

SDS 321: Introduction to Probability and Statistics Lecture 17: Continuous random variables: conditional PDF

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Roadmap

- Two random variables: joint distributions
 - ► Joint pdf
 - ► Joint pdf to a single pdf: Marginalization
 - Conditional pdf
 - Conditioning on an event
 - Conditioning on a continuous r.v
 - ► Total probability rule for continuous r.v's
 - ▶ Bayes theorem for continuous r.v's
 - Conditional expectation and total expectation theorem
 - Independence
- More than two random variables.

Conditional PDFs-conditioning on an event

- ▶ For discrete random variables, we looked at marginal PMFs $p_X(X)$, conditional PMFs $p_{X|Y}(x|y)$, and joint PMFs $p_{X,Y}(x,y)$.
- ▶ These corresponded to the probability of an event, P(A), the conditional probability of an event given some other event, P(A|B), and probability of the intersection of two events, $P(A \cap B)$.
- ▶ We've looked at marginal PDFs, $f_X(x)$ and joint PDFs, $f_{X,Y}(x,y)$.
- ▶ These don't directly give us probabilities of events, but we can use them to calculate such probabilities by integration.
- ▶ We can also look at conditional PDFs! These allow us to calculate the probability of events given extra information.

Conditional PDFs

▶ Recall, the PDF of a continuous random variable X is the non-negative function $f_X(x)$ that satisfies

$$P(X \in B) = \int_B f_X(x) dx$$

for any subset B of the real line.

- ▶ Let A be some event with P(A) > 0
- ▶ The **conditional PDF** of X, given A, is the non-negative function $f_{X|A}$ that satisfies

$$P(X \in B|X \in A) = \int_{B} f_{X|A}(x) dx$$

for any subset B of the real line.

▶ If B is the entire line, then we have

$$\int_{-\infty}^{\infty} f_{X|A}(x) dx = 1$$

▶ So, $f_{X|A}(x)$ is a valid PDF.

4

Conditional PDFs

- The event we are conditioning on can also correspond to a range of values of our continuous random variable.
- Definition-

$$f_{X|\{X\in A\}}(x) = \begin{cases} \frac{f_X(x)}{P(X\in A)} & \text{if } X\in A\\ 0 & \text{otherwise.} \end{cases}$$

In this case, we can write the conditional probability as

$$\int_{B} f_{X|A}(x)dx = \int_{B} \frac{f_{X}(x)1(x \in A)}{P(X \in A)} dx$$

$$= \frac{\int_{A \cap B} f_{X}(x)dx}{P(X \in A)} = \frac{P(\{X \in A\} \cap \{X \in B\})}{P(X \in A)}$$

$$= P(X \in B|X \in A)$$

▶ This is a valid PDF—non-negative and integrates to one. Check?

- $ightharpoonup X \sim Exp(\lambda)$
- $f_X(x) = \lambda e^{-\lambda x}$ when $x \ge 0$, and zero otherwise.
- ▶ P(X > s + t | X > s) = ?

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- P(X > s + t | X > s) = ?
- ▶ Remember the exponential? $F_X(x) = 1 e^{-\lambda x}$.

$$P(X > s + t | X > s) = \frac{P(X > s + t, X > s)}{P(X > s)}$$

$$= \frac{P(X > s + t)}{P(X > s)} = \frac{e^{-\lambda(s+t)}}{e^{-\lambda s}}$$

$$= e^{-\lambda t} = P(X > t)$$

 $ightharpoonup X \sim Exp(\lambda)$

$$f_{X|X>s}(x) = \begin{cases} \frac{\lambda e^{-\lambda x}}{P(X>s)} = \lambda e^{\lambda(x-s)} & \text{if } x>s\\ 0 & \text{Otherwise} \end{cases}$$

▶
$$P(X > s + t | X > s) = ?$$

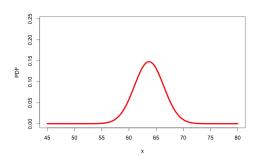
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- ▶ P(X > s + t | X > s) = ?
- ▶ Remember the exponential? $F_X(x) = 1 e^{-\lambda x}$.

$$P(X > s + t | X > s) = \int_{s+t}^{\infty} f_{X|X>s}(x) dx = \lambda \int_{s+t}^{\infty} e^{-\lambda(x-s)} dx$$

$$= \lambda \int_{t}^{\infty} e^{-\lambda u} du = e^{-\lambda t}$$

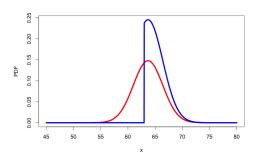
Conditional PDFs: Example

- ► The height X of a randomly picked american woman can be modeled by $X \sim N(63.7, 2.7^2)$
- ▶ Whats the conditional PDF given that the randomly picked woman is at least 63 inches tall?
- ▶ The PDF of heights (X) is shown in red.



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- ▶ Whats the conditional PDF given that the randomly picked woman is at least 63 inches tall?
- ▶ The PDF of heights (*X*) is shown in red.
- ▶ The conditional PDF given X > 63, shown in blue, is the same shape for X > 63... but scaled up to integrate to one.



Conditioning on a different random variable

So far, we conditioned X on an arbitrary event A, or on a range of values of X.

$$P(X \in B|A) = \int_B f_{X|A}(x) dx$$

- We can also condition on the outcome of a second random variable Y.
- ▶ We know we could condition on a range of outcomes of Y, by replacing the arbitrary event A with the event $\{Y \in A\}$

$$P(X \in B | Y \in A) = \int_{B} f_{X|\{Y \in A\}}(x) dx$$

- ▶ What about conditioning on a specific value of Y = y?
- ▶ Even though any outcome Y = y has P(Y = y) = 0, we know that some value has to happen.
 - ▶ Pick some number, say 0.6777, now generate 100 N(0,1) random variables. I will bet a 100\$ that you won't see that number.
 - But when you simulate from the standard normal, you will get a 100 different values, right?

Conditioning on a different random variable

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_{Y}(y)},$$

provided $f_Y(y) > 0$.

What does this mean?

$$f_{X|Y}(x|y)dx = \frac{f(x,y)dxdy}{f(y)dy}$$

$$= \frac{P(x \le X \le x + dx, y \le Y \le y + dy)}{P(y \le Y \le y + dy)}$$

$$= P(x \le X \le x + dx|y \le Y \le y + dy)$$

Multiplication rule: Calculating the joint PDF

- ▶ We can use the same relationship, $f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_{Y}(y)}$, to calculate the joint PDF from the conditional and the marginal PDF.
- i.e., $f_{X,Y}(x,y) = f_{X|Y}(x|y)f_Y(y)$.
- This is a PDF version of our multiplication rule.
- ▶ We can extend it to more than 2 random variables:

$$f_{X,Y,Z}(x,y,z) = f_{Z|X,Y}(z|x,y)f_{Y|X}(y|x)f_{X}(x)$$

We've now got a lot of ways to go between our various PDFs!

- ▶ If we know $f_{X,Y}(x,y)$, we can get $f_{X}(x)$
 - ► How?
- ▶ If we know $f_{X,Y}(x,y)$ and $f_Y(y)$, if $f_Y(y) > 0$ we can get $f_{X|Y}(x|y)$

▶ If we know $f_X(x)$ and $f_{Y|X}(y|x)$, we can get $f_{X,Y}(x,y)$

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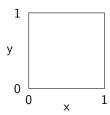
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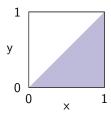
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- ▶ If we know $f_X(x)$ and $f_{Y|X}(y|x)$, we can get $f_{X,Y}(x,y)$
 - ▶ How? multiplication rule! $f_{X,Y}(x,y) = f_X(x)f_{Y|X}(y|x)$

- Let $f_{X,Y}(x,y) = \begin{cases} c & \text{if } 0 \le x \le 1 \text{ and } 0 \le y \le x \\ 0 & \text{otherwise} \end{cases}$
- ▶ What is the conditional PDF of X given Y, $f_{X|Y}(x|y)$?

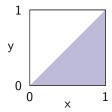
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- ► First things first... what is c? Well, what does our joint PDF look like?



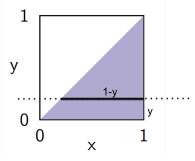
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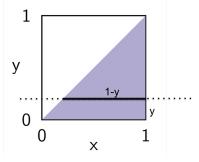
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- ▶ The total area where $0 \le x \le 1$ and $0 \le y \le x$ is 0.5, so c = 2.
- ▶ What is the marginal PDF of Y, $f_Y(y)$?

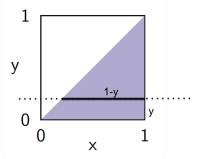


- ➤ To get the marginal PDF of Y, we take the joint PDF and marginalize out X.
- $f_{Y}(y) = \int_{0}^{1} f_{X,Y}(x,y) dx = 2 \int_{0}^{1} \mathbf{1}_{0 \le x \le 1, 0 \le y \le x} dx$



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$$= 2 \int_{x=y}^{1} dx = 2(1-y)$$



► To get the marginal PDF of *Y*, we take the joint PDF and marginalize out *X*.

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$$= 2 \int_{x=y}^{1} dx = 2(1-y)$$

So, the conditional PDF of X given Y = y is

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_{Y}(y)} = \begin{cases} \frac{1}{1-y} & \text{if } y \le x \le 1\\ 0 & \text{otherwise.} \end{cases}$$

Total probability theorem for continuous random variables

- ▶ We know that conditional probabilities must obey the total probability theorem.
- ▶ If $B_1, ..., B_n$ form a partition of Ω, such that $P(B_i) > 0$ for each i, then for any event A,

$$P(A) = \sum_{i=1}^{n} P(B_i)P(A|B_i)$$

▶ In terms of discrete r.v's we have:

$$P(X = x) = \sum_{i} P(X = x|B_i)P(B_i)$$

▶ How about continuous r.v.'s? Replace $P(X = x|B_i)$ by conditional pdf.

$$f_X(x) = \sum_i f_{X|B_i}(x) P(B_i)$$

- ▶ Sometimes our hidden cause is inherently discrete.
 - e.g. I may be interested in whether I have flu or not a binary choice
 - My observation might be my temperature a continuous random variable.
- We want P(A|Y = y) = e.g. P(flu|Y = 100)
- Pretend Y is a discrete r.v.

$$P(A|Y = y) = \frac{P(Y = y|A)P(A)}{P(Y = y|A)P(A) + P(Y = y|A^{c})P(A^{c})}$$

All that changes for a continuous r.v. is:

$$P(A|Y = y) = \frac{f_{Y|A}(y)P(A)}{f_{Y|A}(y)P(A) + f_{Y|A}(y)P(A^{c})}$$

- ▶ The probability that anyone has flu (event A) is 20%.
- ▶ Body temperature is *Y*.
- ▶ Without flu, Y is a normal random variable with $\mu = 98.6$ degrees and $\sigma = .5$.
- ▶ With flu, Y is a normal random variable with $\mu = 102$ and $\sigma = 2$.
- ▶ My temperature is 100. If A is the event "has flu" and Y is temp.

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$$f_{Y|A}(y) = \frac{1}{\sqrt{2\pi \times 4}} \exp{-\frac{(y - 102)^2}{2 \times 4}}$$
$$f_{Y|A}(y) = \frac{1}{\sqrt{2\pi \times .25}} \exp{-\frac{(y - 98.6)^2}{2 \times .25}}$$

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$$P(A|Y=y) = \frac{P(A)f_{Y|A}(y)}{f_{Y}(y)} = \frac{f_{Y|A}(y)P(A)}{f_{Y|A}(y)P(A) + f_{Y|Ac}(y)P(A^{c})}$$

$$P(A|Y=100) = \frac{0.2\frac{1}{2\sqrt{2\pi}}e^{-(100-102)^{2}/8}}{0.2\frac{1}{2\sqrt{2\pi}}e^{-(100-102)^{2}/8} + 0.8\frac{1}{0.5\sqrt{2\pi}}e^{-(100-98.6)^{2}/0.5}} = 0.65$$

Continuous Bayes' rule

▶ Discrete *X*, *Y*.

$$P(X = x | Y = y) = \frac{P(Y = y | X = x)P(X = x)}{\sum_{X} P(Y = y | X = x)P(X = x)}$$

▶ What is $f_{X|Y}(x|y)$?

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$$f_{X|Y}(x|y) = \frac{f_{Y|X}(y|x)f_X(x)}{\int_{-\infty}^{\infty} f_{Y|X}(y|x)f_X(x)dx}$$

Conditional Expectation

- When we were looking at discrete random variables, we looked at conditional expectations.
- ▶ The conditional expectation, E[X|A], of a random variable X given an event A is the value of X we expect to get out, on average, when A is true.
- ▶ We could calculate it by summing over all values x that X can take on, and scaling them by the conditional PMF $p_{X|A}(x) = P(X = x|A)$.

$$E[X|A] = \sum_{x} x p_{X|A}(x)$$

Conditional Expectation

- We can also look at the conditional expectation of a continuous random variable.
- ▶ If $E[X] = \int_{-\infty}^{\infty} x f_X(x) dx$, what do you think the conditional expectation of X given some event A looks like?

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- ▶ How about the conditional expectation of some function g(X) given some event A?
- $E[g(X)|A] = \int_{-\infty}^{\infty} g(x) f_{X|A}(x) dx$

More generally, if $A_1, A_2, ..., A_n$ are a partition of Ω , we have a continuous version of the **total expectation theorem**:

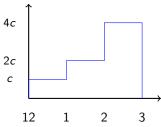
$$E[X] = \sum_{i=1}^{n} P(A_i)E[X|A_i]$$

ightharpoonup Or, if we are conditioning on specific values Y = y,

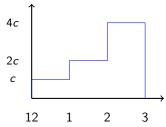
$$E[X] = \int_{-\infty}^{\infty} E[X|Y = y] f_{Y}(y) dy$$

- ▶ I am expecting an email, that will definitely arrive between midday and 3pm.
- ▶ Within a given hour (midday-1, 1-2, 2-3), each time is equally likely.
- ▶ It is twice as likely to arrive between 1 and 2 as it is to arrive between midday and 1.
- ▶ It is twice as likely to arrive between 2 and 3 as it is to arrive between 1 and 2.
- ▶ What does the PDF look like?

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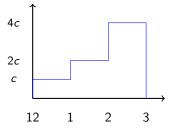


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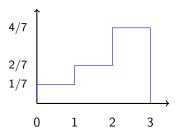


▶ What is *c*?

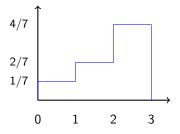
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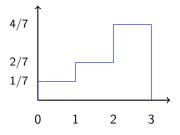
▶ What is c? 1/7



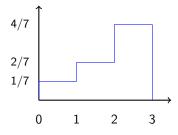
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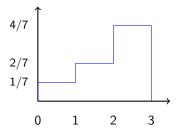


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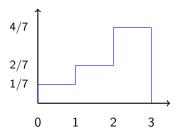
$$f_{X|X>2}(x) = \begin{cases} 1 & \text{if } 2 < x \le 3\\ 0 & \text{otherwise} \end{cases}$$



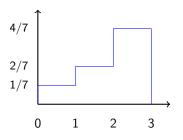
- I wait until 2pm. It still hasn't arrived. What is the expected value of the arrival time?
- ▶ What is the expected time without any conditioning?
- ▶ First, what is the conditional probability, $f_{X|X>2}(x)$?

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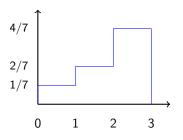
► So,
$$E[X|X>2] = \int_{-\infty}^{\infty} x f_{X|X>2}(x) dx = \int_{2}^{3} x dx = 2.5.$$



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- $P(X > 2) = \int_{2}^{3} f_{X}(x) dx = 4/7$
- ► Similarly, $P(X < 1) = \int_0^1 f_X(x) dx = 1/7$ and $P(1 \le X \le 2) = 2/7$.

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- ▶ By the total probability theorem,

$$f_X(x) = P(X \le 1) f_{X|0 \le X \le 1}(x)$$

+ $P(1 \le X \le 2) f_{X|1 \le X \le 2}(x) + P(X > 2) f_{X|X > 2}(x)$

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So, we can write the total expectation as

$$\begin{split} E[X] &= \int_0^1 x P(X \le 1) f_{X|X \le 1}(x) + \int_1^2 x P(1 \le X \le 2) f_{X|1 \le X \le 2}(x) \\ &+ \int_2^3 x P(X > 2) f_{X|X > 2}(x) \\ &= E[X|X \le 1] P(X \le 1) + E[X|1 \le X \le 2] P(1 \le X \le 2) \\ &+ E[X|X > 2] P(X > 2) \\ &= 0.5 \cdot 1/7 + 1.5 \cdot 2/7 + 2.5 \cdot 4/7 = 27/14 \end{split}$$

- ▶ John's tank holds 15 gallons of gas, and he always refills his tank when he gets down to 5 gallons.
- John's car gets 30MPG on average, with a standard deviation of 2MPG.
- ▶ I plan on borrowing John's car tomorrow. I don't know how much gas he will have. How far should I expect to be able to drive it?

▶ I want *E*[*M*].

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- ▶ Let's set up some reasonable modeling assumptions.
- ▶ Let *G* be the random volume of gas. Assume

$$f_G(g) = egin{cases} 0.1 & ext{if } 5 < g \leq 15 \ 0 & ext{otherwise}. \end{cases}$$

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- ▶ If we have exactly g gallons, what is E[M|G = g]? 30g
- ▶ So, we can use the total expectation theorem to get:

$$E[M] = \int_{-\infty}^{\infty} E[M|G = g] f_G(g) dg = \int_{5}^{15} 30g \times 0.1 \, dg = [1.5g^2]_{5}^{15} = 300$$

Independent random variables

► For discrete random variables, we said two random variables *X* and *Y* are independent if

$$p_{X,Y}(x,y) = p_X(x)p_Y(y) \quad \forall x, y$$

 Just like in the discrete case, we say two continous random variables are independent if

$$f_{X,Y}(x,y) = f_X(x)f_Y(y) \quad \forall x, y$$

- ▶ If $f_Y(y) > 0$, this is the same as saying $f_X(x) = f_{X|Y}(x|y) i.e.$ knowing that Y = y doesn't tell us anything about X.
- ▶ Just like with discrete random variables, we if X and Y are independent we have E[XY] = E[X]E[Y] and var(X + Y) = var(X) + var(Y).
 - For two functions f(X) and g(Y) we have E[f(X)g(Y)] = E[f(X)]E[g(Y)].

▶ For multiple random variables we have:

$$P((X,Y,Z) \in B) = \int_{(x,y,z) \in B)} f_{X,Y,Z}(x,y,z) dxdydz$$

• Marginalization: $f_{X,Y}(x,y) =$

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- ► Conditional PDF: $f_{X,Y|Z}(x,y|z) = \frac{f_{X,Y,Z}(x,y,z)}{f_{Z}(z)}$, For $f_{Z}(z) > 0$

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- ► Multiplication rule:

$$f_{X,Y,Z}(x,y,z) = f_{X|Y,Z}(x|y,z)f_{Y|Z}(y|z)f_{Z}(z), \text{ For } f_{Y,Z}(y,z) > 0$$

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$$f_{X,Y,Z}(x,y,z) = f_{X|Y,Z}(x|y,z)f_{Y|Z}(y|z)f_{Z}(z)$$
, For $f_{Y,Z}(y,z) > 0$
• Independence: $f_{X,Y,Z}(x,y,z) = f_{X}(x)f_{Y}(y)f_{Z}(z)$ For all x,y,z

For two random variables X, Y arising out of the same experiment, we define their CDF as:

$$F_{X,Y}(x,y) = P(X \le x, Y \le y) =$$

- ► How do I get $f_{X,Y}(x,y)$ back? $f_{X,Y}(x,y) = \frac{d^2 F_{X,Y}(x,y)}{dx \ dy}$
- ▶ Let X and Y be jointly uniform on the unit square. $F_{X,Y}(x,y) = xy$ for $0 \le x, y \le 1$
- ▶ What is $f_{X,Y}(x,y)$?. Differentiate! $\frac{d}{dx} \left(\frac{d}{dy}(xy) \right)$
- ▶ This equals 1 for all $0 \le x, y \le 1$!

For two random variables X, Y arising out of the same experiment, we define their CDF as:

$$F_{X,Y}(x,y) = P(X \le x, Y \le y) = \int_{-\infty}^{x} \int_{-\infty}^{y} f_{X,Y}(u,v) du \ dv$$

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Practice problem

- Let $Y = g(X) = X^2$. X is a random variable with a known PDF $f_X(x)$. Whats the PDF of Y?
- ▶ Solution: See example 3.23 of Bertsekas and Tsitsiklis.

Homework, Review, Midterm

- ▶ Next lecture we will work through some extra practice problems.
- ▶ Anything you particularly want to focus on? Email me by this evening and I'll try to also review that in next class.

Some notes about the exam:

- It will be in class a week on Tuesday please be prompt! and last 1hr 15 minutes.
- ▶ You can bring a sheet of paper with notes, formula etc.
- Any calculator is fine.