

Coursework 3: STAT 570

Philip Pham

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1. Consider the Poisson-gamma random effects model given by

$$Y_i \mid \mu_i, \theta_i \sim \text{Poisson}(\mu_i \theta_i) \quad (1)$$

$$\theta_i \sim \text{Gamma}(b, b), \quad (2)$$

which leads to a negative binomial marginal model with the variance a quadratic function of the mean. Design a simulation study, along the lines of that which produced Table 2.3 in the book (overdispersed Poisson example) to investigate the efficiency and robustness under

- a Poisson model;
- quasi-likelihood with $\mathbb{E}[Y] = \mu$ and $\text{Var}(Y) = \alpha\mu$; and
- sandwich estimation.

Use a log-linear model

$$\log \mu_i = \beta_0 + \beta_1 x_i, \quad (3)$$

with $x_i \sim_{\text{iid}} \mathcal{N}(0, 1)$ for $i = 1, 2, \dots, n$, and $\beta_0 = -2$ and $\beta_1 = \log 2$.

Simulate for:

- $b \in \{0.2, 1, 10, 1000\}$.
- $n \in \{10, 20, 50, 100, 250\}$.

Summarize what your take away message is after carrying out these simulations.

Solution: Note that

$$\begin{aligned} \mathbb{P}(Y_i = y \mid \mu_i) &= \int_0^\infty \mathbb{P}(Y_i = y \mid \mu_i, \theta_i = \theta) \mathbb{P}(\theta_i = \theta \mid b) \, d\theta \\ &= \int_0^\infty \left(\frac{(\mu_i \theta)^y}{y!} \exp(-\mu_i \theta) \right) \left(\frac{b^b}{\Gamma(b)} \theta^{b-1} \exp(-b\theta) \right) \, d\theta \\ &= \frac{\mu_i^y b^b}{y! \Gamma(b)} \int_0^\infty \theta^{b+y-1} \exp(-\theta(b + \mu_i)) \, d\theta \\ &= \frac{\Gamma(y+b)}{y! \Gamma(b)} \frac{\mu_i^y b^b}{(\mu_i + b)^{b+y}} = \frac{\Gamma(y+b)}{y! \Gamma(b)} \left(\frac{b}{\mu_i + b} \right)^b \left(\frac{\mu_i}{\mu_i + b} \right)^y \\ &\sim \text{NegativeBinomial} \left(b, \frac{\mu_i}{\mu_i + b} \right). \end{aligned} \quad (4)$$

By properties of the negative binomial distribution, we have that

$$\begin{aligned}\mathbb{E}[Y_i | x_i] &= \mu_i = \exp(\beta_0 + \beta_1 x_i) \\ \text{Var}(Y_i | x_i) &= \mu_i \left(1 + \frac{\mu_i}{b}\right).\end{aligned}\tag{5}$$

Thus, smaller values of b correspond to more dispersion.

Poisson Model

In the Poisson model, we assume that $\text{Var}(Y_i | x_i) = \mu_i$, e.g. $b \rightarrow \infty$.

In this case, the log-likelihood function is

$$l(\beta) = \sum_{i=1}^n \left[y_i (\beta_0 + \beta_1 x_i) - \exp(\beta_0 + \beta_1 x_i) - \sum_{k=1}^{y_i} \log k \right], \tag{6}$$

which gives us the score function

$$S(\beta) = \sum_{i=1}^n \begin{pmatrix} y_i - \exp(\beta_0 + \beta_1 x_i) \\ x_i y_i - x_i \exp(\beta_0 + \beta_1 x_i) \end{pmatrix}. \tag{7}$$

We can estimate β by solving for $S(\hat{\beta}) = \mathbf{0}$, numerically.

We can estimate the variance of the estimates from the Fisher information,

$$\begin{aligned}\text{Var}(\hat{\beta}) &\approx I_n(\hat{\beta})^{-1} \\ &= \left(\sum_{i=1}^n \begin{pmatrix} \exp(\hat{\beta}_0 + \hat{\beta}_1 x_i) & x_i \exp(\hat{\beta}_0 + \hat{\beta}_1 x_i) \\ x_i \exp(\hat{\beta}_0 + \hat{\beta}_1 x_i) & x_i^2 \exp(\hat{\beta}_0 + \hat{\beta}_1 x_i) \end{pmatrix} \right)^{-1} \\ &= \frac{1}{(\sum_{i=1}^n \hat{\mu}_i) (\sum_{i=1}^n x_i^2 \hat{\mu}_i) - (\sum_{i=1}^n x_i \hat{\mu}_i)^2} \begin{pmatrix} \sum_{i=1}^n x_i^2 \hat{\mu}_i & -\sum_{i=1}^n x_i \hat{\mu}_i \\ -\sum_{i=1}^n x_i \hat{\mu}_i & \sum_{i=1}^n \hat{\mu}_i \end{pmatrix},\end{aligned}\tag{8}$$

where $\hat{\mu}_i = \exp(\hat{\beta}_0 + \hat{\beta}_1 x_i)$.