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SDS 384 11: Theoretical Statistics

Lecture 18: Bootstrap

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The bootstrap

- Data $X_1, \ldots, X_n \overset{\text{iid}}{\sim} P$
- Some estimator $\hat{\theta}$ of parameter of interest θ.
- Want:
$$P\left(\hat{\theta} - \kappa_\alpha \hat{\sigma} \leq \theta \leq \hat{\theta} + \kappa_{1-\alpha}\hat{\sigma}\right) \geq 1 - 2\alpha,$$
where $\kappa_\alpha, \kappa_{1-\alpha}$ are the quantiles of $(\hat{\theta} - \theta)/\hat{\sigma}$
- The distribution of $(\hat{\theta} - \theta)/\hat{\sigma}$ depends on P.
- Often this distribution is normal, but with unknown parameters.

Bootstrap: plug in principle

True model	Bootstrapped model
$\hat{\theta}$	$\hat{\theta}^*$
$\hat{\sigma}$	σ
$\frac{\hat{\theta}-\theta}{\hat{\sigma}}$	$\frac{\hat{\theta}^*-\hat{\theta}}{\hat{\sigma}^*}$

Empirical bootstrap

How do you estimate P?

Empirical Bootstrap $\hat{P} = \hat{P}_n$

Generate m samples $(X_1^*, \ldots, X_n^*)^{(j)}$, $j = 1 : m$.
Each giving a $(\hat{\theta}^*, \hat{\sigma}^*)$ pair.
Compute the κ_α quantile
of the distribution of $\frac{\hat{\theta}^* - \hat{\theta}}{\hat{\sigma}^*}$

Parametric bootstrap $\hat{P} = P_{\hat{\theta}}$

Consistency

- We want, as $n \to \infty$,

$$\sup_x \left| P\left(\frac{\hat{\theta} - \theta}{\hat{\sigma}} \leq x\right) - \hat{P}_n\left(\frac{\hat{\theta}^* - \hat{\theta}}{\hat{\sigma}^*} \leq x\right) \right| \xrightarrow{P} 0$$

- We assume:

$$P\left(\frac{\hat{\theta} - \theta}{\hat{\sigma}} \leq x\right) \to F(x)$$

- F is continuous, and so it is enough to show that:

$$\hat{P}_n\left(\frac{\hat{\theta}^* - \hat{\theta}}{\hat{\sigma}^*} \leq x\right) \xrightarrow{P} F(x).$$

Consistency

Theorem

If $X_1, \ldots, X_n$ are iid with mean μ and variance σ^2, then conditioned on the data, for almost every sequence X_1^n

$$\sqrt{n}(\bar{X}_n^* - \bar{X}_n) \xrightarrow{d} N(0, \sigma^2)$$

Proof

- First note that $E[\bar{X}_n^*|X_1^n] = E[X_1^*|X_1^n] = \bar{X}_n$
- Now note that $\text{var}(\bar{X}_n^*|X_1^n) = E[(X_1 - \bar{X}_n)^2|X_1^n] = \hat{\sigma}^2$
- Can we use CLT?
- The $\hat{P}_n$ is changing for each n.
- Need to check the Lindeberg condition.

Lindeberg-feller CLT for triangular arrays

$$X_{11}$$
$$X_{21}, X_{22}$$
$$X_{21}, X_{22}, X_{23}$$
$$\ldots$$

Theorem

For each n *let* $(X_{ni})_{i=1}^{n}$ *be independent random variables with mean zero and variance* σ_{ni}^2*. Let* $Z_n = \sum_{i=1}^{n} X_{ni}$ *and* $B_n^2 = var(Z_n)$*. Then* $Z_n/B_n \xrightarrow{d} N(0,1)$*, as long as the* **Lindeberg condition** *holds.*

The Lindeberg condition

Definition (Lindeberg condition)

For every $\epsilon > 0$,

$$\frac{1}{B_n^2} \sum_{j=1}^{n} E[X_{nj}^2 1(|X_{nj}| \geq \epsilon B_n)] \to 0 \text{ as } n \to \infty \tag{1}$$

Converse: If $\frac{\sigma_{nj}^2}{B_n^2} \to 0$ as $n \to \infty$, i.e. no one variance plays a significant role in the limit, and if $Z_n/B_n \xrightarrow{d} N(0,1)$, then the Lindeberg condition holds.

Necessary and Sufficient: If $\frac{\sigma_{nj}^2}{B_n^2} \to 0$, the the Lindeberg condition is necessary and sufficient to show the CLT.

Does the Lindeberg condition hold?

- Check if $E[(X_i^*)^2 1(|X_i^*| \geq \epsilon\sqrt{n}\hat{\sigma})|X_1^n] \to 0$

$$
\begin{aligned}
E[(X_i^*)^2 1(|X_i^*| \geq \epsilon\sqrt{n}\hat{\sigma})|X_1^n] &= \frac{1}{n\hat{\sigma}^2}\sum_i X_i^2 1(|X_i| \geq \epsilon\sqrt{n}\hat{\sigma}) \\
\text{When } \epsilon\sqrt{n}\hat{\sigma} \geq M \qquad &\leq \frac{C}{n}\sum_i X_i^2 1(|X_i| \geq M) \\
&\overset{a.s.}{\to} E[X_i^2 1(|X_i| \geq M)] \\
&= \text{Arbitrarily small for sufficiently large M}
\end{aligned}
$$

Delta method for bootstrap

Theorem

If we have

- $\hat{\theta} \overset{a.s.}{\to} \theta$
- $\sqrt{n}(\hat{\theta} - \theta) \overset{d}{\to} T$
- *Conditionally, almost surely,* $\sqrt{n}(\hat{\theta}^* - \hat{\theta}) \overset{d}{\to} T$
- ϕ *is continuously differentiable in the neighborhood of* θ*, then conditionally almost surely,*

$$\sqrt{n}(\phi(\hat{\theta}^*) - \phi(\hat{\theta}) \overset{d}{\to} \phi'_\theta(T)$$

- *The traditional delta method gives:*

$$\sqrt{n}(\phi(\hat{\theta}) - \phi(\theta) \overset{d}{\to} \phi'_\theta(T)$$

When does bootstrap fail

Example

Let $X_1, \dots, X_n \overset{\text{iid}}{\sim} F = U([0, \theta])$. $X_{(1)}, \dots, X_{(n)}$ are the order statistics.

- The true limiting distribution

$$P\left(\frac{n(\theta - X_{(n)})}{\theta} > x\right) = P\left(X_{(n)} \leq \theta(1 - x/n)\right) = (1 - x/n)^n \to e^{-x}$$

- The bootstrapped limiting distribution

$$P\left(\frac{n(X_{(n)} - X^*_{(n)})}{X_{(n)}} = 0\right) = P(X^*_{(n)} = X_{(n)}) = \left(1 - (1 - 1/n)^n\right) \to 1 - 1/e$$

When does bootstrap fail

Example

Bootstrap works for U statistics, as long as they are not degenerate.

- Rule of thumb: as long as there is normal convergence bootstrap works.
- Are there more robust methods?

Subsampling

- Draw without replacement Y_i which is a size b subsample of X_1^n
- Repeat for all $\binom{n}{b}$ subsets.
- Calculate $\hat{\theta}_{n,b,i} = \hat{\theta}(Y_i)$ and use the empirical distribution of $\tau_b(\hat{\theta}_{n,b,i} - \hat{\theta})$ to approximate the distribution of $\tau_n(\hat{\theta} - \theta)$
- So we want the following

$$J_n(x, P) = P\left(\tau_n(\hat{\theta} - \theta) \leq x\right)$$

- Which we approximate by:

$$L_{n,b}(x) = \frac{1}{\binom{n}{b}} \sum_i 1\left(\tau_b(\hat{\theta}_{n,b,i} - \hat{\theta}) \leq x\right)$$

Subsampling

Theorem

If $b \to \infty$, $b/n \to 0$, $\tau_b/\tau_n \to 0$ as long as $\tau_n(\hat{\theta} - \theta) \xrightarrow{d} Y$, such that, $P(Y \leq x) = J(x, P)$.

$$J_n(x, P) - L_{n,b}(x) \xrightarrow{P} 0$$

- Since $\hat{\theta}_{n,b,i}$ is just an estimator on a smaller sample from the true distribution,

$$\tau_b(\hat{\theta}_b - \theta) \xrightarrow{d} Y \Rightarrow P(\tau_b(\hat{\theta}_b - \theta) \leq x) \to J(x, P)$$

$$\tau_b(\hat{\theta}_b - \theta) = \tau_b(\hat{\theta}_b - \hat{\theta}_n) + \underbrace{\tau_b(\hat{\theta}_n - \theta)}_{O_P(\tau_b/\tau_n) = o_P(1)}$$

Finishing the proof

- Recall that $\tau_n(\hat{\theta} - \theta) \xrightarrow{d} J(., P)$
- So suffices to show that
  $U_{n,b}(x) := \frac{1}{\binom{n}{b}} \sum_i 1\left(\tau_b(\hat{\theta}_{n,b,i} - \theta) \leq x\right) \xrightarrow{P} J(x, P)$

$$U_{n,b}(x) - J(x, P) = U_{n,b}(x) - E[U_{n,b}(x)] + \underbrace{E[U_{n,b}(x)] - J(x, P)}_{\to 0}$$

- Now recall your HW problem:

$$P\left(U_{n,b} - E[U_{n,b}] \geq \epsilon\right) \leq \exp(-\lfloor \frac{n}{b} \rfloor \epsilon^2) \to 0$$

since $b/n \to 0$.