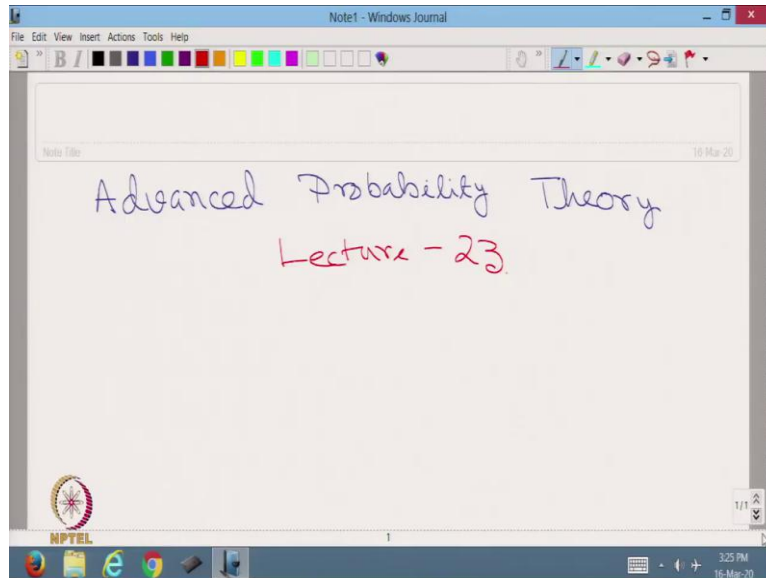


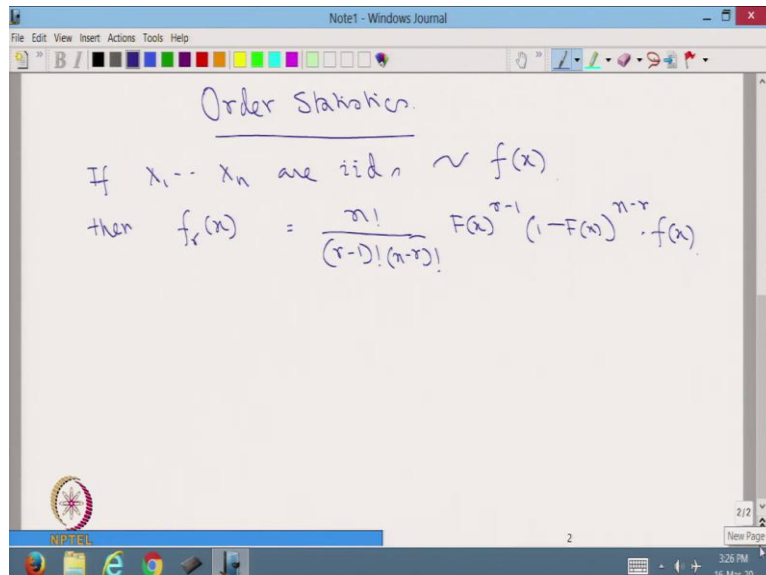
Advanced Probability Theory
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Lecture 23

(Refer Slide Time: 0:24)



Welcome students to MOOCS series of lectures on Advanced Probability Theory. This is lecture number 23.

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If you remember, we are working on Order Statistics. And we have derived that if X_1, X_2, \dots, X_n are iids following parent distribution $f(x)$, then $f_r(x)$ is equal to factorial n upon factorial r

minus 1 into n minus r factorial, $F x$ to the power r minus 1, $1 - F x$ to the power n minus r multiplied by $f x$. Before we proceed any further, let me solve some examples.

(Refer Slide Time: 1:28)

Ex Suppose X_1, X_2, \dots, X_n are iid $U(0,1)$. What is the expectation of (a) $X_{(1)}$, (b) $X_{(n)}$.

Now we need to first compute the pdf of $X_{(1)}$.

Diagram illustrating the distribution of $X_{(1)}$ on the interval $[0, 1]$. A point x is marked on the interval. The segment from 0 to x is labeled 0, and the segment from x to 1 is labeled $n-1$.

$$\therefore f_1(x) = n C_1 f(x) (1 - F(x))^{n-1}$$

$$= n.1.(1-x)^{n-1}$$

$$= n(1-x)^{n-1}$$

Suppose X_1, X_2, X_n are iid uniform 0, 1. What is the expectation of a X_1 and b X_n . Now, we need to first compute the pdf of X_1 . So, we go from the basics 0 1, so we are considering a point x and we are saying that X_1 is at x , what is the pdf? So, before x therefore, there are 0 observations, then one observation is that x and remaining n minus 1 observations are in this region.

Therefore, $f_1 x$ is equal to $nc1$, choosing this 1 element, fx multiplied by $1 - Fx$, whole to the power n minus 1 is equal to n , fx is equal to 1 because it is uniform 0 1 and $1 - Fx$ is equal to $1 - x$ whole to the power n minus 1, is equal to n into $1 - x$ whole to the power n minus 1.

(Refer Slide Time: 3:46)

Handwritten derivation of the expected value $E(X_0)$ for a Beta(2, n) distribution. The derivation is as follows:

$$\begin{aligned} \therefore E(X_0) &= \int_0^1 x \cdot n \cdot (1-x)^{n-1} dx \\ &= n \int_0^1 x^{2-1} (1-x)^{n-1} dx \\ &= \frac{B(2, n)}{B(2, n)} \\ &= \frac{\Gamma(2) \Gamma(n)}{\Gamma(n+2)} \\ \therefore E(X_0) &= n \cdot \frac{\Gamma(2) \Gamma(n)}{\Gamma(n+2)} \\ &= \frac{n!}{(n+1)!} \\ &= \frac{1}{n+1} \end{aligned}$$

Therefore, expected value of X_1 is equal to integration 0 to 1, x into n into 1 minus x , whole to the power n minus 1 , dx is equal to n into integration 0 to 1, x to the power 2 minus 1 into 1 minus x whole to the power n minus 1 dx . Now, this is a beta integral and that is going to give us $\text{beta}(2, n)$ is equal to $\frac{\Gamma(2) \Gamma(n)}{\Gamma(n+2)}$.

Therefore, expected value of X_1 is equal to n into $\frac{\Gamma(2) \Gamma(n)}{\Gamma(n+2)}$ is equal to $\frac{n!}{(n+1)!}$, $\Gamma(n+2)$ is equal to $(n+1)!$ is equal to $n!$ upon $n+1$ is equal to $\frac{1}{n+1}$.

(Refer Slide Time: 5:18)

Handwritten derivation of the expected value $E(X_m)$ for a Beta(n, 1) distribution. The derivation is as follows:

Let us compute $E(X_m)$
 pdf of $X_m = f_m(x)$ we compute as follows:

$$\begin{aligned} \therefore f_m(x) &= \frac{n!}{(n-1)!} (F(x))^{n-1} \cdot f(x) \\ &= n \cdot x^{n-1} \cdot 1 \\ &= n x^{n-1} \end{aligned}$$

Therefore, $E(X_m) = \int_0^1 x \cdot n \cdot x^{n-1} dx$

$$\begin{aligned} &= n \int_0^1 x^n dx \\ &= n \cdot \frac{x^{n+1}}{n+1} \Big|_0^1 \\ &= \frac{n}{n+1} \end{aligned}$$

Let us now compute expected value of X_n the n th order statistic. So, pdf of X_n is equal to $f_n(x)$, we compute as follows. Suppose, it is 0, it is 1, this is x and n th observation is here, so that will give us $f_n(x)$ all other remaining $n - 1$ will be in this region. So, that will give us a $F_n(x)$ to the power $n - 1$ and there should be 0 observations here.

Therefore, $f_n(x)$ is equal to $n \cdot x^{n-1}$, $F_n(x)$ to the power $n - 1$ into $f_n(x)$ is equal to $n \cdot x^n$ to the power $n - 1$. Therefore, expected value of X_n is equal to integration 0 to 1, $x \cdot n \cdot x^{n-1} dx$ is equal to n into 0 to 1, x^n to the power n dx is equal to n into x^{n+1} upon $n + 1$, 0 to 1, which is equal to n upon $n + 1$.

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Handwritten derivation in a Windows Journal window:

$$\text{Therefore, } E(X_{(1)}) = \frac{1}{n+1}$$

$$E(X_{(n)}) = \frac{n}{n+1}$$

Hence the Expected Value of the range.
i.e. Maximum - Minimum = $X_{(n)} - X_{(1)}$

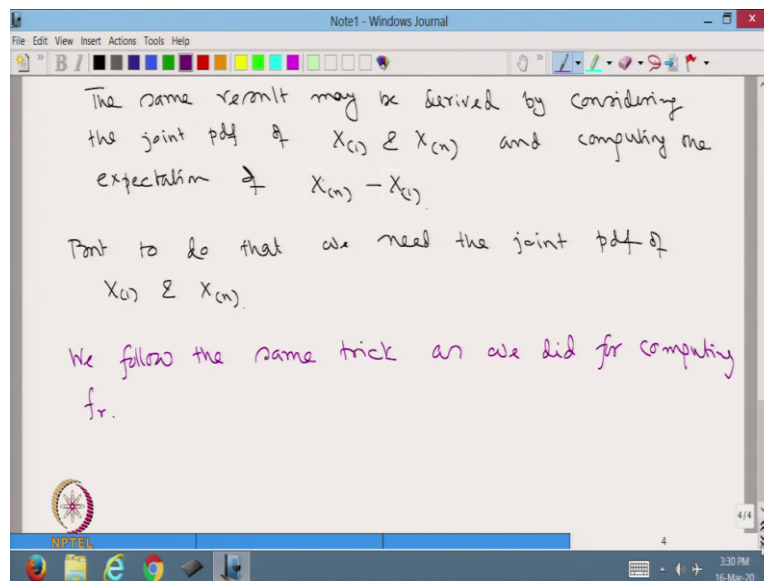
$$\therefore E(\text{Range}) = E(X_{(n)}) - E(X_{(1)}) \quad // \text{ by linearity of Expectation}$$

$$= \frac{n}{n+1} - \frac{1}{n+1}$$

$$= \frac{n-1}{n+1}$$

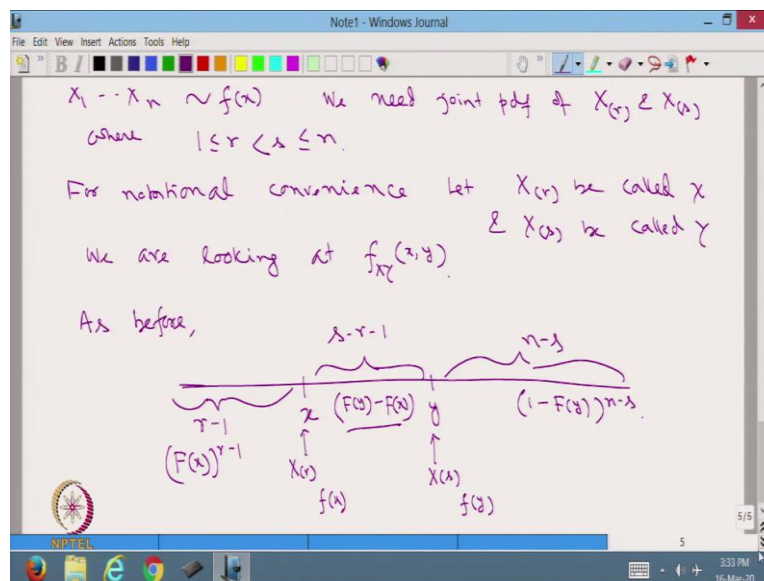
Therefore, expectation of X_1 is equal to 1 upon $n + 1$, expectation of X_n is equal to n upon $n + 1$. Hence, the expected value of the range, what is range? Range is equal to maximum minus minimum is equal to X_n minus X_1 . Therefore, expected value of range is equal to the expected value of X_n minus expected value of X_1 by linearity of expectation is equal to n upon $n + 1$ minus 1 upon $n + 1$ is equal to $n - 1$ upon $n + 1$.

(Refer Slide Time: 9:03)



The same result may be derived by considering the joint pdf of X_1 and X_n and computing the expectation of X_n minus X_1 . But to do that we need the joint pdf of X_1 and X_n . We follow the same trick as we did for computing f_r .

(Refer Slide Time: 10:35)



So, the problem is in X_1, X_2, \dots, X_n follow a pdf f_x we need joint pdf of X_r and X_s , where $1 \leq r < s \leq n$. For notational convenience let X_r be called X and X_s be called Y . We are looking at f_{XY} at x, y . So as before, suppose this is the real line, this is the point x , this is the point y , therefore, the r th order statistic is going to take this value which we know will give f_x the s th order statistic is going to take the value y and that we get from f_y .

Now, r minus 1 many values are here. For them, we shall get F_x to the power r minus 1 s minus r minus 1 values are there. For them we will get F_y minus F_x because they are lying between x and y and for remaining n minus s . We shall get 1 minus F_y to the power n minus s , once we understand that, we can compute the pdf.

(Refer Slide Time: 13:10)

$$\begin{aligned}
 f_{rs}(x, y) &= \frac{n!}{(r-1)!(n-r+1)!} (F(x))^{r-1} f(x) \frac{(n-r)!}{(s-r)! (n-s+1)!} (F(y)-F(x))^{s-r-1} \\
 &\quad \times \frac{(n-s)!}{(n-s)!} (1-F(y))^{n-s} f(y) \\
 &= \frac{n!}{(r-1)!(n-r+1)!} (F(x))^{r-1} f(x) \frac{(n-r)!}{(s-r)! (n-s+1)!} (F(y)-F(x))^{s-r-1} \\
 &\quad \times \frac{(n-s)!}{(n-s)!} (1-F(y))^{n-s} f(y) \\
 &= \frac{n!}{(r-1)!(s-r)! (n-s)!} (F(x))^{r-1} (F(y)-F(x))^{s-r-1} (1-F(y))^{n-s} f(x) f(y)
 \end{aligned}$$

Joint pdf of $X_{(r)}$ & $X_{(s)}$ at (x, y)

Therefore, f_{rs} at x, y , this is the joint pdf of X_r and X_s , is equal to n r minus 1, F_x to the power r minus 1, then out of n minus r plus 1, we choose 1 and we give f_x out of remaining n minus r , we choose s minus r minus 1, we assign the probability F_y minus F_x to the power s minus r minus 1, then out of n minus s plus 1, we choose 1, give the value f_y and the remaining n minus s , they get 1 minus F_y therefore whole to the power n minus s .

Apparently very complicated we try to simplify it, is equal to factorial n upon factorial r minus 1 into factorial n minus r plus 1, F_x to the power r minus 1 multiplied by n minus r plus 1 factorial upon 1 factorial into n minus r factorial and we give the value f_x , then n minus r factorial upon s minus r minus 1 factorial into n minus s plus 1 factorial multiplied by F_y minus F_x , F_y minus F_x to the power s minus r minus 1 multiplied by n minus s plus 1 factorial upon 1 factorial into n minus s factorial multiplied by f_y multiplied by 1 minus F_y whole to the power n minus s . Is equal to, if we look at, this gets canceled with this, n minus r factorial gets cancelled with n minus r factorial s minus r , n minus s plus 1 factorial gets cancelled with n minus s plus 1 factorial.

Therefore, we get n factorial upon r minus 1 factorial into s minus r minus 1 factorial into n minus s factorial and this multiplied by F_x to the power r minus 1 F_y minus F_x to the power s

minus r minus 1 1 minus F_y to the power n minus s into f_x into f_y . So, that is the joint pdf of X_r and X_s at x, y . We are not going to give the mathematical proof, we have already understood the approach. Therefore, from there we are going to solve the problem.

(Refer Slide Time: 17:38)

Problem: To compute $E(X_n - X_1)$ i.e. Expected value of Range.

\therefore Joint pdf of $f_{1,n}(x,y) =$

Diagram: A horizontal line segment from 0 to 1. A point x is marked on the line, with an arrow pointing to it from below labeled X_1 . A point y is marked on the line, with an arrow pointing to it from below labeled X_n . The segment between x and y is bracketed and labeled $n-2$.

$$\begin{aligned}
 &= n_1 f(x) \cdot (n-1) f(y) \cdot (F(y) - F(x))^{n-2} \\
 &= n \cdot (n-1) \cdot f(x) \cdot f(y) \cdot (F(y) - F(x))^{n-2} \\
 &= n(n-1) \cdot 1 \cdot 1 \cdot (y-x)^{n-2} \\
 &= \underline{n(n-1)(y-x)^{n-2}}
 \end{aligned}$$

Since the parent distn is $U(0,1)$
 $\therefore f(x) = 1 = f(y)$
 $\angle F(x) = x$

The problem is to compute expected value of X_n minus X_1 that is expected value of range. Therefore, what we do, joint pdf of X_1 and X_n at the point x, y , therefore we write in a very similar way is equal to this is x X_1 is coming here, this is y X_n is coming here. So, there is nothing on this side, there is nothing on this side and remaining n minus 2 observations are here.

Therefore, the joint pdf we can easily write it as $n-1$ f_x out of the n minus 1, n minus 1 c_1 that we put at f_y multiplied by f_y minus f_x whole to the power n minus 2 is equal to n into n minus 1 into f_x into f_y into F_y minus F_x whole to the power n minus 2. Since, the parent distribution is uniform 0, 1.

Therefore, f_x is equal to 1 is equal to f_y and F_x is equal to x . Hence, putting the value here n into n minus 1 into 1 into 1 into y minus x whole to the power n minus 2 is equal to n into n minus 1 into y minus x , whole to the power n minus 2. So, that is the joint pdf of X_1 and X_n .

(Refer Slide Time: 20:20)

The image shows a handwritten derivation in a Windows Journal window titled "Note1 - Windows Journal". The derivation is as follows:

$$\begin{aligned}
 \therefore E(X_n - X_{n-1}) &= \int_0^1 \int_x^1 n(n-1)(y-x)^{n-2} (y-x) dy dx \\
 &= \int_0^1 \int_x^1 n(n-1)(y-x)^{n-1} dy dx \\
 &= \int_0^1 \left[n(n-1) \frac{(y-x)^n}{n} \right]_x^1 dx \\
 &= \int_0^1 n(n-1) \frac{(1-x)^n}{n} dx \\
 &= \int_0^1 (n-1)(1-x)^n dx
 \end{aligned}$$

On the right side of the derivation, the integral is evaluated:

$$\begin{aligned}
 &= (n-1) \frac{(1-x)^{n+1}}{n+1} \Big|_0^1 \\
 &= - \left((n-1) \left(\frac{0}{n+1} - \frac{1}{n+1} \right) \right) \\
 &= \frac{n-1}{n+1}
 \end{aligned}$$

Below the final result, it is noted: \therefore We get the same result as before.

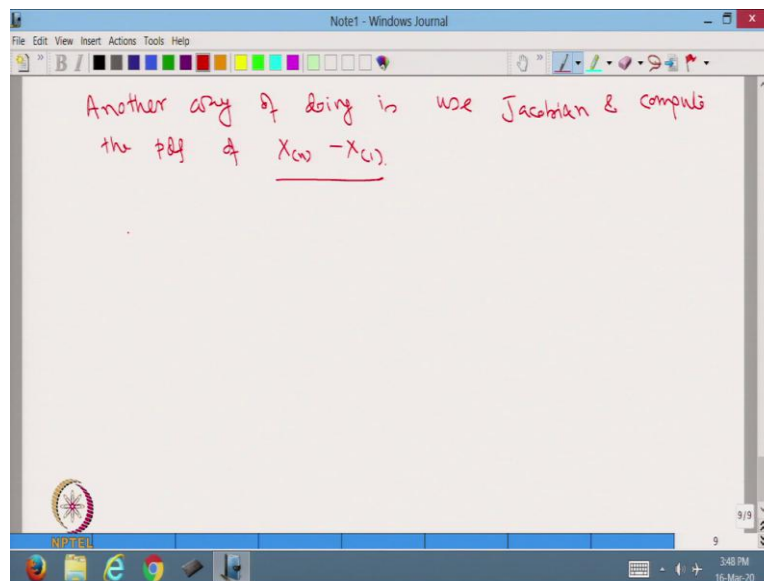
Therefore, expected value of X_n minus X_{n-1} is equal to, now we have to use double integration, n into $n-1$ into $y-x$ whole to the power $n-2$ multiplied by $y-x$, so this is the function for which we are trying to find out the expectation into $dy dx$.

Now, we have to put the range y cannot be less than x , therefore, range of y is equal to x to 1 and x can be anything between 0 to 1 is equal to integration 0 to 1 , integration x to 1 , n into $n-1$ into $y-x$ whole to the power $n-1$ $dy dx$ is equal to 0 to 1 n into $n-1$. Now, we are integrating $y-x$ whole to the power $n-1$ with respect to y .

Therefore, we are getting $y-x$ whole to the power n upon n , x to 1 dx is equal to integration 0 to 1 n into $n-1$. When we put the value 1 here, we get $1-x$ whole to the power n upon n minus this is going to be 0 . If we put x , therefore, we do not need to put that dx is equal to integration 0 to 1 $n-1$ into $1-x$ whole to the power n dx .

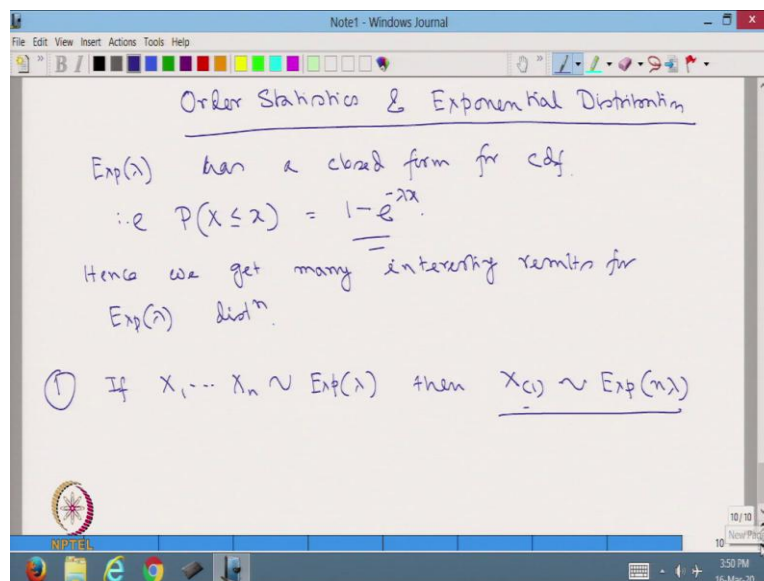
Is equal to $n-1$ goes out, $1-x$ whole to the power $n+1$ upon $n+1$ from 0 to 1 , but it comes with a negative sign because x here is negative is equal to minus $n-1$ into 0 upon $n+1$ minus 1 upon $n+1$ is equal to $n-1$ upon $n+1$. Therefore, we get the same result as before.

(Refer Slide Time: 23:50)



Another way of doing it is use Jacobian and compute the pdf of X_n minus X_1 . I leave that as an exercise.

(Refer Slide Time: 24:23)



Let us now look at order statistics and exponential distribution, as I have said before, exponential lambda has a closed form for cdf that is probability X less than or equal to x is equal to 1 minus e to the power minus lambda x . Hence, we get many interesting results for exponential lambda distribution.

1, that we have already seen that, if X_1, X_2, \dots, X_n follow the exponential lambda then X_1 the order statistic follows exponential with n lambda, this we have already seen, so I am not going to prove it again. So, let us consider this following result.

(Refer Slide Time: 26:15)

② If X_1, X_2, X_3 are iid $\text{Exp}(\lambda)$
 What is the expected value for range $= X_{(3)} - X_{(1)}$
 Since $E(\text{Range}) = E(X_{(3)}) - E(X_{(1)})$
 Let us compute them separately.
 $X_{(1)} \sim \text{Exp}(3\lambda) \therefore E(X_{(1)}) = \frac{1}{3\lambda}$
 $X_{(3)} \sim 3\lambda e^{-\lambda x} (1 - e^{-\lambda x})^2 = 3\lambda e^{-\lambda x} (1 - 2e^{-\lambda x} + e^{-2\lambda x})$
 $= 3\lambda e^{-\lambda x} - 6\lambda e^{-2\lambda x} + 3\lambda e^{-3\lambda x}$
 $\therefore E(X_{(3)}) = \int_0^{\infty} x \cdot 3\lambda e^{-\lambda x} dx - \int_0^{\infty} x \cdot 6\lambda e^{-2\lambda x} dx + \int_0^{\infty} x \cdot 3\lambda e^{-3\lambda x} dx$

If X_1, X_2, X_3 are iid exponential lambda what is the expected value for range is equal to X_3 minus X_1 . Since, expected value of range is equal to expected value of X_3 minus expected value of X_1 . Let us compute them separately, now X_1 follows exponential with 3 lambda that we have seen. Therefore, expected value of X_1 is equal to 1 upon 3 lambda. Now, X_3 what is the pdf, very simple out of 3 , we choose 1 that can be done in 3 ways and put it at the point x that gives us f_x and the remaining two they are less than equal to x .

Therefore, 1 minus e to the power minus lambda x whole square is equal to 3 lambda e to the power minus lambda x into minus $2e$ to the power minus lambda x plus e to the power minus 2 lambda x is equal to 3 lambda e to the power minus lambda x minus 6 lambda e to the power minus 2 lambda x plus 3 lambda e to the power minus 3 lambda x .

Therefore expected value of X_3 is equal to integration 0 to infinity $x \cdot 3$ lambda e to the power minus lambda x dx minus 0 to infinity $x \cdot 6$ lambda e to the power minus 2 lambda x dx plus integration 0 to infinity $x \cdot 3$ lambda e to the power minus 3 lambda x dx .

(Refer Slide Time: 29:51)

Calculate them separately:

$$\int_0^\infty x 3\lambda e^{-\lambda x} dx = 3 E(\text{Exp}(\lambda)) = \frac{3}{\lambda}$$

$$-\int_0^\infty x 6\lambda e^{-2\lambda x} dx = -3 \left(-\int_0^\infty x \cdot 2\lambda e^{-2\lambda x} dx \right) = -\frac{3}{2\lambda}$$

$$\int_0^\infty x 3\lambda e^{-3\lambda x} dx = \frac{1}{3\lambda}$$

$$\therefore E(R) = \frac{3}{\lambda} - \frac{3}{2\lambda} + \frac{1}{3\lambda} - \frac{1}{3\lambda}$$

$$= \frac{3}{\lambda} - \frac{3}{2\lambda} = \frac{1}{\lambda} \left(3 - \frac{3}{2} \right) = \frac{1}{\lambda} \left(1 + \frac{1}{2} \right)$$

So, let us calculate them one by one, integration 0 to infinity $x 3 \lambda e$ to the power minus λx is equal to 3 into expected value of exponential λ is equal to 3 by λ . The second term is, minus integration of 0 to infinity $x 6 \lambda e$ to the power minus $2 \lambda x$ is equal to 3 into minus 0 to infinity x to λe to the power minus $2 \lambda x$ is equal to minus 3 upon 2λ and thirdly integration 0 to infinity $x 3 \lambda e$ to the power minus $3 \lambda x$ is equal to 1 upon 3λ , because it is the expectation of exponential distribution with 3λ .

Therefore, the expected value of range is equal to 3 by λ minus 3 by 2λ plus 1 by 3λ minus 1 by 3λ coming from the expected value of X_1 is equal to 3 by λ minus 3 by 2λ is equal to 1 by λ into 3 minus 3 by 2 is equal to 1 by λ 1 plus 1 by 2.

(Refer Slide Time: 32:24)

Let us calculate the Expected value of range when 4 samples are taken for $Exp(\lambda)$.

$$E(\text{Range}) = E(X_{(4)}) - E(X_{(1)})$$

Now $E(X_{(1)}) = \frac{1}{4\lambda}$ (\because Minimum of 4 observations $\sim Exp(4\lambda)$)

$E(X_{(4)}) = ?$

$$f_4(x) = 4\lambda e^{-\lambda x} (1 - e^{-\lambda x})^3$$

$$= 4\lambda e^{-\lambda x} (1 - 3e^{-\lambda x} + 3e^{-2\lambda x} - e^{-3\lambda x})$$

$$= 4\lambda e^{-\lambda x} - 12\lambda e^{-2\lambda x} + 12\lambda e^{-3\lambda x} - 4\lambda e^{-4\lambda x}$$

Let us now calculate the expected value of range when 4 samples are taken from exponential with lambda. Now, expectation of range is equal to expected value of fourth order statistic minus expected value of the first order statistic, now expected value of X_1 is equal to 1 upon 4 lambda. Since, minimum of 4 observations will follow exponential with 4 lambda.

Now, expectation of X_4 is equal to what? To compute that, let us look at f_4 of x which is going to be out of 4, you choose 1 and put it at the point x with f_x and the remaining 3 will be in less than x . Therefore, f_x to the power 3 is equal to 4 lambda e to the power minus lambda x into 1 minus 3 into 1 square into e to the power minus lambda x plus 3 into 1 into e to the power minus lambda x square minus e to the power minus lambda x whole cube.

Is equal to 4 lambda e to the power minus lambda x into 1 minus 3 e to the power minus lambda x plus 3 e to the power minus 2 lambda x minus e to the power minus 3 lambda x , so that is the pdf for the fourth order statistic, we can write it as is equal to 4 lambda e to the power minus lambda x minus 12 lambda e to the power minus 2 lambda x plus 12 lambda e to the power minus 3 lambda x minus 4 lambda e to the power minus 4 lambda x .

(Refer Slide Time: 36:00)

The image shows a handwritten derivation in a Windows Journal window. The text is as follows:

$$\begin{aligned} \therefore E(X) &= \int_0^{\infty} x 4\lambda e^{-2x} dx - 12 \int_0^{\infty} x \lambda e^{-2x} dx + 12 \int_0^{\infty} x \lambda e^{-3x} dx - \int_0^{\infty} x 4\lambda e^{-4x} dx \\ &= \frac{4}{\lambda} - \frac{6}{2\lambda} + \frac{4}{3\lambda} - \frac{1}{4\lambda} \\ \therefore E(\text{Range}) &= \frac{4}{\lambda} - \frac{6}{2\lambda} + \frac{4}{3\lambda} - \frac{1}{4\lambda} - \frac{1}{4\lambda} \quad \leftarrow E(X) \\ &= \frac{1}{\lambda} \left(4 - 3 + \frac{4}{3} - \frac{1}{2} \right) \\ &= \frac{1}{\lambda} \left(1 + \frac{8-3}{6} \right) = \frac{1}{\lambda} \left(1 + \frac{5}{6} \right) = \frac{1}{\lambda} \left(1 + \frac{1}{2} + \frac{1}{3} \right) \end{aligned}$$

Therefore, expected value of x 4 is equal to integration 0 to infinity x into $4\lambda e^{-2x}$ minus $12 \int_0^{\infty} x \lambda e^{-2x} dx$ plus $12 \int_0^{\infty} x \lambda e^{-3x} dx$ minus integration 0 to infinity $4\lambda e^{-4x}$ is equal to $\frac{4}{\lambda}$ minus $\frac{6}{2\lambda}$ plus $\frac{4}{3\lambda}$ minus $\frac{1}{4\lambda}$.

Therefore, expected value of range is equal to $\frac{4}{\lambda}$ minus $\frac{6}{2\lambda}$ plus $\frac{4}{3\lambda}$ minus $\frac{1}{4\lambda}$ minus $\frac{1}{4\lambda}$, this is coming from expected value of X_1 is equal to $\frac{1}{\lambda} \left(4 - 3 + \frac{4}{3} - \frac{1}{2} \right)$ is equal to $\frac{1}{\lambda} \left(1 + \frac{8-3}{6} \right)$ is equal to $\frac{1}{\lambda} \left(1 + \frac{5}{6} \right)$ is equal to $\frac{1}{\lambda} \left(1 + \frac{1}{2} + \frac{1}{3} \right)$.

(Refer Slide Time: 38:39)

$\therefore E(\text{Range})$ when 3 samples are taken $= \frac{1}{\lambda} (1 + \frac{1}{2})$
 $E(\text{Range})$ when 4 samples are taken $= \frac{1}{\lambda} (1 + \frac{1}{2} + \frac{1}{3})$
 \hookrightarrow A Natural guess is $E(\text{Range})$ when n samples are taken will be $\frac{1}{\lambda} (1 + \frac{1}{2} + \dots + \frac{1}{n-1})$
 To prove this we need some other results.

Therefore, expected value of range when 3 samples are taken is equal to 1 by lambda 1 plus half the expected value of range when 4 samples are taken is equal to 1 by lambda into 1 plus half plus 1 by 3. A natural guess is the expected value of range, when n samples are taken, will be 1 by lambda into 1 plus half plus up to 1 upon n minus 1. How to prove it? To prove this, we need some other results, so here is a result.

(Refer Slide Time: 40:22)

Ex: Show that if $X_1, \dots, X_n \sim \text{iid Exp}(\lambda)$,
 then $X_{(r)}$ & $X_{(r+1)} - X_{(r)}$ are distributed independently.
 The pdf of $X_{(r+1)} - X_{(r)}$ is $\text{Exp}((n-r)\lambda)$
 Ans: We know pdf of $X_{(r)} = \frac{n!}{(r-1)!(n-r)!} (F(x))^{r-1} (1-F(x))^{n-r} f(x)$
 & The joint pdf of $X_{(r)}, X_{(r+1)}$ is:

$$f_{X_{(r)}, X_{(r+1)}}(x, y) = \frac{n!}{(r-1)!(n-r)!} \frac{(n-r)!}{1!(n-r)!} F(x)^{r-1} f(x) (F(y)-F(x))^{n-r-1} f(y)$$

Show that if X_1, X_2, \dots, X_n are iids with exponential lambda, then X_r and $X_{r+1} - X_r$ are distributed independently, that is a very interesting result. And second thing is the pdf of $X_{r+1} - X_r$ is exponential with $n - r$ lambda. So, we want to show this answer. We know pdf of X_r is equal to factorial n upon factorial $r - 1$ into $n - r$ factorial F_x to

the power $r - 1$ $1 - F_x$ to the power $n - r$ into f_x and the joint pdf of X_r and X_{r+1} is how to get that, suppose, X_r takes the value x , X_{r+1} takes the value y , then there are $r - 1$ observations here and there are $n - r - 1$ observations are there and 0 observations in between.

Therefore, $f_r, r+1$ at x, y is equal to $n - r - 1$ F_x to the power $r - 1$, $n - r - 1$ $1 - F_x$ times f_x $n - r - 1$ times f_y and $F_y - F_x$ whole to the power $n - r - 1$ is equal to after simplification, we will get factorial n upon factorial $r - 1$ into factorial $n - r + 1$ into $n - r + 1$ factorial 1 factorial $n - r$ factorial into $n - r$ factorial upon 1 factorial $n - r - 1$ factorial multiplied by F_x to the power $r - 1$ small f_x small f_y into $F_y - F_x$ whole to the power $n - r - 1$.

(Refer Slide Time: 44:39)

Since the parent density is $\text{Exp}(\lambda)$, we have
 $f(x) = \lambda e^{-\lambda x}$ & $F(x) = (1 - e^{-\lambda x})$
 $\therefore f_r(x) = \frac{n!}{(r-1)!(n-r)!} (1 - e^{-\lambda x})^{r-1} \cdot \lambda e^{-\lambda x} \cdot (1 - (1 - e^{-\lambda x}))^{n-r}$
 $= \frac{n!}{(r-1)!(n-r)!} (1 - e^{-\lambda x})^{r-1} \lambda e^{-\lambda x} (e^{-\lambda x})^{n-r+1}$
 $f_{r,r+1}(x,y) = \frac{n!}{(r-1)!(n-r-1)!} (1 - e^{-\lambda x})^{r-1} \cdot \lambda e^{-\lambda x} \cdot \lambda e^{-\lambda y} \cdot (1 - (1 - e^{-\lambda y}))^{n-r-1}$
 $= \frac{n!}{(r-1)!(n-r-1)!} (1 - e^{-\lambda x})^{r-1} e^{-\lambda x} \cdot e^{-\lambda y} \cdot e^{-\lambda(n-r-1)y} \cdot \lambda^2$
 need to find the pdf of $X_{(r+1)} - X_{(r)}$.

Ex: Shows that if $X_1, \dots, X_n \sim \text{iid Exp}(\lambda)$
 then $X_{(r)}$ & $X_{(r+1)} - X_{(r)}$ are distributed independently
 . The pdf of $X_{(r+1)} - X_{(r)}$ is $\text{Exp}((n-r)\lambda)$
 Ans: We know pdf of $X_{(r)} = \frac{n!}{(r-1)!(n-r)!} (F(x))^{r-1} (1-F(x))^{n-r} f(x)$
 & The joint pdf of $X_{(r)}, X_{(r+1)}$ is:
 $f_{r,r+1}(x,y) = \frac{n!}{(r-1)!(n-r-1)!} (F(x))^{r-1} (1-F(x))^{n-r-1} f(x) \cdot (F(y) - F(x))^{n-r-1} f(y)$
 $= \frac{n!}{(r-1)!(n-r-1)!} (F(x))^{r-1} (1-F(x))^{n-r-1} f(x) \cdot (F(y) - F(x))^{n-r-1} f(y)$

Since, the parent density is exponential with lambda, we have f_X is equal to $\lambda e^{-\lambda x}$ and capital F_X the cdf is equal to $1 - e^{-\lambda x}$. Therefore, $f_{X(r)}$ is equal to $n!$ factorial upon $(r-1)!$ factorial into $(n-r)!$ factorial $1 - e^{-\lambda x}$ to the power $r-1$ $\lambda e^{-\lambda x}$ to the power $n-r$ is equal to $n!$ factorial $(r-1)!$ factorial $(n-r)!$ factorial $1 - e^{-\lambda x}$ to the power $r-1$ $\lambda e^{-\lambda x}$ to the power $n-r$ into $(n-r+1)$. So, that is the pdf of the r th order statistic.

What is a $f_{X(r+1)}$ the joint density at x, y by using the formula we derived just now, in which we can cancel this we have $n!$ factorial upon $(r-1)!$ factorial into $(n-r)!$ factorial into $1 - e^{-\lambda x}$ to the power $r-1$ $\lambda e^{-\lambda x}$ to the power $n-r$ $\lambda e^{-\lambda y}$ to the power $n-r-1$ $1 - e^{-\lambda y}$ to the power $r-1$ is equal to $n!$ factorial $(r-1)!$ factorial $(n-r-1)!$ factorial $1 - e^{-\lambda x}$ to the power $r-1$ $e^{-\lambda x}$ to the power $n-r$ $e^{-\lambda y}$ to the power $n-r-1$ y multiplied by λ^2 . So, that is the joint pdf of x, y . We need to find the pdf of $X_{r+1} - X_r$.

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Let us call $U = X_{(r+1)} - X_{(r)}$ $\therefore X_{(r)} = V$ $\therefore |J| = 1$
 $V = X_{(r)}$ $X_{(r+1)} = U + V$

$$f_{UV}(u,v) = f_{XY}(x,y) \cdot |J|$$

expressed in terms of u, v

$$= \frac{n!}{(r-1)!(n-r)!} \lambda^2 (1 - e^{-\lambda v})^{r-1} e^{-\lambda v} e^{-\lambda(u+v)} e^{-\lambda(n-r-1)(u+v)}$$

$$= \frac{n!}{(r-1)!(n-r)!} \lambda^2 (1 - e^{-\lambda v})^{r-1} e^{-\lambda v(2+n-r-1)} e^{-\lambda u(n-r+1)}$$

$$= \frac{n!}{(r-1)!(n-r)!} \lambda^2 (1 - e^{-\lambda v})^{r-1} e^{-\lambda v(n-r+1)} e^{-\lambda u(n-r)}$$

$$= \frac{n!}{(r-1)!(n-r)!} \lambda^2 (1 - e^{-\lambda v})^{r-1} e^{-\lambda v(n-r+1)} \cdot \underbrace{(n-r)\lambda e^{-\lambda(n-r)u}}_{f_U(u)}$$

So, let us call U is equal to $X_{r+1} - X_r$ and V is equal to X_r . Therefore, X_r is equal to V , X_{r+1} is equal to $U + V$. Therefore, Jacobean as we have computed many times is equal to 1. Therefore, f_{UV} at u, v is equal to joint pdf of X_r and X_{r+1} at the point x, y multiplied by 1 expressed in terms of u and v .

Is equal to n factorial upon r minus 1 factorial n minus r minus 1 factorial, λ square 1 minus e to the power minus λv whole to the power r minus 1, e to the power minus λv , e to the power minus λu plus v , e to the power minus λn minus r minus 1 into u plus v multiplied by 1 so that is not needed.

Is equal to n factorial r minus 1 factorial into n minus r minus 1 factorial λ square 1 minus e to the power minus λv whole to the power r minus 1 e to the power minus λv into 1 plus 2 plus n minus r minus 1 multiplied by e to the power minus λu into n minus r minus 1 plus 1.

Is equal to n factorial r minus 1 factorial n minus r minus 1 factorial λ 1 minus e to the power minus λv whole to the power r minus 1 e to the power minus λv to the power n minus r plus 1 multiplied by e to the power minus λu whole to the power n minus r .

Is equal to n factorial r minus 1 factorial n minus r factorial λ 1 minus e to the power minus λv to the power r minus 1 e to the power minus λv to the power n minus r plus 1 into, since we have multiplied by n minus r in the denominator, we take out that n minus r , then one of the two λ s, one we have come here and the other will come here multiplied by e to the power minus λn minus r times u .

(Refer Slide Time: 52:44)

Handwritten notes on a digital whiteboard (Note1 - Windows Journal) showing the factorization of the joint pdf of $X_{(r+1)} - X_{(r)}$ and $X_{(r)}$.

Joint pdf of $X_{(r+1)} - X_{(r)}$ can be factorized into two parts:

- 1st part is pdf of $X = X_{(r)}$
- 2nd part is pdf of $U = X_{(r+1)} - X_{(r)}$

We establish that:

- $X_{(r)} \in X_{(r+1)} - X_{(r)}$ are independent
- pdf of $X_{(r+1)} - X_{(r)}$ is $(n-r)\lambda e^{-\lambda(n-r)u}$

Thus $X_{(r+1)} - X_{(r)} \sim (n-r)\lambda e^{-\lambda(n-r)u}$
i.e. $\text{Exp}((n-r)\lambda)$

Therefore, joint pdf of X_{r+1} minus X_r can be factorized into two parts. The first part is pdf of X , which is the r th order statistic. Therefore, the second part is pdf of U is equal to X_{r+1} minus X_r . Therefore, we establish that 1, X_r and X_{r+1} minus X_r are independent

and secondly, pdf of X_r plus 1 minus X_r is equal to n minus r , λe to the power minus $\lambda(n - r)$. Thus, X_r plus 1 minus X_r is distributed as n minus r λe to the power minus $\lambda(n - r)$, that is exponential with n minus r λ .

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The image shows a handwritten derivation in a Windows Journal window. The text is as follows:

$$\therefore E(X_{(n)} - X_{(r)}) = \frac{1}{(n-r)\lambda}$$

$$\therefore \text{Range} = X_{(n)} - X_{(1)}$$

$$\text{Hence } E(\text{Range}) = E(X_{(n)} - X_{(1)})$$

$$= E(X_{(n)}) - E(X_{(n-1)}) + E(X_{(n-1)}) - E(X_{(n-2)}) - \dots + E(X_{(1)})$$

$$= E(X_{(n)} - X_{(n-1)}) + E(X_{(n-1)} - X_{(n-2)}) + \dots + E(X_{(2)} - X_{(1)})$$

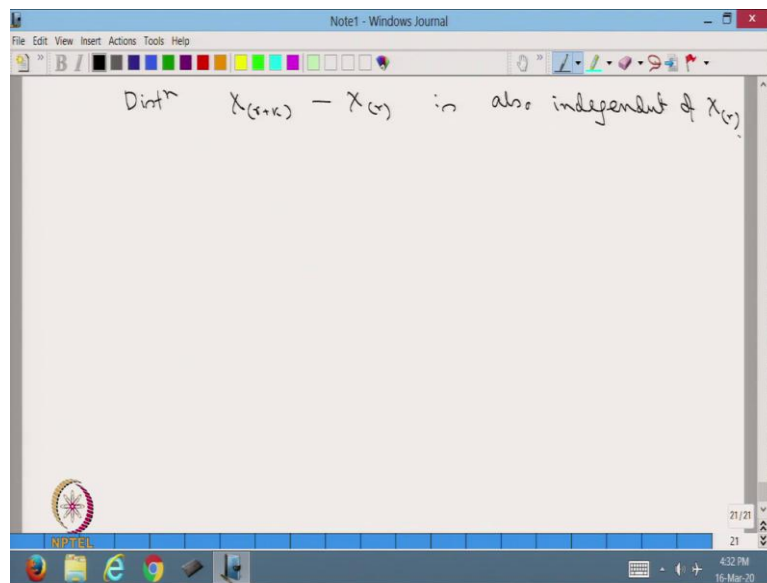
$$= \frac{1}{\lambda} \left(\frac{1}{n-1} + \frac{1}{n-2} + \dots + 1 \right)$$

$$= \frac{1}{\lambda} \left(1 + \frac{1}{2} + \frac{1}{3} + \dots + \frac{1}{n-1} \right)$$

Therefore, X_r plus 1 minus X_r is equal to 1 upon n minus r λ . Therefore, range is equal to X_n minus X_1 . Hence, expected value of range is equal to expected value of X_n minus X_1 is equal to expected value of X_n minus expected value of X_{n-1} plus expected value of X_{n-1} minus expected value of X_{n-2} and like that, if we go, we get plus expected value of X_2 minus expected value of X_1 .

Is equal to expected value of X_n minus X_{n-1} plus expected value of X_{n-1} minus X_{n-2} like that, expectation of X_2 minus X_1 is equal to 1 upon λ 1 upon n minus 1 by taking r is equal to n minus 1 plus 1 upon n minus 2 plus up to 1 is equal to 1 upon λ 1 plus half plus 1 by third 1 by 3 up to 1 upon n minus 1.

(Refer Slide Time: 57:30)



So, that is the result that we wanted to prove okay friends, these are very important result. In fact, the result can be even more generalized, that it is not only successive difference even expectation of, even the distribution of X_r plus k minus X_r is also independent of X_r . I like you to prove that as an exercise. Okay friends, I stop here today, from the next class I shall start with convergence theorems for probability distributions. Okay friends, thank you.