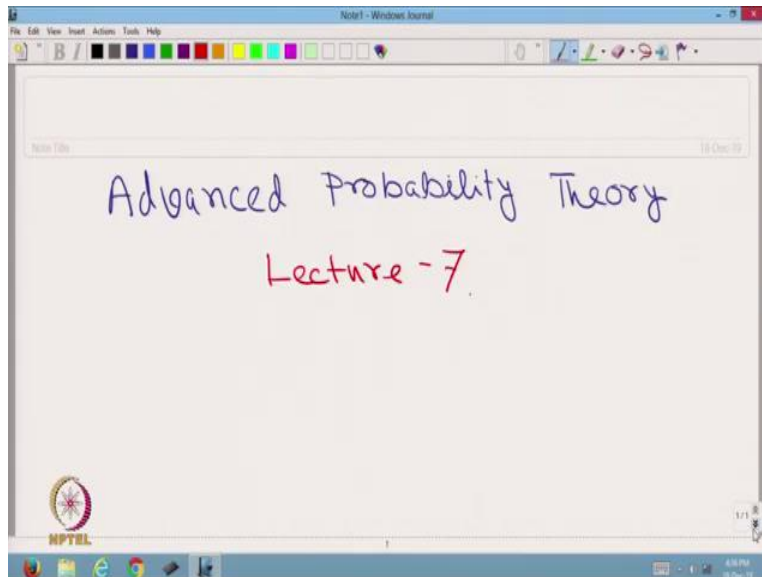


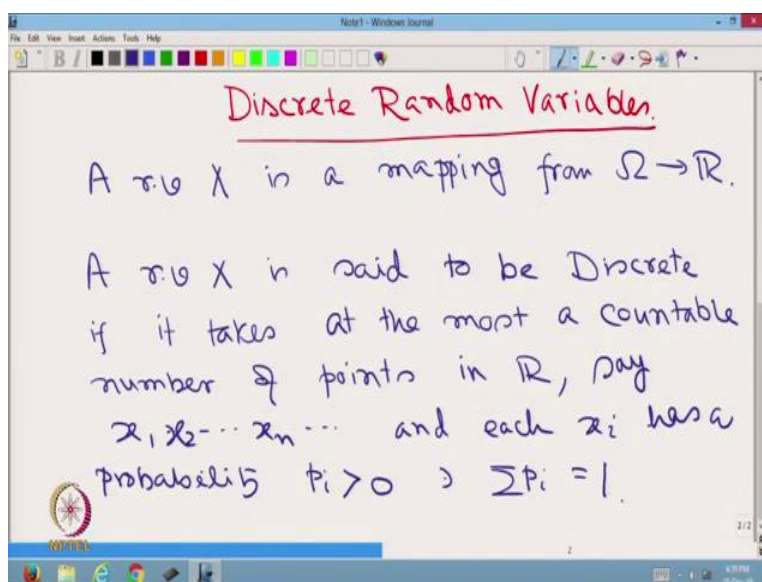
Advanced Probability Theory
Professor Niladri Chatterjee
Department of Mathematics
Indian Institute of Technology, Delhi
Lecture 7

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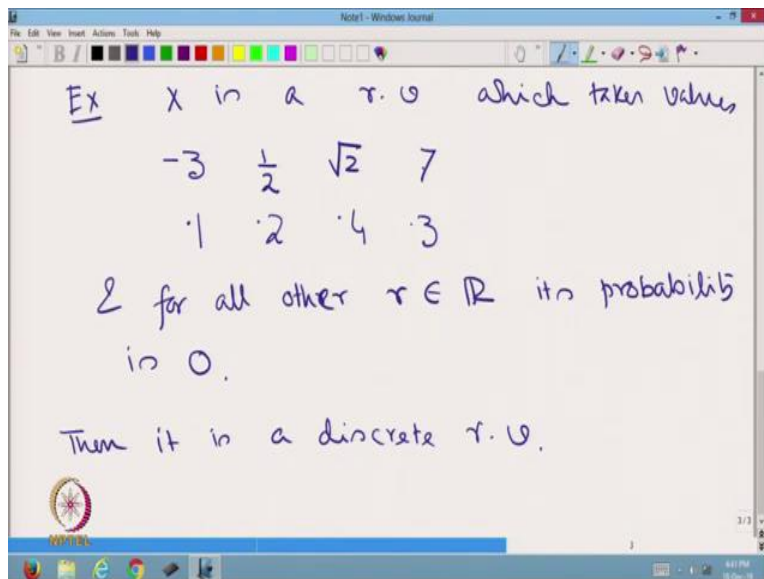
Welcome students. The MOOCs course on Advanced Probability Theory. This is lecture number seven.

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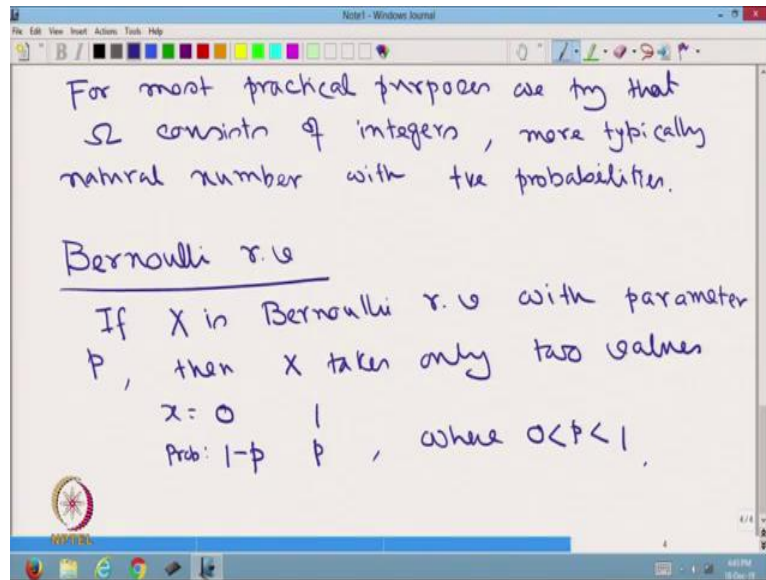
As I discussed in my last class, that in this lecture, we shall start Discrete Random Variables. We know that a random variable X is a mapping from ω to real line, of course, there are certain other properties. I am not going into that. Now, let us look at a discrete random variable. A random variable X is set to be discrete if it takes at the most a countable number of points in \mathbb{R} , say x_1, x_2, \dots, x_n and each x_i has a probability P_i greater than 0 such that $\sum P_i$ is equal to 1.

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Example, suppose X is a random variable which takes values say minus 3, half, root 2 and 7 with probabilities 0.1, 0.2, 0.4 and 0.3 and for all other r belonging to \mathbb{R} its probability is 0, then it is a Discrete Random Variable because it can take only four values and their corresponding probabilities are 0.1, 0.2, 0.4 and 0.3, note that the sum of this is equal to 1.

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For most practical purposes we try that Ω consists of integers, more typically natural numbers with positive probabilities. So, let me now give you several different types of discrete random variables, perhaps, one of the most simplest one is Bernoulli random variable. If X is a Bernoulli random variable with parameter p , then X takes only 2 values 0 and 1 and their probabilities are 1 minus p and p where $0 < p < 1$.

If you remember then we have seen such a random variable, when we were talking about tossing a coin and the corresponding observations were head and tail, when head is mapped into 1 and tail is mapping to 0.

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Consider the expansion of $(a+b)^n$ a, b are real no.

$$(a+b)^n = \binom{n}{0} a^0 b^n + \binom{n}{1} a^1 b^{n-1} + \dots + \binom{n}{n} a^n b^0$$

Now let $a = p$ $0 < p < 1$
 $b = q = 1 - p$

Now let us expand $(q+p)^n$

$$\therefore (q+p) = 1 \quad \therefore (q+p)^n = 1$$

Again: $(q+p)^n = \binom{n}{0} q^0 p^n + \binom{n}{1} q^1 p^{n-1} + \dots + \binom{n}{n} q^n p^0$

$$= 1$$

Now, let us consider the expansion of the a plus b whole to the power n , when a and b are real numbers. We know that a plus b whole to the power n is equal to $\binom{n}{0} a$ to the power 0 b to the power n plus $\binom{n}{1} a$ to the power 1 , b to the power n minus 1 plus up to $\binom{n}{n} a$ to the power n b to the power 0 , all of us have seen such a binomial expansion. Now, let a equal to p , 0 less than p less than 1 , and b is equal to q is equal to 1 minus p .

Now, let us expand q plus p whole to the power n therefore, since, q plus p is equal to 1 . Therefore, q plus p whole to the power n is equal to 1 . Again by using this formula q plus p whole to the power n is equal to $\binom{n}{0} q$ to the power 0 p to the power n plus $\binom{n}{1} q$ to the power 1 p to the power n minus 1 plus up to $\binom{n}{n} q$ to the power n p to the power 0 . Therefore, this sum is going to be equal to 1 .

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This allows us to define a r.v which is called Binomial random variable with parameter p . We shall often denote it by $\text{Bin}(n, p)$ where n is a positive integer > 1 & $0 < p < 1$.

Illustration
 $(q+p)^2 = q^2 + 2pq + p^2$
 Consider a r.v X : $\begin{matrix} 0 & 1 & 2 \\ q^2 & 2pq & p^2 \end{matrix}$
 $\rightarrow \text{Bin}(2, p)$.

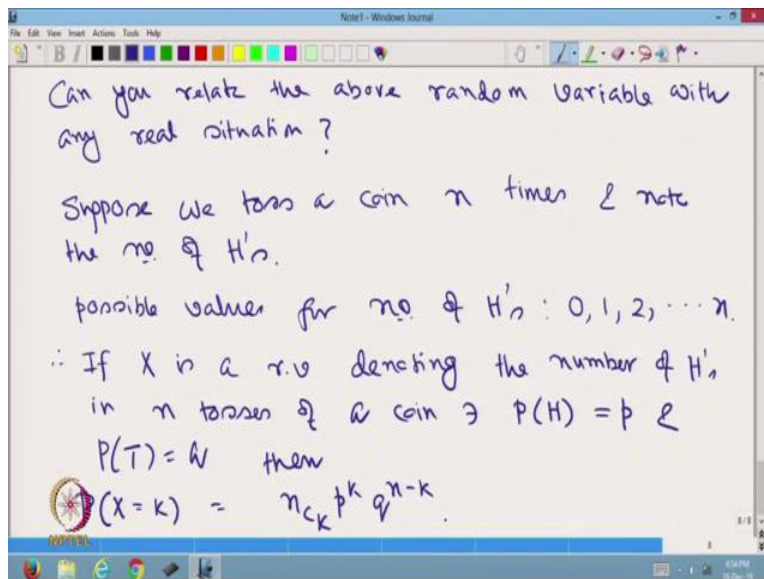
This allows us to define a random variable which is called Binomial random variable with parameter p . We shall often denote it by Binomial n comma P where n is a positive integer greater than 1, and 0 less than p less than 1 . Illustration, q plus P whole square is equal to q square plus $2Pq$ plus P square. Consider a random variable X such that it takes values $0, 1$ and 2 , 0 with probability q square, 1 with probability $2Pq$ and 2 with probability P square. Then this X is a binomial random variable with 2 comma p .

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In a similar way
 If $X \sim \text{Bin}(3, p)$: $X \begin{matrix} 0 & 1 & 2 & 3 \\ q^3 & 3q^2p & 3qp^2 & p^3 \end{matrix}$
 :
 If $X \sim \text{Bin}(n, p)$
 $X : 0 \quad 1 \quad 2 \quad \dots \quad n$
 & $P(X=k) = \underline{{}^nC_k p^k q^{n-k}}$

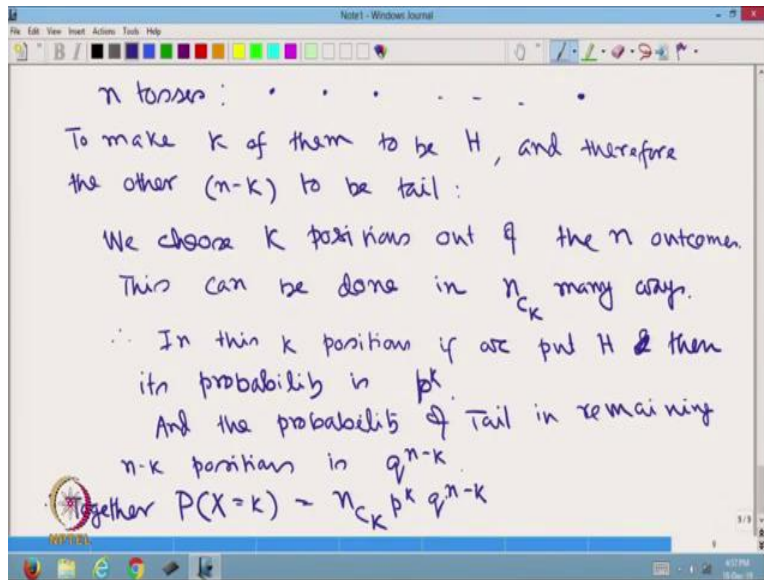
In a similar way, if X is following binomial $3, p$ this notation I will use to denote that X is a random variable, which is following this distribution, then X has the following values 0, 1, 2 and 3 and the corresponding probabilities are q to the power 3, $3q$ square p , $3q$ p square and pq . It is very clear that sum of these is going to be 1. In general if X is a binomial random variable with parameter n comma p then X takes values 0, 1, 2 up to n and probability X is equal to k is equal to nck p to the power k q to the power n minus k .

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Now, if I ask you can you relate the above random variable with any real situation? Perhaps, you can. But still let me explain. Suppose, we toss a coin n times and note the number of heads, then possible values for number of heads can be 0, 1, 2 up to n . Therefore, if X is a random variable denoting the number of heads in n tosses of a coin such that probability of head is equal to p and probability of tail is equal to q . Then probability X is equal to k is nck p to the power k q to the power n minus k .

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Why? Because there are n tosses. So, there may be n outcomes. To make k of them to be H , and therefore, the other n minus k to be tail. What we will do? We choose k positions out of the n outcomes. This can be done in $n_C k$ many ways. Therefore, in this k positions if we put H and its probability is p to the power k because the tosses are independent.

Therefore, to get k heads we have to have the probability p into p into p k times and therefore, it is p to the power k . And the probability of tail in remaining n minus k positions is q to the power n minus k . Therefore, together probability X is equal to k is equal to $n_C k p$ to the power $k q$ to the power n minus k .

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Therefore any r.v. that takes values betⁿ
 $0, 1, 2, \dots, n \Rightarrow P(X=k) = {}^n C_k p^k q^{n-k}$
then X is said to be a Binomial r.v with
parameters n & k .

Suppose there are two coins with probability
of getting $H = p$ for both of them.

If X denotes the no. of H 's in n tosses of
coin 1
 Y denotes the no. of H 's in m tosses of
coin 2

What is the probability of getting k H 's together.

i.e if $X \sim \text{Bin}(n, p)$
& $Y \sim \text{Bin}(m, p)$.

Then how is $X+Y$ distributed?

$X+Y$ is a r.v.

\therefore We are looking at probability $X+Y = k$.
or we are looking at distribution of $X+Y$.

Therefore, any random variable that takes values between $0, 1, 2, \dots, n$ such that probability X is equal to k is equal to ${}^n C_k p^k q^{n-k}$. Then X is said to be a Binomial random variable with parameters n and k . Now, suppose there are two coins with probability of getting a head is equal to p for both of them. If X denotes the number of heads in n tosses of coin 1 and Y denotes the number of heads in m tosses of coin 2, what is the probability of getting k heads together?

If X is equal to Binomial n, p and Y is binomial m, p then how is X plus Y distributed? That is the question. And if you remember, in our last class, we have said that X plus Y is a random variable therefore, we are looking at probability of X plus Y is equal to k or we are looking at distribution of X plus Y .

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It is clear that the total no. of H 's can range betⁿ $0, 1, 2, \dots, m+n$.

Now $X+Y=0$ if $X=0$ & $Y=0$

$$\begin{aligned} \therefore P(X+Y=0) &= P(X=0) \cdot P(Y=0) \\ &= {}^nC_0 p^0 q^n \times {}^mC_0 p^0 q^m \\ &= {}^nC_0 \times {}^mC_0 \cdot q^{m+n} = q^{m+n} \end{aligned}$$

$P(X+Y=1)$?

$$\begin{aligned} &P(X=0) \cdot P(Y=1) + P(X=1) \cdot P(Y=0) \\ &= {}^nC_0 p^0 q^n \times {}^mC_1 p^1 q^{m-1} + {}^nC_1 p^1 q^{n-1} \times {}^mC_0 p^0 q^m \end{aligned}$$

$$\begin{aligned} &= {}^nC_0 p^0 q^n \times {}^mC_1 p^1 q^{m-1} + {}^nC_1 p^1 q^{n-1} \times {}^mC_0 p^0 q^m \\ &= {}^nC_0 \times {}^mC_1 p^1 q^{m+n-1} + {}^nC_1 \times {}^mC_0 p^1 q^{m+n-1} \\ &= (1 \cdot m + n \cdot 1) p^1 q^{m+n-1} = (m+n) p^1 q^{m+n-1} \\ &= {}^{m+n}C_1 p^1 q^{m+n-1} \end{aligned}$$

In general $X+Y=k$ this event can occur if there are

$$\begin{aligned} &P(X=0) \cdot P(Y=k) + \\ &+ P(X=1) \cdot P(Y=k-1) + \\ &+ P(X=2) \cdot P(Y=k-2) + \\ &+ P(X=k) \cdot P(Y=0) \end{aligned}$$

It is clear that the total number of heads can range between 0, 1, 2 up to m plus n , it will take 0 when there is no head in n tosses of X and m tosses of Y it is going to be m plus n if there are all heads in n tosses of X and the m tosses of Y or m tosses of coin 2. Now, X

plus Y is equal to 0 if X is equal to 0, and Y is equal to 0, therefore, probability X plus Y is equal to 0 is equal to probability X is equal to 0 multiplied by probability Y is equal to 0, because these are independent events.

Therefore, we can multiply their probabilities and this is going to be $nC0 p^0 q^n$ multiplied by $mC0 p^0 q^m$ is equal to $nC0$ into $mC0$ into q^{m+n} since, p^0 is equal to 1 is equal to q^{m+n} plus n . since, $nC0$ is equal to 1 and $mC0$ is equal to 1. Let us go one more step. What is the probability X plus Y is equal to 1?

We understand that X plus Y can be 1 in two possible ways. Probability X is equal to 0 multiplied by probability Y is equal to 1 plus probability X is equal to 1 multiplied by probability Y is equal to 0. Because, if we get one head in the total number of tosses, then it will be either there is 0 head from coin 1, but one head from coin 2, or there is 1 head from coin 1 and no heads for coin 2.

So, this probability is going to be $nC0 p^0 q^n$ into $mC1 p^1 q^{m-1}$ plus $mC1 p^1 q^{m-1}$ multiplied by $nC0 p^0 q^n$, which is equal to $nC0 p^0 q^n$ multiplied by $mC1 p^1 q^{m-1}$ plus $nC1 p^1 q^{n-1}$ multiplied by $mC0 p^0 q^m$, q to the power m .

Is equal to $nC0$ into $mC1$ multiplied by p into q^{m+n-1} plus $nC1$ into $mC0$ p into q^{m+n-1} which is equal to $nC0$ is 1 multiplied by m plus $nC1$ is then multiplied by 1 $p q^{m+n-1}$ is equal to m plus n $p q^{m+n-1}$, which we can write it as n plus $nC1 p$ to the power 1 q to the power $m+n-1$.

In general X plus Y is equal to k , this event can occur if there are 0 heads from coin 1, and k heads from coin 2, so we can write it at probability X is equal to 0 into probability Y is equal to k plus probability X is equal to 1 into probability Y is equal to $k-1$ plus probability X is equal to 2 multiplied by probability Y is equal to $k-2$, up to

probability X is equal to k into probability Y is equal to 0. Thus, there are k plus 1 disjoint events which gives rise to the event X plus Y is equal to k .

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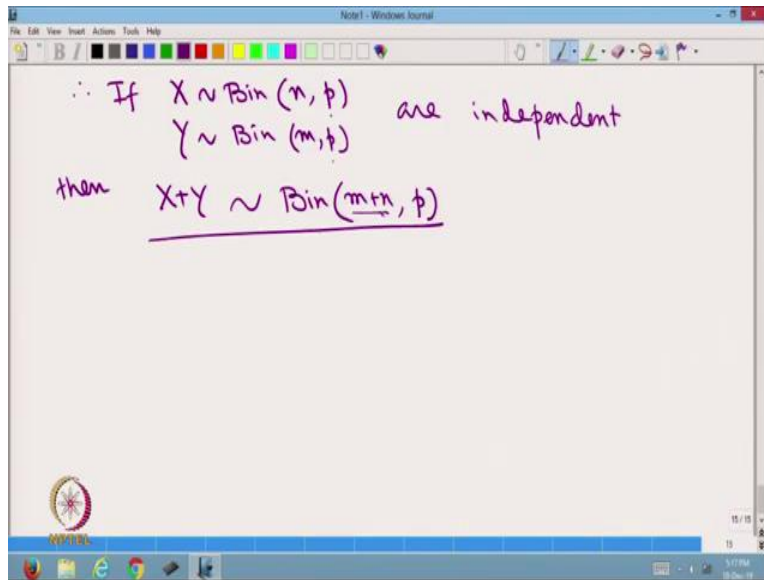
The image shows a handwritten derivation of the probability $P(X+Y=k)$ in a Notepad window. The derivation is as follows:

$$\begin{aligned}
 P(X+Y=k) &= n_0 p^0 q^n * m_{c_k} p^k q^{m-k} \\
 &\quad + n_1 p^1 q^{n-1} * m_{c_{k-1}} p^{k-1} q^{m-k+1} \\
 &\quad + n_2 p^2 q^{n-2} * m_{c_{k-2}} p^{k-2} q^{m-k+2} \\
 &\quad \vdots \\
 &\quad + n_k p^k q^{n-k} * m_{c_0} p^0 q^m \\
 &= p^k q^{m+n-k} (n_{c_0} m_{c_k} + n_{c_1} m_{c_{k-1}} + \dots + n_{c_k} m_{c_0}) \\
 &= \binom{m+n}{k} p^k q^{m+n-k}
 \end{aligned}$$

And therefore, this probability is going to be $n_{c_0} p$ to the power 0 q to the power n multiplied by $m_{c_k} p$ to the power k q to the power m minus k plus $n_{c_1} p$ to the power 1 q to the power n minus 1 multiplied by $m_{c_{k-1}} p$ to the power k minus 1 q to the power m minus k plus 1 plus $n_{c_2} p$ square q to the power n minus 2 multiplied by $m_{c_{k-2}} p$ to the power k minus 2 q to the power m minus k plus 2 up to $n_{c_k} p$ to the power k q to the power n minus k multiplied by $m_{c_0} p$ to the power 0 q to the power m .

Thus, there are k plus 1 terms and if you look at the total power of p is k and total power of q is m plus n minus k , that you can verify for all the k plus 1 terms therefore, we can write it as p to the power k q to the power m plus n minus k multiplied by n_{c_0} into m_{c_k} plus n_{c_1} into $m_{c_{k-1}}$ plus up to n_{c_k} into m_{c_0} . Now, from our high school mathematics we know that this term is $\binom{m+n}{k}$. Therefore, this probability is equal to $\binom{m+n}{k} p^k q^{m+n-k}$.

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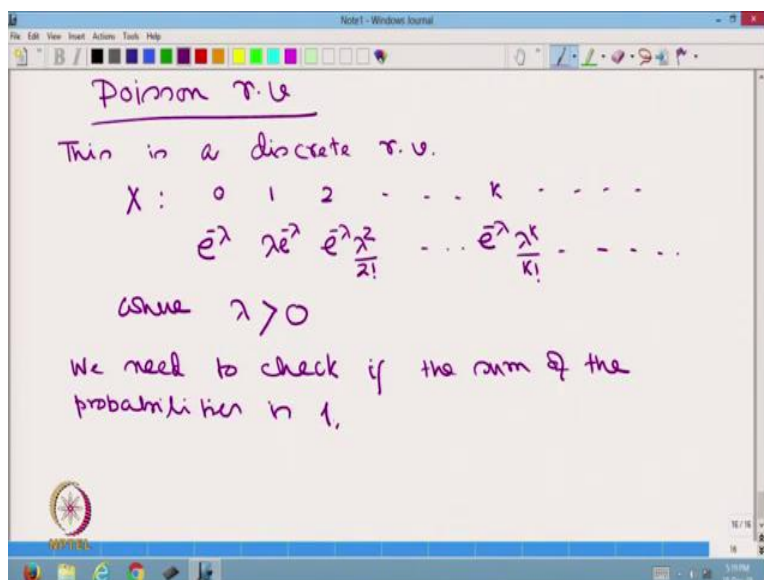


Handwritten note on a digital whiteboard (Notepad - Windows Journal) showing the addition of independent binomial random variables. The text is written in purple ink:

$$\therefore \text{ If } X \sim \text{Bin}(n, p) \text{ are independent} \\ Y \sim \text{Bin}(m, p) \\ \text{then } \underline{X+Y \sim \text{Bin}(m+n, p)}$$

Therefore, what we get? We get that if the X is Binomial n comma p and Y is Binomial m comma p are independent. Then X plus Y follows Binomial m plus n comma p . That is an important result. However, you have to remember that the p has to be the same. The probability of success has to be the same. In that case, the sum of two binomial random variables, if they are independent is going to be a binomial with the first parameter being the sum of their individual values.

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Handwritten note on a digital whiteboard (Notepad - Windows Journal) defining the Poisson distribution. The text is written in purple ink:

Poisson r.v.

This is a discrete r.v.

$$X: 0 \quad 1 \quad 2 \quad \dots \quad k \quad \dots$$
$$e^{-\lambda} \quad \lambda e^{-\lambda} \quad \frac{e^{-\lambda} \lambda^2}{2!} \quad \dots \quad \frac{e^{-\lambda} \lambda^k}{k!} \quad \dots$$

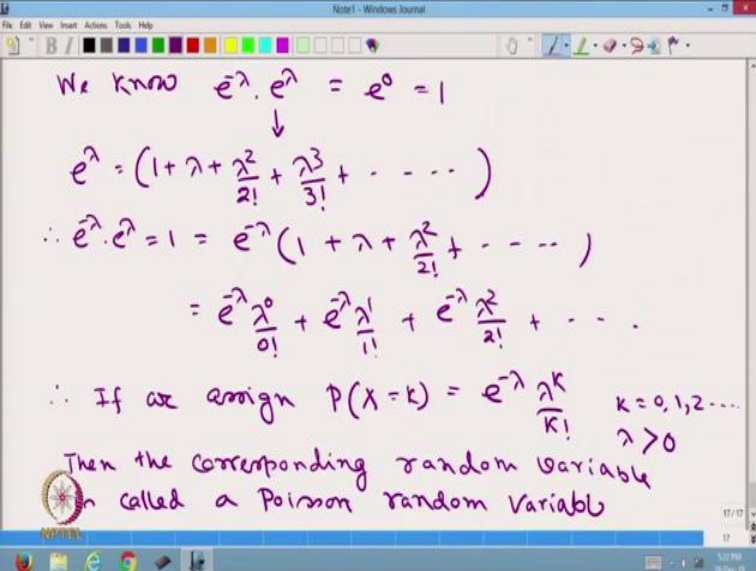
where $\lambda > 0$

We need to check if the sum of the probabilities is 1.

Let us look at another discrete random variable, which is called Poisson random variable. This is a discrete random variable such that X takes values $0, 1, 2, \dots, k$ up to infinity. That means X takes all positive integers or all non-negative integers from 0 to infinity such that probability X is equal to 0 is e to the power minus lambda, probability X is equal to 1 is λe to the power minus lambda, probability X is equal to 2 is e to the power minus lambda λ^2 upon factorial 2 .

Probability X is equal to k is e to the power minus lambda λ^k to the power k upon factorial k . Like that, where lambda is greater than 0 . Since, lambda is greater than 0 , all these individual terms are greater than 0 . Only thing we need to check if the sum of the probability is 1 .

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The image shows a handwritten derivation of the Poisson distribution formula. It starts with the identity $e^{-\lambda} \cdot e^{\lambda} = e^0 = 1$. An arrow points down to the expansion of e^{λ} as a Taylor series: $e^{\lambda} = (1 + \lambda + \frac{\lambda^2}{2!} + \frac{\lambda^3}{3!} + \dots)$. This is then substituted into the identity to get $e^{-\lambda} \cdot e^{\lambda} = 1 = e^{-\lambda} (1 + \lambda + \frac{\lambda^2}{2!} + \dots)$. The next line shows the distribution of terms: $= e^{-\lambda} \frac{\lambda^0}{0!} + e^{-\lambda} \frac{\lambda^1}{1!} + e^{-\lambda} \frac{\lambda^2}{2!} + \dots$. Finally, it concludes that if we assign $P(X=k) = e^{-\lambda} \frac{\lambda^k}{k!}$ for $k=0,1,2,\dots$ and $\lambda > 0$, then the corresponding random variable is called a Poisson random variable.

We know e to the power minus lambda into e to the power lambda is equal to e to the power 0 is equal to 1 . Now, we know that we can expand it to the power lambda as follows e to the power lambda is equal to 1 plus lambda plus lambda square upon factorial 2 plus lambda cube upon factorial 3 up to infinity. Therefore, e to the power minus lambda into e to the power lambda is equal to 1 is equal to e to the power minus lambda into 1 plus lambda plus lambda square upon factorial 2 like that.

Therefore, we are writing it as e to the power minus λ λ^0 upon factorial 0, e to the power minus λ λ^1 upon factorial 1 plus e to the power minus λ λ^2 upon factorial 2 like that. Therefore, if we assign probability X is equal to k is equal to e to the power minus λ λ^k upon factorial k , where k is equal to 0, 1, 2 up to infinity and λ is greater than 0 any real number, then the corresponding random variable is called a Poisson random variable.

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Like Binomial: Suppose $X \sim \text{Poi}(\lambda)$
 $Y \sim \text{Poi}(\mu)$
 question is what is the distribution of $X+Y$.
 Note that $X+Y$ can take values 0, 1, 2, ... ∞

$$P(X+Y=0) = P(X=0) \cdot P(Y=0)$$

$$= e^{-\lambda} \cdot e^{-\mu} = e^{-(\lambda+\mu)}$$

$$P(X+Y=1) = P(X=0)P(Y=1) + P(X=1) \cdot P(Y=0)$$

$$= e^{-\lambda} \cdot e^{-\mu} \mu + e^{-\lambda} \lambda \cdot e^{-\mu}$$

$$= e^{-(\lambda+\mu)} \mu + e^{-(\lambda+\mu)} \lambda = (\lambda+\mu) e^{-(\lambda+\mu)}$$

Natural question is what is the sum of two Poisson random variables? So, like binomial suppose, X is a Poisson random variable with parameter λ and Y is Poisson random variable with say parameter μ , question is what is the distribution of X plus Y ? Note that X plus Y can take values 0, 1, 2 up to infinite. So, probability X plus Y is equal to 0 is equal to probability X is equal to 0 multiplied by probability Y is equal to 0 is equal to e to the power minus λ multiplied by e to the power minus μ , is equal to e to the power minus λ plus μ .

What is probability? X plus Y is equal to 1. We know that the event X plus Y is equal to 1 can happen if X is equal to 0 and Y is equal to 1 or X is equal to 1 and Y is equal to 0. Probability X is equal to 0 into probability Y is equal to 1, because these are independent, we can write as a product plus probability X is equal to 1 multiplied by probability Y is equal to 0.

Is equal to e to the power minus lambda multiplied by e to the power minus mu into mu plus probability X is equal to 1, which is e to the power minus lambda lambda power 1 upon factorial 1 into it to e to the power minus mu, is equal to e to the power minus lambda plus mu multiplied by mu plus e it to the power minus lambda plus mu multiplied by lambda is equal to lambda plus mu into e to the power minus lambda plus mu. I hope now, you are understanding in which direction it is moving. In fact, we are going to move towards that X plus Y is also Poisson random variable with parameters lambda plus mu.

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The image shows a handwritten derivation on a digital whiteboard. The derivation starts with the probability mass function of the sum of two independent Poisson random variables, X and Y , where $X \sim \text{Poisson}(\lambda)$ and $Y \sim \text{Poisson}(\mu)$. The goal is to find $P(X+Y=k)$. The derivation proceeds as follows:

$$\begin{aligned}
 P(X+Y=k) &= P(X=0) \cdot P(Y=k) \\
 &\quad + P(X=1) \cdot P(Y=k-1) \\
 &\quad + P(X=2) \cdot P(Y=k-2) \\
 &\quad \vdots \\
 &\quad + P(X=k) \cdot P(Y=0) \\
 &= e^{-\lambda} \cdot \frac{\lambda^k}{k!} + e^{-\lambda} \lambda \cdot \frac{\lambda^{k-1}}{(k-1)!} + e^{-\lambda} \frac{\lambda^2}{2!} \cdot \frac{\lambda^{k-2}}{(k-2)!} \\
 &\quad \vdots \\
 &\quad + e^{-\lambda} \frac{\lambda^k}{k!} \cdot 1 \\
 &= e^{-(\lambda+\mu)} \left(\frac{\lambda^k}{k!} + \lambda \frac{\lambda^{k-1}}{(k-1)!} + \frac{\lambda^2}{2!} \cdot \frac{\lambda^{k-2}}{(k-2)!} + \dots + \frac{\lambda^k}{k!} \right)
 \end{aligned}$$

So, to verify that let us go one step further then probability X plus Y is equal to k is like the previous example, we are writing probability X is equal to 0 multiplied by probability Y is equal to k plus probability X is equal to 1 multiplied by probability Y is equal to k minus 1 plus probability X is equal to 2 multiplied by probability Y is equal to k minus 2 up to probability X is equal to k multiplied by probability Y is equal to 0.

This is equal to e to the power minus lambda multiplied by e to the power minus mu mu to the power k upon factorial k plus e to the power minus lambda into lambda multiplied by e to the power minus mu mu to the power k minus 1 upon factorial k minus 1 plus e to the power minus lambda lambda square upon factorial 2 multiplied by e to the power minus mu mu to the power k minus 2 upon factorial k minus 2 plus up to e to the power minus lambda lambda to the power k upon factorial k multiplied by e to the minus mu.

Is equal to $e^{-\lambda} \frac{\lambda^k}{k!} + e^{-\lambda} \lambda \frac{\lambda^{k-1}}{(k-1)!} + e^{-\lambda} \frac{\lambda^2}{2!} \frac{\lambda^{k-2}}{(k-2)!} + \dots + e^{-\lambda} \frac{\lambda^k}{k!}$ up to λ^k upon factorial k .

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Handwritten derivation of the probability mass function of the sum of two independent Poisson random variables, X and Y , where $X \sim \text{Poisson}(\lambda)$ and $Y \sim \text{Poisson}(\mu)$. The derivation shows that the sum $X+Y$ follows a Poisson distribution with parameter $\lambda + \mu$.

$$\begin{aligned}
 P(X+Y=k) &= P(X=0) \cdot P(Y=k) \\
 &\quad + P(X=1) \cdot P(Y=k-1) \\
 &\quad + P(X=2) \cdot P(Y=k-2) \\
 &\quad + \dots \\
 &\quad + P(X=k) \cdot P(Y=0) \\
 &= e^{-\lambda} \cdot \frac{\lambda^k}{k!} + e^{-\lambda} \lambda \cdot \frac{\lambda^{k-1}}{(k-1)!} + e^{-\lambda} \frac{\lambda^2}{2!} \frac{\lambda^{k-2}}{(k-2)!} \\
 &\quad + \dots + e^{-\lambda} \frac{\lambda^k}{k!} \cdot e^{-\mu} \\
 &= e^{-(\lambda+\mu)} \left(\frac{\lambda^k}{k!} + \lambda \frac{\lambda^{k-1}}{(k-1)!} + \frac{\lambda^2}{2!} \frac{\lambda^{k-2}}{(k-2)!} + \dots + \frac{\lambda^k}{k!} \right)
 \end{aligned}$$

Handwritten derivation of the probability mass function of the sum of two independent Poisson random variables, X and Y , where $X \sim \text{Poisson}(\lambda)$ and $Y \sim \text{Poisson}(\mu)$. The derivation shows that the sum $X+Y$ follows a Poisson distribution with parameter $\lambda + \mu$.

$$\begin{aligned}
 P(X+Y=k) &= P(X=0) \cdot P(Y=k) \\
 &\quad + P(X=1) \cdot P(Y=k-1) \\
 &\quad + P(X=2) \cdot P(Y=k-2) \\
 &\quad + \dots \\
 &\quad + P(X=k) \cdot P(Y=0) \\
 &= e^{-\lambda} \cdot \frac{\lambda^k}{k!} + e^{-\lambda} \lambda \cdot \frac{\lambda^{k-1}}{(k-1)!} + e^{-\lambda} \frac{\lambda^2}{2!} \frac{\lambda^{k-2}}{(k-2)!} \\
 &\quad + \dots + e^{-\lambda} \frac{\lambda^k}{k!} \cdot e^{-\mu} \\
 &= e^{-(\lambda+\mu)} \left(\frac{\lambda^k}{k!} + \lambda \frac{\lambda^{k-1}}{(k-1)!} + \frac{\lambda^2}{2!} \frac{\lambda^{k-2}}{(k-2)!} + \dots + \frac{\lambda^k}{k!} \right)
 \end{aligned}$$

Now, $e^{-(\lambda+\mu)} \lambda^k$ is equal to $e^{-\lambda} \frac{\lambda^k}{k!} + e^{-\lambda} \lambda \frac{\lambda^{k-1}}{(k-1)!} + e^{-\lambda} \frac{\lambda^2}{2!} \frac{\lambda^{k-2}}{(k-2)!} + \dots + e^{-\lambda} \frac{\lambda^k}{k!}$, is

equal to k factorial into 0 factorial into k factorial λ power 0 μ to the power k plus factorial k factorial upon factorial 1 factorial k minus 1 λ power 1 into μ to the power k minus 1 plus factorial k upon factorial 2 factorial k minus 2 λ square μ to the power k minus 2 up to factorial k upon factorial k into 0 factorial λ power k μ to the power 0 .

Is equal to k factorials into λ power 0 μ to the power k upon k factorial plus λ μ to the power k minus 1 upon 1 factorial into k minus 1 factorial plus λ square μ to the power k minus 2 upon 2 factorial into k minus 2 factorial up to 1 upon k factorial λ power k .

Therefore, if we compared, we see that this term is nothing but the term within the bracket therefore, μ to the power k upon factorial k plus λ μ to the power k minus 1 upon k minus 1 factorial plus λ square μ to the power k minus 2 upon 2 factorial into k minus 2 factorial plus up to λ to the power k upon k factorial this term is nothing but λ plus μ whole to the power k upon factorial k .

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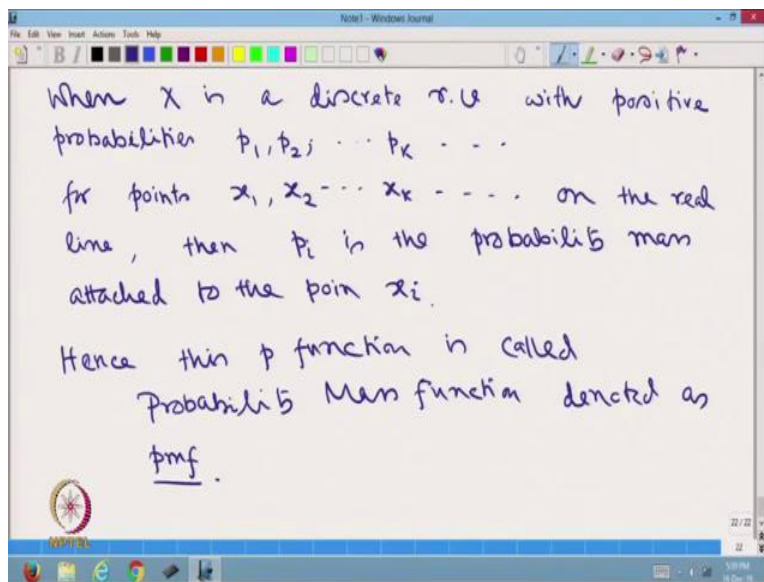
$\therefore P(X+Y=K) = e^{-(\lambda+\mu)} \frac{(\lambda+\mu)^K}{K!}$
 This is true for all $K = 0, 1, 2, \dots$
 \therefore If $\left. \begin{matrix} X \sim \text{Poi}(\lambda) \\ Y \sim \text{Poi}(\mu) \end{matrix} \right\}$ independent
 then $X+Y \sim \text{Poi}(\lambda+\mu)$
 In general if $X_1, X_2, \dots, X_n, \dots$ are independent
 Poisson r.v.s with parameters $\lambda_1, \lambda_2, \dots, \lambda_n, \dots$
 then $\sum_{i=1}^n X_i \sim \text{Poi}(\lambda_1 + \lambda_2 + \dots + \lambda_n)$.

Therefore, if we replace it with λ plus μ to the power k upon k factorial, then we get that probability X plus Y is equal to k is nothing but e to the power minus λ plus μ λ plus μ to the power k upon factorial k . This is true for all k is equal to

0, 1, 2 up to infinity. Therefore, what do you find that if X is Poisson with λ and Y is Poisson with μ are independent then $X + Y$ is Poisson with $\lambda + \mu$.

In general, if X_1, X_2, \dots, X_n are independent Poisson random variables with parameters $\lambda_1, \lambda_2, \dots, \lambda_n$, then $\sum_{i=1}^n X_i$ is distributed as Poisson with $\lambda_1 + \lambda_2 + \dots + \lambda_n$. This is a very interesting result which we shall see later. Because we use Poisson random variables for counting the number of arrivals and we will see that that naturally comes with the summation of the parameters.

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When X is a discrete random variables with positive probabilities that means non-zero probabilities p_1, p_2, \dots, p_k for points x_1, x_2, \dots, x_k on the real line, then p_i is the probability mass attached to the point x_i hence, this function is called Probability Mass Function which we often denote as pmf, for each point the corresponding pmf gives the quantum of probability associated with it.

Okay friends, I stopped here today. So, in this class, we have seen different discrete random variables like Bernoulli, Binomial, and Poisson. In the next class, I shall talk about some more discrete random variables namely, Geometric, Hyper Geometric and Negative Binomial random variables. These are all discrete random variables, which are

very useful for modeling different practical phenomena. Okay friends, thank you so much.