Model-based Statistical Learning



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The good of the first step is to find the
best mixture parameter estimates for the data
ad hard:

Ceann

X

With ETT

A: ..., The

Line of this model without directly oppositive the function.

Starting with the GMR model, we can unite the

Pos. hihelihood of the model:
$$\mathcal{L}(n; \Theta) = \log \left(\frac{m}{11} p(n; \Theta) \right) \\
= \log \left(\frac{m}{11} \frac{k}{2} \pi_{A} N(n; \mu_{A}, \Sigma_{A}) \right)$$

$$= \frac{\log\left(\prod_{i=1}^{n}\sum_{k=1}^{K} \prod_{k} N(n_{i}; \mu_{k}, \Sigma_{k})\right)}{\sum_{i=1}^{n} \log\left(\sum_{k=1}^{K} \prod_{k} N(n_{i}; \mu_{k}, \Sigma_{k})\right)}$$

 $= \int_{a}^{\infty} \log \left(\sum_{k=1}^{N_{K}} \pi_{k} N(n; \mu_{k}, \Sigma_{k}) \right)$ We see here that it is totally possible to oudnote $\mathcal{L}(n; \delta)$ for specific values of n and n but the direct optimization is really difficult due to the $\log (\Xi_{k})$

The idea of the ETT algorithm is to introduce on extra

and non-observed (latent) variable 4, encoding the group memberships. Estimating both Dand 4 fran X is finally easier than jut estimating & fan X.

 $Y_i \in \{91\}^{\prime\prime}$ => $Y_i = (0,0,1,0)$ => x_i belows to the $3^{\prime\prime}$ cluster

 $\begin{cases} Y \sim \mathcal{O}(1; T) \\ \times | Y \sim \mathcal{N}(\mu_A, Z_A) \end{cases}$ indegrate over γ $\Rightarrow p(x) = \sum_{k=1}^{K} \pi_k \mathcal{N}(n; \mu_k, \Sigma_k)$

This allows as to write another log-likelihood, called the complete data log-likelihood, of the comple (x,4):

[] log p (x, 4i); 0) L(x,4,0) = [| log p(x,0) + log p(x,0)] = 2(n;0) + = log p(4; |xi;0) $\Rightarrow \mathcal{L}(\alpha; 0) = \mathcal{L}_{c}(x, \gamma; 0) - \sum_{i=1}^{n} \log p(Y_{i} | x_{i}; 0)$ => Le < L

Ruch: we see here that Le is a lower bound of L and optimizing Le over o will antanahilly lead in the the optimization of L.

Thanks to this remath, Dempster, Card and Rulm pagosed in 1977 the ETT algaillus: - E step: the E step aims at calculating the expected complete log-likelihood: Q(0,0) = E[L(x,4,0) | X; 0)

- M step: the T step aims at maximizing
this farction Q(0;0) over O to provide a
new value for O

Theorem: the sequence of estimates (O) over the iterations of the En algorithm is converging toward a local maximum of the log-liberhard.

The EM algorithm for GMM

In machice:

- e) to avoid being happed in a local maximum, we wouldy do several (10) different random initializations and we keep afterward the OET with the highest likelihood.
- 2) to stop the algorithm, we just monitor the evalution of the Coz. Whethood and we stop when a plateau is detected

The EM algorithm for GMM

Starting with the Estep, we need to focus on
$$Q(0,0) = E[L_c(X,Y;\Theta) | \Theta,X]$$
and $L_c(X,Y;\Theta) = \sum_{i=1}^{m} \log_i p(x_i,Y_i;\Theta)$

$$= \sum_{i=1}^{m} \log_i \sum_{k=1}^{m} y_{ik} \pi_k N(x_i,y_k,\Sigma_k)$$

$$= \sum_{i=1}^{m} \sum_{k=1}^{m} y_{ik} \log_i \pi_k N(x_i,y_k,\Sigma_k)$$

and
$$E[y:k|\vartheta;X] = P(y:k=1|X,\vartheta)$$

Bayes $P(y:k=1|\delta)P(X:|Y:k=1;\vartheta)$
 $Z= \frac{P(X)}{1/2} \times N(X:|Y:k=1;\vartheta)$
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 $E[L_c(x,y|\theta)|\theta^*,X] = \sum_{i=1}^{n} \sum_{k=1}^{n} \frac{E[y;k|\theta^*,X]}{E[y;k|\theta^*,X]} \log \left(\pi_i W(\alpha_i) \mu_i, \Sigma_i \right)$

and Hefore:

The EM algorithm for GMM

In the M step, we just have to ophimize in O

He E[Le(x,4,0) | 0, x):

Max $E[\mathcal{L}_c(X,Y;\theta)|\theta;X] = \sum_{i=1}^{m} \sum_{k=1}^{k} Q(\theta\theta) = \sum_{i=1}^{m} \sum_{k=1}^{k} Q(\theta\theta) = \sum_{i=1}^{m} \sum_{k=1}^{k} Q(\theta) = \sum_{i=1}^{m} \sum_{k=1}^{k} Q(\theta) = \sum_{i=1}^{m} \sum_{k=1}^{m} Q(\theta) = \sum_{i=1}^{m} Q(\theta) = Q(\theta) = \sum_{i=1}^{m} Q(\theta) = Q(\theta) = Q(\theta) = Q(\theta)$ til Pos Th N(xi; M, Zh)

. Finding the update for 1/2:

 $\frac{\partial}{\partial \pi_{k}} Q(\theta; \theta) = \frac{\partial}{\partial \pi_{k}} \left[\sum_{i} \int_{\mathcal{L}} dx_{i} \mathcal{L}_{q_{i}} \mathcal{L}_$