HW1 Peer Assessment

Part A. ANOVA Additional Material: ANOVA tutorial

https://datascienceplus.com/one-way-anova-in-r/

Jet lag is a common problem for people traveling across multiple time zones, but people can gradually adjust to the new time zone since the exposure of the shifted light schedule to their eyes can resets the internal circadian rhythm in a process called "phase shift". Campbell and Murphy (1998) in a highly controversial study reported that the human circadian clock can also be reset by only exposing the back of the knee to light, with some hailing this as a major discovery and others challenging aspects of the experimental design. The table below is taken from a later experiment by Wright and Czeisler (2002) that re-examined the phenomenon. The new experiment measured circadian rhythm through the daily cycle of melatonin production in 22 subjects randomly assigned to one of three light treatments. Subjects were woken from sleep and for three hours were exposed to bright lights applied to the eyes only, to the knees only or to neither (control group). The effects of treatment to the circadian rhythm were measured two days later by the magnitude of phase shift (measured in hours) in each subject's daily cycle of melatonin production. A negative measurement indicates a delay in melatonin production, a predicted effect of light treatment, while a positive number indicates an advance. Raw data of phase shift, in hours, for the circadian rhythm experiment

Phase Shift (hr) **Treatment** Control 0.53, 0.36, 0.20, -0.37, -0.60, -0.64, -0.68, -1.27

Knees 0.73, 0.31, 0.03, -0.29, -0.56, -0.96, -1.61 -0.78, -0.86, -1.35, -1.48, -1.52, -2.04, -2.83 Eyes treatment <- c('Control','Contr

Creating data

'Knees', 'Kn 'Eyes', 'Eyes', 'Eyes', 'Eyes', 'Eyes', 'Eyes') shift < c(0.53, 0.36, 0.20, -0.37, -0.60, -0.64, -0.68, -1.27, 0.73, 0.31, 0.03, -0.29, -0.56, -0.96, -1.61, -0.78, -0.86, -1.35, -1.48, -1.52, -2.04, -2.83) df <- data.frame(treatment, shift)</pre> print(df) treatment shift ## 1 Control 0.53 ## 2 Control 0.36 ## 3 Control 0.20 ## 4 Control -0.37 ## 5 Control -0.60

6 Control -0.64 ## 7 Control -0.68 ## 8 Control -1.27 ## 9 Knees 0.73 ## 10 Knees 0.31 ## 11 Knees 0.03 ## 12 Knees -0.29 ## 13 Knees -0.56 ## 14 Knees -0.96

15 Knees -1.61 ## 16 Eyes -0.78 ## 17 Eyes -0.86 ## 18 Eyes -1.35 Eyes -1.48 ## 19 ## 20 Eyes -1.52 Eyes -2.83 Question A1 - 3 pts Consider the following incomplete R output: **Sum of Squares** F-statistics Source **Mean Squares** p-value ? 3.6122 0.004 **Treatments** 9.415 ? Error **TOTAL** ? Fill in the missing values in the analysis of the variance table. Note: Missing values can be calculated using the corresponding formulas provided in the lectures, or you can build the data frame in R and generate the ANOVA table using the aov() function. Either approach will be accepted.

Df Sum Sq Mean Sq F value Pr(>F) ## treatment 2 7.224 3.612 7.289 0.00447 ** ## Residuals 19 9.415 0.496 ## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

model0 = aov(shift ~ treatment, data = df)

Control Eyes Knees

different mean effect on phase shift

• performance: published relative performance of the CPU

data = read.csv("machine.csv", head = TRUE, sep = ",")

trend (direction and form). Include plots and R-code used.

0

50

Residual standard error: 128.3 on 207 degrees of freedom ## Multiple R-squared: 0.3663, Adjusted R-squared: 0.3632 ## F-statistic: 119.6 on 1 and 207 DF, p-value: < 2.2e-16

The correlation coefficient is 0.6052. This is a fairly strong positive correlation.

in chmax, which is not that good if it will be used as the sole predictor.

c. **1 pts** Fill in the blanks for the degrees of freedom of the ANOVA F-test statistic:

different from the R working directory) and read the data using the R function read.csv().

rep 8.0000 7.000 7.0000

phase shift

F(2, 19)

Read in the data

vendor

<chr>

attach(data)

200

0

Your code here... cor(chmax, performance)

[1] 0.6052093

Coefficients:

confint(model1, level=.95)

chmax

tvalue=10.938 1-pt(tvalue, 207) 2.5 %

3.069251 4.418926

(Intercept) 15.817392 58.633048

more linear relationshipl

transform the data. The function you should use in R is:

Estimate Std. Error t value Pr(>|t|) ## (Intercept) 37.2252 10.8587 3.428 0.000733 *** ## chmax 3.7441 0.3423 10.938 < 2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 128.3 on 207 degrees of freedom ## Multiple R-squared: 0.3663, Adjusted R-squared: 0.3632 ## F-statistic: 119.6 on 1 and 207 DF, p-value: < 2.2e-16

a. 3 pts What are the model parameters and what are their estimates?

97.5 %

assess each one. Each graph may be used to assess one or more model assumptions.

50

b. **3 pts** Residual plot - a plot of the residuals, $\hat{\epsilon}_i$, versus the fitted values, \hat{y}_i

0

suggests the errors are not correlated (related to but not proof of independence assumption)

QQ-Plot of Residuals

0

Theoretical Quantiles

Histogram of resid(model1)

c. **3 pts** Histogram and q-q plot of the residuals

qqline(resid(model1))

800

009

400

200

0

-400

-3

hist(resid(model1), breaks=20)

boxCox(model1,lambda=seq(-1,1,.5))

95%

-1.0

lchmax <- log(chmax+1)</pre> lperf <- log(performance)</pre>

summary(model2)

Call:

model2 = lm(lperf ~ lchmax, data)

-0.5

natural log. Please add one to the predictor before taking the natural log of it

Residual standard error: 0.807 on 207 degrees of freedom ## Multiple R-squared: 0.4103, Adjusted R-squared: 0.4074 ## F-statistic: 144 on 1 and 207 DF, p-value: < 2.2e-16

0 00 00

0

1

000000

2

Based on your interpretation of the model assumptions, is model2 a good fit?

improve the explanatory power somewhat

plot(lchmax, lperf)

9

2

4

3

7

-2

15

10

5

Frequency

-3

hist(resid(model2), breaks=20)

-2

harness more channels, leading to those few leverage/outlier points.

data2 = data[data\$vendor %in% c("honeywell", "hp", "nas"),]

Part C. ANOVA - 8 pts

Filter for honeywell, hp, and nas

data2\$vendor = factor(data2\$vendor)

Your code here...

below.

-2

-1

0

-1400

-1500

log-likelihood

-2

-1

Sample Quantiles

qqnorm(resid(model1), main="QQ-Plot of Residuals")

Model Assumption(s) it checks: Linearity/mean 0 assumption

Interpretation: It appears generally linear, but it is imperfect.

800

plot(chmax, performance)

head(data, 3)

Show the first few rows of data

summary(model0)

-0.7127273

model.tables(model0, type="means") ## Tables of means ## Grand mean

treatment Control Eyes Knees -0.3087 -1.551 -0.3357 ## rep 8.0000 7.000 7.0000 Df

Source **Sum of Squares Mean Squares** F-statistics p-value **Treatments** 2 7.224 . 3.6122 7.289 0.004. 0.496 19 9.415 Error. TOTAL. 21 16.639

Question A2 - 3 pts Use μ_1 , μ_2 , and μ_3 as notation for the three mean parameters and define these parameters clearly based on the context of the topic above (i.e. explain what μ_1 , μ_2 , and μ_3 mean in words in the context of this problem). Find the estimates of these parameters. model.tables(model0, type="means") ## Tables of means ## Grand mean ## ## -0.7127273 ## treatment

 μ_1 : Average phase shift (measured in hours) in each subject's daily cycle of melatonin production for subjects in the control group (no light) Value: μ_2 : Average phase shift (measured in hours) in each subject's daily cycle of melatonin production for subjects in the eyes group Value: -1.551 hours μ_3 Average phase shift (measured in hours) in each subject's daily cycle of melatonin production for subjects in the knees group Value: -0.3357 Question A3 - 5 pts Use the ANOVA table in Question A1 to answer the following questions:

a. **1 pts** Write the null hypothesis of the ANOVA F-test, H_0 H_0 : All three groups (2 treatments + control) will have the same mean effect on

b. **1 pts** Write the alternative hypothesis of the ANOVA F-test, H_A H_1 : At least 1 of the three groups (2 treatments + control) will have a

d. **1 pts** What is the p-value of the ANOVA F-test? p = .00447 e. **1 pts** According the the results of the ANOVA F-test, does light treatment affect phase shift? Use an α -level of 0.05. Yes Part B. Simple Linear Regression We are going to use regression analysis to estimate the performance of CPUs based on the maximum number of channels in the CPU. This data set comes from the UCI Machine Learning Repository. The data file includes the following columns: · vendor: vendor of the CPU • chmax: maximum channels in the CPU

The data is in the file "machine.csv". To read the data in R, save the file in your working directory (make sure you have changed the directory if

chmax

a. 3 pts Use a scatter plot to describe the relationship between CPU performance and the maximum number of channels. Describe the general

<int>

performance

There seems to be a positive

198

269

220

adviser 128 2 32 amdahl 3 amdahl 32 3 rows

Question B1: Exploratory Data Analysis - 9 pts

0 0 0 800 performance

0

100

chmax

150

relationship between performance and max # of channels, though there are some outliers and/or leverage points that have chmax much higher than the bulk of the group. It also appears the data is a "cone" shape so the data is not linear, meaning that chmax alone does not explain all of the variability in performance. b. **3 pts** What is the value of the correlation coefficient between *performance* and *chmax*? Please interpret the strength of the correlation based on the correlation coefficient. # Your code here... $model = lm(performance \sim chmax)$ summary(model) ## Call: ## lm(formula = performance ~ chmax) ## Residuals: Min 1Q Median 3Q ## -486.47 -42.20 -22.20 20.31 867.15 ## Coefficients: Estimate Std. Error t value Pr(>|t|)## (Intercept) 37.2252 10.8587 3.428 0.000733 *** 3.7441 0.3423 10.938 < 2e-16 *** ## chmax ## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Your code here... $model1 = lm(performance \sim chmax, data)$ summary(model1) ## ## Call: ## lm(formula = performance ~ chmax, data = data) ## ## Residuals: Min 1Q Median 3Q Max ## -486.47 -42.20 -22.20 20.31 867.15

The model parameters are chmax and the intercept. The estimate of chmax is 3.7441 and the estimate of the intercept is 37.2252

c. **2 pts** Interpret the estimated value of the β_1 parameter in the context of the problem. β_1 , which is chmax in this case, is 3.7441. This

b. **2 pts** Write down the estimated simple linear regression equation. $y = 3.7441 \cdot chmax + 37.2252 + e$

d. **2 pts** Find a 95% confidence interval for the β_1 parameter. Is β_1 statistically significant at this level?

suggests that for every additional maximum channel in the CPU, the performance metric will imporve by 3.7441

The 95% confidence interval for β_1 is (3.069251,4.418926). That makes β_1 statistically significant (non zero) at that level.

e. **2 pts** Is β_1 statistically significantly positive at an α -level of 0.01? What is the approximate p-value of this test?

c. 2 pts Based on this exploratory analysis, would you recommend a simple linear regression model for the relationship? I would say no. The coefficient of determination (r-squared) is .3663, indicating that only one third of the variability in performance can be explained by variability

d. 1 pts Based on the analysis above, would you pursue a transformation of the data? Do not transform the data. Yes, the funnel shape and the clustering of points at the lower end of chmax suggests to me that a logarithmic transformation may compress the data and lead to a

Question B2: Fitting the Simple Linear Regression Model - 11 pts

Fit a linear regression model, named *model1*, to evaluate the relationship between performance and the maximum number of channels. *Do not*

[1] 0 #confint(model1, level=.95) The p-value of the one sided test is so small that R interprets it as 0, so yes the β_1 is statistically significantly positive at an α -level of 0.01 Question B3: Checking the Assumptions of the Model - 8 pts

Create and interpret the following graphs with respect to the assumptions of the linear regression model. In other words, comment on whether there are any apparent departures from the assumptions of the linear regression model. Make sure that you state the model assumptions and

a. **2 pts** Scatterplot of the data with *chmax* on the x-axis and *performance* on the y-axis plot(chmax, performance) 0 1000 800 performance 900 0 400 0 200 0 0 0

100

chmax

plot(fitted(model1), resid(model1), main="Fitted vs Residuals", xlab="Fitted Values")

Fitted vs Residuals

150

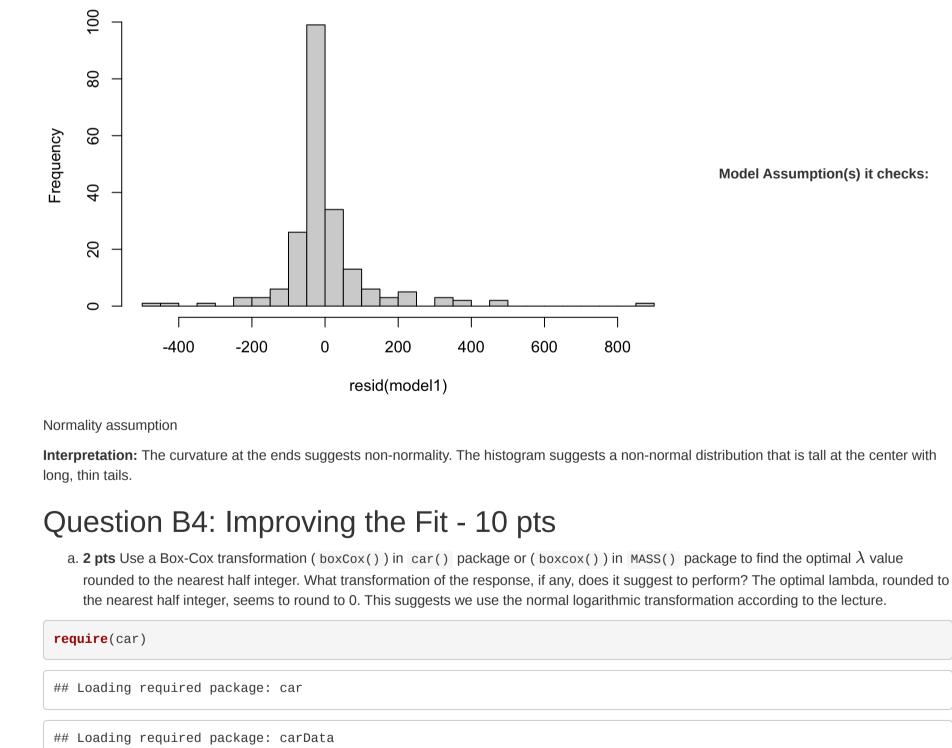
009 400 0 resid(model1) 200 0 -400 100 200 600 700 300 400 500 Fitted Values **Model Assumption(s)** it checks: Constant variance assumption Independence assumption (not really, we can only check for uncorrelated errors) Interpretation: There is no clear shape to the distribution (such as a curve) that suggests there is a non-linear relationship, but we see a

megaphone effect on the residuals, which means the constant variance assumption does not hold. There are no clusters of residuals, which

 ∞ $^{\circ}$

2

Model Assumption(s) it checks:



Profile Log-likelihood

0.0

λ

lm(formula = lperf ~ lchmax, data = data) ## Residuals: Min **1**Q Median 3Q ## -2.22543 -0.59429 0.01065 0.59287 1.85995 ## Coefficients: Estimate Std. Error t value Pr(>|t|)<2e-16 *** ## (Intercept) 2.47655 0.14152 17.5 <2e-16 *** 0.64819 12.0 ## lchmax 0.05401 ## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

c. **2 pts** Compare the R-squared values of *model1* and *model2*. Did the transformation improve the explanatory power of the model? Model 1: Multiple R-squared: 0.3663, Adjusted R-squared: 0.3632 Model 2: Multiple R-squared: 0.4103, Adjusted R-squared: 0.4074 Yes, it did

d. **4 pts** Similar to Question B3, assess and interpret all model assumptions of *model2*. A model is considered a good fit if all assumptions hold.

0 0

0

0

0

5

@ 000 @

3

00 **0**

Ichmax

Fitted vs Residuals

plot(fitted(model2), resid(model2), main="Fitted vs Residuals", xlab="Fitted Values")

0.5

b. **2 pts** Create a linear regression model, named *model2*, that uses the log transformed *performance* as the response, and the log

transformed *chmax* as the predictor. Note: The variable *chmax* has a couple of zero values which will cause problems when taking the

1.0

0 resid(model2) 0 00 0 -2 0 2.5 3.0 3.5 4.0 5.0 5.5 4.5 Fitted Values qqnorm(resid(model2), main="QQ-Plot of Residuals") qqline(resid(model2)) **QQ-Plot of Residuals** Sample Quantiles 0 7

0

Theoretical Quantiles

Histogram of resid(model2)

0

and log(chmax) shows strong linearity The scatterplot of the fitted values vs. the residual looks balanced and evenly distributed, suggesting

resid(model2)

2

The scatterplot of log(performance)

constance variance and independence. The Q-Q plot is close to linear, with limited variance on the tails, and the histogram looks much more like a normal distribution, suggesting normality. Based on these charts, model2 seems like a better fit than model 1 Question B5: Prediction - 3 pts Suppose we are interested in predicting CPU performance when chmax = 128. Please make a prediction using both *model1* and *model2* and provide the 95% prediction interval of each prediction on the original scale of the response, performance. What observations can you make about the result in the context of the problem? predict() new = data.frame(chmax = 128)predict(model1, new, interval='prediction', level=.95) fit lwr upr ## 1 516.4685 252.2519 780.6851 new2 = data.frame(lchmax = log(c(128)))exp(predict(model2, new2, interval='prediction', level=.95)) fit lwr ## 1 276.3256 54.90877 1390.594 #predict(model2, log(new+1)) #buch[50]-exp(predict(model.red, new)) The model1 prediction is almost twice as high as the model2 prediction. Visually, when I look at the scatterplot of the original data, it seems that the leverage points that have high performance may be influencers that drag up the slope of the fitted linear regression model, resulting in larger estimates. Since model2 performed better on all the metrics we tested (better met assumptions, higher R-squared), I believe the model2 prediction to be more reliable. It is also worth noting that chmax of 128 is fairly large in the sample we trained these on, and other than 2 leverage points, all the points at this chmax or greater are much lower. Without knowing too much about the underlying problem space, it may be the case that on average, large chmax doesn't continue to lead to performance growth except in a few cases where something about the technology allows it to

400 Performance The boxplots suggest that hp has the 200 100

We are going to continue using the CPU data set to analyse various vendors in the data set. There are over 20 vendors in the data set. To simplify the task, we are going to limit our analysis to three vendors, specifically, honeywell, hp, and nas. The code to filter for those vendors is provided

boxplot(data2\$performance~as.factor(data2\$vendor), main="Boxplot of performance by vendor",xlab="Vendor",ylab="Pe

1. **2 pts** Using data2, create a boxplot of *performance* and *vendor*, with *performance* on the vertical axis. Interpret the plots.

Boxplot of performance by vendor

honeywell hp nas Vendor lowest performance products, honeywell is somewhat higher with a few outliers with moderate performance, and nas specializes in high performance models, including the extreme outliers. 2. **3 pts** Perform an ANOVA F-test on the means of the three vendors. Using an α -level of 0.05, can we reject the null hypothesis that the means of the three vendors are equal? Please interpret. model3 = aov(performance ~ vendor, data = data2) summary(model3) Df Sum Sq Mean Sq F value Pr(>F) ## vendor 2 154494 77247 6.027 0.00553 ** ## Residuals 36 461443 12818 ## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1 model.tables(model3, type="means") ## Tables of means ## Grand mean ## 112.8718 ## ## vendor honeywell hp nas 60.46 36.43 176.9

13.00 7.00 19.0 one of the means is different than the others. different from each other? TukeyHSD(model3, conf.level=.95) Tukey multiple comparisons of means ## 95% family-wise confidence level ## Fit: aov(formula = performance ~ vendor, data = data2)

140.46617 18.11095 262.8214 0.0214092

nas-honeywell and nas-hp are statistically significantly different from each other according to the Tukey pairwise comparison.

nas-hp

Yes, the P value of the F-test is .00553, which is significant at an alpha level of .05. This means that there is strong statistical evidence that at least 3. **3 pts** Perform a Tukey pairwise comparison between the three vendors. Using an α -level of 0.05, which means are statistically significantly ## \$vendor diff lwr upr p adj ## hp-honeywell -24.03297 -153.76761 105.7017 0.8934786