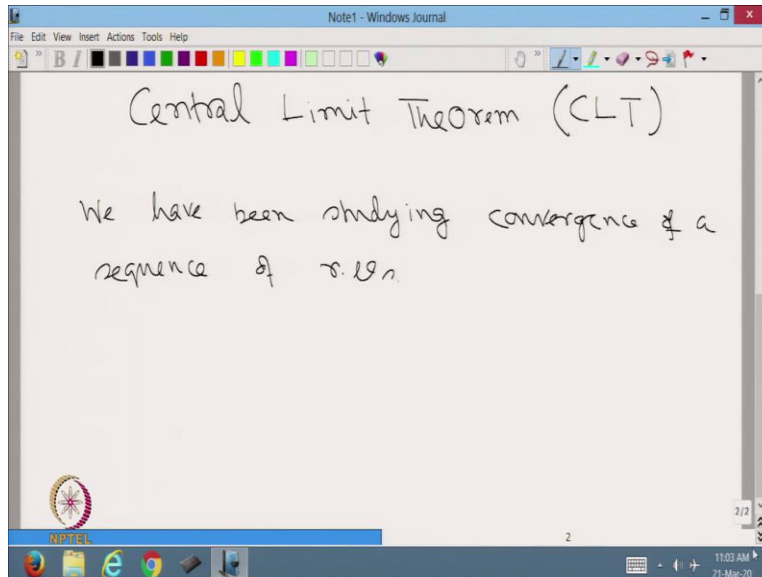


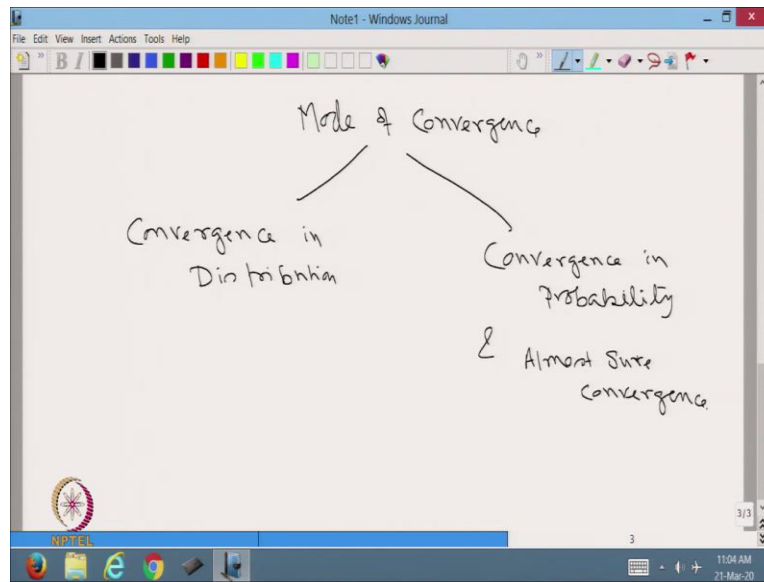
Advanced Probability Theory
Professor Niladri Chatterjee
Department of Mathematics
Indian Institute of Technology Delhi
Lecture 29

(Refer Slide Time: 00:34)



Welcome students to the mock series of lectures on Advanced Probability Theory, this is lecture number 29. As I said at the end of the last class that today we shall be studying Central Limit Theorem, a very important concept of probability which in short we call CLT. Over the last few lectures we have been studying convergence of a sequence of random variables.

(Refer Slide Time: 01:19)



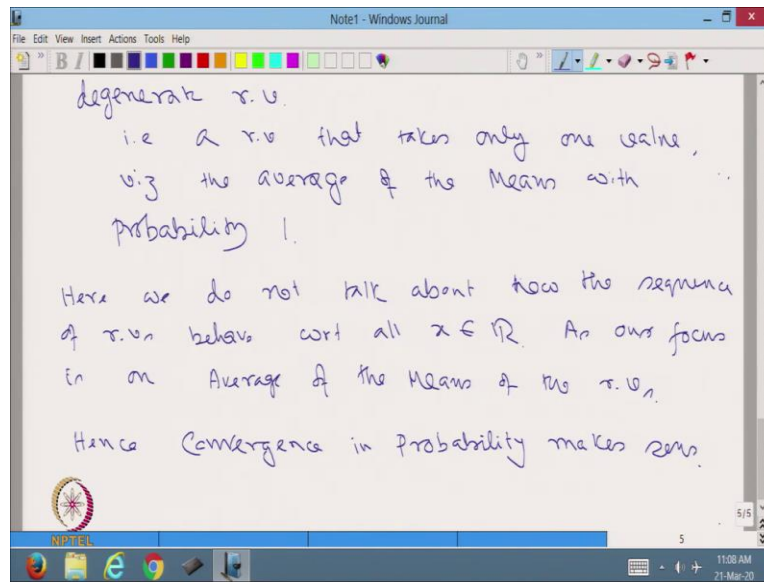
So, first we studied mode of convergence, here we looked at convergence in distribution and also we have looked at convergence in probability and almost sure convergence.

(Refer Slide Time: 02:08)

The handwritten text reads: 'We have Law of Large Numbers. say, WLLN. Here we have checked that the average of a sequence of r.v.s converge in Probability to the average of their Means. If we look at from Convergence towards a r.v. then we find that the average of the sequence of r.v.s converge in Probability to a.' The text is written in black ink, with the last sentence starting in blue ink.

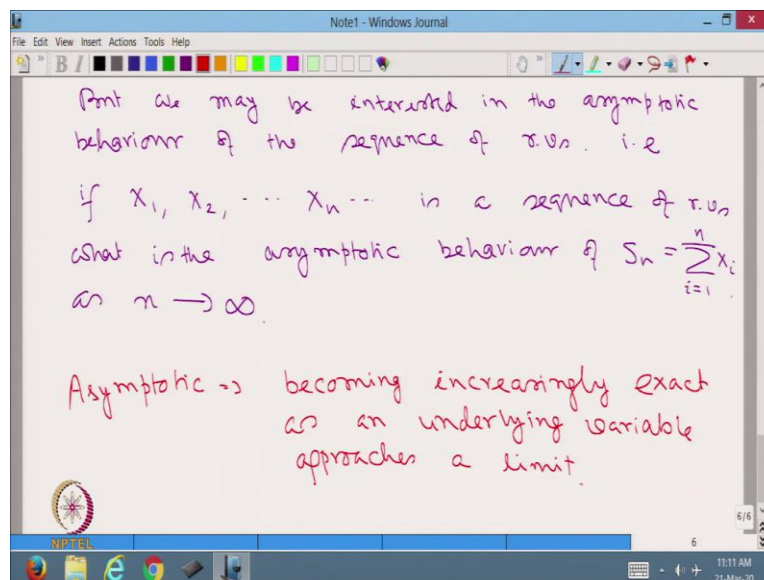
If we look at your application, we have seen laws of large numbers in particular say weak law of large numbers, here we have checked that the average of a sequence of random variables converge in probability to the average of their means. If we look at it from convergence towards a random variables, random variable then we find that the average of the sequence of the random variables converge in probability to a degenerate random variable.

(Refer Slide Time: 4:23)



That is, a random variable that takes only one value namely the average of the means with probability 1. Here we do not talk about how the sequence of random variables behave with respect to all x belonging to \mathbb{R} as our focus is on average of the means of the random variables. Hence convergence in probability makes sense.

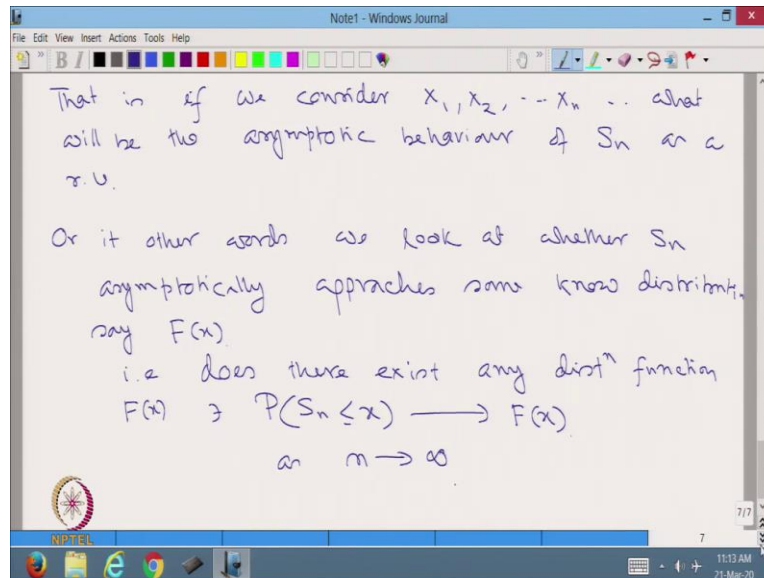
(Refer Slide Time: 06:11)



But we may be interested in the asymptotic behaviour of the sequence of random variables that is, if X_1, X_2, X_n is a sequence of random variable, what is the asymptotic behaviour of S_n is equal to $\sum_{i=1}^n X_i$ as n goes to infinity. Now, some of you may not know the

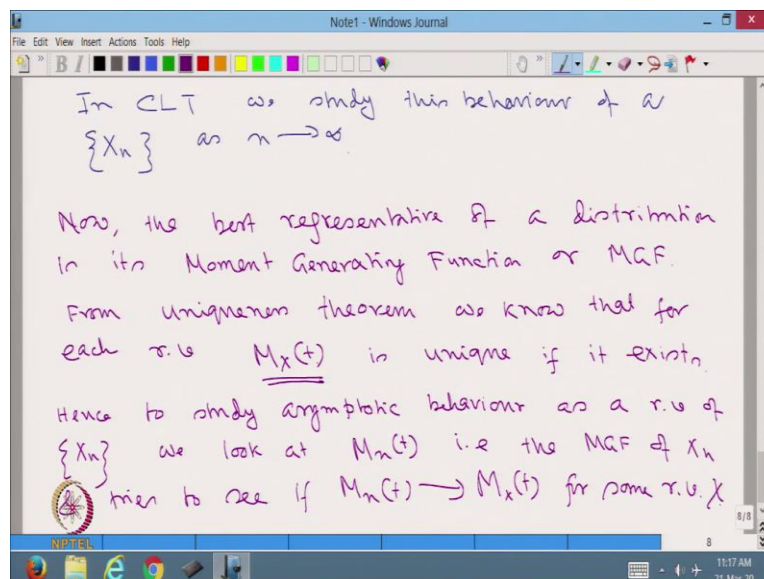
meaning of the word asymptotic, so here I define for you, asymptotic means becoming increasingly close, increasingly exact as an underlying variable approaches a limit.

(Refer Slide Time: 8:25)



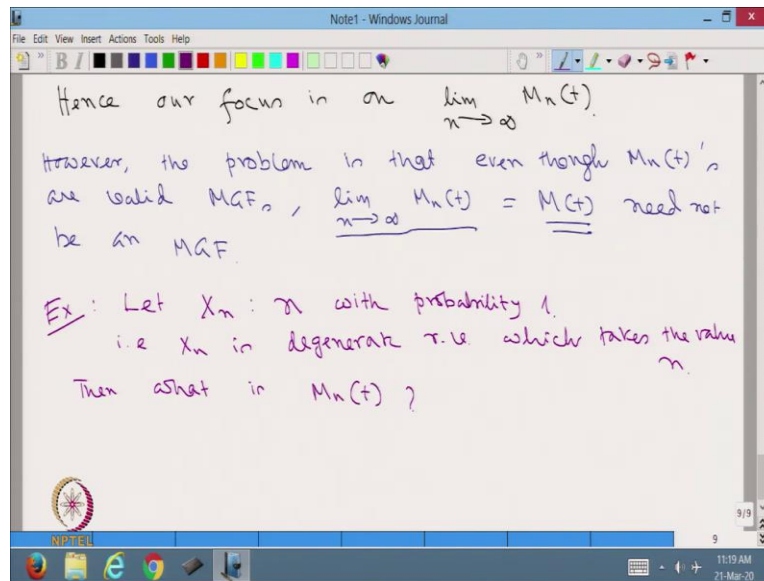
That is if we consider X_1, X_2, X_n , what will be the asymptotic behaviour of S_n as a random variable? Or in other words we look at whether S_n asymptotically approaches some known distribution say F_x that is does there exist any distribution function F_x such that probability S_n less than equal to X converges to F_x as n goes to infinity.

(Refer Slide Time: 10:20)



In central limit theorem we study this behaviour of a sequence of random variable X_n as n goes to infinity. Now, the best representative of a distribution is its moment generating function or MGF, from uniqueness theorem we know that for each random variable X $M_X(t)$ that is moment generating function of X at t is unique if it exists. Hence to study asymptotic behaviour as a random variable of X_n is sequence of random variables we look at $M_{X_n}(t)$ that is the MGF of X_n and tries to see if $M_{X_n}(t)$ converges to $M_X(t)$ for some random variable X .

(Refer Slide Time: 13:07)



Hence, our focus is on limit n going to infinity $M_n(t)$. However, the problem is that even though $M_n(t)$'s are valid moment generating functions limit n going to infinity $M_n(t)$ is equal to $M(t)$ need not be an MGF. That is the sequence of moment generating functions converging to something which is not a moment generating function itself. Example, let X_n with distributed as follows it takes the value n with probability 1, that is X_n is a degenerate random variable, which takes the value n . Then what is $M_n(t)$?

(Refer Slide Time: 15:13)

$M_n(t) = E(e^{tx}) = \begin{cases} e^{tn} & \text{if } t \neq 0 \\ 1 & \text{if } t = 0 \end{cases}$
 \therefore we are looking at $M(t) = \lim_{n \rightarrow \infty} e^{tn}$
 Now if $t > 0$ $e^{tn} \rightarrow \infty$ i.e. does not converge
 if $t < 0$ $e^{tn} \rightarrow 0$
 $\therefore M(t) = \begin{cases} 0 & \text{if } t < 0 \\ 1 & \text{if } t = 0 \\ \infty & \text{if } t > 0 \end{cases}$
 $\therefore M(t)$ cannot be an MGF.
 Since to be an MGF we need $t > 0 \exists$ it converges for $(-t, +t)$.

$M_n(t)$ is equal to expected value of e to the power tx equal to e to the power tn if t not equal to 0 and is equal to 1 if t is equal to 0, therefore we are looking at $M(t)$ is equal to limit n going to infinity e to the power tn . Now, if t is greater than 0, e to the power tn goes to infinity that is does not converge and if t is less than 0, e to the power tn goes to 0. Therefore, $M(t)$ is equal to 0 if t is less than 0, 1 if t is equal to 0 and infinity if t is greater than 0, therefore $M(t)$ cannot be an MGF, since to be an MGF we need t greater than 0 such that it converges for minus t to plus t .

(Refer Slide Time: 17:25)

Hence we look at situations where $\lim_{n \rightarrow \infty} M_n(t)$ exists.
Ex: Let $\{X_n\}$ be a sequence of Binomial r.v.s
 i.e. $X_n \sim \text{Bin}(n, p)$ \exists the expectation np remains constant.
 Does X_n converge to some r.v.?
Ans: Let λ be the constant expectation.
 $\therefore np = \lambda \quad \therefore p = \frac{\lambda}{n} \quad \therefore q = 1 - \frac{\lambda}{n}$
 We know that the MGF of $\text{Bin}(n, p) = (q + pe^t)^n$

Hence, we look at situations when limit n going to infinity $M_n(t)$ exist. Example, let X_n be a sequence of binomial random variables that is X_n is distributed as binomial n comma p such that the expectation np remains constant. Does X_n converge to some random variable? That is the question. So, let λ be the constant expectation, therefore np is equal to λ , therefore p is equal to λ/n , therefore q is equal to $1 - \lambda/n$. We know that the MGF of binomial n, p is equal to $q + pe^{t}$ to the power n .

(Refer Slide Time: 19:49)

∴ In the above scenario :

$$M_n(t) = \left(1 - \frac{\lambda}{n} + \frac{\lambda}{n} e^t\right)^n = \left(1 - \frac{\lambda}{n}(1 - e^t)\right)^n$$

We know that $\lim_{n \rightarrow \infty} \left(1 - \frac{x}{n}\right)^n \rightarrow e^{-x}$.

∴ $M_n(t) \xrightarrow{\lim n \rightarrow \infty} e^{-\lambda(1 - e^t)}$

↑ This is the MGF of Poisson r.v. with parameter $= \lambda$.

Therefore, in the above scenario $M_n(t)$ is equal to $1 - \frac{\lambda}{n}$ that is q plus $\frac{\lambda}{n}$ by n e^t to the power t whole to the power n is equal to $1 - \frac{\lambda}{n}$ into $1 - e^t$ to the power t whole to the power n , we know that limit n going to infinity $1 - \frac{x}{n}$ whole to the power n converges to e^{-x} , therefore this $M_n(t)$ converges to limit n going to infinity $e^{-\lambda(1 - e^t)}$, this is the MGF of Poisson random variable with parameter is equal to λ .

(Refer Slide Time: 21:47)

Therefore, in the above case with the restriction on $np = \lambda$ we see the sequence of r.v.s $\{X_n\}$

$\rightarrow \text{Poi}(\lambda)$

Therefore, in the above case with the restriction on np is equal to λ we see the sequence of random variables X_n converges to Poisson with λ .

(Refer Slide Time: 22:30)

EX Suppose $X_1, X_2, \dots, X_n, \dots$ is a sequence of r.v.s s.t. $X_n \sim \text{Poi}(n)$.

Question is does X_n converge to some known distⁿ?

Ans: Let us consider $\{Y_n\}$ to be a sequence of r.v.s $\Rightarrow Y_n = \frac{X_n - n}{\sqrt{n}}$

We know that if $X \sim \text{Poi}(n)$ then $E(X) = \text{Var}(X) = n$

$\therefore \frac{X - n}{\sqrt{n}}$ is the Normalized form of X .

Let us, now consider another example, suppose X_1, X_2, X_n is a sequence of random variables such that X_n is distributed as Poisson with parameter n , question is, does X_n converge to some known distribution? Answer, let us consider Y_n to be a sequence of random variables such that Y_n is equal to X_n minus n over root n . We know that if X is distributed as Poisson with λ , then expectation of X is equal to variance of X is equal to λ . Therefore, x minus λ over upon root over λ is the normalized form of X . Because it is variable minus mean divided by standard deviation.

(Refer Slide Time: 24:48)

$\therefore Y_n = \frac{X_n - n}{\sqrt{n}}$ is standardized X_n .
 What is the MGF of Y_n ?
 We know that if X has the MGF $M_X(t)$, then for constants a, b the MGF of $M_{ax+b}(t) = e^{bt} \cdot M_X(at)$.
 \therefore MGF $Y_n = M_{\frac{X_n - n}{\sqrt{n}}}(t)$ Note that:
 \therefore MGF $Y_n = e^{-t\sqrt{n}} \cdot M_{X_n}\left(\frac{t}{\sqrt{n}}\right)$ $b = -\sqrt{n}$ & $a = \frac{1}{\sqrt{n}}$.

Therefore, Y_n which is equal to X_n minus n over root over n is standardized or normalized X_n . What is the MGF of Y_n ? We know that, if X has the moment generating function $M_X t$, then for constants a and b the MGF of ax plus b at t is equal to e to the power bt into MG of X at the point $a t$, therefore MGF of Y_n is equal to moment generating function of X_n minus n upon root n at t , note that b is equal to minus root n and a is equal to 1 upon root n , therefore MGF of Y_n is equal to e to the power minus t root n into MGF of X_n at t upon root n .

(Refer Slide Time: 27:20)

Now MGF of $Poi(\lambda) = e^{\lambda(1-e^{-t})}$ at t .
 \therefore MGF of X_n at $\frac{t}{\sqrt{n}} = e^{-n(1-e^{t/\sqrt{n}})}$
 \therefore Here $\lambda = n$.
 $= \text{Exp}\left(-n(1 - e^{t/\sqrt{n}})\right)$
 $= \text{Exp}\left(-n\left(x - \left(x + \frac{t}{\sqrt{n}} + \frac{t^2}{2n} + \frac{t^3}{3!n^{3/2}} + \dots\right)\right)\right)$
 $= \text{Exp}\left(-n\left(-\frac{t}{\sqrt{n}} - \frac{t^2}{2n} - \frac{t^3}{6n^{3/2}} - \dots\right)\right)$ Plus terms having n^2 or higher of n in the denominator.
 $= \text{Exp}\left(\frac{t\sqrt{n}}{1} + \frac{t^2}{2} + \frac{t^3}{6\sqrt{n}} + \dots\right)$ terms with n or its higher power in the denominator.

Now, MGF of Poisson random variable with parameter lambda is equal to $e^{\lambda(e^t - 1)}$. Therefore, MGF of X_n at t/\sqrt{n} is equal to $e^{\lambda(e^{t/\sqrt{n}} - 1)}$, since here lambda is equal to n, then is equal to $e^{n(e^{t/\sqrt{n}} - 1)}$. I am writing this as exponential function minus n into $1 - e^{t/\sqrt{n}}$. Now let us expand $e^{t/\sqrt{n}}$ to the power t upon root n, this is equal to $1 + t/\sqrt{n} + \frac{t^2}{2n} + \frac{t^3}{6n^{3/2}} + \dots$ upon factorial 3 n to the power 3 by 2 plus term having n square or higher power of n in the denominator is equal to exponentiation of minus n into this one cancels with this, therefore minus t by root n minus t square upon 2n minus t cube upon 6 n to the power 3 by 2 minus terms with higher power of n in the denominator is equal to exponentiation of t root n plus t square upon 2 plus t cube upon 6 square root of n plus terms with n or its higher power in the denominator.

(Refer Slide Time: 30:46)

The image shows a handwritten derivation in a Windows Journal window titled "Note1 - Windows Journal". The derivation is as follows:

$$\begin{aligned} \therefore \text{MGF of } Y_n &= e^{-\lambda} \cdot E\left(\lambda^n + \frac{\lambda^2}{2} + \frac{\lambda^3}{6!n^{3/2}} + \dots\right) \\ &= e^{-\lambda} \left(\lambda^n + \lambda^n + \frac{\lambda^2}{2} + \frac{\lambda^3}{6!n^{3/2}} + \dots \right) \\ &= e^{-\lambda} \left(\lambda^n + \frac{\lambda^2}{2} + \dots \right) \\ \therefore \lim_{n \rightarrow \infty} M_{Y_n}(t) &= e^{t^2/2} \\ \therefore \text{The MGF of } Y_n(t) &\rightarrow e^{t^2/2} \text{ which is the MGF of } N(0, 1) \end{aligned}$$

The derivation is written in purple ink on a light blue background. The final result shows that the MGF of $Y_n(t)$ converges to $e^{t^2/2}$, which is the MGF of a standard normal distribution $N(0, 1)$.

Note1 - Windows Journal

Now MGF of $Poi(\lambda) = e^{-\lambda(1-e^t)}$ at t .

\therefore MGF of X_n at $\frac{t}{\sqrt{n}} = e^{-n(1-e^{\frac{t}{\sqrt{n}}})}$

\therefore Here $\lambda = n$.

$= \text{Exp}\left(-n(1-e^{\frac{t}{\sqrt{n}}})\right)$

$= \text{Exp}\left(-n\left(1 - \left(1 + \frac{t}{\sqrt{n}} + \frac{t^2}{2n} + \frac{t^3}{3!n^{3/2}} + \dots\right)\right)\right)$

$= \text{Exp}\left(-n\left(-\frac{t}{\sqrt{n}} - \frac{t^2}{2n} - \frac{t^3}{6n^{3/2}} - \dots\right)\right)$ Plus terms having n^2 & higher of n in the denominator.

$= \text{Exp}\left(\frac{t\sqrt{n}}{1} + \frac{t^2}{2} + \frac{t^3}{6\sqrt{n}} + \dots\right)$ terms with n or its higher power in the denominator.

Therefore, MGF of Y_n is equal to e to the power minus root n into t multiplied by exponentiation of this term that is t root n plus t square upon 2 plus t cube upon 6 into n to the power half plus terms with higher power of n in the denominator, is equal to e to the power minus root n t plus root n t plus t square by 2 plus t cube upon 6 n to the power half plus other terms is equal to e to the power t square by 2 plus t cube into 6 this is 6 equal to factorial 3, n to the power half plus other terms. Therefore, limit n going to infinity M_{Y_n} at t is equal to e to the power t square by 2, because all these terms are going to 0 as n goes to infinity. Therefore, the MGF of Y_n t converges to e to the power t square by 2, which is the MGF of standard normal with parameter 0, 1.

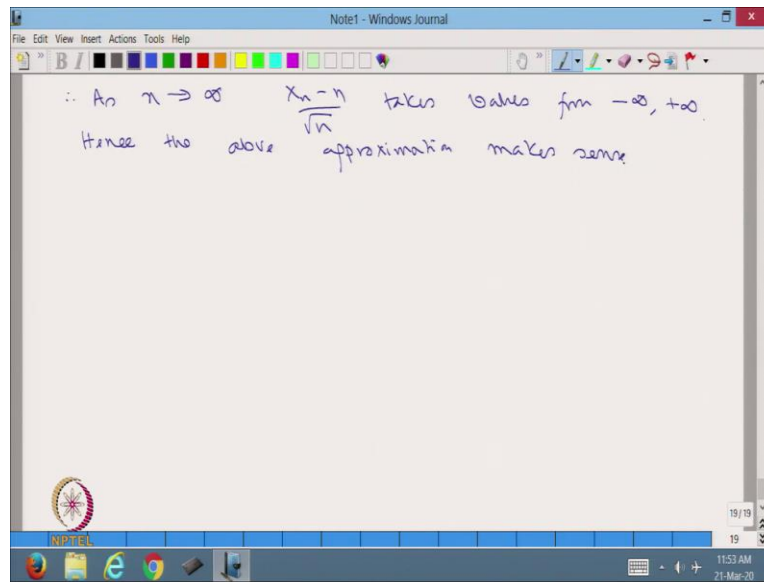
(Refer Slide Time: 33:07)

\therefore If X_n is a sequence of r.v.s st $X_n \sim \text{Poi}(n)$
then as $n \rightarrow \infty$ $X_n \xrightarrow{L} N(0,1)$

Note that Poisson is +ve r.v.
Whereas $N(0,1)$ takes values in $(-\infty, \infty)$
Hence we need to standardize X_n with $\frac{X-n}{\sqrt{n}}$
So what happens in the following
 $X: 0, 1, 2, \dots$
 $\therefore \frac{X-n}{\sqrt{n}}$ takes values from $-\frac{n}{\sqrt{n}}$ to ∞
when $X=0$

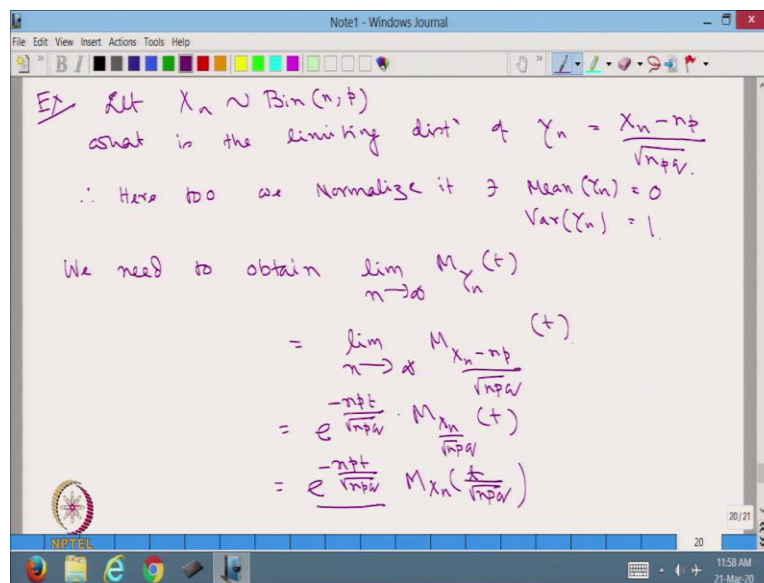
Therefore, if X_n is a sequence of random variables, such that X_n is distributed as Poisson with parameter n then as n goes to infinity X_n converges in distribution to normal $0, 1$. Note that, Poisson is a positive random variable, whereas normal $0, 1$ takes values in minus infinity to plus infinity. Hence, we need to standardize or normalize X_n with X_n minus n upon root over n . So, what happens is the following, X takes values $0, 1, 2$, etcetera, therefore X minus n upon root n takes values from minus n upon root n when X is equal to 0 to infinite.

(Refer Slide Time: 35:20)



Therefore, as n goes to infinity X_n minus np upon root n takes values from minus infinity to infinity, hence the above approximation makes sense.

(Refer Slide Time: 35:58)



Another example, let X_n is distributed as binomial n, p , what is the limiting distribution of Y_n is equal to X_n minus np upon root over npq . Therefore, here too we normalize it such that mean of Y_n is equal to 0 and variance of Y_n is equal 1. Therefore, we need to obtain limit n going to infinity M_{Y_n} at a point t is equal to limit n going to infinity $M_{X_n - np}$ upon root of over npq at the point t is equal to e to the power minus npt upon root over npq into M_{X_n} upon root

over npq at the point t is equal to e to the power minus npt upon root over npq into M_{Xn} at t upon root over npq .

(Refer Slide Time: 38:12)

The image shows a handwritten derivation of the Moment Generating Function (MGF) for a binomial random variable Y_n in a Windows Journal window. The derivation is as follows:

$$\begin{aligned}
 &\text{We know that the MGF for a Binomial r.v.} \\
 &= (q + pe^t)^n \\
 \therefore \text{MGF}_{Y_n} &= e^{-\frac{npt}{\sqrt{npq}}} (q + pe^{\frac{t}{\sqrt{npq}}})^n \\
 &= \left(e^{-\frac{pt}{\sqrt{npq}}} (q + pe^{\frac{t}{\sqrt{npq}}}) \right)^n \\
 &= \left(qe^{-\frac{pt}{\sqrt{npq}}} + p \cdot e^{-\frac{pt}{\sqrt{npq}} + \frac{t}{\sqrt{npq}}} \right)^n \\
 &= \left(qe^{-\frac{pt}{\sqrt{npq}}} + p e^{\frac{t(1-p)}{\sqrt{npq}}} \right)^n \\
 &= \left(qe^{-\frac{pt}{\sqrt{npq}}} + p e^{\frac{qt}{\sqrt{npq}}} \right)^n
 \end{aligned}$$

We know that, the moment generating function for a binomial random variable is equal to q plus $p e$ to the power t whole to the power n . Therefore, MGF of Y_n is equal to e to the power minus npt root over npq multiplied by q plus $p e$ to the power t upon root over npq whole to the power n , is equal to e to the power minus pt root over npq multiplied by q plus $p e$ to the power t upon root over npq whole to the power n is equal to $q e$ the power minus pt root over npq plus p times e to the power minus pt root over npq plus t root over npq whole to the power n , is equal to q times e to the power minus pt upon root over npq plus p times e to the power t into 1 minus p upon root over npq whole to the power n , is equal to q times e to the power minus pt root over npq plus p times e to the power qt root over npq whole to the power n .

(Refer Slide Time: 40:43)

$$\therefore \text{Question is what is the limit of } \left(q e^{-\frac{pt}{\sqrt{npq}}} + p e^{\frac{qt}{\sqrt{npq}}} \right)^n ?$$

$$= \left(q \cdot \left(1 - \frac{pt}{\sqrt{npq}} + \frac{p^2 t^2}{2! npq} - \frac{p^3 t^3}{3! n^{3/2} (pq)^{3/2}} + \dots \right) + p \cdot \left(1 + \frac{qt}{\sqrt{npq}} + \frac{q^2 t^2}{2! npq} + \frac{q^3 t^3}{3! n^{3/2} (pq)^{3/2}} + \dots \right) \right)^n$$

$$= \left((q+p) + \frac{qp^2 t^2 + pq^2 t^2}{2 npq} + \text{terms with } n^{3/2} \text{ or higher powers of } n \text{ in the denominator} \right)^n$$

$$= \left(1 + \frac{t^2}{2n} + \dots \right)^n$$

Therefore, question is, what is the limit of q times e to the power minus pt root over npq plus p times e to the power qt root over npq whole to the power n . So we do this by expanding this exponential terms, therefore we are writing it as q times 1 minus pt upon root over npq plus p square t square upon 2 into npq minus p cube t cube upon n to the power 3 by 2 pq to the power 3 by 2 into factorial 3 plus terms with n square or higher power in denominator plus p times now we are expanding this, 1 plus qt upon root over npq plus q square t square upon factorial 2 into npq plus q cube t cube upon factorial 3 n to the power 3 by 2 pq to the power 3 by 2 plus like as before higher powers of n square in the denominator whole to the power n is equal to, now we are adding so we get q plus p minus qpt plus qpt , therefore they get cancelled plus qp square t square plus pqp square t square upon $2npq$ plus terms with n to the power 3 by 2 or higher powers of n in the denominator, is equal to 1 plus if we take pq common it is t square upon 2 and that pq cancels with this plus n plus other terms whole to the power n .

(Refer Slide Time: 44:22)

The image shows a handwritten derivation in a Windows Journal window. The text is as follows:

$$\therefore M_{Y_n}(t) = \left(1 + \frac{t^2}{2n} + \text{terms with } n^{3/2} \text{ or higher powers on } n \text{ in the denominator}\right)^n$$

$$\therefore \log M_{Y_n}(t) = n \log \left(1 + \frac{t^2}{2n} + \dots\right)$$

$$= n \left(\frac{t^2}{2n} + \text{terms with higher powers of } n \text{ in denominator}\right)$$

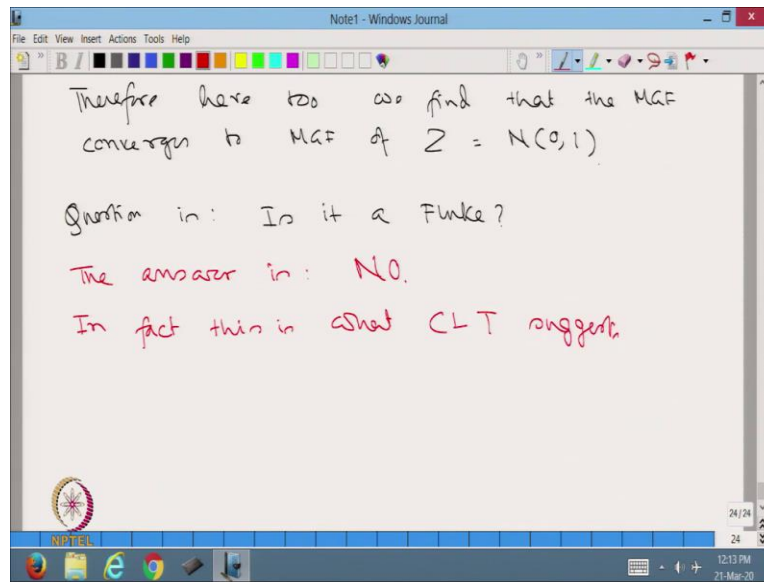
$$= \frac{t^2}{2} + \text{terms with power of } n \text{ in the denominator.}$$

$$\therefore \lim_{n \rightarrow \infty} \log M_{Y_n}(t) = \frac{t^2}{2} \quad \text{as other terms go to 0}$$

$$\Rightarrow \lim_{n \rightarrow \infty} M_{Y_n}(t) = e^{\frac{t^2}{2}} \quad \text{which is the MGF of } N(0, 1)$$

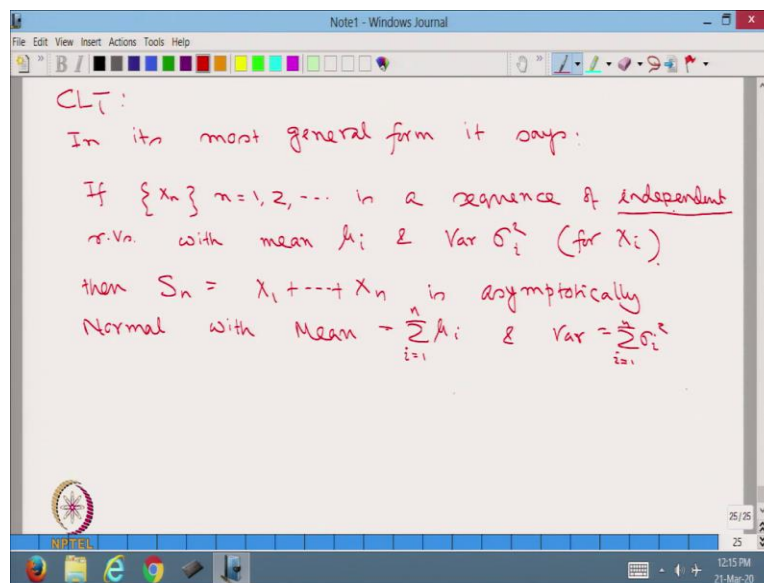
Therefore, moment generating function of Y_n is equal to what we get from here is 1 plus t square upon $2n$ plus terms with n to the power 3 by 2 or higher powers of n in the denominator whole to the power n . Therefore, M_{Y_n} at t , here also there should have been a t is equal to n times log of 1 plus t square upon $2n$ plus such terms is equal to n times t square upon 2 by n plus terms with higher powers of n in denominator is equal to t square 2 plus terms with power of n in the denominator. Therefore, limit n going to infinity log of M_{Y_n} at t is equal to t square upon 2 as other terms go to 0, implies limit n going to infinity M_{Y_n} to the power of t is equal to e to the power t square by 2 which is the MGF of standard normal 0, 1.

(Refer Slide Time: 47:04)



Therefore, here too we find that the MGF converges to MGF of Z which is equal to normal 0, 1. Question is, is it a fluke? The answer is, no, in fact this is what central limit theorem suggest.

(Refer Slide Time: 48:06)



So, what is central limit theorem? In its most general form, it says if X_n n is equal to 1, 2 etcetera is a sequence of independent random variables with mean μ_i and variance σ_i^2 for X_i then S_n is equal to X_1 plus X_2 up to X_n is asymptotically normal with mean is equal to $\sum_{i=1}^n \mu_i$ and variance is equal to $\sum_{i=1}^n \sigma_i^2$ is equal to 1 to n , here also i is equal to 1 to n . So, this is the most general form when the only assumption is that X_1, X_2, X_n are

independent. In this class we shall not prove this, what we should do that if they are not only independent, but identically distributed then same result holds. Okay students, I stop here today, in the next class I shall start with this statement. Thank you.