

9/4/2019 [new]

## LDA: Linear Fisher Discriminator. - Derivation

①

LDA is a special / specific case of MDA.

- only two classes
- and one approach to separate two classes is called Fisher Discriminator, also called as Linear Fisher Discriminator.

### Fisher Discriminator:

Given:

$n$   $d$ -dimensional feature vectors

$$\left. \begin{array}{l} x_1 = (x_{11}, \dots, x_{1d}) \\ x_2 = (x_{21}, \dots, x_{2d}) \\ \vdots \\ x_n = (x_{n1}, \dots, x_{nd}) \end{array} \right\} \begin{array}{l} \text{out of these} \\ n_1 \text{ no. of F.V.} \in \omega_1 \\ n_2 \text{ no. of F.V.} \in \omega_2 \end{array}$$

$n_1$  = no. of F.V. from class  $\omega_1$

$n_2$  = no. of F.V. from class  $\omega_2$ .

$w$  = direction of projection.

$\|w\| = 1$  (unit vector in the direction of projection line)

$\Rightarrow \omega_1 = (x_1, x_2, \dots, x_{n_1}) = n_1 \text{ no. of samples}$

$\omega_2 = (x_1, x_2, \dots, x_{n_2}) = n_2 \text{ no. of samples}$

Suppose if I take orthogonal projection on data  $x_i$  and I get  $y_i$

$$y_i = w^T x_i$$

↓

$$y_i = w^T x_i$$

$$[w_1 \ w_2] \begin{bmatrix} x_1 \\ x_2 \end{bmatrix}$$

$w_1 x_1 + w_2 x_2 = 0$  and this passes through origin.

$$x_1 = x_2$$

$$x_1 - x_2 = 0$$

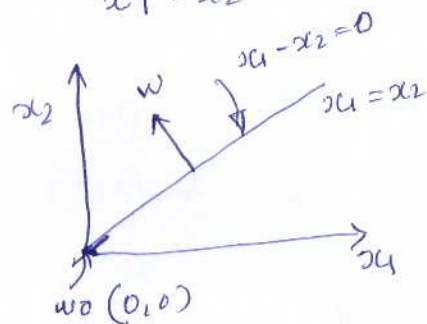
$$w_1 x_1 - w_2 x_2 = 0$$

$$1 \cdot x_1 - 1 \cdot x_2 = 0$$

$$(1 \ -1) \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} = 0$$

$w^T$

$$x_1 - x_2 = 0$$



If I take projections on  $d$ -dimensional F.V  $x_1, \dots, x_n$   
 $y_1, y_2, \dots, y_n$  [these are projected vectors].

$$\underbrace{y_1, y_2, \dots, y_n}_{n_1 \in w_1} \quad \underbrace{\phantom{y_1, y_2, \dots, y_n}}_{n_2 \in w_2}$$

(2)

let  $m_1$  is the mean of class  $w_1$   $\forall x \in w_1$  and class

$$m_1 = \frac{1}{n_1} \sum_{\forall x \in w_1} x$$

$w_2 \forall x \in w_2$

$$\left[ \text{set } X = (\underbrace{x_1, x_2, \dots, x_{n_1}}_{n_1 \text{ no. of samples}}) \right] \text{ and}$$

$$m_2 = \frac{1}{n_2} \sum_{\forall x \in w_2} x$$

$$\text{set } X = \{x_1, x_2, \dots, x_{n_2}\}$$

$n_2$  no. of samples

$$w_1 = (x_1, x_2, \dots, x_{n_1})$$

$$w_2 = (x_1, x_2, \dots, x_{n_2})$$

when I compute the projection of the mean vector,  
the projected mean vector  $m_1$  can be written as  
 $\tilde{m}_1$ .

$$\tilde{m}_1 = \frac{1}{n_1} \sum_{\forall y \in w_1} y \quad \leftarrow \text{projected vector of } x.$$

$y = w^t x$

$$\tilde{m}_1 = \frac{1}{n_1} \sum_{\forall x \in w_1} w^t x$$

$$= \frac{1}{n_1} w^t \sum_{\forall x \in w_1} x$$

$$= w^t \frac{1}{n_1} \sum_{\forall x \in w_1} x = w^t m_1$$

$$\boxed{\tilde{m}_1 = w^t m_1}$$

$\tilde{m}_1 \Rightarrow$  Projection of mean of vectors of <sup>class</sup> set  $w_1$  is nothing but  $w^t m_1$ .

In the same manner, I can compute, the projection of mean of samples of class  $w_2$  which is nothing but

$$\tilde{m}_2 = w^t m_2$$

now, what is the distance between these two projected mean?

Distance between projected means:

$$|\tilde{m}_1 - \tilde{m}_2| = |w^t m_1 - w^t m_2|$$

$$= |w^t (m_1 - m_2)|$$

vector  $w$ .

from, this I can ~~simply~~ increase the distance between  $|\tilde{m}_1 - \tilde{m}_2|$  by simply ~~scaling~~ scaling of ' $w$ '.

Here I assume  $\|w\| = 1$ . If the  $\|w\| > 1$  (more than 1)

then the distance between  $|\tilde{m}_1 - \tilde{m}_2|$  will go on

increasing.

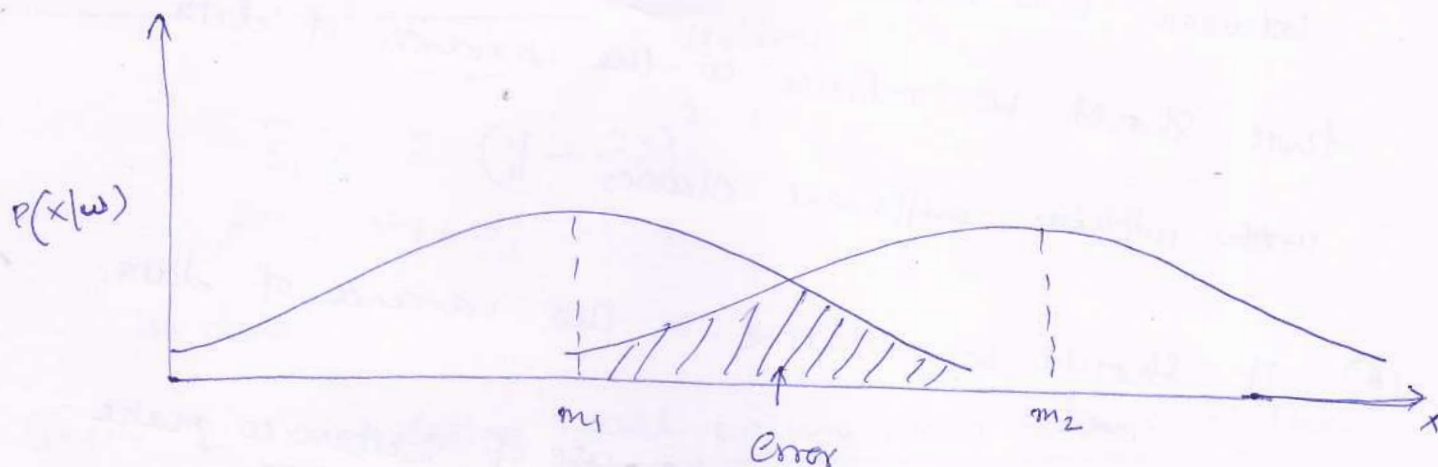
\* This ensures that I can increase the separability between two classes

But, the question is how much should I increase to separate the two classes?



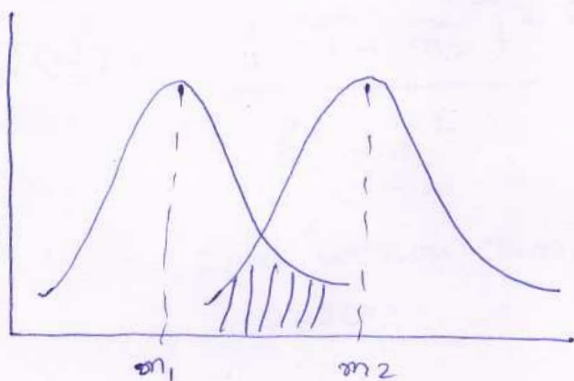
### Illustration:

If the variance is large:



We should have the difference between  $m_1$  and  $m_2$  should be quite large so that the error of classification is reduced.

If the variance is small:



⊛ In this case, we don't need that much separation which we need in the above case.

⊛ The separation between  $m_1$  and  $m_2$  is much less than the above ~~one~~ case. (even then the error is minimized)

⑧ So how much should be the difference between two projected mean in the reduced space, that should be relative to the variance of data ~~with~~ within different classes.

⑨ It should be relative to the variance of data.

⑩ Accordingly, we can make use of Scatter to make some Criterion function.

Scatter of projected data:

$$\tilde{s}_i^2 = \overbrace{\sum_{y \in w_i} (y - \tilde{m}_i)^2}^{\text{variance}} \Rightarrow \text{we have not normalized } \frac{1}{n}.$$

$\nearrow$   
 its class

## Fisher Discriminant method:

④

### Scatter of projected data:

$$\tilde{S}_i = \sum_{y \in w_i} (y - \tilde{m}_i)^2$$

Variance.

↑  
y ∈ w<sub>i</sub>

i<sup>th</sup> class

Then, I can define total within class scatter of the projected samples.

$$S = \tilde{S}_1^2 + \tilde{S}_2^2$$

We want the distance between two mean should be relative to this total scatter 'S'. So we can define the criterion function as follows.

$$J(w) = \frac{|\tilde{m}_1 - \tilde{m}_2|^2}{\tilde{S}_1^2 + \tilde{S}_2^2}$$

difference b/w projected mean

⇒ as large as possible w.r.t the total within class scatter.

total within class scatter.

We want to maximize  $J(w)$  which is the ratio of

$|\tilde{m}_1 - \tilde{m}_2|^2$  upon  $\tilde{S}_1^2 + \tilde{S}_2^2$ .  $J(w)$  is ratio of inter class scatter to intra class scatter.

$$J(w) = \frac{\text{Inter class}}{\text{Intra class}} = \frac{|\tilde{m}_1 - \tilde{m}_2|^2}{\tilde{S}_1^2 + \tilde{S}_2^2}$$



Fisher Discriminant maximize the

$$J(w) = \frac{|\tilde{m}_1 - \tilde{m}_2|^2}{s_1^2 + s_2^2}$$

The value of 'w' which maximizes  $J(w)$  is the projection direction.

$J(\cdot)$  in terms of 'w'

$S_i$   
↑

Scatter within ith class

$S_w \rightarrow$  total within class scatter.

$$S_i = \sum_{x \in w_i} (x - m_i)(x - m_i)^t$$

Scatter for individual class  
(ith class)

we can define, total within class scatter as

$$S_w = \sum_i S_i \quad \text{here, we consider only two classes}$$

$$S_w = S_1 + S_2$$



scatter of projected samples:

⑤

$$\tilde{S}_i = \sum_{\forall y \in w_i} (y - \tilde{m}_i)^2$$

$$\tilde{S}_i^2 = \sum_{\forall x \in w_i} (w^t x_i - w^t m_i)^2$$

$$y_i = w^t x_i$$

by rearranging this,

$$\tilde{S}_i^2 = \sum_{\forall x \in w_i} w^t (x_i - m_i) (x_i - m_i)^t w$$

$$\tilde{S}_i^2 = \sum_{\forall x \in w_i} (w^t x - w^t m_i)^2$$

$$= \sum_{\forall x \in w_i} (w^t x - w^t m_i) \cdot (w^t x - w^t m_i)$$

$$= \sum w^t (x - m_i) \cdot \underbrace{w^t (x - m_i)}$$

$$= \sum w^t (x - m_i) \cdot (x - m_i)^t w$$

$$= w^t \left[ \sum_{\forall x \in w_i} (x - m_i) (x - m_i)^t \right] \cdot w$$

$$\boxed{\tilde{S}_i^2 = w^t S_i w}$$

then the sum of  $\tilde{s}_1^2 + \tilde{s}_2^2$

$$\tilde{s}_1^2 + \tilde{s}_2^2 = w^t S_1 w + w^t S_2 w$$

$$= w^t (S_1 + S_2) w$$

$$= w^t \underbrace{S_w}_{\substack{\uparrow \\ \text{total within class scatter} \text{ [refer page 4]}}} w$$

In the same manner,

separation of the projected means.

$$(\tilde{m}_1 - \tilde{m}_2)^2 = (w^t m_1 - w^t m_2)^2$$

$$= (w^t m_1 - w^t m_2) \cdot (w^t m_1 - w^t m_2)$$

$$= w^t (m_1 - m_2) \cdot \underbrace{w^t (m_1 - m_2)}_{\substack{\downarrow \\ \text{column vector}}}$$

$$= w^t \underbrace{(m_1 - m_2)}_{\substack{\downarrow \\ \text{row vector}}} \cdot \underbrace{(m_1 - m_2)^t}_{\substack{\downarrow \\ \text{column vector}}} w$$

$$= w^t \underbrace{(m_1 - m_2)(m_1 - m_2)^t}_{\substack{\text{outer product of 2 vectors} \\ \text{rank is at most 1}}} w$$

$$= w^t \underbrace{(S_B w)}_{\substack{\downarrow \\ \text{between class scatter}}} \Rightarrow \text{is a vector in the direction of } (m_1 - m_2)$$

Observation:  $w^t \underbrace{S_B w}_{\substack{\downarrow \\ \text{between class scatter}}} \Rightarrow \text{is a vector in the direction of } (m_1 - m_2)$

In particular, for any  $w$ ,  $S_B w$  is in the direction of  $m_1 - m_2$ , and  $S_B$  is quite singular.

⑥

$$J(w) = \frac{|\tilde{m}_1 - \tilde{m}_2|^2}{\tilde{s}_1^2 + \tilde{s}_2^2}$$

in terms of 'w'

$$J(w) = \frac{\overset{\substack{\downarrow \text{ between class scatter}}}{w^t S_B w}}{\underset{\substack{\uparrow \text{ total within class scatter}}}{w^t S_W w}} \Rightarrow \text{Generalized Rayleigh Coefficient.}$$

maximize this ratio by varying vector 'w' and 'w' gives the projection direction.

Take a derivative w.r.t 'w' and equal it to zero.

$$J(w) = \frac{w^t S_B w}{w^t S_W w} = \frac{u}{v} = \frac{v \cdot u' - u \cdot v}{v^2}$$

$$\frac{d}{dw} J(w) = \frac{w^t S_W w \cdot \frac{d}{dw} (w^t S_B w) - w^t S_B w \frac{d}{dw} (w^t S_W w)}{(w^t S_W w)^2}$$

$$0 = \frac{w^t S_W w [2 S_B w] - w^t S_B w [2 S_W w]}{(w^t S_W w)^2}$$

Divide  
 $w^t S_W w$

$$w^t S_W w [2 S_B w] = w^t S_B w [2 S_W w]$$

$$2 S_B w = \left( \frac{w^t S_B w}{w^t S_W w} \right) 2 S_W w.$$



$$S_B w = \lambda S_w w$$

$$\boxed{S_B w = \lambda S_w w} \Rightarrow \text{generalized eigenvalue problem}$$

If  $S_w$  is non-singular then

$S_w^{-1}$  exists.

Singular  $\equiv$  if the determinant of matrix

$$|A| = 0$$

then singular

$$A^{-1} = \frac{1}{|A|} [ \quad ]$$

Inverse does not exist.

for singular matrix

$$\boxed{S_w^{-1} S_B w = \lambda w} \Rightarrow \text{eigenvalue problem}$$

$(m_1 - m_2)$

eigen vector of  $S_w^{-1} S_B$

and  $\lambda$  is the corresponding eigenvalue.

It is unnecessary to solve for the eigenvalues & vectors of  $S_w^{-1} S_B$  due to the fact that  $S_B w$  is always in the direction of  $m_1 - m_2$  (in  $\lambda w$ )

The scaling factor of 'w' is not important [in  $\lambda w$ ]

$[S_B w = (m_1 - m_2)]$  but the direction of 'w' is so important.

$$S_w^{-1} S_B w = \lambda w$$

$$S_w^{-1} K (m_1 - m_2) = \lambda w$$

$$S_w^{-1} \left( \frac{K}{\lambda} \right) (m_1 - m_2) = w$$

$$S_w^{-1} (K) (m_1 - m_2) = w$$

$$\boxed{w = S_w^{-1} (m_1 - m_2)}$$