A Comparison of Skew-Normal and Skew-t Regression Models

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Objectives

We examine the properties of skew-normal and skew-t models from both a Bayesian and frequentist perspective, and investigate the computational tools available for fitting these models. We apply skew-normal and skew-t models to data from the Nurture study, a cohort of mothers who gave birth between 2013 and 2016.

Introduction

In many applications of classical linear regression, the distribution of residuals exibits non- normal qualities such as skewness or heavy tails, making the assumption of normal error terms difficult to justify. The common statistical suggestion in these cases is to implement a transformation of the response variable, but this can result in a loss of interpretability. The skew-elliptical family is a broad class of probability distributions that contain the normal distribution as a special case and allow for flexible modeling when data exhibit skewness.

Definitions

Let ϕ and Φ be the standard normal pdf and cdf, respectively. Azzalini (1985) defined the density of a skew-normal random variable Z follows.

$$f(z; \lambda) = 2\phi(z)\Phi(\lambda z)$$

Similar to the construction of the familiar student's t random variable, we (cite) can define a skew-t random variable as the ratio of a skew normal and the square root of a χ^2 divided by its degrees of freedom. The resulant density is

$$t(x; \lambda, \nu) = 2t_0(x; \nu)T_0(\lambda x \sqrt{\frac{\nu+1}{\nu+x^2}}; \nu+1)$$

where t_0 and T_0 are the density and mass functions of the student's t distribution, respectively.

Motivation

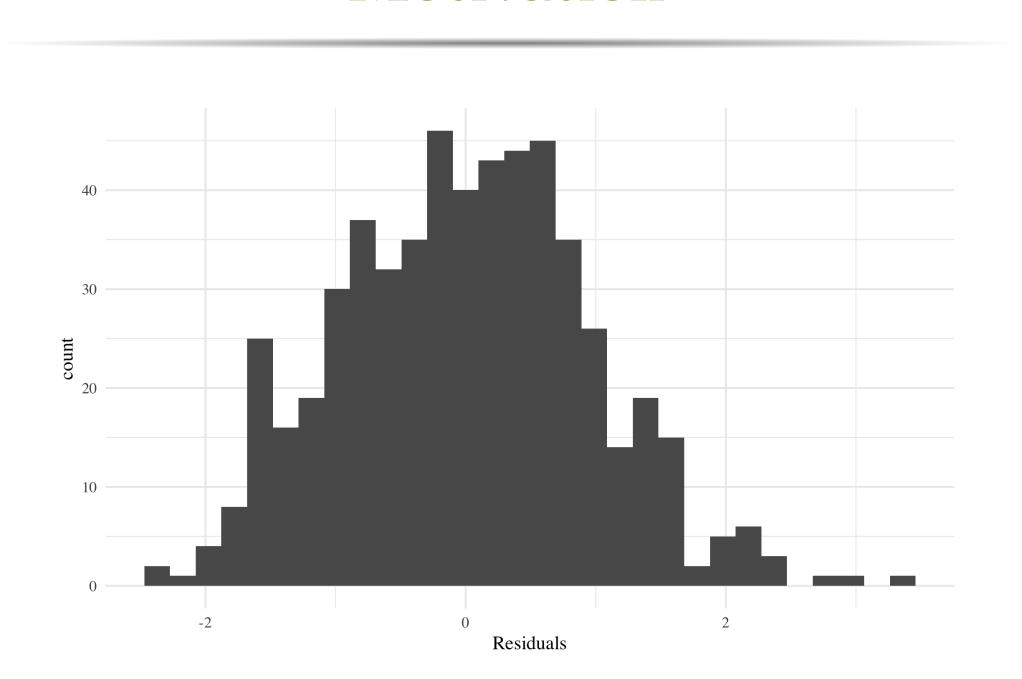


Figure 1: Distribution of residuals exhibiting skewness. Pearson's skewness = 0.17. Shapiro-Wilk p-value = 0.08.

Comparison Criteria

- Ability to faithfully reproduce underlying model parameters: Using simulated test data generated according to a deterministic equation with known parameters, is the imputation method able to faithfully replicate those underlying parameters when artificial censoring and missingness is introduced?
- Predictive Bias: $E(\hat{Y} Y)$, where \hat{Y} is the predicted value and Y is the true value.
- $Predictive \ Variance: \ Var(\hat{Y} Y)$
- Convergence Properties: Does the imputation method converge to any stable state?
- **6** Computational Efficiency: Is the algorithm scalable to larger data sets?

Important Results

Enterococcus counts are significantly influenced by precipication, tidal stage, and location. The best performing imputation method of the three tested employs sampling from a truncated normal distribution.

Imputation Methods

- A naive approach: Censored values are sampled from a Uniform(0, 10) and missing values are replaced with the overall mean Enterococcus for 2013-2015.
- values are sampled from a truncNorm() distribution, and missing values are predicted by the model made from all non-missing values.
- Sexpected value of a Truncated Normal: Let Y = Enterococcus concentrations. If $Y \sim logNormal(\mu, \sigma^2) \implies ln(Y) \sim Normal(\mu, \sigma^2)$, it can be shown that $E[ln(Y)|a < ln(Y) < b] = \mu + \sigma \frac{\phi(\frac{a-\mu}{\sigma}) \phi(\frac{b-\mu}{\sigma})}{\Phi(\frac{b-\mu}{\sigma}) \Phi(\frac{a-\mu}{\sigma})}$. When $a = -\infty$ and b = ln(10), $E(ln(Y)|0 < ln(Y) < ln(10)) = \mu \sigma \frac{\phi(\frac{ln(10)-\mu}{\sigma})}{\Phi(\frac{ln(10)-\mu}{\sigma})}$

Imputation Results



Figure 2: Simulated model parameter estimates method 2

| Meth | nod Abs. | Err. Bias | Variance |
|------|----------|-----------|----------|
| 1 | 8.417 | 0.576 | 0.229 |
| 2 | 2.187 | -0.009 | 0.002 |
| 3 | 9.476 | -0.493 | 1.1619 |

Modeling Results

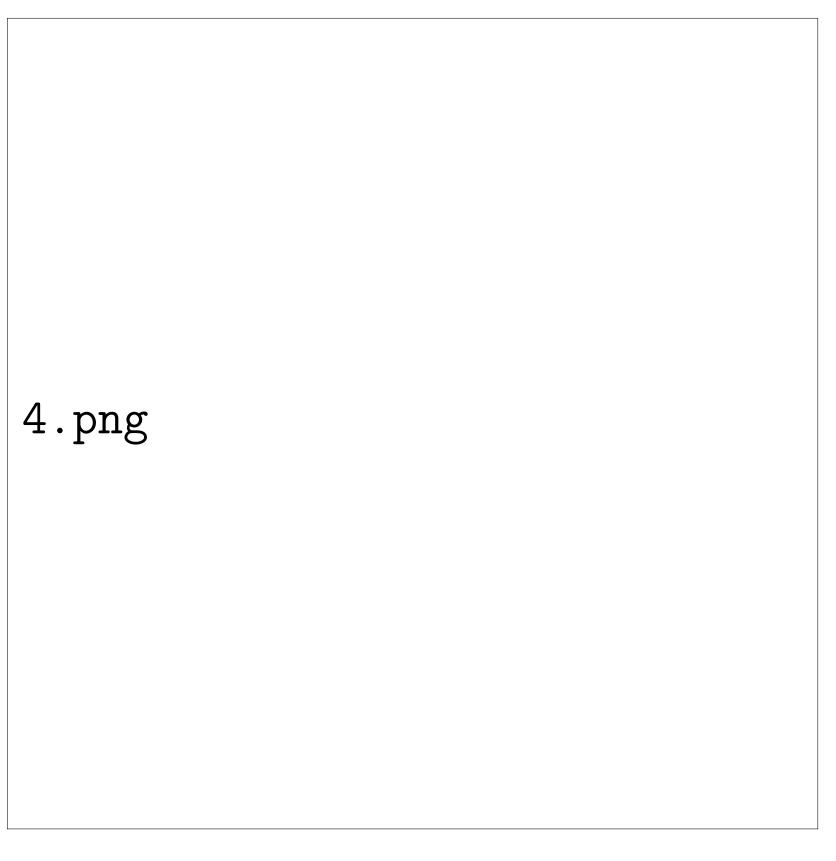


Figure 3: Model summary

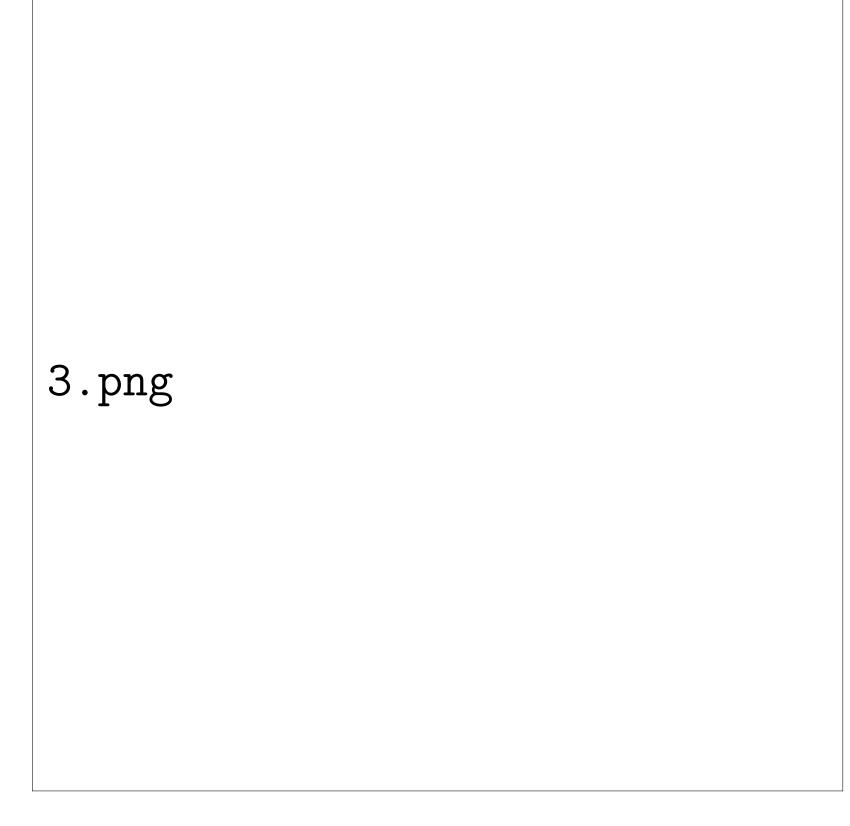


Figure 4: Residual analysis

Conclusion

Missing and censored values are best imputed by using Method 2, Sampling from a Truncated Normal. A general lineral model can be built after imputing according to Method 2, and used in the future to predict Enterococcus concentrations in Charleston, SC.

Further Resources

https://carter-allen.github.io