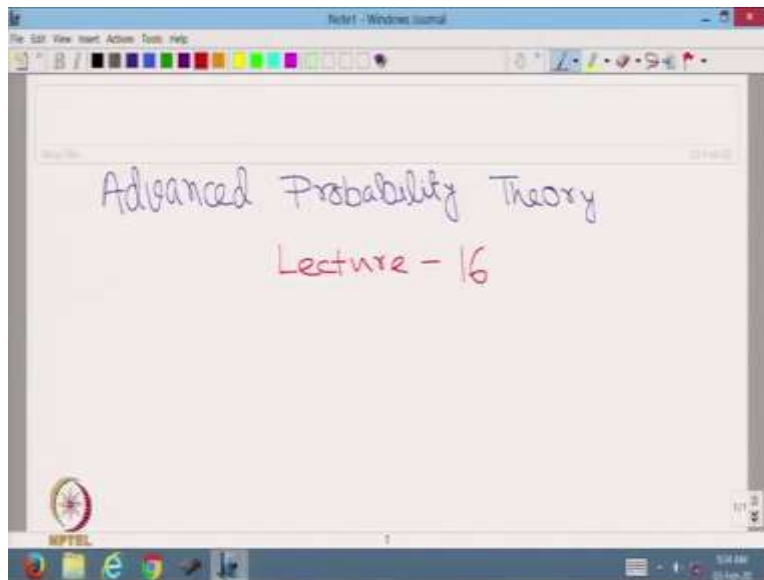


**Advanced Probability Theory**  
**Professor Niladri Chatterjee**  
**Department of Mathematics**  
**Indian Institute of Technology, Delhi**  
**Lecture 16**

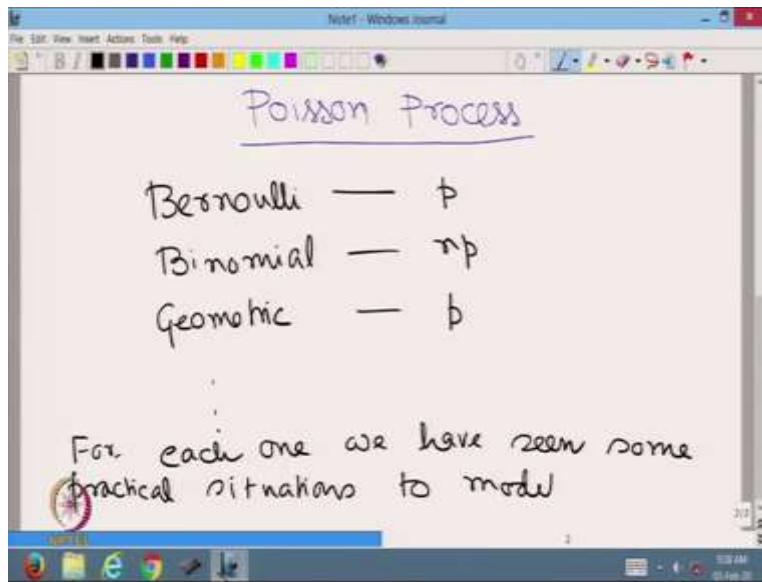
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Welcome students to the MOOC lecture series on Advanced Probability Theory this is lecture number 16. If you notice that over the last 6 weeks, we have studied the basic theory of probability, then random variables, discrete continuous types, and also different moments of different random variables in detail. And each week, we have focused on one particular topic in this week. However, we shall not focus on a particular topic rather, what we shall do, we shall touch upon some basic interesting results that can come from different distributions.

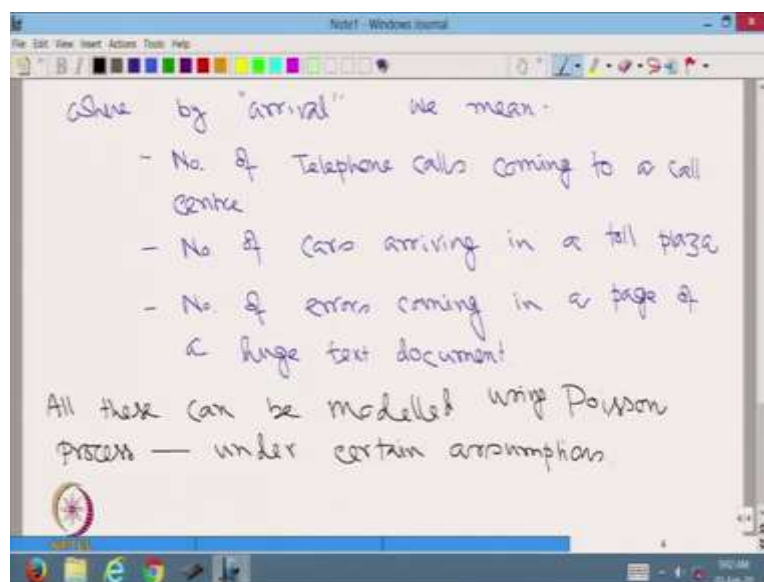
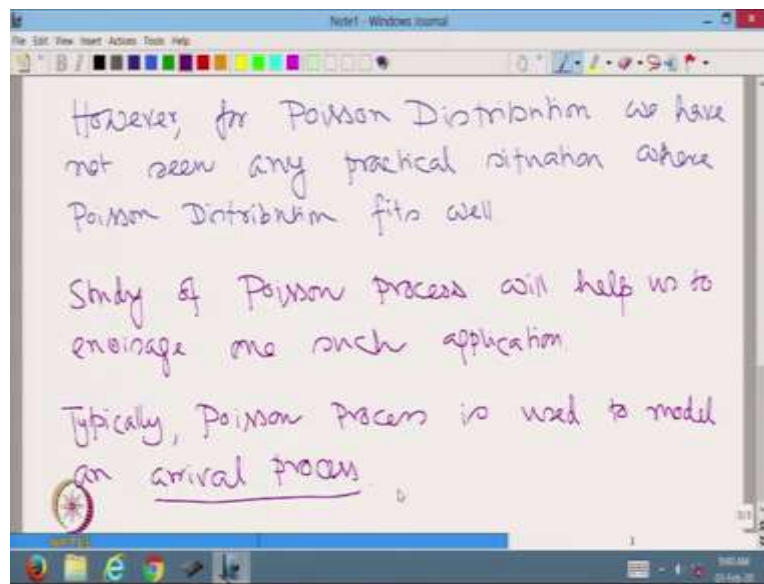
Also we shall study by variate normal distribution, which is a very important probability distribution, when you deal with multivariate data, we shall talk about that later.

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For today's lecture, the focus is on what is called Poisson process. We have studied many discrete distributions, say binomial, say Bernoulli with parameter  $p$  binomial with parameter  $np$  geometric with  $P$  etc. For each one of them we have seen some practical situations to model them.

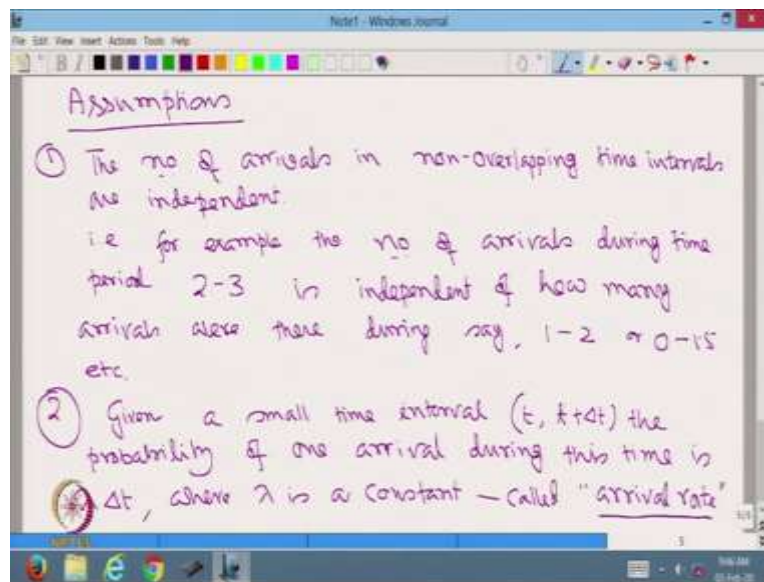
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However, for Poisson distribution we have not yet seen any practical situation where Poisson distribution fits well. Study of Poisson process will help us to envisage one such application. Typically Poisson process is used to model an arrival process, this is very important that a very specific type of practical problems can be modeled with Poisson distribution or Poisson process where by arrival we mean many situations say number of telephone calls coming to a call center.

It may mean number of cars arriving in a toll plaza it may mean number of errors coming in a page of a huge text document all these can be modelled using Poisson process under certain assumptions, okay. So, let us first see what are the basic assumptions.

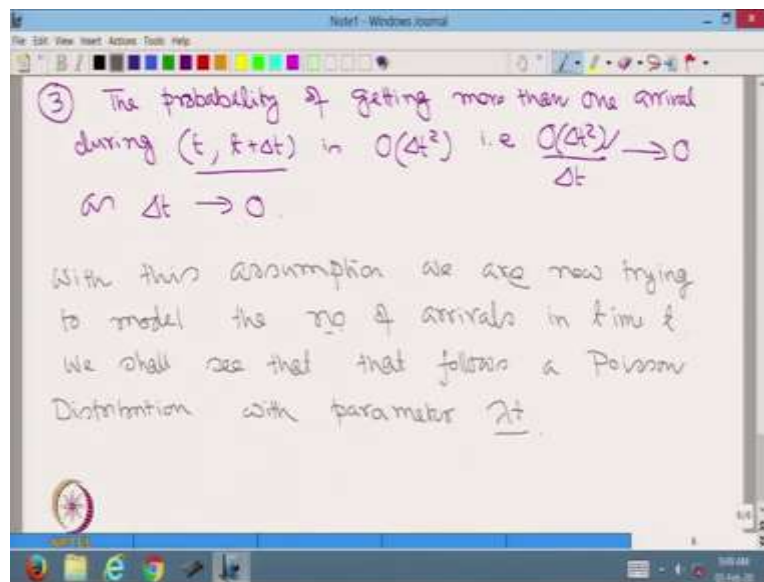
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One, the number of arrivals in non-overlapping time intervals are independent that is, say for example, the number of arrivals during time period 2 to 3 is independent of how many arrivals were there during say 1 to 2 or 0 to 1.5 etc.

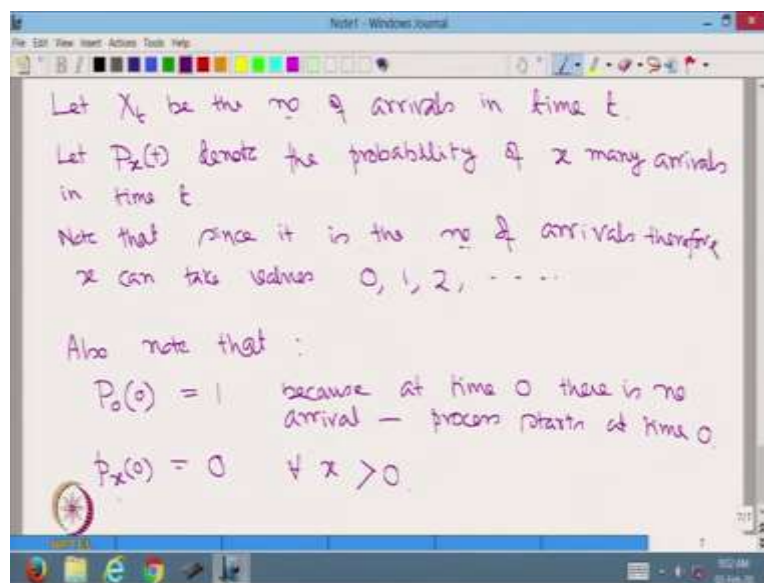
Second assumption is that given a small time interval  $t$  to  $t$  plus  $\Delta t$  the probability of one arrival during this time is  $\lambda \Delta t$ , where  $\lambda$  is a constant and we call it arrival rate.

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Assumption 3, the probability of getting more than one arrival during  $t$  to  $t$  plus  $\Delta t$ , a small interval is order of  $\Delta t$  square that is order of  $\Delta t$  square upon  $\Delta t$  will go to 0, as  $\Delta t$  goes to 0 that means that this is the very-very small quantity even in comparison with  $\Delta t$  with this assumption. We are now trying to model the number of arrivals in time  $t$  we shall see that, that follows a Poisson distribution with parameter  $\lambda t$ .

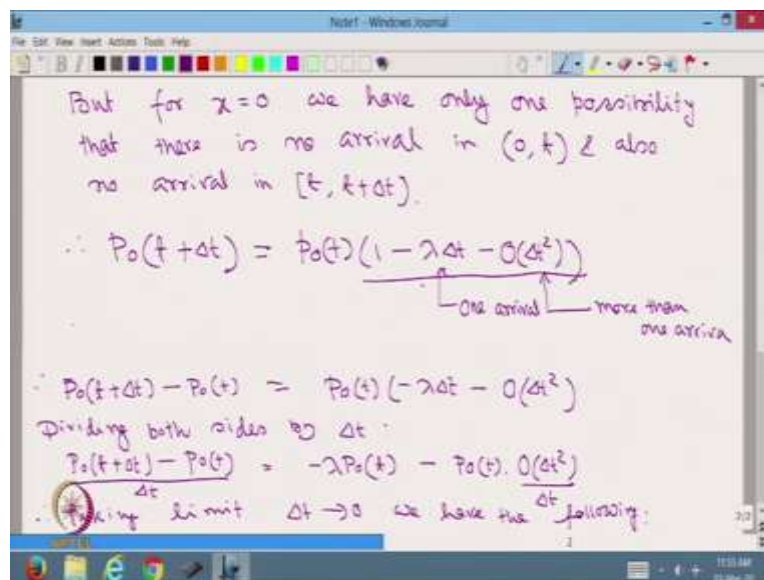
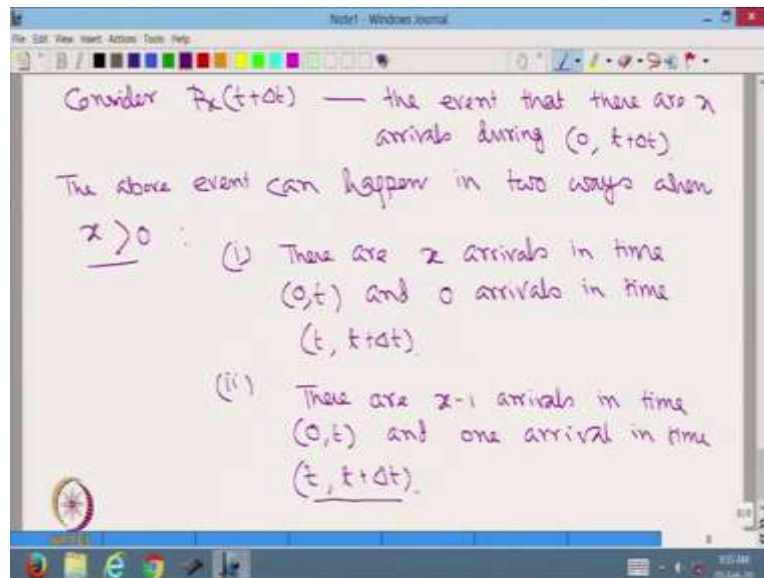
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So, let us start, let  $x_t$  be the number of arrivals in time  $t$ , let  $P_{xt}$  denote the probability of  $x$  many arrivals in time  $t$ . Note that since it is the number of arrivals therefore  $x$  can take values 0, 1, 2,

etc. Also note that  $P_{00}$  is equal to 1 because at time 0 there is no arrival that is process starts at time 0 and  $p_x 0$  is to 0 for all  $x$  greater than 0, because the same reason that at times 0 there is no arrival.

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So consider  $P_{xt}$  plus  $\Delta t$  that is the event, there are  $x$  arrivals during 0 to  $t$  plus  $\Delta t$ , the above event can happen in two ways when  $x$  is greater than 0, what are they? The first one is that there are  $x$  arrivals in time 0 to  $t$  and 0 arrivals in time  $t$  to  $t$  plus  $\Delta t$ . Other way is that there are  $x$  minus 1 arrivals in time 0 to  $t$  and one arrival in time  $t$  to  $t$  plus  $\Delta t$ .

And we have seen that or we have assumed and we have assumed that there cannot be more than arrival more than one arrival in this period. Therefore, these are the only two possibilities when  $x$  is greater than 0.

But for  $x$  is equal to 0 we have only one possibility that there is no arrival in 0 to  $t$  and also no arrival in  $t$  to  $t + \Delta t$ . Therefore,  $P_0(t + \Delta t)$  is equal to  $P_0(t) \cdot (1 - \lambda \Delta t - \text{order of } \Delta t^2)$  because in time interval  $t$  to  $t + \Delta t$ , this is the probability of one arrival. This is the probability of more than one arrival.

Therefore, one minus this quantity gives us the probability of zero arrivals during this period. Therefore,  $P_0(t + \Delta t) - P_0(t)$  is equal to  $P_0(t) \cdot (-\lambda \Delta t - \text{Big O of } \Delta t^2)$ . Dividing both sides by  $\Delta t$ ,  $\frac{P_0(t + \Delta t) - P_0(t)}{\Delta t}$  is equal to  $-\lambda P_0(t) - \text{Big O of } \Delta t$ . Therefore, taking limit  $\Delta t$  is going to 0 we have the following.



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Handwritten derivation on a digital notepad:

$$\begin{aligned} \text{LHS} &= p_0'(t) \\ \text{RHS} &= -\lambda p_0(t) - p_0(t) \lim_{\Delta t \rightarrow 0} \frac{O(\Delta t^2)}{\Delta t} \\ &= -\lambda p_0(t) - 0 = -\lambda p_0(t) \end{aligned}$$

$\therefore$  We get  $p_0'(t) = -\lambda p_0(t)$

or  $\frac{p_0'(t)}{p_0(t)} = -\lambda$

$\therefore$  After integration we have  $\log p_0(t) = -\lambda t + C$

At  $t=0$ ,  $p_0(0) = 1$ ,  $\log p_0(0) = 0 = -\lambda \cdot 0 + C$

$\Rightarrow C = 0$

Handwritten derivation on a digital notepad:

$\therefore$  we have  $\log p_0(t) = -\lambda t$

$\therefore$  hence  $p_0(t) = e^{-\lambda t}$

i.e. probability of zero arrivals in time  $t$

$$= e^{-\lambda t} = e^{-\lambda t} \frac{(\lambda t)^0}{0!}$$

Thus we find

$$P(0 \text{ arrivals in time } t) = e^{-\lambda t} \frac{(\lambda t)^0}{0!}$$

LHS is equal to  $P_0'(t)$  and RHS is equal to  $-\lambda p_0(t) - p_0(t) \lim_{\Delta t \rightarrow 0} \frac{O(\Delta t^2)}{\Delta t}$ . Therefore, this is equal to  $-\lambda p_0(t) - 0$  is equal to  $-\lambda p_0(t)$ . Therefore, we get  $P_0'(t) = -\lambda p_0(t)$ . Or  $\frac{p_0'(t)}{p_0(t)} = -\lambda$ .

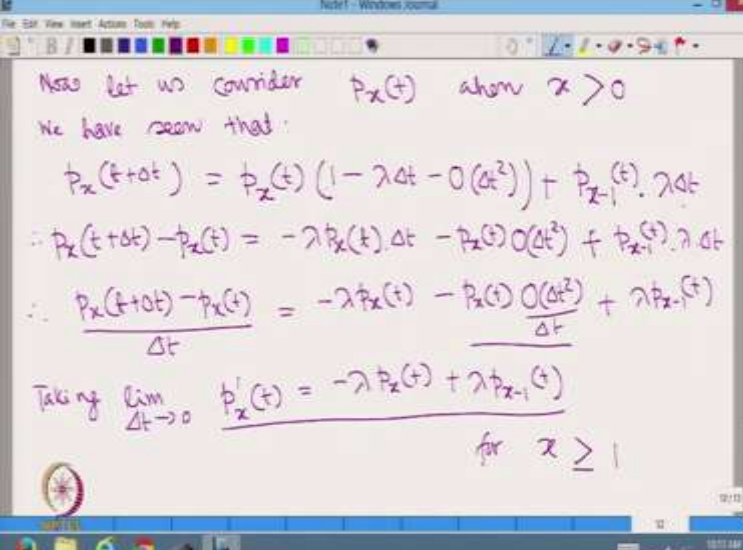
Therefore, after integration we have  $\log p_0(t) = -\lambda t + C$  at  $t=0$ ,  $p_0(0) = 1$ . Therefore,  $\log p_0(0) = 0 = -\lambda \cdot 0 + C$  implies  $C = 0$ . Therefore, we have  $\log p_0(t) = -\lambda t$  and hence  $p_0(t) = e^{-\lambda t}$  that is probability of 0 arrivals in time  $t$  is equal to  $e^{-\lambda t}$ .



to the power minus lambda t which is equal to e to the power minus lambda t, lambda t to the power 0 upon factorial 0.

Thus, we find probability zero arrivals in time t is equal to e to the power minus lambda t, lambda t to the power 0 upon factorial 0.

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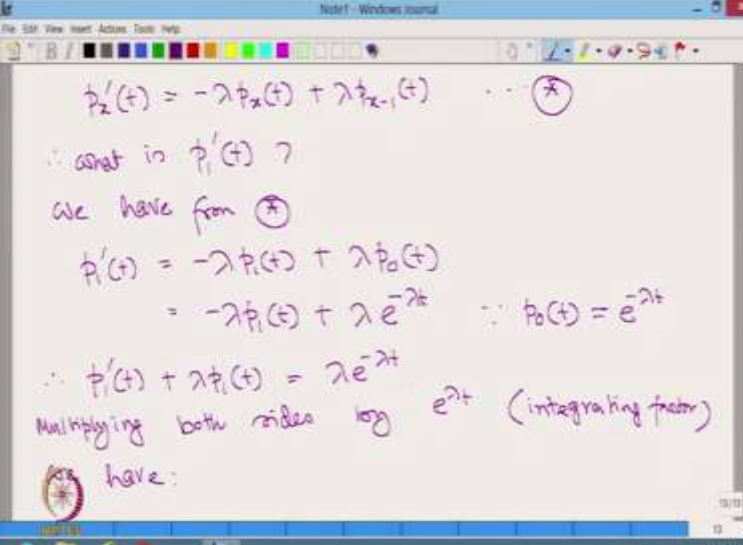
Now let us consider  $P_x(t)$  where  $x > 0$   
 we have seen that:

$$P_x(t + \Delta t) = P_x(t) (1 - \lambda \Delta t - O(\Delta t^2)) + P_{x-1}(t) \cdot \lambda \Delta t$$

$$\therefore P_x(t + \Delta t) - P_x(t) = -\lambda P_x(t) \Delta t - P_x(t) O(\Delta t^2) + P_{x-1}(t) \cdot \lambda \Delta t$$

$$\therefore \frac{P_x(t + \Delta t) - P_x(t)}{\Delta t} = -\lambda P_x(t) - \frac{P_x(t) O(\Delta t^2)}{\Delta t} + \lambda P_{x-1}(t)$$

Taking  $\lim_{\Delta t \rightarrow 0}$   $P'_x(t) = -\lambda P_x(t) + \lambda P_{x-1}(t)$   
 for  $x \geq 1$



$$P'_x(t) = -\lambda P_x(t) + \lambda P_{x-1}(t) \quad \dots (*)$$

$\therefore$  what is  $P'_1(t)$ ?

We have from (\*)

$$P'_1(t) = -\lambda P_1(t) + \lambda P_0(t)$$

$$= -\lambda P_1(t) + \lambda e^{-\lambda t} \quad \therefore P_0(t) = e^{-\lambda t}$$

$$\therefore P'_1(t) + \lambda P_1(t) = \lambda e^{-\lambda t}$$

Multiplying both sides by  $e^{\lambda t}$  (integrating factor)

we have:

Handwritten derivation of the probability  $p_1(t)$  for a Poisson process:

$$e^{\lambda t} p_1'(t) + \lambda e^{\lambda t} p_1(t) = \lambda$$

$$\Rightarrow \frac{d}{dt} (e^{\lambda t} p_1(t)) = \lambda$$

After integration:  $e^{\lambda t} p_1(t) = \lambda t + C$

At time  $t=0$  we know  $e^{\lambda \cdot 0} p_1(0) = \lambda \cdot 0 + C$

Since  $p_1(0) = 0$ , we have  $C = 0$

We have:  $e^{\lambda t} p_1(t) = \lambda t$  or  $p_1(t) = \frac{\lambda t}{e^{\lambda t}} = e^{-\lambda t} \lambda t$

Handwritten derivation of the probability  $p_2(t)$  for a Poisson process:

We get probability of getting two arrivals in time  $t = e^{-\lambda t} \lambda^2 t^2 / 2!$  ✓

Let us now consider  $p_2(t)$

We have in a similar way:

$$p_2'(t) = -\lambda p_2(t) + \lambda p_1(t)$$

$$\text{or } p_2'(t) + \lambda p_2(t) = \lambda \cdot e^{-\lambda t} \lambda t = \lambda^2 e^{-\lambda t} t$$

Again using  $e^{\lambda t}$  as integrating factor:

$$e^{\lambda t} p_2'(t) + \lambda e^{\lambda t} p_2(t) = \lambda^2 t \quad \text{or} \quad \frac{d}{dt} (e^{\lambda t} p_2(t)) = \lambda^2 t$$

Now let us consider  $P_x(t)$  when  $x$  is greater than 0. We have seen that  $P_x(t) + \Delta t$  is equal to is probability of  $x$  arrivals in time  $t$  multiplied by 0 arrivals in time  $t$  to  $t + \Delta t$  plus probability of  $x - 1$  arrivals in time  $t$ , multiplied by one arrival in time  $t$  to  $t + \Delta t$ . This we have already seen.

Therefore,  $P_x(t) + \Delta t - P_x(t)$  is equal to minus  $\lambda P_x(t) \Delta t$  minus  $P_x(t)$  order of Big O of  $\Delta t^2$  plus  $P_{x-1}(t) \lambda \Delta t$ . Therefore,  $P_x(t) + \Delta t - P_x(t)$  upon  $\Delta t$  is equal to minus  $\lambda P_x(t)$  minus  $P_x(t)$  Big O of  $\Delta t$  upon  $\Delta t$  plus  $\lambda P_{x-1}(t)$  taking limit  $\Delta t$  going to 0, we have  $P_x'(t)$  is equal to minus  $\lambda P_x(t)$  this goes to 0 plus  $\lambda P_{x-1}(t)$ .

So, this is the result that we have for  $x$  greater than equal to 1. Therefore, what we have? We have  $p_x$  prime  $t$  is equal to minus  $\lambda p_x t$  plus  $\lambda p_x$  minus 1  $t$ , therefore, what is  $p_1$  prime  $t$ ? This is we have say let us call it star  $p_1$  prime  $t$  is equal to minus  $\lambda p_1 t$  plus  $\lambda p_0 t$  is equal to minus  $\lambda p_1 t$  plus  $\lambda$  times  $e$  to the power minus  $\lambda t$ . Since,  $p_0 t$  we have already obtained is equal to  $e$  to the power minus  $\lambda t$ .

Therefore,  $p_1$  prime  $t$  plus  $\lambda$  times  $p_1 t$  is equal to  $\lambda e$  to the power minus  $\lambda t$  multiplying both sides by  $e$  to the power  $\lambda t$ , which is the integrating factor we have  $e$  to the power  $\lambda t$  into  $p_1$  prime  $t$  plus  $\lambda e$  to the power  $\lambda t$   $p_1 t$  is equal to  $\lambda$  or  $d dt$  of  $e$  to the power  $\lambda t$   $p_1 t$  is equal to  $\lambda$ .

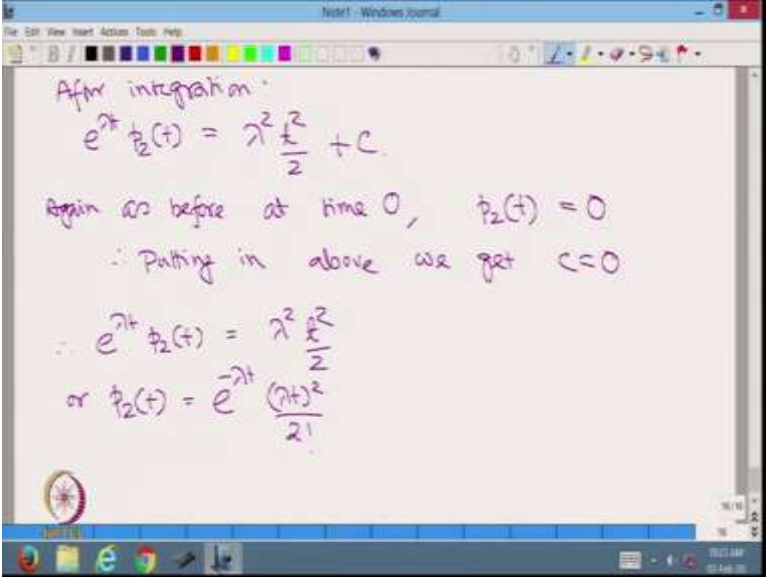
Therefore, after integration what we get is  $e$  to the power  $\lambda t$  into  $p_1 t$  is equal to  $\lambda t$  plus  $C$  at time  $t$  is equal to 0. We know  $e$  to the power of  $\lambda 0$  into  $p_1$  at 0 is equal to  $\lambda 0$  plus  $C$ . Now this is 0, we have already seen at that at time 0, there would not be any arrival implies  $c$  is equal to 0.

Therefore, we have  $e$  to the power  $\lambda t$   $p_1 t$  is equal to  $\lambda t$  or  $p_1 t$  is equal to  $e$  to the power minus  $\lambda t$   $\lambda t$  is equal to  $e$  to the power minus  $\lambda t$   $\lambda t$  to the power one upon factorial 1. Therefore, we get probability of getting one arrival in time  $t$  is equal to  $e$  to the power minus  $\lambda t$  into  $\lambda t$ .

Let us now consider  $P_2 t$ , we have in a similar way  $p_2$  prime  $t$  is equal to minus  $\lambda p_2 t$  plus  $\lambda p_1 t$  or  $P_2$  prime  $t$  plus  $\lambda p_2 t$  is equal to  $\lambda$  into  $e$  to the power minus  $\lambda t$   $\lambda t$ . This is the value of  $p_1 t$  is equal to  $\lambda^2 e$  to the power minus  $\lambda t$  into  $t$ .

Again using  $e$  to the power  $\lambda t$  as integrating factor we have  $e$  to the power  $\lambda t$  into  $p_2$  prime  $t$  plus  $\lambda e$  to the power  $\lambda t$   $P_2 t$  is equal to  $\lambda^2 t$  or  $d dt$  of  $e$  to the power  $\lambda t$   $p_2 t$  is equal to  $\lambda^2 t$ .

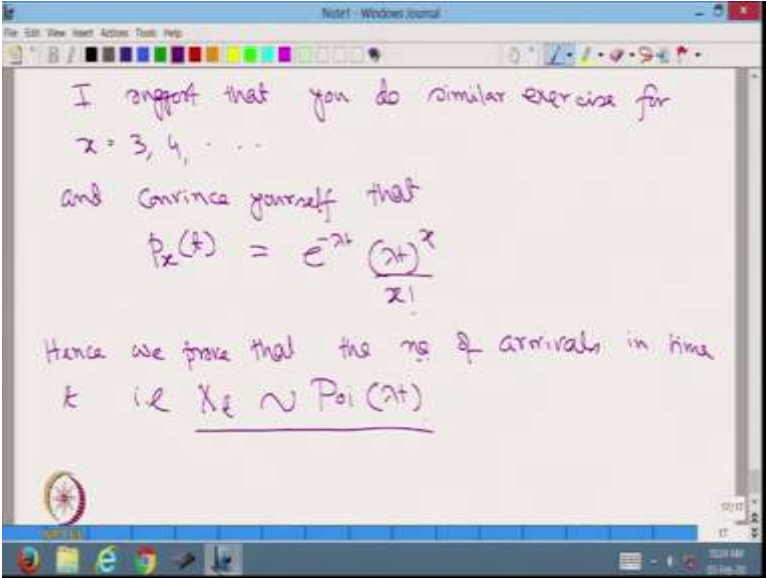
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A screenshot of a Notepad window titled "Notepad - Windows Journal". The window contains handwritten text in purple ink. The text reads: "After integration:", followed by the equation 
$$e^{\lambda t} \frac{t^2}{2} = \lambda^2 \frac{t^2}{2} + C$$
. Below this, it says "Again as before at time 0,  $p_2(t) = 0$ ", followed by "Putting in above we get  $C = 0$ ". Then, it shows 
$$\therefore e^{\lambda t} \frac{t^2}{2} = \lambda^2 \frac{t^2}{2}$$
 and finally 
$$\text{or } p_2(t) = e^{-\lambda t} \frac{(\lambda t)^2}{2!}$$
. The window has a standard toolbar at the top and a taskbar at the bottom.

After integration what we have  $e^{\lambda t} P_2(t)$  is equal to  $\lambda^2 t^2$  by 2 plus  $c$ . Again as before at time 0  $P_2(t)$  is equal to 0. Therefore, putting in above we get  $c$  is equal to 0. Therefore,  $e^{\lambda t} P_2(t)$  is equal to  $\lambda^2 t^2$  by 2 or  $P_2(t)$  is equal to  $e^{-\lambda t} \frac{(\lambda t)^2}{2!}$ .

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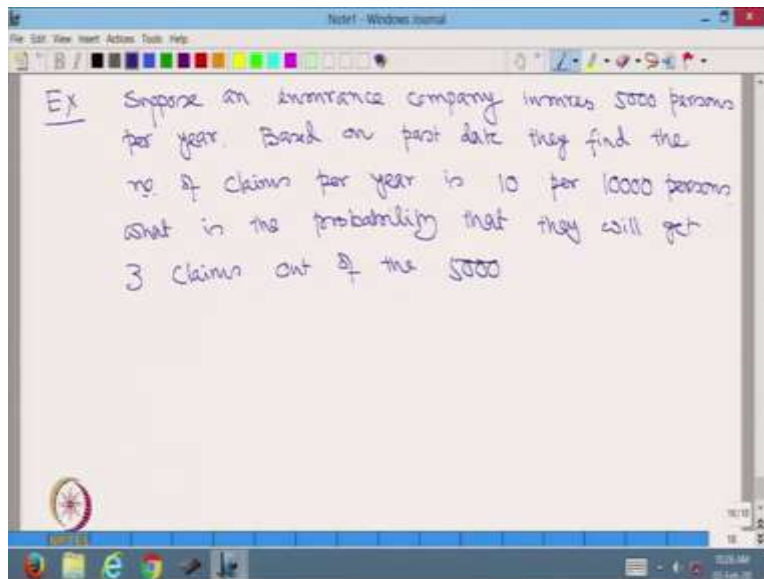
A screenshot of a Notepad window titled "Notepad - Windows Journal". The window contains handwritten text in purple ink. The text reads: "I suggest that you do similar exercise for  $x = 3, 4, \dots$  and convince yourself that", followed by the equation 
$$p_x(t) = e^{-\lambda t} \frac{(\lambda t)^x}{x!}$$
. Below this, it says "Hence we prove that the no. of arrivals in time  $t$  i.e.  $X_t \sim \text{Poi}(\lambda t)$ ". The window has a standard toolbar at the top and a taskbar at the bottom.

I suggest that you do similar exercise for  $x$  is equal to 3, 4, etc. and convince yourself that probability of  $x$  many arrivals in time  $t$  is equal to  $e^{-\lambda t} \frac{(\lambda t)^x}{x!}$ .

power  $x$  upon factorial  $x$ . Hence, we proved that the number of arrivals in time  $t$  that is  $x \propto t$  the random variable is distributed as well so with  $\lambda t$ .

So, very interesting result that we can model the number of arrivals in time  $t$  using a Poisson distribution under of course certain assumptions as we have mentioned earlier.

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Example, suppose an insurance company insures 5000 persons per year based on past data they find the number of claims per year is 10 per 10000 persons. What is the probability that they will get three claims out of the 5000?

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The claim arrival rate can be computed as:

$$\frac{10}{10,000} = 0.001$$

So that is the rate of arrival or  $\lambda$ .

Let us consider each of the new insurance to be one arrival.

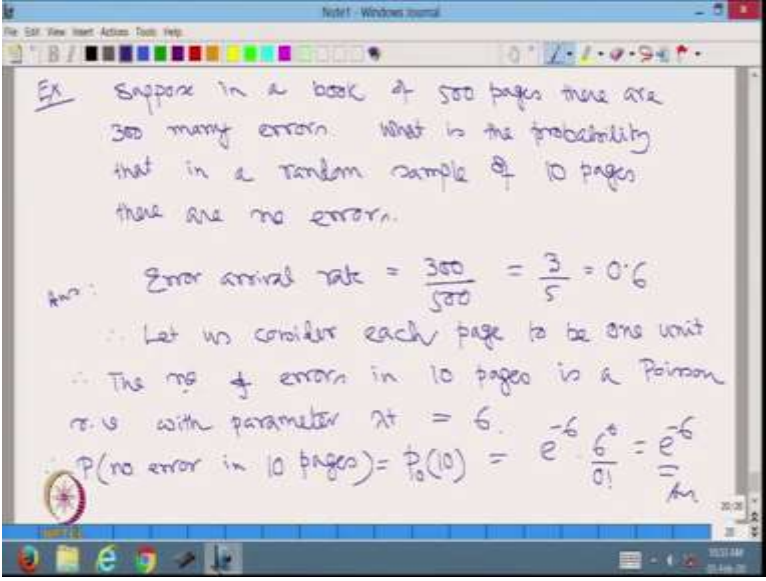
$\therefore$  We want to compute the probability of 3 arrivals out of 5000 or the probability of 3 arrivals when total unit is 5000.

$$\therefore \lambda t = 0.001 \times 5000 = 5$$
$$\therefore P(3 \text{ claim arrival in 5000 new policies}) = e^{-5} \frac{5^3}{3!} = e^{-5} \cdot 20.8$$

Thus we can get the probability using Poisson Distribution.

Solution the claim arrival rate can be computed as 10 upon 10000 is equal to 0.001 so, that is the rate of arrival or lambda. Let us consider each of the new insurance to be one arrival. Therefore, we want to compute the probability of 3 arrivals out of 5000 or the probability of three arrivals when total unit is 5000. Therefore, lambda t is equal to 0.001 multiplied by 5000 is equal to 5, therefore, probability 3 claim arrival in 5000 new policies is equal to e to the power minus 5, 5 to the power 3 upon factorial 3 is equal to e to the power minus 5 into something like say 20.8. Hence, thus we can get the probability using Poisson distribution.

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The image shows a handwritten solution in a Notepad window. The text is as follows:

Ex Suppose in a book of 500 pages there are 300 many errors. What is the probability that in a random sample of 10 pages there are no errors.

Ans: Error arrival rate =  $\frac{300}{500} = \frac{3}{5} = 0.6$

∴ Let us consider each page to be one unit

∴ The no. of errors in 10 pages is a Poisson r.v. with parameter  $\lambda t = 6$ .

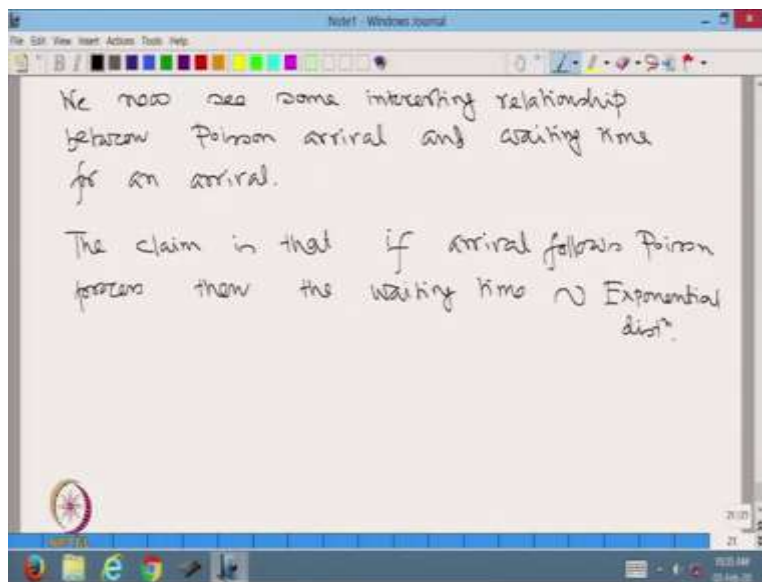
$P(\text{no error in 10 pages}) = P_0(10) = \frac{e^{-6} 6^0}{0!} = \frac{e^{-6}}{1} = \frac{1}{e^6}$

Another example, example suppose in a book of 500 pages there are 300 many errors, what is the probability that in a random sample of 10 pages there are no errors? Answer, we can see that errors arrival rate is equal to 300 upon 500 is equal to 3 by 5 is equal to 0.6. Therefore, let us consider each page to be 1 unit, therefore the number of errors in 10 pages is a Poisson random variable with parameter lambda t is equal to 6.

Therefore, probability no errors in 10 pages is equal to  $P_0$  of 10 is equal to  $e$  to the power minus 6, 6 to the power 0 upon factorial 0 is equal to  $e$  to the power minus 6 that is the answer.

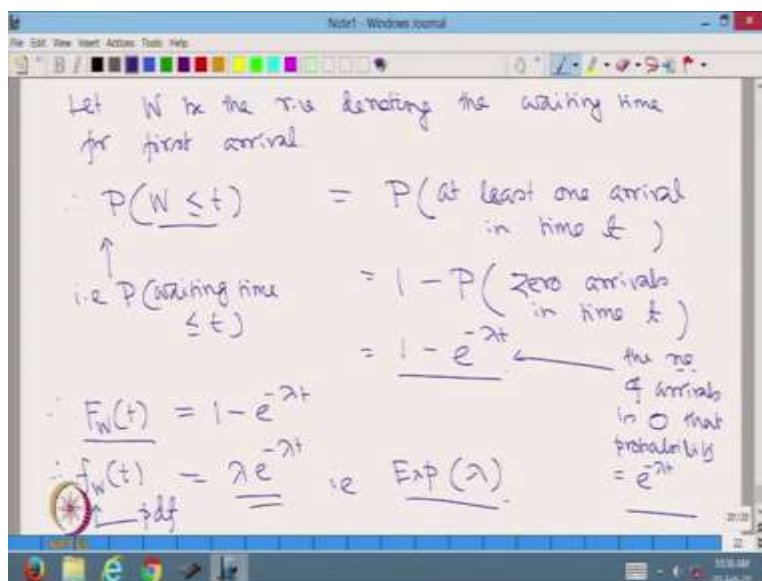


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We now see some interesting relationship between Poisson arrival and waiting time for an arrival, the claim is that if arrival follows Poisson process then the waiting time is exponential distribution.

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Let  $w$  be the random variable denoting the waiting time for first arrival, therefore, probability  $w$  is less than equal to  $t$  that is probability waiting time less than equal to  $t$  is equal to probability of at least one arrival in time  $t$ , because if there is one arrival before time  $t$ , then waiting time has to be less than equal to  $t$  is equal to 1 minus probability of 0 arrivals in time  $t$  is equal to 1 minus  $e$

to the power minus lambda t. This is because the number of arrivals is 0 that probability is e to the power minus lambda t that we have already seen.

Therefore, if w of t the cumulative distribution function is equal to 1 minus e to the power minus lambda t. Therefore f w of t that is the PDF is equal to lambda e to the power minus lambda t that is exponential with lambda. So, this is a very interesting relationship between exponential and Poisson distribution.

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What is the waiting time for k arrivals?  
 Let  $W_k$  be the waiting time for k arrivals.  
 $\therefore F_{W_k}(t) = P(\text{The time required for k arrivals} \leq t)$   
 $= P(\text{At least k arrivals till time t})$   
 $\therefore$  if there are k or more arrivals in time t, then the  $W_k \leq t$

$= 1 - P(0 \text{ arrivals in } t)$   
 $- P(1 \text{ arrival in time } t)$   
 $- \dots$   
 $- P(k-1 \text{ arrivals in time } t)$   
 $= 1 - \left( e^{-\lambda t} - \lambda t e^{-\lambda t} - \frac{e^{-\lambda t} (\lambda t)^2}{2!} - \dots - \frac{e^{-\lambda t} (\lambda t)^{k-1}}{(k-1)!} \right)$   
 $\therefore f_{W_k}(t) = \frac{d}{dt} (e^{-\lambda t}) - \frac{d}{dt} \left( \sum_{i=1}^{k-1} e^{-\lambda t} \frac{(\lambda t)^i}{i!} \right)$



Therefore,  $f_w k t$  this we get by differentiating this term is equal to  $d dt$  of  $\text{minus } e \text{ to the power } \text{minus } \lambda t \text{ minus } d dt \text{ of } \sigma_i$  is equal to  $1 \text{ to } k \text{ minus } 1 e \text{ to the power } \text{minus } \lambda t$  into  $\lambda t \text{ to the power } i \text{ upon factorial } i$ . So, I have separated out this from the rest, because it has only one term involving  $t$  others have two terms involving  $t$ .

Therefore,  $f_w k t$  is equal to  $\text{minus } d dt \text{ of } e \text{ to the power } \text{minus } \lambda t \text{ minus } \sigma_i$  is equal to  $1 \text{ to } k \text{ minus } 1 d dt \text{ of } e \text{ to the power } \text{minus } \lambda t \lambda t \text{ to the power } i \text{ upon factorial } i$ , because, with the derivative of the sum is equal to sum of derivatives. Now,  $d dt \text{ of } e \text{ to the power } \text{minus } \lambda t$  is equal to  $\text{minus } \lambda e \text{ to the power } \text{minus } \lambda t$  therefore,  $\text{minus } d dt \text{ of } e \text{ to the power } \text{minus } \lambda t$  is equal to  $\lambda e \text{ to the power } \text{minus } \lambda t$ .

Now,  $d dt \text{ of } e \text{ to the power } \text{minus } \lambda t, \lambda t \text{ to the power } i \text{ upon factorial } i$  is equal to  $e \text{ to the power } \text{minus } \lambda t, \lambda t \text{ to the power } i \text{ minus } 1 \text{ upon } i \text{ factorial}$  then  $i$  will come into picture because  $i$  into  $\lambda t \text{ to the power } i \text{ minus } 1$  multiplied by  $\lambda$  plus  $\lambda t \text{ to the power } i \text{ upon factorial } i$  into  $\text{minus } \lambda e \text{ to the power } \text{minus } \lambda t$  is equal to  $\lambda e \text{ to the power } \text{minus } \lambda t, \lambda t \text{ to the power } i \text{ minus one upon } i \text{ minus one factorial}$ .

Because, this  $i$  cancels with  $1 i$ ,  $\text{minus } \lambda e \text{ to the power } \text{minus } \lambda t \lambda t \text{ to the power } i \text{ upon factorial } i$  therefore,  $d dt \text{ of } \lambda t e \text{ to the power } \text{minus } \lambda t$  is equal to  $\lambda e \text{ to the power } \text{minus } \lambda t \text{ minus } \lambda e \text{ to the power } \text{minus } \lambda t \text{ into } \lambda t \text{ to the power } 1 \text{ upon factorial } 1$ . Therefore,  $d dt \text{ of } e \text{ to the power } \text{minus } \lambda t \lambda t \text{ to the power } 2 \text{ upon factorial } 2$  is equal to  $\lambda e \text{ to the power } \text{minus } \lambda t \text{ into } \lambda t \text{ to the power } 1 \text{ upon factorial } 1 \text{ minus } \lambda e \text{ to the power } \text{minus } \lambda t \lambda t \text{ to the power } 2 \text{ upon factorial } 2$ .

In a similar way  $d dt \text{ of } e \text{ to the power } \text{minus } \lambda t^2, \lambda t \lambda t^2 \text{ upon factorial } 3$  is equal to  $\lambda e \text{ to the power } \text{minus } \lambda t \text{ into } \lambda t^2 \text{ upon factorial } 2 \text{ minus } \lambda e \text{ to the power } \text{minus } \lambda t \lambda t^2 \text{ upon factorial } 3$  like that if we go then the last term is  $d dt \text{ of } e \text{ to the power } \text{minus } \lambda t, \lambda t \text{ to the power } k \text{ minus } 1 \text{ upon } k \text{ minus } 1 \text{ factorial}$  is equal to.

Let us write the last term, this is going to be  $\lambda e \text{ to the power } \text{minus } \lambda t \lambda t \text{ to the power } k \text{ minus } 1 \text{ upon } k \text{ minus } 1 \text{ factorial}$ . Now, we note that, this cancels with this, this will cancel with this finally this term will be cancelled. Therefore, the whole result is coming out to

be  $\lambda e^{-\lambda t}$  to the power  $k - 1$  minus  $\lambda e^{-\lambda t}$  to the power  $k - 1$  upon  $(k - 1)!$ .

Now, let us look at this entire term we have computed this derivative but it will have a minus sign and this  $d/dt$  of  $\lambda e^{-\lambda t}$  that we have already calculated to be  $-\lambda e^{-\lambda t}$  to the power  $k - 1$ . Therefore, minus of summation  $i$  is equal to  $1$  to  $k - 1$   $d/dt$  of  $\lambda e^{-\lambda t}$  to the power  $k - 1$  upon  $(k - 1)!$  is equal to minus  $\lambda e^{-\lambda t}$  to the power  $k - 1$  plus  $\lambda e^{-\lambda t}$  to the power  $k - 1$  upon  $(k - 1)!$ .

Now, this term will cancel with this because this is  $\lambda e^{-\lambda t}$  to the power  $k - 1$  and this is minus  $\lambda e^{-\lambda t}$  to the power  $k - 1$ .

(Refer Slide Time: 58:06)

The image shows a handwritten derivation on a digital whiteboard. The steps are as follows:

$$f_{W_k}(t) = \frac{\lambda e^{-\lambda t} (\lambda t)^{k-1}}{(k-1)!}$$

$$= \frac{\lambda^k}{(k-1)!} e^{-\lambda t} t^{k-1}$$

Then, it states:  $N! (k-1)! = \Gamma k$

$$\therefore f_{W_k}(t) = \frac{\lambda^k}{\Gamma k} e^{-\lambda t} t^{k-1}$$

To the right of the boxed equation, it says: "i.e. waiting time for k arrivals ~  $\Gamma(\lambda, k)$ " with a checkmark below it.

Therefore, if  $w_k(t)$  is equal to what we get is equal to  $\lambda e^{-\lambda t}$  to the power  $k - 1$  minus  $\lambda e^{-\lambda t}$  to the power  $k - 1$  upon  $(k - 1)!$  is equal to  $\lambda$  to the power  $k$  upon  $(k - 1)!$   $e^{-\lambda t}$  to the power  $k - 1$   $t$  to the power  $k - 1$ . Now  $(k - 1)!$  is equal to  $\Gamma k$  this we have seen earlier.

Therefore,  $f_{W_k}(t)$  is equal to  $\lambda$  to the power  $k$  upon  $\Gamma k$   $e^{-\lambda t}$  to the power  $k - 1$   $t$  to the power  $k - 1$  that is waiting time for  $k$  arrivals distributed as gamma with  $\lambda$  and  $k$ .

This is not very surprising, since sum of exponential with same parameter  $\lambda$  follows gamma distribution. Therefore, this was something we would expect and we have got that result.

Okay friends, I stopped here today. In the next class, I shall start with conditional expectation and variance and also we shall see some important inequalities like Chebyshev's Inequality and Markov Inequality. Okay then thank you so much.