



**Applied Statistics**  
**Homework-5**

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PROBLEM 1:-

Applied Statistics

Hw-5

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Q1) I don't think this conclusion is appropriate. We can say that older mothers & higher IQ are associated. But I think we should also consider other factors that might help depend on the higher IQ. We can also take into account factors like education & quality, medical facility rather than just considering age.



PROBLEM 2:-

Q.2) 12408 -

$$n_1 = 73$$

$$\bar{x}_1 = 6.22$$

$$s_1 = 1.62$$

$$n_2 = 105$$

$$\bar{x}_2 = 5.81$$

$$s_2 = 1.43$$

$$n_3 = 240$$

$$\bar{x}_3 = 5.77$$

$$s_3 = 1.24$$

$$n_4 = 1080$$

$$\bar{x}_4 = 5.47$$

$$s_4 = 1.31$$

a)

$$S_w^2 = \frac{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2 + (n_3 - 1)s_3^2 + (n_4 - 1)s_4^2}{n_1 + n_2 + n_3 + n_4 - 4}$$

$$= \frac{(73 - 1)1.62^2 + (105 - 1)1.43^2 + (240 - 1)1.24^2 + (1080 - 1)1.31^2}{73 + 105 + 240 + 1080 - 4}$$

$$S_w^2 = \underline{\underline{1.75}}$$

$$\text{Now, } \bar{X} = \frac{n_1\bar{x}_1 + n_2\bar{x}_2 + n_3\bar{x}_3 + n_4\bar{x}_4}{n_1 + n_2 + n_3 + n_4}$$

$$= \frac{73 \times 6.22 + 105 \times 5.81 + 240 \times 5.77 + 1080 \times 5.47}{73 + 105 + 240 + 1080}$$

$$= \underline{\underline{5.58}}$$



$$S_B^2 = \frac{n_1(\bar{x}_1 - \bar{x})^2 + n_2(\bar{x}_2 - \bar{x})^2 + n_3(\bar{x}_3 - \bar{x})^2 + n_4(\bar{x}_4 - \bar{x})^2}{k-1}$$

$$= \frac{73(6.22 - 5.58)^2 + 105(5.81 - 5.58)^2 + 240(5.77 - 5.58)^2 + 1080(5.67 - 5.58)^2}{3}$$

$$= \underline{\underline{19.06}}$$

$$F = \frac{S_B^2}{S^2_W} = \frac{19.06}{1.75}$$

$$= \underline{\underline{10.89}}$$

$$\text{Degree of freedom} = k-1 = 4-1$$

$$= \underline{\underline{3}}$$

$$\text{and } n-k = 1498-4$$

$$= \underline{\underline{1494}}$$



b) For  $F_{3,1494}$

b) From the table, the critical value,

$$F_{3,1494} = 2.6108$$

at significance  
level = 0.05

$$\Rightarrow P\text{-value} < 0.001$$

Hence we reject  $H_0$ , so we conclude that  
the mean LDL cholesterol is different  
from the four population.

c) The assumptions are:-

The data within one group represents  
a random sample from a ~~pop~~ population.

The population are independent within  
the group, the observations are  
normally distributed with mean  $\mu$

Variance  $\sigma^2$  is same for all groups.

#### PART D:- R CODE AND OUTPUT:-

```
1 library(readxl)
2 library("psych")
3 #contr1 <- c(1,-1/3,-1/3, -1/3)
4 contr1 <- c(1, 1,1,-3)
5 ldl.data <- read_excel("hw5.xlsx")
6 k<-4
7 (contr1.est <- sum(contr1*ldl.data[, 'mean']))
8 (MSE <- sum((ldl.data[, 'n']-1)*ldl.data[, 'sd']^2)/sum(ldl.data[, 'n']-1))
9 contr1.se <- sqrt(sum(contr1^2/ldl.data[, 'n'])*MSE)
10 contr1.t <- contr1.est/contr1.se
11 contr1.p <- 2*pt(-abs(contr1.t),df=sum(ldl.data[, 'n'])-k)
12 c(contr1.est,contr1.se, contr1.t,contr1.p)
13 |
14
15
```

13:1 (Top Level) ↕

Console	Terminal ×	Jobs ×
~/LAB1.1/ ↗		
<pre>&gt; library("readxl") &gt; library("psych") &gt; #contr1 &lt;- c(1,-1/3,-1/3, -1/3) &gt; contr1 &lt;- c(1, 1,1,-3) &gt; ldl.data &lt;- read_excel("hw5.xlsx") &gt; k&lt;-4 &gt; (contr1.est &lt;- sum(contr1*ldl.data[, 'mean'])) [1] 1.39 &gt; (MSE &lt;- sum((ldl.data[, 'n']-1)*ldl.data[, 'sd']^2)/sum(ldl.data[, 'n']-1)) [1] 1.754207 &gt; contr1.se &lt;- sqrt(sum(contr1^2/ldl.data[, 'n'])*MSE) &gt; contr1.t &lt;- contr1.est/contr1.se &gt; contr1.p &lt;- 2*pt(-abs(contr1.t),df=sum(ldl.data[, 'n'])-k) &gt; c(contr1.est,contr1.se, contr1.t,contr1.p) [1] 1.390000e+00 2.503289e-01 5.552696e+00 3.324252e-08 &gt;  </pre>		



d) from R,

est	se	t-test	p-value
1.39	$2.50 \times 10^{-1}$	S.S	$3.324 \times 10^{-8}$

Reject  $H_0$  at  $\alpha = 0.05$ , since  $3.324 \times 10^{-8} < 0.05$ .

With Bonferroni,

$$= \frac{0.05}{\left(\frac{4}{2}\right)} = 0.0083$$

Hence we still reject.

For Scheffe,

$t_{1494, 0.25} = -1.96$  &  $S.S > -1.96$   
we reject.

Hence No correction at all.

## PART E:- CODE AND OUTPUT:-

```
1 library("readxl")
2 library("psych")
3 contr1 <- c(1,-1/3,-1/3, -1/3)
4 #contr1 <- c(1, 1,1,-3)
5 ld1.data <- read_excel("hw5.xlsx")
6 k<-4
7 (contr1.est <- sum(contr1*ld1.data[, 'mean']))
8 (MSE <- sum((ld1.data[, 'n']-1)*ld1.data[, 'sd']^2)/sum(ld1.data[, 'n']-1))
9 contr1.se <- sqrt(sum(contr1^2/ld1.data[, 'n'])*MSE)
10 contr1.t <- contr1.est/contr1.se
11 contr1.p <- 2*pt(-abs(contr1.t),df=sum(ld1.data[, 'n'])-k)
12 c(contr1.est,contr1.se, contr1.t,contr1.p)
13 |
14
15
```

13:1 (Top Level) ↕

onsole

Terminal ×

Jobs ×

./LAB1.1/ ↗

```
library("readxl")
library("psych")
contr1 <- c(1,-1/3,-1/3, -1/3)
#contr1 <- c(1, 1,1,-3)
ld1.data <- read_excel("hw5.xlsx")
k<-4
(contr1.est <- sum(contr1*ld1.data[, 'mean']))
L] 0.5366667
(MSE <- sum((ld1.data[, 'n']-1)*ld1.data[, 'sd']^2)/sum(ld1.data[, 'n']-1))
L] 1.754207
contr1.se <- sqrt(sum(contr1^2/ld1.data[, 'n'])*MSE)
contr1.t <- contr1.est/contr1.se
contr1.p <- 2*pt(-abs(contr1.t),df=sum(ld1.data[, 'n'])-k)
c(contr1.est,contr1.se, contr1.t,contr1.p)
L] 0.536666667 0.163948582 3.273384000 0.001086971
|
```



~~Q)~~

e) From R,

est	sd	t-test	p-value
0.536	0.16	3.27	0.0010

Because  $0.0010 < 0.05$ , we reject  $H_0$

For Bonferroni, as above  $0.0003 > p$ , so we again reject.

For Scheffe,  $t_{149, 0.25} = -1.61$ , so we still reject as  $-1.61 < 3.27$ .

Hence we reject the null hypothesis

### PROBLEM 3:-

#### CODE:-

```
1 attach(airquality)
2 airquality[is.na(airquality)] = 0
3
4
5 pairwise.t.test(Ozone, Month, p.adjust.method = "bonf")
6
7 pairwise.t.test(Ozone, Month, p.adjust.method = "fdr")
8
9 airquality$Month<-as.factor(airquality$Month)
10 p<-aov(Ozone~Month,data=airquality)
11 TukeyHSD(p,conf.level = 0.95)
12
13 detach()
14
```

#### OUTPUT:-

##### Bonferroni:-

```
> pairwise.t.test(Ozone, Month, p.adjust.method = "bonf")

      Pairwise comparisons using t tests with pooled SD

data:  Ozone and Month
   5      6      7      8
6 1.0000 -      -      -
7 0.0015 4.4e-06 -      -
8 0.0011 2.9e-06 1.0000 -
9 1.0000 0.0625 0.1399 0.1088

P value adjustment method: bonferroni
```

##### FDR:-

```
> pairwise.t.test(Ozone, Month, p.adjust.method = "fdr")

      Pairwise comparisons using t tests with pooled SD

data:  Ozone and Month
   5      6      7      8
6 0.19069 -      -      -
7 0.00037 2.2e-06 -      -
8 0.00035 2.2e-06 0.92620 -
9 0.19069 0.01251 0.01998 0.01813

P value adjustment method: fdr
> airquality$Month<-as.factor(airquality$Month)
```



#### TUKEY:-

```
Tukey multiple comparisons of means
95% family-wise confidence level

Fit: aov(formula = Ozone ~ Month, data = airquality)

$Month
      diff      lwr      upr    p adj
6-5 -10.9731183 -32.27095900 10.324722 0.6139469
7-5  29.7741935   8.65164668 50.896740 0.0013894
8-5  30.4838710   9.36132410 51.606418 0.0009868
9-5  10.5935484 -10.70429233 31.891389 0.6454439
7-6  40.7473118  19.44947111 62.045153 0.0000044
8-6  41.4569892  20.15914853 62.754830 0.0000029
9-6  21.5666667   0.09496314 43.038370 0.0484120
8-7   0.7096774 -20.41286945 21.832224 0.9999830
9-7 -19.1806452 -40.47848588  2.117196 0.0990957
9-8 -19.8903226 -41.18816330  1.407518 0.0795001
```

No, the results are not the same. I think Tukey is the more appropriate conclusion because we are using a pairwise comparison.

#### PROBLEM 4:-

##### CODE:-

```
1 #import data set
2 data <- read.table(file="lowbwt.txt", header = TRUE)
3
4 data$sex<-as.factor(data$sex)
5 data$tox<-as.factor(data$tox)
6
7 ##Anova with bocking
8 x <- aov(sbp~sex+tox, data=data)
9 summary(x)
10
11 ##Anova without bocking
12 x<-aov(sbp~sex,data=data)
13 summary((x))
14
```

##### OUTPUT:-

```
> data$tox<-as.factor(data$tox)
> ##Anova with bocking
> x <- aov(sbp~sex+tox, data=data)
> summary(x)
      Df Sum Sq Mean Sq F value Pr(>F)
sex      1      48   48.25    0.367   0.546
tox      1      67   66.76    0.508   0.478
Residuals 97  12758  131.53
> ##Anova without bocking
```

```

> ##Anova without bocking
> x<-aov(sbp~sex,data=data)
> summary(x)

```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
sex	1	48	48.25	0.369	0.545
Residuals	98	12825	130.87		

```

> |

```

- We fail to reject Null hypothesis at .05 significant level. There is no significant difference for the mean systolic blood pressure for low birth weight boys and girls.
- p-value =0.5461 for gender effects on blood pressure with blocking  
p-value =0.5451 for gender effects on blood pressure without blocking.
- The ANOVA F test without blocking equivalent to the independent two sample test.  
The ANOVA F-test with blocking is same as paired two sample test.

#### PROBLEM 5:-

CODE for a,b,c:-

```

1 rt <- read.csv("response_times.csv", header = TRUE, sep = ",")
2 #a
3 anova_result<-aov(time~size, data=rt)
4 summary(anova_result)
5
6 #b
7 TukeyHSD(anova_result)
8
9 #c
10
11 pairwise.t.test(rt$time,rt$size, p.adjust.method = "none")
12

```

#### PART A):-

```

> rt <- read.csv("response_times.csv", header = TRUE, sep = ",")
> #a
> anova_result<-aov(time~size, data=rt)
> summary(anova_result)

```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
size	3	1.929	0.6431	4.234	0.00882 **
Residuals	60	9.113	0.1519		

```

---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
> #b

```

Based on the ANOVA test, we fail to reject the null hypothesis. There is no significant difference among response times.



## PART B) :-

### HSD :-

```
> #D
> TukeyHSD(anova_result)
  Tukey multiple comparisons of means
    95% family-wise confidence level

Fit: aov(formula = time ~ size, data = rt)

$size
      diff      lwr      upr    p adj
Medium-Large  0.1295625 -0.23454756  0.49367256 0.7833229
Small-Large   0.3094375 -0.05467256  0.67354756 0.1227683
XLarge-Large  -0.1641250 -0.52823506  0.19998506 0.6347743
Small-Medium  0.1798750 -0.18423506  0.54398506 0.5631358
XLarge-Medium -0.2936875 -0.65779756  0.07042256 0.1549644
XLarge-Small  -0.4735625 -0.83767256 -0.10945244 0.0057723
```

### LSD :-

```
> pairwise.t.test(rt$time,rt$size, p.adjust.method = "none")
      Pairwise comparisons using t tests with pooled SD

data:  rt$time and rt$size

      Large  Medium Small
Medium 0.3508 -      -
Small  0.0284 0.1967 -
XLarge 0.2383 0.0372 0.0011

P value adjustment method: none
> |
```

We came to same conclusion in both tests. Fail to reject null hypothesis.