

# AI Programming: Lecture 4

## Variational Inference

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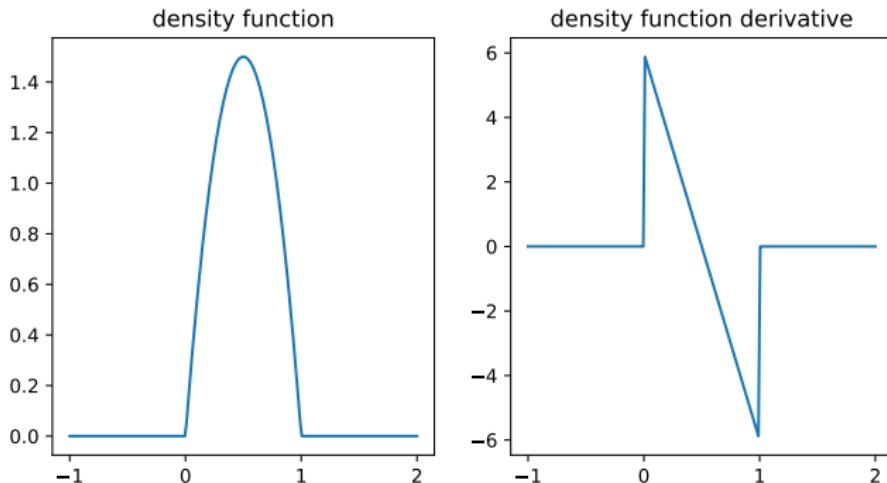
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## From Constrained to Unconstrained Model

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# From Constrained to Unconstrained Model



## From Constrained to Unconstrained Model

Every density that is defined for constrained continuous variables can be transformed to an unconstrained density. For  $Y = H(X)$

$$p_Y(y) = p_X(H^{-1}(y)) \cdot \left| \frac{d}{dy} H^{-1}(y) \right| \quad \text{defined for } y \in H^{-1}(\text{support}(X))$$

$$p_X(x) = p_Y(H(x)) \cdot \left| \frac{d}{dx} H(x) \right| \quad \text{defined for } x \in \text{support}(X)$$

## From Constrained to Unconstrained Model

Example:

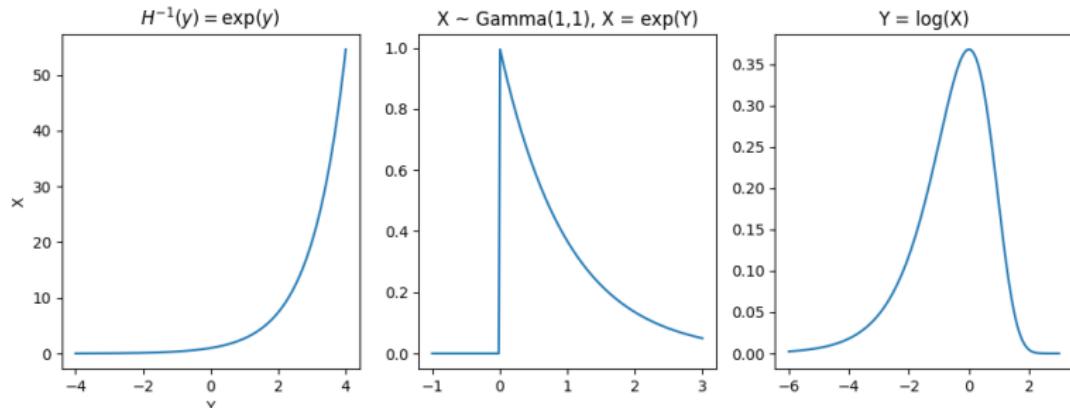
$X \sim \text{Gamma}(1, 1)$  is supported on  $[0, \infty)$

Choose  $y = H(x) = \log(x)$

$$p_Y(y) = p_X(\exp(y)) \cdot \exp(y)$$

is supported on  $\mathbb{R}$ .

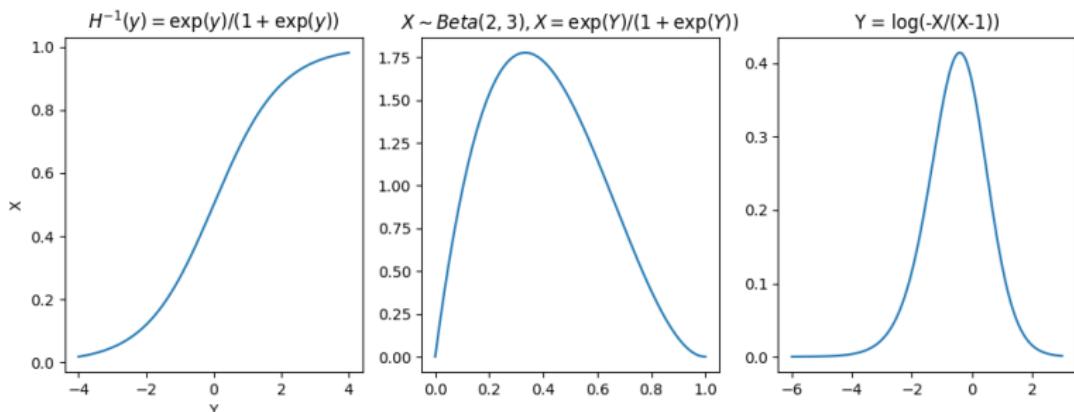
# From Constrained to Unconstrained Model



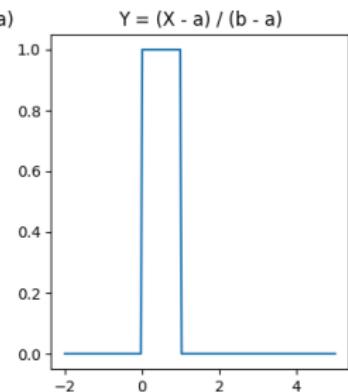
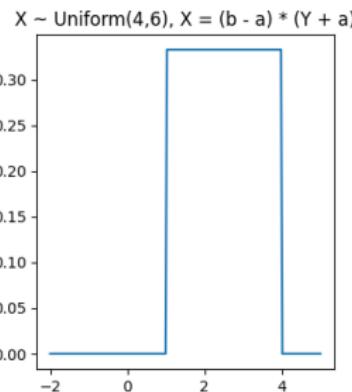
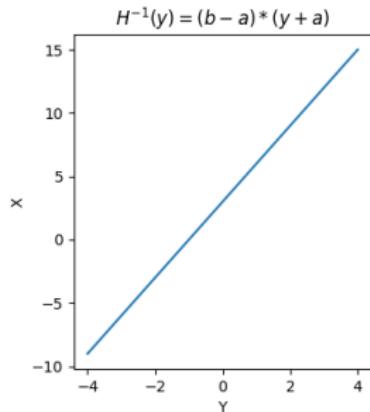
With pytorch:

```
T = dist.transform_to(dist_constrained.support)
dist_unconstrained = dist.TransformedDistribution(dist_constrained, T.inv)
```

# From Constrained to Unconstrained Model



# From Constrained to Unconstrained Model



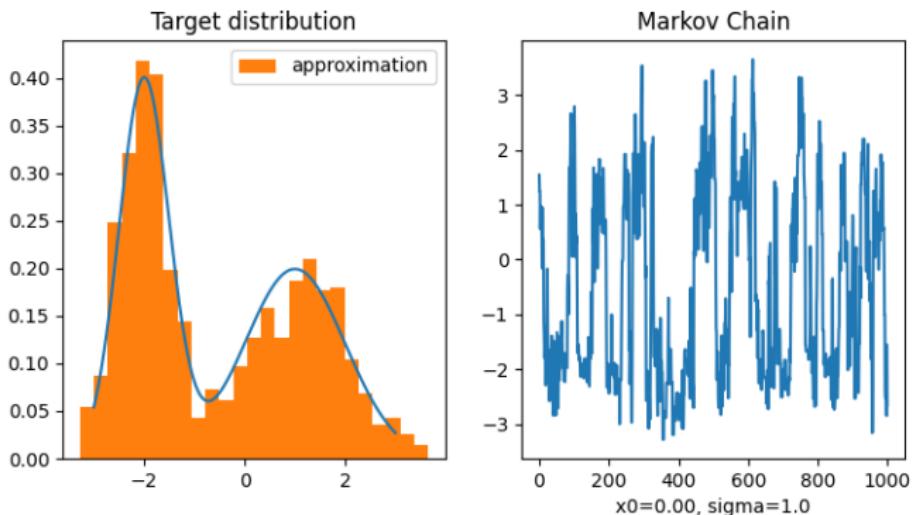
# Variational Inference

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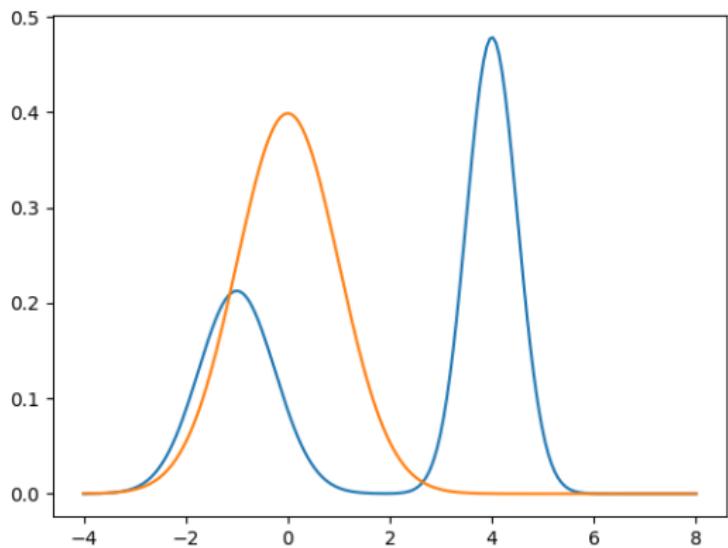
# Variational Inference

- Assumption: Finite number of continuous (unconstrained) random variables
- Idea: Approximate the posterior with a fixed-form parameterised variational distribution.
- Optimise the parameters of the variational distribution to be "close" to true posterior.

# Variational Inference



# Variational Inference



Adjust mean and deviation of orange Normal distribution to "fit" blue distribution.

How to measure distance between two distributions?

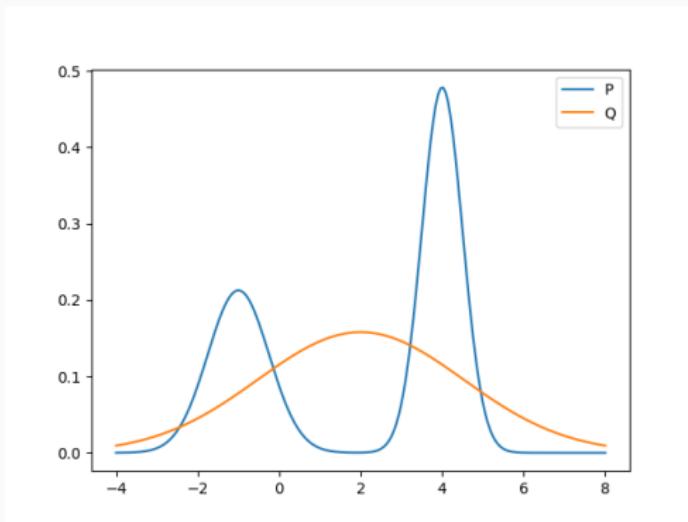
Kullback-Leibler Divergence

$$\begin{aligned} D_{\text{KL}}(P \parallel Q) &= \mathbb{E}_{X \sim P} \left[ \log \left( \frac{P(X)}{Q(X)} \right) \right] \\ &= \underbrace{\mathbb{E}_{X \sim P} [\log P(X)]}_{=-\text{entropy}} - \mathbb{E}_{X \sim P} [\log Q(X)] \end{aligned}$$

It is the amount of information lost when  $Q$  is used to approximate  $P$ .

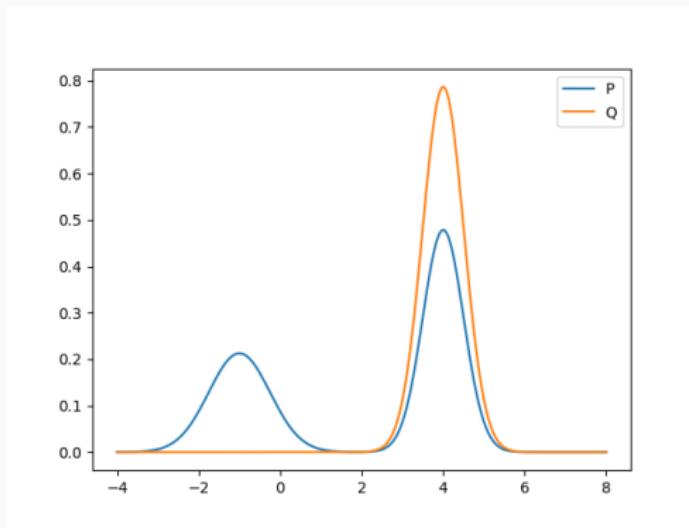
# Variational Inference: Kullback Leibler Divergence

$$\min_Q D_{\text{KL}}(P \parallel Q) = \min_Q \mathbb{E}_{X \sim P} \left[ \log \left( \frac{P(X)}{Q(X)} \right) \right]$$



# Variational Inference: Kullback Leibler Divergence

$$\min_Q D_{\text{KL}}(Q \parallel P) = \min_Q \mathbb{E}_{X \sim Q} \left[ \log \left( \frac{Q(X)}{P(X)} \right) \right]$$



KL-Divergence is not symmetric!

# Variational Inference: Kullback Leibler Divergence

How to compute  $D_{\text{KL}}(P \parallel Q)$ ?

- Discrete Variables:

$$D_{\text{KL}}(P \parallel Q) = \sum_x P(x) \log \left( \frac{Q(x)}{P(x)} \right)$$

- Continuous Variables:

$$D_{\text{KL}}(P \parallel Q) = \int_{-\infty}^{\infty} P(x) \log \left( \frac{Q(x)}{P(x)} \right) dx$$

- Sampling!

$$D_{\text{KL}}(P \parallel Q) \approx \frac{1}{N} \sum_{i=1}^N \log \left( \frac{Q(x_i)}{P(x_i)} \right) \quad x_i \sim P$$

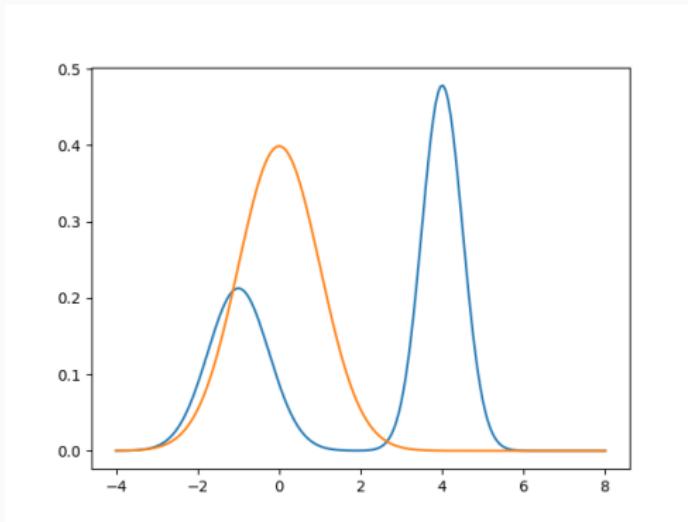
# Variational Inference: Kullback Leibler Divergence

```
# numerical integration
def integrate(func: Callable[[torch.Tensor], torch.Tensor], a, b, N=500):
    x = torch.linspace(a, b, N)
    y = func(x)
    return torch.trapz(y, x)

def kl_divergence_1(p: dist.Distribution, q: dist.Distribution, a, b, N=500):
    def integrand(x: torch.Tensor) -> torch.Tensor:
        px = p.log_prob(x).exp()
        qx = q.log_prob(x).exp()
        l = px * torch.log(px / qx)
        l[px == 0] = 0
    return integrate(integrand, a, b, N)

# sampling
def kl_divergence_2(p: dist.Distribution, q: dist.Distribution, N=500):
    x = p.sample((N,))
    px = p.log_prob(x).exp()
    qx = q.log_prob(x).exp()
    l = torch.log(px / qx)
    return l.mean()
```

# Variational Inference: Kullback Leibler Divergence



```
kl_divergence_1(normal_mixture, normal, -5, 7, N=10000)
```

Returns:

```
tensor(4.5455)
```

```
kl_divergence_2(normal_mixture, normal, N=1000000)
```

Returns:

```
tensor(4.5453)
```

# Variational Inference: Kullback Leibler Divergence

For Bayesian inference  $P = P(\Theta|X)$ .

We cannot sample from posterior, so we are bound to

$D_{\text{KL}}(Q \parallel P) = \mathbb{E}_{\Theta \sim Q} \left[ \log \left( \frac{Q(\Theta)}{P(\Theta|X)} \right) \right]$ , where we sample from the variational distribution.

We cannot evaluate  $P(\Theta|X)$  and we rewrite to joint.

$$\begin{aligned} D_{\text{KL}}(Q \parallel P) &= \mathbb{E}_{\Theta \sim Q} \left[ \log \left( \frac{Q(\Theta)}{P(\Theta|X)} \right) \right] = \mathbb{E}_{\Theta \sim Q} \left[ \log \left( \frac{Q(\Theta)P(X)}{P(\Theta, X)} \right) \right] \\ &= \mathbb{E}_{\Theta \sim Q} [\log Q(\Theta) - \log P(\Theta, X)] + \mathbb{E}_{\Theta \sim Q} [\log P(X)] \\ &= \underbrace{-\mathbb{E}_{\Theta \sim Q} [\log P(\Theta, X) - \log Q(\Theta)]}_{\text{ELBO}(Q, P)} + \underbrace{\log P(X)}_{\text{log-evidence}} \\ \underbrace{\log P(X)}_{\text{constant}} &= D_{\text{KL}}(Q \parallel P) + \text{ELBO}(Q, P) \end{aligned}$$

Maximising the ELBO w.r.t.  $Q$  minimizes  $D_{\text{KL}}(Q \parallel P)$ !

# Variational Inference: Example

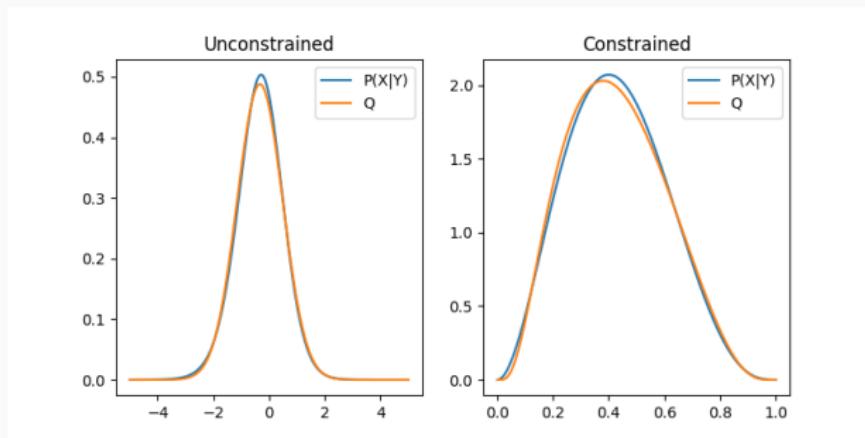
Coin model:

```
flips = torch.tensor([0.,1.,1.,0.,0.])
prior = dist.Uniform(0,1)
posterior = dist.Beta(1+flips.sum(), 1+(1-flips).sum()) # analytical solution
def likelihood(p):
    return dist.Bernoulli(p).log_prob(flips.reshape(-1,1)).sum(dim=0).exp()

# transform to unconstrained model
T = dist.transforms.transform_to(posterior.support) # sigmoid: R -> [0,1]
posterior_unconstrained = dist.TransformedDistribution(posterior, T.inv)
prior_unconstrained = dist.TransformedDistribution(prior, T.inv)
```

# Variational Inference: Example

```
def DKL_QP(X):
    # variational distribution is Normal parameterised with mean and sigma
    mu, sigma = X
    q = dist.Normal(mu, sigma)
    return kl_divergence(q, posterior_unconstrained, mu-4*sigma, mu+4*sigma)
# minimize by brute force (grid search)
mu, sigma = brute(DKL_QP, [(-3,3), (0.1,1.)], Ns=100)
variational = dist.Normal(mu, sigma)
```



# Variational Inference

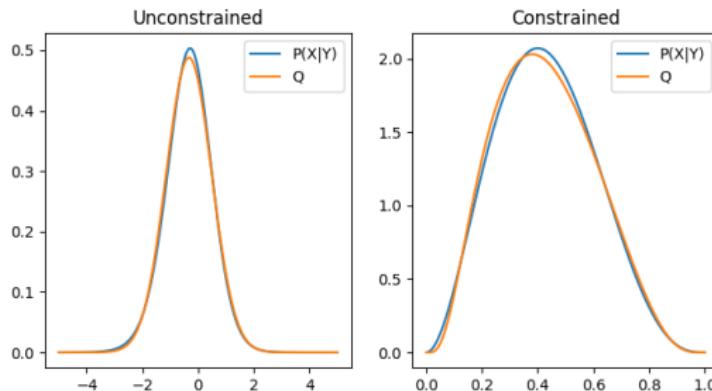
Why don't we fit Gaussian to constrained posterior?

$$D_{\text{KL}}(Q \parallel P) = \mathbb{E}_{\Theta \sim Q} \left[ \log \left( \frac{Q(\Theta)}{P(\Theta|X)} \right) \right]$$

Only well-defined if  $P(\theta|X) = 0 \implies Q(\theta) = 0$ .

# Variational Inference: Example

```
def ELBO_loss(X):
    mu, sigma = X
    q = dist.Normal(mu, sigma)
    def integrand(x: torch.Tensor) -> torch.Tensor:
        px = likelihood(T(x)) * prior_unconstrained.log_prob(x).exp()
        qx = q.log_prob(x).exp()
        l = qx * torch.log(px / qx)
        l[px == 0] = 0
    return -integrate(integrand, mu-4*sigma, mu+4*sigma) # -ELBO
mu, sigma = brute(ELBO_loss, [(-3,3), (0.1,1.)], Ns=100) # minimize
```

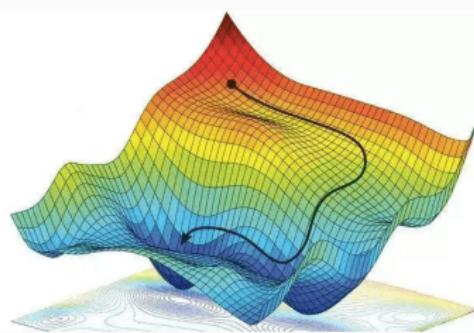
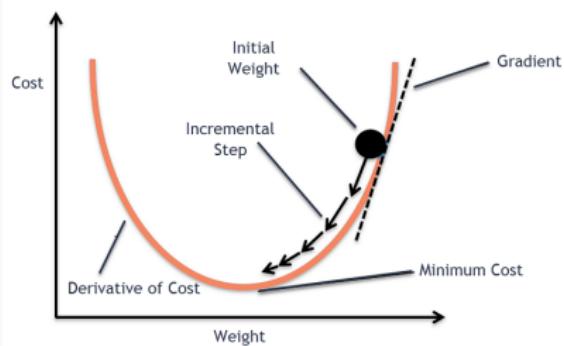


# Automatic Differentiation Variational Inference (ADVI)

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- Computing  $D_{KL}$  with numerical integration → estimate with sampling
- Optimise parameters with grid-search → perform gradient descent
- Single variable → multiple variables

# ADVI: Gradient Descent



# ADVI: Automatic Differentiation (AD)

```
X = torch.tensor(2., requires_grad=True)  
Y = X**2 + torch.log(X)  
Y.backward() # compute gradients  
X.grad.item(), (2*X + 1/X).item()
```

Returns:

(4.5, 4.5)

# ADVI: Variational Distribution

Mean-field Gaussian:

Independent Gaussian for each of  $K$  variables:

$$Q(\boldsymbol{\theta}; \phi) = \text{MvNormal}(\boldsymbol{\theta}; \boldsymbol{\mu}, \text{diag}(\boldsymbol{\sigma}^2)) = \prod_{k=1}^K \text{Normal}(\theta_k; \mu_k, \sigma_k^2)$$

$Q$  parameterised by unconstrained  $\phi = (\boldsymbol{\mu}, \boldsymbol{\omega})$  with  $\boldsymbol{\sigma} = \exp(\boldsymbol{\omega})$ .

Full-rank Gaussian:

Correlated Gaussians

$$Q(\boldsymbol{\theta}; \phi) = \text{MvNormal}(\boldsymbol{\theta}; \boldsymbol{\mu}, \boldsymbol{\Sigma})$$

$Q$  parameterised by unconstrained  $\phi = (\boldsymbol{\mu}, L)$  with  $\boldsymbol{\Sigma} = LL^T$ .

## ADVI: ELBO

Assumption:

- Unconstrained model for simpler notation.
- Mean-field Gaussian:  $Q(\boldsymbol{\theta}; \boldsymbol{\phi}) = \text{MvNormal}(\boldsymbol{\theta}; \boldsymbol{\mu}, \text{diag}(\boldsymbol{\sigma}^2))$

$$\text{ELBO}(\boldsymbol{\phi}) = \mathbb{E}_{\boldsymbol{\theta} \sim Q(\cdot; \boldsymbol{\phi})} [\log P(\boldsymbol{\theta}, X)] + \underbrace{\mathbb{E}_{\boldsymbol{\theta} \sim Q(\cdot; \boldsymbol{\phi})} [-\log Q(\boldsymbol{\theta}; \boldsymbol{\phi})]}_{\text{Entropy of } Q}$$

Entropy can be determined analytically

$$\mathbb{H}[Q(\boldsymbol{\theta}; \boldsymbol{\phi})] = \frac{K}{2} \log(2\pi e) + \frac{1}{2} \sum_{k=1}^K \log \sigma_k^2$$

## ADVI: ELBO Gradient: Reparametrisation Trick

$$\begin{aligned}\nabla_{\phi} \mathbb{E}_{\theta \sim Q(\cdot; \phi)} [\log P(\theta, X)] &= \nabla_{\phi} \mathbb{E}_{\theta \sim \text{MvNormal}(\mu, \text{diag}(\exp(\omega)^2))} [\log P(\theta, X)] \\ &= \nabla_{\phi} \mathbb{E}_{\eta \sim \text{MvNormal}(0, I)} [\log P(\exp(\omega) \cdot \eta + \mu, X)] \\ &= \mathbb{E}_{\eta \sim \text{MvNormal}(0, I)} \left[ \underbrace{\nabla_{\phi} \log P(\exp(\omega) \cdot \eta + \mu, X)}_{\text{Compute with AD}} \right]\end{aligned}$$

Putting all together:

For  $\eta_i \sim \text{MvNormal}(0, I)$

$$\nabla_{\mu} \text{ELBO}(\phi) \approx \frac{1}{L} \sum_{i=1}^L \nabla_{\mu} \log P(\exp(\omega) \cdot \eta_i + \mu, X)$$

$$\nabla_{\omega} \text{ELBO}(\phi) \approx \frac{1}{L} \sum_{i=1}^L \nabla_{\omega} \log P(\exp(\omega) \cdot \eta_i + \mu, X) + 1$$

Derivative of entropy

Typically,  $1 \leq L \leq 10$ .

## ADVI: constrained case

Transformation

$$T : \text{support}(P(\theta)) \rightarrow \mathbb{R}^K, \quad \zeta = T(\theta)$$

The transformed density becomes

$$P(\zeta, X) = P(T^{-1}(\zeta)) |\det J_{T^{-1}}(\zeta)|$$

and we fit the Gaussians to the unconstrained density

$$\text{ELBO}(\phi) = \mathbb{E}_{\zeta \sim Q(\cdot; \phi)} [\log P(T^{-1}(\zeta), X) + \log |\det J_{T^{-1}}(\zeta)|] + \underbrace{\mathbb{H}[Q(\zeta; \phi)]}_{\text{Entropy of } Q}$$

```
T = dist.transforms.transform_to(dist_constrained.support)
dist_unconstrained = dist.TransformedDistribution(dist_constrained, T.inv)
```

1. Initialise  $\phi$  (e.g.  $\phi = \mathbf{0}$ )
2. Repeat
  - Draw  $L$  samples  $\boldsymbol{\eta}_i \sim \text{MvNormal}(\mathbf{0}, I)$
  - Approximate gradient  $\nabla_{\phi} \text{ELBO}(\phi)$  with  $\boldsymbol{\eta}_i$ .
  - Update  $\phi$  with gradient (e.g. Adagrad update rule)

Hard math - easier implementation.

See Assignment 4.

# ADVI in Turing (Linear Regression)

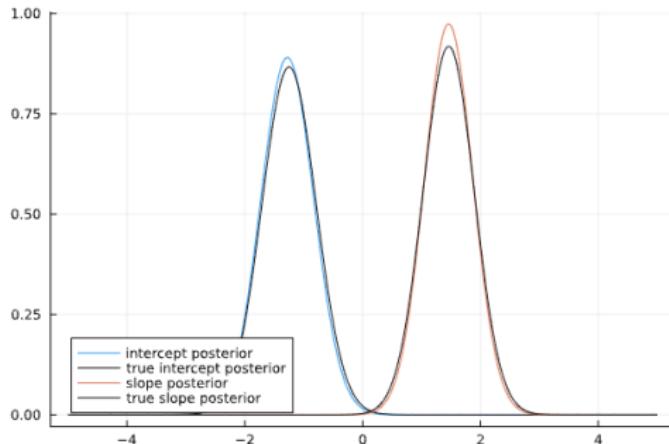
```
model = linear_regression(x,y; broad_prior_params...);

advi = ADVI(25, 10_000); # (samples_per_step, max_iters)
Q = Variational.meanfield(model); # assumption: posterior distribution consists of Normal distributions
result = vi(model, advi, Q; optimizer=Variational.DecayedADAGrad(1e-2, 1.1, 0.9))

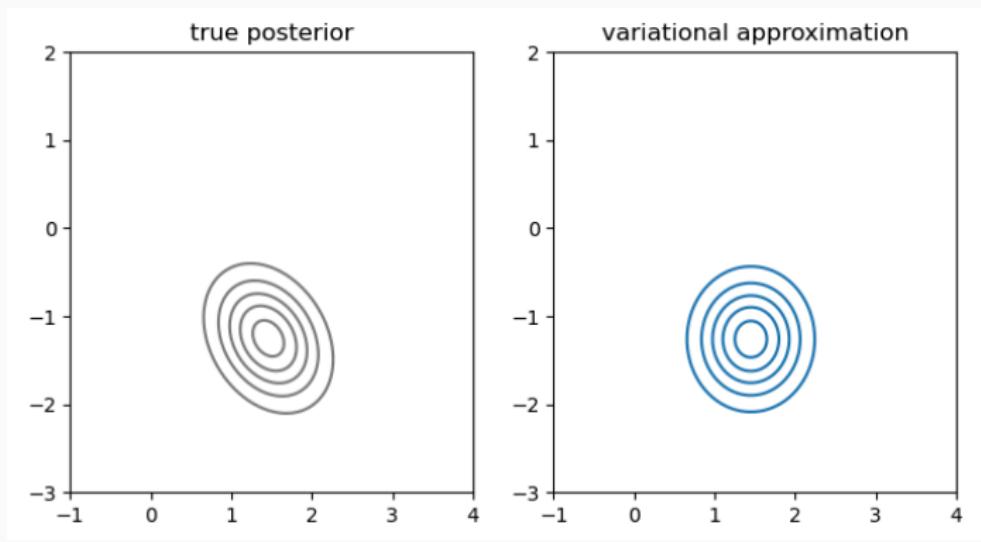
[ADVI] Optimizing... 100% Time: 0:00:01
```

MultivariateTransformed{DistributionsAD.TuringDiagMvNormal{Vector{Float64}, Vector{Float64}}}, Stacked{Vector{Vector{Float64}}} dist: DistributionsAD.TuringDiagMvNormal{Vector{Float64}, Vector{Float64}}}(m=[1.459082954777685, -1.27957085], transform: Stacked([identity, identity], UnitRange{Int64}[1:1, 2:2], UnitRange{Int64}[1:1, 2:2]))

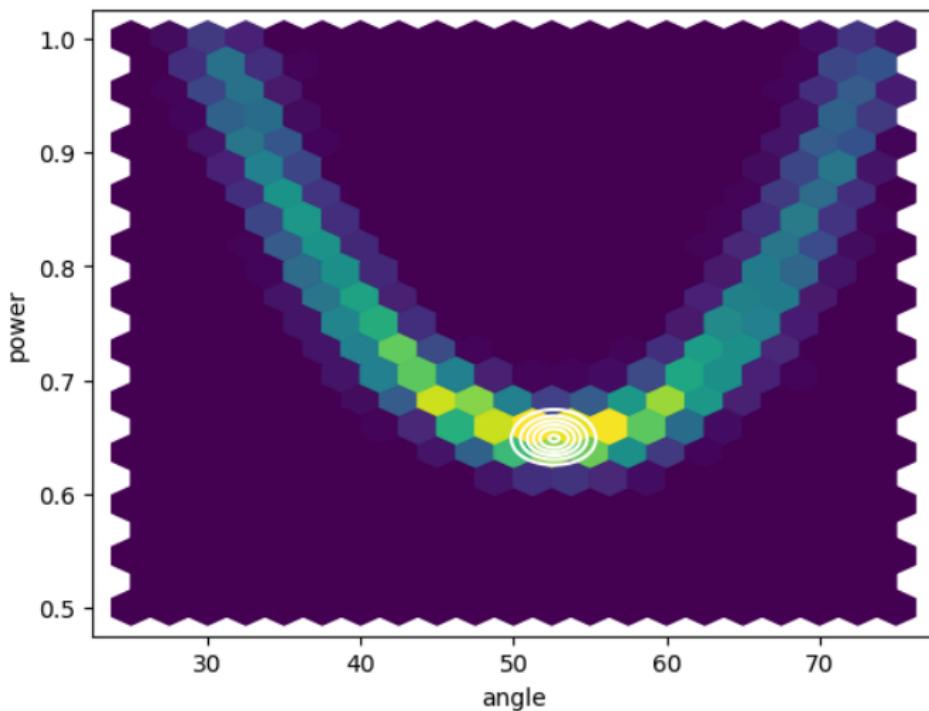
```
plot(x -> pdf(Normal(result.dist.m[2], result.dist.o[2]), x), label="intercept posterior", color=1)
plot!(x -> pdf(true_intercept_posterior, x), label="true intercept posterior", color="black")
plot!(x -> pdf(Normal(result.dist.m[1], result.dist.o[1]), x), label="slope posterior", color=2)
plot!(x -> pdf(true_slope_posterior, x), label="true slope posterior", color="black", legend=:bottomleft)
```



# ADVI in Turing (Linear Regression)



# ADVI in Turing (Golf)



```
https:
//pyro.ai/examples/bayesian_regression_ii.html
https://pyro.ai/examples/vae.html
```

# Stochastic Variational Inference

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# Stochastic Variational Inference

The complexity of ADVI is  $\mathcal{O}(NLK)$ , where  $N$  is the number of data points, since

$$\log P(\Theta, X) = \sum_{i=1}^N \log P(\Theta, X_i).$$

Thus, ADVI as presented, does not scale well to large datasets.

## Subsampling

Suppose we have  $X_j$  are independent and identically distributed (iid). We are interested in the quantity

$$\sum_{j=1}^N f(X_j).$$

Draw an index uniformly at random  $J \sim \text{DiscreteUniform}(1, \dots, N)$ , then

$$\sum_{j=1}^N f(X_j) = N \cdot \mathbb{E}_J [f(X_J)].$$

The right hand side can be approximated with the Monte Carlo method,  $j_m \sim \text{DiscreteUniform}(1, \dots, N)$  iid,

$$\sum_{j=1}^N f(X_j) \approx \frac{N}{M} \sum_{m=1}^M f(X_{j_m}).$$

# Subsampling

Thus, from a large dataset we can subsample a mini-batch by selecting data points at random,  $\{X_{j_1}, \dots, X_{j_M}\}$ , and the sum will be approximately equal to the sum over the entire dataset

$$\sum_{j=1}^N f(X_j) \approx \frac{N}{M} \sum_{m=1}^M f(X_{j_m}).$$

This approximation is more accurate for larger mini-batches.

# Stochastic Variational Inference

In Stochastic Variational Inference, we doubly approximate the gradient of the ELBO:

For  $\eta_i \sim \text{MvNormal}(\mathbf{0}, I)$ ,  $j_m \sim \text{DiscreteUniform}(1, \dots, N)$

$$\begin{aligned}\nabla_{\mu} \text{ELBO}(\phi) &\approx \frac{1}{L} \sum_{i=1}^L \nabla_{\mu} \log P(\exp(\omega) \cdot \eta_i + \mu, X) \\ &\approx \frac{N}{M} \frac{1}{L} \sum_{m=1}^M \sum_{i=1}^L \nabla_{\mu} \log P(\exp(\omega) \cdot \eta_i + \mu, X_{j_m})\end{aligned}$$

Typically, for  $M \approx 100$  large enough,  $L$  can be set to 1.

The complexity of SVI is  $\mathcal{O}(MLK)$ , where  $M \ll N$  is the mini-batch size,  $L$  is the number of samples per iteration, and  $K$  is the number of variables.

# Resources

KL Divergence - Clearly explained!

[https://www.youtube.com/watch?v=9\\_eZHt2qJs4](https://www.youtube.com/watch?v=9_eZHt2qJs4)

Variational Inference + ELBO Intuition

<https://www.youtube.com/watch?v=HxQ94L8n0vU>

Automatic Differentiation and Gradient Descent

<https://medium.com/@rhome/automatic-differentiation-26d5a993692b>

Paper: Automatic Differentiation Variational Inference

<https://arxiv.org/pdf/1603.00788.pdf>

Paper: Stochastic Variational Inference

<https://arxiv.org/pdf/1312.6114.pdf>