

1) We know that

$$L = |y - \hat{y}|^2$$

$$E(L) = \int \int |y - \hat{y}|^2 p(x, y) dx dy.$$

$$= \int \int |y - \hat{y}(x)|^2 p(x, y) dx dy.$$

For the Averaged loss to be minimum,  
differentiate  $E(L)$  w.r.t.  $\hat{y}(x)$ .

$$\frac{\partial E(L)}{\partial \hat{y}(x)} = 2 \int (y - \hat{y}(x)) p(x, y) dx dy$$

$$\Rightarrow \int y p(x, y) dy - \int \hat{y}(x) p(x, y) dy = 0$$

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$$\Rightarrow \hat{y}(x)p(x) = \int y p(x,y) dy$$

$$\Rightarrow \hat{y}(x) \cancel{p(x)} = \int y \frac{p(x,y)}{\cancel{p(x)}} dy$$

$$\Rightarrow \boxed{\hat{y}(x) = E[\cancel{y}/x]}$$

$$2) E[(y - \hat{y})^2] = E[(y - y^* + y^* - \hat{y})^2]$$

$$= E[(y - y^*)^2] + E[(y^* - \hat{y})^2] +$$

$$+ 2E[(y - y^*)(y^* - \hat{y})]$$

Simplifying the third term,

$$E[(y - y^*)(y^* - \hat{y})] = \iint (\hat{y}(x) - y^*(x)) \cdot (y^*(x) - y) p(x,y) dx dy$$

$$= \int (\hat{y}(x) - y^*(x)) \int (y^*(x) - y) p(x,y) dy dx$$

But here,  $\int (y^*(x) - y) p(x,y) dy dx$

$$= \int (y^*(x) - y) p(y/x) p(x) dx.$$



$$= \int y^*(x) p(x) dx - \underbrace{\int \left( \int y p(y/x) \right) p(x) dy dx}_{\int E(y/x) p(x) dx}$$

$$= 0$$

$$\therefore E[(y - \hat{y}(x))^2] = E[(y - y^*(x))^2] + E[(\hat{y}(x) - y^*(x))^2]$$

Expanding the 2nd term in a dataset  $D$ ,

$$E_D[(\hat{y}_D(x) - y^*(x))^2] = E_D[(y_D(x) - E_D(\hat{y}(x)))^2] + E_D[(y^*(x) - E_D(\hat{y}(x)))^2]$$

$$= E_D[(y_D(x) - E_D(\hat{y}(x)) + E_D(\hat{y}(x)) - y^*(x))^2]$$

$$= E_D[(y_D(x) - E_D(\hat{y}(\vec{x})))^2] + \text{Variance}$$

$$+ E_D[(y^*(x) - E_D(\hat{y}(\vec{x})))^2] \rightarrow \text{Bias}^2$$

~~Qy~~

$$\therefore E[(y - \hat{y}(x))^2] = \text{Variance} + (\text{bias})^2 + \text{noise}$$

### 3) Least Squares Solution for k-class discriminant classifier.

Let  $X$  be  $d$ -dimensional and  $y$  be labels which belong to one of the classes.

$$\tilde{X} = (I \ X^T)^T, \quad \tilde{X}^T = \begin{pmatrix} 1 & 1 & \dots & 1 \\ x_{11} & x_{12} & \dots & x_{1N} \\ \vdots & \vdots & \ddots & \vdots \\ x_{d1} & x_{d2} & \dots & x_{dN} \end{pmatrix}$$

Let  $\hat{y}(x) = \tilde{X} \tilde{W}$  when  $\tilde{W}^T$  represents the set of coefficients of dimensions  $k$  w.r.t to each class.

$$\tilde{W} = \begin{pmatrix} w_{10} & \dots & w_{k0} \\ w_{11} & \dots & w_{k1} \\ \vdots & \ddots & \vdots \\ w_{1d} & \dots & w_{kd} \end{pmatrix}$$

Sum of squares  $E = (\tilde{X} \tilde{W} - y)^T (\tilde{X} \tilde{W} - y)$   
Minimising it,

$$\frac{\partial E}{\partial \tilde{W}} = 0 \Rightarrow \frac{\partial}{\partial \tilde{W}} (\tilde{W}^T \tilde{X}^T - y^T) (\tilde{X} \tilde{W} - y) = 0$$

$$\Rightarrow \frac{\partial}{\partial \tilde{W}} (\tilde{W}^T \tilde{X}^T \tilde{X} \tilde{W} - y^T \tilde{X} \tilde{W} - \tilde{W}^T \tilde{X}^T y + y^T y) = 0$$



$$\Rightarrow \tilde{X}^T \tilde{X} \tilde{W} + \tilde{X}^T \tilde{X} \tilde{W} - \tilde{X}^T y - \tilde{X}^T y = 0.$$

$$\Rightarrow 2 \tilde{X}^T \tilde{X} \tilde{W} - 2 \tilde{X}^T y = 0.$$

$$\Rightarrow \boxed{\tilde{W}^* = (\tilde{X}^T \tilde{X})^{-1} \tilde{X}^T y}$$

Using the  $\tilde{W}^*$ , calculating  $\hat{y} = \tilde{X} \tilde{W}^*$  and finding the maximum  $\hat{y}_k$ , that is the ~~maximum~~ class of the respective input.

4)  $X$  is a  $d$ -dimensional input,  $Y$  is 2-class output label.

$N$ -samples of which  $n_1$  belong to class 1 and  $n_2$  to class 2.

$V$  be a unit vector along the imaginary line boundary. Distance of  $x$  from the line will be  $-V^T x$ .

Therefore mean of projection onto class '1' will be

$$\mu_1 = \frac{1}{n_1} \sum_{x_i \in C_1} V^T x_i = V^T \left( \frac{1}{n_1} \sum_{x_i \in C_1} x_i \right)$$

$$\boxed{\mu_1 = V^T \mu_1'}$$

Similarly,  $\mu_2 = v^T \mu'_2$

Let the variance be the measure of scatter.

$$S = \sum_{i=1}^n (z_i - \mu_2)^2 \quad \text{Variance}$$

$$\begin{aligned} S_1 &= \sum_{y_i \in C_1} (y_i - \mu_1)^2 \\ &= \sum_{x_i \in C_1} (v^T x_i - v^T \mu'_1)^2 \end{aligned}$$

Similarly,

$$S_2 = \sum_{y_i \in C_2} (y_i - \mu_2)^2 = \sum_{x_i \in C_2} (v^T x_i - v^T \mu'_2)^2$$

In Fisher's discriminant process, we will maximise the separation between the two classes and minimize the scattering

$$J(v) = \frac{(\mu_1 - \mu_2)^2}{S_1 + S_2}$$



$$\text{Let } S_1 = \sum_{x_i \in C_1} (x_i - \mu_1') (x_i - \mu_1')^T$$

$$S_2 = \sum_{x_i \in C_2} (x_i - \mu_2') (x_i - \mu_2')^T$$

$$\therefore S_1 = \sum_{y \in C_1} (V^T x_i - V^T \mu_1')^2$$

$$= \sum_{x_i \in C_1} V^T (x_i - \mu_1') (x_i - \mu_1')^T V$$

$$= V^T S_1 V$$

$$\text{Similarly, } S_2 = V^T S_2 V$$

$$\text{① + ② } S_1 + S_2 = V^T (S_1 + S_2) V$$

$$= V^T S_w V$$

$$\text{Let } d = (\mu_1' - \mu_2') (\mu_1' - \mu_2')^T$$

$$\text{Consider } (\mu_1' - \mu_2')^2 = V^T (\mu_1' - \mu_2') (\mu_1' - \mu_2')^T V$$

$$= V^T d V$$

$$\therefore J(V) = \frac{V^T d V}{V^T S_w V}$$

$$\text{For maximising, } \nabla J(V) = 0$$

$$\Rightarrow \frac{(V^T S_w V)(2dV) - (2S_w V)V^T dV}{(V^T S_w V)^2} = 0$$

$$\Rightarrow v^T S_w v dv = v^T \otimes dv S_w v$$

$$\Rightarrow dv = v^T dv$$

$$\Rightarrow dv = \boxed{(v^T S_w v)^{-1} v^T dv} S_w v$$

$\lambda$

$$\Rightarrow dv = \lambda S_w v$$

$$\Rightarrow \lambda v = (S_w)^{-1} dv$$

Consider  $dxv = (\mu'_1 - \mu'_2)(\mu'_1 - \mu'_2)^T v$

$$= (\mu'_1 - \mu'_2) (\underbrace{v^T \mu'_1}_{\text{const}} - \underbrace{v^T \mu'_2}_{\text{const}})^T$$

$$dv = \text{const} \times (\mu'_1 - \mu'_2)$$

$$\therefore \lambda v = (S_w)^{-1} (\mu'_1 - \mu'_2) \times \text{const}$$

$$\Rightarrow v_{op} \propto S_w^{-1} (\mu'_1 - \mu'_2)$$

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5) Given loss function

$$L(\hat{y}, y) = \begin{cases} 0 & y = \hat{y} \\ 1 & \text{else} \end{cases}$$

$$\begin{aligned} E_{x,y}[L(\hat{y}, y)] &= E_x[E_{y/x}[L(\hat{y}, y)]] \\ &= E_x\left[\sum_{k \in C_k} L(y, \hat{y}) \cdot p(y = \hat{y}/x)\right] \end{aligned}$$

Let  $y = y^*$  be the optimal label we know  
that  $\sum_{k=1}^K p(y = y^*/x) = 1$ .

$$\Rightarrow E_{x,y}[L(\hat{y}, y)] = E_x[1 - p(y = y^*/x)]$$

Find  $y^*$  such that  $\arg\min_y E_x[1 - p(y = y^*/x)]$

$\Rightarrow p(y = y^*/x)$  should be maximum.

If  $p(y = y^*/x)$  is maximum for a class then it will be predicted class.

$y^* = C_k$  (class  $k$ ) if

$$p(C_k/x) > p(C_j/x) \quad \forall k \neq j$$