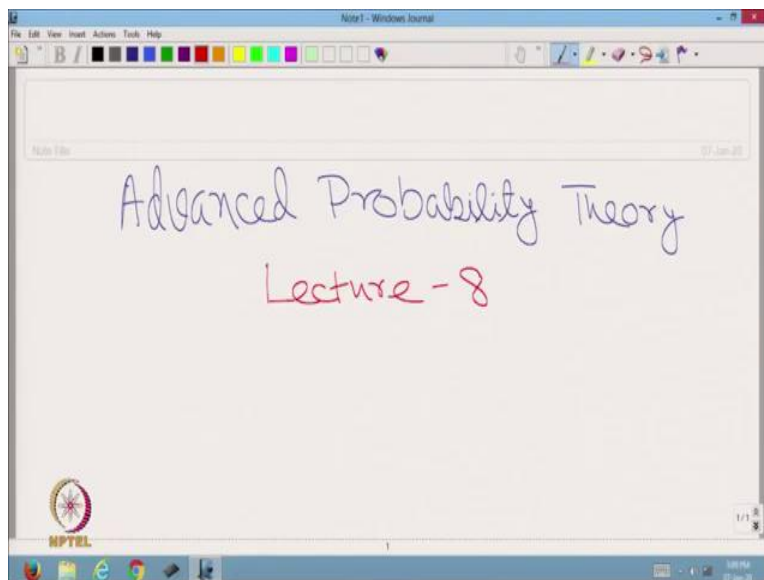
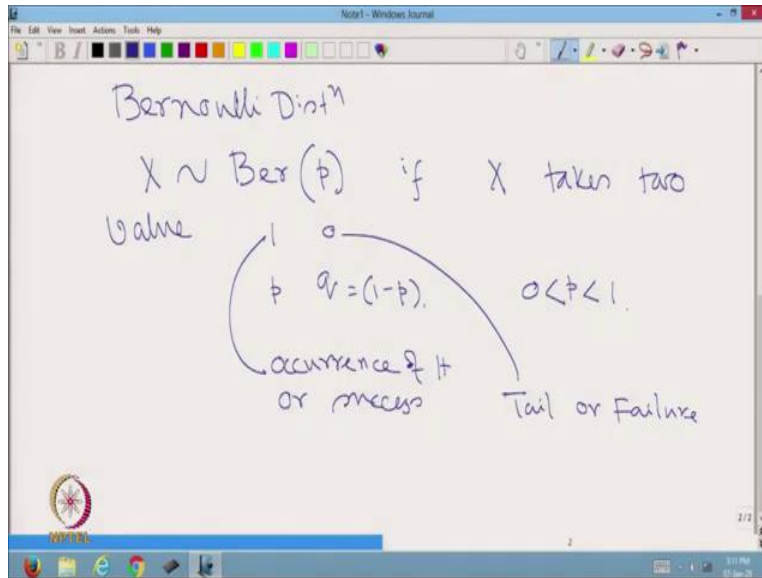


**Advanced Probability Theory**  
**Professor Niladri Chatterjee**  
**Department of Mathematics,**  
**Indian Institute of Technology, Delhi**  
**Lecture 8**

Welcome students. The MOOC's lecture series on Advanced Probability Theory. This is lecture number eight. In the last lecture, if you remember, we have started discussions on discrete random variables. In particular, we have studied binomial distribution and Poisson distribution. In today's lecture, we shall look at a few more discrete random variables or discrete distributions, which are very important from mathematical as well as application point of view.

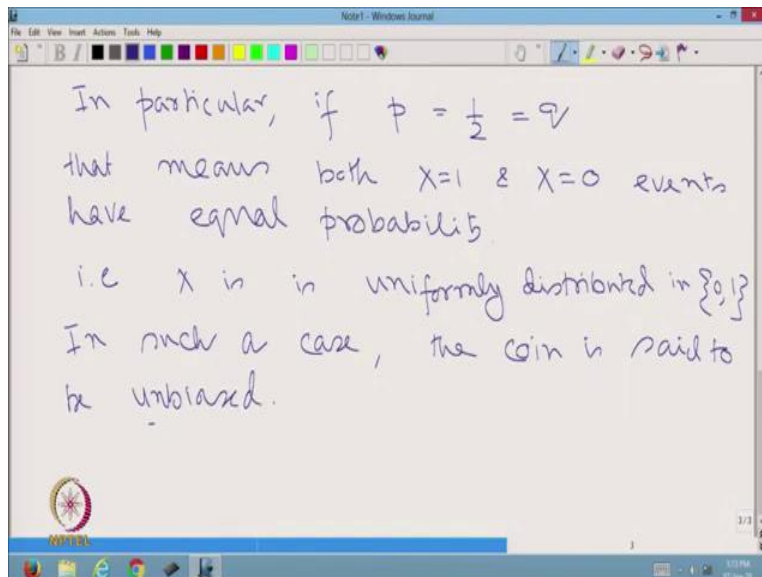
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So, if you remember, we started with Bernoulli distribution that is  $x$  is distributed as Bernoulli  $p$  if  $x$  takes 2 values, one with probability  $p$  and 0 with probability  $q$ , which is equal to 1 minus  $p$ , 0 less than  $p$  less than 1. So, this we have used to model the toss of a coin, whereby 1 we may occurrence of head or a success and by 0 we mean tail or failure. This we have discussed already.

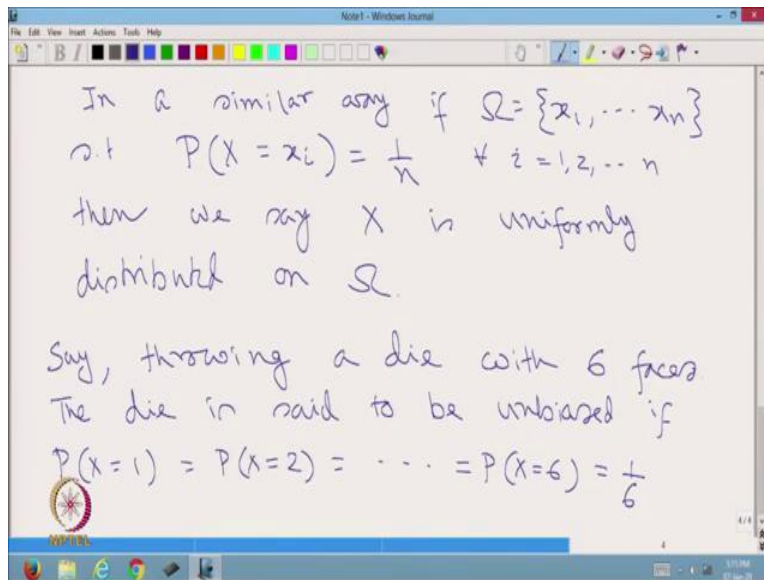
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In particular, if  $p$  is equal to half is equal to  $q$  that means both  $x$  is equal to 1, and  $x$  is equal to 0. These two events have equal probability or in other words,  $x$  is uniformly

distributed in 0 and 1, this is a set of two values and in such a case the coin is called, the coin mean the coin that is being used for tossing unbiased.

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In a similar way, if  $\omega$  is equal to  $x_1, x_2, \dots, x_n$  such that probability  $X$  is equal to  $x_i$  is equal to  $\frac{1}{n}$  for all  $i$  is equal to  $1, 2$  up to  $n$ . Then we say  $X$  is uniformly distributed on  $\omega$ . Say for example, you are throwing a die with 6 faces. The die is said to be unbiased if probability  $X$  is equal to 1 is equal to probability  $X$  is equal to 2 is equal to probability  $X$  is equal to 6 is equal to  $\frac{1}{6}$ .

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Geometric Distribution.

Let  $X$  be a r.v defined over all positive integers:

$X:$	1	2	3	...	K	...
	$p$	$pq$	$pq^2$	...	$pq^{K-1}$	...

We can see that  $P(X=x)$  is  $p_x$  is  $> 0$   
 $\forall x = 1, 2, \dots, K, \dots$

$0 < p < 1$  &  $q = 1 - p$

Now, let us study under distribution, which is called Geometric Distribution. Let  $X$  be a random variable defined over all positive integers. So,  $x$  takes the values 1, 2, 3,  $K$  and it goes to infinity, probability  $X$  is equal to 1 is  $p$ , probability  $X$  is equal to 2 is  $pq$ , probability  $X$  is equal to 3 is  $pq$  square like that probability  $X$  is equal to  $K$  is  $pq$  to the power  $K$  minus 1 here, the first question is it a valid probability distribution?

So, let us assume that  $0 < p < 1$  and  $q$  is equal to  $1 - p$ . Therefore, we can see that probability  $X$  is equal to  $x$  which we may denote as  $p_x$  is greater than 0 for all  $x$  is equal to 1 to  $K$  up to infinity, because  $p$  is positive,  $q$  is positive, so their product is positive.

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The image shows a Notepad window with handwritten text in purple ink. The text is as follows:

We need to show  $\sum_{x=1}^{\infty} p_x = 1$

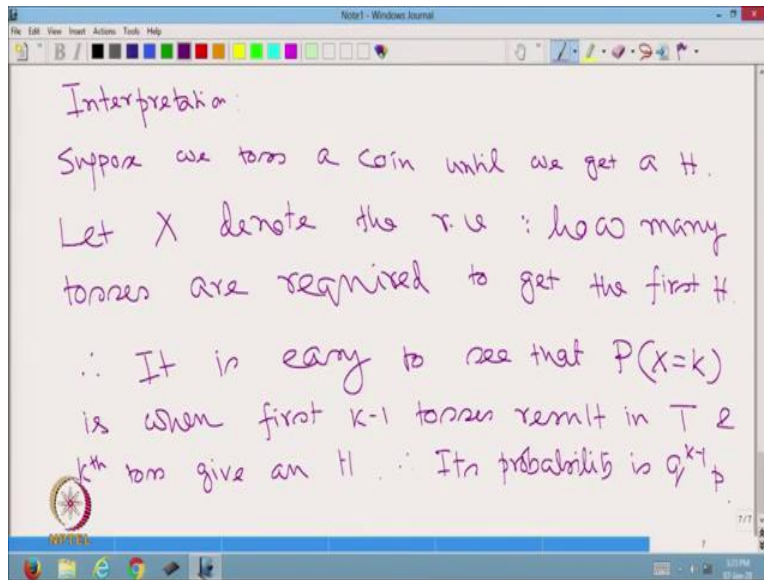
Since  $p_x = p q^{x-1}$

$$\therefore \sum_{x=1}^{\infty} p_x = p + pq + \dots + pq^{x-1} + \dots$$
$$= p(1 + q + q^2 + \dots + q^{x-1} + \dots)$$
$$= p \cdot \frac{1}{1-q} = p \cdot \frac{1}{p} = 1$$

$\therefore$  The above is a valid probability Distribution

Next we need to show  $\sum_{x=1}^{\infty} p_x$  is equal to 1, that is very easy to show since  $p_k$  is equal to  $p q$  to the power  $K$  minus 1. Therefore,  $\sum_{x=1}^{\infty} p_x$  is equal to  $p$  plus  $pq$  plus the  $p q$  to the power  $K$  minus 1 up to infinity is equal to  $p$  into  $1 + q + q$  square plus  $q$  to the power  $K$  minus 1 plus up to infinity. Now, this is a geometric series. Therefore, we know that its sum is going to be  $1$  upon  $1$  minus  $q$  is equal to  $p$  into  $1$  upon  $p$  is equal to  $1$ . Therefore, the above is a valid probability distribution.

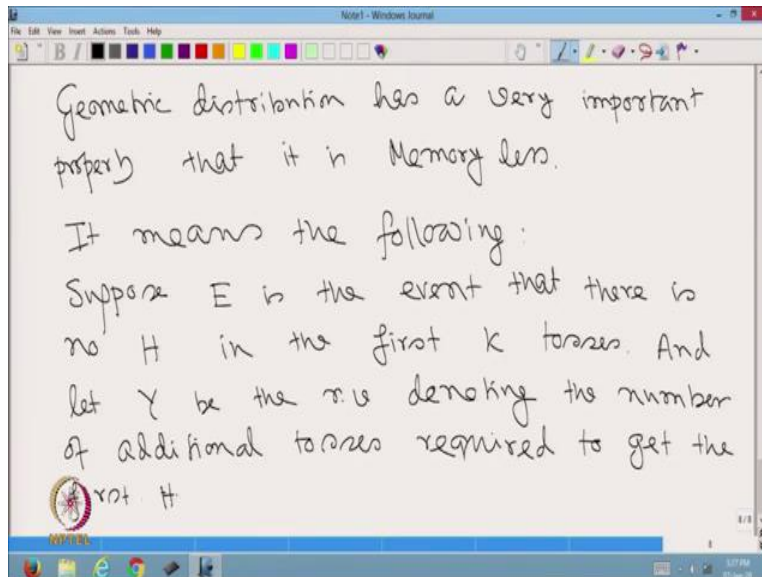
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So, its interpretation. Suppose, we toss a coin until we get a Head. So, how many tosses are required? Let  $X$  denote the random variable, how many tosses are required to get the first Head. So, it may take 1, when the very first toss we get head and obviously its probability is  $p$ , if we get Tail at the first toss that probability is  $q$  and then a head in the second toss then that probability is  $p$ . Therefore, we need two tosses to get the first head and its probability is therefore  $p q$ .

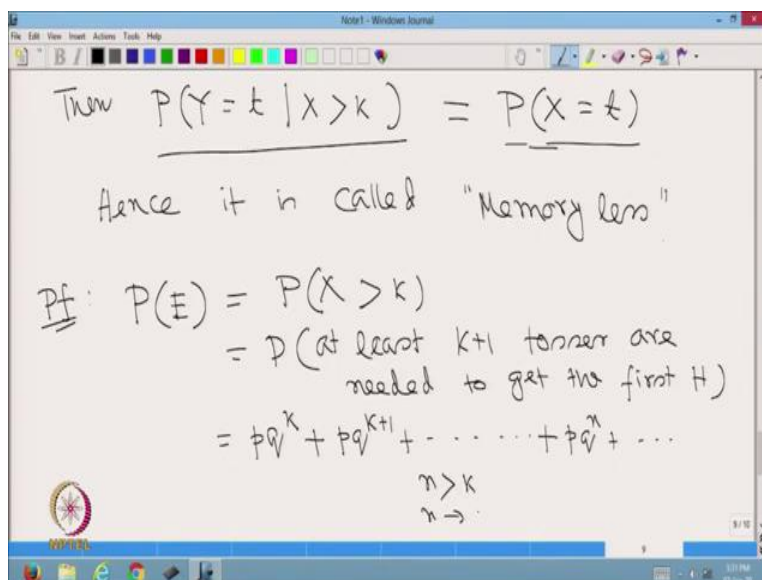
Similarly, if the value of  $X$  is  $K$  that means that we are tossing the coin  $K$  times to get the first head implies that first  $q$  minus 1 tosses are going to be tail and the  $K^{\text{th}}$  toss is going to be head. Therefore, it is easy to see that probability  $X$  is equal to  $K$  is when first  $K$  minus 1 tosses resulting tail and  $K^{\text{th}}$  toss gives them H therefore, its probability is  $q$  to the power  $K$  minus 1 into  $p$ . So, that is the basic interpretation of Geometric distribution.

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Now, geometric is very important, has a very important property that it is memory less. What does it mean? It means the following: Suppose  $E$  is the event that there is no head in the first  $K$  tosses. And let  $Y$  be the random variable denoting the number of additional tosses required to get the first head.

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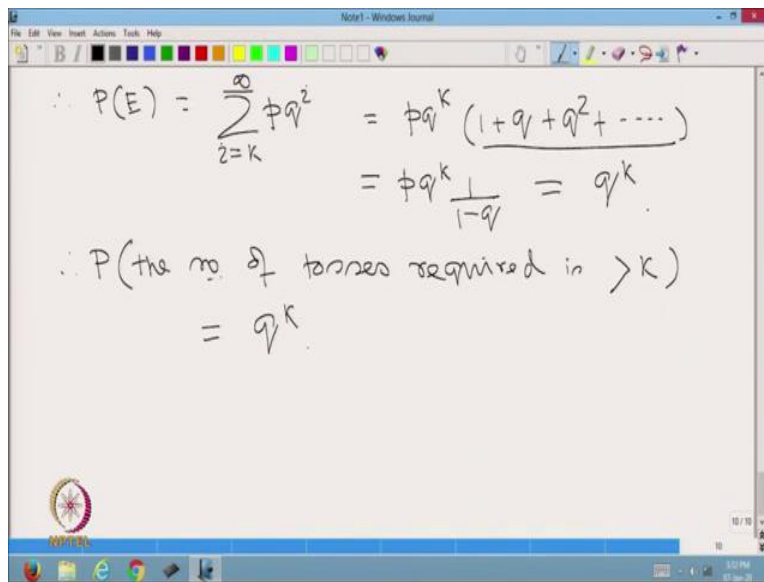


Then probability Y is equal to t given x is greater than k is same as probability X is equal to t. So, let me explain this expression, given that no head has occurred in the first K tosses. Therefore, X has to be greater than K and Y denotes the number of additional tosses needed to get the first head. So, it says that the probability of having another t tosses to get the first head, given that X is greater than k.

That is same as from the beginning we need t tosses to get the first head or in other words, that we have already consumed K many tosses, it has no effect on the probability of what is going to happen subsequently. Hence, it is called memory less. So, let us prove that so, what is the probability of event E? It is probability X is greater than K is equal to probability at least K plus 1 tosses are needed to get the first head.

Is equal to pq to the power K that we need K plus 1 tosses to get the first head plus pq to the power K plus 1 denoting that we need K plus too many tosses to get the first head, where n is greater than K and n is going to infinity.

(Refer Slide Time: 17:20)



$$\begin{aligned} \therefore P(E) &= \sum_{i=K}^{\infty} p q^i = p q^K (1 + q + q^2 + \dots) \\ &= p q^K \frac{1}{1-q} = q^K. \end{aligned}$$

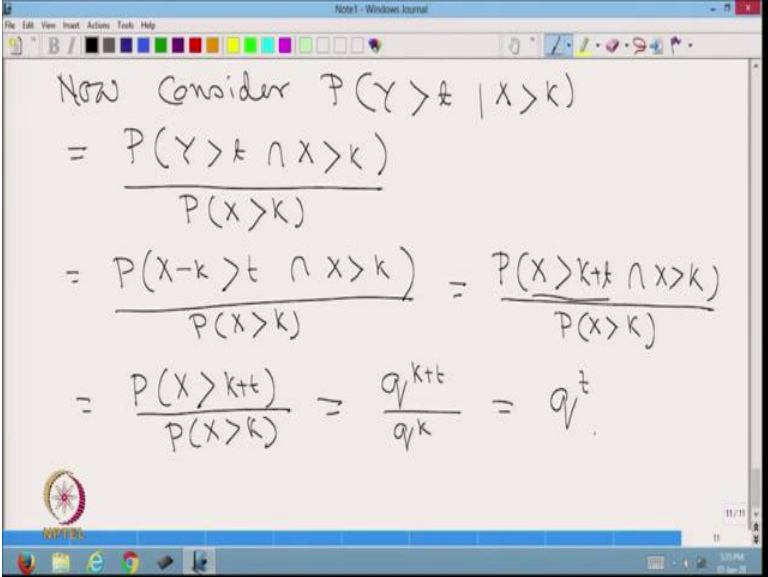
$$\therefore P(\text{the no. of tosses required is } > K) = q^K.$$

Therefore, probability of E is equal to sigma pq to the power i, i is equal to K to infinity is equal to pq to the power K into 1 plus q plus q square up to infinity is equal to pq to the power K, this sum we already know is equal to 1 upon 1 minus q which is equal to q to



the power K. Therefore, probability the number of tosses required is greater than K is equal to q to the power K.

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Now Consider  $P(Y > t | X > k)$

$$= \frac{P(Y > t \cap X > k)}{P(X > k)}$$

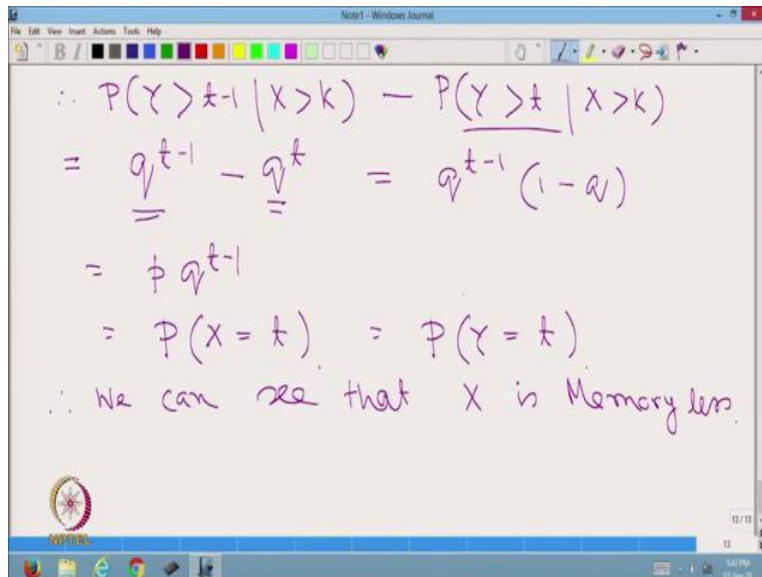
$$= \frac{P(X - k > t \cap X > k)}{P(X > k)} = \frac{P(X > k + t \cap X > k)}{P(X > k)}$$

$$= \frac{P(X > k + t)}{P(X > k)} = \frac{q^{k+t}}{q^k} = q^t$$

Now, let us consider probability Y is greater than t given that X is greater than K. This conditional probability we know is equal to probability Y greater than t and X is greater than K upon probability X greater than K is equal to probability X minus K greater than t and X greater than K upon probability X greater than K is equal to probability X greater than K plus t and X greater than K divided by probability X greater than K. Now, probability X greater than K plus t and X greater than K, these means they are intersection is this event.

Therefore, this is going to be probability X greater than K plus t divided by probability X greater than K. Now, probability X greater than K, we have already calculated to be q to the power K. Therefore, probability X greater than K plus t is going to be q to the power K plus t divided by q to the power K is equal to q to the power t. So, probability Y greater than t given X is greater than K, that probability is coming out to be q to the power t.

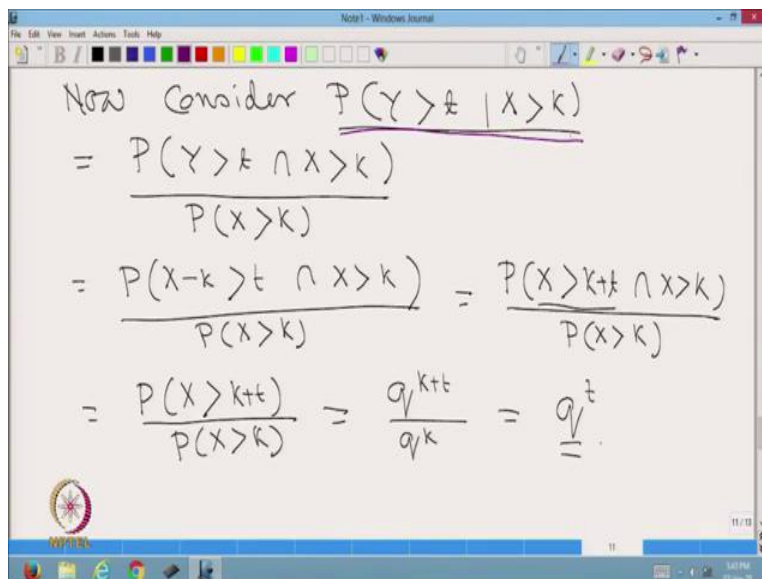
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Handwritten derivation in a Notepad window:

$$\begin{aligned} \therefore P(Y > t-1 | X > k) - P(Y > t | X > k) \\ &= \underline{q^{t-1}} - \underline{q^t} = q^{t-1}(1-q) \\ &= p q^{t-1} \\ &= P(X = t) = P(Y = t) \end{aligned}$$

$\therefore$  We can see that  $X$  is Memoryless.



Handwritten derivation in a Notepad window:

Now Consider  $P(Y > t | X > k)$

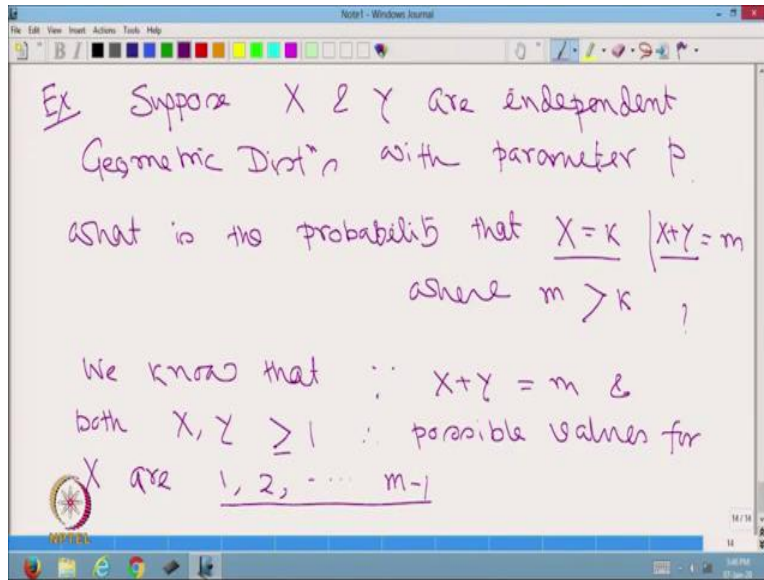
$$\begin{aligned} &= \frac{P(Y > t \cap X > k)}{P(X > k)} \\ &= \frac{P(X - k > t \cap X > k)}{P(X > k)} = \frac{P(X > k+t \cap X > k)}{P(X > k)} \\ &= \frac{P(X > k+t)}{P(X > k)} = \frac{q^{k+t}}{q^k} = \underline{q^t} \end{aligned}$$

Therefore, probability  $Y$  greater than  $t$  minus 1 given  $X$  greater than  $K$  minus probability  $Y$  greater than  $t$  given  $X$  greater than  $K$  is equal to  $q$  to the power  $t$  minus 1 minus  $q$  to the power  $t$ . This is because we have seen that probability  $Y$  greater than  $t$  given  $X$  greater than  $K$  is  $q$  to the power  $t$ .

Therefore,  $Y$  greater than  $t$  minus 1 given  $X$  greater than  $K$  is  $q$  to the power  $t$  minus 1, in a very similar way, this is  $q$  to the power  $t$ . Therefore, this is  $q$  to the power  $t$  minus 1 to 1

minus  $q$  is equal to  $pq$  to the power  $t$  minus 1 which is equal to probability  $X$  is equal to  $t$  is equal to probability  $Y$  is equal to  $t$ . Therefore, we can see that  $X$  is memory less.

(Refer Slide Time: 22:03)



Another example suppose,  $X$  and  $Y$  are independent geometric distribution with parameter  $P$ . What is the probability that  $X$  is equal to  $K$  given that  $X$  plus  $Y$  is equal to  $m$  where  $m$  is greater than  $K$ ? So, we are given that the sum of two geometric random variables  $X$  and  $Y$  is  $m$ , given that what is the probability that  $X$  is equal to  $K$ ? We know that since,  $X$  plus  $Y$  is equal to  $m$  and both  $X$  and  $Y$  are greater than equal to 1, therefore, possible values for  $X$  are 1, 2, 3 up to  $m$  minus 1. So,  $X$  can take one of the possible values between 1, 2, 3 up to  $m$  minus 1.

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$$P(X=1 | X+Y=m) = ?$$

$$X+Y \text{ can take value } m \text{ in the following way:}$$

$$X=1 \text{ \& } Y=m-1 \quad X=2 \text{ \& } Y=m-2 \quad \dots \quad X=m-1 \text{ \& } Y=1$$

$$\downarrow$$

$$p \cdot pq^{m-2} \quad pq \cdot pq^{m-3} \quad \dots \quad pq^{m-2} \cdot p$$

$$= p^2 q^{m-2} \quad p^2 q^{m-2} \quad \dots \quad p^2 q^{m-2}$$

$$\therefore P(X+Y=m) = \sum_{x=1}^{m-1} p^2 q^{m-2} = (m-1) p^2 q^{m-2}$$

So, probability  $X$  is equal to 1, given  $X$  plus  $Y$  is equal to  $m$  is equal to how much?  $X$  plus  $Y$  can take value in the following ways.  $X$  is equal to 1, and  $Y$  is equal to  $m$  minus 1,  $X$  is equal to 2, and  $y$  is equal to  $m$  minus 2, up to  $X$  is equal to  $m$  minus 1 and  $Y$  is equal to 1. So, what is this probability? It is  $p$  multiplied by  $pq$  to the power  $m$  minus 2, this is  $pq$  multiplied by  $pq$  to the power  $m$  minus 3.

This is  $pq$  to the power  $m$  minus 2 multiplied by  $p$  is equal to, if you note that all of them are  $p$  square  $q$  to the power minus 2, this is also  $p$  square  $q$  to the power  $m$  minus 2 and this is also  $p$  square  $q$  to the power minus 2. Therefore, probability  $X$  plus  $Y$  is equal to  $m$  has the probability  $\sigma p$  square  $q$  to the power  $m$  minus 2,  $x$  is equal to 1 to  $m$  minus 1 is equal to  $m$  minus 1 into  $p$  square  $q$  to the power  $m$  minus 2.

(Refer Slide Time: 26:36)

$$\begin{aligned} \therefore P(X=1 | X+Y=m) &= \frac{P(X=1 \wedge Y=m-1)}{P(X+Y=m)} = \frac{p \cdot pq^{m-2}}{(m-1)p^2q^{m-2}} = \frac{1}{m-1} \end{aligned}$$

In a similar way :

$$\begin{aligned} P(X=2 | X+Y=m) &= \frac{P(X=2 \wedge Y=m-2)}{P(X+Y=m)} \\ &= \frac{pq \cdot pq^{m-3}}{(m-1)p^2q^{m-2}} = \frac{1}{m-1} \end{aligned}$$

$$P(X=1 | X+Y=m) = ?$$

$X+Y$  can take value  $m$  in the following way :

$X=1 \text{ \& } Y=m-1$	$X=2 \text{ \& } Y=m-2$	$\dots$	$X=m-1 \text{ \& } Y=1$
$\downarrow$			
$p \cdot pq^{m-2}$	$pq \cdot pq^{m-3}$	$\dots$	$pq^{m-2} \cdot p$
$= p^2q^{m-2}$	$p^2q^{m-2}$	$\dots$	$p^2q^{m-2}$

$$\therefore P(X+Y=m) = \sum_{x=1}^{m-1} p^2q^{m-2} = (m-1)p^2q^{m-2}$$

Therefore, probability  $X$  is equal to 1 given  $X$  plus  $Y$  is equal to  $m$  is equal to probability  $X$  is equal to 1, and  $Y$  is equal to  $m$  minus 1 divided by probability  $X$  plus  $Y$  is equal to  $m$  is equal to  $p$  multiplied by  $p$  into  $q$  to the power  $m$  minus 2 divided by this probability  $m$  minus 1 into  $p$  square into  $q$  minus 2. So,  $m$  minus 1 into  $p$  square into  $q$  to the power  $m$  minus 2 is equal to 1 upon  $m$  minus 1.

In a similar way probability  $X$  is equal to 2 given  $X$  plus  $Y$  is equal to  $m$  is same as probability  $X$  is equal to 2 and  $Y$  is equal to  $m$  minus 2 upon probability  $X$  plus  $Y$  is

equal to  $m$  is equal to  $pq$  multiplied by  $pq$  to the power  $m$  minus 3 upon  $m$  minus 1 into  $p$  square into  $q$  to the power  $m$  minus 2 is equal to 1 upon  $m$  minus 1.

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Ex Suppose  $X$  &  $Y$  are independent Geometric Dist<sup>n</sup> with parameter  $p$ .  
 what is the probability that  $X=k \mid X+Y=m$  where  $m > k$  ?  
 We know that  $X+Y=m$  & both  $X, Y \geq 1$ . possible values for  $X$  are 1, 2, ..., m-1

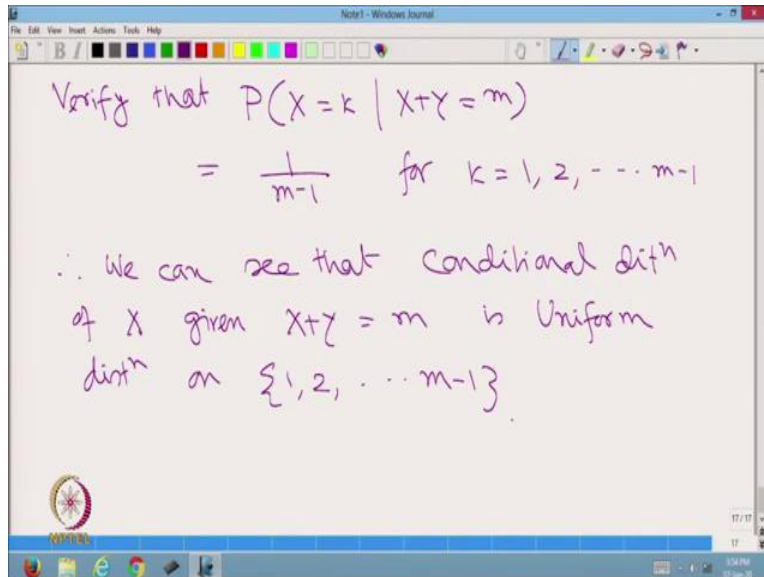
We have observed that  $X$  can take values between 1, 2, 3 up to  $m$  minus 1.

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$\therefore P(X=1 \mid X+Y=m)$   
 $= \frac{P(X=1 \wedge Y=m-1)}{P(X+Y=m)} = \frac{p \cdot pq^{m-2}}{(m-1)p^2q^{m-2}} = \frac{1}{m-1}$   
 In a similar way :  
 $P(X=2 \mid X+Y=m) = \frac{P(X=2 \wedge Y=m-2)}{P(X+Y=m)}$   
 $= \frac{pq \cdot pq^{m-3}}{(m-1)p^2q^{m-2}} = \frac{1}{m-1}$

Now, we have seen that probability  $X$  is equal to 1 given  $X$  plus  $Y$  is equal to  $m$  is  $1$  upon  $m$  minus  $1$ , probability  $X$  is equal to 2 given  $X$  plus  $Y$  is equal to  $m$  that is also  $1$  upon  $m$  minus  $1$ .

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The image shows a digital whiteboard interface with a menu bar (File, Edit, View, Insert, Actions, Tools, Help) and a toolbar with various drawing tools. The main area contains handwritten text in purple ink. The text reads: 'Verify that  $P(X=k | X+Y=m) = \frac{1}{m-1}$  for  $k=1, 2, \dots, m-1$ '. Below this, it says: '∴ We can see that Conditional dist<sup>n</sup> of  $X$  given  $X+Y=m$  is Uniform dist<sup>n</sup> on  $\{1, 2, \dots, m-1\}$ '. At the bottom left of the whiteboard is a small circular logo with a star. The bottom of the screen shows a Windows taskbar with icons for Start, Internet Explorer, and other applications, along with a system clock showing 17:17 on 17/11/2014.

Verify that  $P(X=k | X+Y=m)$   
 $= \frac{1}{m-1}$  for  $k=1, 2, \dots, m-1$

∴ We can see that Conditional dist<sup>n</sup>  
of  $X$  given  $X+Y=m$  is Uniform  
dist<sup>n</sup> on  $\{1, 2, \dots, m-1\}$ .

Verify that probability  $X$  is equal to  $k$  given  $X$  plus  $Y$  is equal to  $m$  is equal to  $1$  upon  $m$  minus  $1$ , for  $k$  is equal to  $1, 2, 3$  up to  $m$  minus  $1$ . Therefore, we can see conditional distribution of  $X$  given  $X$  plus  $Y$  is equal to  $m$  is uniform distribution on  $1, 2$  up to  $m$  minus  $1$ . So, that is another interesting property for geometric distribution.



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We now consider another discrete r.v.  
Negative Binomial

A r.v.  $X$  is said to have a Negative Binomial Distribution, denoted as  $NBD(p, k)$  if it takes the values as follows:

$X = 0 \quad 1 \quad 2 \quad \dots \quad n \quad \dots \quad \text{i.e all possible integers } \geq 0$

$P(X=n)$

$\binom{n+k-1}{k-1} p^k q^n$

Now, let me move further. We now consider another discrete random variable, namely Negative Binomial. So, a random variable  $X$  is said to have a negative binomial distribution denoted as NBD, with two parameters  $p$  and  $k$  as follows. So,  $X$  takes the values 0, 1, 2 then up to infinity that is all possible integers greater than equal to 0. So,  $X$  takes the value of 0, 1, 2 up to  $n$ , all possible integers and the probability  $X$  is equal to  $n$  is  $n$  plus  $k$  minus 1, see  $k$  minus 1  $p$  to the power  $k$   $q$  to the power  $n$ .

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How do we interpret?

$X$  takes the value  $n$  with probability

$\binom{n+k-1}{k-1} p^k q^n$

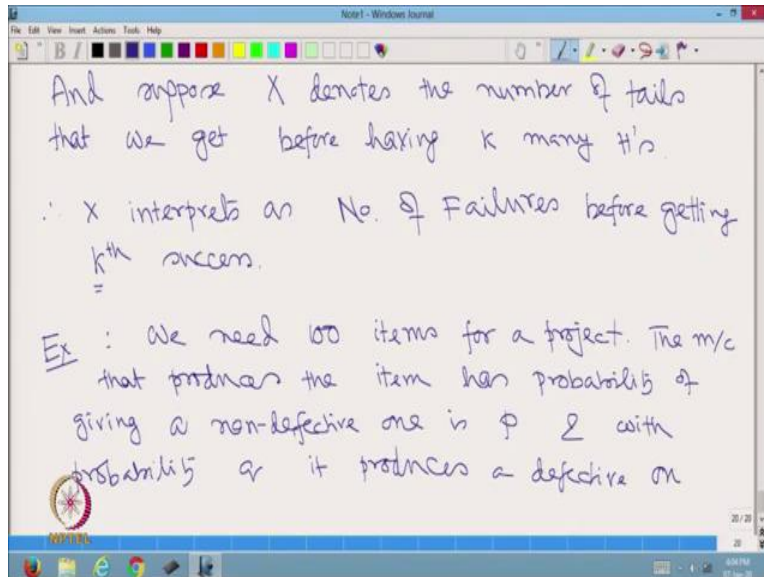
Note that  $k \geq 1$   $p$  are parameters of the distn

Suppose we have a coin with Prob of getting H is  $p$ .



What does it mean? So,  $X$  takes the value  $n$  with probability  $n$  plus  $k$  minus 1, see  $k$  minus 1,  $p$  to the power  $k$   $q$  to the power  $n$ . Note that  $k$  and  $p$  are parameters of the distribution. So, suppose we have a coin with probability of getting head is  $p$ .

(Refer Slide Time: 34:16)



And suppose  $X$  denotes the number of failures or number of tails that we get before having  $k$  many heads. So, we have a coin, we keep on tossing and we keep noting the outcomes, we keep doing that until we get  $k$  heads and  $X$  denotes the number of tails that we have got in that process.

Therefore,  $X$  interprets as number of failures before getting  $k^{\text{th}}$  success so  $k$  is parameter of the distribution, so that has been fixed. Say, for example, we need 100 items for a project. The machine that produces the item has probability of giving a non-defective one is  $p$  and with probability  $q$  it produces a defective one.

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We want how many defective items are produced before getting 100 non-defective items.

$\uparrow$   
 $K = 100$

$\therefore X = 0 \Rightarrow$  all first 100 items are non-defective  
 $\therefore$  Its probability is  $p^{100}$

$$\binom{0+K-1}{K-1} p^{100} q^0 = p^{100}$$

$\rightarrow P(X=0) = \binom{0+K-1}{K-1} p^{100} q^0$

We want to know how many defective items are produced before getting 100 non-defective items. So, value of  $K$  is equal to 100. Therefore,  $X$  is equal to 0 implies all first 100 items are non-defective and its probability is  $p$  to the power 100. So, by our definition, 0 plus  $K$  minus 1, see  $K$  minus 1,  $p$  to the power 100  $q$  to the power 0 is equal to  $p$  to the power 100. Therefore, probability  $X$  is equal to 0 is equal to 0 plus  $K$  minus 1 see  $K$  minus 1  $p$  to the power 100,  $q$  to the power 0.

(Refer Slide Time: 40:58)

In a similar way we can check for all  $k$   
 $k = 0, 1, 2, \dots$

The logic is that:

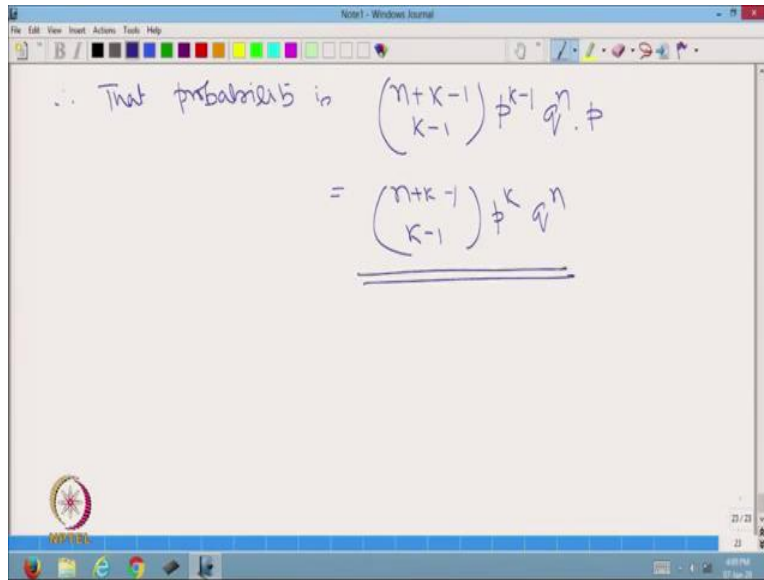
So if there are  $n$  failures then together we need  $n+k$  many trials to get  $k$  successes.

$\therefore$  Obviously the  $(n+k)^{\text{th}}$  trial is a success.

$\therefore$  Out of  $n+k-1$  many earlier trials there are  $k-1$  many success.

In a similar way we can check for all K. K is equal to 0, 1, 2 up to infinity. The logic is that so if there are n failures then together we need n plus K many tosses to get K success which you can say as H. Therefore, obviously the n plus Kth trial is a success. Therefore, out of n plus K minus 1 many earlier trials there are K minus 1 many success.

(Refer Slide Time: 40:59)



The screenshot shows a Windows Journal window with the following handwritten text:

$$\therefore \text{That probability is } \binom{n+K-1}{K-1} p^{K-1} q^n \cdot p$$

$$= \underline{\underline{\binom{n+K-1}{K-1} p^K q^n}}$$

Therefore, that probability is out of the total n plus K minus 1, you are choosing K minus 1, all of them have been a success that has the probability p to the power K minus 1. The remaining ones are failures that has probability q to the power n finally, the last one is a success. Therefore, this is coming out to be n plus K minus 1 see K minus 1 p to the power K q to the power n.

(Refer Slide Time: 41:53)

How do we interpret?

$X$  takes the value  $n$  with probability

$$\binom{n+k-1}{k-1} p^k q^n$$

Note that  $k$  &  $p$  are parameters of the dist<sup>n</sup>.

Suppose we have a coin with Prob of getting H is  $p$ .

So, that is how we get these probabilities as we have indicated here.

(Refer Slide Time: 41:58)

Question: Is it a valid pmf?

We know that:

$$\binom{n+k-1}{k-1} p^k q^n = \frac{(n+k-1)(n+k-2) \dots k(k-1) \dots 1}{(k-1)! n!} p^k q^n$$

$$= \frac{(n+k-1)(n+k-2) \dots k}{n!} p^k q^n$$

$$= \frac{(-1)^n (-k)(-k-1) \dots (-k-(n-1))}{n!} p^k q^n$$

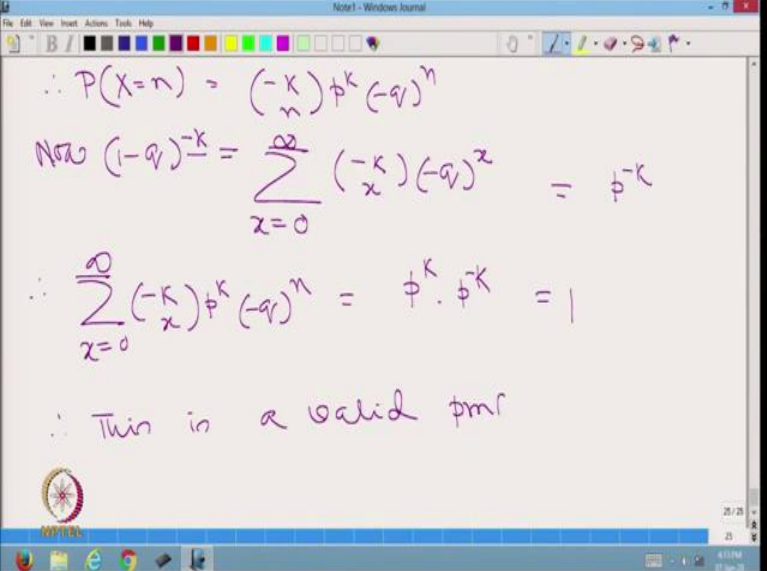
$$= \frac{(-k)_n}{n!} p^k (-q)^n$$

Question is, is it a valid pmf? We know that  $n$  plus  $k$  minus 1 see  $k$  minus 1  $p$  to the power  $k$   $q$  to the power  $n$  is equal to  $n$  plus  $k$  minus 1 into  $n$  plus  $k$  minus 2 up to  $k$ ,  $k$  minus 1 up to 1 divided by  $k$  minus 1 factorial,  $n$  factorial into  $p$  to the power  $k$   $q$  to the

power  $n$ , which is equal to if we cancel this part we get  $n$  plus  $k$  minus 1 into  $n$  plus  $k$  minus 2 up to  $k$  divided by  $n$  factorial  $p$  to the power  $k$   $q$  to the power  $n$ .

Is equal to minus 1 to the power  $n$ , note that there are  $n$  many terms here,  $k$  plus 0,  $k$  plus 1 up to  $k$  plus  $n$  minus 1. So, we take minus 1 to the power  $n$  common and then we get minus  $k$ , minus  $k$  minus 1 up to minus  $k$  minus  $n$  minus 1 divided by  $n$  factorial into  $p$  to the power  $k$   $q$  to the power  $n$  and this we can write it as minus  $k$  see  $n$   $p$  to the power  $k$  minus  $q$  to the power  $n$ . You have to understand binomial expansion with negative coefficients, which I am sure you have done in your class 11 or 12 algebra. So, this is the expansion that we are getting for the  $k$ th term.

(Refer Slide Time: 44:41)



The image shows a digital whiteboard with the following handwritten content:

$$\therefore P(X=n) = \binom{-k}{n} p^k (-q)^n$$

$$\text{Now } (1-q)^{-k} = \sum_{x=0}^{\infty} \binom{-k}{x} (-q)^x = p^{-k}$$

$$\therefore \sum_{x=0}^{\infty} \binom{-k}{x} p^k (-q)^n = p^k \cdot p^{-k} = 1$$

$\therefore$  This is a valid pmf

Therefore, probability  $X$  is equal to  $n$ , we can write it as minus  $k$  see  $n$   $p$  to the power  $k$  minus  $q$  to the power  $n$ , now  $1$  minus  $q$  whole to the power minus  $k$  is equal to sigma  $x$  is equal to  $0$  to infinity, note that it is a negative coefficient. Therefore, there will be infinitely many terms minus  $k$  see  $x$  minus  $q$  whole to the power  $x$ , which is equal to  $p$  to the power minus  $k$ . Therefore, sigma  $x$  is equal to  $0$  to infinity minus  $k$  see  $x$   $p$  to the power  $k$  minus  $q$  to the power  $x$  is equal to  $p$  to the power  $k$  into  $p$  to the power minus  $k$  is equal to  $1$  therefore, this is a valid pmf.

(Refer Slide Time: 46:10)

Hyper Geometric Dist<sup>n</sup>

$X \sim \text{HyperGeometric}(M, N)$

Suppose a box contains  $N$  balls of which  $M$  are white & remaining  $N-M$  are Black. We draw  $n$  balls out of the box without replacement.

We want to find the probability  $x$  of them are white.

Before we stop, I give you another discrete random distribution, Hyper Geometric Distribution. Here, it has two parameters  $m$  and  $n$ . So, suppose a box contains  $n$  balls of which  $m$  are white and remaining  $n$  minus  $m$  are black. We draw  $n$  balls out of the box without replacement, we want to find the probability  $x$  of them are white.

(Refer Slide Time: 48:07)

Now 
$$P(X=x) = \frac{\binom{M}{x} \binom{N-M}{n-x}}{\binom{N}{n}}$$

white  
 $x: 0, 1, 2, \dots, \min(n, M)$

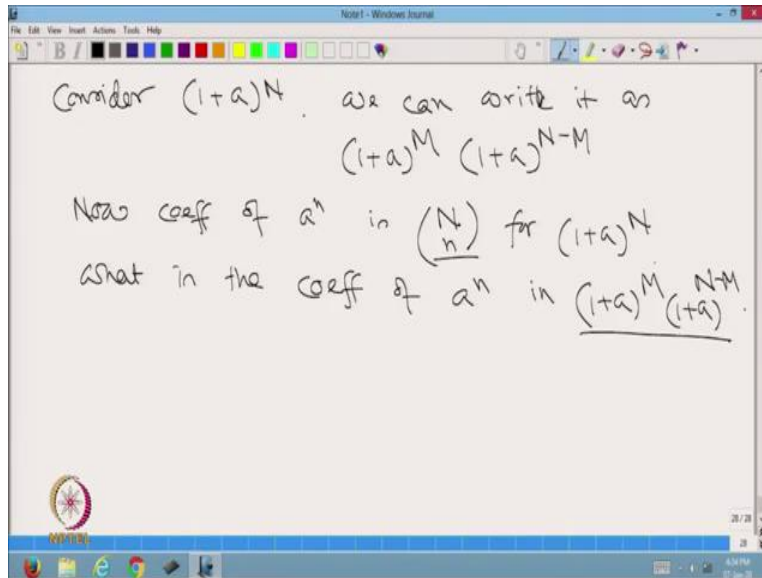
and 0 otherwise.

Obviously  $P_x \geq 0 \quad \forall x$

To check that it is a valid pmf we need to see that the sum is 1.

Now, probability  $X$  is equal to  $x$  is there are  $m$  white balls out of them  $x$  can be chosen in  $M$  see  $x$  ways out of the remaining  $N$  minus  $M$  black balls. Therefore, I choose  $n$  minus  $x$  are blacks, and total number of selection is  $N$  see  $n$  where  $x$  can take the value  $0, 1, 2$  up to a minimum of  $n$  and  $M$  and  $0$  otherwise. Obviously, these terms  $P_x$  is greater than equal to  $0$  for all  $x$  to show it is a valid pmf we need to see that the sum is  $1$ .

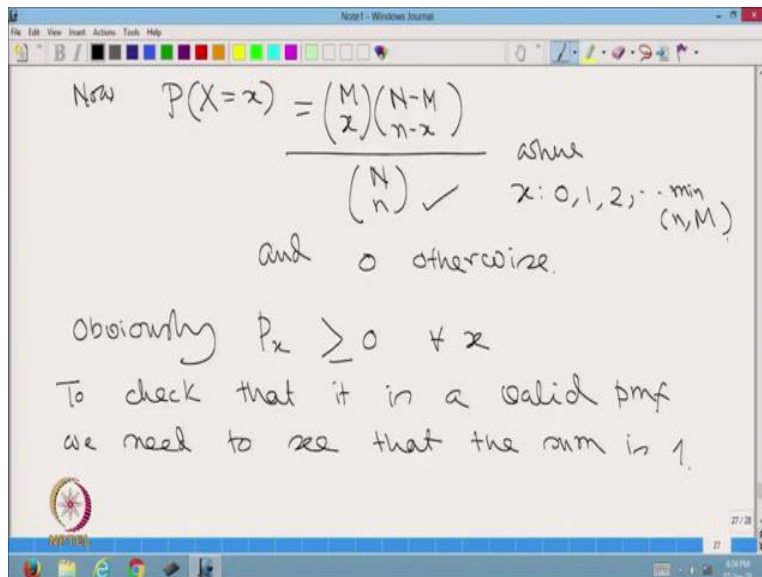
(Refer Slide Time: 49:42)



Consider  $(1+a)^N$ . we can write it as  $(1+a)^M (1+a)^{N-M}$

Now coeff of  $a^n$  is  $\binom{N}{n}$  for  $(1+a)^N$

what in the coeff of  $a^n$  in  $(1+a)^M (1+a)^{N-M}$



Now  $P(X=x) = \frac{\binom{M}{x} \binom{N-M}{n-x}}{\binom{N}{n}}$  where  $x: 0, 1, 2, \dots, \min(n, M)$

and  $0$  otherwise.

Obviously  $P_x \geq 0 \forall x$

To check that it is a valid pmf we need to see that the sum is  $1$ .

I give you a hint, considered  $1$  plus  $a$  whole to the power  $n$ , we can write it as  $1$  plus  $a$  whole to the power  $m$  into  $1$  plus  $a$  whole to the power  $N$  minus  $M$ . Now, coefficient of  $a$

to the power  $n$  is  $N \text{ see } n$  for  $1$  plus a whole to the power  $N$ . What is the coefficient of  $a$  to the power  $n$  in  $1$  plus a whole to the power  $M$  into  $1$  plus a whole to the power  $N$  minus  $M$ ? I want you to calculate that and see that this actually is giving the sum of the probabilities that we get from the Hyper Geometric distribution.

Therefore, that divided by  $N \text{ see } n$  is going to give us  $1$  and  $N \text{ see } n$  is the denominator here. So, effectively, I want you to check from here that the coefficient of  $a$  to the power  $n$  is actually the sum of these terms for  $x$  is equal to  $0$  to minimum over  $n$   $M$ . Therefore, that will give us a pmf.

Okay friends, I stopped here today. So, in this last two lectures, we have covered different discrete distributions, namely, Bernoulli, Binomial, Poisson, Geometric, Negative binomial and Hyper Geometric. In the next class, I shall introduce you to the concept of Continuous Random Variables. Thank you.