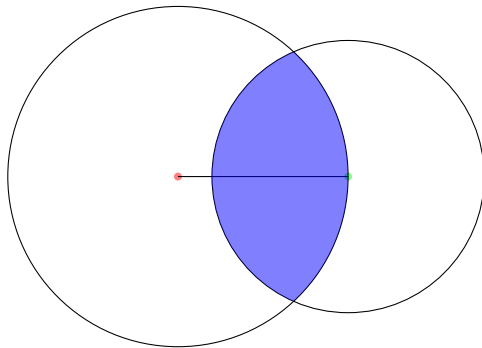


Sojourn Time of Moving Relays in Dual-Hop Cooperative Communication

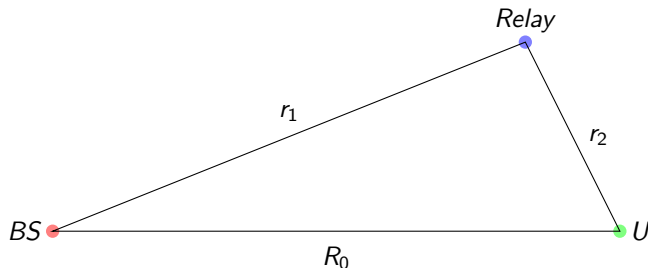
Prudhvi Porandla
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June 20, 2016

Problem



Downlink Cooperation policy



$$\{P_b r_1^{-\gamma} > P_b R_0^{-\gamma}, P_r r_2^{-\gamma} > P_b R_0^{-\gamma}\}$$

P_b and P_r are the transmission powers of base station and relay respectively.

$$\{r_1 < R_0, r_2 < R_2\} \text{ where } R_2 = cR_0, c = \left(\frac{P_r}{P_b}\right)^{\frac{1}{\gamma}}$$

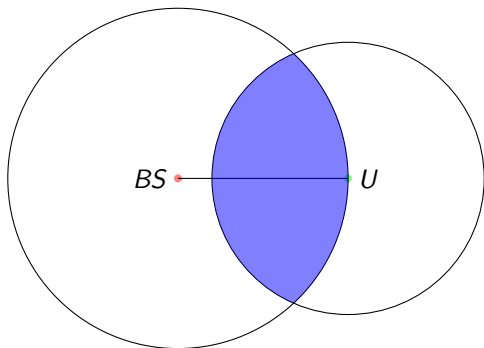


Figure: Feasible region

Motivation

- ▶ Relay selection

Motivation

- ▶ Relay selection
- ▶ Effect on the whole network

Mobility Model

The n th transition of a node is defined by the quadruple

$$(\mathbf{X}_{n-1}, \mathbf{X}_n, V_n, S_n)$$

X_{n-1} and X_n are the starting and destination waypoints.

Velocity - V_n , Pause time or Thinking time at destination - S_n

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Mobility model

- ▶ Distribution of transition length ($L_n = \|\mathbf{X}_{n-1} - \mathbf{X}_n\|$)
- ▶ Distribution of angle made by the vector $\mathbf{X}_n - \mathbf{X}_{n-1}$ w.r.t x-axis.
- ▶ Distributions of V_n and S_n

Rayleigh RWP

Angle is chosen uniformly from $[0, 2\pi]$

Transition length is Rayleigh distributed with parameter λ .

$$P(L > l) = \exp(-\lambda \pi l^2), l \geq 0$$

$$V \equiv v \text{ and } S \equiv 0$$

Rayleigh RWP

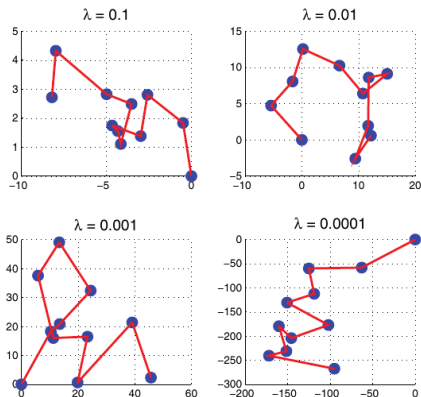


Figure: Sample traces of Rayleigh RWP¹

¹Image Credits: Xingqin Lin et al.

Rayleigh RWP

The mean transition length and time are as follows:

$$E[L] = \frac{1}{2\sqrt{\lambda}}$$

$$E[T] = \frac{1}{2v\sqrt{\lambda}}$$

Existing Literature

- ▶ Sojourn Time and Handover Rate in cellular networks, Xingqin Lin et al.

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Node distribution during one movement period

$$f(r, \theta) = \frac{\sqrt{\lambda}}{\pi r} \exp(-\lambda \pi r^2)$$

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- Analysis of Cell Sojourn Time in Heterogeneous Networks With Small Cells, S. Shin et al.

Expected amount of time a mobile spends in macro-cell-only-area (MoA)

Sojourn Time

If we know the expected number of transitions $E[N]$ in which a node moves out of the region, then sojourn time can be given by

$$S_T = (E[N] - 1)E[T] + E[T_{last}]$$

where T_{last} is the time spent inside the region during the last transition. Since it is difficult to characterize T_{last} ,

$$S_T \approx (E[N] - 1/2)E[T]$$

$E[N]$

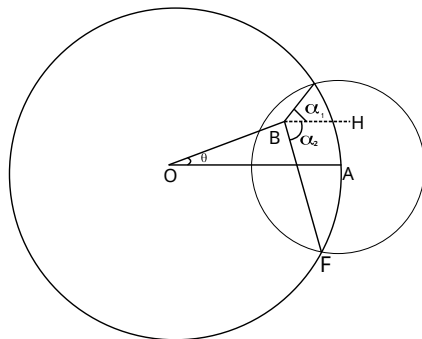
$$E(N) = \sum_{k=1}^{\infty} kPr(r, \theta, k)$$

where

$$Pr(r, \theta, k) = \int_{S-A} \int_A \dots \int_A f_{X_1/X_0}(x_1/x_0) \dots f_{X_k/X_{k-1}}(x_k/x_{k-1}) dA_1 \dots dA_k$$

$f_{X_n/X_{n-1}}(x_n/x_{n-1})$ is the probability density of the destination X_n given that the node's current position is X_{n-1}

Leaving in One Transition



$$\begin{aligned}\rho(r, \theta) &= Pr(-\alpha_2 < \alpha < \alpha_1, r_1 > r_{11}) + Pr(\alpha_1 < \alpha < 2\pi - \alpha_2, r_1 > r_{12}) \\ &= \int_{-\alpha_2}^{\alpha_1} \int_{r_{11}}^{\infty} f_{r_1, \alpha}(r_1, \alpha) dr_1 d\alpha + \int_{\alpha_1}^{2\pi - \alpha_2} \int_{r_{12}}^{\infty} f_{r_1, \alpha}(r_1, \alpha) dr_1 d\alpha\end{aligned}$$

This is a general expression that can be used for any mobility model.

In case of RWP, r_1 and α are chosen independently. Therefore,
 $f_{r_1, \alpha}(r_1, \alpha) = f_{r_1}(r_1)f_{\alpha}(\alpha)$

$$\begin{aligned}\rho(r, \theta) &= \int_{-\alpha_2}^{\alpha_1} f_{\alpha}(\alpha) \int_{r_{11}}^{\infty} f_{r_1}(r_1) dr_1 d\alpha + \int_{\alpha_1}^{2\pi - \alpha_2} f_{\alpha}(\alpha) \int_{r_{12}}^{\infty} f_{r_1}(r_1) dr_1 d\alpha \\ &= \int_{-\alpha_2}^{\alpha_1} \frac{1}{2\pi} e^{-\lambda \pi r_{11}^2} d\alpha + \int_{\alpha_1}^{2\pi - \alpha_2} \frac{1}{2\pi} e^{-\lambda \pi r_{12}^2} d\alpha\end{aligned}$$

Leaving in One Transition

The probability with which a node at (r, θ) moves out of the region of interest during the next transition is

$$\rho(r, \theta) = \int_{-\alpha_2}^{\alpha_1} \frac{1}{2\pi} e^{-\lambda\pi r_{11}^2} d\alpha + \int_{\alpha_1}^{2\pi-\alpha_2} \frac{1}{2\pi} e^{-\lambda\pi r_{12}^2} d\alpha$$

Where

$$r_{11} = -r \cos(\theta - \alpha) + \sqrt{R_0^2 - r^2 \sin^2(\theta - \alpha)}$$

$$r_{12} = R_0 \cos \alpha - r \cos(\theta - \alpha) + \sqrt{[R_0 \cos \alpha - r \cos(\theta - \alpha)]^2 - r^2 \sin^2 \theta - [R_0 - r \cos \theta]^2 + R_0^2}$$

Leaving in One Transition

$$\beta_2 = \cos^{-1} \left(1 - \frac{R_2^2}{2R_0^2} \right) \quad \beta_1 = \beta_2 - \theta$$

$$\alpha_1 = \theta + \tan^{-1} \left(\frac{R_0 \sin \beta_1}{R_0 \cos \beta_1 - r} \right)$$

$$\alpha_2 = \tan^{-1} \left(\frac{r \sin \theta + R_0 \sin \beta_2}{R_0 \cos \beta_2 - r \cos \theta} \right)$$

Simulations

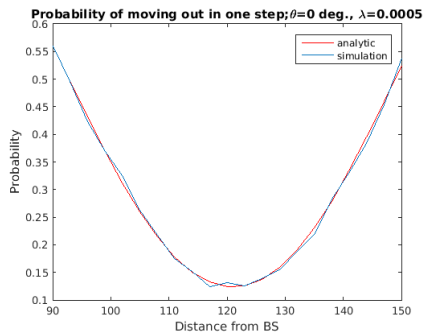
For all simulations, a 1000×1000 square is used to represent the whole 2-D plane. $R_0 = 150$, $R_2 = 0.4R_0$

To draw a length from Rayleigh distribution, the following method is used:

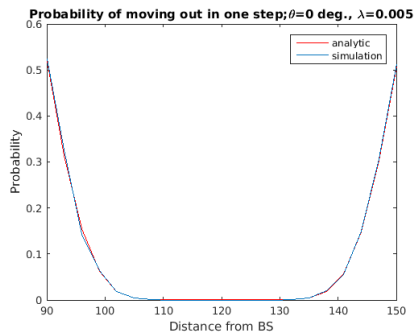
1. Draw a number N from Poisson distribution with density λA where A is the area of the square.
2. Distribute these N points uniformly on the square
3. Of these N points, choose the point that is closest to the point under consideration.

This leads to Rayleigh distribution of transition length and can be proved using null probability of a Poisson Point Process.

Simulations



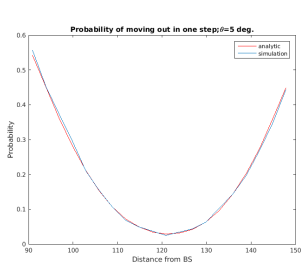
(a) $\lambda = 0.0005$



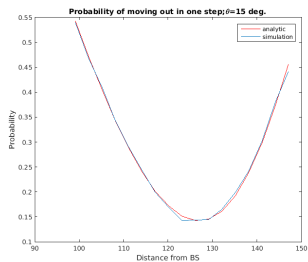
(b) $\lambda = 0.005$

Figure: $\rho(r, 0)$ vs. λ

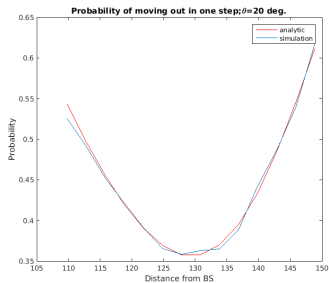
Simulations



(a) $\theta = 5^\circ$



(b) $\theta = 15^\circ$



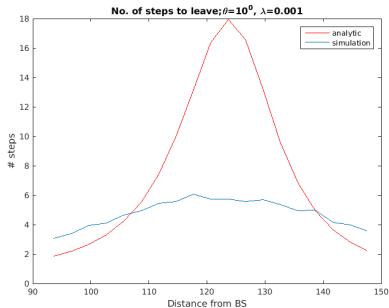
$E[N]$

Let us make a gross approximation and use the same probability of leaving for all waypoints in the path. Then the average number of steps a node starting at (r, θ) takes to leave the region is given by

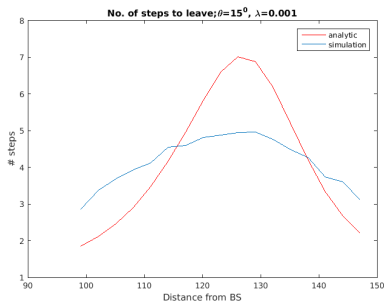
$$\begin{aligned} E[N] &= \sum_{k=1}^{\infty} k(1 - \rho(r, \theta))^{k-1} \rho(r, \theta) \\ &= \frac{1}{\rho(r, \theta)} \end{aligned}$$

$E[N]$

The plots are for points along the radius at angles $\theta = 10^\circ$ and $\theta = 15^\circ$. We can see that the analytical formula agrees better for narrower regions.



(d) $\theta = 10^\circ$



(e) $\theta = 15^\circ$

Figure: Expected number of transitions

Ongoing Work

Markov Chain Process

The transitions in this mobility model have Markovian property

$$f_{X_n/X_{n-1}, X_{n-2}, \dots, X_0}(x_n/x_{n-1}, x_{n-2}, \dots, x_0) = f_{X_n/X_{n-1}}(x_n/x_{n-1})$$

Where X_{n-1} is the current waypoint and X_n is the next waypoint.

Ongoing Work

Discretizing the State Space

- ▶ The state space of the Markov chain is continuous
- ▶ Discretise the space into $n+1$ states
 n states lie inside the region of interest and the $n + 1$ th state represents space outside the region

Ongoing Work

Discretizing the State Space

- ▶ The state space of the Markov chain is continuous
- ▶ Discretise the space into $n+1$ states
 n states lie inside the region of interest and the $n + 1$ th state represents space outside the region
- ▶ Model the motion as an Absorbing Markov Chain
The transition probabilities among first n states depend on the distances between the states and the transition probabilities from these n states to the absorbing state is $\rho(r_i, \theta_i)$

The transition matrix is

$$P = \begin{pmatrix} Q & R \\ \mathbf{0} & 1 \end{pmatrix}$$

$Q_{n \times n}$ is the transition matrix of n non-absorbing states

$R_{n \times 1}$ contains the probabilities of moving out in one step from each of those n states.

Expected number of transitions $\mathbf{t}_{n \times 1}$

$$\mathbf{t} = F\mathbf{1}$$

and the variance vector V is given by

$$V = (2F - I)\mathbf{t} - \mathbf{t}_{sq}$$

where $F = \sum_{k=0}^{\infty} Q^k = (I - Q)^{-1}$

References

- ▶ X. Lin, R. K. Ganti, P. J. Fleming, and J. G. Andrews, Towards understanding the fundamentals of mobility in cellular networks, IEEE Transactions on Wireless Communications, vol. 12, 2013.
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- ▶ H. Elkotby and M. Vu, Interference and throughput analysis of uplink user-assisted relaying in cellular networks, PIMRC, 2014.
- ▶ R. K. Ganti, Stochastic geometry and wireless networks. SPCOM, July 2012.
- ▶ Wikipedia, Absorbing Markov Chain - wikipedia, the free encyclopedia,.