

Homework Assignment 2

Due Friday March 3rd midnight

SDS 384-11 Theoretical Statistics

1. Remember Hoeffding's Lemma? We proved it with a weaker constant in class using a symmetrization type argument. Now we will prove the original version. Let X be a bounded r.v. in $[a, b]$ such that $E[X] = \mu$. Let $f(\lambda) = \log E[e^{\lambda(X-\mu)}]$. Show that $f''(\lambda) \leq (b-a)^2/4$. Now use the fundamental theorem of calculus to write $f(\lambda)$ in terms of $f''(\lambda)$ and finish the argument.
2. Bernstein's inequality for bounded i.i.d sequences of random variables $\{X_i\}$ with $|X_i| \leq M$ gives: $P(|\sum_i (X_i - E[X_i])| \geq t) \leq 2 \exp\left(\frac{-t^2/2}{\sum_i \text{var}(X_i) + Mt/3}\right)$. There is another better inequality called Bennett's inequality, which we will prove here.

- (a) Consider zero mean r.v.s X_i such that $|X_i| \leq b$ and $\text{var}(X_i) = \sigma_i^2$. Prove that

$$\log E[\exp(\lambda X_i)] \leq \sigma_i^2 \lambda^2 \left(\frac{e^{\lambda b} - 1 - \lambda b}{(\lambda b)^2} \right) \quad \forall \lambda \in \mathbb{R}.$$

- (b) Given independent r.v.s $X_i, i = 1, \dots, n$ satisfying the above condition prove

$$\text{(Bennett's inequality)} \quad P\left(\sum_i X_i \geq n\delta\right) \leq \exp\left(-\frac{n\sigma^2}{b^2} h(b\delta/\sigma^2)\right),$$

where $\sigma^2 = \frac{\sum_i \sigma_i^2}{n}$ and $h(t) := (1+t) \log(1+t) - t$ for $t \geq 0$.

- (c) Show that Bennett's inequality is at least as good as Bernstein's inequality.

3. Given a scalar random variable X , suppose that there are positive constants c_1, c_2 such that,

$$P(|X - E[X]| \geq t) \leq c_1 \exp(-c_2 t^2) \quad \forall t \geq 0.$$

- (a) Prove that $\text{var}(X) \leq \frac{c_1}{c_2}$
- (b) A median m_X is any number such that $P(X \geq m_X) \geq 1/2$ and $P(X \leq m_X) \geq 1/2$. Show by example that the median does not need to be unique.
- (c) Show that if the mean concentration bound holds, then for any median m_X , \exists positive constants c_3, c_4 such that

$$P(|X - m_X| \geq t) \leq c_3 \exp(-c_4 t^2),$$

Extra credit Conversely, show that whenever the above median concentration holds, then mean concentration holds with positive constants $c_1 = 2c_3$ and $c_2 = c_4/8$. (Note the change from $c_4/4$ to $c_4/8$). We will prove this using the following steps.

- i. Create a iid copy Y of X . Now show that $\forall \lambda > 0$,

$$P(|X - EX| \geq t) \leq \exp(-\lambda^2 t^2) E_{X,Y} \exp(\lambda^2 (X - Y)^2)$$

Hint: use Jensen's inequality

- ii. Now show that for a suitable choice of λ ,

$$E[\exp(\lambda^2 (X - Y)^2)] \leq 2c_3$$

Hint: use the trick to go from tail bound to moment bound. Use $P(|X - Y| \geq t) \leq P(|X - m_X| \geq t/2) + P(|Y - m_Y| \geq t/2)$, where $m_X = m_Y$ are medians of X and Y .

- iii. Now finally put the above two parts to show the final result.

4. Given a positive semidefinite matrix $Q \in \mathbb{R}^{n \times n}$, consider $Z = \sum_{i,j} Q_{ij} X_i X_j$. When $X_i \sim N(0, 1)$, prove the Hanson-Wright inequality.

$$P(Z \geq \text{trace}(Q) + t) \leq \exp\left(-\min\left\{c_1 t / \|Q\|_{op}, c_2 t^2 / \|Q\|_F^2\right\}\right),$$

where $\|Q\|_{op}$ and $\|Q\|_F$ denote the operator and frobenius norms respectively. *Hint: The rotation-invariance of the Gaussian distribution and sub-exponential nature of χ^2 -variables could be useful.*