Non-parametric Inference on Global CO_2 Emission from Cement Production

Sampriti Dutta Sankhadeep Mitra Soumya Paul



 $\begin{tabular}{ll} Under\ the\ guidance\ of:\\ Prof.\ Sharmishtha\ Mitra\\ Professor,\ Department\ of\ Mathematics\ and\ Statistics,\ IIT\ Kanpur \\ \end{tabular}$

Test for Independence

Data Processing

Here is the link of the dataset we are going to use Cement Emissions.

Test for Independence

Data Processing

Here is the link of the dataset we are going to use Cement Emissions.

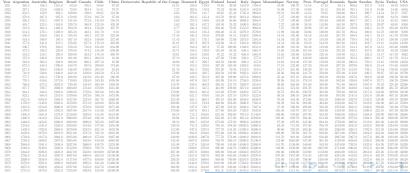
• We have subsetted the data from 1931 onward and dropped any columns with all *NAs* or 0s.

Test for Independence

Data Processing

Here is the link of the dataset we are going to use Cement Emissions.

- We have subsetted the data from 1931 onward and dropped any columns with all *NAs* or 0s.
- Column named Global, which actually represents the total CO₂
 emissions for different time stamps is deleted that as it is not of
 our interest.



Dataset Description Pictorial Visualizatio

Kruskal-Wallis test Test for Independence

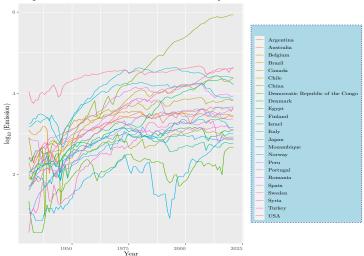
Data Information

Data are collected for 91 years and across 24 different countries.

```
'data.frame': 91 obs. of 25 variables:
##
    $ Year
                                                1931 1932 1933 1934 1935 ...
                                        : num
    $ Argentina
                                               265 247 254 280 356 ....
                                          miim
##
    $ Australia
                                               196 124 160 207 276 ...
                                          num
##
                                               1218 1039 963 938 1087 ...
    $ Belgium
                                        : miim
    $ Brazil
                                               83.6 72.7 109 159.9 181.6 ...
##
                                          nıım
##
    $ Canada
                                               800 363 189 273 273 ...
                                          num
##
    $ Chile
                                               50.9 54.5 69 101.7 141.6 ...
                                          num
##
    $ China
                                               97.9 79.6 113.2 97.9 156.1 ...
                                        : num
    $ Democratic Republic of the Congo:
                                          num
                                               21.81 7.27 3.63 3.63 3.63 ...
##
    $ Denmark
                                               251 204 273 382 374 ...
                                        : num
                                               120 120 142 145 189 ...
    $ Egypt
                                          num
    $ Finland
                                               80 76.2 80 112.7 134.3 ...
##
                                          num
##
    $ Israel
                                               40 50.9 65.4 69 90.9 ...
                                          num
##
    $ Italy
                                               1519 1545 1756 2013 2086 ...
                                          num
##
                                               1788 1843 2366 2304 2904 ...
    $ Japan
                                          nıım
                                               10.9 10.9 10.18 7.27 7.27 ...
##
    $ Mozambique
                                          num
##
    $ Norway
                                          num
                                               110 117 111 124 131 ...
##
    $ Peru
                                               14.5 10.9 14.5 21.8 29.1 ...
                                          nıım
    $ Portugal
                                               47.2 58.1 80 90.9 105.4 ...
                                          num
##
    $ Romania
                                               98.1 105.4 109 156.3 189 ...
                                          num
    $ Spain
                                               807 705 694 672 669 ...
##
                                          num
##
    $ Sweden
                                                258 241 201 287 367 ...
                                          num
                                               3.63 10.9 14.54 10.9 14.54 ...
    $ Syria
                                        : miim
```

CO_2 emission for different countries

CO₂ Emission for different countries over different period of time

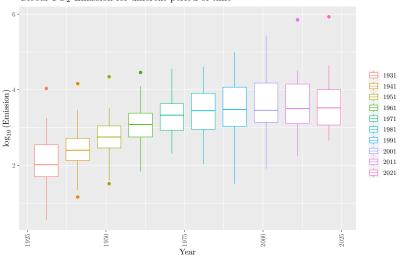


Sign Test (Test for Location) Wilcoxon Rank-Sum (Mann-Whitney U) Test

> Mood's Test Kruskal-Wallis test Test for Independence

Decades-long global CO_2 emission

Global CO_2 Emission for different period of time





• The paired observations may be designated (X, Y). For comparisons of paired observations (X, Y),

- The paired observations may be designated (X, Y). For comparisons of paired observations (X, Y),
- Define $Z_i = X_i Y_i, \forall i = 1, ..., n$,

- The paired observations may be designated (X, Y). For comparisons of paired observations (X, Y),
- Define $Z_i = X_i Y_i, \forall i = 1, ..., n$,
- The basic assumptions are:
 - For a fixed p, 0 the distribution of <math>Z has a common (unknown) quantile, say θ , such that $F_i(\theta) = P(Z_i < \theta) = p, \forall i$.

- The paired observations may be designated (X, Y). For comparisons of paired observations (X, Y),
- Define $Z_i = X_i Y_i, \forall i = 1, ..., n$,
- The basic assumptions are:
 - For a fixed p, 0 the distribution of <math>Z has a common (unknown) quantile, say θ , such that $F_i(\theta) = P(Z_i < \theta) = p, \forall i$.
 - F is continuous in the neighborhood of the unknown median of Z, say M, i.e. $P(Z_i < M) = P(Z_i > M) = \frac{1}{2}, P(Z_i = M) = 0$.

• Here our

$$H_0: M = 0$$
 against $H_A: M \ngeq 0$

• Here our

$$H_0: M = 0 \text{ against } H_A: M \ngeq 0$$

• Define Kas the number of observations for which $z_i > 0$, then under H_0 , $K \sim Bin(n, \frac{1}{2})$, ignoring the zero differences.

• Here our

$$H_0: M = 0 \text{ against } H_A: M \ngeq 0$$

- Define Kas the number of observations for which $z_i > 0$, then under H_0 , $K \sim Bin(n, \frac{1}{2})$, ignoring the zero differences.
- For exact test,
 - Observe the value of K as K_{obs}

Here our

$$H_0: M = 0 \text{ against } H_A: M \ngeq 0$$

- Define Kas the number of observations for which $z_i > 0$, then under H_0 , $K \sim Bin(n, \frac{1}{2})$, ignoring the zero differences.
- For exact test,
 - Observe the value of K as K_{obs}
 - Obtain the *p-value* in the following way

$$\begin{cases} 2\min\left\{\mathbb{P}_{H_0}\left(K \leq K_{obs}\right), \mathbb{P}_{H_0}\left(K \geq K_{obs}\right)\right\} & \text{for } H_A: M \neq 0\\ \mathbb{P}_{H_0}\left(K \geq K_{obs}\right) & \text{for } H_A: M > 0\\ \mathbb{P}_{H_0}\left(K \leq K_{obs}\right) & \text{for } H_A: M < 0 \end{cases}$$

Here our

$$H_0: M = 0 \text{ against } H_A: M \ngeq 0$$

- Define Kas the number of observations for which $z_i > 0$, then under H_0 , $K \sim Bin(n, \frac{1}{2})$, ignoring the zero differences.
- For exact test,
 - Observe the value of K as K_{obs}
 - Obtain the *p-value* in the following way

$$\begin{cases} 2\min\left\{\mathbb{P}_{H_0}\left(K\leq K_{obs}\right), \mathbb{P}_{H_0}\left(K\geq K_{obs}\right)\right\} & \text{for } H_A: M\neq 0\\ \mathbb{P}_{H_0}\left(K\geq K_{obs}\right) & \text{for } H_A: M>0\\ \mathbb{P}_{H_0}\left(K\leq K_{obs}\right) & \text{for } H_A: M<0 \end{cases}$$

• For α level of significance, reject H_0 if p-value $\leq \alpha$.

Here our

$$H_0: M = 0 \text{ against } H_A: M \ngeq 0$$

- Define Kas the number of observations for which $z_i > 0$, then under H_0 , $K \sim Bin(n, \frac{1}{2})$, ignoring the zero differences.
- For exact test,
 - Observe the value of K as K_{obs}
 - Obtain the *p-value* in the following way

$$\begin{cases} 2\min\left\{\mathbb{P}_{H_0}\left(K \leq K_{obs}\right), \mathbb{P}_{H_0}\left(K \geq K_{obs}\right)\right\} & \text{for } H_A: M \neq 0 \\ \mathbb{P}_{H_0}\left(K \geq K_{obs}\right) & \text{for } H_A: M > 0 \\ \mathbb{P}_{H_0}\left(K \leq K_{obs}\right) & \text{for } H_A: M < 0 \end{cases}$$

- For α level of significance, reject H_0 if p-value $\leq \alpha$.
- For the asymptotic test we conclude in the same way as exact test, just with the addition of the assumption of normality and finite population correction.

For decades-long global CO_2 emission:

We first conduct both sided test for year 1931 and 1941.

As we can see that p-value is less than 0.05, we reject the null hypothesis that CO_2 emission doesn't differ at 5% level of significance.

Moreover, we are interested to know whether CO_2 emission for year 1931 is less than the CO_2 emission for year 1941.

```
##
## One-sample Sign-Test
##
## data: z
## S = 4, number of differences = 23, p-value = 0.0013
## alternative hypothesis: true median is less than 0
## 96.8 percent confidence interval:
## -Inf -18.18
## sample estimates:
## median of the differences
## -70.85
```

Here, p-value is less than 0.05 and we reject the null hypothesis that CO_2 emission doesn't differ at 5% level of significance. Thus we can conclude that there is a significant increase in CO_2 emission.

The table below shows the decisions for all pairs possible. Note that we have performed exact test here.

	Decision made
1931 vs 1941	There is significant increase in CO_2 emission
1941 vs 1951	There is significant increase in CO_2 emission
1951 vs 1961	There is significant increase in CO_2 emission
1961 vs 1971	There is significant increase in CO_2 emission
1971 vs 1981	There is no significant change in CO_2 emission
1981 vs 1991	There is no significant change in CO_2 emission
1991 vs 2001	There is significant increase in CO_2 emission
2001 vs 2011	There is no significant change in CO_2 emission
2011 vs 2021	There is no significant change in CO_2 emission

Table: Decision using Sign test

For Countries:

We firstly conduct both sided test for Argentina and Australia.

```
##
##
    One-sample Sign-Test
##
  data:
  S = 55, number of differences = 91, p-value = 0.05857
  alternative hypothesis: true median is not equal to 0
   96.5 percent confidence interval:
##
     -4.0148.9
## sample estimates:
## median of the differences
                         87.2
##
```

As we can see that p-value is greater than 0.05 we fail to reject the null hypothesis that CO_2 emission doesn't differ at 5% level of significance.

The table below shows the decisions for first ten pairs possible. Note that we have performed asymptotic test here.

Argentina vs Australia
Argentina vs Belgium
Argentina vs Brazil
Argentina vs Canada
Argentina vs Chile
Argentina vs China
Argentina vs Democratic Republic of the Congo
Argentina vs Denmark
Argentina vs Egypt
Argentina vs Finland

Decision made

There is no significant change in CO_2 emission There is significant increase in CO_2 emission There is significant increase in CO_2 emission There is significant increase in CO_2 emission There is significant decrease in CO_2 emission There is significant increase in CO_2 emission There is significant decrease in CO_2 emission There is significant decrease in CO_2 emission There is no significant change in CO_2 emission There is no significant decrease in CO_2 emission There is significant decrease in CO_2 emission

Table: Decision from sign test

• $X_1, X_2,, X_n$ are random samples from absolutely continuous distribution function F_X (.)

- $X_1, X_2,, X_n$ are random samples from absolutely continuous distribution function F_X (.)
- $Y_1, Y_2,, Y_n$ are random samples from absolutely continuous distribution function F_Y (.)

- $X_1, X_2,, X_n$ are random samples from absolutely continuous distribution function F_X (.)
- $Y_1, Y_2,, Y_n$ are random samples from absolutely continuous distribution function $F_Y(.)$
- \bullet X's and Y's are independent

• Here our

$$H_0: F_Y(x) = F_X(x)$$
 against $H_A: F_Y(x) \ngeq F_X(x)$

• Here our

$$H_0: F_Y(x) = F_X(x)$$
 against $H_A: F_Y(x) \not\supseteq F_X(x)$

 Define Was the sum of ranks of the X observations from the combined ordered arrangement of the sample.

Here our

$$H_0: F_Y(x) = F_X(x)$$
 against $H_A: F_Y(x) \ngeq F_X(x)$

- Define Was the sum of ranks of the X observations from the combined ordered arrangement of the sample.
- For exact test,
 - Observe the value of W as W_{obs}

• Here our

$$H_0: F_Y(x) = F_X(x)$$
 against $H_A: F_Y(x) \ngeq F_X(x)$

- Define Was the sum of ranks of the X observations from the combined ordered arrangement of the sample.
- For exact test,
 - Observe the value of W as W_{obs}
 - Obtain the *p-value* in the following way

$$\begin{cases} 2\min\left\{\mathbb{P}_{H_{0}}\left(W\leq W_{obs}\right),\mathbb{P}_{H_{0}}\left(W\geq W_{obs}\right)\right\} & \text{ for } H_{A}:F_{Y}\left(x\right)\neq F_{X}\left(x\right)\\ \mathbb{P}_{H_{0}}\left(W\geq W_{obs}\right) & \text{ for } H_{A}:F_{Y}\left(x\right)>F_{X}\left(x\right)\\ \mathbb{P}_{H_{0}}\left(W\leq W_{obs}\right) & \text{ for } H_{A}:F_{Y}\left(x\right)< F_{X}\left(x\right) \end{cases}$$

• Here our

$$H_0: F_Y(x) = F_X(x)$$
 against $H_A: F_Y(x) \ngeq F_X(x)$

- Define Was the sum of ranks of the X observations from the combined ordered arrangement of the sample.
- For exact test,
 - Observe the value of W as W_{obs}
 - Obtain the *p-value* in the following way

$$\begin{cases} 2\min\left\{\mathbb{P}_{H_{0}}\left(W\leq W_{obs}\right),\mathbb{P}_{H_{0}}\left(W\geq W_{obs}\right)\right\} & \text{ for } H_{A}:F_{Y}\left(x\right)\neq F_{X}\left(x\right)\\ \mathbb{P}_{H_{0}}\left(W\geq W_{obs}\right) & \text{ for } H_{A}:F_{Y}\left(x\right)>F_{X}\left(x\right)\\ \mathbb{P}_{H_{0}}\left(W\leq W_{obs}\right) & \text{ for } H_{A}:F_{Y}\left(x\right)< F_{X}\left(x\right) \end{cases}$$

• For α level of significance, reject H_0 if p-value $\leq \alpha$.

Here our

$$H_0: F_Y(x) = F_X(x)$$
 against $H_A: F_Y(x) \ngeq F_X(x)$

- Define Was the sum of ranks of the X observations from the combined ordered arrangement of the sample.
- For exact test,
 - Observe the value of Was W_{obs}
 - Obtain the *p-value* in the following way

$$\begin{cases} 2\min\left\{\mathbb{P}_{H_{0}}\left(W\leq W_{obs}\right),\mathbb{P}_{H_{0}}\left(W\geq W_{obs}\right)\right\} & \text{ for } H_{A}:F_{Y}\left(x\right)\neq F_{X}\left(x\right)\\ \mathbb{P}_{H_{0}}\left(W\geq W_{obs}\right) & \text{ for } H_{A}:F_{Y}\left(x\right)>F_{X}\left(x\right)\\ \mathbb{P}_{H_{0}}\left(W\leq W_{obs}\right) & \text{ for } H_{A}:F_{Y}\left(x\right)< F_{X}\left(x\right) \end{cases}$$

- For α level of significance, reject H_0 if p-value $\leq \alpha$.
- For the asymptotic test we conclude in the same way as exact test, just with the addition of the assumption of normality and finite population correction.

For decades-long global CO_2 emission:

Here we have shown the output for the years 1931 and 1941.

As the p-value is greater than 0.05, we fail to reject H_0 at 5% level of significance.

The table below shows the results for all decades.

```
Decision made
1931 vs 1941
               There is no significant change in CO_2 emission
1941 vs 1951
               There is significant increase in CO_2 emission
1951 vs 1961
               There is significant increase in CO_2 emission
1961 vs 1971
               There is no significant change in CO_2 emission
1971 vs 1981
               There is no significant change in CO_2 emission
1981 vs 1991
               There is no significant change in CO_2 emission
               There is no significant change in CO_2 emission
1991 vs 2001
2001 vs 2011
               There is no significant change in CO_2 emission
2011 vs 2021
               There is no significant change in CO_2 emission
```

Table: Decision values using Wilcoxon Rank-Sum test

For Countries:

Here we give a snapshot for Australia and Argentina.

```
##
##
    Wilcoxon rank sum test with continuity correction
##
  data: x and y
  W = 4277.5, p-value = 0.7009
## alternative hypothesis: true location shift is not equal to 0
   95 percent confidence interval:
  -308.0 399.7
##
## sample estimates:
## difference in location
                 72.70007
##
```

The p-value for the both sided test is greater than 0.05. So we can accept H_0 that there is indeed no significant difference between the CO₂ emission of Australia and Argentina, at 5% level of significant.

The below table shows the result for some pairs of countries.

Decision made
There is no significant change in CO_2 emission
There is no significant change in CO_2 emission
There is significant increase in CO_2 emission
There is significant increase in CO_2 emission
There is significant decrease in CO_2 emission
There is significant increase in CO_2 emission
There is significant decrease in CO_2 emission
There is significant decrease in CO_2 emission
There is no significant change in CO_2 emission
There is significant decrease in CO_2 emission

Table: Decision values using Wilcoxon Rank-Sum test

• The (unknown) medians (or means) of the two samples are equal or the sample observations can be adjusted to have equal locations

Assumptions

- The (unknown) medians (or means) of the two samples are equal or the sample observations can be adjusted to have equal locations
- $X_1, X_2,, X_n$ are random samples from absolutely continuous distribution function F_X (.)

Assumptions

- The (unknown) medians (or means) of the two samples are equal or the sample observations can be adjusted to have equal locations
- $X_1, X_2,, X_n$ are random samples from absolutely continuous distribution function F_X (.)
- $Y_1, Y_2,, Y_n$ are random samples from absolutely continuous distribution function $F_Y(.)$

Assumptions

- The (unknown) medians (or means) of the two samples are equal or the sample observations can be adjusted to have equal locations
- $X_1, X_2,, X_n$ are random samples from absolutely continuous distribution function F_X (.)
- $Y_1, Y_2,, Y_n$ are random samples from absolutely continuous distribution function $F_Y(.)$
- X's and Y's are independent

Exploratory Data Analysis
Sign Test (Test for Location)
Wilcoxon Rank-Sum (Mann-Whitney U) Test
Mood's Test
Kruskal-Wallis test
Test for Independence

Methodology

Here our

 $H_0: F_Y$ and F_X have equal dispersion against

 H_A : dispersion of $Y \not\stackrel{>}{\not\equiv}$ dispersion of X

Here our

 H_0 : F_Y and F_X have equal dispersion against

 H_A : dispersion of $Y \not\supseteq dispersion$ of X

• Here a linear rank statistic M is defined as, $M = \sum_{i=1}^{N} \left(i - \frac{N+1}{2}\right)^2 Z_i$, where

 $Z_i=1,$ if the i th observation of the ordered combined sample is an X observation and $Z_i=0,$ otherwise.

Here our

 $H_0: F_Y$ and F_X have equal dispersion against

 H_A : dispersion of $Y \not\stackrel{>}{\not\equiv}$ dispersion of X

 \bullet Here a linear rank statistic M is defined as, $M = \sum_{i=1}^N \left(i - \frac{N+1}{2}\right)^2 Z_i,$ where

 $Z_i=1,$ if the i th observation of the ordered combined sample is an X observation and $Z_i=0,$ otherwise.

- For exact test,
 - Observe the value of Mas M_{obs}

Here our

 H_0 : F_Y and F_X have equal dispersion against

 H_A : dispersion of $Y \not\stackrel{>}{\not\equiv}$ dispersion of X

 \bullet Here a linear rank statistic M is defined as, $M = \sum_{i=1}^N \left(i - \frac{N+1}{2}\right)^2 Z_i,$ where

 $Z_i = 1$, if the *i* th observation of the ordered combined sample is an *X* observation and $Z_i = 0$, otherwise.

- For exact test,
 - Observe the value of M as M_{obs}
 - Obtain the p-value in the following way

$$\begin{cases} 2 \min \left\{ \mathbb{P}_{H_0} \left(M \leq M_{obs} \right), \mathbb{P}_{H_0} \left(M \geq M_{obs} \right) \right\} & \text{for } H_A \text{ : dispersion of } Y \neq \text{dispersion of } X \\ \mathbb{P}_{H_0} \left(M \geq M_{obs} \right) & \text{for } H_A \text{ : dispersion of } Y < \text{dispersion of } X \\ \mathbb{P}_{H_0} \left(M \leq M_{obs} \right) & \text{for } H_A \text{ : dispersion of } Y > \text{dispersion of } X \end{cases}$$

Here our

 H_0 : F_Y and F_X have equal dispersion against

 H_A : dispersion of $Y \not\supseteq$ dispersion of X

 \bullet Here a linear rank statistic M is defined as, $M = \sum_{i=1}^N \left(i - \frac{N+1}{2}\right)^2 Z_i, \text{where}$

 $Z_i=1,$ if the i th observation of the ordered combined sample is an X observation and $Z_i=0,$ otherwise.

- For exact test,
 - Observe the value of M as M_{obs}
 - Obtain the p-value in the following way

$$\begin{cases} 2 \min \left\{ \mathbb{P}_{H_0} \left(M \leq M_{obs} \right), \mathbb{P}_{H_0} \left(M \geq M_{obs} \right) \right\} & \text{for } H_A : \text{dispersion of } Y \neq \text{dispersion of } X \\ \mathbb{P}_{H_0} \left(M \geq M_{obs} \right) & \text{for } H_A : \text{dispersion of } Y < \text{dispersion of } X \\ \mathbb{P}_{H_0} \left(M \leq M_{obs} \right) & \text{for } H_A : \text{dispersion of } Y > \text{dispersion of } X \end{cases}$$

• For α level of significance, reject H_0 if p-value $\leq \alpha$.

Here our

 H_0 : F_Y and F_X have equal dispersion against

 H_A : dispersion of $Y \not\supseteq$ dispersion of X

 \bullet Here a linear rank statistic M is defined as, $M = \sum_{i=1}^N \left(i - \frac{N+1}{2}\right)^2 Z_i,$ where

 $Z_i = 1$, if the i th observation of the ordered combined sample is an X observation and $Z_i = 0$, otherwise.

- For exact test,
 - Observe the value of M as M_{obs}
 - Obtain the p-value in the following way

$$\begin{cases} 2 \min \left\{ \mathbb{P}_{H_0} \left(M \leq M_{obs} \right), \mathbb{P}_{H_0} \left(M \geq M_{obs} \right) \right\} & \text{for } H_A \text{ : dispersion of } Y \neq \text{dispersion of } X \\ \mathbb{P}_{H_0} \left(M \geq M_{obs} \right) & \text{for } H_A \text{ : dispersion of } Y < \text{dispersion of } X \\ \mathbb{P}_{H_0} \left(M \leq M_{obs} \right) & \text{for } H_A \text{ : dispersion of } Y > \text{dispersion of } X \end{cases}$$

- For α level of significance, reject H_0 if p-value $\leq \alpha$.
- For the asymptotic test we conclude in the same way as exact test, just with the addition of the assumption of normality and finite population correction.

For decades-long global CO_2 emission:

From the results of Wilcoxon Rank,-Sum only for four pairs (1971,1981),(1981,1991); (2001,2011); (2011,2021) we have performed Mood's Test. First the result for the years 1971 and 1981 are shown below:

```
##
## Mood two-sample test of scale
##
## data: x and y
## Z = 1.1714, p-value = 0.2414
## alternative hypothesis: two.sided
```

The p-value is much larger than 0.05 and hence we accept the null hypothesis at 5% level of significance and conclude that there is no significant difference between the variability of the year 1971 and 1981.

The results for all some pairs are presented here.

	Decision made
1961 vs 1971	There is no significant change in ${\cal C}{\cal O}_2$ emission
1971 vs 1981	There is no significant change in CO_2 emission
1981 vs 1991	There is no significant change in CO_2 emission
1991 vs 2001	There is no significant change in ${\cal C}{\cal O}_2$ emission
2001 vs 2011	There is no significant change in CO_2 emission
2011 vs 2021	There is no significant change in CO_2 emission

Table: Decisions using Mood's test

For Countries:

From the results of Wilcoxon Rank-Sum, 35 pairs of countries have no significant difference in CO_2 emission like, (Australia, Argentina); (Belgium. Romania) etc.

We have shown the result for Australia and Argentina.

```
##
## Mood two-sample test of scale
##
## data: x and y
## Z = -0.16643, p-value = 0.8678
## alternative hypothesis: two.sided
```

As the p-value is much larger than 0.05, so we conclude that there is no significant difference in variability of Australia and Argentina.

Exploratory Data Analysis
Sign Test (Test for Location)
Wilcoxon Rank-Sum (Mann-Whitney U) Test
Mood's Test
Kruskal-Wallis test
Test for Independence

Result

Also the results for some pairs are presented,

Decision made

There is no significant change in CO_2 emission Argentina vs Australia Argentina vs Belgium There is significant decrease in CO_2 emission Argentina vs Egypt There is significant increase in CO_2 emission Argentina vs Romania There is significant increase in CO_2 emission Australia vs Belgium There is significant decrease in CO_2 emission Australia vs Egypt There is significant increase in CO_2 emission Australia vs Romania There is significant increase in CO_2 emission There is significant increase in CO_2 emission Belgium vs Egypt Belgium vs Romania There is significant increase in CO_2 emission Brazil vs Spain There is significant decrease in CO_2 emission Brazil vs Turkey There is significant increase in CO_2 emission Canada vs Egypt There is significant increase in CO_2 emission Canada vs Turkey There is significant increase in CO_2 emission Chile vs Finland There is significant decrease in CO_2 emission There is significant increase in CO_2 emis Chile vs Israel Sampriti Dutta, Sankhadeep Mitra, Soumya Paul

Assumption

• All k sets of independent samples are drawn from absolutely continuous distribution function $F_i(.)$ $\forall i = 1(1)k$.

Assumption

- All k sets of independent samples are drawn from absolutely continuous distribution function $F_i(.)$ $\forall i = 1(1)k$.
- In the location alternative setup the populations are identical in every respect, except the location parameter.

• Here our

$$H_0: F_1(x) = F_2(x) = \dots = F_k(x)$$
 against H_A : not H_0

• Here our

$$H_0: F_1(x) = F_2(x) = \dots = F_k(x)$$
 against H_A : not H_0

Rank all data from all groups together. Assign any tied values
the average of the ranks they would have received had they not
been tied.

Here our

$$H_0: F_1(x) = F_2(x) = ... = F_k(x) \text{ against } H_A : \text{not } H_0$$

- Rank all data from all groups together. Assign any tied values the average of the ranks they would have received had they not been tied.
- The test statistic is given by the equation below

$$H = \frac{12}{N(N-1)} \sum_{i=1}^{k} \frac{R_i^2}{n_i} - 3(N+1)$$

Where N is the total sample size, k is the number of groups we are comparing, R_i is the sum of ranks for group i, and n_i is the sample size of group i.

- For exact test,
 - Observe the value of $Has\ H_{obs}$

- For exact test,
 - Observe the value of $Has\ H_{obs}$
 - Obtain the *p-value* in the following way

$$\mathbb{P}_{H_0} (H \geq H_{obs}) \text{ for } H_A : \text{not } H_0$$

- For exact test,
 - Observe the value of $Has\ H_{obs}$
 - \bullet Obtain the *p-value* in the following way

$$\mathbb{P}_{H_0} (H \ge H_{obs}) \text{ for } H_A : \text{not } H_0$$

• For α level of significance, reject H_0 if p-value $\leq \alpha$.

- For exact test,
 - Observe the value of Has H_{obs}
 - \bullet Obtain the *p-value* in the following way

$$\mathbb{P}_{H_0} (H \ge H_{obs}) \text{ for } H_A : \text{not } H_0$$

- For α level of significance, reject H_0 if p-value $\leq \alpha$.
- Asymptotically H follows χ_{k-1}^2 distribution. So, we perform the asymptotic test accordingly.

Exploratory Data Analysis Sign Test (Test for Location) Wilcoxon Rank-Sum (Mann-Whitney U) Test Mood's Test Kruskal-Wallis test Test for Independence

Result

First we found the sample standard deviation. Then we have performed Kruskal-Wallis test for those group (size> 2), for which the standard deviations are close enough. The countries Argentina, Peru, Portugal have close standard deviation.

	Standard Deviation		
Argentina	1327.8789		
Australia	1245.9507		
Belgium	885.7870		
Brazil	7882.2382		
Canada	2393.8287		
Chile	343.2244		
China	266589.4357		
Democratic Republic of the Congo	104.2538		
Denmark	392.7684		
Egypt	7270.9458		
Finland	251.4007		
Israel		≣ '	9

	Standard Deviation
Italy	6102.5108
Japan	14658.8081
Mozambique	246.2908
Norway	352.0404
Peru	1365.9741
Portugal	1368.9408
Romania	2203.5197
Spain	5345.5497
Sweden	448.3529
Syria	736.1023
Turkey	12109.3962
USA	10951.2528

For decades we haven't found any set of interest as the standard deviations are not that much close.

	Standard Deviation
1931	2194.929
1941	2939.747
1951	4412.352
1961	6129.387
1971	9117.920
1981	11615.976
1991	21798.270
2001	54726.254
2011	143328.878
2021	172785.052

The p-value of this test for these 3 countries is very smaller than 0.05, so we reject H_0 at 5% level of significance and conclude that there is significant difference in CO_2 emission between Argentina, Peru, Portugal.

```
##
## Kruskal-Wallis rank sum test
##
## data: dat_gather_2$Emission and as.factor(dat_gather_2$Count
## Kruskal-Wallis chi-squared = 23.74, df = 2, p-value = 6.997e-
```

• Let $(X_1, Y_1), (X_2, Y_2), ...(X_n, Y_n)$, be a set of observations from the joint distribution function $F_{(X,Y)}$

- Let $(X_1, Y_1), (X_2, Y_2), ...(X_n, Y_n)$, be a set of observations from the joint distribution function $F_{(X,Y)}$
- We test for the null hypothesis $H_0: F_{X,Y}(x,y) = F_X(x)F_Y(y)$ against H_A : Not H_0 .

- Let $(X_1, Y_1), (X_2, Y_2), ...(X_n, Y_n)$, be a set of observations from the joint distribution function $F_{(X,Y)}$
- We test for the null hypothesis $H_0: F_{X,Y}(x,y) = F_X(x)F_Y(y)$ against H_A : Not H_0 .
- Let, π_c , be the probability of concordance and π_d , be the probability of discordance.

- Let $(X_1, Y_1), (X_2, Y_2), ...(X_n, Y_n)$, be a set of observations from the joint distribution function $F_{(X,Y)}$
- We test for the null hypothesis $H_0: F_{X,Y}(x,y) = F_X(x)F_Y(y)$ against H_A : Not H_0 .
- Let, π_c , be the probability of concordance and π_d , be the probability of discordance.
- Then **Kendall's** τ is defined as; $\tau = \pi_c \pi_d$.

• For a sample of size N, the sample estimate of τ is defined as:

$$T = \frac{\sum\limits_{1 \le i < j \le N} A_{ij}}{\binom{N}{2}},$$

, where $A_{ij} = sign(X_j - X_i) sign(Y_j - Y_i)$ which to equivalent to saying

$$A_{ij} = \begin{cases} 1 & \text{, if pairs are concordant with probability } \pi_c \\ 0 & \text{, if atleast one component is tied with probability } \pi_t = 1 - \pi_c - \pi_d \\ -1 & \text{, if pairs are discordant with probability } \pi_d \end{cases}$$

- For exact test,
 - Observe the value of T as T_{obs}

- For exact test,
 - Observe the value of T as T_{obs}
 - Obtain the *p-value* in the following way

$$\begin{cases} 2 \min \left\{ \mathbb{P}_{H_0} \left(T \leq T_{obs} \right), \mathbb{P}_{H_0} \left(T \geq T_{obs} \right) \right\} & \text{for } H_A \text{ : there is dependence between two pop } \\ \mathbb{P}_{H_0} \left(T \geq T_{obs} \right) & \text{for } H_A \text{ : there is positive association} \\ \mathbb{P}_{H_0} \left(T \leq T_{obs} \right) & \text{for } H_A \text{ : there is negative association} \end{cases}$$

- For exact test,
 - Observe the value of T as T_{obs}
 - Obtain the *p-value* in the following way

$$\begin{cases} 2 \min \left\{ \mathbb{P}_{H_0} \left(T \leq T_{obs} \right), \mathbb{P}_{H_0} \left(T \geq T_{obs} \right) \right\} & \text{for } H_A \text{ : there is dependence between two pop } \\ \mathbb{P}_{H_0} \left(T \geq T_{obs} \right) & \text{for } H_A \text{ : there is positive association} \\ \mathbb{P}_{H_0} \left(T \leq T_{obs} \right) & \text{for } H_A \text{ : there is negative association} \end{cases}$$

• For α level of significance, reject H_0 if p-value $\leq \alpha$.

- For exact test,
 - Observe the value of T as T_{obs}
 - Obtain the *p-value* in the following way

$$\begin{cases} 2 \min \left\{ \mathbb{P}_{H_0} \left(T \leq T_{obs} \right), \mathbb{P}_{H_0} \left(T \geq T_{obs} \right) \right\} & \text{for } H_A \text{ : there is dependence between two pop } \\ \mathbb{P}_{H_0} \left(T \geq T_{obs} \right) & \text{for } H_A \text{ : there is positive association} \\ \mathbb{P}_{H_0} \left(T \leq T_{obs} \right) & \text{for } H_A \text{ : there is negative association} \end{cases}$$

- For α level of significance, reject H_0 if p-value $\leq \alpha$.
- For the asymptotic test we conclude in the same way as exact test, just with the addition of the assumption of normality and finite population correction.

Sampriti Dutta, Sankhadeep Mitra, Soumya Paul

Result

For Countries: Here we give a snapshot for *Australia* and *Argentina*.

```
##
##
    Kendall's rank correlation tau
##
## data: x and y
  z = 9.661, p-value < 2.2e-16
  alternative hypothesis: true tau is not equal to 0
## sample estimates:
##
         tan
## 0.6884085
##
##
    Kendall's rank correlation tau
##
## data: x and y
  z = 9.661, p-value < 2.2e-16
## alternative hypothesis. true tan is greater than (
```

Exploratory Data Analysis
Sign Test (Test for Location)
Wilcoxon Rank-Sum (Mann-Whitney U) Test
Mood's Test
Kruskal-Wallis test
Test for Independence

Result

The p-value for the both sided test is less than 0.05. So we can reject H_0 and also for greater than test p-value is less than 0.05 which implies H_0 is rejected also and thus we can conclude that there is indeed positive association between the CO_2 emission of Australia and Argentina, at 5% level of significant.

The below table shows the result for some pairs of countries.

Decision made
There is positive association

Table: Decisions from Kendall Rank Correlation Coefficient

Notice that there is only positive association present in all the pairs.

```
## Decision made
## There is positive association
## 276
```

Thus, we can conclude that CO_2 emission is moving in same direction for all the countries.

Thank You