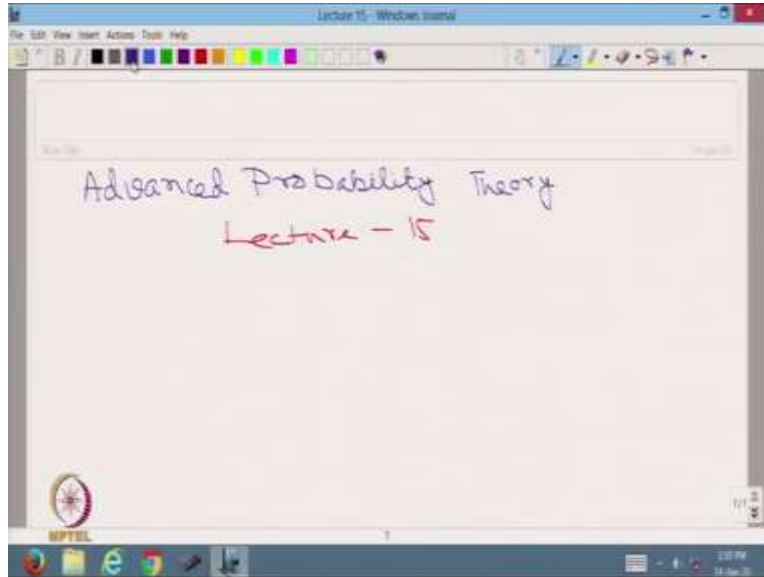


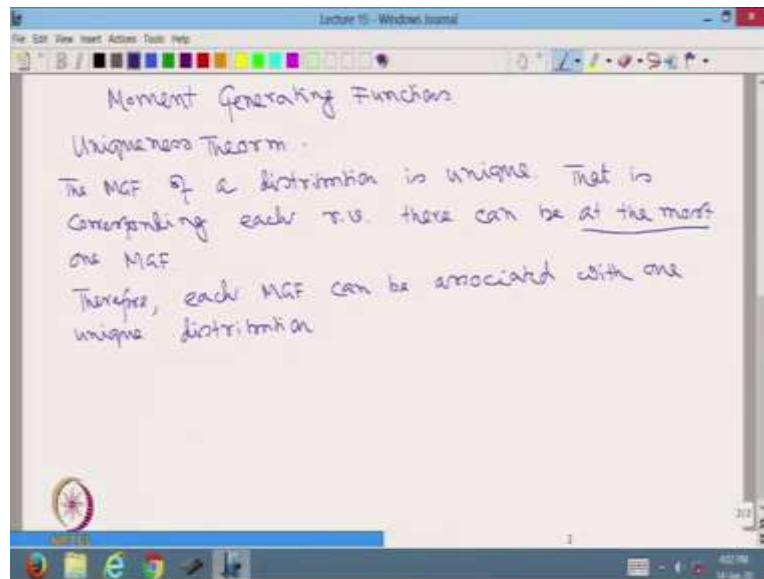
Advanced Probability Theory
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Lecture 15

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Welcome students to the MOOC lecture series on Advanced Probability Theory, this is lecture number 15.

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In the last class, we have seen Moment Generating Functions and some properties of it. These are important because of the uniqueness theorem which says that the moment generating function of a distribution is unique. That is corresponding to each random variable there can be at the most one MGF. I am saying at the most, because there may be distributions, which do not have the moment generating function. Therefore each moment generating function can be associated with one unique distribution.

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The advantage of the Uniqueness Theorem of the MGF is that if we identify the pattern of the MGF for some variable X , then we can identify its distribution.

Ex: $X \sim \text{Bin}(n, p) \quad \therefore \text{MGF}_X(t) = (q + pe^t)^n$
 $Y \sim \text{Bin}(m, p) \quad \therefore \text{MGF}_Y(t) = (q + pe^t)^m$

Now suppose X & Y are independent.
 $\therefore \text{MGF of } Z = X + Y = \text{MGF}_Z(t) = \frac{(q + pe^t)^{m+n}}{\text{MGF Bin}(m+n, p)}$

Therefore we can see that the sum of two independent $\text{Bin}(n, p) + \text{Bin}(m, p)$ is distributed as $\text{Bin}(n+m, p)$.

This is one result we have proved earlier. Now we can find very easily.

The advantage of the uniqueness theorem of the moment generating function is that if we identify the pattern of the moment generating function for some variable x , then we can identify its distribution. For example, suppose x is binomial n comma p . Therefore MGF of x at t is equal to q plus p , e to the power t , whole to the power n . Suppose y is distributed as binomial m p , therefore MGF, MGF of y t is equal to q plus p e to the power t whole to the power m .

Now, suppose x and y are independent therefore MGF of z is equal to x plus y is equal to, we know that it is the product of their individual moment generating functions that is q plus p e to

the power t whole to the power m plus n which is the MGF of binomial m plus n comma p . Therefore, we can see that this sum of 2 independent binomial n comma p and binomial m comma p is distributed as binomial n plus m comma p , this is one result we have proved earlier. But now, we can find very easily.

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In a similar way:

$$X \sim N(\mu_1, \sigma_1^2) \quad \therefore MGF_X(t) = e^{t\mu_1 + \frac{1}{2}\sigma_1^2 t^2}$$

$$Y \sim N(\mu_2, \sigma_2^2) \quad \therefore MGF_Y(t) = e^{t\mu_2 + \frac{1}{2}\sigma_2^2 t^2}$$

\therefore If X & Y are independent:

$$\text{then } MGF_{X+Y}(t) = e^{(\mu_1 + \mu_2)t + \frac{1}{2}(\sigma_1^2 + \sigma_2^2)t^2}$$

i.e. $X+Y \sim N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$

again a result earlier we proved by convolution, but now get very easily

In a similar way, suppose x is normal with μ_1 sigma 1 square therefore its MGF at t is equal to e to the power $\mu_1 t$ plus half sigma 1 square t square suppose y is normal μ_2 comma sigma 2 square. Therefore MGF of y is equal to e to the power $\mu_2 t$ plus half sigma 2 square t square. Therefore, if x and y are independent then MGF of x plus y at t is equal to their product and it is equal to e to the power μ_1 plus $\mu_2 t$ plus half sigma 1 square plus sigma 2 square into t square that is x plus y is distributed as normal with μ_1 plus μ_2 and variance is equal to sigma 1 square plus sigma 2 square. Again a result, earlier we proved by convolution, but now get very easily.

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Let us now focus on some deficiencies.

Ex. Consider $p(x)$ to be pmf of the following form:

$$p(x) = \frac{1}{x(x+1)} \quad x=1, 2, 3, \dots$$

$$\sum_{x=1}^{\infty} p(x) = \frac{1}{1 \cdot 2} + \frac{1}{2 \cdot 3} + \frac{1}{3 \cdot 4} + \dots$$

$$= 1 - \frac{1}{2} + \frac{1}{2} - \frac{1}{3} + \frac{1}{3} - \frac{1}{4} + \dots$$

$$= 1 \quad \text{as the sum} \rightarrow \infty$$

\therefore This is a valid pmf.

\therefore What is $E(X)$?

$$= \sum_{x=1}^{\infty} x \cdot p(x) = \sum_{x=1}^{\infty} \frac{x}{x(x+1)} = \frac{1}{2} + \frac{1}{3} + \dots = \sum_{i=2}^{\infty} \frac{1}{i}$$

\therefore We know that this is a divergent series.

Hence its first moment does not exist.

Let us now focus on some deficiencies for example, consider p_x to be a PMF of the following form p_x is equal to $\frac{1}{x(x+1)}$ upon x is equal to 1, 2, 3. Therefore, $\sum_{x=1}^{\infty} p_x$ is equal to 1 to infinity is equal to, which is equal to $1 - \frac{1}{2} + \frac{1}{2} - \frac{1}{3} + \frac{1}{3} - \frac{1}{4} + \dots$ and if we notice that alternating terms keep on canceling. Therefore, this is going to be 1 as the sum goes to infinity, therefore this is a valid PMF.

Therefore, what is expected value of x ? This is equal to $\sum_{x=1}^{\infty} x \cdot p_x$ upon x is equal to 1 to infinity is equal to $\sum_{x=1}^{\infty} \frac{x}{x(x+1)}$ is equal to $\sum_{i=2}^{\infty} \frac{1}{i}$ is equal to half plus

one third up to infinity is equal to sigma 1 over i, i is equal to 2 to infinity and we know that this is a divergent series. Hence, its first moment does not exist.

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The image shows a handwritten derivation on a digital whiteboard. At the top, it defines a random variable X with values $2^0, 2^1, 2^2, \dots, 2^n, \dots$. An arrow points from 2^n down to $\frac{e^{-1}}{n!}$. Below this, the probability mass function is calculated as a sum from $x=0$ to ∞ of $P(X=x)$, which is $e^{-1} \left(\frac{1}{0!} + \frac{1}{1!} + \frac{1}{2!} + \dots \right) \frac{e^{-1}}{n!}$. This simplifies to $e^{-1} \cdot e = 1$, with a note: "It is a valid pdf". Then, the moment generating function (MGF) is given as $\sum_{x=0}^{\infty} \frac{e^{-1}}{x!} e^{tx}$. A note states: "This sum does not converge for any $t > 0$ ". Finally, it concludes: "MGF does not exist".

Another example, x is a random variable, which takes values 2 to the power 0, 2 to the power 1, 2 to the power 2, etc with probability of 2 to the power n is equal to e to the power minus 1 into n factorial. Therefore, sigma $p x$, x is equal to 0 to infinity is equal to e to the power minus 1 into 0 factorial plus 1 upon 1 factorial plus 1 upon 2 factorial plus etc is equal to e to the power minus 1 into e is equal to 1. Therefore, it is a valid PDF.

Therefore, its MGF is equal to sigma e to the power minus 1 x factorial into e to the power t into 2 to the power x , x is equal to 0 to infinity. This sum does not converge for any t greater than 0 therefore, moment generating function does not exist. Like that, we can show many examples, where there are problems with respect to MGF. However MGF is important because of its uniqueness.

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Hence a natural extension is to compute $E(e^{itX})$
 i.e. $E(\cos tX + i \sin tX)$
 $\therefore e^{itX}$ is a complex number $\therefore |e^{itX}| = 1$
 Hence the series obtained from $E(e^{itX})$
 i.e. $E\left(1 + itX + \frac{(it)^2}{2!}X^2 + \frac{(it)^3}{3!}X^3 + \dots\right)$
 $= 1 + itE(X) - \frac{t^2}{2!}E(X^2) - \frac{it^3}{3!}E(X^3) + \dots$
 This always converges.
 Then $E(e^{itX})$ for a real no t is called the characteristic
 function of a r.v X often written as $\phi_X(t)$

Hence a natural extension is to compute expected value of e to the power itx that is expected value of $\cos t x$ plus i sine $t x$. Since e to the power itx is a complex number. Therefore, modulus of e to the power itx is equal to 1 hence this series obtained from expected value of e to the power itx , that is expected value of 1 plus itx plus $i t$ whole square upon factorial 2 x square plus $i t$ whole cube upon factorial 3 x cube which is equal to 1 plus $i t$ times expected value of x minus t square upon factorial 2 expected value of x square minus t cube upon factorial 3 x cube.

Like that, we get an infinite series, this always converges, this expectation of e to the power itx for a real number t is called the characteristic function of a random variable x which we often write as, as ϕ_x .

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Characteristic function of different r.v.s are as follows:

- Binomial (n, p) : $(q + pe^{it})^n$
- Poisson (λ) : $\sum_{x=0}^{\infty} e^{-\lambda} \frac{\lambda^x}{x!} e^{itx} = e^{-\lambda} \sum_{x=0}^{\infty} \frac{(\lambda e^{it})^x}{x!} = e^{-\lambda} e^{\lambda e^{it}} = e^{\lambda(e^{it} - 1)}$
- Exponential (λ) : $\int_0^{\infty} e^{itx} \lambda e^{-\lambda x} dx = \lambda \int_0^{\infty} e^{-(\lambda - it)x} dx = \frac{\lambda}{\lambda - it}$
- Normal (μ, σ^2) : $\mu t + \frac{1}{2}(\sigma^2 t^2) = \mu t + \frac{1}{2}\sigma^2 t^2$

Since we have already computed moment generating function for different random variables we can easily find that characteristic functions of different random variables as follows binomial n comma p it is going to be q plus p e to the power it whole to the power n Poisson λ is equal to $e^{-\lambda}$ e to the power $\lambda(e^{it} - 1)$.

λ power x upon factorial x , e to the power itx is equal to $\sum_{x=0}^{\infty} e^{-\lambda} \frac{\lambda^x}{x!} e^{itx}$ is equal to $e^{-\lambda} \sum_{x=0}^{\infty} \frac{(\lambda e^{it})^x}{x!}$ is equal to $e^{-\lambda} e^{\lambda e^{it}}$ is equal to $e^{-\lambda} e^{\lambda(e^{it} - 1)}$ is equal to $e^{\lambda(e^{it} - 1)}$ that is the characteristic function of Poisson random variable.

Similarly, for exponential random variable this is going to be $\int_0^{\infty} e^{itx} \lambda e^{-\lambda x} dx$ is equal to $\lambda \int_0^{\infty} e^{-(\lambda - it)x} dx$ is equal to $\lambda \left[\frac{e^{-(\lambda - it)x}}{-(\lambda - it)} \right]_0^{\infty} = \frac{\lambda}{\lambda - it}$. Finally, for normal μ comma σ^2 is moment generating function is $\mu t + \frac{1}{2}\sigma^2 t^2$.

Therefore, here we should get $\frac{1}{2}\sigma^2 t^2$ is equal to $\mu t + \frac{1}{2}\sigma^2 t^2$, okay friends in a similar way, we can calculate the characteristic function of different random variables. Again like the moment generating function, the characteristic

function is also unique and given any distribution, it will have a unique characteristic function and given any characteristic function, it should correspond with one particular distribution.

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One distribution which is very useful is Cauchy distⁿ.

$X \sim C(1, 0)$ if $f(x) = \frac{1}{\pi} \frac{1}{1+x^2}$ $-\infty < x < \infty$

Its more general form is $C(r, \mu)$

$f(x) \sim \frac{1}{\pi} \frac{r}{r^2 + (x-\mu)^2}$

For this lecture we shall stick to $C(1, 0)$

$\therefore \int_{-\infty}^{\infty} \frac{1}{\pi} \frac{1}{1+x^2} dx = \frac{1}{\pi} \tan^{-1} x \Big|_{-\infty}^{\infty} = \frac{1}{\pi} (\frac{\pi}{2} - (-\frac{\pi}{2}))$

$= 1$ \therefore it can be considered as a pdf.

One particular distribution, where it is very useful is Cauchy distribution, x is distributed as Cauchy with 1 comma 0 if f_x is equal to $\frac{1}{\pi} \frac{1}{1+x^2}$. Its more general form is $C(r, \mu)$ and its f_x is equal to $\frac{1}{\pi} \frac{r}{r^2 + (x-\mu)^2}$. For this lecture we shall stick to $C(1, 0)$ since, so, I missed a point let us write that minus infinity less than x less than infinity.

That is it is defined over the entire r . Since integration minus infinity to infinity, $\frac{1}{\pi} \frac{1}{1+x^2} dx$ is equal to $\frac{1}{\pi} \tan^{-1} x$ from minus infinity to infinity is equal to $\frac{1}{\pi} \left(\frac{\pi}{2} - (-\frac{\pi}{2}) \right)$ is equal to 1 therefore, it can be considered as a PDF.

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What is expectation?

$$\frac{1}{\pi} \int_{-\infty}^{\infty} \frac{x}{1+x^2} dx$$

Analytically it is not convergent

But for most practical applications we consider the principal which is equal to 0

\therefore Therefore, we often write $E(x) = 0$

What about $E(x^2)$?

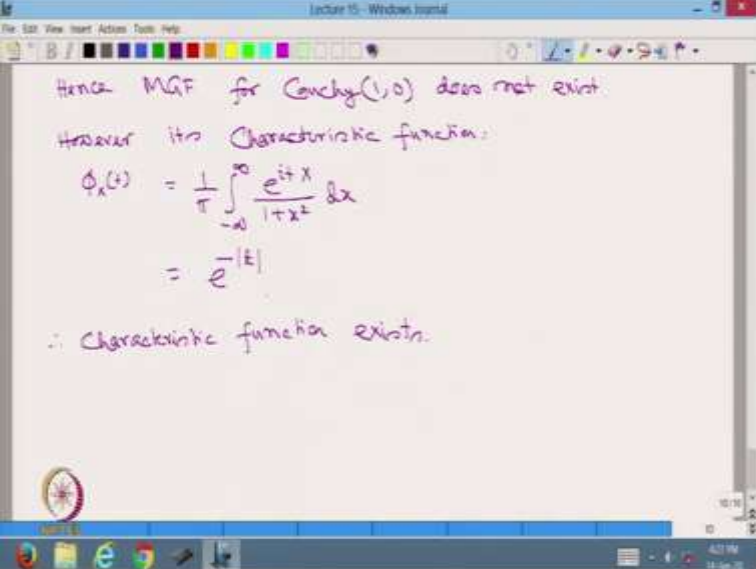
$$E(x^2) = \frac{1}{\pi} \int_{-\infty}^{\infty} \frac{x^2}{1+x^2} dx = 2 \int_0^{\infty} \frac{x^2}{1+x^2} dx$$

this is surely not convergent.

So, what is its expectation? $\frac{1}{\pi} \int_{-\infty}^{\infty} x \text{ upon } 1 \text{ plus } x^2 \text{ dx}$ analytically it is not convergent. But for most practical applications we consider with the principle value, which is equal to 0, and therefore, you often find that expected value of x is equal to 0. What about expected value of x^2 ?

Expected value of x^2 is equal to $\frac{1}{\pi} \int_{-\infty}^{\infty} x^2 \text{ upon } 1 \text{ plus } x^2 \text{ dx}$ is equal to $2 \int_0^{\infty} x^2 \text{ upon } 1 \text{ plus } x^2 \text{ dx}$ and this is surely not convergent. Thus we can show that higher order moments for a Cauchy distribution do not exist.

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Handwritten notes on a digital whiteboard:

Hence MGF for Cauchy(1,0) does not exist.

However its Characteristic function:

$$\phi_X(t) = \frac{1}{\pi} \int_{-\infty}^{\infty} \frac{e^{itx}}{1+x^2} dx$$
$$= e^{-|t|}$$

\therefore Characteristic function exists.

Hence, moment generating function for Cauchy 1 comma 0 does not exist. However, its characteristic function which is defined as $\phi_X(t)$ is equal to $\frac{1}{\pi} \int_{-\infty}^{\infty} e^{itx} \frac{1}{1+x^2} dx$ and I am not computing it, but this converges to $e^{-|t|}$. Therefore, characteristic function exists. Although its moment generating function does not exist.

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Properties of characteristic function

- 1) $\phi(0) = 1$ obvious
- 2) $|\phi(t)| \leq 1$ $\therefore |\phi(t)| = \left| \int_{-\infty}^{\infty} e^{itx} f(x) dx \right|$
 $\leq \int_{-\infty}^{\infty} |e^{itx}| f(x) dx = \int_{-\infty}^{\infty} f(x) dx = 1$
- 3) $\phi_x(t)$ and $\phi_x(-t)$ are complex conjugates.
 $\phi_x(t) = E(\cos tx + i \sin tx) = E(\cos tx) + i E(\sin tx)$
 $\therefore \overline{\phi_x(t)} = E(\cos tx) - i E(\sin tx)$
 $= E(\cos tx - i \sin tx) = E(e^{-itx})$
 $= \phi_x(-t)$

4. $\phi_{a+bx}(t) = e^{iat} \phi_x(bt)$
 This is straightforward.

Properties of characteristic function, 1, $\phi(0)$ is equal to 1 at the point 0 it is 1 which is obvious that is why I am not doing it. 2, Modulus of $\phi(t)$ is less than equal to 1, since modulus of $\phi(t)$ is equal to modulus of $\int_{-\infty}^{\infty} e^{itx} f(x) dx$ which is less than equal to $\int_{-\infty}^{\infty} |e^{itx}| f(x) dx$ which is equal to $\int_{-\infty}^{\infty} f(x) dx$ is equal to 1.

3, $\phi_x(t)$ and $\phi_x(-t)$ are complex conjugates, $\phi_x(t)$ is equal to expected value of $\cos tx$ plus i sine tx is equal to expected value of $\cos tx$ plus i times expected value of sine tx .

Therefore, $\phi_X(t)$ is conjugate is equal to expected value of $\cos tx$ minus i expected value of $\sin tx$ is equal to expected value of $\cos tx$ minus i $\sin tx$ is equal to expected value of e^{itx} to the power minus i $t x$ which is equal to ϕ_X at minus t . Thus, we get the result. 4, $\phi_X(a + bt)$ is equal to e^{iat} into $\phi_X(bt)$ this is straight forward and I leave it as an exercise.

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Probability Generating Function

This comparatively simple, and works for discrete integer valued r.v.

Suppose X is a r.v. taking values:

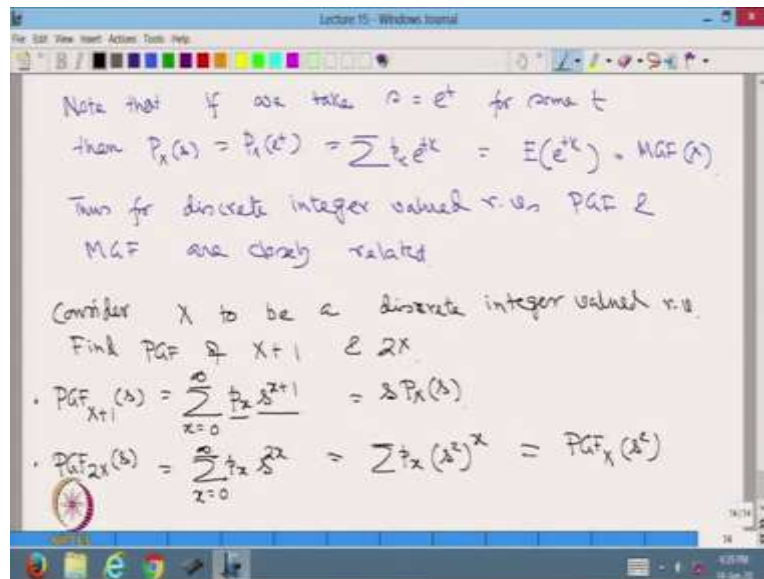
0	1	2	...	k	...
p_0	p_1	p_2	...	p_k	...

Then the PGF $P_X(s)$ is defined as:

$$P_X(s) = \sum_{i=0}^{\infty} p_i s^i \quad \text{if it converges for } -1 < s < 1 \text{ around } 0.$$

The next generating function that we study is Probability Generating Function. This is comparatively simple and works for discrete integer value random variables. So, suppose x is a random variable taking values 0, 1, 2, k like that, and probability of x taking the value i is equal to say p_i , then the probability generating function of x is defined as p_X at s is equal to $\sum p_i s^i$ to the power i , i is equal to 0 to infinity if it converges for some interval minus s not less than s less than s not around 0.

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Note that if we take $s = e^t$ for some t
 then $P_X(s) = P_X(e^t) = \sum p_k e^{tk} = E(e^{tk}) = \text{MGF}(X)$
 Thus for discrete integer valued r.v.s PGF & MGF are closely related

Consider X to be a discrete integer valued r.v.
 Find PGF of $X+1$ & $2X$

- $\bullet \text{PGF}_{X+1}(s) = \sum_{k=0}^{\infty} p_k s^{k+1} = s P_X(s)$
- $\bullet \text{PGF}_{2X}(s) = \sum_{k=0}^{\infty} p_k s^{2k} = \sum p_k (s^2)^k = \text{PGF}_X(s^2)$

Note that if we take s is equal to e to the power t for some t , then P_X of s is equal to p_X at e to the power t is equal to $\sum p_k e^{tk}$ is equal to the expected value of e to the power tk is equal to moment generating function of x thus for discrete integer valued random variables, probability generating function and moment generating function are closely related. Now consider x to be a discrete integer valued random variable.

Find PGF of x plus 1 and $2x$, PGF of x plus 1 at s is equal to $\sum_{k=0}^{\infty} p_k s^{k+1}$ because x plus 1 takes the value say 5 when x is equal to 4, therefore s to the power 5 will be multiplied by the probability of 4, same is for all x is equal to s times probability generating function of x around the point s PGF of $2x$ at a point s is equal to $\sum_{k=0}^{\infty} p_k s^{2k}$ is equal to $\sum p_k (s^2)^k$ is equal to PGF of x at the point s^2 .

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Similarly if we need PGF for X

$$PGF_X(s) = \sum_{x=0}^{\infty} p_x s^x = PGF_X(s^k)$$

Similarly, if we need PGF for Kx that is PGF of K times x at a point s is equal to sigma x is equal to 0 to infinity p_x into s to the power kx is equal to PGF of x at s the power k .

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Now suppose X is a r.v taking values $0, 1, 2, \dots$

$X =$	0	1	2	\dots	k	\dots
P_x	p_0	p_1	p_2	\dots	p_k	\dots

Consider $q_k = P(X > k)$

$$\begin{aligned} q_0 &= p_1 + p_2 + p_3 + \dots \\ q_1 &= p_2 + p_3 + \dots \\ q_2 &= p_3 + \dots \end{aligned}$$

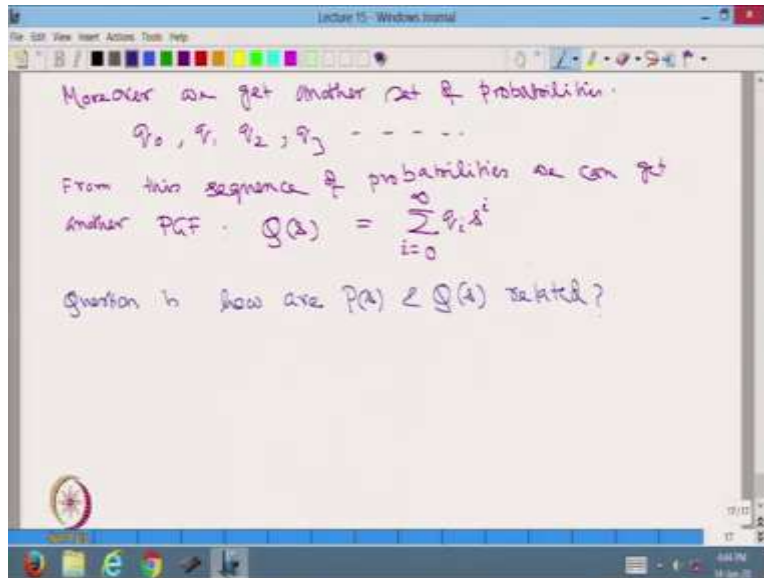
$$\sum_{k=0}^{\infty} q_k = 1p_1 + 2p_2 + 3p_3 + \dots + kp_k + \dots = E(X)$$

Thus for discrete integer valued r.v's we find $\sum_{k=0}^{\infty} P(X > k) = E(X)$

Now, suppose x is a random variable taking values $0, 1, 2$ up to infinity like this x takes the values $0, 1, 2, K$ with p_0, p_1, p_2, p_k . Consider q_k is equal to probability x greater than k . Therefore, q_0 is equal to p_1 plus p_2 plus p_3 up to infinity q_1 is equal to p_2 plus p_3 up to infinity, q_2 is equal to p_3 plus up to infinity.

Like that if we sum then we get $\sum_{i=0}^{\infty} q_i$ is equal to 1 times p_1 plus 2 times p_2 plus 3 times p_3 plus k times p_k plus up to infinity. This is nothing but expected value of x thus for discrete integer valued PMFs we find $\sum_{x \text{ greater than } x} x$ is equal to 0 to infinity is equal to expected value of x . So this is an interesting property of such variables.

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But for me, for the time being, we find that another set of probabilities namely, q_0, q_1, q_2, q_3 like that, from this sequence of probabilities we can get another probability generating function namely $Q(s)$ is equal to $\sum_{i=0}^{\infty} q_i s^i$, i is equal to 0 to infinity, question is how are $P(s)$ and $Q(s)$ related that is a very interesting question and we solve it as follows.

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Handwritten derivation on a digital whiteboard:

Note that $q_i = P(X > x_i)$
 $\therefore p_i = P(X > x_{i-1}) - P(X > x_i)$

$$P(s) = \sum_{x=0}^{\infty} p_x s^x = p_0 + \sum_{i=1}^{\infty} p_i s^i$$

$$= p_0 + \sum_{i=1}^{\infty} (q_{i-1} - q_i) s^i$$

$$= p_0 + \sum_{i=1}^{\infty} q_{i-1} s^i - \sum_{i=1}^{\infty} q_i s^i$$

$$= p_0 + s \sum_{i=1}^{\infty} q_{i-1} s^{i-1} - \sum_{i=0}^{\infty} q_i s^i + q_0$$

$$= (p_0 + q_0) + s \sum_{j=0}^{\infty} q_j s^j - \sum_{i=0}^{\infty} q_i s^i$$

$$= (p_0 + q_0) + Q(s)(s-1)$$

Now $p_0 + q_0 = 1$
 $= p_0 + p_1 + p_2 + p_3 + \dots$
 $= 1$

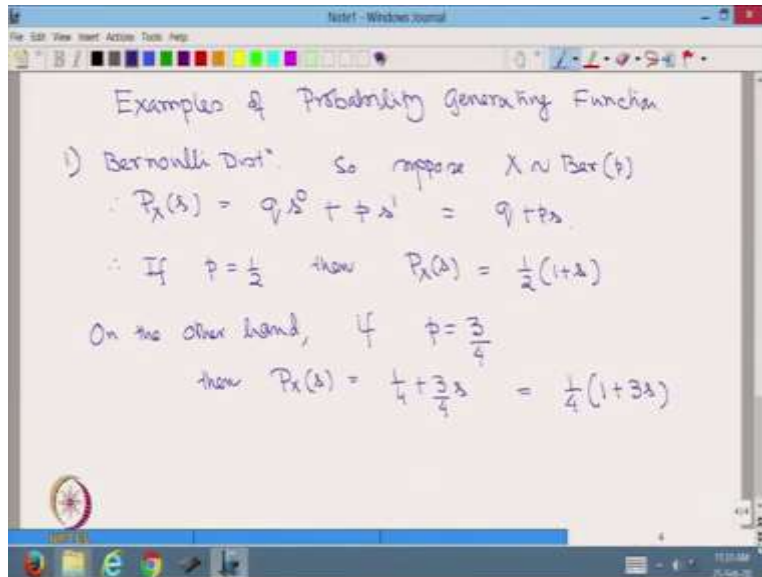
$P(s) = 1 + Q(s)(s-1)$
 $Q(s) = \frac{1 - P(s)}{1 - s}$

Q_i is equal to probability x greater than x_i . Therefore, p_i is equal to probability x greater than x_i minus 1 minus probability x greater than x_i . Therefore, P_s of x is equal to $\sum_{x=0}^{\infty} p_x s^x$ is equal to p_0 plus $\sum_{i=1}^{\infty} p_i s^i$ is equal to p_0 plus $\sum_{i=1}^{\infty} (q_{i-1} - q_i) s^i$ is equal to p_0 plus $\sum_{i=1}^{\infty} q_{i-1} s^i$ minus $\sum_{i=1}^{\infty} q_i s^i$ is equal to p_0 plus $s \sum_{i=1}^{\infty} q_{i-1} s^{i-1}$ minus $\sum_{i=0}^{\infty} q_i s^i$ plus q_0 is equal to $(p_0 + q_0) + s \sum_{j=0}^{\infty} q_j s^j - \sum_{i=0}^{\infty} q_i s^i$ is equal to $(p_0 + q_0) + Q(s)(s-1)$.

Since we have added the terms q_0 to the power 0, which is q_0 with a minus sign to compensate, we had one q_0 is equal to p_0 plus q_0 plus s times now, we understand that since summing from i is equal to 1 to infinity, q_i minus 1 into s to the power i , i minus 1, we can write it as j is equal to 0 to infinity, q_j s to the power j minus $\sum_{i=0}^{\infty} q_i s^i$ this is equal to p_0 plus q_0 plus this is q_s and this is also q_s . So, we can write it as q_s into s minus 1.

Now, p_0 plus q_0 is equal to p_0 plus probability is greater than 0, which is equal to p_1 plus p_2 plus p_3 up to infinity is equal to 1. Therefore, P_s let me write it as P_s is equal to 1 plus Q_s into s minus 1. Therefore, Q_s is equal to $1 - P_s$ upon $1 - s$. So, these two different probability generating functions are related by this relationship. So, this is an interesting generating function which is comparatively simpler than moment generating function or characteristic function.

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Examples of Probability Generating Function

1) Bernoulli Dist'. So suppose $X \sim \text{Ber}(p)$

$$\therefore P_X(s) = q s^0 + p s^1 = q + ps$$

\therefore If $p = \frac{1}{2}$ then $P_X(s) = \frac{1}{2}(1+s)$

On the other hand, if $p = \frac{3}{4}$

$$\text{then } P_X(s) = \frac{1}{4} + \frac{3}{4}s = \frac{1}{4}(1+3s)$$

Let us now see examples of probability generating function, 1, Bernoulli distribution so, suppose x is distributed as Bernoulli P therefore, $P_X(s)$ at s is equal to q times s to the power 0 plus p times s to the power 1 is equal to $q + ps$. Therefore, if p is equal to half then $P_X(s)$ is equal to half times 1 plus s . On the other hand if P is equal to say 3 by 4, then $P_X(s)$ is equal to 1 by 4 plus 3 by 4 times s is equal to 1 by 4 into 1 plus 3 s .

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The image shows a screenshot of a Windows Journal window titled "Notet - Windows Journal". It contains handwritten mathematical derivations for the probability mass functions of the Binomial and Poisson distributions.

2) Binomial (n, p)

$$\begin{aligned}
 P_x(s) &= \sum_{x=0}^n \binom{n}{x} p^x q^{n-x} s^x \\
 &= \sum_{x=0}^n \binom{n}{x} (ps)^x q^{n-x} \\
 &= (q + ps)^n
 \end{aligned}$$

3) Poisson (λ)

$$\begin{aligned}
 P_x(s) &= \sum_{x=0}^{\infty} e^{-\lambda} \frac{\lambda^x}{x!} s^x \\
 &= e^{-\lambda} \sum_{x=0}^{\infty} \frac{(\lambda s)^x}{x!} = e^{-\lambda} e^{\lambda s} = e^{-\lambda(1-s)}
 \end{aligned}$$

Binomial distribution with parameter n comma p therefore, P_x of s is equal to sigma x is equal to 0 to n , n choose x p to the power x , q to the power n minus x into s to the power x is equal to sigma x is equal to 0 to n , n choose x ps to the power x q to the power n minus x which is equal to $q + ps$ whole to the power n .

Poisson distribution with λ therefore, P_x of s is equal to sigma x is equal to 0 to infinity e to the power minus λ , λ power x upon factorial x into s to the power x which is equal to e to the power minus λ sigma x is equal to 0 to infinity λ s to the power x upon factorial x which is equal to e to the power minus λ into e to the power λs which is equal to e to the power minus λ into 1 minus s .

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How do we get back the probabilities p_x from $P(s)$?

$$P_X(s) = p_0 + p_1 s + p_2 s^2 + \dots + p_n s^n + \dots$$

May be finite or infinite series.

$$P'_X(s) = p_1 + 2p_2 s + 3p_3 s^2 + 4p_4 s^3 + \dots$$

Now $p_X(0) = p_0 = \text{Prob}(X=0)$
 $p'_X(0) = p_1 = \text{Prob}(X=1)$
 $p''_X(0) = 2p_2 + 3 \cdot 2 p_3 s + 4 \cdot 3 p_4 s^2 + \dots \Big|_0$
 $= 2p_2 = 2 \text{ times } P(X=2)$

$\therefore p'_x(s) \Big|_0 = 2 p_2$
 In a similar way
 $p^{(3)}_x(s) = 3 \cdot 2 \cdot p_3 + 4 \cdot 3 \cdot 2 \cdot p_4 s + \dots$
 $\therefore p^{(3)}_x(0) = 3! p_3 = 3! P(X=3)$
 In a similar way $p^{(r)}_x(s) = r! P(X=r)$
 by setting $s=0$

How do we get back our probabilities? P_x from the probability generating function $p(s)$ of x that is very simple $p_x(s)$ is equal to p_0 plus $p_1 s$ plus $p_2 s^2$ up to say $p_n s^n$ $p_n s$ to the power n . It may be finite or infinite series $P'_x(s)$ is equal to p_1 plus $2 p_2 s$ plus $3 p_3 s^2$ plus $4 p_4 s^3$, like that.

Now, P_x at $x=0$ is equal to p_0 is equal to probability x is equal to 0, p'_x at $x=0$ is equal to p_1 is equal to probability x is equal to 1. P''_x at $x=0$ is equal to $2p_2 + 3p_3$ is equal to probability x is equal to 2.

plus 4 into 3 into p_4 s square plus, etc at 0 is equal to 2 p_2 is equal to 2 times probability x is equal to 2.

Therefore, $p''(x)$ at 0 is equal to 2 times p_2 . In a similar way the third derivative of probability generating function is equal to 3 into 2 into p_3 plus 4 into 3 into 2 into p_4 s, etc or $p_x^{(3)}$ at 0 is equal to factorial 3 into p_3 is equal to factorial 3 into probability x is equal to 3. In a similar way, we can find the r^{th} derivative of the probability generating function will give us r factorial into probability x is equal to r by setting s equal to 0.

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Let us observe some important properties:

$$P_X(s) \Big|_{s=1} = p_0 + p_1 + p_2 + \dots = 1$$

$$P'_X(s) \Big|_{s=1} = p_1 + 2p_2 + 3p_3 + \dots = E(X) \checkmark$$

$$P''_X(s) \Big|_{s=1} = 2p_2 + 3 \cdot 2 \cdot p_3 + 4 \cdot 3 \cdot p_4 + \dots$$

$$= \sum_{x=2}^{\infty} x(x-1) p_x = E(X(X-1))$$

$$= E(X^2) - E(X)$$

$$\sigma^2 = P''_X(s) \Big|_{s=1} + P'_X(s) \Big|_{s=1} - (P'_X(s) \Big|_{s=1})^2 = V(X) \checkmark$$

Let us now observe some important properties p_x at s at s is equal to 1 is equal to p_0 plus p_1 plus p_2 is equal to 1, $p'(x)$ at s is equal to 1 is equal to p_1 plus 2 p_2 plus 3 p_3 etc is equal to expectation of x . Similarly, $p''(x)$ at s at s is equal to 1 is equal to 2 p_2 plus 3 into 2 into p_3 plus 4 into 3 into p_4 like that is equal to sigma x is equal to 2 to infinity x into x minus 1 into p_x which is equal to expected value of x into x minus 1 which is equal to expected value of x square minus expected value of x .

Therefore, $p''(x)$ of s at s is equal to 1 plus $p'(x)$ at s is equal to 1 minus $(p'(x) \text{ at } s \text{ is equal to 1})^2$ is equal to variance of x . I like you to verify this results for different discrete distributions.

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Cumulant Generating function

The cumulative generating function of a r.v. X denoted as $K_X(t)$ is defined as $\ln M_X(t)$
 i.e. it is the natural log of the MGF.

Since $M_X(t) = 1 + \mu_1 t + \frac{\mu_2^2}{2!} t^2 + \frac{\mu_3^3}{3!} t^3 + \dots$
 \therefore What is $\ln(M_X(t))$?

The last topic on this is called Cumulant Generating Function, the cumulative generating function of a random variable x denoted as k_x of t is defined as \log of M_{xt} to the base e . That is, it is the natural log of the moment generating function since M_{xt} is equals 1 plus μ_1 prime t plus t square upon factorial 2 μ_2 prime plus t cube upon factorial 3 μ_3 prime like that. Therefore, what is \log of M_{xt} ? That is the question.

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We know $\ln(1+x) = x - \frac{x^2}{2} + \frac{x^3}{3} - \frac{x^4}{4} + \dots$

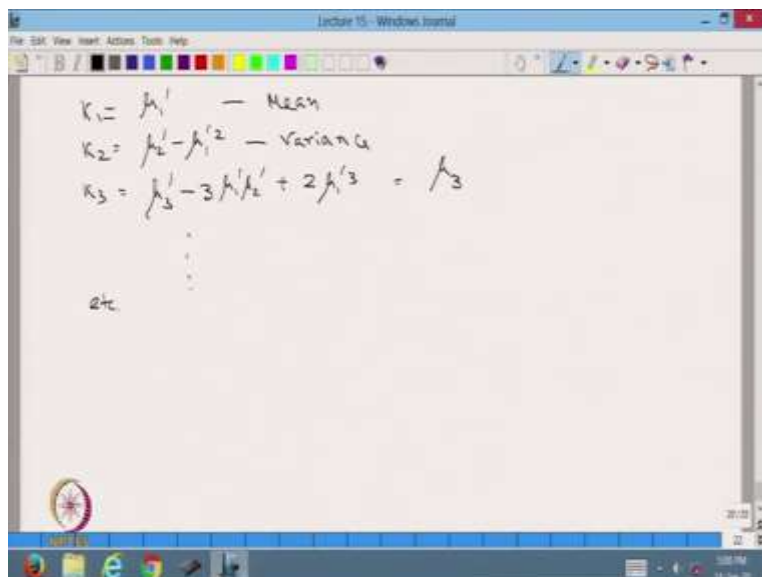
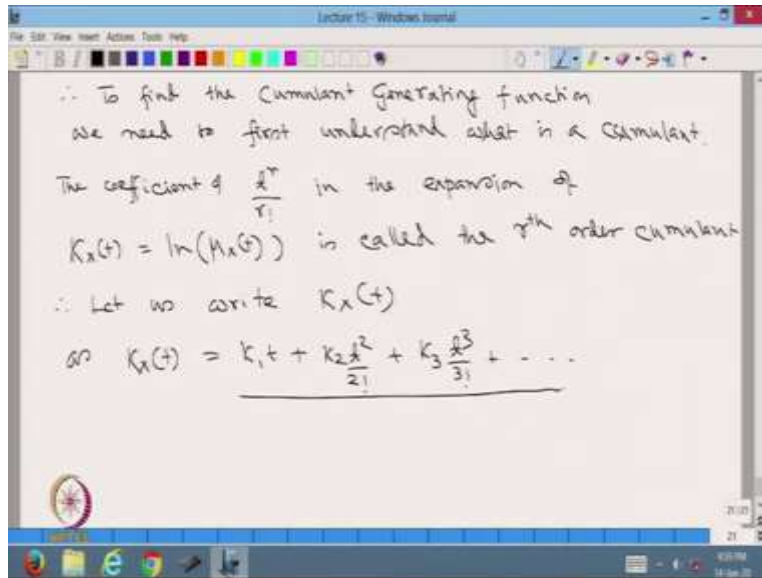
$\therefore \ln\left(1 + \mu_1 t + \frac{\mu_2^2}{2!} t^2 + \frac{\mu_3^3}{3!} t^3 + \frac{\mu_4^4}{4!} t^4 + \dots\right)$

$= \left[\mu_1 t + \frac{\mu_2^2}{2!} t^2 + \frac{\mu_3^3}{3!} t^3 + \dots \right]$

$\quad - \frac{1}{2} \left(\mu_1 t + \frac{\mu_2^2}{2!} t^2 + \frac{\mu_3^3}{3!} t^3 + \dots \right)^2$

$\quad + \frac{1}{3} \left(\mu_1 t + \frac{\mu_2^2}{2!} t^2 + \frac{\mu_3^3}{3!} t^3 + \dots \right)^3$

$\quad \vdots$



We know natural log of 1 plus x is equal to x minus x square upon 2 plus x cube upon 3 minus x 4 upon 4 etc therefore, log of 1 plus $\mu_1' t$ plus t^2 upon 2 minus $\mu_2' t^2$ plus t^3 upon factorial 3 minus $\mu_3' t^3$ plus t^4 upon factorial 4 minus $\mu_4' t^4$ etc is equal to $t \mu_1' + \frac{t^2}{2} \mu_2' - \frac{t^3}{6} \mu_3' + \frac{t^4}{24} \mu_4' - \dots$ like that, we can have the infinite sum.

Therefore, to find the Cumulant Generating Function we need to first understand what is a cumulant. The coefficient of t to the power r upon r factorial in the expansion of $K_X(t)$, which is

And if we equate this with this expression that we got here, what we are getting is k_1 is equal to μ_1' , which is the mean k_2 is equal to $\mu_2' - \mu_1'^2$, which is equal to variance k_3 is equal to $\mu_3' - 3\mu_1'\mu_2'$, $\mu_2' + 2\mu_1'^3$, which is equal to the third order central moment etc. Thus from the Cumulant generating function also we can get moments of different orders.

Ex $N(\mu, \sigma^2)$: $M_X(t) = e^{j\mu + \frac{1}{2}\sigma^2 k^2}$
 $\therefore k_X(t) = \mu + \frac{1}{2}\sigma^2 k^2$
 $\therefore k_1(t) = \mu$
 $k_2(t) = \sigma^2$
 $k_r(t) = 0 \quad r > 2$

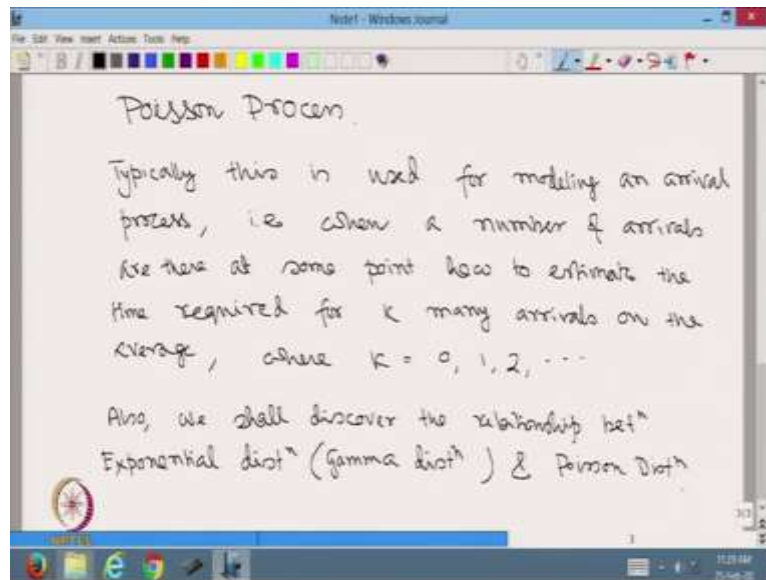
Ex $\Gamma(1, \alpha)$
 $\therefore \text{MGF} = \left(\frac{1}{1-t}\right)^{-\alpha} \therefore \ln M_X(t) = -\alpha \ln(1-t)$
 $= -\alpha \left(-t - \frac{t^2}{2} - \frac{t^3}{3} - \dots \right)$
 $= \alpha \left(t + \frac{t^2}{2} + \frac{t^3}{3} + \dots \right)$
 $\therefore k_1(t) = \alpha \quad k_2(t) = \alpha \quad k_3(t) = \underline{\underline{\alpha}} \quad k_4(t) = \underline{\underline{\alpha}} \dots$

Similarly, another example say gamma with 1 comma alpha therefore MGF is equal to 1 upon 1 minus t whole to the power minus alpha therefore, log of Mxt is equal to minus alpha into log of 1 minus t is equal to minus alpha into minus t minus t square by 2 minus t cube by 3 is equal to alpha times t plus t square by 2 plus t cube by 3 etc.

Therefore, $K_1 t$ is equal to α , $k_2 t$ is equal to α , $k_3 t$ is equal to 2α , $k_4 t$ is equal to 6α etc. This is very simple, because we have to consider the coefficient of t to the power r

upon factor r and to compensate for that factorial we need this extra coefficients. Okay friends, I stop here today.

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In the next class we shall discuss an important concept of probability which is called Poisson process. Typically this is used for modeling an arrival process that is when a number of arrivals are there at some point how to estimate the time required for K many arrivals on the average, where K can be 0, 1, 2 etc. Also we shall discover the relationship between exponential distribution or gamma distribution in a more generalized way and Poisson distribution. Thank you.