Time series in epidemiology

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Objectives of this module

- ► To understand the limitations of standard linear regressions models for time series data.
- ► To model temporal trends (both seasonal and non-seaonal) through the use of explanatory variables.
- ▶ To understand and apply basic models for time series analysis.

1. Review of linear regression

Linear regression model

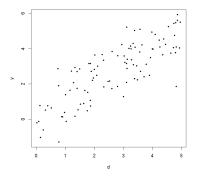
$$Y_i = \beta_0 + \beta_1 x_{1i} + ... + \beta_k x_{ki} + Z_i : i = 1, ..., n$$

- ightharpoonup explanatory variables/factors $x_{1i},...,x_{ki}$
- ▶ regression parameters $\beta_0, ..., \beta_k$
- ightharpoonup residuals Z_i
- ► responses *Y*_i

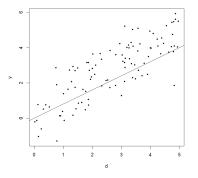
```
fit1<-lm(y x); summary(fit1)
xsq<-x*x; fit2<-lm(y x+xsq); summary(fit2)
names(fit2)
xx<-0.1*(0:200); beta<-fit2$coef;
ff<-beta[1]+beta[2]*xx+beta[3]*xx*xx
par(mfrow=c(1,1)); plot(x,y,pch=19,xlim=c(0,20));
lines(xx,ff,col="blue",lwd=3)</pre>
```

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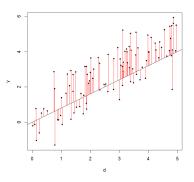
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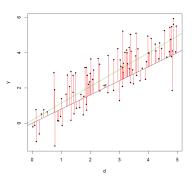


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$$RSS = \sum_{i=1}^{n} (y_i - \beta_1 - \beta_2 x_i)^2$$

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General linear model: things to remember

ightharpoonup Transformation of x and/or Y widens the scope of the model

Examples:

- $Y_i = \beta_1 + \beta_2 \log x_i + Z_i$
- Normality of the Z_i is less important than independence and constant variance
- ▶ But if the Z_i are iid $N(0, \sigma^2)$, likelihood-based inference is straightforward (and least squares estimates are maximum likelihood estimates)
- Diagnostic checking of residuals is an important part of model-building, but requires subjective judgement.

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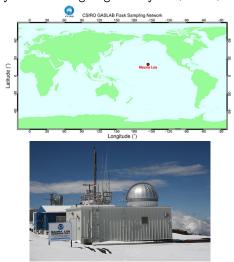
- $ightharpoonup Y^{\top} = (Y_1, \ldots, Y_t)$
- Our model for nature,

$$[Y] = [Y_1] \times [Y_2|Y_1] \times \ldots \times [Y_t|Y_{t-1}, Y_{t-2}, \ldots, Y_1].$$

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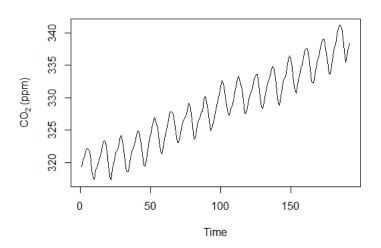
Example: CO₂ time series

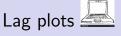
http://yearsoflivingdangerously.com/watch/season-1/

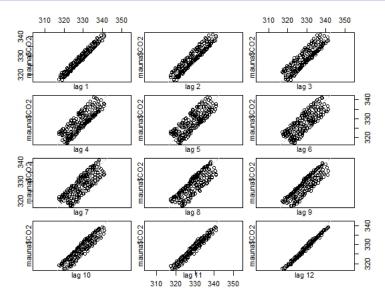


Example: CO_2 time series









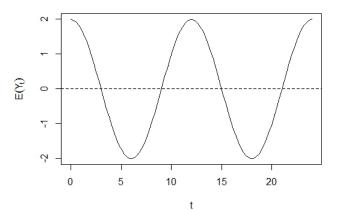
$$E[Y_t] = \beta_1 + A\sin\{2\pi ft + \phi\}$$

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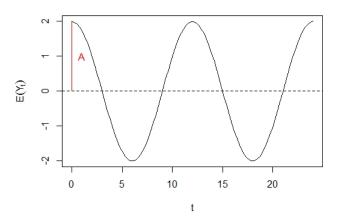
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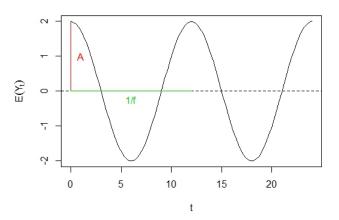
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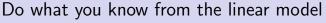
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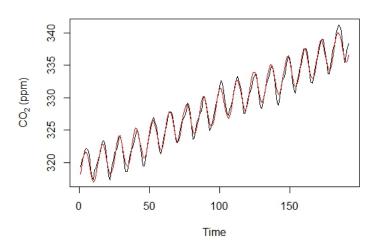
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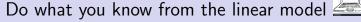
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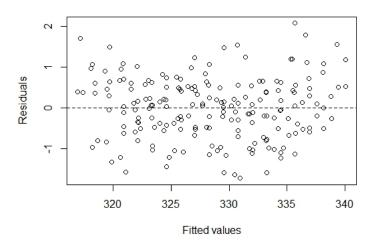






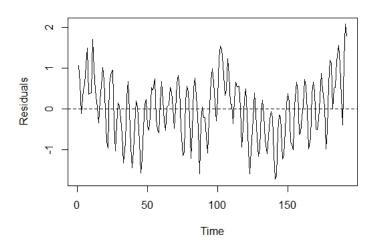


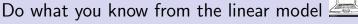




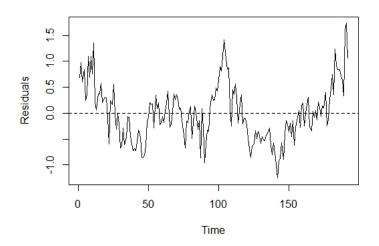
Do what you know from the linear model

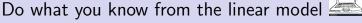






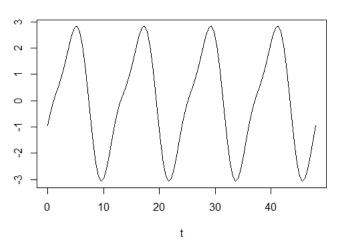








Seasonal trend



Stationarity

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► First order stationarity

$$E[Y_t] = \mu$$
, for all t .

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Second order stationarity

$$V[Y_t] = \sigma^2$$
, for all t

and

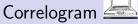
$$COV(Y_t, Y_{t-k}) = \gamma_k = \sigma^2 \rho_k.$$

▶ What value does γ_0 take?

Correlogram



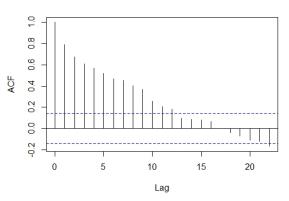
$$\hat{\gamma}_k = \frac{1}{n} \sum_{t=1}^{n-k} (y_t - \bar{y})(y_{t-k} - \bar{y})$$





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Correlogram of the residuals



Autoregressive models

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- $Y_{t} = \sum_{k=1}^{p} \phi_{k} B^{k} Y_{t} + Z_{t} \iff (1 \phi_{1} B \phi_{2} B^{2} \dots B^{p} \phi_{p}) Y_{t} = Z_{t}$

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Stationarity

An AR(p) process is stationary if and only if the roots of the equation

$$1 - \phi_1 x - \phi_2 x^2 - \dots - \phi_p x^p = 0$$

are in absolute value smaller than 1.

$$Y_t = \phi Y_{t-1} + Z_t$$

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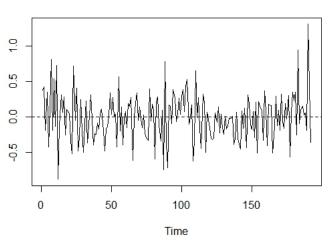
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$$\log L(\theta) = -\frac{1}{2} \left[n \log \{\sigma^2\} - \log(1 - \phi^2) + (1 - \phi^2) \frac{y_1^2}{\sigma^2} + \sum_{t=2}^n \frac{(y_t - \phi y_{t-1})^2}{\sigma^2} \right]$$

CO₂ time series (continue)



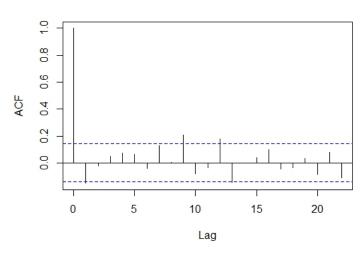
AR(1) Residuals



CO₂ time series (continue)



Series AR(1) residuals



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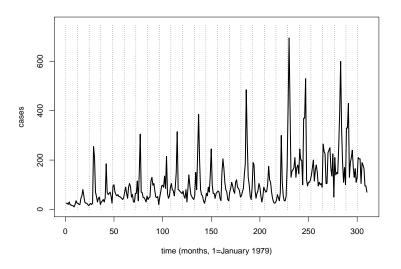
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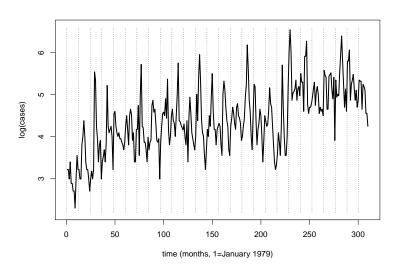
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- A model for the data

$$Y(t_i) = d(t_i)^{\top} \beta + W(t_i) + Z(t_i).$$

Malaria cases in Kericho, Kenya



Malaria cases in Kericho, Kenya

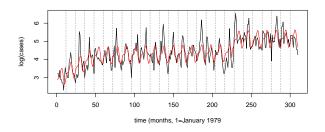


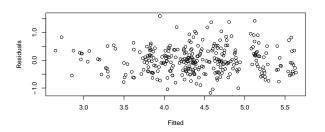
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- Let Y(t) denote the log-transformed number of cases.
- A model for the data:

$$Y(t) = \beta_0 + \beta_1 t + \beta_2 \max\{t - 50, 0\} + \beta_3 I(t > 225) + \beta_2 \sin(2\pi t/12) + \beta_3 \cos(2\pi t/12) + \beta_4 \sin(2\pi t/6) + \beta_5 \cos(2\pi t/6) + Z(t)$$
(1)





Assumption: W(t) is a stationary zero-mean stochatic process with covariance function

$$Cov\{W(t), W(t')\} = \sigma^2 \rho(h), h = |t - t'|$$

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- Assumption: Z(t) are i.i.d. variables with mean 0 and variance τ^2
- Theoretical varioram (definition)

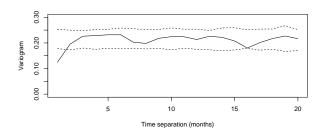
$$v(h) = \frac{1}{2}E[(W(t) + Z(t) - W(t') - Z(t'))^{2}]$$

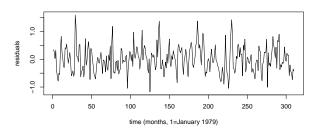
= $\tau^{2} + \sigma^{2}(1 - \rho(h))$

Let Z(t) denote the residuals from a standard linear regression model.

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- ► The empirical variogram (definition)

$$\hat{v}(h) = \frac{1}{2}(\hat{Z}(t) - \hat{Z}(t'))^2, h = |t - t'|.$$
 (2)





► Stationary covariance function

$$\operatorname{cov}\{W(t), W(t')\} - \sigma^2 \rho(h), \tag{3}$$

Stationary covariance function

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Matern covariance functions

$$\rho(h) = \{2^{\kappa - 1} \Gamma(\kappa)\}^{-1} (h/\phi)^{\kappa} \mathcal{K}_{k}(h/\phi), h > 0$$
(4)

Stationary covariance function

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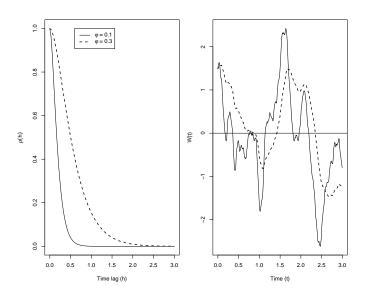
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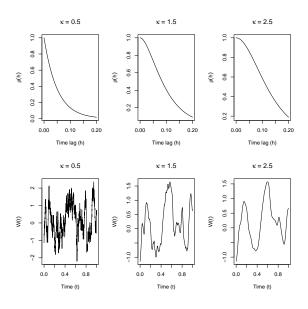
▶ Special case $(\kappa = 1/2)$

$$\rho(h) = \exp\{-h/\phi\}. \tag{5}$$

The scale parameter ϕ



The smoothness parameter κ



• Unkown parameters to estimate: $\theta = (\beta, \sigma^2, \phi, \tau^2)$.

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- The vector $(Y(t_1), \ldots, Y(t_n))$ follows a multivariate Gaussian distribution with mean $D\beta$ and covariance matrix $\Omega = \sigma^2 R + \tau^2 I$ where

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▶ The likelihood function

$$I(\theta) = -\frac{1}{2} \left\{ n \log \sigma^2 + \log |\Omega| + (y - D\beta)^{\mathsf{T}} \Omega^{-1} (y - D\beta) / \sigma^2 \right\},\tag{6}$$

Kericho data: parameter estimation

	$\kappa = 0.5$	
Parameter	Point estimate	95% CI
β_0	2.892	(2.535, 3.248)
eta_1	0.137	(0.038, 0.235)
eta_2	-0.345	(-0.444, -0.247)
eta_3	0.162	(0.086, 0.238)
$eta_{ extsf{4}}$	0.126	(0.050, 0.202)
eta_{5}	0.026	(0.018, 0.035)
eta_{6}	-0.025	(-0.035, -0.016)
eta_{7}	0.700	(0.397, 1.004)
σ^2	0.214	(0.177, 0.258)
ϕ	1.149	(0.870, 1.516)
$ au^2$	-	