

# Influence of ecological traits on mammal extinction risk

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## Abstract

## 1 Introduction

## 2 Methods

### 2.1 Bioprovince occupancy

Bipartite network of species occurrence and spatial unit. Spatial units were defined as 2x2 latitude-longitude grid cells.

### 2.2 Survival model

Marginal posterior of the parameter estimates given the data.

$$p(\theta|y) \propto p(y|\theta)p(\theta) \tag{1}$$

$\theta$  is a vector of parameters.

#### 2.2.1 Sampling distribution

$$\alpha \in \mathbb{R}^+$$

$$\sigma \in \mathbb{R}^+$$

$$y \in [0, \infty)$$

$$\begin{aligned}
p(y|\alpha, \sigma) &= \text{Weibull}(y|\alpha, \sigma) \\
&= \frac{\alpha}{\sigma} \left(\frac{y}{\sigma}\right)^{\alpha-1} \exp\left(-\left(\frac{y}{\sigma}\right)^\alpha\right)
\end{aligned} \tag{2}$$

$$\sigma_{ij} = \frac{\exp(-(r_j + \beta_0 + \sum_{b=1}^B \beta_b X_{i,b}))}{\alpha} \tag{3}$$

The effect of origination cohort ( $j$ ) was included as a random effect  $r$  in the parameterization of  $\sigma$ . The most recent temporal bin, 0 - 2 Mya, was excluded, leaving 32 different cohorts.

Each cohort was considered exchangeable samples of a shared general “cohort effect.” The individual cohort effects was estimated in a hierarchical framework where the between cohort variance constrained the individual cohort effect estimate. This is done by setting a hyperprior on the scale parameter of the prior on  $r$ .

### 2.2.2 Censoring

Right censored observations are observations where the point of extinction has not yet been observed. In this case, this means taxa that are still extant. For each right censored observation, the log probability is incremented by the complementary cumulative density function evaluated at the observed duration.

Left censored observations, on the other hand, correspond to observations that went extinct any time between 0 and some known point. In this study, taxa occurring in only a single time bin were left censored. Because of the minimum resolution of the record, we cannot observe if these taxa went extinct in less than that single bin or not. For each left censored observation, the log probability is incremented by the cumulative density function evaluated at the observed duration.

### 2.2.3 Priors

$$\begin{aligned}
\tau &\sim \text{half Cauchy}(0, 2.5) \\
r_j &\sim \text{Normal}(0, \tau) \\
\beta_0 &\sim \text{Normal}(0, 10) \\
\beta_b &\sim \text{Normal}(0, 10) \\
\alpha &\sim \text{half Cauchy}(0, 2.5)
\end{aligned}$$

#### 2.2.4 Posterior inference

The joint and marginal posteriors of all parameters were approximated using a Markov-chain Monte Carlo (MCMC) routine implemented in the Stan programming language CITATION. Stan implements a Hamiltonian Monte Carlo using a No-U-Turn sampler CITATION.

Posterior approximation was done using four parallel MCMC chains. Chain convergence was evaluated using the scale reduction factor,  $\hat{R}$  CITATION. Values of  $\hat{R}$  close to 1, or less than or equal to 1.1, indicate approximate convergence. Convergence means that the chains are approximately stationary and the samples are well mixed CITATION.

Each chain was evaluated for 2000 steps, with the first 1000 used as warm-up and the last 1000 as samples from the posterior. In total, this yields 4000 samples from the posterior for all estimated parameters.

#### 2.2.5 Posterior predictive checks

$y$

$y^{rep}$

## 3 Results

## 4 Discussion