

Deterministic Signals : Amplitude described by
math formula
Random Signals : Precise description
difficult

Let $S = \{\zeta_1, \zeta_2, \dots\}$ be universal set

\mathcal{F} = Collection of subsets of 'S' with S itself
is called σ field.

eg. $\mathcal{F} = \{(\zeta_1, \zeta_2), (\zeta_1, \zeta_2, \zeta_3)\}$

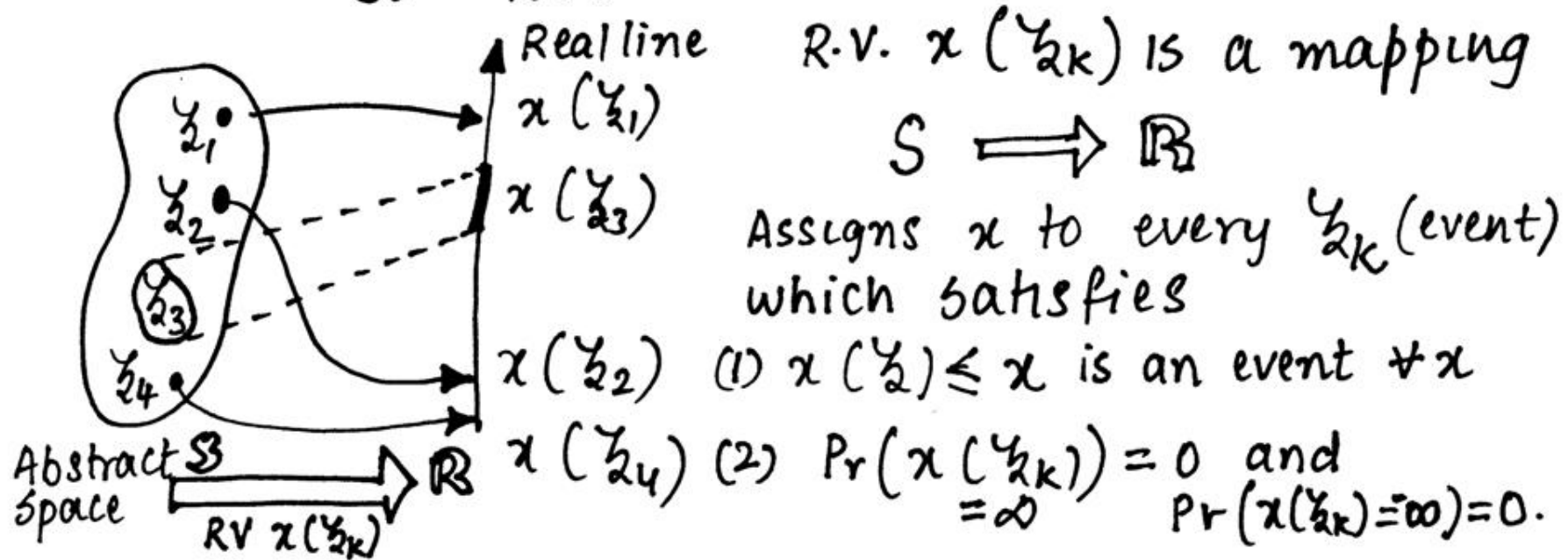
Elements of \mathcal{F} called Events

$Pr\{\zeta_k\}$, $k=1, 2, \dots$ is called Probability of
event ζ_k

The Probability Space

Probability space $= (S, \mathcal{F}, Pr)$

- * Basic Space contains abstract events/outcomes difficult to manipulate
- * $Pr\{\cdot\}$ function difficult to manipulate
So 'RV'.



Complex Random Variable

Complex RV $x(\omega) = \underbrace{x_R(\omega)}_{\text{real valued RV}} + j \underbrace{x_I(\omega)}_{\text{real valued RV}}$

RV - is neither 'R' nor 'V', but a mapping

However $\underbrace{x(\omega)}_{\text{RV}} = \underbrace{x}_{\text{value of RV}}$

if x is discrete valued $\{x_k\}$; Discrete RV
else Continuous RV.
There are Mixed RV.

CDF, PDF, and PMF

• CDF of $x(\zeta)$: $F_x(x) = \Pr \{ x(\zeta) < x \}$

where $\Pr \{ x(\zeta) \leq x \}$ is a function of the set $\{ x(\zeta) \leq x \}$

Pdf of $x(\zeta)$ $f_x(x) = \frac{d}{dx} \underbrace{F_x(x)}_{\text{CDF}}$

But $f_x(x) \underbrace{\Delta x}_{\text{interval}} = \Delta F_x(x) = F_x(x + \Delta x) - F_x(x)$
 $= \underbrace{\Pr(x < x(\zeta) \leq x + \Delta x)}_{\text{Probability}}$

Properties

$f_X(x) = \frac{d}{dx} F_X(x)$; Integrate both sides

$$F_X(x) = \int_{-\infty}^x f_X(v) dv$$

For discrete RV: pmf $p_k = \Pr\{X(\omega) = x_k\}$

Properties:

$$0 \leq F_X(x) \leq 1 ; F_X(-\infty) = 0 ; F_X(\infty) = 1 ;$$

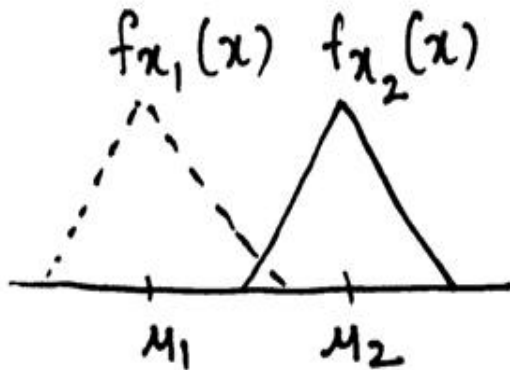
$$f_X(x) \geq 0 ; \int_{-\infty}^{\infty} f_X(x) dx = 1$$

$$\therefore \Pr\{x_1 < X(\omega) \leq x_2\} = F_X(x_2) - F_X(x_1) = \int_{x_1}^{x_2} f_X(x) dx$$

Moments of RV/PDF

RVs characterized by ~~PDFs~~ Pdf
Pdf represented by Moments

$$E \{ x(\zeta) \} = \underline{\mu_x} = \sum_{k=-\infty}^{\infty} x_k p_k \quad ; \quad x(\zeta) \text{ discrete}$$
$$\int_{-\infty}^{\infty} x f_x(x) dx \quad ; \quad x(\zeta) \text{ Continuous}$$



$f_x(x)$ is symmetric abt. 'a'
then $\mu_x = a$

Moments

Note : $E [\alpha x(\frac{Y}{2}) + \beta] = \alpha \mu(x) + \beta$

* $E \left[\underbrace{g(x(\frac{Y}{2}))}_{y(\frac{Y}{2})\text{-which is a transformation of } x(\frac{Y}{2})} \right] = \int_{-\infty}^{\infty} g(x) f_x(x) dx$

$$E [x^m(\frac{Y}{2})] = \int_{-\infty}^{\infty} x^m f_x(x) dx = r_x^{(m)}$$

$r_x^{(m)}$ = m^{th} order moment of $x(\frac{Y}{2})$

$r_x^{(2)}$ = Mean Squared value. $[E(x^2(\frac{Y}{2})) \neq E^2(x(\frac{Y}{2}))]$

Central Moments

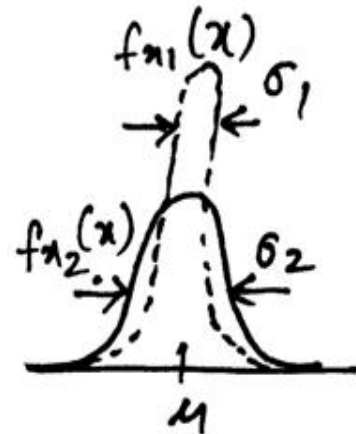
• central moments : $\gamma_x^{(m)} = E\{[x(\frac{1}{2}) - \mu_x]^m\}$

mth order central : $\gamma_x^{(m)} = \int (x - \mu_x)^m f_x(x) dx$.
moment

$$\gamma_x^{(2)} = \sigma_x^2 = \text{var}[x(\frac{1}{2})] = E\{[x(\frac{1}{2}) - \mu_x]^2\}$$

$$\sigma_x = \sqrt{\gamma_x^{(2)}} = \text{Std. Deviation}$$

Also $\gamma_x^{(m)} = \sum_{k=0}^m \underbrace{\binom{m}{k} (-1)^k \mu_x^k \gamma_x^{(n-k)}}_{\text{moments}}$
central moment



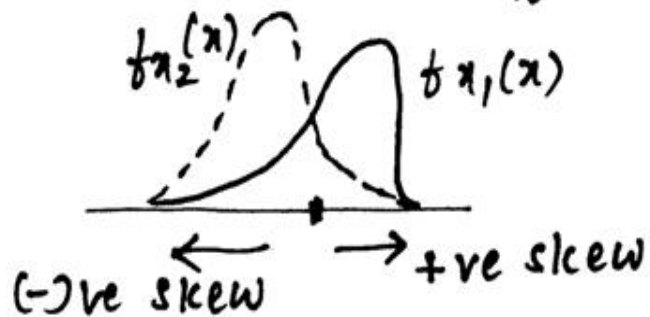
$$\sigma_x^2 = \gamma_x^{(2)} - \mu_x^2 = E[x^2(\frac{1}{2})] - E^2[x(\frac{1}{2})]$$

Higher Order Moments

Skewness : Degree of asymmetry around ' μ ' of the ~~dist~~

Normalized 3rd order central moment

$$\text{Skew } \tilde{k}_x^{(3)} = E \left\{ \left[\frac{x(\frac{1}{2}) - \mu_x}{\sigma_x} \right]^3 \right\} = \frac{\gamma_x^{(3)}}{\sigma_x^3}.$$



Kurtosis: Relative flatness of a distn. about μ

$$\text{4th order: } \tilde{k}_x^{(4)} = E \left\{ \left[\frac{x(\frac{1}{2}) - \mu_x}{\sigma_x} \right]^4 \right\} - 3 = \frac{\gamma_x^{(4)}}{\sigma_x^4} - 3$$

Characteristic Function

• characteristic function of $\underbrace{x(\xi)}_{RV}$

$$\underbrace{\phi_x(\xi)}_{\text{F.T}} = E \left\{ e^{j\xi x(\xi)} \right\} = \int_{-\infty}^{\infty} f_x(x) e^{j\xi x} dx$$

Replace ξ by s , we have

moment gen. fn. $\underbrace{\bar{\phi}_x(s)}_{\text{L.T}} = E \left\{ e^{s x(\xi)} \right\} = \int_{-\infty}^{\infty} f_x(x) e^{sx} dx$

$$\underbrace{x_x^{(m)}}_{\text{mth order moment of } x(\xi)} = \left. \frac{d^m}{ds^m} [\bar{\phi}_x(s)] \right|_{s=0} = (-j)^m \left. \frac{d^m}{d\xi^m} [\phi_x(\xi)] \right|_{\xi=0}$$

where $m=1, 2, \dots$

Cumulants of RV

Cumulant generating function

$$\bar{\psi}_x(s) = \ln \bar{\phi}_x(s) = \ln E \{ e^{s x(\frac{1}{2})} \}$$

$\psi_x(\xi)$ = and characteristic function

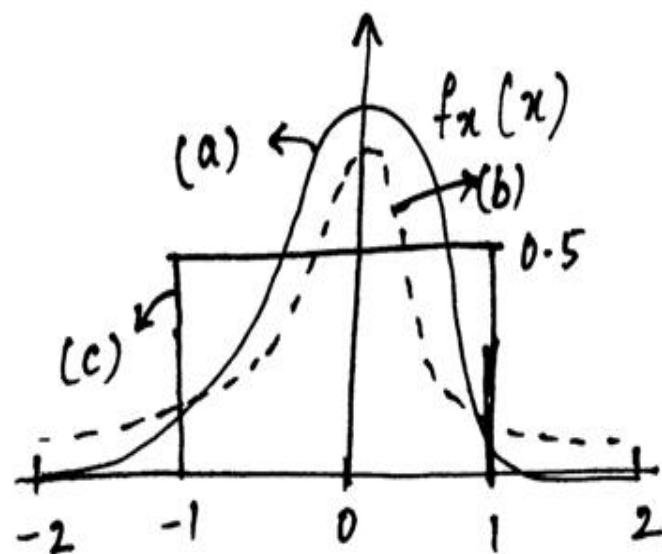
Cumulants of RV $x(\frac{1}{2})$

$$k_x^{(m)} = \left. \frac{d^m}{ds^m} [\bar{\psi}_x(s)] \right|_{s=0} = (-j)^m \left. \frac{d^m}{d\xi^m} [\psi_x(\xi)] \right|_{\xi=0} \quad \text{for } m = 1, 2, 3, \dots$$

$$\underline{k}_x^{(0)} = 0 ; \quad \underline{k}_x^{(1)} = \sigma_1(x) = \mu_x = 0 ; \quad \underline{k}_x^{(2)} = \gamma_x^{(2)} = \sigma_x^2$$

$$\underline{k}_x^{(3)} = \gamma_x^{(3)} ; \quad \underline{k}_x^{(4)} = \gamma_x^{(4)} - 3\sigma_x^4 ; \quad \dots$$

Normal, Cauchy, Uniform RV



$$N(\mu_X, \sigma_X^2)$$

$$(a) f_X(x) = \frac{1}{\sqrt{2\pi\sigma_X^2}} \exp\left[-\frac{1}{2}\left(\frac{x-\mu_X}{\sigma_X}\right)^2\right]$$

$$-\infty < \mu < \infty; \quad \sigma \geq 0$$

$$\phi_X(\xi) = \exp\left(j\mu_X\xi - \frac{1}{2}\sigma_X^2\xi^2\right)$$

$$\gamma_X^{(m)} = E\{[X(\xi) - \mu_X]^m\}$$

$$= \begin{cases} 1 \cdot 3 \cdot 5 \dots (m-1) \sigma_X^m & m \text{ even} \\ 0 & m \text{ odd} \end{cases}$$

But $\gamma_X^{(4)} = \underline{3} \sigma_X^4$

$\therefore \text{kurtosis} = 0$

(b) Cauchy RV and (c) Uniform

Random Vectors

R.V. ✓, now let's see Random Vector (\vec{RV})

$$\vec{RV} = \vec{x}(\omega) = [x_1(\omega), x_2(\omega), \dots, x_M(\omega)]^T$$

\vec{RV} characterized by its 'cdf' as

$$F_x(x_1, \dots, x_M) = \Pr \{ x_1(\omega) \leq x_1, \dots, x_M(\omega) \leq x_M \}$$

'or' $F_x(x) = \Pr \{ x(\omega) \leq x \}$

\vec{RV} characterized by its pdf as

$$\underbrace{f_x(x)}_{\text{pdf}} = \frac{\partial}{\partial x_1} \dots \frac{\partial}{\partial x_M} F_x(x)$$

Individual RV $\therefore f_{x_j}(x_j) = \int \dots \int_{(M-1)} f_x(x) dx_1 \dots dx_j \dots dx_M$

Characterization of Random Vectors

$$F_X(x) = \int_{-\infty}^{x_1} \cdots \int_{-\infty}^{x_M} f_X(v) dv_1 \cdots dv_M = \int_{-\infty}^x f_X(v) dv$$

* If $P_r \{x_1(\xi) \leq x_1, x_2(\xi) \leq x_2\}$
 $= P_r \{x_1(\xi) \leq x_1\} P_r \{x_2(\xi) \leq x_2\}$

$\Rightarrow F_{x_1, x_2}(\cancel{x}) = F_{x_1}(x_1) F_{x_2}(x_2) \neq$ and
 $F_{x_1, x_2}(x_1, x_2)$
 $f_{x_1, x_2}(x_1, x_2) = f_{x_1}(x_1) f_{x_2}(x_2)$

Complex Random Vectors

Complex RV and \vec{RV}
Map space 'S' \rightarrow complex space 'G'

$$\mu = E[x(\xi)] = E[x_R(\xi) + j x_I(\xi)] = \mu_{x_R} + j \mu_{x_I}$$

$$\sigma_x^2 = E\{|x(\xi) - \mu_x|^2\} = E\{|x(\xi)|^2\} - |\mu_x|^2$$

$$\text{complex } \vec{RV} \quad x(\xi) = x_R(\xi) + j x_I(\xi) = \begin{bmatrix} x_{R1} \\ \vdots \\ x_{RM} \end{bmatrix} + j \begin{bmatrix} x_{I1}(\xi) \\ \vdots \\ x_{IM}(\xi) \end{bmatrix}$$

cdf, marginal pdf, etc for complex \vec{RV} is
simple extension to scalar case.

Statistical Description of Random Vectors

• statistical description of \vec{RV}

(a) Mean Vector $\mu_x = E \{x(\underline{z})\} = \begin{bmatrix} E \{x_1(\underline{z})\} \\ E \{x_2(\underline{z})\} \\ \vdots \\ E \{x_M(\underline{z})\} \end{bmatrix} = \begin{bmatrix} \mu_1 \\ \mu_2 \\ \vdots \\ \mu_M \end{bmatrix}$

(b) Second order moments

AutoCorrelation matrix

$$R_x = E \{x(\underline{z}) x^H(\underline{z})\} = \begin{bmatrix} r_{11} & \dots & r_{1M} \\ \vdots & & \vdots \\ r_{M1} & & r_{MM} \end{bmatrix}$$

2nd order moments : $r_{ii} = E \{ |x_i(\underline{z})|^2 \} = r_{x_i}^{(2)}$

$r_{ij} = E \{ x_i(\underline{z}) x_j^*(\underline{z}) \} = r_{ji}^*, \quad i \neq j$
 is the correlation and matrix $R_x = R_x^H$

Co-variance between RV

AutoCovariance Matrix

$$\mathbf{\Gamma}_x = E \{ [x(\frac{1}{2}) - \mu_x] [x(\frac{1}{2}) - \mu_x]^H \} = \begin{bmatrix} \gamma_{11} & \dots & \gamma_{1M} \\ \vdots & \ddots & \vdots \\ \gamma_{M1} & \dots & \gamma_{MM} \end{bmatrix}$$

$$\underbrace{\gamma_{ii}}_{\text{self variance } (\sigma_{x_i}^2)} = E \{ |x_i(\frac{1}{2}) - \mu_i|^2 \} \quad i=1, 2, \dots, M$$

Covariance between $x_i(\frac{1}{2})$ and $x_j(\frac{1}{2})$

$$\begin{aligned} \gamma_{ij} &= E \{ [x_i(\frac{1}{2}) - \mu_i] [x_j(\frac{1}{2}) - \mu_j]^* \} \quad \text{--- (a)} \\ &= E [x_i(\frac{1}{2}) x_j^*(\frac{1}{2})] - \mu_i \mu_j^* = \gamma_{ji}^* \end{aligned}$$

Also correlation Co-eff $\rho_{ij} = \frac{\gamma_{ij}}{\sigma_i \sigma_j} = \rho_{ji}^*$ $\left[\begin{array}{l} \rho_{ij} = 1 ; \text{perfect corr.} \\ \rho_{ij} = 0 ; \text{uncorrelated} \end{array} \right]$

Correlation between RV

$$\Gamma_x = E \{ [x(\frac{1}{2}) - \mu_x] [x(\frac{1}{2}) - \mu_x]^H \}$$

$$\Gamma_x = \underbrace{R_x}_{\text{Auto Cov.}} - \underbrace{\mu_x \mu_x^H}_{\text{Auto Corr.}}$$

* uncorrelated does not imply independent
But vice versa is true

$$E [x_i(\frac{1}{2}) x_j^*(\frac{1}{2})] = E [x_i(\frac{1}{2})] E [x_j^*(\frac{1}{2})]$$

$$\text{From (a) } \underline{r_{ij} = 0}$$

* RV 'Orthogonal' if correlation

$$r_{ij} = E [x_i(\frac{1}{2}) x_j^*(\frac{1}{2})] = 0; i \neq j$$

Linear transformation of \vec{RV}

$$y(\underline{z}) = g[x(\underline{z})] = \underset{[L \times M]}{\underbrace{A}} x(\underline{z})$$

\vec{RV} $y(\underline{z})$ is completely characterized by $f_Y(y)$ the pdf.

Assume ' $L=M$ ' and ' $\underset{\text{Matrix}}{A}$ ' non singular, real valued

$$\therefore f_Y(y) = \frac{f_X(g^{-1}(y))}{|J|} ; J = \begin{bmatrix} \frac{\partial y_1}{\partial x_1} & \dots & \frac{\partial y_M}{\partial x_1} \\ \vdots & \ddots & \vdots \\ \frac{\partial y_1}{\partial x_M} & \dots & \frac{\partial y_M}{\partial x_M} \end{bmatrix} = |\det A|$$

Statistical Description of LTRV

$$\begin{aligned}\therefore f_Y(\mathbf{y}) &= \frac{f_X(\mathbf{A}^{-1}\mathbf{y})}{|\det \mathbf{A}|}; \quad (\text{Real } \vec{RV}) \\ &= \frac{f_X(\mathbf{A}^{-1}\mathbf{y})}{|\det \mathbf{A}|}; \quad (\text{Complex } \vec{RV})\end{aligned}$$

(a) Mean : $\mu_Y = E\{Y(\frac{1}{2})\} = A E\{X(\frac{1}{2})\} = A \mu_X$

(b) AutoCorr : $R_Y = E\{Y Y^H\} = E\{A X X^H A^H\} = A R_X A^H$

(c) AutoCov : $\Gamma_Y = A \Gamma_X A^H$

(d) Cross Correlation : $R_{XY} = E\{X(\frac{1}{2}) X^H(\frac{1}{2}) A^H\} = R_X A^H$
Matrix

Normal \vec{RV} (Real valued)

$$f_{\mathbf{x}}(\mathbf{x}) = \frac{1}{(2\pi)^{M/2} |\mathbf{\Gamma}_{\mathbf{x}}|^{1/2}} \exp \left[-\frac{1}{2} (\mathbf{x} - \mu_{\mathbf{x}})^T \mathbf{\Gamma}_{\mathbf{x}}^{-1} (\mathbf{x} - \mu_{\mathbf{x}}) \right]$$

Normal \vec{RV} (Complex valued)

$$f_{\mathbf{x}}(\mathbf{x}) = \frac{1}{\pi^M |\mathbf{\Gamma}_{\mathbf{x}}|} \exp \left[-(\mathbf{x} - \mu_{\mathbf{x}})^H \mathbf{\Gamma}_{\mathbf{x}}^{-1} (\mathbf{x} - \mu_{\mathbf{x}}) \right]$$

$\mu_{\mathbf{x}}$: Mean ; $\mathbf{\Gamma}_{\mathbf{x}}$: Covariance Matrix

Plugging in mean of \vec{RV} $x(\frac{1}{2}) = \mu_{\mathbf{x}}$ and $\text{var} = \sigma_{\mathbf{x}}^2$
gives scalar case (check!)

Properties of normal distribution of a \vec{RV}

- (a) Pdf is completely specified by mean and Cov. matrix
- (b) If Components of x ($\frac{1}{2}$) are mutually uncorrelated then they are also independent
- (c) Linear transformation of a normal \vec{RV} is also normal (Plug y in place of x)
- (d) Fourth order moment can be expressed in terms of second order moments.

Sums of Independent Random Variables

Sums of Independent RV ($Y(\frac{1}{2})$ as $\{x_k(\frac{1}{2})\}_1^M$)

$$y(\frac{1}{2}) = c_1 x_1(\frac{1}{2}) + c_2 x_2(\frac{1}{2}) + \dots + c_M x_M(\frac{1}{2})$$

$$y(\frac{1}{2}) = \boxed{\text{sum}} \sum_{k=1}^M c_k x_k(\frac{1}{2})$$

Mean:

$$\mu_y = \sum_{k=1}^M c_k \mu_{x_k}$$

Variance:

$$\sigma_y^2 = E \left\{ \left| \sum_{k=1}^M c_k [x_k(\frac{1}{2}) - \mu_{x_k}] \right|^2 \right\}$$

$$\sigma_y^2 = \sum_{k=1}^M |c_k|^2 \sigma_{x_k}^2$$

[using statistical
independence]
of variables

Pdf:

Let $y(\frac{1}{2}) = x_1(\frac{1}{2}) + x_2(\frac{1}{2})$

Characteristic Functions

1st characteristic function: $\phi_y(\xi) = E [e^{j\xi y(\xi)}]$

$$= E [e^{j\xi (x_1(\xi) + x_2(\xi))}] \stackrel{\text{ind.}}{=} E [e^{j\xi x_1(\xi)}] E [e^{j\xi x_2(\xi)}]$$

$\phi_y(\xi) = \phi_{x_1}(\xi) \phi_{x_2}(\xi)$; using convolution property we have

$$f_y(y) = f_{x_1}(y) * f_{x_2}(y)$$

2nd characteristic function: $\psi_y(\xi) = \psi_{x_1}(\xi) + \psi_{x_2}(\xi)$

mth Order Cumulant of $y(\xi)$

$$\kappa_m^{(y)} = \kappa_m^{(x_1)} + \kappa_m^{(x_2)}$$

- If $f_{X_1}(x)$ is "PDF of" a uniform random variable $[x_k(\frac{1}{2})]_{k=1}^4$ and $Y_M(\frac{1}{2}) = \sum_{k=1}^M x_k$; $M=2, 3, 4$.

If we start $f_{Y_2}(y)$ and go on to $f_{Y_4}(y)$ i.e., as 'M' increases pdf gets closer to Gaussian pdf.

- Stable and Infinitely divisible distributions as assignment (reading)
- Stable \rightarrow Means preserved

- stable distn. \rightarrow distn. preserved under convolution (self reproduce)
eg. Gaussian $\overset{c}{p}df$ has finite variance and stable

* Central Limit theorem (CLT)

If $y(\frac{1}{2}) = \sum_{k=1}^M c_k x_k(\frac{1}{2})$; then does the cdf converge as $M \rightarrow \infty$. If each $x(\frac{1}{2})$ is IID-stable then it does. Else?
 and $\mu_{x_k} < \infty, \sigma_{x_k}^2 < \infty$

$$CLT: \quad Y_M(\frac{1}{2}) = \frac{\sum_{k=1}^M x_k(\frac{1}{2}) - \mu_{Y_M}}{\sigma_{Y_M}} \quad \left[\begin{array}{l} \text{Distrn. of} \\ \text{Normalized} \\ \text{Sum} \end{array} \right]$$

converges to that of a "Normal RV" with zero mean and unit SD as $M \rightarrow \infty$

Stochastic Processes

Extend concept of RV and \vec{RV} to "sequences".

Sample space $S = \{\omega_1, \omega_2, \dots\}$ occurring with a probability $Pr\{\omega_k\}$, $k=1, 2, \dots$

Define a Rule that assigns each ω_k to a sequence $x(n, \omega_k)$, $-\infty < n < \infty$

$\{S, Pr, x(n, \omega_k)\}$ constitute a stochastic process DT

Defn: $x(n, \omega_k)$, $-\infty < n < \infty$, is a random sequence if for a fixed n_0 , $x(n_0, \omega_k)$ is a RV

Summary

Set of all $\{x(n, \zeta)\}$ is called ensemble

Each $x(n, \zeta_k)$ is called realization of the sample sequence.

		n	ζ
$x(n, \zeta)$	RV	fixed	variable
	SS	variable	fixed
	N	fixed	fixed
	SP	variable	variable

RV: Rand. variable

SS: Sample Seq.

N: Number

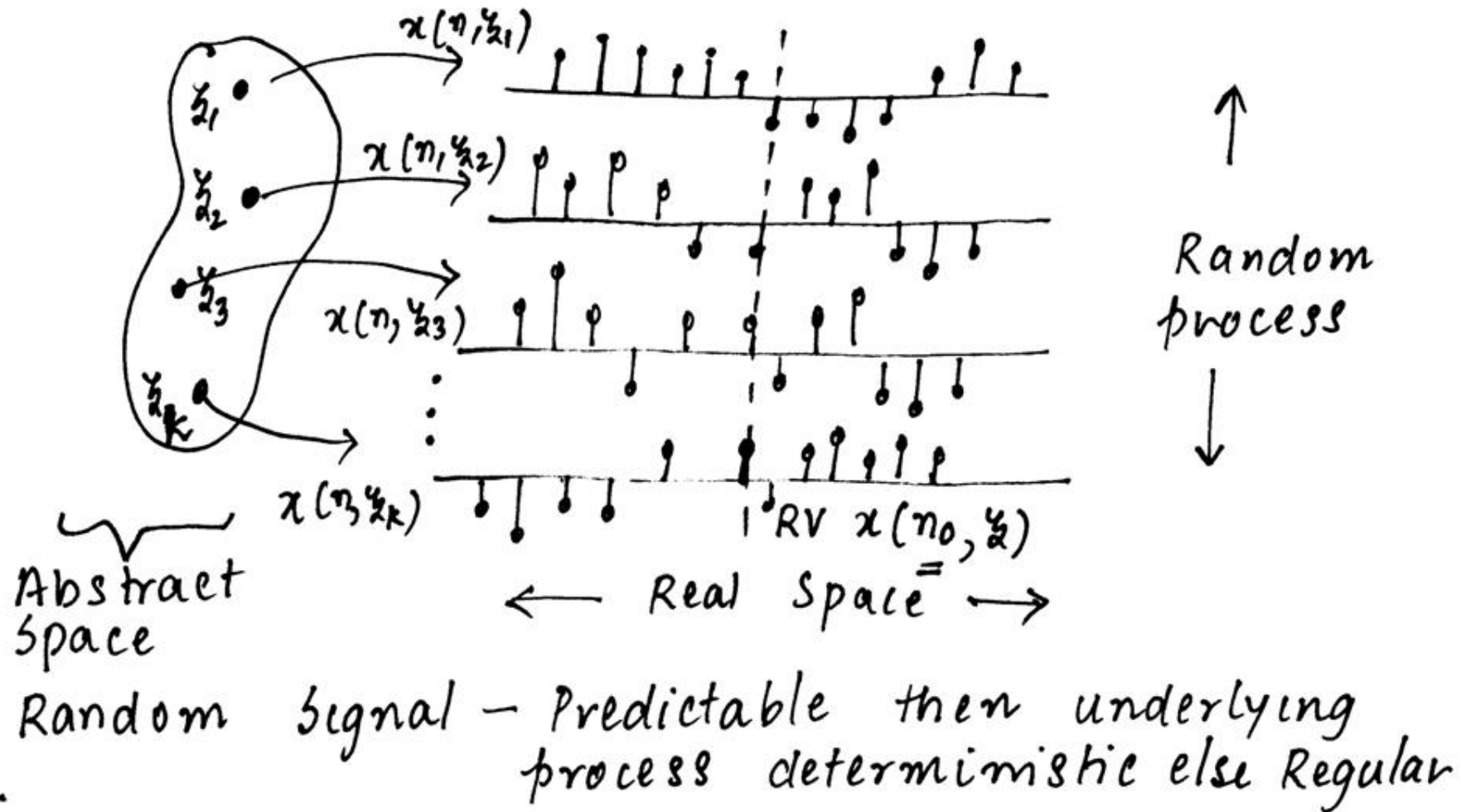
SP: Stoch. Process

Random Sequence is also called time series

Random/Stochastic process are same

Random Sequences

Realization of a Random Process = Random Sequence



Description of Random Processes

- Random signal characterized by
 - 1) Pdf
 - 2) Correlation
 - 3) Signal power
 - 4) Histograms

Description with Cdf (kth order)

$$F_x(x_1, \dots, x_k; n_1, \dots, n_k) = \Pr \{ x(n_1) \leq x_1, \dots, x(n_k) \leq x_k \}$$

with Pdf

$$f_x(x_1, \dots, x_k; n_1, \dots, n_k) = \frac{\partial^{2k} F_x(x_1, \dots, x_k; n_1, \dots, n_k)}{\partial x_{n_1} \dots \partial x_{n_k}}$$

However random processes described by 'averages'.

Description of Random Processes

- Statistical Description of random process
Notation: use $x(n)$ to describe $x(n, \omega)$ and also single realization $x(n)$

Second order statistical description

$$\mu_x(n) = E[x(n)] = E\{x_R(n) + j x_I(n)\}$$

$$\sigma_x^2(n) = E[|x(n) - \mu_x(n)|^2] = E[|x(n)|^2] - |\mu_x(n)|^2$$

Autocorrelation sequence $r_{xx}(n)$ described as

$$r_{xx}(n_1, n_2) = E[x(n_1) x^*(n_2)]$$

Description of Random Processes

- Auto Covariance of $x(n)$

$$\gamma_{xx}(n_1, n_2) = E \left[(x(n_1) - \mu_x(n_1)) (x(n_2) - \mu_x(n_2))^* \right]$$

$$\gamma_x(n_1, n_2) = r_{xx}(n_1, n_2) - \mu_x(n_1) \mu_x^*(n_2)$$

$$\text{Cross correlation: } r_{xy}(n_1, n_2) = E [x(n_1) y^*(n_2)]$$

$$\text{Cross covariance: } \gamma_{xy}(n_1, n_2) = r_{xy}(n_1, n_2) - \mu_x(n_1) \mu_y^*(n_2)$$

$$\text{Normalized Cross correlation: } \rho_{xy}(n_1, n_2) = \frac{\gamma_{xy}(n_1, n_2)}{\sigma_x(n_1) \sigma_y(n_2)}$$

Properties of Stochastic processes

Independent process

$$a) f_x(x_1, \dots, x_k; n_1, \dots, n_k) = f_1(x_1; n_1) \dots f_k(x_k; n_k) \\ \forall k, n_i, i = 1, \dots, k$$

b) If all random variables have same pdf $f(x)$, $\forall k$ then $x(n)$ is I.I.D

$$c) \text{Uncorrelated process: } \gamma_x(n_1, n_2) = \begin{cases} \sigma_x^2(n_2); & n_1 = n_2 \\ 0; & n_1 \neq n_2 \end{cases}$$

$$d) \text{Orthogonal Process: } \gamma_x(n_1, n_2) = \begin{cases} \sigma_x^2(n_1) + |\mu_x(n_1)|^2; & n_1 = n_2 \\ 0; & n_1 \neq n_2 \end{cases}$$

Properties of Stochastic Processes

- Wide sense periodic : ^{if} $\mu_x(n) = \mu_x(n+N), \forall n$
and $r_x(n_1, n_2) = r_x(n_1+N, n_2) = r_x(n_1, n_2+N)$
 $= r_x(n_1+N, n_2+N)$

joint stochastic Processes (Properties)

- $f_{xy}(x, y; n_1, n_2) = f_x(x; n_1) f_y(y; n_2) \Rightarrow$ Independent
- $r_{xy}(n_1, n_2) = 0 ; \forall n_1 \neq n_2 \Rightarrow$ Uncorrelated
- $r_{xy}(n_1, n_2) = 0 ; \forall n_1 \neq n_2 \Rightarrow$ Orthogonal

Stationarity of Random process

Statistical description of $x(n)$ = S.D. of $x(n+k)$
 , $\forall k$

a) Order 'N' stationary

$$f_x(x_1, \dots, x_N; n_1, \dots, n_N) = f_x(x_1, \dots, x_N; n_{1+k}, \dots, n_{N+k})$$

If $x(n)$ is stationary for all orders $N=1, 2, \dots$ then
"Strict sense stationary"

b) WSS (Wide Sense Stationary)

stationary upto order $N=2$.

• \nrightarrow

Stationarity (Contd.)

- A random signal WSS if
 - i) Mean is a constant, independent of 'n'
and Var $E\{x(n)\} = \mu_x$.
 $Var[x(n)] = \sigma_x^2$.
 - ii) AutoCorrelation depends on the lag
 $\lambda = n_1 - n_2$
$$\begin{aligned} r_x(n_1, n_2) &= r_x(n_1 - n_2) = r_x(\lambda) = E\{x(n+\lambda)x^*(n)\} \\ &= E\{x(n)x^*(n-\lambda)\} \end{aligned}$$
- Alternately AutoCovariance $\gamma_x(\lambda) = r_x(\lambda) - |\mu_x|^2$.

Stationarity (Contd.)

Weiner Process : $x(n)$ is running sum of independent steps or increments

Jointly WSS ($x(n)$ and $y(n)$) if

$$r_{xy}(l) = E \{ x(n) y^*(n-l) \} = r_{xy}(l) - \mu_x \mu_y^*$$

Properties of Autocorrelation sequences

$$\left. \begin{aligned} (i) \quad r_x(0) &= \sigma_x^2 + |\mu_x|^2 \geq 0 \\ r_x(0) &\geq |r_x(l)|, \quad \forall l \end{aligned} \right] \begin{aligned} |\mu_x|^2 &= \text{Av. DC Power} \\ \sigma_x^2 &= r_x(0); \text{ Av. AC Power} \\ r_x(0) &= \text{Total Av. power.} \end{aligned}$$

$$ii) \quad r_x(-l) = r_x^*(l) : \text{conjugate symmetric about lag.}$$

Ergodicity

Asymptotic Stationary : If statistics of $x(n)$ and $x(n+k)$ become stationary as $k \rightarrow \infty$

Ergodicity: Ergodic means all statistical info can be obtained from any single representative member of the ensemble

So we now replace operations or Expectation over an ensemble to a single realization called Time Average = $\lim_{N \rightarrow \infty} \frac{1}{2N+1} \sum_{n=-N}^N (\cdot)$

Ergodicity

Time Averages on single realizations of a random process

$$\text{Mean} = \langle x(n) \rangle ; \text{Mean Square } \langle |x(n)|^2 \rangle$$

$$\text{Variance} = \langle |x(n) - \langle x(n) \rangle|^2 \rangle$$

$$\text{AutoCorrelation} = \langle x(n) x^*(n-l) \rangle$$

$$\text{AutoCovariance} = \langle [x(n) - \langle x(n) \rangle] [x(n-l) - \langle x(n) \rangle]^* \rangle$$

Ergodic Random Process : Random process with a single sufficient realization

$x(n)$ is Ergodic if its "Ensemble Average = Time Av."

Ergodicity

* Ergodic in mean: if $\langle x(n) \rangle = E \{ x(n) \}$

* Ergodic in correlation: if $\langle x(n) x^*(n-l) \rangle = E \{ x(n) x^*(n-l) \}$

* Joint Ergodicity: Two random signals are jointly ergodic if i) they are individually ergodic

$$\text{ii) } \langle x(n) y^*(n-l) \rangle = E \{ x(n) y^*(n-l) \}$$

True Estimate in practice of Time Average

$$= \frac{1}{2N+1} \sum_{n=-N}^N (\cdot)$$

ERGODICITY: \rightarrow
" ONE REALIZATION OF THE RANDOM SIGNAL $x(n)$ AS $n \rightarrow \infty$ TAKES ON VALUES WITH THE SAME STATISTIC AS $x(n_1)$ AT $n = n_1$ "

- Discuss Random signal variability with Figure 3.8 in Kogon.

Frequency domain description of Stationary Processes

PSD: Fourier transform of its Autocorrelation sequence $r_x(l)$

Zero Mean
Non periodic

$$R_x(e^{j\omega}) = \sum_{l=-\infty}^{\infty} r_x(l) e^{-j\omega l} \quad [\omega: \text{freq. in radians per sample}]$$

Non Zero/Zero Mean Periodic

$$R_x(e^{j\omega}) = \sum_i 2\pi A_i \delta(\omega - \omega_i)$$

A_i : Amplitudes at frequencies at ω_i

and

$$r_x(l) = 1/2\pi \int_{-\pi}^{\pi} R_x(e^{j\omega}) e^{j\omega l} d\omega$$

Properties of PSD

- (i) $R_x(e^{j\omega})$ is a real valued periodic function
- (ii) If $x(n)$ is real, then $R_x(e^{j\omega}) = R_x(e^{-j\omega})$
- (iii) $R_x(e^{j\omega}) \geq 0$; Non negative
- (iv) Area under $R_x(e^{j\omega}) = \text{Av. Power of } x(n)$
$$\frac{1}{2\pi} \int_{-\pi}^{\pi} R_x(e^{j\omega}) d\omega = r_x(0) = E\{|x(n)|^2\} \geq 0$$

Other Random Sequences

White Noise

$$w(n) \sim WN(\mu_w, \sigma_w^2)$$

'iff' $E\{w(n)\} = \mu_w$ and

$$r_w(l) = E\{w(n) w^*(n-l)\} = \sigma_w^2 \delta(l)$$

$$\text{or } R_w(e^{j\omega}) = \sigma_w^2; \quad -\pi \leq \omega \leq \pi$$

Strict White Noise

$$w(n) \sim IID(\mu_w, \sigma_w^2)$$

Harmonic Process

$$x(n) = \sum_{k=1}^M A_k \cos(\omega_k n + \phi_k); \quad \omega_k \neq 0$$

. $A_k, \{\omega_k\}_1^M$ are constants, $\{\phi_k\}_1^M$ is a RV in $[0, 2\pi]$

Harmonic Process (contd.)

$$* \quad E \{ x(n) \} = 0 ; \quad \overset{* \text{ ACF}}{\gamma_x(l)} = \frac{1}{2} \sum_{k=1}^N A_k^2 \cos \omega_k l ;$$

$$\text{PSD: } R_x(e^{j\omega}) = \sum_{k=-M}^M 2\pi \left(\frac{A_k^2}{4} \right) \delta(\omega - \omega_k) = \sum_{k=-M}^M \frac{\pi}{2} A_k^2 \delta(\omega - \omega_k)$$

$-\infty < l < \infty$

Generalized Harmonic Process

$$x(n) = \sum_{k=1}^M A_k e^{j(\omega_k n + \phi_k)}$$

and

$$\gamma_x(l) = \sum_{k=1}^M |A_k|^2 e^{j\omega_k l}$$

Power Spectrum
= M impulses
at ω_k with amp.
 $2\pi |A_k|^2$.

- Cross PSD (ZM and jointly stationary stoch. processes)

$$R_{xy}(e^{j\omega}) = \sum_{\tau=-\infty}^{\infty} r_{xy}(\tau) e^{-j\omega\tau}$$

$$\text{and } r_{xy}(\tau) = \frac{1}{2\pi} \int_{-\pi}^{\pi} R_{xy}(e^{j\omega}) e^{j\omega\tau} d\omega$$

$$\text{Since } r_{xy}(\tau) = r_{yx}^*(-\tau); R_{xy}(e^{j\omega}) = R_{yx}^*(e^{j\omega})$$

Normalized Cross Spectrum (Coherence)

$$\gamma_{xy}(e^{j\omega}) = \frac{R_{xy}(e^{j\omega})}{\sqrt{R_x(e^{j\omega})} \sqrt{R_y(e^{j\omega})}}$$

- Cross Spectral density functions (complex)

$$R_x(z) = \sum_{\forall l} r_x(l) z^{-l} \quad ; \text{ complex PSD}$$

$$R_{xy}(z) = \sum_{\forall l} r_{xy}(l) z^{-l} \quad ; \text{ complex CPSD}$$

assuming $z = e^{j\omega}$, is within the ROC
and $r_x(l)$ and $r_{xy}(l)$ are
absolutely summable

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