

1. (a) A 90% C.I. is  $\hat{\mu} \pm c \frac{\sigma}{\sqrt{n}}$ ,  $C \sim N(0,1)$

$$= 20.59 \pm 1.645 \frac{1.2}{\sqrt{15}}$$

$$= [20.08, 21.10]$$

(b) We want a 90% C.I. with width 0.1

$$1.645 \frac{1.2}{\sqrt{n}} \leq 0.1$$

$$\sqrt{n} \geq 1.645 \frac{1.2}{0.1}$$

$$\geq 19.74$$

$$n \geq 389.66$$

So  $n$  has to be at least 390

2. Let  $\bar{y}_A$  be the mean in diet A and  $\bar{y}_B$  be the mean in diet B. Let  $\bar{y}_d = \bar{y}_A - \bar{y}_B$

$$\bar{y}_d \pm c \frac{s_d}{\sqrt{n_d}}, C \sim t_8$$

$$= 14.889 - 12.444 \pm 2.306 \frac{8.308}{\sqrt{9}}$$

$$= [-3.94, 8.83]$$

Since we are 95% confident that the true value of the difference lies in this interval. We conclude that the two diets do not have much difference

3.  $H_0: \mu_b > \mu_c$   $H_a: \mu_b \leq \mu_c$

$$\hat{\mu}_b - \hat{\mu}_c = \frac{1+5}{2} = 2$$

$$c \frac{s_d}{\sqrt{n_d}} = 3, C \sim t_{16}$$

$$c = 2.12$$

$$\sqrt{\frac{s_b^2}{n_b} + \frac{s_c^2}{n_c}} = \frac{3}{2.12}$$

$$d = \frac{2}{\frac{3}{2.12}} = 1.413$$

$$P(D \geq d) = P(D \geq 1.413) \text{ where } D \sim t_{16}$$

$$= 1 - P(D < 1.413)$$

$$= 0.088$$

There is evidence reject  $H_0$ .

$$\begin{aligned} 4.(a) \text{ Let } W &= \sum_{ij} Y_{ij}^2 + \lambda (\tau_1 + \tau_2) \\ &= \sum_{ij} (y_{ij} - \mu - \tau_i)^2 + \lambda (\tau_1 + \tau_2) \\ &= \sum_j (y_{1j} - \mu - \tau_1)^2 + \sum_j (y_{2j} - \mu - \tau_2)^2 + \lambda (\tau_1 + \tau_2) \end{aligned}$$

$$\frac{\partial W}{\partial \mu} = -2 \sum_{ij} (y_{ij} - \mu - \tau_i)$$

$$\frac{\partial W}{\partial \tau_1} = -2 \sum_j (y_{1j} - \mu - \tau_1) + \lambda$$

$$\frac{\partial W}{\partial \tau_2} = -2 \sum_j (y_{2j} - \mu - \tau_2) + \lambda$$

$$\frac{\partial W}{\partial \lambda} = \tau_1 + \tau_2$$

$$\begin{aligned} (b) E(\tilde{\tau}_1) &= E(\bar{Y}_{1+} - \bar{Y}_{++}) \\ &= E\left(\sum_{j=1}^5 \frac{Y_{1j}}{5} - \sum_{ij} \frac{Y_{ij}}{10}\right) \\ &= \frac{1}{5} E\left(\sum_{j=1}^5 Y_{1j}\right) - \frac{1}{10} E\left(\sum_{ij} Y_{ij}\right) \\ &= \frac{1}{5} \sum_{j=1}^5 E(\mu + \tau_1 + R_{1j}) - \frac{1}{10} \sum_{ij} E(\mu + \tau_i + R_{ij}) \\ &= \frac{1}{5} \left(\sum_{j=1}^5 \mu + \tau_1\right) - \frac{1}{10} \left(\sum_{j=1}^5 \mu + \tau_1 + \sum_{j=1}^5 \mu + \tau_2\right) \quad \text{Since } E(R_{ij}) = 0 \quad \forall i, j \\ &= \frac{1}{5} (5\mu + 5\tau_1) - \frac{1}{10} 10\mu \quad \text{Since } \tau_1 + \tau_2 = 0 \\ &= \tau_1 \end{aligned}$$

Since  $E(\tilde{\tau}_1) = \tau_1$ ,  $\tilde{\tau}_1$  is unbiased.

$$(c) \text{Var}(\tilde{\tau}_1 - \tilde{\tau}_2)$$

$$= \text{Var}(\bar{Y}_{1+} - \bar{Y}_{++} - \bar{Y}_{2+} + \bar{Y}_{++})$$

$$= \text{Var}(\bar{Y}_{1+} - \bar{Y}_{2+})$$

$$= \text{Var}(\bar{Y}_{1+}) + \text{Var}(\bar{Y}_{2+}) \quad \text{Since independent}$$

$$= \frac{1}{25} \text{Var}\left(\sum_{j=1}^5 \mu + \tau_1 + R_{1j}\right) + \frac{1}{25} \text{Var}\left(\sum_{j=1}^5 \mu + \tau_2 + R_{2j}\right)$$

$$= \frac{1}{25} \text{Var}\left(\sum_{j=1}^5 R_{1j}\right) + \frac{1}{25} \text{Var}\left(\sum_{j=1}^5 R_{2j}\right) \quad R_{ij} \sim N(0, \sigma^2)$$

$$= \frac{1}{25} \sum_{j=1}^5 \text{Var}(R_{1j}) + \frac{1}{25} \sum_{j=1}^5 \text{Var}(R_{2j})$$

$$= \frac{g^2}{5} + \frac{g^2}{5}$$

$$\approx \frac{2}{5} g^2$$