Project Assignment #2

FEM1230 Estimation Theory

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Outline

- Data Model
- ► Maximum Likelihood Estimation
- Mean Square Error Bounds With Affine Bias
- ▶ Mean Square Error Of SNR Affine Biased Estimator

Data Model

Assume a data model

$$x[n] = A + w[n]$$
 $n = 1, 2, ..., N - 1$ (1)

where A is the unknown DC level and $w[n] \sim \mathcal{N}(0, \sigma^2)$ is additive noise with unknown variance σ^2 ,

$$\theta = \left[A \, \sigma^2 \right] \tag{2}$$

▶ The SNR $\alpha = \frac{A^2}{\sigma^2}$ is lower bounded for any unbiased estimator $\hat{\alpha}$ by CRLB,

$$var(\alpha) \ge \frac{4\alpha + 2\alpha^2}{N}.$$
 (3)

Problem: Signal-to-noise ratio (SNR) estimation of the data model with unknown A and σ^2 .

Maximum Likelihood Estimation (MLE)

Using the probability density function of the given data model

$$\rho(\mathbf{x};\theta) = \frac{1}{(2\pi\sigma^2)^{\frac{N}{2}}} \exp\left[\frac{-1}{2\sigma^2} \sum_{n=0}^{N-1} (x[n] - A)^2\right],\tag{4}$$

where $\mathbf{x} = [x[0] \dots x[N-1]]^T$ and by using the invariance property of the MLE, the MLE of SNR α is given by

$$\hat{\alpha}_{ML} = \widehat{\left(\frac{A^2}{\sigma^2}\right)}_{ML} = \frac{\hat{A}_{ML}^2}{\hat{\sigma}_{ML}^2} = \frac{\left(\frac{1}{N}\sum_{n=0}^{N-1}x[n]\right)^2}{\frac{1}{N}\sum_{n=0}^{N-1}\left(x[n] - \frac{1}{N}\sum_{n=0}^{N-1}x[n]\right)^2}.$$
 (5)

Maximum Likelihood Estimation (MLE)

- ► The PDF of \hat{A}_{ML}^2 is non-central chi-squared $\mathcal{X'}_{\kappa_1}^2$ with degree of freedom $\kappa_1 = 1$ and non-centrality parameter $\lambda = N\alpha$ because $\hat{A}_{ML} \sim \mathcal{N}(A, \sigma^2/N)$.
- ▶ The PDF of $\hat{\sigma}^2$ is central chi-squared $\mathcal{X}^2_{\kappa_2}$ with degree of freedom $\kappa_2 = (N-1)$.
- ▶ Therefore the PDF of $\hat{\alpha}_{ML}$ is a non-central chi-squared,

$$f(\hat{\alpha}_{ML}) = \begin{cases} \frac{e^{-\frac{N\alpha}{2}}}{N-1} \sum_{i=0}^{\infty} \frac{(N\alpha/2)^{i} \Gamma\left(i + \frac{N}{2}\right)}{i! \Gamma\left(\frac{1+2i}{2}\right) \Gamma\left(\frac{N-1}{2}\right)} \frac{\hat{\alpha}^{(i-\frac{1}{2})}}{[1+\hat{\alpha}]^{(i+\frac{N}{2})}}, \quad \alpha \ge 0 \\ 0 \quad \alpha < 0 \end{cases}$$

with the expected value

$$E\{\hat{\alpha}_{ML}\} = \frac{N}{N-3} \left(\alpha + \frac{1}{N}\right) \quad N > 3$$
 (6)

and variance

$$var\{\hat{\alpha}_{ML}\} = \frac{\left[2\alpha^2 + 4\alpha\left(1 - \frac{2}{N}\right) + \frac{2}{N}\left(1 - \frac{2}{N}\right)\right]}{N(1 - \frac{3}{N})^2(1 - \frac{5}{N})} \quad N > 5$$
 (7)

Maximum Likelihood Estimation (MLE)

▶ For $N \to \infty$, The ML estimator $\hat{\alpha}_{\textit{ML}}$ is asymptotically efficient,

$$var{\{\hat{\alpha}_{ML}\}} \stackrel{N \to \infty}{\longrightarrow} CRLB$$
 (8)

and the PDF of $\hat{\alpha}_{ML}$ is obey a Gaussian distribution for $N \to \infty$ from the central limit theorem.

▶ The following figure compare the ML MSE to the CRLB as a function of α for finite N.

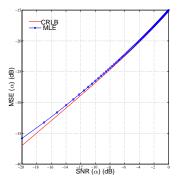


Figure: The CRLB and the ML MSE as the function of α for N = 200.

Mean Square Error Bounds (MSEB) With Affine Bias

▶ CRLB on the MSE of an unbiased estimator $\hat{\theta}$ of θ is

$$MSEB(0, 0, \theta) = E\{\|\hat{\theta} - \theta\|^2\} \ge Tr(J^{-1}(\theta)),$$
 (9)

where $J(\theta)$ is the Fisher information matrix.

▶ Affine MSE bound on the MSE of an estimator with bias $E\{\hat{\theta}\} - \theta = M\theta + u$ is

$$MSEB(M, u, \theta) = ||E\{\hat{\theta}\} - \theta||^2 + Tr((I + D(\theta))J^{-1}(\theta)(I + D(\theta))^H)$$

= $(M\theta + u)^H(M\theta + u) + Tr(I + M)J^{-1}(\theta)(I + M)^H,$ (10)

for some matrix M and $u \in \mathcal{U}$, where $D(\theta) = \frac{\partial (\mathcal{E}\{\hat{\theta}\} - \theta)}{\partial \theta}$.

Mean Square Error Bounds (MSEB) With Affine Bias

Theorem 1 [Y. C. Eldar (2008)]

For a random vector \mathbf{x} with PDF $p(\mathbf{x}, \theta)$, \hat{M} and \hat{u} are unique and admissible solution of

$$(\hat{M}, \hat{u}) = \arg\min_{M, u} \sup_{\theta \in \mathcal{U}} \{ \textit{MSEB}(M, u, \theta) - \textit{MSEB}(0, 0, \theta) \}. \tag{11}$$

and for $\hat{M} \neq 0$ or $\hat{u} \neq 0$, we have

$$MSEB(\hat{M}, \hat{u}, \theta) < MSEB(0, 0, \theta), \quad \forall \ \theta \in \mathcal{U},$$
 (12)

with affine bias estimator $\hat{\theta}_b$ which achieves the affine MSE bound and have smaller MSE thanc $\hat{\theta}$ \forall $\theta \in \mathcal{U}$,

$$\hat{\theta}_b = (1+M)\hat{\theta} + u. \tag{13}$$

Mean Square Error Bounds (MSEB) With Affine Bias

Proposition 1 [Y. C. Eldar (2008)]

Let a random vector \mathbf{x} with pdf $p(\mathbf{x}, \theta)$ has the Fisher information of form

$$J^{-1}(\theta) = b^2 \theta_0^2 + 2c\theta_0 + a. \tag{14}$$

Then the minimax M and u that are the solution to

$$(\hat{M}, \hat{u}) = \arg \min_{M, u} \sup_{\alpha \in \mathcal{U}} \{MSEB(M, u, \theta) - MSEB(0, 0, \theta)\}$$
 (15)

with $\mathcal{U} = \mathbb{R}$ are given by

$$M = \begin{cases} -\frac{2b^2}{1+b^2}, & |b| < 1\\ -1, & |b| \ge 1, \end{cases} \quad u = \begin{cases} -\frac{2c}{1+b^2}, & |b| < 1\\ -\frac{c}{b^2}, & |b| \ge 1. \end{cases}$$
 (16)

Furthermore, if there exists an efficient estimator $\hat{\theta}$, then

$$\hat{\theta}_b = \begin{cases} \frac{1 - b^2}{1 + b^2} - \frac{2c}{1 + b^2}, & |b| < 1\\ -\frac{c}{b^2}, & |b| \ge 1, \end{cases}$$
 (17)

has smaller MSE than $\hat{\theta} \forall \theta$.

- ▶ The $\hat{\alpha}_{MI}$ does not achieve the CRLB.
- \blacktriangleright To achieve lower bound than CRLB, one need to find the scalar \hat{M} and \hat{u} that are the solution to

$$(\hat{M}, \hat{u}) = \arg \min_{M, u} \sup_{\alpha \ge 0} \{ (M\alpha + u)^2 + ((1+M)^2 - 1)J^{-1}(\alpha) \}$$
 (18)

• Affine biased estimator of the SNR α is

$$\hat{\alpha}_b = (1+M)\hat{\alpha}_{ML} + u,\tag{19}$$

The optimal M and u, that lead to lower MSE than CRLB, can be found using the Semi-definite programming formulation.

Since, for finite N, we have

$$var\{\hat{\alpha}_{ML}\} = b^2\alpha^2 + 2c\alpha + a \ge J^{-1}(\alpha)$$
 (20)

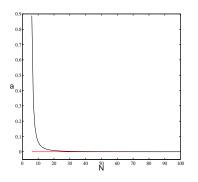
where

$$a = \frac{\frac{2}{N}(1 - \frac{2}{N})}{N(1 - \frac{3}{N})^2(1 - \frac{5}{N})}$$

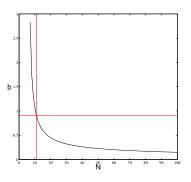
$$b = \sqrt{\frac{2}{N(1 - \frac{3}{N})^2(1 - \frac{5}{N})}}$$
(21)

$$c = \frac{2(1 - \frac{2}{N})}{N(1 - \frac{3}{N})^2(1 - \frac{5}{N})}$$
 (22)

with a > 0 for N > 5.



(a) The parameter a as the function of N.



(b) The parameter b as the function of N.

As $\hat{\alpha}_{ML}$ is an efficient estimator of α , the affine biased estimator of α is, given by Proposition 1,

$$\hat{\alpha}_{b} = (1 + M)\hat{\alpha}_{ML} + u$$

$$= \begin{cases} \frac{(N(1 - \frac{5}{N})^{2}(1 - \frac{5}{N}) - 2)\hat{\alpha}_{ML} - 4(1 - \frac{2}{N})}{N(1 - \frac{3}{N})^{2}(1 - \frac{5}{N}) + 2}, & |b| < 1(i.e., N > 10) \\ -(1 - \frac{2}{N}), & |b| \ge 1, \end{cases}$$
(23)

and

$$MSEB(M, u, \hat{\alpha}_b) < MSE(0, 0, \hat{\alpha}_{ML})$$
 (24)

for a > 0.

Conclusion

- ML estimator is asymptotically efficient for large N.
- Affine biased estimator analytically outperform the ML estimator in terms of MSE for finite N.
- ▶ The suggested affine biased estimator is realizable for N > 5.
- One can obtain an MSE that is lower than the (unbiased) CRLB over some range of α and N, as described by Y. C. Eldar.

References



S. M. Kay, Fundamentals of Statistical Signal Processing: Estimation Theory, Prentice Hall, Upper Saddle River, New Jersey 07458, 1993.



Y. C. Eldar, "Uniformly Improving the Cramér-Rao Bound and Maximum-Likelihood Estimation," IEEE Trans. Signal Process., vol. 54, pp. 2943-2956, Aug. 2006.



Y. C. Eldar, "MSE Bounds With Affine Bias Dominating the Cramér-Rao Bound," IEEE Trans. Signal Process., Vol. 56, pp. 3824-3836, Aug. 2008.