

1. (a) Assume $Y_n = i$ where $i > 0$

$$Y_{n+1} = i+1 \text{ if } X_{n+1} = i+1 \text{ or } -i-1$$

$$Y_{n+1} = i-1 \text{ if } X_{n+1} = i-1 \text{ or } -i+1$$

$$P(Y_{n+1} = 1 | Y_n = 0) = 1$$

For $i > 0$

$$P(Y_{n+1} = i+1 | Y_n = i) = P(|X_n| = i+1 | |X_n| = i)$$

$$= \begin{cases} P & \text{if } X_n = i \\ 1-P & \text{if } X_n = -i \end{cases}$$

← only depends on X_n and Y_n

$$P(Y_{n+1} = i-1 | Y_n = i) = P(|X_n| = i-1 | |X_n| = i)$$

$$= \begin{cases} 1-P & \text{if } X_n = i \\ P & \text{if } X_n = -i \end{cases}$$

If $\{Y_n\}_{n=0,1,2,\dots}$ is a DTMC, $1-P=P \Rightarrow P=\frac{1}{2}$

~~Since $P_{i,i+1} = P_{i,i-1}$ and $P_{i,i} = 0$ for $i > 0$~~

$$(b) P((X_{n+1}, X_{n+2}) = (j_1, j_2) | (X_n, X_{n+1}) = (i_1, i_2), (X_{n-1}, X_n) = (i_0, i_1), \dots, (X_0, X_1) = (i_0, i_1))$$

$$= \begin{cases} 0 & \text{if } j_2 \neq j_1 \\ P_{j_2, j_1} (P_{i_2, j_2}) & \text{o.w.} \end{cases}$$

since the prob. only depends on (X_n, X_{n+1}) , $\{(X_n, X_{n+1})\}_{n=0,1,\dots}$ is also a DTMC

Let S' be the state space of $\{X_n\}_{n=0,1,\dots}$

S' be the state space of $\{(X_n, X_{n+1})\}_{n=0,1,\dots}$

$$\pi((i, j)) = \sum_{(m, n) \in S'} \pi((m, n)) P'_{(m, n)}(i, j)$$

$$= \sum_{k \in S} \pi((k, i)) P_{ij}$$

$$= P_{ij} \sum_{k \in S} \pi((k, i))$$

Let $\pi(i, j) = \pi(i) P_{ij}$

$$\pi(i, j) = \pi(i) P_{ij}$$

$$P_{ij} \sum_{k \in S} \pi((k, i)) = P_{ij} \sum_{k \in S} \pi(k) P_{ki} = P_{ij} \pi(i) \text{ by defn}$$

$$\sum_{(i, j) \in S'} \pi((i, j)) = \sum_{(i, j) \in S'} \pi(i) P_{ij} = \sum_{j \in S} \sum_{i \in S} \pi(i) P_{ij} = \sum_{j \in S} \pi(j) = 1$$

Hence $\pi(i, j) = \pi(i) P_{ij}$ is a stationary distribution of $\{(X_n, X_{n+1})\}_{n=0,1,\dots}$

2. (a) consider a DTMC with transition matrix $P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix}$

Both states have period 2

so $P^2 = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$

clearly, P is reducible

(b) since $\gcd\{n \geq 1: p_{xx}^{(n)}\} = d \forall x \in S$ (irreducible)

$\Rightarrow \gcd\{n \geq 1: p_{xx}^{(n)}\} = d$

$\gcd\{n \geq 1: (p_{xx}^{(n)})^{\frac{1}{d}}\} = 1$ by the property of gcd

By defn for all the states in $\{x_n\}_{n=0,1,\dots}$, the period is 1.

3. $S = \{1, 2, 3\}$

$$P = \begin{matrix} & \begin{matrix} 1 & 2 & 3 \end{matrix} \\ \begin{matrix} 1 \\ 2 \\ 3 \end{matrix} & \begin{bmatrix} 0.8 & 0.1 & 0.1 \\ 0.6 & 0.2 & 0.2 \\ 0.4 & 0.3 & 0.3 \end{bmatrix} \end{matrix}$$

Since $P_{ij} > 0 \forall i, j \in S$, the chain is irreducible

Since the chain is finite and irreducible, all states are recurrent

By theorem

$$\begin{cases} \pi(1) = 0.8\pi(1) + 0.6\pi(2) + 0.4\pi(3) \\ \pi(2) = 0.1\pi(1) + 0.2\pi(2) + 0.3\pi(3) \\ \pi(3) = 0.1\pi(1) + 0.2\pi(2) + 0.3\pi(3) \\ \pi(1) + \pi(2) + \pi(3) = 1 \end{cases} \Rightarrow \begin{cases} \pi(1) = \frac{5}{7} \\ \pi(2) = \frac{1}{7} \\ \pi(3) = \frac{1}{7} \end{cases}$$

Hence stationary distribution exists

By theorem

$\frac{N_n(y)}{n} \rightarrow \pi(y)$

so $\frac{N_n(1)}{n} \rightarrow \frac{5}{7}$

$\frac{N_n(2)}{n} \rightarrow \frac{1}{7}$

$\frac{N_n(3)}{n} \rightarrow \frac{1}{7}$

4. (a) Since this chain has a ^{transition} triangular matrix, it is detail balanced.

By theorem, this chain is time-reversible

(b) since it is time-reversible

$\{Y_m\}_{m=0,1,2}$ has the same distribution as $\{X_n\}_{n=0,1,2}$ where $Y_m = X_{2-m}$

$$P(X_1=2 | X_2=3) = P(Y_1=2 | Y_0=3)$$

$$= P_{32}$$

$$= 0.6$$

$$(c) P(X_2=3, X_4=1 | X_3=2)$$

$$= P(Y_2=3, Y_0=1 | Y_1=2)$$

$$= \frac{P(Y_2=3, Y_1=2, Y_0=1)}{P(Y_1=2)}$$

$$= \frac{P_{12} P_{23} \pi(1)}{\pi(2)}$$

$$= \frac{0.2 \cdot 0.3 \pi(1)}{\pi(2)}$$

$$= \frac{0.06 \pi(1)}{\pi(2)}$$

now:

By part B:

$$\frac{\pi(1)}{\pi(2)} = \frac{0.1 \pi(0) + 0.7 \pi(2)}{\pi(2)}$$

$$= \frac{0.1 \pi(0)}{\pi(2)} + 0.7$$

$$= \frac{0.8 \pi(1)}{\pi(2)} + 0.7$$

$$= \frac{0.8 \pi(1)}{\pi(2)} + 0.7$$

$$= 3.5$$

$$\text{Hence } P(X_2=3, X_4=1 | X_3=2) = 3.5 \times 0.06 = 0.21$$

5. (a) This chain is ~~not~~ irreducible

since ~~there exists~~ $i \rightarrow j \forall i, j \in S$

~~This chain can be divided into 2 classes~~ $d(25)=1$ since $P_{44} > 0$

~~$\{0,1,2,4\}$ $d(\{0,1,2,4\})=0$ since $P_{44} > 0$~~

~~$\{0,1,2,4\}$ $d(\{0,1,2,4\})=0$ since $P_{44} > 0$~~

$$(b) \begin{cases} \pi(0) = \frac{1}{2} \pi(3) + \frac{1}{2} \pi(4) \\ \pi(1) = \frac{1}{3} \pi(0) \\ \pi(2) = \frac{2}{3} \pi(0) \\ \pi(3) = \frac{1}{2} \pi(1) + \frac{1}{4} \pi(2) \\ \pi(4) = \frac{1}{2} \pi(1) + \frac{3}{4} \pi(2) + \frac{1}{2} \pi(4) \end{cases}$$

Then

$$\Rightarrow \pi = \begin{bmatrix} \frac{0}{\frac{3}{11}} & \frac{1}{\frac{1}{11}} & \frac{2}{\frac{2}{11}} & \frac{3}{\frac{1}{11}} & \frac{4}{\frac{4}{11}} \end{bmatrix}$$

(c) ~~By theorem~~ since I, S

$$\text{By theorem } E_0(T_0) = \frac{1}{\pi_0} = \frac{11}{3}$$

(d) $C = \{3\}$ $A = \{0,1,2,4\}$

$$\begin{cases} g(0) = 1 + \frac{1}{3} g(1) + g(2) \cdot \frac{2}{3} \\ g(1) = 1 + \frac{1}{2} g(4) \\ g(2) = 1 + \frac{3}{4} g(4) \end{cases}$$

$$g(1) = 1 + \frac{1}{2} g(4)$$

$$g(2) = 1 + \frac{3}{4} g(4)$$

$$\begin{cases} g(3) = 0 \\ g(4) = \frac{1}{2}g(0) + \frac{1}{2}g(4) \end{cases}$$

$$\Rightarrow g = \begin{bmatrix} 0 & 1 & 2 & 3 & 4 \\ 1 & 0 & 6 & \frac{7}{2} & 0 & 12 \end{bmatrix}$$

$$E_0(V_A) = 10$$

$$(e) A = \{4\}, R = \{2\}, C = \{0, 1, 3\}$$

$$\begin{cases} h(0) = \frac{1}{5}h(1) \\ h(1) = \frac{1}{2}h(3) + 1 - \frac{1}{2} \\ h(2) = 0 \\ h(3) = h(0) \end{cases}$$

$$h = \begin{bmatrix} \frac{1}{5} & \frac{3}{5} & \frac{3}{5} \\ 0 & 1 & 0 \end{bmatrix}$$

$$P(V_A < V_B | X_0 = 0) = \frac{1}{5}$$

b. Assume that the expected number of visits to state j before the first visit to state i is k with $k \neq 1$

case 1, $k \neq 0$

contradicts the fact that

$$b. E_i(N_{ij} | T_j^{(n)} < T_i) = \sum_{n=1}^{\infty} n \cdot P(T_j^{(n)} < T_i | X_0 = i)$$

$$= P(T_j < T_i | X_0 = i) P(T_j > T_i | X_0 = j) + 2 \cdot P(T_j < T_i | X_0 = i) P(T_j < T_i | X_0 = j) + \dots$$

by Markov Property

$$= P(T_j < T_i | X_0 = i)^2 + 2 \cdot P(T_j < T_i | X_0 = i)^2 P(T_j < T_i | X_0 = j) + \dots$$

since i, j are symmetric
 $1 - P(T_j < T_i | X_0 = i) = P(T_i < T_j | X_0 = j)$

$$P(T_i < T_j | X_0 = i) = P(T_j < T_i | X_0 = j) \quad \text{since } P(T_i = T_j) = 0$$

$$\Rightarrow \sum_{n=1}^{\infty} P(T_j < T_i | X_0 = i)^2 n \cdot P(T_j < T_i | X_0 = j)^{n-1}$$

$$= P(T_j < T_i | X_0 = i)^2 \sum_{n=1}^{\infty} n \cdot P(T_j < T_i | X_0 = j)^{n-1} \quad \text{Let } P(T_i < T_j | X_0 = i) = p$$

$$= (1-p)^2 \left(\frac{1}{1-p} + \frac{p}{1-p} + \frac{p^2}{1-p} + \dots \right)$$

each column is a geometric series with $0 \leq p \leq 1$

$$= (1-p)^2 \sum_{n=0}^{\infty} \frac{p^n}{1-p} = (1-p) \sum_{n=0}^{\infty} p^n = (1-p) \frac{1}{1-p} = 1$$