14 - Quadratic Classification and Classification for Several Populations

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Classification of Two Normal Populations When Covariances are Unequal

- Consider the multivariate normal densities with $\Sigma_i, i=1,2$, replacing Σ .
- Substituting multivariate normal densities with different covariance matrices gives

$$R_1 : -\frac{1}{2}\mathbf{x}'(\mathbf{\Sigma}_1^{-1} - \mathbf{\Sigma}_2^{-1})\mathbf{x} + (\boldsymbol{\mu}_1'\mathbf{\Sigma}_1^{-1} - \boldsymbol{\mu}_2'\mathbf{\Sigma}_2^{-1})\mathbf{x} - k$$

$$\geq \ln\left(\frac{c(1|2)}{c(2|1)} \cdot \frac{p_2}{p_1}\right)$$

$$R_2 : -\frac{1}{2}\mathbf{x}'(\mathbf{\Sigma}_1^{-1} - \mathbf{\Sigma}_2^{-1})\mathbf{x} + (\boldsymbol{\mu}_1'\mathbf{\Sigma}_1^{-1} - \boldsymbol{\mu}_2'\mathbf{\Sigma}_2^{-1})\mathbf{x} - k$$

$$< \ln\left(\frac{c(1|2)}{c(2|1)} \cdot \frac{p_2}{p_1}\right)$$

where

$$k = \frac{1}{2} \ln \left(\frac{|\boldsymbol{\Sigma}_1|}{|\boldsymbol{\Sigma}_2|} \right) + \frac{1}{2} (\boldsymbol{\mu}_1' \boldsymbol{\Sigma}_1^{-1} \boldsymbol{\mu}_1 - \boldsymbol{\mu}_2' \boldsymbol{\Sigma}_2^{-1} \boldsymbol{\mu}_2)$$

- The classification regions are define by quadratic functions of \mathbf{x} .
- When $m{\Sigma}_1 = m{\Sigma}_2$, the term $-\frac{1}{2} \mathbf{x}' (m{\Sigma}_1^{-1} m{\Sigma}_2^{-1}) \mathbf{x}$ disappears.

Quadratic Classification Rule

• Allocate to \mathbf{x}_0 to π_1 if

$$-\frac{1}{2}\mathbf{x}_{0}'(\boldsymbol{S}_{1}^{-1} - \boldsymbol{S}_{2}^{-1})\mathbf{x}_{0} + (\bar{\boldsymbol{x}}_{1}'\boldsymbol{S}_{1}^{-1} - \bar{\boldsymbol{x}}_{2}'\boldsymbol{S}_{2}^{-1})\mathbf{x}_{0} - k$$

$$\geq \ln\left(\frac{c(1|2)}{c(2|1)} \cdot \frac{p_{2}}{p_{1}}\right)$$

• Allocate to \mathbf{x}_0 to π_2 otherwise.

Play with the IRIS data

```
library(dplyr) # for data manipulation
# select variables, filter species
iris23 <- iris %>%
   select(Sepal.Length, Petal.Length, Species) %>%
   filter(Species != "setosa") %>%
   droplevels() # drop empty level (e.g. setosa)
```

Using qda() for Quadratic Classification/Discriminant Analysis

```
#
    versicolor virginica
# versicolor 47 3
# virginica 3 47
```

Quadratic and Linear Discrimination produces similar results for the IRIS data.

Example 11.1: Discriminating owners from nonowners of riding mowers

\$ size : num 18.4 16.8 21.6 20.8 23.6 19.2 17.6 22.4

\$ riding: Factor w/ 2 levels "owner", "nonowner": 1 1 1 1

\$ income: num 90 115.5 94.8 91.5 117 ...

#

Linear Classification/Discriminant Analysis

```
# owner 109 20
# nonowner 87 18

mower.pred.lda <- predict(mower.lda)
table(mower.pred.lda$class, mower$riding)</pre>
```

```
# owner nonowner
# owner 11 2
# nonowner 1 10
```

income size

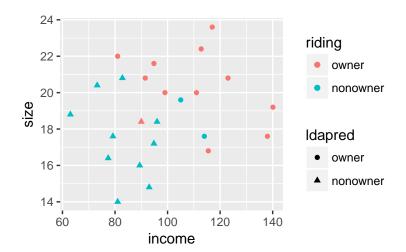
#

Add preicted class column to data

```
# add new predicted class column, save
mower <- mower %>%
  mutate(ldapred = mower.pred.lda$class)
head(mower)
```

```
#
   income size riding ldapred
# 1
      90
           18
              owner nonowner
# 2
     116 17
              owner
                      owner
# 3
    95 22
              owner
                      owner
# 4 92 21
              owner
                      owner
# 5 117 24
              owner
                      owner
# 6
      140 19
              owner
                      owner
```

Plots



Classification with Several Populations

• Let $f_i(\mathbf{x})$ be the density associated with population π_i , $i = 1, 2, \dots, g$.

$$p_i$$
 = the prior prob'y of population $\pi_i, i=1,\ldots,g$ $c(k|i)=$ cost of allocating an item to π_k when it belongs to $\pi_i, k\neq i$

• Let R_k be the set of x's classified as π_k and

$$P(k|i) = P(\text{classifying item as } \pi_k|\pi_i) = \int_{R_k} f_i(\mathbf{x}) d\mathbf{x}$$

$$P(i|i) = 1 - \sum_{\substack{k=1\\k\neq i}}^g P(k|i)$$

Conditional Expected Cost of Misclassifying (CECM)

• The conditional expected cost of misclassifying an \mathbf{x} from π_1 into π_2 , or π_3, \ldots , or π_k is

$$ECM(1) = P(2|1)c(2|1) + P(3|1)c(3|1) + \dots + P(g|1)c(g|1)$$
$$= \sum_{k=2}^{g} P(k|1)c(k|1)$$

- The conditional expected cost occurs with prior proby p_1 , the proby of π_1 .
- Similarly, we can obtain $ECM(2), \dots, ECM(g)$.

Overall ECM

The overall ECM is

$$ECM = p_1 ECM(1) + \dots + p_g ECM(g)$$
$$= \sum_{i=1}^{g} p_i \left(\sum_{\substack{k=1\\k \neq i}}^{g} P(k|i) c(k|i) \right)$$

• Choose mutually exclusive and exhaustive classification regions R_1,\ldots,R_g such that ECM is a minimum.

Minimum ECM Classification Rule with Equal Misclassification Costs

- Suppose misclassification costs are equal (say all equal to 1).
- Allocate \mathbf{x}_0 to π_k if

$$p_k f_k(\mathbf{x}) > p_i f_i(\mathbf{x}) \text{ for all } i \neq k$$
 (1)

or
$$\ln p_k f_k(\mathbf{x}) > \ln p_i f_i(\mathbf{x})$$
 (2)

 Note that this misclassification rule is identical to the one that maximizes the "posterior" probability

$$\begin{split} P(\pi_k|\mathbf{x}) &= P(\mathbf{x} \, \text{comes from} \ \ \, \pi_k \, \text{given that} \, \mathbf{x} \, \text{was observed}) \\ &= \frac{p_k f_k(\mathbf{x})}{\sum_{i=1}^g p_i f_i(\mathbf{x})} = \frac{(\text{prior}) \times (\text{likelihood})}{\Sigma[(\text{prior}) \times (\text{likelihood})]} \end{split}$$

for $k = 1, \ldots, g$

Classification with Normal Populations

- Let $f_i(\mathbf{x}) \sim N_p(\boldsymbol{\mu}_i, \boldsymbol{\Sigma}_i)$, possibly unequal $\boldsymbol{\Sigma}_i$, $i = 1, \dots, g$
- Allocate \mathbf{x} to π_k if

$$\ln p_k f_k(\mathbf{x}) = \ln p_k - \left(\frac{p}{2}\right) \ln(2\pi) - \frac{1}{2} \ln |\mathbf{\Sigma}_k|$$
$$- \frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_k)' \mathbf{\Sigma}_k^{-1} (\mathbf{x} - \boldsymbol{\mu}_k)$$
$$= \max_i \ln p_i f_i(\mathbf{x})$$

• Let $d_i^Q(\mathbf{x})$ be the quadratic discriminant score given by

$$d_i^Q(\mathbf{x}) = -\frac{1}{2} \ln |\mathbf{\Sigma}_i| - \frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_i)' \mathbf{\Sigma}_i^{-1} (\mathbf{x} - \boldsymbol{\mu}_i) + \ln p_i,$$

where $i = 1, \ldots, g$

Quadratic Discriminant Analysis (QDA)

- Using Minimum Total Probability of Misclassification (TPM)
 Rule for Normal Populations
- (QDA) Allocate ${\bf x}$ to π_k if the quadratic discriminant score

$$d_k^Q(\mathbf{x}) = \max\left\{d_1^Q(\mathbf{x}), \dots, d_g^Q(\mathbf{x})\right\}$$

where

$$d_i^Q(\mathbf{x}) = -\frac{1}{2} \ln |\mathbf{\Sigma}_i| - \frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_i)' \mathbf{\Sigma}_i^{-1} (\mathbf{x} - \boldsymbol{\mu}_i) + \ln p_i,$$

Linear Discriminant Analysis (LDA)

• When $\Sigma_i = \Sigma$, for i = 1, ..., g, the quadratic discriminator score becomes

$$d_i^Q(\mathbf{x}) = -\frac{1}{2} \ln |\mathbf{\Sigma}| - \frac{1}{2} \mathbf{x}' \mathbf{\Sigma}^{-1} \mathbf{x} + \boldsymbol{\mu}_i' \mathbf{\Sigma}^{-1} \mathbf{x} - \frac{1}{2} \boldsymbol{\mu}_i' \mathbf{\Sigma}^{-1} \boldsymbol{\mu}_i + \ln p_i$$
$$= -\frac{1}{2} \ln |\mathbf{\Sigma}| - \frac{1}{2} \mathbf{x}' \mathbf{\Sigma}^{-1} \mathbf{x} + d_i(\mathbf{x})$$

where $d_i(\mathbf{x})$ is called the **linear disciminant score**,

$$d_i(\mathbf{x}) = \boldsymbol{\mu}_i' \boldsymbol{\Sigma}^{-1} \mathbf{x} - \frac{1}{2} \boldsymbol{\mu}_i' \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_i + \ln p_i$$

- The first two terms are the same for all $d_i^Q(\mathbf{x})$ so we ignore these terms.
- (LDA) Allocate ${\bf x}$ to π_k if the linear discriminant score

$$d_k(\mathbf{x}) = \max\{d_1(\mathbf{x}), \dots, d_q(\mathbf{x})\}\$$

Sample Version of linear discriminant scores

• We estimate $d_i(\mathbf{x})$ by

$$\hat{d}_i(\mathbf{x}) = \bar{\boldsymbol{x}}_i' \boldsymbol{S}_p^{-1} \mathbf{x} - \frac{1}{2} \bar{\boldsymbol{x}}_i' \boldsymbol{S}_p^{-1} \bar{\boldsymbol{x}}_i + \ln p_i, \text{ for } i = 1, \dots, g$$

• Equivalently, we can also consider the squared distances (estimate of the second term of $d_i^Q(\mathbf{x})$)

$$D_i^2(\mathbf{x}) = (\mathbf{x} - \bar{\mathbf{x}}_i)' \mathbf{S}_p^{-1} (\mathbf{x} - \bar{\mathbf{x}}_i)$$

Sample Linear Discriminant Analysis

- Estimated TPM Rule for Equal-Covariance Normal Populations
- Allocate x to π_k if the **estimated linear discriminant score**

$$\hat{d}_k(\mathbf{x}) = \max \left\{ \hat{d}_1(\mathbf{x}), \dots, \hat{d}_g(\mathbf{x}) \right\}$$

• Or, (in terms of squared distances) allocate ${f x}$ to π_k if

$$-rac{1}{2}D_k^2(\mathbf{x}) + \ln p_k$$
 is largest

• Or, (when priors are all equal) allocate \mathbf{x} to π_k if $D_k^2(\mathbf{x})$ is smallest (observation is assigned to closest population)

Example 11.11: Classifying a potential business-school graduate students

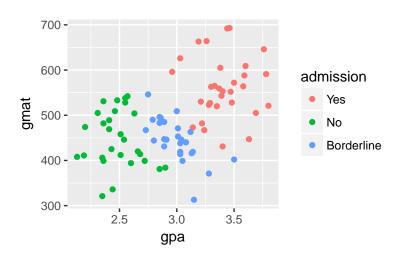
'data.frame': 85 obs. of 3 variables:

\$ gmat : int 596 473 482 527 505 693 626 663 447 58
\$ admission: Factor w/ 3 levels "Yes", "No", "Borderline"

head(bsgrad.df)

```
#
    gpa gmat admission
# 1 3.0 596
                   Yes
# 2 3.1 473
                   Yes
# 3 3.2
        482
                   Yes
 4 3.3
         527
                   Yes
# 5 3.7
         505
                   Yes
# 6 3.5
         693
                   Yes
```

```
bsgrad.df %>%
  ggplot(aes(x = gpa, y = gmat, colour = admission)) +
  geom_point()
```



Linear Discriminant Analysis, Grad School

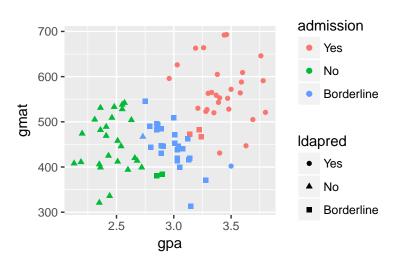
```
# Yes No Borderline
# Yes 28 0 1
# No 0 26 1
# Borderline 3 2 24
```

```
# accuracy rate
sum(diag(cm.bs.lda))/nrow(bsgrad.df)
```

```
# [1] 0.92
```

LDA, Business School

LDA result, Business School



Quadratic Discriminant Analysis, Business School

```
# Yes No Borderline
# Yes 28 0 1
# No 0 26 1
# Borderline 3 2 24
```

```
# accuracy rate
sum(diag(cm.bs.qda))/nrow(bsgrad.df)
```

```
# [1] 0.92
```

LDA and QDA, Business School

Similar Result for LDA and QDA

```
#
# Yes No Borderline
# Yes 28 0 1
# No 0 26 1
# Borderline 3 2 24
cm.bs.qda
```

QDA, Business School

QDA result, Grad School

