

Homework Assignment 3

Due April 12th by midnight.

SDS 384-11 Theoretical Statistics

1. In this question we consider the Jackknife estimate of variance of a symmetrical measurable function of $n - 1$ variables S . Let $X_1, \dots, X_n - 1$ be i.i.d. Consider $S = S(X_1, \dots, X_{n-1})$. Now let

$$S_i = S(X_1, \dots, X_{i-1}, X_{i+1}, \dots, X_n)$$

So $S = S_n$. If S has finite variance, then the Jackknife estimate of its variance is given by:

$$\text{var}_{JACK}(S) = \sum_i \left(S_i - \frac{\sum_j S_j}{n} \right)^2$$

In Efron and Stein's Annals of Statistics paper in 1981 the following remarkable result was proven.

$$\text{var}(S) \leq E(\text{var}_{JACK}(S)) \tag{1}$$

This is what we will prove here today. First define $V_i = E[S|X_1, \dots, X_i] - E[S|X_1, \dots, X_{i-1}]$.

- (a) Prove that $\text{var}(S) = \sum_{i=1}^{n-1} EV_i^2$
 - (b) Prove that $E\text{var}_{JACK}(S) = (n - 1)E[(S_1 - S_2)^2]/2$
 - (c) Now prove Eq 1.
2. In this question we will look at the Gaussian Lipschitz theorem. Consider $X_1, \dots, X_n \stackrel{iid}{\sim} N(0, 1)$.
- (a) Prove that the order statistics are 1-Lipschitz.
 - (b) Now show that, for large enough n ,

$$c\sqrt{\log n} \leq E[\max_i X_i] \leq \sqrt{2 \log n}$$

where c is some universal constant.

- i. For the upper bound, let $Y = \max_i X_i$. First show that $\exp(tE[Y]) \leq \sum_i E \exp(tX_i)$. Now pick a t to get the right form.
- ii. For the lower bound, do the following steps.
 - A. Show that $E[Y] \geq \delta P(Y \geq \delta) + E[\min(Y, 0)]$
 - B. Now show that $E[\min(Y, 0)] \geq E[\min(X_1, 0)]$
 - C. Finally, relate $P(Y \geq \delta)$ to $P(X_1 \geq \delta)$ by using independence.

- D. Now show that $P(X_1 \geq \delta) \geq \exp(-\delta^2/\sigma^2)/c$, for some universal constant c .
- E. Choose the parameter δ carefully to have $P(X_1 \geq \delta) \geq 1/n$, for large enough n .
3. Let \mathcal{P} be the set of all distributions on the real line with finite first moment. Show that there does not exist a function $f(x)$ such that $Ef(X) = \mu^2$ for all $P \in \mathcal{P}$ where μ is the mean of P , and X is a random variable with distribution P .
4. Let g_1 and g_2 be estimable parameters within \mathcal{P} with respective degrees m_1 and m_2 .
- (a) Show $g_1 + g_2$ is an estimable parameter with degree $\leq \max(m_1, m_2)$.
- (b) Show $g_1 g_2$ is an estimable parameter with degree at most $m_1 + m_2$.
5. A continuous distribution with CDF $F(x)$, on the real line is symmetric about the origin if, and only if, $1 - F(x) = F(-x)$ for all real x . This suggests using the parameter,

$$\theta(F) = \int (1 - F(x) - F(-x))^2 dF(x) \quad (2)$$

$$= \int ((1 - F(-x))^2 dF(x) - 2 \int (1 - F(-x))F(x) dF(x) + \int F(x)^2 dF(x) \quad (3)$$

as a nonparametric measure of how asymmetric the distribution is. Find a kernel h , of degree 3, such that $E_F h(X_1, X_2, X_3) = \theta(F)$ for all continuous F . Find the corresponding U statistic.