

# Probability Theory and Random Processes (MA225)

LECTURE SLIDES  
Lecture 09



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# Expectation of Function of RV

**Example 1:** Let the random variable  $X$  be a DRV with PMF

$$f_X(x) = \begin{cases} \frac{1}{7} & \text{if } x = -2, -1, 0, 1 \\ \frac{3}{14} & \text{if } x = 2, 3 \\ 0 & \text{otherwise.} \end{cases}$$

Let  $Y = X^2$ . Find the expectation of  $Y$ .

# Expectation of Function of RV

**Theorem:** Let  $X$  be a DRV with PMF  $f_X(\cdot)$  and support  $S_X$ . Let  $g : \mathbb{R} \rightarrow \mathbb{R}$ . Then

$$E[g(X)] = \sum_{x \in S_X} g(x)f_X(x) \quad \text{provided} \quad \sum_{x \in S_X} |g(x)|f_X(x) < \infty.$$

**Theorem:** Let  $X$  be a CRV with PDF  $f_X(\cdot)$ . Let  $g : \mathbb{R} \rightarrow \mathbb{R}$ . Then

$$E[g(X)] = \int_{-\infty}^{\infty} g(x)f_X(x)dx \quad \text{provided} \quad \int_{-\infty}^{\infty} |g(x)|f_X(x)dx < \infty.$$

# Expectation of Function of RV

**Theorem:** Let  $X$  be a RV (either DRV or CRV). Then

- 1 Let  $A \subset \mathbb{R}$ . Then  $E(I_A(X)) = P(X \in A)$ .
- 2  $h_1(x) \leq h_2(x)$ , for all  $x \in \mathbb{R}$ , then  $E[h_1(X)] \leq E[h_2(X)]$ , provided all the expectations exist.
- 3  $a < b$  are two real numbers such that  $S_X \subset [a, b]$ , then  $a \leq E(X) \leq b$ , provided the expectation exists.
- 4  $E(a + bX) = a + bE(X)$ , where  $a$  and  $b$  are two real numbers.
- 5 Let  $h_1(\cdot), \dots, h_p(\cdot)$  be real valued function of real numbers such that  $E(h_i(X))$  exists for all  $i = 1, 2, \dots, p$ , then

$$E\left(\sum_{i=1}^p h_i(X)\right) = \sum_{i=1}^p E(h_i(X)).$$

# Remarks

- For  $r = 1, 2, \dots$ ,  $\mu_r = E(X^r)$  is called  $r$ th raw moment of  $X$ , if the expectation exists.
- $\mu'_r = E[(X - E(X))^r]$  is called  $r$ th central moment of  $X$ , if the expectations exist.
- $\mu'_2 = E[(X - E(X))^2]$  is called variance of  $X$  when it exists and is denoted by  $Var(X)$ .
- $Var(X) = E(X^2) - (E(X))^2$ .

# Moment Generating Function

**Def:** The moment generating function of random variable  $X$  is defined by

$$M_X(t) = E(e^{tX})$$

provided there exists a real number  $a > 0$  such that the expectation exists for all  $t \in (-a, a)$  (the expectation exists in a neighbourhood of the origin).

**Example 2:**  $X \sim \text{Bin}(n, p)$ , then  $M_X(t) = (1 - p + pe^t)^n$  for all  $t \in \mathbb{R}$ .

**Example 3:**  $X \sim \text{Exp}(\lambda)$ , then  $M_X(t) = (1 - \frac{t}{\lambda})^{-1}$  for all  $t < \lambda$ .

**Example 4:**  $X \sim N(\mu, \sigma^2)$ , then  $M_X(t) = e^{\mu t + \frac{t^2 \sigma^2}{2}}$  for all  $t \in \mathbb{R}$ .

**Def:**  $X$  and  $Y$  are said to be same in distribution if  $F_X(x) = F_Y(x)$  for all  $x \in \mathbb{R}$ .

**Theorem:** Let  $X$  and  $Y$  be two random variables having MGFs  $M_X(\cdot)$  and  $M_Y(\cdot)$ , respectively. Suppose that there exists a positive real number  $a$  such that  $M_X(t) = M_Y(t)$  for all  $t \in (-a, a)$ . Then  $X$  and  $Y$  are same in distribution.

**Example 5:** Let  $X \sim N(\mu, \sigma^2)$ . Find the distribution of  $Y = a + bX$ .

The above is Technique 3 of Transformation of Random Variable.

**Remark:** If the MGF  $M_X(t)$  exist for  $t \in (-a, a)$  for some  $a > 0$ , the derivatives of all order exist at  $t = 0$  and

$$E(X^k) = \left. \frac{d^k}{dt^k} M_X(t) \right|_{t=0}$$

for all positive integer  $k$ .