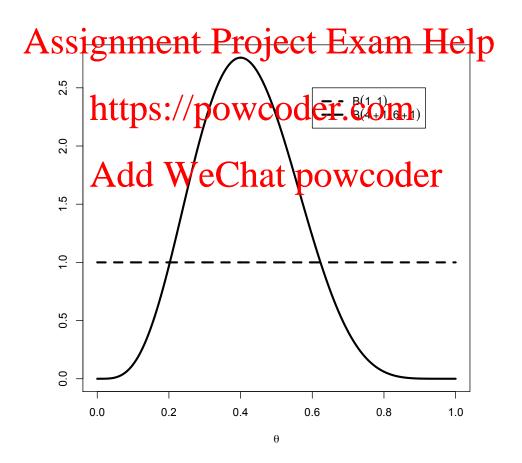
Now you perhaps can guess why  $g(\vartheta)$  was selected  $B(\alpha, \beta)$ : merely out of convenience. Which distribution is  $\propto \vartheta^{x+\alpha-1}(1-\vartheta)^{n-x+\beta-1}$ ? It is B( $\alpha + x$ ,  $\beta + n - x$ )

There is no need to calculate  $\binom{n}{x}\vartheta^x(1-\vartheta)^{n-x}\frac{\vartheta^{\alpha-1}(1-\vartheta)^{\beta-1}}{B(\alpha,\beta)}d\vartheta$ 

#### For instance

Out of 10 trials, we observed 4 ones and 6 zeros For  $g(\vartheta)$ , we put B(1,1) - assuming "we know nothing":  $g(\vartheta) \propto \vartheta^{1-1}(1-\vartheta)^{1-1} = 1$ 

- which is nothing but uniform distribution on [0, 1]



#### So what?

Well, these are somewhat lucky, and not that frequent circumstances; albeit in this case it can be argued that the family of beta distributions has many members, allowing thus for flexible "expression of initial uncertainty" about  $\vartheta$ , it it is quite clear that there will be many more instance which will not go that smoothly

Once again, first steps are easy: once  $g(\vartheta)$  and  $f(x|\vartheta)$  are specified, then they are multiplied ment Project Exam Help

But the results has to be integrated, which may be difficult - even numerically (and even the combine of the co

Yes, but why do we **Agentive (thappotovite der**all? The product is proportional to  $g(\vartheta|x)$  - so it may perhaps tell us something about  $\vartheta$  itself?

Well, it can indicate the shape ("where  $\vartheta$  is concentrated"), but if we need to calculate something out of  $g(\vartheta|x)$ 

For instance, the mean:

it is known that the mean of  $B(\alpha, \beta)$  is  $\frac{\alpha}{\alpha + \beta}$ 

## If only we could do Monte Carlo

Imagine that we know  $g(\vartheta|x)$ , and can generate random numbers Z<sub>i</sub> out of this distribution

Then we can get estimates of its expected value: 
$$\frac{1}{n}\sum_{i} Z_{i}$$

Well, we know  $g(\vartheta|x)$  up to a constant... Assignment Project Exam Help ... but neither inversion nor acceptance/rejection method would work without the knowledge of the constant! https://powcoder.com

But Markov chain Month Chlappithms det!!

## **Another example**

Suppose the data come from the distribution

$$f(x|\vartheta) = \frac{1}{c\sqrt{2\pi}} e^{-\frac{(x-\vartheta)^2}{2c}}$$
 that is,  $N(\vartheta, c^2)$ 

and the prior distribution of  $\vartheta$  is

$$g(\mu) = \frac{1}{s\sqrt{25}} \operatorname{sign}^{\frac{(\vartheta - m)^2}{25}} \operatorname{Project}^{\frac{1}{25}} \operatorname{Exian}^{\frac{(\vartheta - m)^2}{25}}$$

where c, m, s are consittered/prowoodenstants

A possible story: measuring that power at in power at the first power at m = 100 (by definition), with s = 15 (approximately; US 1970's). We are measuring IQ by tests: their outcomes are  $x_1, \ldots, x_n$ , each with standard deviation c

(In practice we often observe note one x, but several of those. However, the notation is easier with one.)

#### So... the result is sort of intuitive

$$f(x|\vartheta) \propto e^{-\frac{(x-\vartheta)^2}{2c^2}}$$

$$g(\vartheta)$$
  $\propto$   $e^{-\frac{(\vartheta-m)^2}{2s^2}}$ 

$$f(x,\vartheta) \propto e^{-\frac{(x-\vartheta)^2}{2c^2}} e^{-\frac{(\vartheta-m)^2}{2s^2}} = e^{-\frac{(x-\vartheta)^2}{2c^2} - \frac{(\vartheta-m)^2}{2s^2}} = e^{A\vartheta^2 + B\vartheta + C}$$

with A < 0 - another normal distribution

The result is N 
$$\left(\frac{\underset{\overline{c^2}+\overset{1}{s^2}}{\underbrace{\frac{1}{c^2}+\overset{1}{s^2}}},\underset{\overline{c^2}+\overset{1}{\underbrace{\frac{1}{c^2}+\overset{1}{s^2}}}}{\underbrace{\frac{1}{c^2}+\overset{1}{\underbrace{\frac{1}{c^2}+\overset{1}{s^2}}}}}\right) = N \left(\frac{\underset{\overline{c^2}+\overset{1}{\underline{s^2}}}{\underbrace{\frac{1}{c^2}+\overset{1}{\underline{s^2}}}}}{\underbrace{\frac{1}{c^2}+\overset{1}{\underline{s^2}}}}\right)$$

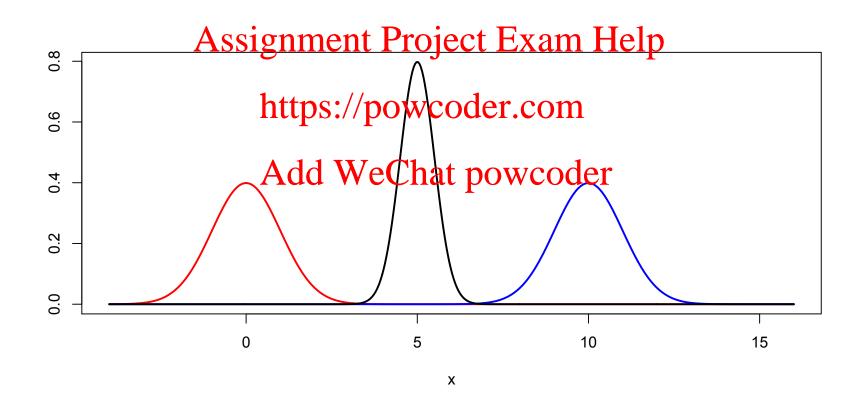
More generally, if we have n the stantipow todores  $x_1, x_2, \ldots, x_n$ , each with standard deviation c, then the result is

$$N\left(\frac{\frac{n\bar{x}}{c^2} + \frac{m}{s^2}}{\frac{n}{c^2} + \frac{1}{s^2}}, \frac{1}{\frac{n}{c^2} + \frac{1}{s^2}}\right) = N\left(\frac{\frac{\bar{x}}{c^2} + \frac{m}{ns^2}}{\frac{1}{c^2} + \frac{1}{ns^2}}, \frac{1}{\frac{n}{c^2} + \frac{1}{s^2}}\right)$$

where  $\bar{x}$  is the average of the  $x_i\sp{'s}$  - and this has a natural interpretation

## But sometimes a bit funny

Suppose that  $ns=c^2$ ; then the result is  $N\left(\frac{\overline{x}+m}{2},\frac{c^2}{2}\right)$  If, say, n=1, s=c=1, m=0, and x=10 then the result looks like this



Have you got a good story for that? (There are bad stories...)

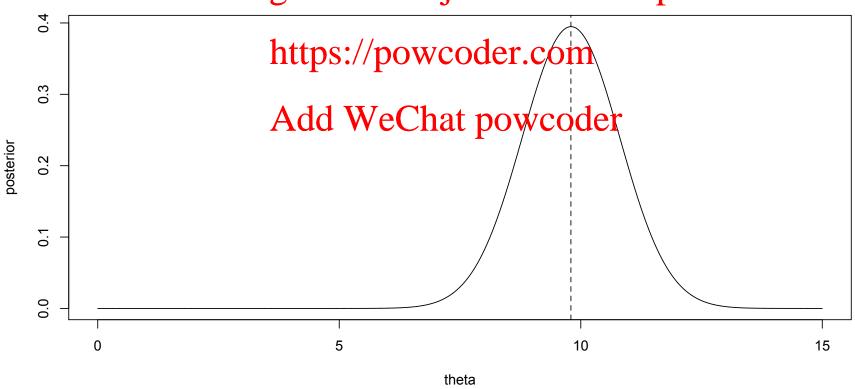
#### A fix?

If, instead of the normal distribution used before,  $g(\vartheta)$  is a t distribution, say, with 1 degree of freedom (the Cauchy distribution)

$$g(\vartheta) = \frac{1}{\pi(1+\vartheta^2)}$$
 (this distribution is centered at 0)

and as before, x = 10, where x is  $N(\vartheta, 1)$ , then the result is

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## Markov chain Monte Carlo: concluding remarks

But such fixes have to be done numerically: and Markov chain Monte Carlo is a convenient vehicle for that, as it allows to generate random numbers following a density  $c\ell(x)$  without knowing c.

It is still a methodology in development; compared to classical Monte Carlo methods, there are still a couple of aspects that have to be paid attention to, and adequately tuned.

Burn-in. The Assignment Project Example follow the desired distribution closely only later in the sequence - thus the initial random numbers are discarded the weyer it is hard to say exactly how many.

Convergence. Unlike in the independent case, there are no performance guarant eddlike Chabye wcode! Hoeffing inequalities. To assess whether the generated sequence is long enough, some empirical criteria have been proposed. (The book of Rizzo mentions Gelman-Rubin.)

There are also other, more specific aspects to be tuned. For instance, when Metropolis algorithm uses random walk as  $q_x(y)$  (a "proposal transition distribution"), then it is important to scale it (set the length of the average step) so that it is neither too small nor too big (which is expressed by the rejection rate).

## Monte Carlo: concluding remarks

We bid a farewell to random numbers at this point - in a hope that what we have seen so far reinforced certain principles, rather than gave mere recipes, how these numbers can be effectively used in numerical computations. In particular, we hope that we understand more the following aspects now:

- how these numbers are generated by computers
- how they can Assigned in Project dax amplifients that aim at exploration of the properties of probabilistic phenomena that are hard to be obtained in the properties of probabilistic phenomena that are methods
- how they can be used for numerical calculations that may be hard to accomplish by classical numerical methods, like the computation of integrals; and how these methods can be improved
- how to obtain, in certain cases, validation criteria, precision estimates, and performance guarantees
- and also, what are the principles of some more specific, and upto-date technologies, applicable in specific situations - like Markov chain Monte Carlo methods