

# SDS 384 11: Theoretical Statistics

## **Lecture 2: Stochastic Convergence**

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# Convergence of expectations: exchanging limit and integral

# Theorem (Monotone convergence theorem) If $0 \le X_1 \le X_2 \le \cdots \le X_n \uparrow X$ , then

$$E[X_n] \rightarrow E[X]$$

#### Lemma (Fatou's lemma)

If  $X_n \geq Y$   $\forall n$  for some random variable Y with  $E|Y| < \infty$  then

$$\liminf_{n\to\infty} E[X_n] \geq E[\liminf_n X_n]$$

## Theorem (Dominated convergence theorem)

If  $X_n \stackrel{a.s.}{\to} X$  and  $|X_n| \le Y$  with  $E[|Y|] < \infty$ , then

$$E[X_n] \rightarrow E[X]$$

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# Remember liminf and limsup

- $\liminf_{n\to\infty} a_n = \lim_{n\to\infty} \inf_{m\geq n} a_m$
- $\limsup_{n\to\infty} a_n = \lim_{n\to\infty} \sup_{m\geq n} a_m$

#### MCT→ Fatou

#### Proof.

Consider the random variable X<sub>n</sub> − Y. These are positive. For all m ≥ n,

 $\inf_{k \ge n} (X_k - Y) \le X_m - Y$ 

$$E\left[\inf_{k\geq n}(X_k-Y)\right]\leq E[X_m-Y]\quad \text{Take } E[] \text{ of both sides}$$
 
$$\leq \inf_{m\geq n}E[X_m-Y]$$
 
$$\lim_{n\to\infty}E\left[\inf_{k>n}(X_k-Y)\right]\leq \lim_{n\to\infty}\inf_{m>n}E[X_m-Y]= \liminf_{n\to\infty}E[X_n-Y]$$

- All that is left, is to exchange limit and integral on LHS.
- Note that  $\inf_{k \ge n} (X_k Y)$  is an increasing positive sequence. This converges to  $\liminf_{n \to \infty} (X_n Y)$ . Apply MCT.

### **Fatou** → **DCT**

#### Proof.

- Note that random variable  $-Y \le X_n \le Y$ .
- Also note that  $E[\liminf_{n\to\infty} X_n] = E[X] = E[\limsup_{n\to\infty} X_n].$
- Apply Fatou on  $X_n$  and  $-X_n$ .

$$E[\liminf_{n\to\infty} X_n] \leq \liminf_{n\to\infty} E[X_n] \leq \limsup_{n\to\infty} E[X_n] \leq E[\limsup_{n\to\infty} X_n]$$

• But both ends equal E[X] and so the middle two quantities must be equal and hence proved.

## Things you should know

Consider *n* i.i.d. random variables  $X_i \sim F$ .

Definition (Empirical distribution function) The empirical distribution function is defined as:

$$F_n(x) = \frac{1}{n} \sum_i 1(X_i \le x).$$

**Theorem (Glivenko-Cantelli)** The random variable  $\sup |F_n(x) - F(x)|$  almost surely converges to zero.

$$P\left(\limsup_{n} \sup_{x} |F_n(x) - F(x)| \to 0\right) = 1$$

## Things you should know

Let  $X_1, \ldots X_n$  be i.i.d random variables with  $E[|X_1|] \leq \infty$ , mean  $\mu$ .

Theorem (Weak law of large numbers)

$$\bar{X}_n \stackrel{P}{\rightarrow} \mu$$

Theorem (Strong law of large numbers)

$$\bar{X}_n \stackrel{\textit{a.s.}}{\rightarrow} \mu$$

Theorem (Central limit theorem)

If 
$$E[X_i^2] = \sigma^2$$
,  $\sqrt{n}(\bar{X}_n - \mu) \stackrel{d}{\rightarrow} N(0, \sigma^2)$ .

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# Things you should know

Let  $X_1, ..., X_n$  be i.i.d random variables with mean  $\mu$ .

## Theorem (Berry Esseen)

If  $E[X_i^2] = \sigma^2$ , and  $E[|X_i|^3] = \rho < \infty$ ,

$$\sup_{x} \left| P\left( \frac{\sqrt{n}(\bar{X}_{n} - \mu)}{\sigma} \le x \right) - \Phi(x) \right| \le \frac{C\rho}{\sigma^{3}\sqrt{n}} \qquad \forall n,$$

where  $\Phi(x)$  is the CDF of the standard normal and c is an universal constant known to be greater than 0.4097 and less that 0.7975.

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# Lindeberg-feller CLT for triangular arrays

$$X_{11}$$
 $X_{21}, X_{22}$ 
 $X_{31}, X_{32}, X_{33}$ 

#### **Theorem**

For each n let  $(X_{ni})_{i=1}^n$  be independent random variables with mean zero and variance  $\sigma_{ni}^2$ . Let  $Z_n = \sum_{i=1}^n X_{ni}$  and  $B_n^2 = var(Z_n)$ . Then  $Z_n/B_n \stackrel{d}{\to} N(0,1)$ , as long as the **Lindeberg condition** holds.

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# The Lindeberg condition

**Definition (Lindeberg condition)** For every  $\epsilon > 0$ ,

$$\frac{1}{B_n^2} \sum_{j=1}^n E[X_{nj}^2 1(|X_{nj}| \ge \epsilon B_n)] \to 0 \text{ as } n \to \infty$$
 (1)

**Converse:** If  $\frac{\sigma_{nj}^2}{B_n^2} \to 0$  as  $n \to \infty$ , i.e. no one variance plays a significant role in the limit, and if  $Z_n/B_n \stackrel{d}{\to} N(0,1)$ , then the Lindeberg condition holds.

**Necessary and Sufficient:** If  $\frac{\sigma_{nj}^2}{B_n^2} \to 0$ , the the Lindeberg condition is necessary and sufficient to show the CLT.

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Let  $X_1, \ldots, X_n$  be independent random variables with mean zero and variance one. Do you think  $\sqrt{n}\bar{X}_n \stackrel{d}{\to} N(0,1)$ ?

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$$X_{nj} = \begin{cases} 2j & \text{w.p. } \frac{1}{8j^2} \\ 0 & \text{w.p. } 1 - \frac{1}{4j^2} \\ -2j & \text{w.p. } \frac{1}{8j^2} \end{cases}$$

Let  $X_1, \ldots, X_n$  be independent random variables with mean zero and variance one. Do you think  $\sqrt{n}\bar{X}_n \stackrel{d}{\to} N(0,1)$ ?

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•  $E[X_{nj}] = 0$  and  $var(X_{nj}) = 1$ .  $B_n^2 = n$ .

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- $E[X_{nj}] = 0$  and  $var(X_{nj}) = 1$ .  $B_n^2 = n$ .
- Lets check the Lindeberg condition with  $\epsilon = 1$ .

$$\frac{1}{n}\sum_{j}E[X_{nj}^{2}1(|X_{nj}|\geq\sqrt{n})]=\frac{1}{n}\sum_{j}2\times4j^{2}1(2j\geq\sqrt{n})\frac{1}{8j^{2}}=\frac{1}{n}\sum_{j\geq\sqrt{n}/2}1\to1$$

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- Lets check the Lindeberg condition with  $\epsilon = 1$ .

$$\frac{1}{n}\sum_{j}E[X_{nj}^{2}1(|X_{nj}| \geq \sqrt{n})] = \frac{1}{n}\sum_{j}2\times 4j^{2}1(2j \geq \sqrt{n})\frac{1}{8j^{2}} = \frac{1}{n}\sum_{j>\sqrt{n}/2}1 \rightarrow 1$$

• Since  $\sigma_{nj}^2/B_n^2=1/n\to 0$ , this implies that the CLT does not hold for the sum.

#### **Permutation Tests**

Consider 2n paired experimental units with measurement  $(X_i, Y_i)_{i=1}^n$  in which  $X_j$  is the result of the treatment and  $Y_j$  is the result of control.

•  $H_0$  is that the treatment has had no effect, i.e.  $Z_j = X_j - Y_j$  conditioned on the magnitude  $|Z_j|$  is symmetric, i.e.  $P(Z_i = |z_i|) = P(Z_i = -|z_i|) = 1/2$ .

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- Thus, under  $H_0$ ,  $(Z_1, ..., Z_n)$  has  $2^n$  possible values  $(\pm |z_1|, ..., \pm |z_n|)$ .
- Conditioned on the magnitudes of the differences,  $B_n^2 = \sum_i z_i^2$ .

  Assume that  $\max_i z_i^2/B_n^2 \to 0$ . Then  $\sum_i Z_i/B_n \stackrel{d}{\to} N(0,1)$  using the Lindeberg-feller theorem.

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## Permutation tests: proof

#### Proof.

• Lets check the Lindeberg condition:

$$\begin{split} &\frac{\sum_{j=1}^{n} E[Z_{j}^{2} 1(|Z_{j}| \geq \epsilon B_{n})||Z_{1}| = z_{1}, \ldots, |Z_{n}| = z_{n}]}{B_{n}^{2}} = \frac{\sum_{j} z_{j}^{2} 1(z_{j} \geq \epsilon B_{n})}{B_{n}^{2}} \\ &\leq \frac{(\sum_{j} z_{j}^{2}) 1(\max_{j} z_{j} \geq \epsilon B_{n})}{B_{n}^{2}} \\ &= 1(\max_{j} z_{j} \geq \epsilon B_{n}) \end{split}$$

• Since  $\max_{i} z_{i}^{2}/B_{n}^{2} \to 0$ , the above is zero for all sufficiently large n.

## Tail probabilities

- Most often, we are interested in the question, how far is an empirical quantity from its "population variant"?
- This empirical quantity can be an eigenvalue of a matrix, or the weight vector learned using linear regression and so on.
- For rest of today, and a few more lectures we will brush up on tail inequalities.
- Lets start with the mean?

- How will you bound  $P(|\bar{X}_n \mu| \ge t)$ ? Central limit theorem works under regularity conditions, but its only asymptotic.
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  - Martingale based methods: Azuma-Hoeffding, McDiarmid