

Homework 3

Exercise 1

THE JOINT PDF OF Y_1 AND Y_2 IS:

$$f(y_1, y_2) = f(y_1) \cdot f(y_2) = e^{-y_1} \cdot e^{-y_2} = e^{-(y_1 + y_2)}$$

(INDEPENDENT)

$$\text{LET } U = \frac{Y_1}{Y_1 + Y_2} \quad \text{AND} \quad V = Y_2$$

$$\text{IT FOLLOWS THAT } Y_2 = V \quad \text{AND} \quad Y_1 = \frac{UV}{1-U}$$

JACOBIAN:

$$J = \begin{vmatrix} \frac{\partial V}{\partial Y_1} & \frac{\partial V}{\partial Y_2} \\ \frac{\partial U}{\partial Y_1} & \frac{\partial U}{\partial Y_2} \end{vmatrix} = \begin{vmatrix} 1 & 1 \\ \frac{Y_2}{(Y_1 + Y_2)^2} & -\frac{Y_1}{(Y_1 + Y_2)^2} \end{vmatrix} = \frac{(1-U)^2}{V}$$

THEREFORE THE JOINT PDF OF U AND V IS:

$$f(u, v) = e^{-\left(\frac{uv}{1-u} + v\right)} \cdot \frac{v}{(1-u)^2} = e^{-\frac{v}{1-u}} \cdot \frac{v}{(1-u)^2}$$

TO FIND PDF OF U INTEGRATE THE JOINT PDF W.R.T. V :

$$f(u) = \int_0^{\infty} e^{-\frac{v}{1-u}} \cdot \frac{v}{(1-u)^2} dv = 1.$$

EXERCISE 2 :

(a) $X \sim N(0,1)$, $Y \sim N(0,1)$ X, Y ARE INDEPENDENT

JOINT PDF OF X, Y :

$$f(x,y) = f(x) \cdot f(y) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}y^2}$$
$$= \frac{1}{2\pi} e^{-\frac{1}{2}x^2 - \frac{1}{2}y^2}$$

$U = X+Y$
 $V = X-Y$ $\left\{ \begin{array}{l} \rightarrow X = \frac{U+V}{2} \text{ (AM)} \\ \rightarrow Y = \frac{U-V}{2} \end{array} \right.$

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JACOBIAN :

$$J = \begin{vmatrix} \frac{\partial U}{\partial X} & \frac{\partial U}{\partial Y} \\ \frac{\partial V}{\partial X} & \frac{\partial V}{\partial Y} \end{vmatrix} = \begin{vmatrix} 1 & 1 \\ 1 & -1 \end{vmatrix} = -2$$

JOINT PDF OF U, V :

$$f(u,v) = \frac{1}{2\pi} e^{-\frac{1}{2}\left(\frac{u+v}{2}\right)^2} \cdot e^{-\frac{1}{2}\left(\frac{u-v}{2}\right)^2} \cdot |-2|$$

THIS IS SIMPLIFIED TO :

$$f(u,v) = \frac{1}{\sqrt{2}\sqrt{2\pi}} e^{-\frac{1}{2}\frac{u^2}{2}} \cdot \frac{1}{\sqrt{2}\sqrt{2\pi}} e^{-\frac{1}{2}\frac{v^2}{2}} \rightarrow$$

EXERCISE 3 :

$$X \sim \text{beta}(\alpha, \beta) \quad , \quad Y \sim \text{beta}(\alpha + \beta, \gamma)$$

$$f(x, y) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1} \cdot \frac{\Gamma(\alpha + \beta + \gamma)}{\Gamma(\alpha + \beta)\Gamma(\gamma)} x^{\alpha + \beta - 1} (1-x)^{\gamma-1}$$

$$\begin{cases} u = xy \\ v = x \end{cases} \quad \begin{cases} y = \frac{u}{v} \\ x = v \end{cases}$$

$$J = \begin{vmatrix} \frac{\partial u}{\partial x} & \frac{\partial u}{\partial y} \\ \frac{\partial v}{\partial x} & \frac{\partial v}{\partial y} \end{vmatrix} = \begin{vmatrix} y & x \\ 1 & 0 \end{vmatrix} = -x$$

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$$f(u, v) = \frac{\Gamma(\alpha + \beta + \gamma)}{\Gamma(\alpha)\Gamma(\beta)\Gamma(\gamma)} v^{\alpha-1} v^{\beta-1} \left(\frac{u}{v}\right)^{\alpha + \beta - 1} \left(\frac{u}{v}\right)^{\gamma-1} \frac{1}{v}$$

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Exercise 4

(a) CDF of R :

$$P(R \leq r) = 1 - P(R > r) = 1 - P\left(X = 0 \text{ PARTICLES IN SPHERE WITH VOLUME } \frac{4}{3}\pi r^3\right)$$

$$\text{HERE } \lambda = \rho \frac{4}{3}\pi r^3$$

$$= 1 - \frac{e^{-\rho \frac{4}{3}\pi r^3}}{1} = 1 - e^{-\rho \frac{4}{3}\pi r^3}$$

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$$f(r) = 4\rho\pi r^2 e^{-\rho \frac{4}{3}\pi r^3}$$



$$(b) \text{ Let } U = R^3$$

$$F_U(u) = P(U \leq u) = P(R^3 \leq u)$$

$$= P(R \leq u^{1/3}) = F_R(u^{1/3})$$

$$f_U(u) = \frac{1}{3} u^{-2/3} 4\lambda\pi u^{2/3} e^{-\lambda \frac{4}{3}\pi u}$$

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$$f_U(u) = \frac{4}{3}\lambda\pi e^{-\frac{4}{3}\lambda\pi u}$$

\therefore

$$U \sim \text{Exp}\left(\frac{4}{3}\lambda\pi\right)$$

$$\text{And } E U = \frac{3}{4\lambda\pi}$$



Exercise 5

$$X_1 \sim \chi_{r_1}^2$$

$$X_2 \sim \chi_{r_2}^2$$

$$f(x_1, x_2) = f(x_1) f(x_2) = \frac{x_1^{\frac{r_1}{2}-1} e^{-x_1/2}}{\Gamma(\frac{r_1}{2}) 2^{r_1/2}} \cdot \frac{x_2^{\frac{r_2}{2}-1} e^{-x_2/2}}{\Gamma(\frac{r_2}{2}) 2^{r_2/2}}$$

$$J = \begin{vmatrix} \frac{\partial y_1}{\partial x_1} & \frac{\partial y_1}{\partial x_2} \\ \frac{\partial y_2}{\partial x_1} & \frac{\partial y_2}{\partial x_2} \end{vmatrix} = \begin{vmatrix} \frac{1}{x_2} & -\frac{x_1}{x_2^2} \\ 1 & 1 \end{vmatrix} = \frac{1}{x_2} + \frac{x_1}{x_2^2} = \frac{x_2 + x_1}{x_2^2} = \frac{y_2}{x_2^2}$$

$$y_1 = \frac{x_1}{x_2} \quad \left. \begin{array}{l} \\ \\ \end{array} \right\} \quad x_1 = \frac{y_1 y_2}{1+y_1}$$

$$y_2 = \frac{x_2}{x_1+x_2} \quad \left. \begin{array}{l} \\ \\ \end{array} \right\} \quad x_2 = \frac{y_2}{\frac{1}{1+y_1} + 1} = \frac{y_2}{1+y_1}$$

$$f(y_1, y_2) = \frac{\left(\frac{y_1 y_2}{1+y_1}\right)^{\frac{r_1}{2}-1} e^{-\frac{y_1 y_2}{2(1+y_1)}} \left(\frac{y_2}{1+y_1}\right)^{\frac{r_2}{2}-1} e^{-\frac{y_2}{2(1+y_1)}}}{\Gamma(\frac{r_1}{2}) 2^{r_1/2} \Gamma(\frac{r_2}{2}) 2^{r_2/2} (1+y_1)^2 y_2} \cdot \frac{\Gamma(\frac{r_1+r_2}{2})}{\Gamma(\frac{r_1}{2}) \Gamma(\frac{r_2}{2})}$$

$$= \frac{y_1^{\frac{r_1}{2}-1} (1+y_1)^{-\frac{r_1+r_2}{2}} \Gamma(\frac{r_1+r_2}{2})}{\Gamma(\frac{r_1}{2}) \Gamma(\frac{r_2}{2})} \cdot \frac{y_2^{\frac{r_1+r_2}{2}-1} e^{-y_2/2}}{\Gamma(\frac{r_1+r_2}{2}) 2^{r_1+r_2}}$$

$$= f(y_1) \cdot f(y_2)$$

$\therefore y_1, y_2$ ARE INDEPENDENT

$$\text{AND } y_2 \sim \chi_{r_1+r_2}^2$$

SINCE $\frac{x_1}{x_2}$ IND OF X_3 (GIVEN) $\left. \begin{array}{l} \\ \\ \end{array} \right\} \frac{x_1/r_1}{x_2/r_2}$ IND OF $\frac{X_3/r_3}{x_1+x_2/(r_1+r_2)}$

$$\text{AND } \frac{x_1}{x_2} \text{ IND OF } x_1+x_2$$

Exercise 6

$$\begin{aligned}
 (a), P(X_1, \dots, X_{r-1} | X_r) &= \frac{P(X_1, \dots, X_r)}{P(X_r)} \quad \text{NOTE: } X_r \sim \text{bin}(n, p_r) \\
 &= \frac{\frac{n!}{x_1! x_2! \dots x_r!} p_1^{x_1} p_2^{x_2} \dots p_r^{x_r}}{\frac{n!}{(n-x_r)! x_r!} p_r^{x_r} (1-p_r)^{n-x_r}} \\
 &= \frac{(n-x_r)!}{x_1! x_2! \dots x_{r-1}!} \left(\frac{p_1}{1-p_r} \right)^{x_1} \dots \left(\frac{p_{r-1}}{1-p_r} \right)^{x_{r-1}}
 \end{aligned}$$

$\therefore \sim \text{Multinomial}$ with $n-x_r$ trials
 AND PROBABILITIES $\frac{p_1}{1-p_r}, \frac{p_2}{1-p_r}, \dots, \frac{p_{r-1}}{1-p_r}$

Exercise 7

$$X_i | p_i \sim \text{Bernoulli}(p_i), \quad p_i \sim \text{beta}(\alpha, \beta)$$

$$(a) \quad 1. E X_i = E(E(X_i | p_i)) = E p_i = \frac{\alpha}{\alpha + \beta}$$

$$E \sum X_i = n \frac{\alpha}{\alpha + \beta}$$

$$2. \text{VAR}(X_i) = E(\text{VAR}(X_i | p_i)) + \text{VAR}(E(X_i | p_i))$$

$$= E(p_i(1-p_i)) + \text{VAR}(p_i)$$

$$= E p_i - E p_i^2 + \text{var}(p_i)$$

$$= \frac{\alpha}{\alpha+b} - \left(\frac{\alpha^2}{(\alpha+b)^2} + \frac{\alpha b}{(\alpha+b)^2(\alpha+b+1)} \right) + \frac{\alpha b}{(\alpha+b)^2(\alpha+b+1)}$$

$$= \frac{\alpha(\alpha+b)(\alpha+b+1) - \alpha^2(\alpha+b+1) - \alpha b + \alpha b}{(\alpha+b)^2(\alpha+b+1)} = \frac{\alpha b}{(\alpha+b)^2}$$

$$\therefore \text{var}\left(\sum X_i\right) = n \frac{\alpha b}{(\alpha+b)^2}$$

(b)

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$X_i | p_i \sim b(n_i, p_i)$, $p_i \sim \text{beta}(\alpha, b)$

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$$E Y = \frac{\alpha}{\alpha+b} \sum n_i$$

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$$\text{var}(Y) = \sum \text{var}(X_i)$$

$$\text{where } \text{var}(X_i) = n_i \frac{\alpha b (\alpha+b+n_i)}{(\alpha+b)^2 (\alpha+b+1)}$$

Exercise 8

CDF of $Q = \prod X_i$

$$F_Q(q) = P(Q \leq q)$$

$$= P(\prod X_i \leq q)$$

$$= P(-\ln \prod X_i > \ln q)$$

$$= P(-\ln X_1 - \dots - \ln X_n > \ln q)$$

Now each $-\ln X_i \sim \text{Exp}(1)$

$$(1 - F_T(-\ln q))$$

Then $T = -\sum \ln X_i \sim \Gamma(n, 1)$

THUS, $F_Q(q) = 1 - F_T(-\ln q)$

$$f(q) = -f_T(-\ln q) = \frac{1}{q} \frac{(-\ln q)^{n-1}}{\Gamma(n)} e^{-(-\ln q)}$$