

Advanced Probability Theory
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Lecture 28

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WLLN

Theorem: If $X_1, X_2, \dots, X_n, \dots$ is a sequence of r.v.s
 $\exists E(X_i) = \mu_i < \infty$ & $Var(X_i) = \sigma_i^2$
then the sequence $\frac{S_n}{n} \xrightarrow{P} \frac{\sum \mu_i}{n}$
if $Var(\sum X_i)$ is $\exists \frac{Var(\sum X_i)}{n^2} \rightarrow 0$
as $n \rightarrow \infty$
Here $S_n = X_1 + X_2 + \dots + X_n$

Welcome students to the mock series of lectures on Advanced Probability Theory, this is lecture number 28, if you remember in the last class we were discussing weak law of large numbers which states that, if X_1, X_2, \dots, X_n is a sequence of random variables such that expectation of X_i is equal to μ_i which is finite and variance of X_i is equal to σ_i^2 then the sequence S_n upon n convergence in probability to $\frac{\sum \mu_i}{n}$ if variance of $\sum X_i$ is such that variance of $\sum X_i$ upon n^2 goes to 0 as n goes to infinity, here S_n is equal to X_1 plus X_2 plus here S_n is equal to X_1 plus X_2 plus X_n .

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In essence it says that the average of n r.v.s converges in Probability to the average of their respective Means or Expectations.
under certain condition.

If $\{X_n\}$ is a sequence iid, i.e. independent & identically distributed r.v.s then the condition required for WLLN to hold is $E(|X_i|) < \infty$

In that case $\frac{\sum_{i=1}^n X_i}{n} \xrightarrow{P} \mu$ which is the common Expectation.

So, in essence it says that the average of the n random variables converges in probability to the average of their respective means or expectations under certain condition. Now, if X_n is a sequence of iid's that is independent and identically distributed random variable then the condition required for weak law of large numbers to hold is expected value of modulus of X_i is finite. And in that case $\frac{\sum_{i=1}^n X_i}{n}$ converges in probability to μ which is the common expectation, in essence what we are saying that this one condition there that expected value of mod X_i is finite that takes care of the rather more stringent condition that we need for a general scenario when $X_1 X_2 X_n$ are any arbitrary sequence of random variables.

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Ex If X_1, X_2, \dots, X_n are iid $U(0,1)$, then
 $G_n = \text{Geometric Mean of } X_1, \dots, X_n \xrightarrow{P} c$, where
 c is constant.
• Show that the above statement is True.
• Also, determine the value of c .

So, before we move further, let us consider one example, if X_1, X_2, X_n are iid uniform 0, 1, then G_n which is the geometric mean of X_1, X_2, X_n converges in probability to a constant C , show that the above statement is correct, true and also determine the value of C .

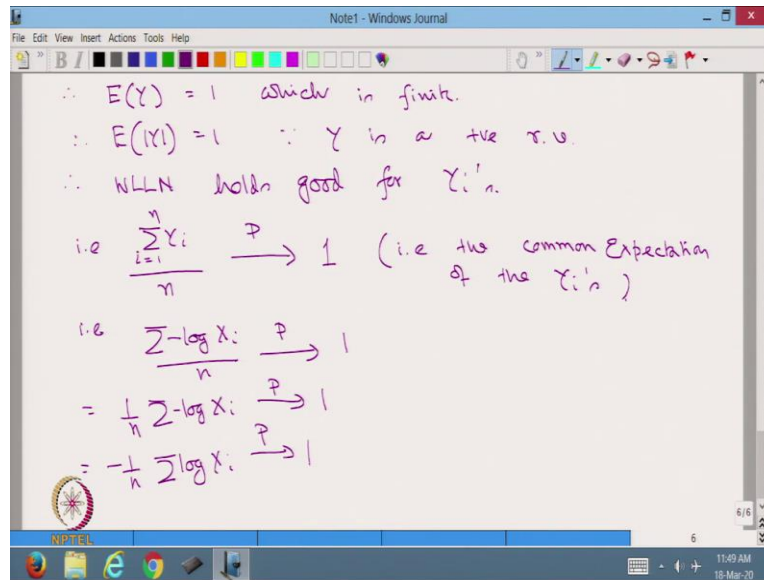
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Ans Let Y_i be $-\log X_i$
 $\therefore P(Y_i \leq y) = P(-\log X \leq y) = P(\log X \geq -y)$
 $= P(X \geq e^{-y}) = 1 - P(X \leq e^{-y})$
 $= \frac{1 - e^{-y}}{1}$
 \therefore Together we get that
 $Y \sim \text{Exp}(1)$
 $\therefore Y_1, Y_2, \dots, Y_n \dots$
are iid $\text{Exp}(1)$
This is because
 $X_i \sim U(0,1)$
 $\therefore \log X \in (-\infty, 0)$
ie $-\log X \in (0, \infty)$
 $\rightarrow Y \in (0, \infty)$

Answer, let Y_i be minus log of X_i , therefore probability Y_i less than equal to y is equal to probability minus of log X less than equal to y is equal to probability log X greater than minus y is equal to probability X greater than e to the power minus y is equal to 1 minus probability X less than equal to e to the power minus y is equal to 1 minus e to the power minus y . This is

because X is or X_i is uniform $0, 1$, therefore $\log X$ belongs to minus infinity to 0 that is minus $\log X$ belongs 0 to infinity implies Y belongs to 0 to infinity. Therefore, together we get that Y is distributed as exponential with parameter 1 , as this is the CDF of exponential 1 . Therefore, Y_1, Y_2, Y_n are iid exponential 1 .

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Handwritten mathematical derivations in a Windows Journal window:

$$\begin{aligned} \therefore E(Y) &= 1 \text{ which is finite.} \\ \therefore E(X_i) &= 1 \quad \therefore Y \text{ is a tve r.v.} \\ \therefore \text{WLLN holds good for } X_i\text{'s.} \\ \text{i.e. } \frac{\sum_{i=1}^n X_i}{n} &\xrightarrow{P} 1 \quad (\text{i.e. the common expectation of the } X_i\text{'s}) \\ \text{i.e. } \frac{\sum -\log X_i}{n} &\xrightarrow{P} 1 \\ &= \frac{1}{n} \sum -\log X_i \xrightarrow{P} 1 \\ &= -\frac{1}{n} \sum \log X_i \xrightarrow{P} 1 \end{aligned}$$

Therefore expected value of Y is equal to 1 , which is finite, therefore expected value of $\log Y$ is equal to 1 , since Y is a positive random variable, therefore weak law of large numbers holds good for Y_i 's that is $\sum_{i=1}^n Y_i$ is equal to 1 to n divided by n converges in probability to 1 , that is the common mean, common expectation of the Y_i 's, that is $\sum \log X_i$ upon n converges in probability to 1 , implies 1 by n $\sum \log X_i$ converges in probability to 1 , implies $-\frac{1}{n} \sum \log X_i$ converges in probability to 1 .

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Handwritten mathematical derivation in a Windows Journal window:

$$\Rightarrow -\frac{1}{n} \left(\log \left(\prod_{i=1}^n X_i \right) \right) \xrightarrow{P} 1$$

$$\Rightarrow -\left(\log \left(\left(\prod_{i=1}^n X_i \right)^{\frac{1}{n}} \right) \right) \xrightarrow{P} 1$$

$$\Rightarrow -\log G_n \xrightarrow{P} 1$$

$$\Rightarrow \log G_n \xrightarrow{P} -1$$

$$\Rightarrow G_n \xrightarrow{P} e^{-1}$$

Therefore:

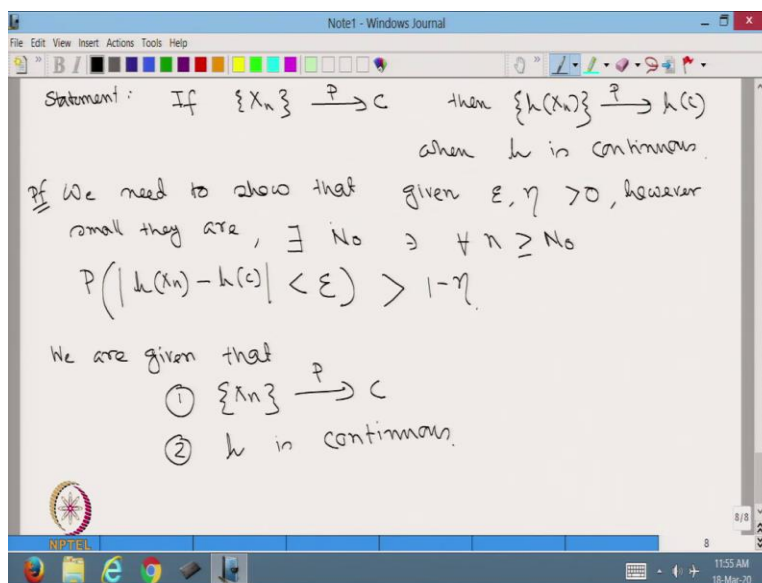
- ① $G_n \xrightarrow{P}$ to a constant C
- ② The value of C is $\frac{1}{e}$.

A boxed note on the right states:

If $X_n \xrightarrow{P} C$
 then $h(X_n) \xrightarrow{P} h(C)$
 if h is continuous

Implies minus 1 by n log of product of X_i i is equal to 1 to n converges in probability to 1, implies minus log of product of X_i i is equal to 1 to n to the power 1 by n converges in probability to 1, implies minus log of the geometric mean of X_1, X_2, X_n converges in probability to 1 implies log G_n converges in probability to minus 1 implies G_n converges in probability to e to the power minus 1, therefore one, G_n converges in probability to a constant C , two, the values of C is 1 upon e. Now, I have used a result at this point, the result is that if X_n converges in probability to a constant C then h of X_n converges in probability to h of C , if h is continuous. So, this is one result which I have used, so let me prove it.

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Statement: If $\{X_n\} \xrightarrow{P} C$ then $\{h(X_n)\} \xrightarrow{P} h(C)$ when h is continuous.

If we need to show that given $\epsilon, \eta > 0$, however small they are, $\exists N_0 \ni \forall n \geq N_0$

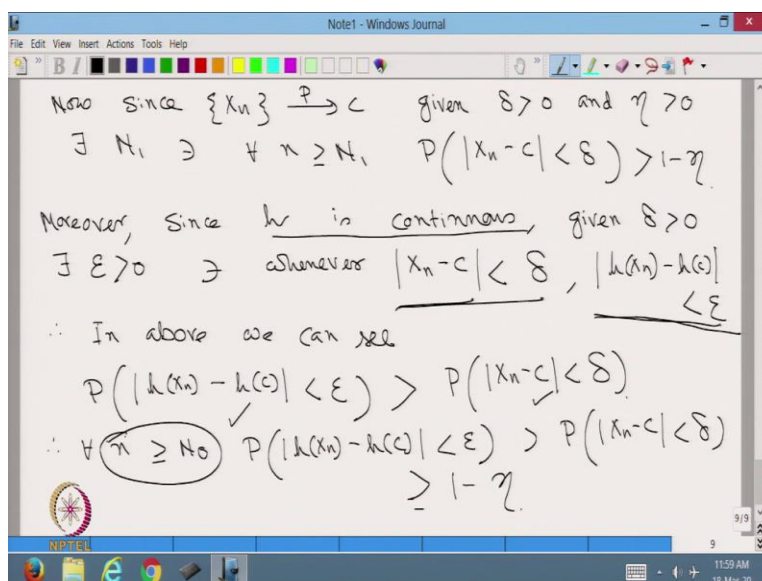
$$P(|h(X_n) - h(C)| < \epsilon) > 1 - \eta$$

We are given that

- ① $\{X_n\} \xrightarrow{P} C$
- ② h is continuous.

So, statement, if X_n converges in probability to C then the sequence h of X_n converges in probability to h of C when h is continuous. Proof, we need to show that given epsilon and eta greater than 0, however small they are, there exist N naught such that for all n greater than equal to N naught probability of modulus of $h X_n$ minus h of c less than epsilon is greater than 1 minus eta. We are given that one, X_n converges in probability to C and two, h is continuous

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Now since $\{X_n\} \xrightarrow{P} C$ given $\delta > 0$ and $\eta > 0$

$$\exists N_1 \ni \forall n \geq N_1, P(|X_n - C| < \delta) > 1 - \eta$$

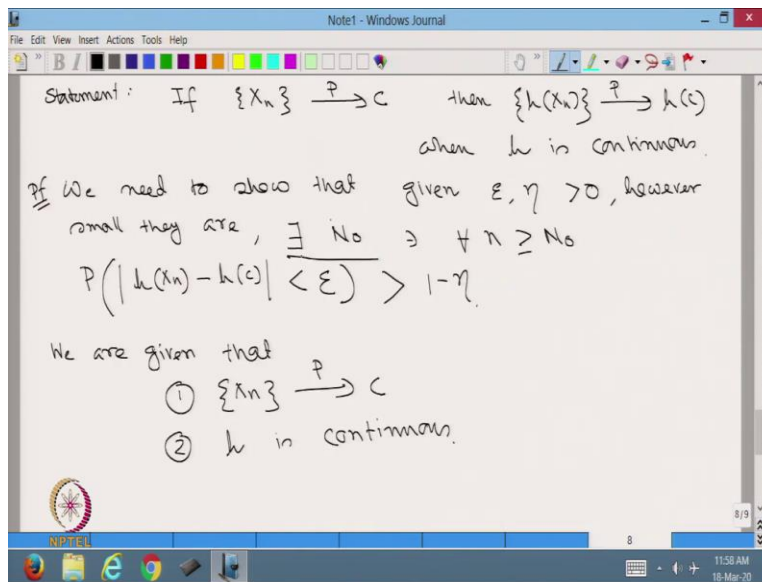
Moreover, since h is continuous given $\epsilon > 0$

$$\exists \delta > 0 \ni \text{whenever } |X_n - C| < \delta, |h(X_n) - h(C)| < \epsilon$$

\therefore In above we can see

$$P(|h(X_n) - h(C)| < \epsilon) > P(|X_n - C| < \delta)$$

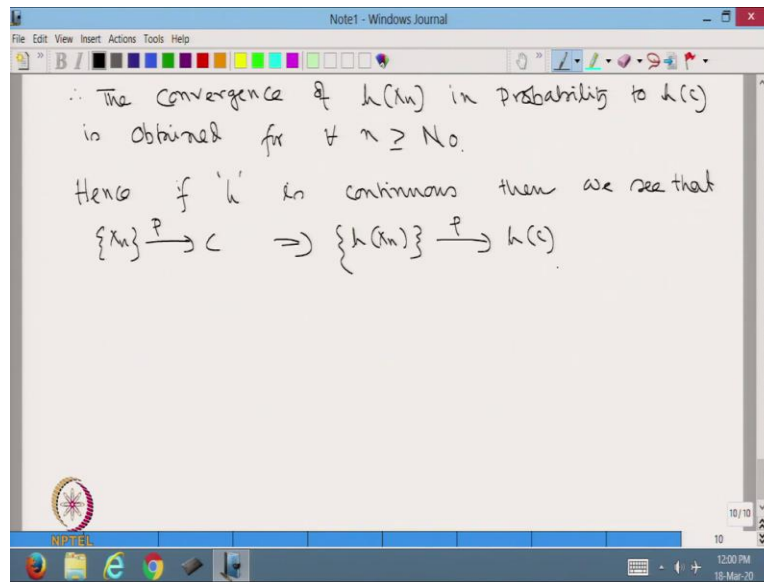
$\therefore \forall n \geq N_0, P(|h(X_n) - h(C)| < \epsilon) > P(|X_n - C| < \delta) \geq 1 - \eta$



Now, since X_n converges in probability to C , given δ greater than 0 and η greater than 0 there exist N_1 such that for all n greater than equal to N_1 probability modulus of X_n minus C less than δ is greater than $1 - \eta$. Moreover, since h is continuous, given δ greater than 0 there exist ϵ greater than 0 such that whenever modulus of X_n minus C is less than δ modulus of h of X_n minus h of C is less than ϵ . So, because h is continuous we get a neighbourhood around C in which we get this outcome, we get this result.

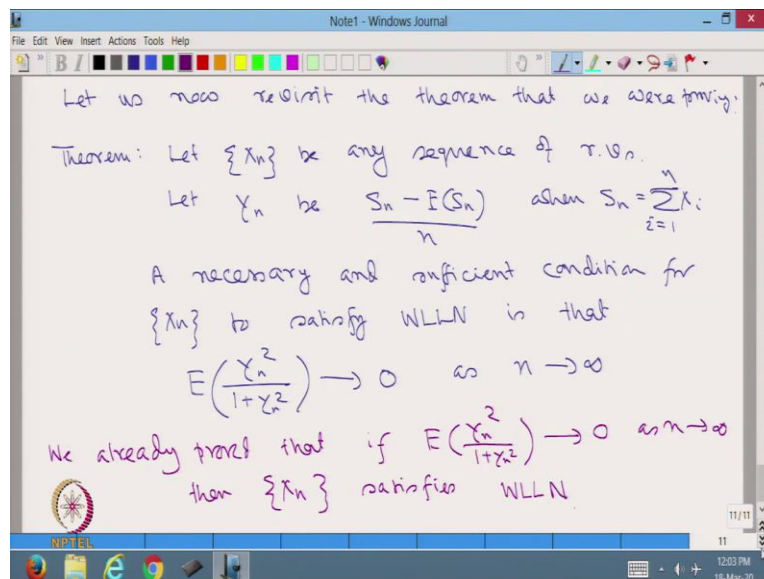
Therefore, in above we can see probability modulus of h of X_n minus h of C less than ϵ is greater than probability modulus of X_n minus C less than δ , because this event implies this event therefore this has a higher probability, then this event therefore for all n greater than equal to N naught let us go back this is the N naught now I am talking about probability modulus of h of X_n minus h of C less than ϵ is greater than probability modulus of X_n minus C less than δ which is greater than equal to $1 - \eta$, thus we get one N naught such that this property.

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Therefore, the convergence of h of X_n in probability to h of C is obtained for all n greater than equal to N_0 . Hence, if h is continuous then we see that X_n converging in probability to C implies h of X_n converges in probability into h of C

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So, let us now revisit the theorem that we were proving, the theorem was that let X_n be any sequence of random variables, let Y_n be S_n minus expected value of S_n divided by n when S_n is equal to $\sum_{i=1}^n X_i$, a necessary and sufficient condition for X_n to satisfy weak law of large numbers is that expected value of Y_n square upon $1 + Y_n$ square goes to 0 as n

goes to infinity. Note that we already proved in the last class, if expected value of Y_n square upon 1 plus Y_n square goes to 0 as n goes infinity then the sequence of random variables X_n satisfies weak law of large number.

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The image shows a digital whiteboard with handwritten text in purple ink. The text reads: "We now prove the converse, i.e. Suppose $\{X_n\}$ satisfies WLLN, we need to show that $E\left(\frac{X_n^2}{1+X_n^2}\right) \rightarrow 0$ as $n \rightarrow \infty$." Below this, the proof is given as: "Pf $E\left(\frac{X_n^2}{1+X_n^2}\right) = \int_{-\infty}^{\infty} \frac{y^2}{1+y^2} f_n(y) dy$ " with a note underneath: "where $f_n(y)$ is the pdf of X_n ." The whiteboard interface includes a menu bar (File, Edit, View, Insert, Actions, Tools, Help), a toolbar with various drawing tools, and a taskbar at the bottom with icons for NPTL, a folder, and a browser, along with a system clock showing 12:06 PM on 18-Mar-20.

We now prove the converse, suppose X_n satisfies weak law of large numbers, we need to show that expected value of Y_n square upon 1 plus Y_n square goes to 0 as n goes to infinity. Proof expected value of Y_n square upon 1 plus Y_n square is equal to integration minus infinity to infinity y square upon 1 plus y square $f_n(y) dy$, where $f_n(y)$ is the pdf of Y_n , this we know because we are computing the expectation of a function of a random variable and we have done this many times before.

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$$E\left(\frac{Y_n^2}{1+Y_n^2}\right) = \int_A \frac{y^2}{1+y^2} f_n(y) dy + \int_{A^c} \frac{y^2}{1+y^2} f_n(y) dy$$

where $A = \{\omega \mid |Y_n(\omega)| > \epsilon\}$
 $\therefore A^c = \{\omega \mid |Y_n(\omega)| \leq \epsilon\}$

Note that $\frac{y^2}{1+y^2} \leq 1 \quad \because y^2 \geq 0$
 $\therefore y^2 \leq (1+y^2) \quad \forall y$

Therefore, expected value of Y_n square upon 1 plus Y_n square is equal to integration over A y square upon 1 plus y square f of f_n of y dy plus integration of y square upon 1 plus y square f_n y dy on A complement, where A is the set of ω such that modulus of Y_n ω is greater than epsilon and therefore A complement is the set of ω such that modulus of Y_n ω is less than equal to epsilon. Now, note that y square upon 1 plus y square is less than equal to 1, since y square is greater than equal to 0. Therefore, y square is less than equal to 1 plus y square for all y . Now, we are going to plug in this expression.

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$$E\left(\frac{Y_n^2}{1+Y_n^2}\right) = \int_A \frac{y^2}{1+y^2} f_n(y) dy + \int_{A^c} \frac{y^2}{1+y^2} f_n(y) dy$$

$$\leq \left(\int_A 1 \cdot f_n(y) dy \right) + \int_{A^c} \frac{y^2}{1+y^2} f_n(y) dy$$

$$\leq P(A) + \int_{A^c} \epsilon^2 f_n(y) dy$$

$$= P(A) + \epsilon^2 P(A^c)$$

$$= P(A) + \epsilon^2 \quad (\because P(A^c) \leq 1)$$

This because $A: |Y_n| > \epsilon \therefore \text{on } A^c |Y_n| \leq \epsilon \therefore Y_n^2 \leq \epsilon^2$

Therefore, let us rewrite expectation of Y_n square upon $1 + Y_n$ square is equal to integration over A y square upon $1 + y$ square $f_n(y) dy$ plus integration over A^c y square upon $1 + y$ square $f_n(y) dy$, since this is less than 1 we can write it as this is less than equal to A into $1 f_n(y) dy$ plus integration over A^c y square into $f_n(y) dy$, this is because y square is less than $1 + y$ square therefore y square upon $1 + y$ square is less than y square, less than equal to probability of A because if we take 1 out this gives the probability of A plus integration over A^c y square into $f_n(y) dy$.

This is because A is equal to such that modulus of Y_n is greater than ϵ , therefore on A^c modulus of Y_n is less than equal to ϵ , therefore Y_n square is less than equal to ϵ square is equal to probability of A plus ϵ square into probability of A^c because if we take ϵ square out we get probability of A^c is equal to probability of A plus ϵ square since probability of A^c is less than equal to 1.

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The image shows a digital notepad with the following handwritten content:

$$\frac{Y_n^2}{1+Y_n^2} \leq P(A) + \epsilon^2$$

$$= P(|Y_n| > \epsilon) + \epsilon^2 \quad \checkmark$$

Now $Y_n = \frac{S_n - E(S_n)}{n} = \frac{\sum_{i=1}^n X_i - E(\sum_{i=1}^n X_i)}{n}$

And we are given that $\{X_i\}$ satisfies WLLN.

$$\therefore P\left(\left|\frac{\sum_{i=1}^n X_i - \sum_{i=1}^n \mu_i}{n}\right| > \epsilon\right) \xrightarrow{P} 0 \quad \text{as } n \rightarrow \infty$$

where $\mu_i = E(X_i)$

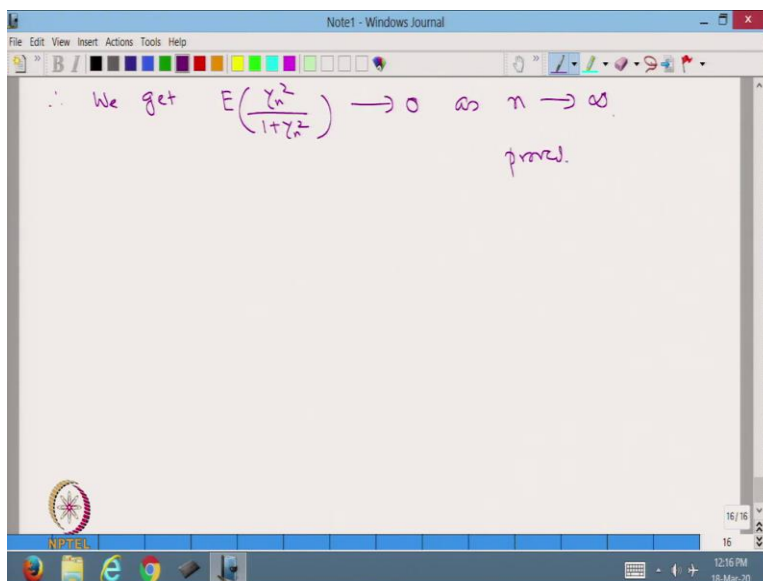
\therefore We get that the value of $\frac{Y_n^2}{1+Y_n^2}$ can be made arbitrarily small as $n \rightarrow \infty$.

$\therefore \epsilon$ can be taken arbitrarily small positive quantity.

Therefore, Y_n square upon $1 + Y_n$ square is less than equal to probability of A plus ϵ square is equal to probability modulus of Y_n greater than ϵ plus ϵ square. Now, Y_n is equal to S_n minus expected value of S_n divided by n is equal to $\sum_{i=1}^n X_i$ minus expected value of $\sum_{i=1}^n X_i$ divided by n and we are given that X_i satisfies weak law of large numbers. Therefore, probability modulus of $\sum_{i=1}^n X_i$ minus $\sum_{i=1}^n \mu_i$ divided by n greater than ϵ converges in probability to 0 as n goes to

infinity, where μ_i is equal to the expected value X_i . Therefore, from this result and this result we get that the value of Y_n^2 upon $1 + Y_n^2$ can be made arbitrarily small as n goes to infinity. Since ϵ can be taken arbitrarily small positive quantity.

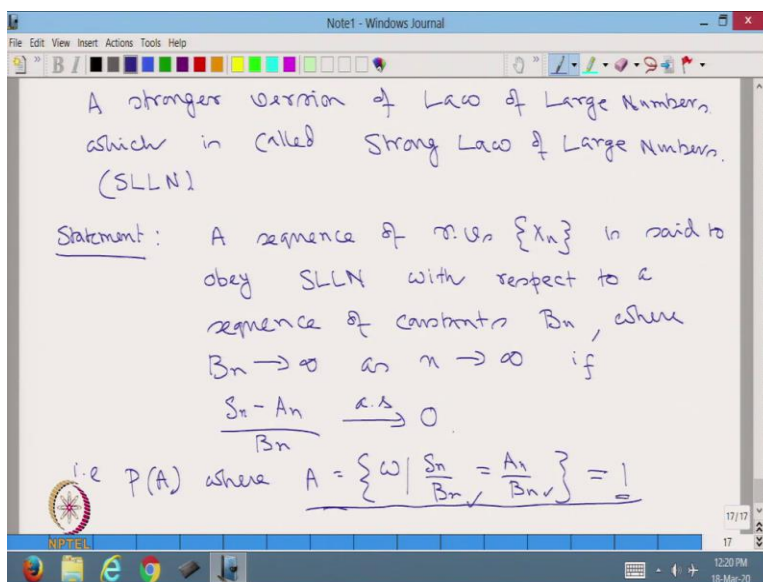
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A screenshot of a Windows Journal window titled "Note1 - Windows Journal". The window contains handwritten text in purple ink. The text reads: "∴ We get $E\left(\frac{Y_n^2}{1+Y_n^2}\right) \rightarrow 0$ as $n \rightarrow \infty$ proved." The window has a standard toolbar with various drawing tools and a taskbar at the bottom showing the time as 12:16 PM on 16-Mar-20.

Therefore, we get expected value of Y_n^2 upon $1 + Y_n^2$ goes to 0 as n goes to infinity, proved.

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A screenshot of a Windows Journal window titled "Note1 - Windows Journal". The window contains handwritten text in purple ink. The text reads: "A stronger version of Law of Large Numbers which is called Strong Law of Large Numbers (SLLN).
Statement: A sequence of r.v.s $\{X_n\}$ is said to obey SLLN with respect to a sequence of constants B_n , where $B_n \rightarrow \infty$ as $n \rightarrow \infty$ if $\frac{S_n - A_n}{B_n} \xrightarrow{a.s.} 0$.
i.e $P(A)$ where $A = \left\{ \omega \mid \frac{S_n}{B_n} = \frac{A_n}{B_n} \right\} = 1$ " The window has a standard toolbar with various drawing tools and a taskbar at the bottom showing the time as 12:20 PM on 16-Mar-20.

Now, let us consider a stronger version of law of large numbers which is called strong law of large numbers or SLLN. So, statement a sequence of random variables X_n is said to obey strong

law of large numbers with respect to a sequence of constants B_n , where B_n goes to infinity as n goes to infinity if S_n minus A_n upon B_n converges almost surely to 0, that is probability of A where A is equal to the set of ω on which S_n upon B_n is equal to A_n upon B_n is equal to 1. That means the measure of the set on which S_n upon B_n equals A_n upon B_n that has the probability 1.

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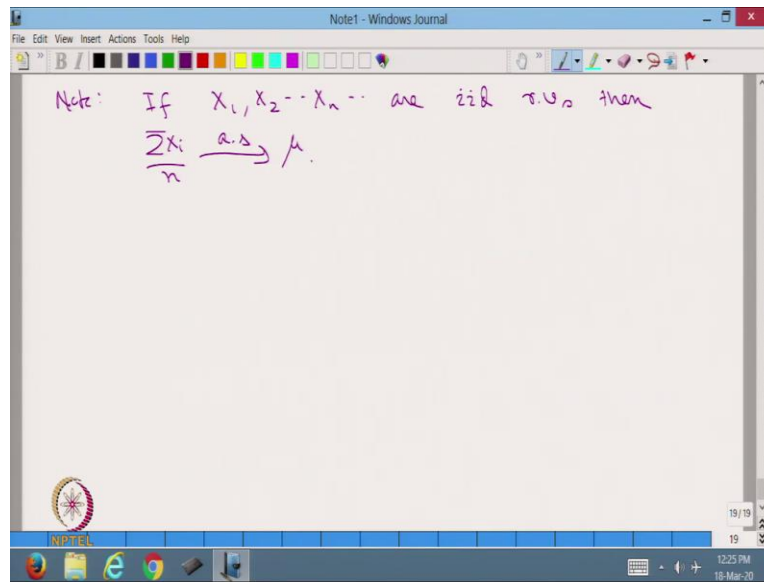
Note: The A_n 's are called centering constants.
The B_n 's are called Norming constants.

In general form of WLLN we can use $\{A_n\}$ & $\{B_n\} \Rightarrow B_n \rightarrow \infty$ as $n \rightarrow \infty$
 ✓ and we say that $\{X_n\}$ obeys WLLN wrt $\{B_n\}$ of constants if there exists $\{A_n\}$ of real nos $\Rightarrow \frac{S_n - A_n}{B_n} \xrightarrow{P} 0$ as $n \rightarrow \infty$.

In general we consider $A_n = \sum \mu_i$ & $B_n = n$.

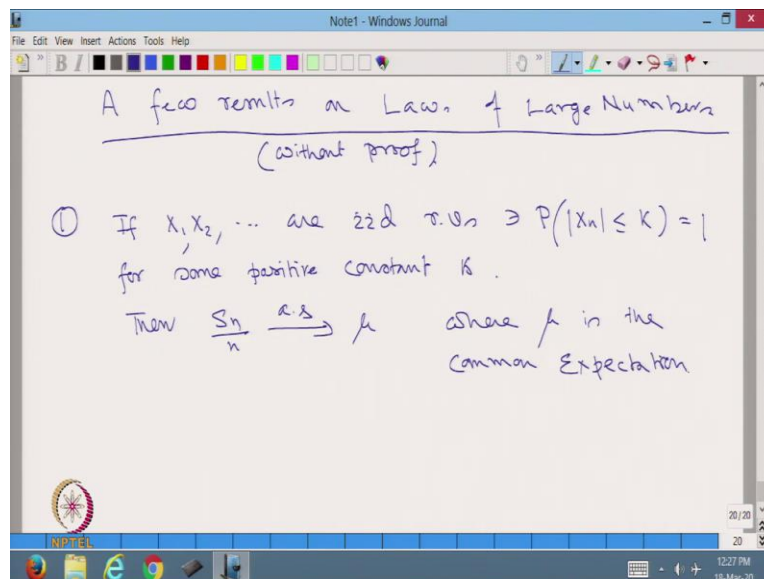
Note, the A_n 's are called centering constants and the B_n 's are called Norming constant. In a general form of weak law of large numbers, we can use this sequence of numbers A_n 's and sequence of numbers B_n 's such that B_n is going to infinity as n goes to infinity and we say that X_n obeys weak law of large numbers with respect to the sequence B_n of constants if there exists a sequence A_n of real numbers such that S_n minus A_n upon B_n converges in probability to 0 as n goes to infinity. Note that this is a general form of weak law of large numbers, in general we consider A_n is equal to sigma over mu i and B_n is equal to n . And put this values here you will get that we have defined weak law of large numbers in this general form with this particular value of A_n and B_n .

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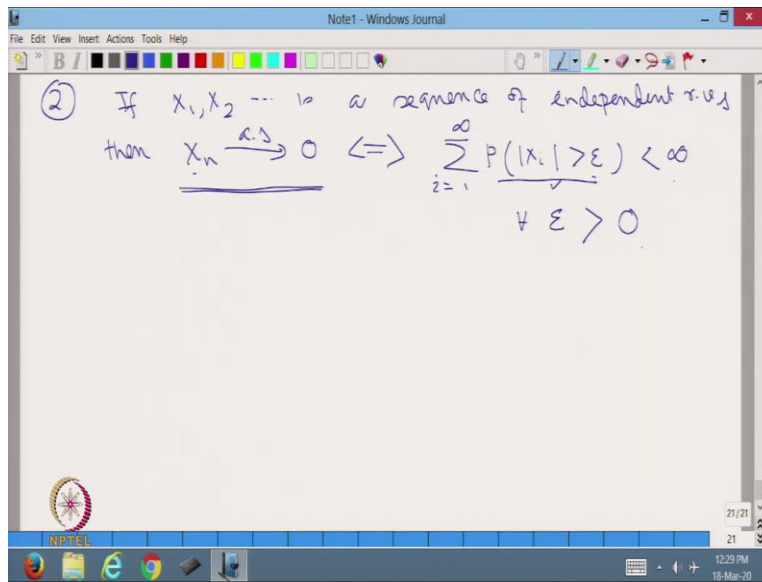
Note, if X_1, X_2, X_n are iid random variables then $\sum X_i$ upon n converges almost surely to μ . So, these results can be proved with some more knowledge of mathematics and measure theory, but that is not within the gamut of this course, so I am not proving this result.

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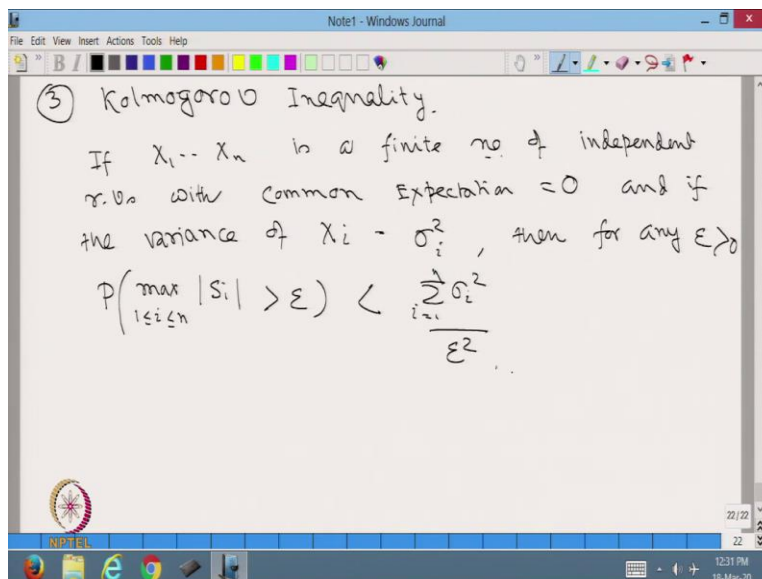
So, let us take this results and I stop this lecture by stating a few more important results on law of large numbers. One, if X_1, X_2 are iid random variables, such that probability modulus of X_n is less than equal to k is equal to 1 for some positive constant k , then S_n upon n converges almost surely to μ , where μ is the common expectation.

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Two, if X_1, X_2 , etcetera is a sequence of independent random variable, then X_n converges almost surely to 0 implies and implied by $\sum_{i=1}^{\infty} P(|X_i| > \varepsilon) < \infty$ for all $\varepsilon > 0$, it is very easy to visualize as the probability that X_n is taking the value of 0 is going to 1, therefore whatever positive ε we take the modulus, the absolute value of X_i to be greater than that has to be finite, because if that is infinite then X_n cannot converge to 0 with probability 1.

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Three, this is called Kolmogorov inequality, which states that if X_1, X_2, \dots, X_n is a finite number of independent random variables with common expectation is equal to 0 and if the variance of X_i is equal to σ_i^2 , then for any ϵ greater than 0 probability maximum over 1 less than equal to $\sum_{i=1}^n \sigma_i^2$ greater than ϵ^2 is less than $\sum_{i=1}^n \sigma_i^2$ is equal to 1 to n divided by ϵ^2 .

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④ If $X_1, X_2, \dots, X_n, \dots$ is a sequence of r.v.s
 $\Rightarrow \sum_{i=1}^{\infty} \text{Var}(X_i) < \infty$ then $\sum_{i=1}^{\infty} (X_i - E(X_i)) \xrightarrow{\text{a.s.}} 0$

Corollary If $\{X_n\}$ is a sequence of independent r.v.s & if $\sum_{k=1}^{\infty} \frac{\text{Var}(X_k)}{B_k^2} < \infty$
 and $B_k \rightarrow \infty$ as $k \rightarrow \infty$
 then $\frac{S_n - E(S_n)}{B_n} \xrightarrow{\text{a.s.}} 0$

Four, if X_1, X_2, \dots, X_n is a sequence of random variables such that $\sum_{i=1}^{\infty} \text{Var}(X_i)$ is finite, then $\sum_{i=1}^{\infty} (X_i - E(X_i))$ converges almost surely to 0. A corollary to the above is that, if X_n is a sequence of independent random variables and if $\sum_{k=1}^{\infty} \frac{\text{Var}(X_k)}{B_k^2}$ is finite and B_k goes to infinity as k goes to infinity, then $\frac{S_n - E(S_n)}{B_n}$ converges almost surely to 0.

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⑤ Borel's SLLN

For a sequence of Bernoulli trials with constant probability of success 'p' SLLN holds with $A_n = np$ & $B_n = n$

i.e. $\frac{\sum X_i}{n} \xrightarrow{a.s.} p$

Five, Borel's strong law of large numbers. For a sequence of Bernoulli trials with constant probability of success p , the strong law of large numbers holds with A_n is equal to np and B_n is equal to n that is $\sum X_i$ upon n converges almost surely to p .

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⑥ Poisson's WLLN

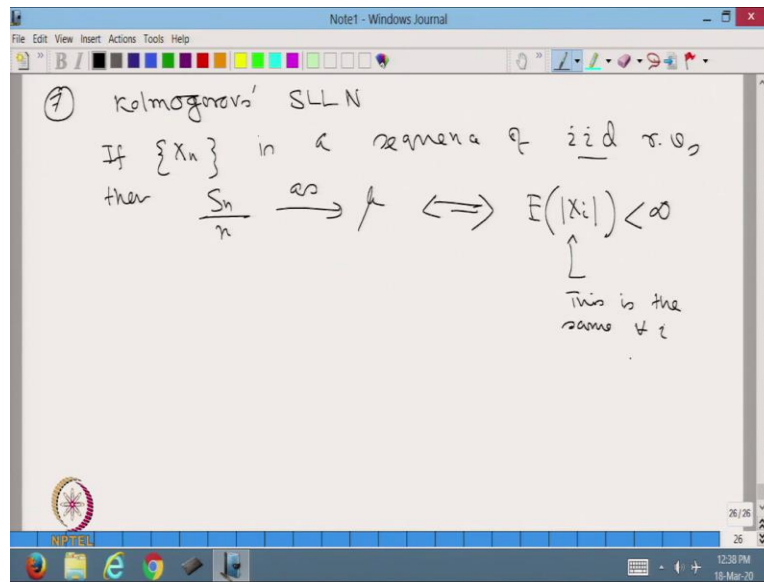
Also on Bernoulli trials but here 'p' is not a constant, rather it can change for each X_k :

$$X_k = \begin{cases} 1 & p_k \\ 0 & 1-p_k \end{cases}$$

then $\frac{\sum X_i}{n} \xrightarrow{p} \frac{\sum p_i}{n}$

Number six, Poisson's weak law of large numbers also on Bernoulli trials but here p is not a constant rather it can change for each X_k such that X_k takes the value 1 with probability p_k and 0 with $1 - p_k$ with probability p_k and 0 with $1 - p_k$, then $\sum X_i$ upon n converges in probability to $\sum p_i$ upon n .

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And finally Kolmogorov strong law of large numbers which say that if X_n is a sequence of iid random variables then S_n upon n converges almost surely to the common mean μ implies and implied by expected value of modulus of X_i is finite, note that this is the same for all i . Okay friends I stop here today, so I have given you a lot of results consingning convergence of sequence of random variables when the number of such variables is large and there are several theorems stating different results, they can be proved mathematically using measure theory and analysis. But as I said that in this course, you are not going into the details of those proof, but it is better for our applications to remember this results with that I conclude my talk on law of large numbers from the next class we shall start the most fundamental result of probability or one of the most fundamental results of probability mainly central limit theorems. Okay friends, thank you.