

PSET-11 Q1)

- Pick eigen vectors corresponding to largest eigen value of the covariance matrix Σ .
- Reduce to one dimension (PCA)
- To reduce to two dimensions → pick ~~second largest~~ ^{second largest} var eigen vector corresponding to second largest eigen value.
- We implement PCA in order to find best set of directions such that variability of data is maximized in the reduced dimensional space.

$$\text{i.e. } x' = Xp$$

↓
new matrix after dimensional reduction. (PCA)

$$\max \|x'\|^2 = \max \|Xp\|^2$$

$$P^T P = 1$$

$$= \max \text{tr}((Xp)^T (Xp))$$

$$\text{such that } P^T P = 1$$

$$= \max (p^T X^T X p)$$

$$\text{such that } P^T P = 1$$

$$= \max (p^T S p)$$

$$\text{such that } P^T P = 1$$

where P is the direction.

S = Scatter matrix

(Un normalized
Covariance matrix)

Using Lagrangian multipliers:-

$$\max (p^T S p) \quad \text{s.t. } P^T P = 1$$

$$\Rightarrow \max (p^T S p) - \lambda (P^T P - 1) \quad (\lambda > 0)$$

This is used to solve constrained optimization by introducing λ to combine $p^T S p$ and the constraint

Then solve for stationary point:

$$\frac{\partial L(p, \lambda)}{\partial \lambda} = 0 \quad \text{where } L(p, \lambda) = p^T S p - \lambda(p^T p - 1)$$

$$\rightarrow p^T p = 1$$

$$\frac{\partial L}{\partial p} \rightarrow S p = \lambda p$$

$$\therefore L(p, \lambda) = p^T S p - \lambda(p^T p - 1) = \lambda p^T p = \lambda$$

\therefore optimization reduces to max λ such that $\lambda p = S p$

This means eigenvector of p of S is found corresponding to maximum eigenvalue λ .

To find k directions, find eigenvectors corresponding to each of k eigenvalues sorted in descending order.

i.e. take p_1, p_2 which $S p = \lambda p$

p_1 is the eigenvector corresponding to highest λ .

\therefore To maximize value, p_2 should correspond to second highest eigenvalue.

$$\lambda = \begin{bmatrix} \lambda_1 & 0 & \dots & \dots \\ 0 & \lambda_2 & \dots & \dots \\ 0 & 0 & \lambda_3 & \dots \\ 0 & 0 & 0 & \dots \end{bmatrix}$$

②

Let $x_i \in \mathbb{R}^n$ be a sample
 & let X be the matrix of all samples x_i .

$$\Rightarrow X \in \mathbb{R}^{m \times n}$$

Dimensionality reduction using PCA:-

On applying PCA to X , we obtain a matrix of principle components, $V \in \mathbb{R}^{n \times d}$ (assuming $d = \text{rank}(X)$)

$$\Rightarrow \text{Reconstructed matrix } \boxed{X' = X V V^T}$$

$$\text{Loss function } J = \frac{1}{m} \sum_{i=1}^m (\text{original matrix} - \text{reconstructed matrix})^2 \quad \forall x_i \in X$$

$$= \frac{1}{m} \sum_{i=1}^m (x_i' - x_i) (x_i' - x_i)^T$$

$$= \frac{1}{m} \sum_{i=1}^m (x_i V V^T - x_i) (x_i V V^T - x_i)^T$$

Minimising the objective function J using gradient descent:-

J is a function of V ($\because X' = X V V^T$)

$$\theta_j \leftarrow \theta_j - \alpha \frac{\partial}{\partial \theta_j} J(\theta_j)$$

Here $\theta_j \equiv V$

$$V \leftarrow V - \alpha \frac{\partial}{\partial V} \left[\frac{1}{m} \sum_{i=1}^m (x_i V V^T - x_i) (x_i V V^T - x_i)^T \right]$$

$$\frac{dJ(V)}{dV} = \frac{d}{dV} \left(\frac{1}{m} \sum_{i=1}^m (x_i V V^T - x_i) (x_i V V^T - x_i)^T \right)$$

$$= \frac{1}{m} \sum_{i=1}^m \left[2 (x_i V V^T - x_i)^T (2 x_i V) \right]$$

$$= \frac{4}{m} \sum_{i=1}^m \left[x_i (V V^T - I) \right]^T (x_i V) = \frac{4}{m} \sum_{i=1}^m (V V^T - I) (x_i^T x) V$$

Hence, update function

$$V \leftarrow V - \frac{\eta \kappa}{m} \sum_{i=1}^m (V V^T - I) (x_i^T x) V$$

Hence, update function

$$V \leftarrow V - \frac{4\alpha}{m} \sum_{i=1}^m (VV^T - I)(x_i^T x) V$$

3. In case of PCA

Loss function = reconstruction error.

each sample $x_i \in \mathbb{R}^n$

Given sample data $X \in \mathbb{R}^{m \times n}$

Reconstructed data $X' = XVV^T$

where $V \in \mathbb{R}^{n \times d}$ is matrix of principal components.

$$RSS = J = \frac{1}{m} \sum (\text{error})^2 \quad \forall \text{ samples.}$$

$$= \frac{1}{m} \sum_{i=1}^m (x_i' - x_i)(x_i' - x_i)^T \quad \forall x_i \in X$$

$$\Rightarrow J(V) = \frac{1}{m} \sum_{i=1}^m (x_i V V^T - x_i)(x_i V V^T - x_i)^T \quad \forall x_i \in X$$

$$\frac{dJ(V)}{dV} = \frac{4}{m} \sum_{i=1}^m (VV^T - I)(x_i^T x) V$$

(i) For L_1 regularisation (Lasso regression) $\Rightarrow RSS + L_1\text{-norm}$

$$L^{\text{Lasso}}(\lambda) = (J(V) + \lambda \|V\|_1)$$

$$\Rightarrow \min \left(\frac{1}{m} \sum_{i=1}^m [x_i V V^T - x_i][x_i V V^T - x_i]^T + \lambda \|V\|_1 \right)$$

$$\Rightarrow \frac{\partial}{\partial V} L^{\text{Lasso}}(\lambda) = 0$$

$$\Rightarrow \frac{4}{m} \sum_{i=1}^m (VV^T - I)(x_i x_i^T) V + \lambda (\text{sign}(V)) = 0$$

(ii) For L_2 regularisation (Ridge regression) \Rightarrow RSS + L_2 -norm.

$$L^{\text{Ridge}}(\lambda) = (J(\mathbf{V}) + \lambda \|\mathbf{V}\|_2)$$

$$\Rightarrow \min \left(\frac{1}{m} \sum_{i=1}^m [\mathbf{x}_i \mathbf{V} \mathbf{V}^T - \mathbf{x}_i] [\mathbf{x}_i \mathbf{V} \mathbf{V}^T - \mathbf{x}_i]^T + \lambda \|\mathbf{V}\|_2 \right)$$

$$\Rightarrow \frac{\partial}{\partial \mathbf{V}} L^{\text{Ridge}}(\lambda) = 0$$

$$\Rightarrow \frac{1}{m} \sum_{i=1}^m (\mathbf{V} \mathbf{V}^T - \mathbf{I}) (\mathbf{x}_i^T \mathbf{x}_i) \mathbf{V} + \lambda \frac{\mathbf{V}}{\|\mathbf{V}\|_2} = 0$$