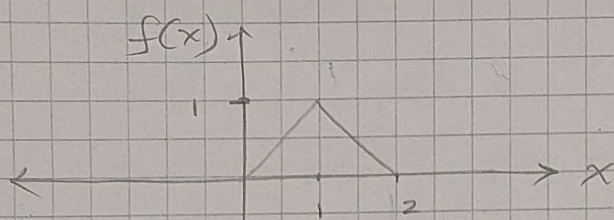


## Lab 1 : Theory.

1)



Put CDF of the above as  $F$ .

If  $t < 0$ , clearly  $F(t) = 0$

If  $t \in [0, 1]$ ,

$$F(t) = \int_{-\infty}^t x \, dx$$

$$= \left[ \frac{x^2}{2} \right]_0^t = \frac{t^2}{2}$$

To elaborate,

$$F(t) = \int_{-\infty}^t f(x) \, dx = \int_0^1 x \, dx$$

If  $t \in [1, 2]$ ,

$$F(t) = \int_{-\infty}^t f(x) \, dx = \int_0^1 x \, dx + \int_1^t (2-x) \, dx$$

$$= \frac{1}{2} + \left[ 2x - \frac{x^2}{2} \right]_1^t$$

$$= \frac{1}{2} + \left( 2t - \frac{t^2}{2} \right) - \left( 2 - \frac{1}{2} \right)$$

$$= 2t - \frac{t^2}{2} - 1$$

$$\therefore F(t) = \begin{cases} 0 & \text{if } t < 0 \\ t^2/2 & \text{if } t \in [0, 1] \\ 2t - t^2/2 - 1 & \text{if } t \in [1, 2] \\ 1 & \text{else} \end{cases}$$

2) Denote this distribution as  $IP$ .

Mean of  $IP = \mu(IP) = \mu$

$$= \mu(IP) = \sum_{x \in [1, 2]} x \cdot IP(x) = 1 \cdot (1-p) + 2 \cdot p$$

$$= 1 + p$$

Variance of  $IP$

$$= \text{Var}(IP) = \sum (x - \mu(IP))^2 \cdot IP(x)$$



$$= (1 - 1 - p)^2 (1 - p) + (2 - 1 - p)^2 (p)$$

~~$$= (p^2 - p) + (p - p^2) = 0$$~~

$$= p^2 (1 - p) + (1 - p)^2 p$$

$$= p(1 - p)(p + 1 - p)$$

$$= p(1 - p)$$

NOTE: The variance is the same as for a regular Bernoulli distribution; shifting the mean does not change the variance!

b) Mean of triangular distr. from (1)

$$= \int_{-\infty}^{\infty} x f(x) dx$$

$$= \int_0^1 x^2 dx + \int_1^2 (2x - x^2) dx$$

$$= \left[ \frac{x^3}{3} \right]_0^1 + \left[ x^2 - \frac{x^3}{3} \right]_1^2$$

$$= \frac{1}{3} + \left( 4 - \frac{8}{3} \right) - \left( 1 - \frac{1}{3} \right)$$

$$= 1 \quad (\text{easily seen in the plot earlier})$$

c) Std. dev. of Uniform(a, b) is

$$\sqrt{\text{Var}(\text{Uniform}(a, b))}$$

$$\text{Var}(\text{Uniform}(a, b))$$

$$= \int_{-\infty}^{\infty} \left( x - \frac{a+b}{2} \right)^2 \frac{1}{b-a} dx$$

$$= \int_a^b \left( x - \frac{a+b}{2} \right)^2 \frac{1}{b-a} dx$$

$$= \left[ \frac{\left( x - \frac{a+b}{2} \right)^3}{3} \right]_a^b \cdot \frac{1}{b-a}$$

$$= \frac{1}{3} \left( \frac{(b-a)^3}{3} - \frac{(a-b)^3}{3} \right) \frac{1}{b-a}$$

$$= \frac{1}{3} ((b-a)^2 + (b-a)^2)$$



$$= \frac{2(b-a)^2}{3}$$

$$\Rightarrow \text{Std. dev.} = (b-a) \sqrt{\frac{2}{3}}$$

3) ~~By~~ By definition of pushforward:

$$f_* \mathbb{P}(-\infty, t] = \mathbb{P}(\{x \mid x^2 \in (-\infty, t]\})$$

(where  $\mathbb{P} = \text{Uniform}(0, 1)$ ) is the probability measure with density  $\text{Uniform}(0, 1)$

$$= \mathbb{P}((-\infty, \sqrt{t}])$$

$$= \int_{-\infty}^{\sqrt{t}} \text{Uniform}(0, 1)(x) dx$$

$$= \int_0^{\sqrt{t}} 1 dx$$

$$= [x]_0^{\sqrt{t}} = \sqrt{t}$$

NOTE: PDF is not the same as probability. If something has prob. density as 1, does not mean it is certain to occur

∴ CDF of  $f_* \mathbb{P}$  is  $\sqrt{t}$  if  $t \in [0, 1]$

if  $t > 1$ , CDF = 1

if  $t < 0$ , CDF = 0

consider only +ve root i.e. assume  $\sqrt{t} = |\sqrt{t}|$

Hence

$$\text{CDF}_{f_* \mathbb{P}}(t) = \begin{cases} 0 & \text{if } t < 0 \\ \sqrt{t} & \text{if } t \in [0, 1] \\ 1 & \text{if } t > 1 \end{cases}$$

To get PDF  $f_* \mathbb{P}$ , we must differentiate

$$\text{CDF}_{f_* \mathbb{P}} : \frac{d \text{CDF}_{f_* \mathbb{P}}}{dt} \text{ is ...}$$

Case 1:  $t < 0$

$$= 0$$

Case 2:  $t \in [0, 1]$

$$= \frac{d \sqrt{t}}{dt} = \frac{1}{2} (t)^{-\frac{1}{2}} = \frac{1}{2\sqrt{t}}$$

Case 3:  $t > 1$

$$= 0$$



$$\therefore \text{PDF}_{f \star P}(x) = \begin{cases} \frac{1}{2^j} & \text{if } t \in [0, 1] \\ 0 & \text{else} \end{cases}$$

4) PDF of Uniform  $(0, 1]$

$$= \begin{cases} 1 & \text{if } x \in (0, 1] \\ 0 & \text{else} \end{cases}$$

Consider

$$A_i = \left( \frac{1}{2^{i+1}}, \frac{1}{2^i} \right], \quad 0 \leq i$$

1st to show all  $A_i$ 's are pairwise disjoint...

Consider  $A_i$  &  $A_j$ , where  $i < j$

Let  $t \in A_i$  &  $t \in A_j$

$$\Rightarrow t \in \left( \frac{1}{2^{i+1}}, \frac{1}{2^i} \right] \text{ & } t \in \left( \frac{1}{2^{j+1}}, \frac{1}{2^j} \right]$$

$$\Rightarrow t < \min \left( \frac{1}{2^{i+1}}, \frac{1}{2^{j+1}} \right) = \frac{1}{2^{j+1}}$$

$$\text{& } t \geq \max \left( \frac{1}{2^i}, \frac{1}{2^j} \right) = \frac{1}{2^i}$$

$$\therefore t \in \left( \frac{1}{2^{j+1}}, \frac{1}{2^i} \right]$$

$$\text{But } A_i \subset \left( \frac{1}{2^{i+1}}, \frac{1}{2^i} \right] \text{ & }$$

$$A_j \subset \left( \frac{1}{2^{j+1}}, \frac{1}{2^j} \right] \text{ & } A_i \neq A_j$$

~~$\Rightarrow A_i$  &  $A_j$~~

$$\Rightarrow \frac{1}{2^{j+1}} < \frac{1}{2^j} \leq \frac{1}{2^{i+1}} < \frac{1}{2^i}$$

$\Rightarrow$  There exist elements in  $A_i$  not in  $A_j$  and vice versa, since their ranges are not equal if  $i \neq j$ .

Consider  $A_i$ .

$$\begin{aligned} A_i \cup A_{i+1} &= \left( \frac{1}{2^{i+1}}, \frac{1}{2^i} \right] \cup \left( \frac{1}{2^{i+2}}, \frac{1}{2^{i+1}} \right] \\ &= \left( \frac{1}{2^{i+2}}, \frac{1}{2^i} \right] \end{aligned}$$



$$\begin{aligned}
 &\Rightarrow A_i \cup A_{i+1} \cup A_{i+2} \\
 &= \left( \frac{1}{2^{i+2}}, \frac{1}{2^i} \right] \cup \left( \frac{1}{2^{i+3}}, \frac{1}{2^{i+2}} \right] \\
 &= \left( \frac{1}{2^{i+3}}, \frac{1}{2^i} \right]
 \end{aligned}$$

In general,

$$\bigcup_{i=0}^n A_i = \left( \frac{1}{2^{n+1}}, \frac{1}{2^0} \right] = \left( \frac{1}{2^{n+1}}, 1 \right]$$

$$\begin{aligned}
 \therefore \lim_{n \rightarrow \infty} \bigcup_{i=0}^n A_i &= \lim_{n \rightarrow \infty} \left( \frac{1}{2^{n+1}}, 1 \right] \\
 &= [0, 1]
 \end{aligned}$$

$$\Rightarrow \bigcup_{i=0}^{\infty} A_i = [0, 1]$$