

# Homework Assignment 1

## Due in class, Wednesday Feb 7th

SDS 384-11 Theoretical Statistics

- Given densities  $p_n$  and  $q_n$  with respect to some measure  $\mu$ , let  $X$  be distributed according to the distribution with density  $p_n$ . Define the likelihood ratio  $L_n(X)$  as  $L_n(X) = q_n(X)/p_n(X)$  for  $p_n(X) > 0$ .  $L_n(X) = 1$  if  $p_n(X) = q_n(X) = 0$  and  $L_n(X) = \infty$  otherwise. Show that the likelihood ratio is a uniformly tight sequence.  $E[|L_n(X)|] = E[L_n(X)] = \int_{x:p_n(x)>0} \frac{q_n(x)}{p_n(x)} p_n(x) dx \leq 1$ . So for  $\epsilon > 0$ , take  $P(L_n \geq 1/\epsilon) \leq \epsilon$  for all  $n$ . So its UT.
- Consider a sequence of iid random variables  $\{X_n\}$  such that  $X_i \sim \text{Beta}(\theta, 1)$ , where  $\theta > 0$ . Let  $\bar{X}_n$  denote the sample mean. The method of moments estimator of  $\theta$  is  $\hat{\theta}_n = \bar{X}_n/(1 - \bar{X}_n)$ . Derive the asymptotic distribution of  $\sqrt{n}(\hat{\theta}_n - \theta)$ . Recall that the expectation of a  $\text{beta}(\beta, 1)$  random variable is  $\theta/(1 + \theta)$ . So  $\bar{X}_n \xrightarrow{P} \theta/(1 + \theta)$  and variance  $\sigma^2 = \frac{\theta}{(\theta+1)^2(\theta+2)}$ . Now

$$\begin{aligned} \sqrt{n}(\hat{\theta}_n - \theta) &= \sqrt{n} \left( \frac{\bar{X}_n}{1 - \bar{X}_n} - \theta \right) \\ &= \sqrt{n}(1 + \theta) \frac{\bar{X}_n - \frac{\theta}{1+\theta}}{1 - \bar{X}_n} \end{aligned}$$

Using CLT we have  $\sqrt{n}(\bar{X}_n - \frac{\theta}{1+\theta}) \xrightarrow{d} N(0, \sigma^2)$ .  $1 - \bar{X}_n \xrightarrow{P} 1/(1 + \theta)$ . So  $\sqrt{n}(\hat{\theta}_n - \theta) \xrightarrow{d} N(0, \theta(\theta + 1)^2/(\theta + 2))$ .

- Derive the following one sided improvement of Chebyshev's inequality for a random variable  $X$  with variance  $\sigma^2$  and any  $t > 0$

$$P(X - E[X] \geq t) \leq \frac{\sigma^2}{\sigma^2 + t^2} \quad (1)$$

Take  $E[X] = 0$  WLOG.

$$P(X \geq t) = P((X + u)^2 \geq (t + u)^2) \leq \frac{\sigma^2 + u^2}{(t + u)^2}$$

Minimizing the RHS w.r.t  $u \geq 0$  gives the answer.

- If  $X_n \xrightarrow{d} X \sim \text{Poisson}(\lambda)$ , is it necessarily true that  $E[g(X_n)] \rightarrow E[g(X)]$ ?  
 (a)  $g(x) = 1(x \in (0, 10))$   $g(x)$  is discontinuous at 0 and 10, and  $P(X \in \{0, 10\}) \neq 0$ . So our theorem does not apply. We will create a counter-example. Let  $X_n = X + 1/n$ .  $E[g(X_n)] = P(X \in (-1/n, 10 - 1/n)) \rightarrow P(X \in [0, 9])$ . On the other hand  $E[g(X)] = P(X \in [1, 9])$ .

- (b)  $g(x) = e^{-x^2}$  is continuous and bounded, so the Portmanteau theorem proves that the expectation converges.
- (c)  $g(x) = \text{sgn}(\cos(x))$  [ $\text{sgn}(x) = 1$  if  $x > 0$ ,  $-1$  if  $x < 0$  and  $0$  if  $x = 0$ .]  $g(x)$  is discontinuous at  $(\pi/2, 3\pi/2, \dots)$ . However since  $X$  takes values in integers, we can safely say that  $E[g(X_n)] \rightarrow E[g(X)]$
- (d)  $g(x) = x$   $g(x)$  is continuous but unbounded. So let's find a counter example. Let  $X_n = X$  with probability  $1 - 1/n$  and  $X_n = n$  with probability  $1/n$ . So  $X_n \xrightarrow{P} X$  and  $X_n \xrightarrow{d} X$ . But  $E[g(X_n)] = \lambda(1 - 1/n) + 1 \rightarrow \lambda + 1$ , but  $E[g(X)] = \lambda$ .
5. Consider  $n$  i.i.d random variables  $\{X_n\}$  uniformly distributed on the set of  $n$  points  $\{1/n, 2/n, \dots, 1\}$ . Show that  $X_n \xrightarrow{d} X$  where  $X \sim \text{Uniform}(0, 1)$ . Does  $X_n \xrightarrow{P} X$ ?

WLOG let  $x \in [i/n, (i+1)/n]$ .  $P(X_n \leq x) = i/n \rightarrow x$  as  $n \rightarrow \infty$ . So  $X_n$  is converging in distribution to a Uniform r.v. Remember that convergence in probability also implies convergence in the Cauchy sense, i.e.  $P(|X_m - X_n| \geq \epsilon) \leq P(|X_m - X| + |X_n - X| \geq \epsilon) \leq P(|X_m - X| \geq \epsilon/2) + P(|X_n - X| \geq \epsilon/2) \rightarrow 0$  for  $m, n \rightarrow \infty$  if  $X_n \xrightarrow{P} X$ .

In this case however,  $P(|X_m - X_n| \geq \epsilon) \not\rightarrow 0$  since  $X_m, X_n$  are independent.

So it is clear, that the sequence is non convergent in a Cauchy sense, since the sequence is iid. Another way to think about this is to look at a partial converse of convergence in probability. This states that a sequence converges in probability iff for every sequence, there is a subsequence that converges a.s. In our case, there is no such subsequence, and hence the sequence does not converge in probability.