

1. I have carefully read the integrity statement
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$$\begin{cases} \frac{1}{2}\pi(0) + \frac{1}{4}\pi(1) = \pi(0) \\ \frac{1}{2}\pi(0) + \frac{1}{4}\pi(1) + \frac{1}{6}\pi(2) = \pi(1) \\ \frac{1}{2}\pi(1) + \frac{5}{6}\pi(2) = \pi(2) \\ \pi(0) + \pi(1) + \pi(2) = 1 \end{cases}$$

$$\Rightarrow \pi = \left[\frac{1}{4} \quad \frac{1}{2} \quad \frac{1}{4} \right]$$

~~the stationary distribution~~

Since π is the only solution, it is unique.

(b) Since $i \rightarrow j \forall i, j \in S$, the chain is irreducible

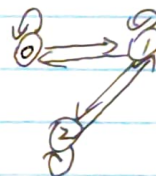
Since $P_{ii} > 0 \forall i \in S$, the chain is aperiodic

By part (a), stationary distribution exists

By theorem,

$$\lim_{n \rightarrow \infty} P(X_n = j | X_0 = i) = \pi(j) \quad \forall i, j \in S$$

$$= \begin{cases} \frac{1}{4} & j=0 \\ \frac{1}{2} & j=1 \\ \frac{1}{4} & j=2 \end{cases}$$



(c) Since I (irreducible), S (stationary distribution exists)

$$E_y(T_y) = \frac{1}{\pi(y)}$$

$$E_0(T_0) = \frac{1}{\pi(0)} = 4$$

$$3. (a) P(X_0=0, X_2=1) = P(X_2=1 | X_0=0) P(X_0=0) = P_{01}^2 \mu(0) = \frac{77}{120} \times \frac{1}{2} = \frac{77}{240}$$

$$(b) P(X_{11}=0 | X_9=2, X_0=3) = \frac{P(X_{11}=0, X_9=2, X_0=3)}{P(X_9=2, X_0=3)} = P(X_{11}=0 | X_9=2) = P_{20}^2 = \frac{11}{20}$$

Markov property

~~the Markov property~~

(c) Assume this MC is at state X_n ,

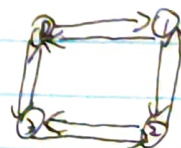
X_{n+1} is even if X_n is odd

X_{n+1} is odd if X_n is even

$\sum_{n=0}^9 X_n$ has odd number of odd terms if $X_0=3$

Hence $\sum_{n=0}^9 X_n$ is odd

$$P(\sum_{n=0}^9 X_n \text{ is even} | X_0=3) = 0$$



4. (a) $\{0, 1\}, \{2, 3\}$

~~every is transient~~

Since $\{2, 3\}$ is closed and $P_{03} > 0$

$P_{00} < 1$ ($0 \rightarrow 3$ and never exits $\{2, 3\}$)

So $\{0, 1\}$ is transient

Since $\{2, 3\}$ is closed and finite,

$\{2, 3\}$ is recurrent by corollary, $\{2, 3\}$ is positive recurrent

(b) $A = \{2\}, B = \{3\}, C = \{0, 1\}$

$$\begin{cases} h(0) = \frac{1}{2}h(0) + \frac{1}{3}h(1) \\ h(1) = \frac{1}{4}h(0) + \frac{1}{2}h(1) + \frac{1}{6} \end{cases} \quad \text{where } h(2) = 1 \text{ and } h(3) = 0$$

$$h = \begin{bmatrix} \frac{1}{3} & \frac{1}{2} \end{bmatrix}$$

The prob. is $\frac{1}{3}$

(c) $A = \{2, 3\}, C = \{0, 1\}$

$$\begin{cases} g(0) = 1 + \frac{1}{2}g(0) + \frac{1}{3}g(1) \\ g(1) = 1 + \frac{1}{4}g(0) + \frac{1}{2}g(1) \end{cases}$$

$$g = \begin{bmatrix} 5 & 9 \end{bmatrix}$$

The expected time is $\frac{9}{2}$

(d) ~~Let~~ $A = \{3\}, C = \{0, 1, 2\}$

$$\begin{cases} g(0) = 1 + \frac{1}{2}g(0) + \frac{1}{3}g(1) \\ g(1) = 1 + \frac{1}{4}g(0) + \frac{1}{2}g(1) + \frac{1}{6}g(2) \\ g(2) = \frac{2}{3}g(2) \end{cases}$$

$$g = \begin{bmatrix} 6 & 6 & 3 \end{bmatrix}$$

Let T_i be the i -th time that it visits state 3.

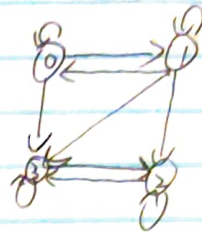
$$E_1(T_3^3) = 1 + \frac{1}{2}E_1(T_3^3) + \frac{1}{3}E_1(T_3^3) + \frac{1}{6}E_1(T_3^3)$$

$$\begin{aligned} &= 1 + \frac{1}{2}E_1(T_3^3) + \frac{1}{3}E_1(T_3^3) + \frac{1}{6}E_1(T_3^3) \\ &= 1 + \frac{1}{2}(3 + 1 + \frac{1}{2}E_1(T_3^3)) \\ &= 1 + \frac{1}{2}(4 + \frac{1}{2}E_1(T_3^3)) \\ &= \frac{7}{2} + \frac{1}{4}E_1(T_3^3) \\ &= \frac{7}{2} \end{aligned}$$

$$E_1(T_3^3) = 6 + \frac{1}{3}E_2(T_3^2) + \frac{1}{3}E_3(T_3^1) + 1$$

$$= 7 + \frac{1}{3}(3 + 1 + \frac{1}{3}E_2(T_3^2)) + \frac{1}{3}(1 + \frac{1}{3}E_2(T_3^2))$$

$$= 10$$



5. (a) Assume $P_{ij} \neq 0$

since $P_{ij} > 0$ (P_{ij} is a probability, $P_{ij} \in [0, 1]$)

and i is recurrent

$P_{ji} = 1$ by theorem

Since $P_{ij} > 0$ and $P_{ji} = 1 > 0$

i and j should be in the same class

But i is null recurrent and j is positive recurrent

$\Rightarrow i$ and j are not in the same class

By contradiction $P_{ij} = 0$

(b) $P_{ij} = P(T_j < \infty | X_0 = j)$

$$= \sum_{i \in S \setminus \{j\}} P(T_j < \infty | X_0 = j, X_1 = i) P(X_1 = i | X_0 = j) + P(T_j < \infty | X_0 = j, X_1 = j) P(X_1 = j | X_0 = j)$$

$$= \sum_{i \in S \setminus \{j\}} P(T_j < \infty | X_1 = i) P_{ji} + P(T_j < \infty | X_0 = j, X_1 = j) P_{jj}$$

If $P_{jj} = 0$

$$P_{ij} = \sum_{i \in S \setminus \{j\}} P_{ij} P_{ji} \quad \text{Markov property}$$

$$= \sum_{i \in S \setminus \{j\}} P_{ji} \quad (P_{ij} = 1)$$

$$= 1 \quad \text{by defn and } P_{jj} = 0$$

If $P_{jj} > 0$

$$P_{ij} = \sum_{i \in S \setminus \{j\}} P_{ij} P_{ji} + P_{jj} \quad (P(T_j < \infty | X_0 = j, X_1 = j) = 1 \text{ if } P_{jj} > 0)$$

$$= \sum_{i \in S \setminus \{j\}} P_{ji} + P_{jj}$$

$$= \sum_{i \in S} P_{ji}$$

$$= 1 \quad \text{by defn}$$

Hence j is recurrent

~~not a martingale~~
 b.(a) $P(X_1=0) = p^3$ (since independence)

(b) $E(Y) = 1 \cdot r + 2 \cdot q = r + 2q$

~~blue extinction~~

extinction happens for sure $\Rightarrow E(Y) \leq 1$

$$E(Y) = r + 2q = 1 - p - q + 2q = 1 - p + q$$

$$1 - p + q \leq 1$$

$$p \geq q$$

$$\varphi(s) = p + rs + qs^2$$

$$p + rs + qs^2 = s$$

$$qs^2 + (r-1)s + p = 0$$

$$qs^2 - (q+p)s + p = 0$$

$$s = \frac{(q+p) \pm \sqrt{(q+p)^2 - 4qp}}{2q}$$

$$= \frac{q+p \pm \sqrt{(q-p)^2}}{2q}$$

$$= \frac{q+p \pm (q-p)}{2q} \quad \text{since } q-p > 0$$

$$s = 1 \text{ or } \frac{p}{q}$$

since $s \in (0, 1)$

$$s = \frac{p}{q}$$

(c) $S = \{0, 1, 2, 4, 6, \dots\}$

$\{0\}$ is the absorbing state

$\{0\}$ is recurrent

~~Since~~ Since $P_{ii} = 0$ (The # of offsprings is even)

$\{1\}$ is transient

Let $i \in S \setminus \{0, 1\}$,

since $P_{i0} > 0$ and 0 is the absorbing state

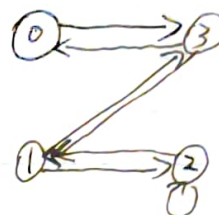
$$P_{ii} < 1$$

so $\{2, 4, 6, \dots\}$ is transient

7. (a) Let X_n be the # of umbrellas at current location

$$S = \{0, 1, 2, 3\}$$

$$P = \begin{matrix} & \begin{matrix} 0 & 1 & 2 & 3 \end{matrix} \\ \begin{matrix} 0 \\ 1 \\ 2 \\ 3 \end{matrix} & \begin{bmatrix} 0 & 1 & 0 & 0 \\ 0 & 0 & 1-p & p \\ 0 & 1-p & p & 0 \\ 1-p & p & 0 & 0 \end{bmatrix} \end{matrix}$$



$$\begin{cases} \pi(0) = (1-p)\pi(3) \\ \pi(1) = (1-p)\pi(2) + p\pi(3) \\ \pi(2) = (1-p)\pi(1) + p\pi(2) \\ \pi(3) = \pi(0) + p\pi(1) \\ \pi(0) + \pi(1) + \pi(2) + \pi(3) = 1 \end{cases}$$

$$\pi = \left[\frac{1-p}{4-p}, \frac{1}{4-p}, \frac{1}{4-p}, \frac{1}{4-p} \right]$$

Since $i \rightarrow j \forall i, j \in S$,
this MC is irreducible

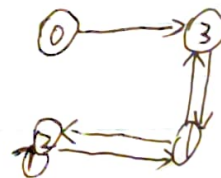
$$I, S \Rightarrow R$$

By theorem $\frac{N_n(0)}{n} = \pi(0) = \frac{1-p}{4-p}$

long-run percentage of time that the professor gets wet is

$$p \frac{1-p}{4-p} = \frac{p-p^2}{4-p}$$

$$(c) P = \begin{matrix} & \begin{matrix} 0 & 1 & 2 & 3 \end{matrix} \\ \begin{matrix} 0 \\ 1 \\ 2 \\ 3 \end{matrix} & \begin{bmatrix} 0 & 0 & 0 & 1 \\ 0 & 0 & 1-p & p \\ 0 & 1-p & p & 0 \\ 0 & 1 & 0 & 0 \end{bmatrix} \end{matrix}$$



$$\begin{cases} \pi(0) = 0 \\ \pi(1) = (1-p)\pi(2) + \pi(3) \\ \pi(2) = (1-p)\pi(1) + p\pi(2) \\ \pi(3) = \pi(0) + p\pi(1) \\ \pi(0) + \pi(1) + \pi(2) + \pi(3) = 1 \end{cases}$$

$$\pi = \left[0, \frac{1}{2+p}, \frac{1}{2+p}, \frac{p}{2+p} \right]$$

Since the state 0 is only visited initially and it does not affect the long-run percentage, we can just remove it and

$$\pi = \left[\frac{1}{2+p}, \frac{1}{2+p}, \frac{p}{2+p} \right]$$

Since $i \rightarrow j \forall i, j \in \{1, 2, 3\}$

this MC is irreducible

$$I, S \Rightarrow R$$

$$\begin{aligned} & \pi(3) + p\pi(1) + p\pi(2) \\ &= \frac{p}{2+p} + \frac{p}{2+p} + \frac{p}{2+p} \\ &= \frac{3p}{2+p} \end{aligned}$$

By theorem $\frac{N_n(0)}{n} = \pi(0)$

The long-run percentage of time that the professor gets wet is

8. (a) $\{0\}$ ~~is~~ is positive recurrent

since $P_{00}=1$ and $E_0(T_0)=1$

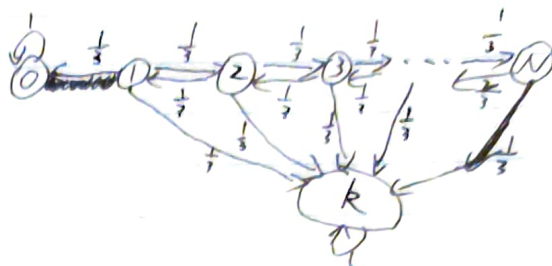
$\{k\}$ is positive recurrent

since $P_{kk}=1$ and $E_k(T_k)=1$

$\{1, 2, 3, \dots, N\}$ is transient

since $\forall i \in \{1, 2, \dots, N\} P_{ik} > 0$ and k is absorbing

so $P_{ii} < 1$



(b) $P(X_4=1 | X_0=1) = P_{11}^4 = \sum_{i \in S} P_{1i}^2 P_{ii}^2$

$P_{11}^2 = \begin{bmatrix} \frac{1}{3} & \frac{1}{9} & 0 & \frac{1}{9} & 0 & \dots & 0 & \frac{4}{9} \end{bmatrix}$ if $N > 3$

$P_{11}^2 = \begin{bmatrix} 0 & \frac{1}{9} & 0 & \frac{1}{9} & 0 & \dots & 0 \end{bmatrix}^T$ if $N > 3$

$P_{11}^4 = \frac{1}{81} + \frac{1}{81}$ (if $N > 3$)

$= \frac{2}{81}$

$P_{11}^2 = \begin{bmatrix} \frac{1}{3} & \frac{1}{9} & 0 & \frac{1}{9} & \frac{4}{9} \end{bmatrix}$ if $N=3$

$P_{11}^2 = \begin{bmatrix} 0 & \frac{1}{9} & 0 & \frac{2}{9} & 0 \end{bmatrix}$ if $N=3$

$P_{11}^4 = \frac{1}{81} + \frac{2}{81}$
 $= \frac{1}{27}$

(c) $P = \begin{matrix} & \begin{matrix} 0 & 1 & 2 & 3 & k \end{matrix} \\ \begin{matrix} 0 \\ 1 \\ 2 \\ 3 \\ k \end{matrix} & \begin{bmatrix} 1 & 0 & 0 & 0 & 0 \\ \frac{1}{3} & 0 & \frac{1}{3} & 0 & \frac{1}{3} \\ 0 & \frac{1}{3} & 0 & \frac{1}{3} & \frac{1}{3} \\ 0 & 0 & \frac{2}{3} & 0 & \frac{1}{3} \\ 0 & 0 & 0 & 0 & 1 \end{bmatrix} \end{matrix}$

$A = \{0\}$ $B = \{k\}$ $C = \{1, 2, 3\}$

$h(1) = \frac{1}{3} + \frac{1}{3}h(2)$

$(h(0)=1, h(k)=0)$

$h(2) = \frac{1}{3}h(1) + \frac{1}{3}h(3)$

$h(3) = \frac{2}{3}h(2)$

$h = \begin{bmatrix} \frac{7}{18} & \frac{2}{6} & \frac{1}{9} \end{bmatrix}$

The probability is $\frac{1}{6}$

$$9. (a) P(X_3 \neq 0, X_2 = 0 | X_0 = 1)$$

$$= P(X_3 = 1, X_2 = 0 | X_0 = 1) + P(X_3 = 2, X_2 = 0 | X_0 = 1)$$

$$= \frac{P(X_3 = 1, X_2 = 0, X_0 = 1) + P(X_3 = 2, X_2 = 0, X_0 = 1)}{P(X_0 = 1)}$$

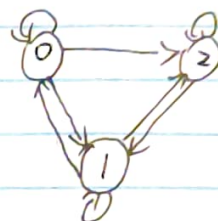
$$= \frac{P(X_3 = 1 | X_2 = 0) P(X_2 = 0 | X_0 = 1) P(X_0 = 1) + P(X_3 = 2 | X_2 = 0) P(X_2 = 0 | X_0 = 1) P(X_0 = 1)}{P(X_0 = 1)}$$

Markov Property

$$= P_{01} P_{10}^2 + P_{02} P_{10}^2$$

$$= \frac{1}{4} \times \frac{3}{16} + \frac{1}{4} \times \frac{3}{16}$$

$$= \frac{3}{32}$$



$$(b) \begin{cases} \pi(0) = \frac{1}{2} \pi(0) + \frac{1}{4} \pi(1) \\ \pi(1) = \frac{1}{4} \pi(0) + \frac{1}{4} \pi(1) + \frac{5}{8} \pi(2) \\ \pi(2) = \frac{1}{4} \pi(0) + \frac{1}{2} \pi(1) + \frac{3}{8} \pi(2) \end{cases}$$

$$\pi = \begin{bmatrix} \pi(0) \\ \pi(1) \\ \pi(2) \end{bmatrix}^T = \begin{bmatrix} 1 \\ 2 \\ 2 \end{bmatrix}^T \text{ where } X \in \mathbb{R} \text{ (This gives all the stationary measure of } P)$$

since this chain is irreducible ($i \rightarrow j \forall i, j \in S$)

and recurrent (finite and closed)

By theorem, $\mu_0(1) = E(\sum_{n=0}^{\infty} \mathbb{1}_{\{X_n = 1\}})$

$$x=1 \quad \mu_0(0)=1$$

Since π is still a stationary measure and $\pi(0) = 1 = \mu_0(0)$

$$\mu_0 E(\sum_{n=0}^{\infty} \mathbb{1}_{\{X_n = 1\}}) = 2$$

$$(c) P(X_n \neq 1, X_{n-1} \neq 1, \dots, X_1 \neq 1 | X_0 = 1)$$

$$= (\frac{1}{2})^n + \sum_{k=0}^{n-1} (\frac{1}{2})^k \frac{1}{4} (\frac{3}{8})^{n-k-1}$$

$$= 2^{-n} + \frac{1}{4} (\frac{3}{8})^{n-1} \sum_{k=0}^{n-1} (\frac{4}{3})^k$$

$$= 2^{-n} + \frac{1}{4} (\frac{3}{8})^{n-1} \frac{1 - (\frac{4}{3})^n}{-\frac{1}{3}}$$

$$= 2^{-n} - \frac{3}{4} (\frac{3}{8})^{n-1} + 2^{-n+1}$$

$$= (\frac{1}{2})^n \cdot 3 - \frac{3}{4} (\frac{3}{8})^n \cdot \frac{2}{3}$$

$$= (\frac{1}{2})^n (3 - 2 (\frac{3}{4})^n)$$

$$\text{Have } a = \frac{1}{2} \quad b = \frac{3}{4}$$

all the time
(the chain stays at 0 or stays at 0 k times, move to 2 and stays at 2 for the rest of the times)

10(a) The chain is ~~irred~~ irreducible

Since ~~the~~ S is finite

By theorem, at least one of the states is positive recurrent
Hence since it is irreducible

~~All~~ S is ~~rec~~ positive recurrent

By observation, all the cycles have lengths

So the # of steps it takes to go back ^{even} to the same state is even
Hence the period of all states is 2

It is not aperiodic

(b) $A = \{7\}$ $C = \{0, 1, 2, 3, 4, 5, 6\}$

$$\begin{cases} g(0) = 1 + \frac{1}{3}g(1) + \frac{1}{3}g(3) + \frac{1}{3}g(5) \\ g(1) = 1 + \frac{1}{3}g(0) + \frac{1}{3}g(2) + \frac{1}{3}g(4) \\ g(2) = 1 + \frac{1}{3}g(1) + \frac{1}{3}g(3) + \frac{1}{3}g(6) \\ g(3) = 1 + \frac{1}{3}g(0) + \frac{1}{3}g(2) + \frac{1}{3}g(6) \\ g(4) = 1 + \frac{1}{3}g(1) + \frac{1}{3}g(5) + \frac{1}{3}g(6) \\ g(5) = 1 + \frac{1}{3}g(0) + \frac{1}{3}g(4) + \frac{1}{3}g(6) \\ g(6) = 1 + \frac{1}{3}g(3) + \frac{1}{3}g(5) + \frac{1}{3}g(6) \\ g(7) = 1 \end{cases}$$

$g(0) = 10$ \swarrow starting from 0

The expected number of steps until the chain reaches state 7 ^{for} the first time is 10

11. (a) Define a distribution π where $\pi(i) = P_{ii}$ ($i \in S$)

$$\sum_{i \in S} \pi(i) P_{ij} = \sum_{i \in S} P_{ii} P_{ij}$$

$$= \sum_{i \in S} P_{ii} P_{ij}$$

$$= P_{aj}$$

$$= P_{aj} \quad (P=P^2)$$

$$= \pi(j)$$

$$\sum_{i \in S} \pi(i) = \sum_{i \in S} P_{ii} = 1 \text{ by defn of DTM}$$

Hence π is a stationary distribution

Since it is irreducible

By theorem, $\{X_n\}_{n=0,1,2,\dots}$ is positive recurrent

(b) Since $P=P^2$

$$\Rightarrow P^n = P \quad \forall n \in \mathbb{N}$$

By part (a), all states are positive recurrent

$$\exists n, \text{ s.t. } P_{ii}^n > 0 \quad \forall i \in S$$

$$\text{Hence } P_{ii} > 0 \quad \forall i \in S$$

\Rightarrow it is aperiodic

Since A (aperiodic), I (irreducible), S (stationary distribution exists from part (a))

$$\lim_{n \rightarrow \infty} P_{xy}^n = \pi(y)$$

$$\pi(x) P_{xy} = \pi(x) \lim_{n \rightarrow \infty} P_{xy}^n = \pi(x) \pi(y) = \pi(y) \lim_{n \rightarrow \infty} P_{yx}^n = \pi(y) P_{yx}$$

By theorem, it is time reversible

(c) By theorem, the chain can be partitioned into T, U_1, U_2, \dots where T are all the transient states and $U_i, i=1,2,\dots$ are closed recurrent classes

$i, j \in S$ are

① Assume i is not positive recurrent

$$\Rightarrow \pi(i) = 0 = \pi(j) \text{ by theorem}$$

$$\Rightarrow \pi(i) P_{ij} = \pi(j) P_{ji}$$

Which satisfies detail balance condition

② Assume i is positive recurrent and j is not

$$\Rightarrow \pi(i) P_{ij} = \pi(i) \cdot 0 = 0 \text{ since the class that } i \text{ is in is closed} \Rightarrow i \neq j$$

$$\Rightarrow \pi(j) P_{ji} = 0 \cdot P_{ji} = 0 \text{ by theorem}$$

③ Assume i and j are positive recurrent

If i and j are not in the same class, $\pi(i) P_{ij} = 0 = \pi(j) P_{ji}$ (closed) condition

else, since the class is irreducible, by part (b) it satisfies the detail balance

Hence, the chain is time reversible

$$- 12 (a) \pi(i) = \sum_{j=0}^{\infty} \pi(j) p_{ji}$$

$$= \sum_{j=0}^{\infty} \pi(j) p_{ji}$$

Since $p_{ji} = p_{i-j}$, $0 < i \leq j+1$, and $i > 0$

$$\Rightarrow p_{ji} = \begin{cases} p_{i-j} & -1 < i-j \leq j \\ 0 & \text{otherwise} \end{cases}$$

$$\Rightarrow \pi(i) = \sum_{j=i-1}^{\infty} \pi(j) p_{ji}$$

(b) $\Phi(\alpha) = q_0 + q_1 \alpha + q_2 \alpha^2 + \dots$ is the generating function of $\{q_n\}$
 $= p_1 + p_0 \alpha + p_1 \alpha^2 + \dots$

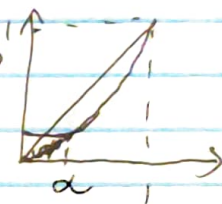
$\Rightarrow \Phi'(\alpha) = q_1 + 2q_2 \alpha + 3q_3 \alpha^2 + \dots > 0$ since all terms are non-negative and at least one term is positive, otherwise, p_1, p_0, p_1, \dots wouldn't be a distribution

$\Phi''(\alpha) = 2q_2 + 6q_3 \alpha + 12q_4 \alpha^2 + \dots > 0$ for the same reason as above

$\Rightarrow \Phi(\alpha)$ is increasing and convex $\leftarrow \Phi(1) = \sum_{k=0}^{\infty} q_k = 1$

$$\Phi'(1) = p_0 + 2p_1 + 3p_2 + \dots = \sum_{k=0}^{\infty} k q_k > 1 \text{ since } k \in [0, \infty), q_k > 0$$

Hence there exists an $\alpha \in (0, 1)$, s.t. $\alpha = q_0 + q_1 \alpha + q_2 \alpha^2 + \dots$



(c) Let $\pi \sim \text{Geo}(1-\alpha)$

By part (a) $\pi(i) = \sum_{j=1}^{\infty} \pi(j) p_{ji}$, for $i > 0$

$$\pi(i) = \pi(i-1) p_1 + \pi(i) p_0 + \pi(i+1) p_1 + \dots$$

$$(1-\alpha) \alpha^i = (1-\alpha) \alpha^{i-1} q_0 + (1-\alpha) (1-\alpha)^{i-1} q_1 + \dots$$

$$\alpha = q_0 + \alpha q_1 + \alpha^2 q_2 + \dots$$

Since α is the solution, $\pi(i) = (\pi P)(i)$ for $i > 0$

$$\text{For } i=0, \pi(0) = \pi(0) \sum_{j=0}^{\infty} p_{j0} + \pi(1) \sum_{j=0}^{\infty} p_{j1} + \dots$$

$$(1-\alpha) \sum_{j=0}^{\infty} p_{j0} + (1-\alpha) \alpha \sum_{j=0}^{\infty} p_{j1} + (1-\alpha) \alpha^2 \sum_{j=0}^{\infty} p_{j2} + \dots$$

$$= \sum_{j=0}^{\infty} \sum_{i=0}^{\infty} p_{ji} \alpha^i \cdot (1-\alpha)$$

$$= \begin{bmatrix} p_0 + p_1 \alpha + p_2 \alpha^2 + p_3 \alpha^3 + \dots \\ + p_1 + p_2 \alpha + p_3 \alpha^2 + \dots \\ + p_2 + p_3 \alpha + \dots \end{bmatrix} (1-\alpha)$$

$$= (1-\alpha) \sum_{i=1}^{\infty} q_i \sum_{j=0}^{i-1} \alpha^j$$

$$= (1-\alpha) \sum_{i=1}^{\infty} q_i \frac{1-\alpha^i}{1-\alpha}$$

$$= \sum_{i=1}^{\infty} q_i - \sum_{i=1}^{\infty} q_i \alpha^i$$

$$= 1 - q_0 - (\alpha - q_0) \text{ by part (b)}$$

$$= 1 - \alpha$$

$$= \pi(0)$$

Hence $\pi P = \pi$

Since π is already a distribution

$$\sum_{i=0}^{\infty} \pi(i) = 1$$

Hence $\pi \sim \text{Geo}(1-\alpha)$ is a stationary distribution

$$13. (a) P(Y_{n+1}=j | Y_n=i)$$

$$= P(X_{E_1} + \dots + E_n + E_{n+1} = j | X_{E_1} + \dots + E_n = i)$$

$$= \sum_{k=1}^{\infty} P(X_{E_1} + \dots + E_n + E_{n+1} = j | X_{E_1} + \dots + E_n = i, E_{n+1} = k) P(E_{n+1} = k)$$

$$= \sum_{k=1}^{\infty} P_{ij}^k P_k$$

(b) since $\{X_n\}_{n=0,1,\dots}$ is irreducible

$$\exists n, \text{ s.t. } P_{ij}^n > 0, \forall i, j \in S$$

Let P' be the transition matrix of $\{Y_n\}_{n=0,1,\dots}$

$$P'_{ij} = \sum_{k=1}^{\infty} P_{ij}^k P_k \geq P_{ij}^n P_n > 0 \quad (P_n > 0)$$

$i \rightarrow j, \forall i, j \in S \Rightarrow \{Y_n\}_{n=0,1,\dots}$ is irreducible

By the proof from part (c) and the chain is irreducible

$\Rightarrow \{Y_n\}_{n=0,1,\dots}$ is positive recurrent

(c) The stationary distribution of $\{Y_n\}_{n=0,1,\dots}$ is also π

$$\sum_{i \in S} \pi(i) P'_{ij}$$

$$= \sum_{i \in S} \pi(i) \sum_{k=1}^{\infty} P_{ij}^k P_k$$

$$= \sum_{k=1}^{\infty} P_k \sum_{i \in S} \pi(i) P_{ij}^k$$

$$= \sum_{k=1}^{\infty} P_k \pi(j)$$

$$(\pi = \pi P^n \quad \forall n \in \mathbb{N})$$

$$= \pi(j)$$

$$\left(\sum_{k=1}^{\infty} P_k = 1 \right)$$

$$\text{Clearly } \sum_{i \in S} \pi(i) = 1$$

Hence, the stationary distribution of $\{Y_n\}_{n=0,1,\dots}$ is π