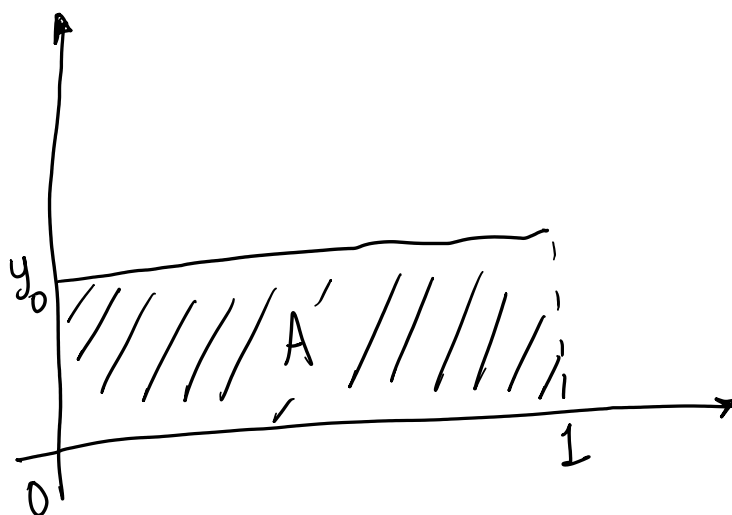


①

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a)

$f_0 \rightarrow$  uniform over  $[0, 1]$



$$A = 1 = y_0 \times 1$$

$$\Rightarrow y_0 = 1$$

$$\therefore f_0(\theta) = \begin{cases} 1, & 0 \leq \theta \leq 1 \\ 0, & \text{else} \end{cases}$$

b) Data :  $\left\{ \text{Continuous r.v. } X \text{ is uniform over } [0, \theta] \right\}$

$$0 \leq \theta \leq 1$$

$X \sim \text{uniform on } [0, \theta]$

$\therefore$  PDF of  $x$  is

$$f(x|\theta) = \begin{cases} \frac{1}{\theta} & , \quad 0 \leq x \leq \theta \\ 0 & , \quad \text{else} \end{cases}$$

$\therefore$  Likelihood,

$$f(D|\theta) = f(x|\theta)$$

$$= \begin{cases} \frac{1}{\theta} & , \quad 0 \leq x \leq \theta \\ 0 & , \quad \text{else} \end{cases}$$

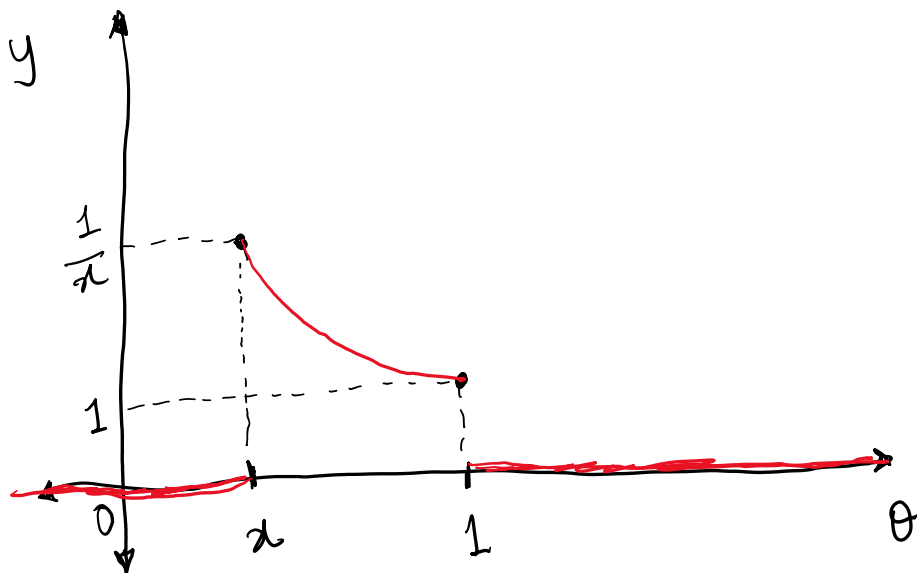
c) Posterior PDF

$$f_1(\theta|D) = \frac{f(x_1|\theta) f_0(\theta)}{Z}$$

$$= \begin{cases} \frac{\frac{1}{\theta} x^1}{Z}, & 0 \leq x_1 \leq \theta \\ 0, & \text{else} \end{cases}$$

$$= \begin{cases} \frac{1}{Z \theta}, & 0 \leq x_1 \leq \theta \\ 0, & \text{else} \end{cases}$$

$$y = f(x_1 | \theta) f_0(\theta) = \frac{1}{\theta}$$



d)

$$Z = \int_0^1 f(x_1 | \theta) f_0(\theta) d\theta$$

$$= \int_0^x f(x_1 | \theta) \times \underset{\substack{\uparrow \\ \theta \geq x}}{0} d\theta + \int_x^1 f(x_1 | \theta) \times 1 d\theta$$

$$= \int_x^1 \frac{1}{\theta} \times 1 \times d\theta$$

$$= -\frac{1}{\ln(x)}$$

$$\therefore f_1 = \begin{cases} \frac{1}{Z \theta} = -\frac{1}{\theta \ln(x)} & x \leq \theta \leq 1 \\ 0 & , \text{ else} \end{cases}$$

(2)

$X \sim \text{uniform on } [0, \theta]$

$$f_0(\theta) = \begin{cases} 1, & \text{if } \begin{matrix} 0 \leq x_1 \leq \theta \\ 0 \leq x_2 \leq \theta \end{matrix} \\ 0, & \text{else } 0 \leq \theta \leq 1 \end{cases}$$

Likelihood:

$$f(D|\theta) = f(x_1|\theta) \times f(x_2|\theta)$$

← because independent measurements

$$= \begin{cases} \frac{1}{\theta^2}, & \begin{matrix} 0 \leq x_1 \leq \theta \\ 0 \leq x_2 \leq \theta \end{matrix} \\ 0, & \text{else} \end{cases}$$

Posterior:

$$f_1(\theta|D) = \begin{cases} \frac{1}{Z \theta^2}, & \begin{matrix} 0 \leq x_1 \leq \theta \\ 0 \leq x_2 \leq \theta \end{matrix} \\ 0, & \text{else} \end{cases}$$

let,  $x = \max(x_1, x_2)$

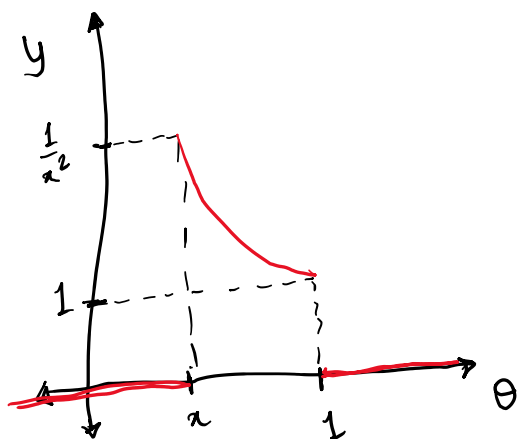
$$Z = \int_0^1 f(x_1|\theta) \times f(x_2|\theta) \times f_0(\theta) \times d\theta$$

$$Z = \int_x^1 \frac{1}{\theta^2} \times 1 \times d\theta$$

$\nwarrow$   
 $x = \max(x_1, x_2)$

$$= \frac{1}{x} - 1$$

$$y = f(x_1|\theta) \times f(x_2|\theta) \times f_0(\theta) = \begin{cases} \frac{1}{\theta^2} & 0 \leq x_1 \leq \theta \\ 0 & 0 \leq x_2 \leq \theta \end{cases}$$



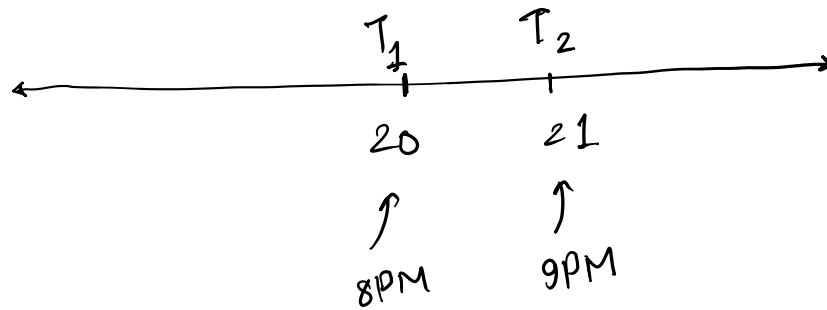
$$f_1(\theta|x) = \begin{cases} \frac{y}{Z}, & x \leq \theta \leq 1 \\ 0, & \text{else} \end{cases}$$

$$= \begin{cases} \frac{1}{\theta^2} \times \frac{x}{1-x}, & x \leq \theta \leq 1 \\ 0, & \text{else} \end{cases}$$

3

$$\lambda = 2 \text{ hr}^{-1}$$

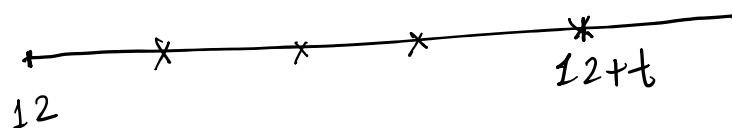
(a)



$$T_2 - T_1 = 1 \text{ hr}$$

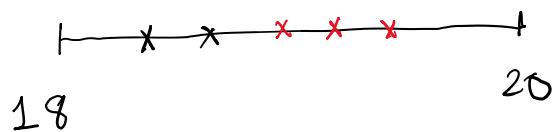
$$\begin{aligned} \therefore P(N(1) = 0) &= \frac{(\lambda \times 1)^0 e^{-\lambda \times 1}}{0!} \\ &= e^{-\lambda} \end{aligned}$$

(b)



$$\begin{aligned} E(T_4) &= 12 + \frac{4}{\lambda} = 12 + \frac{4}{2} = 14 \text{ hr} \\ &= 2 \text{ PM} \end{aligned}$$

(c)



$$\begin{aligned} P(N(20-18) \geq 2) &= 1 - P(N(2) < 2) \\ &= 1 - P(N(2) = 0) - P(N(2) = 1) \\ &= 1 - \frac{e^{-2 \times 2} \times (2 \times 2)^0}{0!} - \frac{e^{-2 \times 2} \times (2 \times 2)^1}{1!} \\ &= 1 - 5e^{-4} \end{aligned}$$

④

$$P(Y=1) = 0.4$$

$$P(Y=2) = 0.3$$

$$P(Y=3) = 0.2$$

$$P(Y=4) = 0.1$$

a)

$$P(Y=4) = 0.1$$



$$\begin{aligned}
 \text{b)} \quad \lambda' &= \text{rate of full cars} = \lambda P(Y=4) \\
 &= 10 \times 0.1 \text{ per minute} \\
 &= 1 \text{ per minute}
 \end{aligned}$$

$$\begin{aligned}
 P(N(1) = 2) &= \frac{e^{-1 \times 1} \times (1)^2}{2!} \\
 &= \frac{1}{2e}
 \end{aligned}$$

$$\begin{aligned}
 \text{c)} \quad E\left(\sum_{i=1}^{N(1)} Y_i\right) &= E(N(1)) \times E(Y) \\
 &= (10 \times 1) \times \left(1 \times 0.4 + 2 \times 0.3 + 3 \times 0.2 + 4 \times 0.1\right) \\
 &= 20
 \end{aligned}$$

d)

$$P\left(N\left(\frac{1}{6}\right) \geq 2\right) = 1 - P\left(N\left(\frac{1}{6}\right) < 2\right)$$

$$= 1 - P\left(N\left(\frac{1}{6}\right) = 0\right) - P\left(N\left(\frac{1}{6}\right) = 1\right)$$

$$= 1 - \frac{e^{-10 \times \frac{1}{6}} \times \left(10 \times \frac{1}{6}\right)^0}{0!} - \frac{e^{-10 \times \frac{1}{6}} \times \left(10 \times \frac{1}{6}\right)^1}{1!}$$

$$= 1 - \frac{8}{3}e^{-\frac{5}{3}}$$


---

⑤

$$p_1 = \frac{\mu}{\mu + \lambda} = \frac{3}{5}$$

$$p_2 = \frac{\lambda}{\mu + \lambda} = \frac{2}{5}$$

This can be mapped to coin-tossing problem

$$P(\text{at least 2 men before 2 women})$$

$$= 1 - P(1 \text{ man before 2 women})$$

$$- P(0 \text{ men before 2 women})$$

$$= 1 - \underbrace{(P(MWW) + P(WMW))}_{1 \text{ man}} - \underbrace{P(WW)}_{0 \text{ men}}$$

$$= 1 - \left( 2 \times \left( \frac{2}{5} \right)^2 \times \left( \frac{3}{5} \right) \right) - \left( \frac{2}{5} \right)^2$$

$$= \frac{81}{125}$$

6

(a)

$$P\left(N\left(\frac{1}{3}\right)=2 \mid N(1)=2\right)$$

$$= \frac{P\left(N\left(\frac{1}{3}\right)=2 \cap N(1)=2\right)}{P(N(1)=2)}$$

$$= \frac{P\left(N\left(\frac{1}{3}\right)=2 \cap N\left(\frac{2}{3}\right)=0\right)}{P(N(1)=2)}$$

memory less property

$$= \frac{P\left(N\left(\frac{1}{3}\right)=2\right) P\left(N\left(\frac{2}{3}\right)=0\right)}{P(N(1)=2)}$$

$$= \frac{\frac{e^{-\frac{\lambda}{3}} \times \left(\frac{\lambda}{3}\right)^2}{2!} \times \frac{e^{-\frac{2\lambda}{3}} \times \left(\frac{2\lambda}{3}\right)^0}{0!}}{e^{-\lambda} \frac{\lambda^2}{2!}}$$

$$= \boxed{\frac{1}{9}}$$

$$(b) \quad P\left(N\left(\frac{1}{3}\right) \geq 1 \mid N(1)=2\right) = 1 - P\left(N\left(\frac{1}{3}\right)=0 \mid N(1)=2\right)$$

$$= 1 - \frac{P\left(N\left(\frac{1}{3}\right)=0 \cap N\left(\frac{2}{3}\right)=2\right)}{P(N(1)=2)}$$

$$= 1 - \frac{\frac{e^{-\frac{\lambda}{3}} \left(\frac{\lambda}{3}\right)^0}{0!} \times \frac{e^{-2\lambda} \left(\frac{2\lambda}{3}\right)^2}{2!}}{\frac{e^{-\lambda} \lambda^2}{2!}}$$

$$= 1 - \frac{4}{9} = \boxed{\frac{5}{9}}$$