

# Mediation and Moderation

GIVEN THE PACE OF  
TECHNOLOGY, I PROPOSE  
WE LEAVE MATH TO THE  
MACHINES AND GO PLAY  
OUTSIDE.



# **Logistics**

# **Things that came up**

# Great use of gganimate



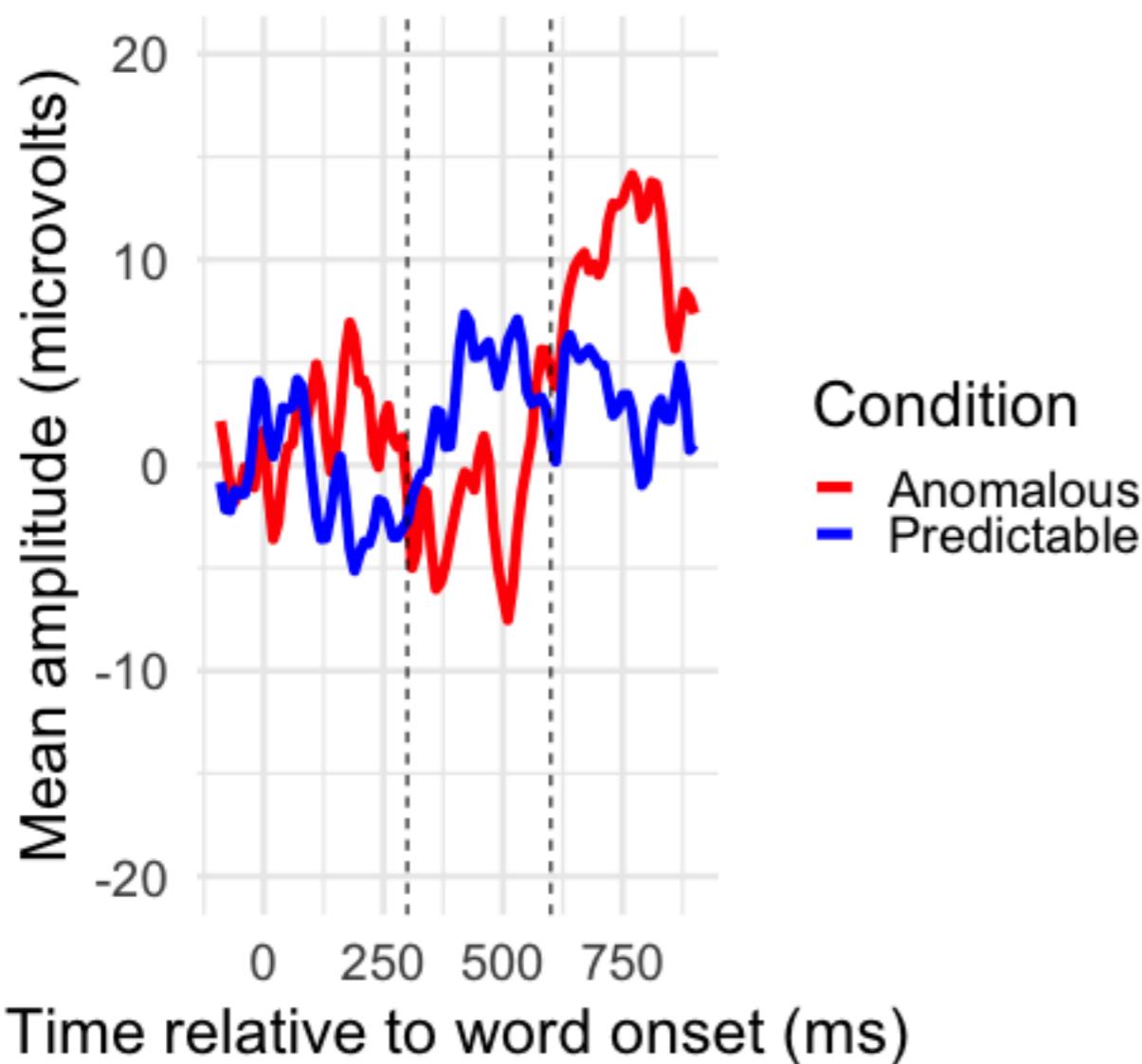
Rachel A. Ryskin

[ryskin@mit.edu](mailto:ryskin@mit.edu)

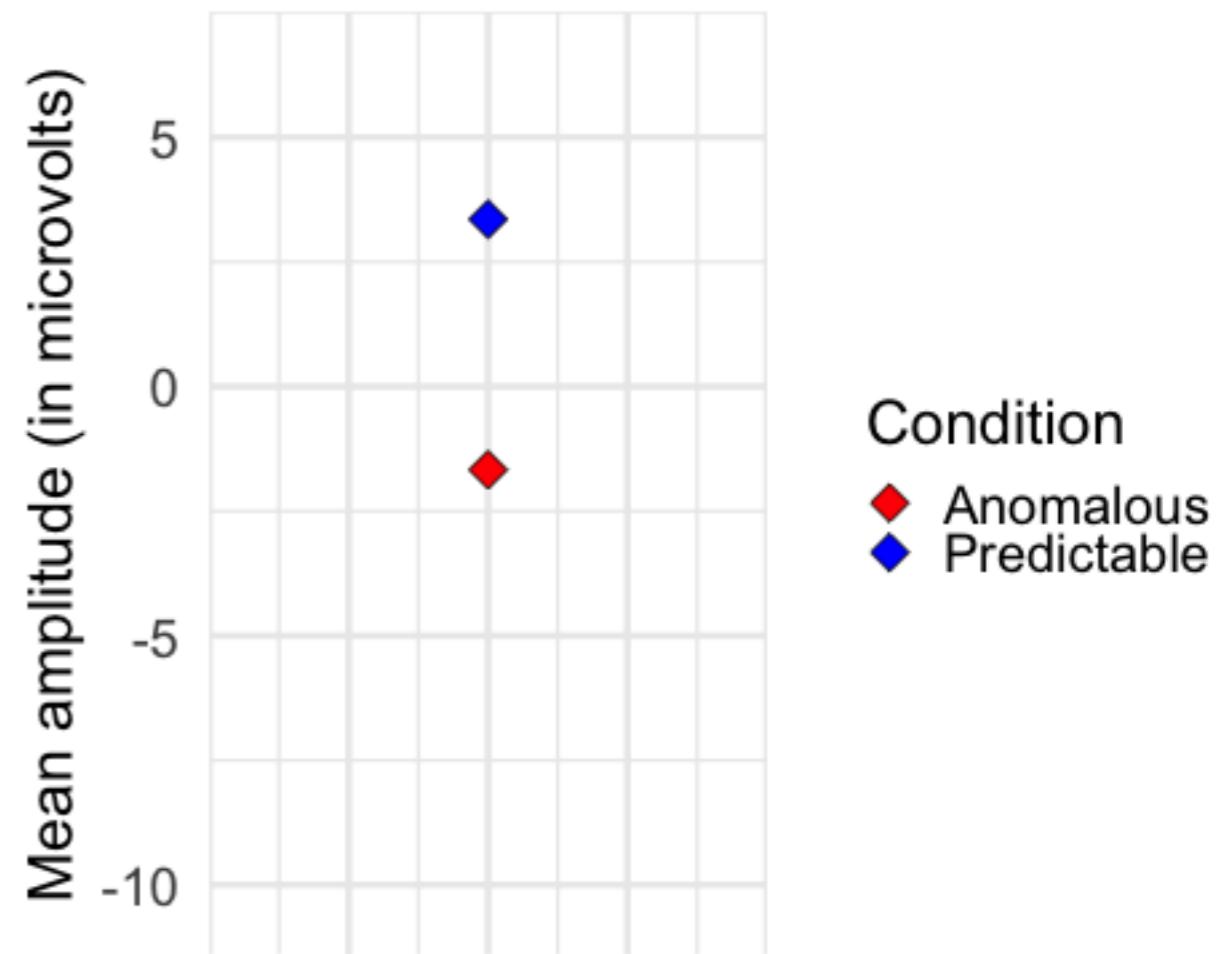
[Curriculum Vitae](#)

[Google Scholar](#) | [GitHub](#) | [Twitter](#)

Number of subjects: 1

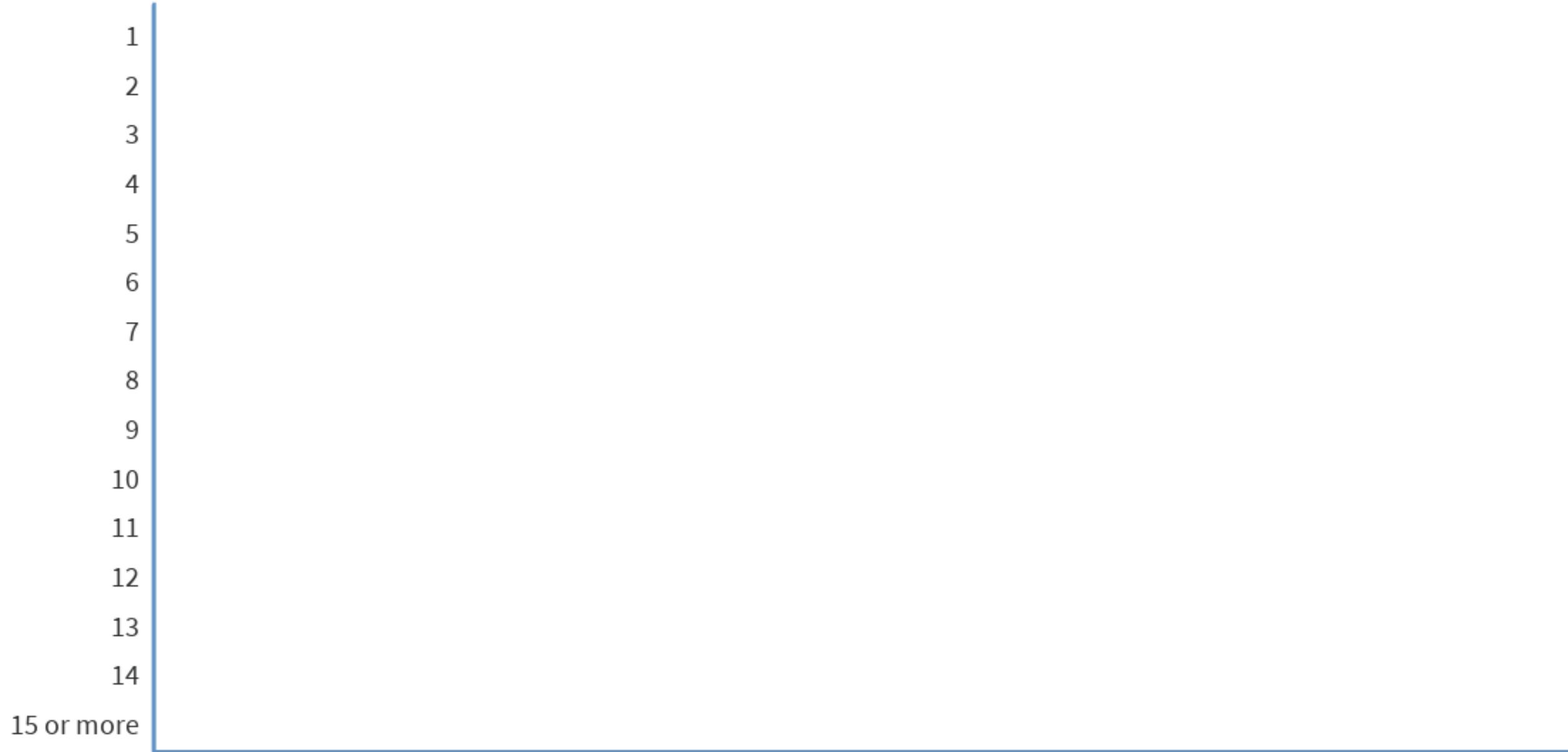


Mean over 300-600ms window  
(N=1)



# **Homework 4**

# How many hours did it take you to complete Homework 4?



# Midterm

# 2 small errors in earliest release

if you downloaded the midterm before 3pm

## Question 12

**Question 14:** Make two residual plots (for the model fitted in Question 5). One that shows model predictions on the x-axis and residuals on the y-axis, and another that shows a density plot of the residuals. Do these residual plots look ok? Describe briefly what to look for when checking the model assumptions.

# YOUR CODE HERE

Your answer:

## Question 12

**Question 15:** Compare the augmented model (which includes whether a person has any kids as a predictor) with the model you've fit in Question X. Report the results with PRE as your effect size.

# YOUR CODE HERE

Your answer:

# Feel free to add packages

## Load libraries

```
library("knitr")      # for knitting  
library("pwr")        # for power analysis  
library("broom")       # for tidying model fits  
library("effectsize") # for effect size measures  
library("tidyverse")   # for everything else
```

# code chunk for question 19

**Question 19** (4 points): What could be a reason that age is not a significant predictor of life satisfaction when controlling for the number of kids? Give a brief explanation.

```
# YOUR CODE HERE
```

Your answer:



**not necessary to use this code chunk**

# Plan for today

- Questions? Answers!
  - degrees of freedom
  - t-test vs. permutation test
  - overlapping confidence intervals
- Power analysis (continued)
- Controlling for variables
- Mediation
- Moderation

# **Questions? Answers!**

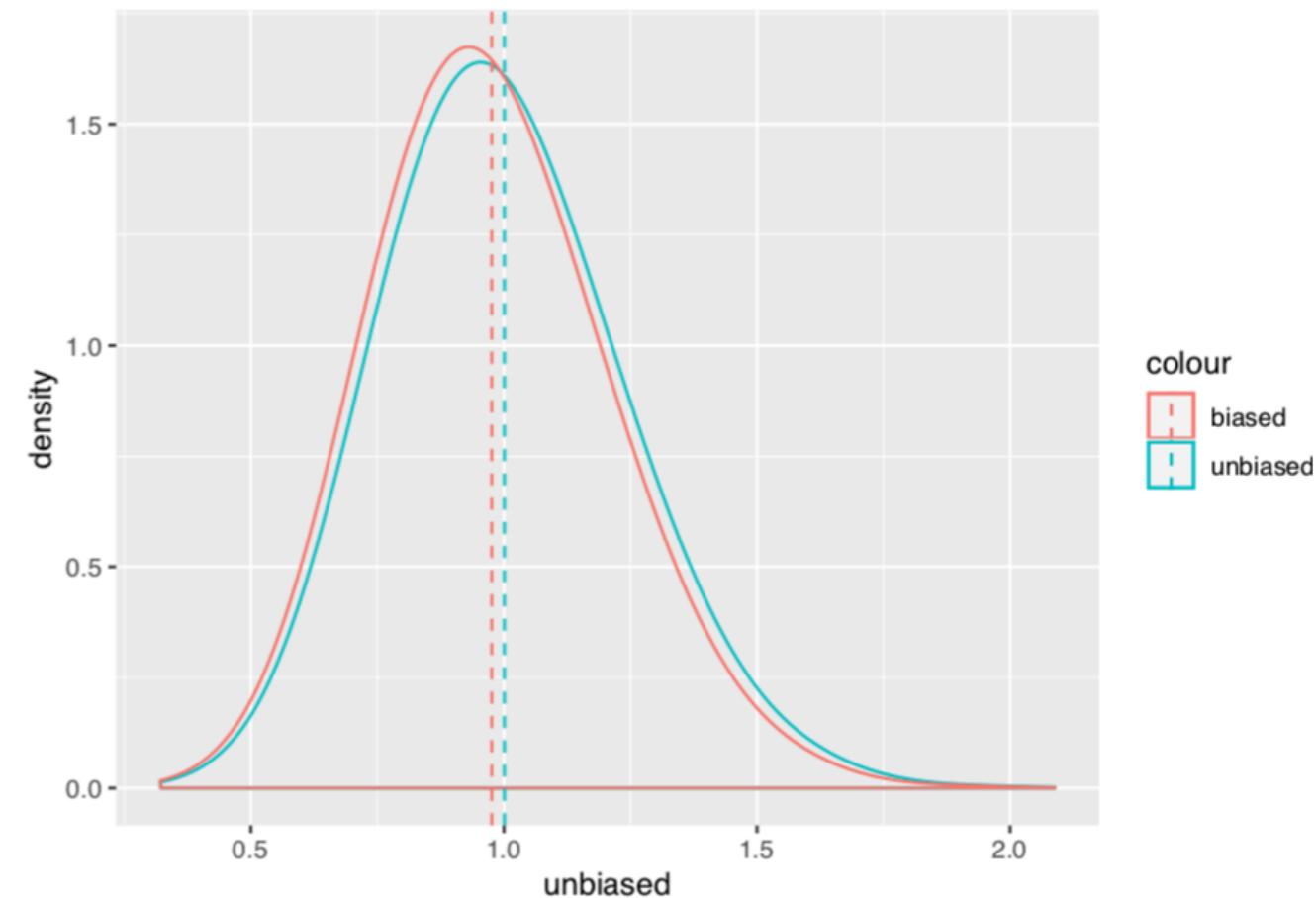
# **degrees of freedom**

# degrees of freedom

$$sd(Y) = \sqrt{\sum_{i=1}^n \frac{(Y_i - \bar{Y})^2}{n - 1}}$$

Why?

makes sure that the sample standard deviation is an unbiased estimator of the population standard deviation



```
# 7. Calculate which of the two measures  $s^2$  or  $s'^2$  is a better
#   estimator of the true population variance
sample_variances %>%
  summarize(error.unbiased = population_variance - mean(unbiased),
           error.biased = population_variance - mean(biased))

## # A tibble: 1 x 2
##   error.unbiased error.biased
##       <dbl>      <dbl>
## 1     -0.000786    0.0242
```

# degrees of freedom

$$sd(Y) = \sqrt{\sum_{i=1}^n \frac{(Y_i - \bar{Y})^2}{n - 1}}$$

estimating the population mean

In statistics, the number of degrees of freedom is the **number of values** in the final calculation of a statistic **that are free to vary**.

[https://en.wikipedia.org/wiki/Degrees\\_of\\_freedom\\_\(statistics\)](https://en.wikipedia.org/wiki/Degrees_of_freedom_(statistics))

# degrees of freedom

How many degrees of freedom does a **line** have?



# degrees of freedom

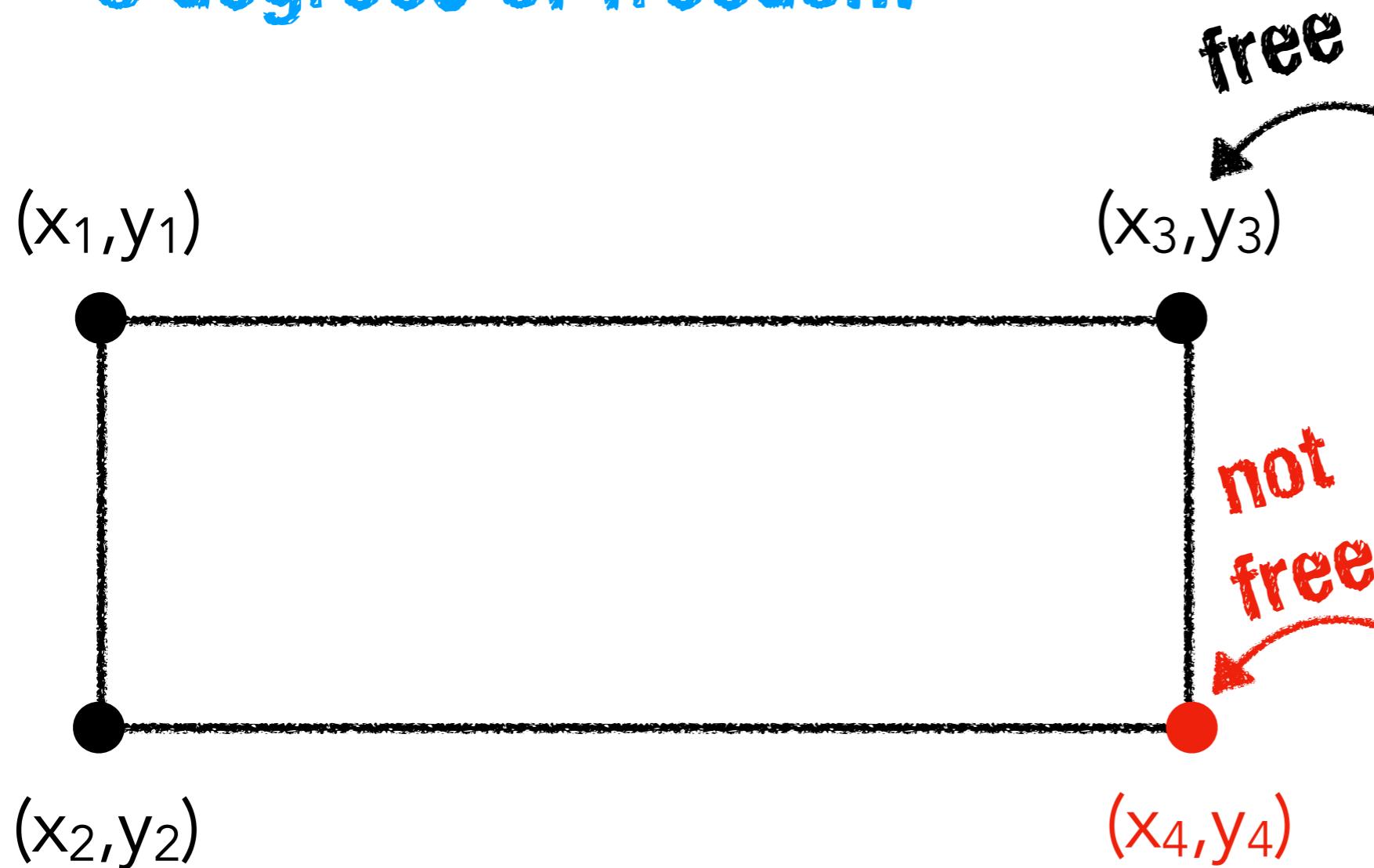
How many degrees of freedom does a **line** have?



# degrees of freedom

How many degrees of freedom does a **rectangle** have?

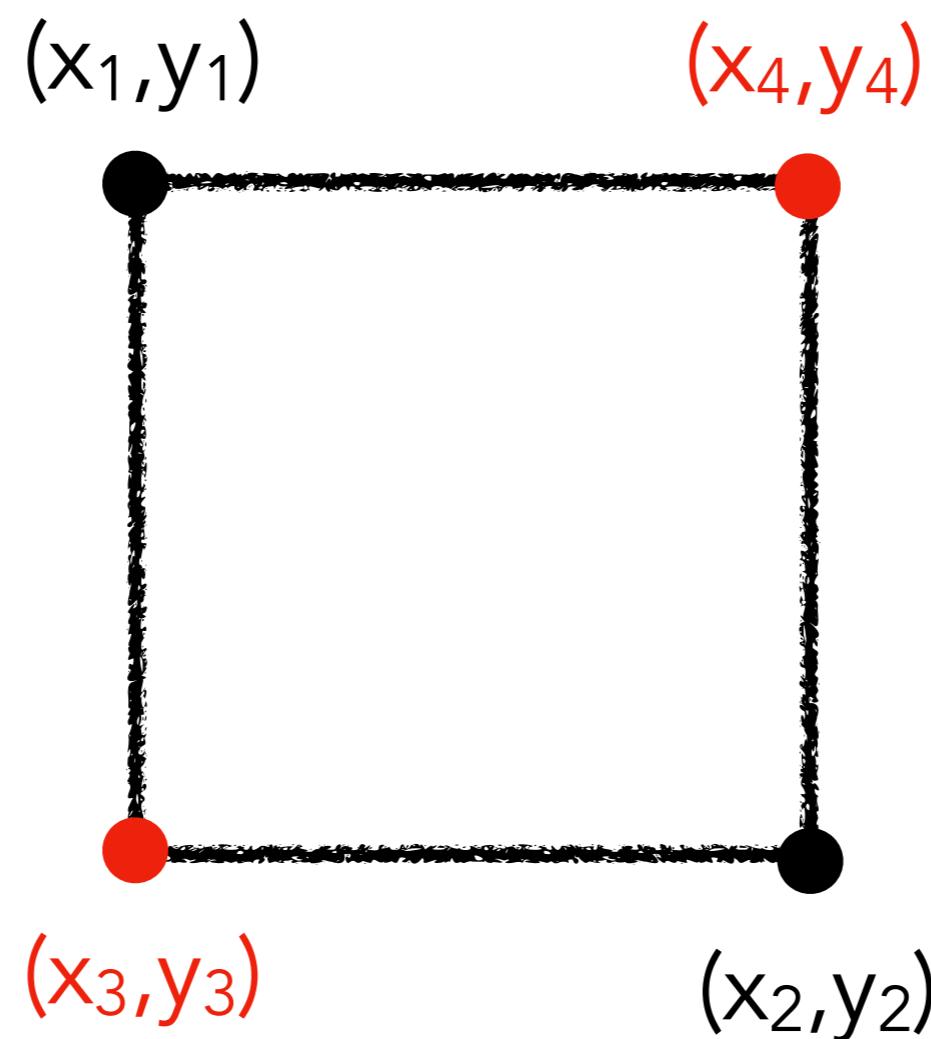
**3 degrees of freedom**



# degrees of freedom

How many degrees of freedom does a **square** have?

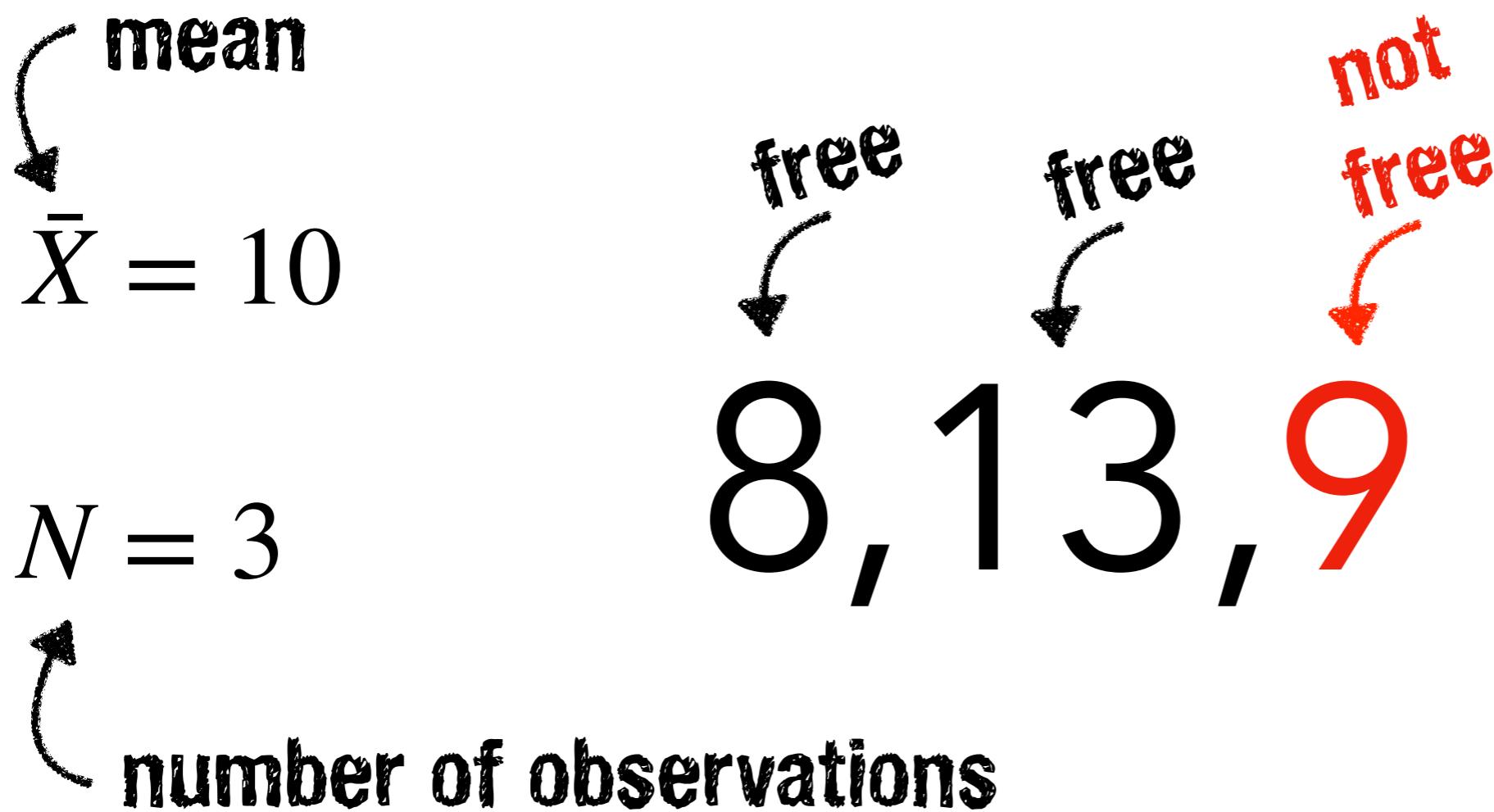
**2 degrees of freedom**



# degrees of freedom

How many degrees of freedom does a **mean** have?

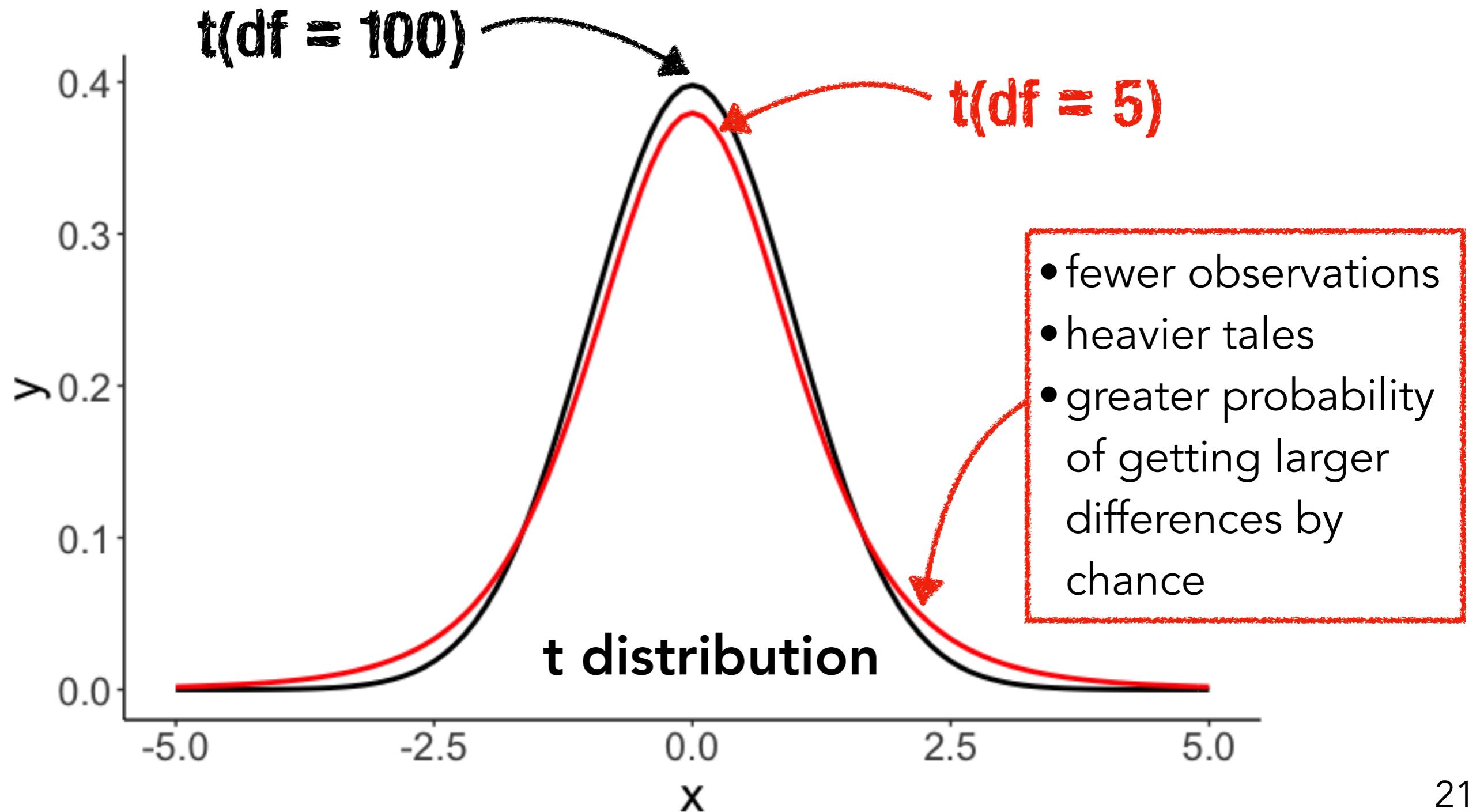
**the mean as an estimator has  $n - 1$  degrees of freedom**



# degrees of freedom

determine the shape of the sampling distributions for our test statistic

the degrees of freedom generally equals the number of observations (or pieces of information) minus the number of parameters estimated



# **t-test vs. permutation**

# t-test vs. permutation test

The screenshot shows a Reddit post in the [r/statistics](#) subreddit. The post is titled "Permutation Test vs. T Test". The original poster (OP) is u/[deleted] and it was posted 2 years ago. The post has 5 upvotes. The OP's comment reads: "Is there ever a time when a T Test would be preferable to a permutation test, all other things being equal? Permutation seems like the superior option in most cases." Below the post are options to "Share", "Save", "Hide", and "Report". A note indicates "This thread is archived" and "New comments cannot be posted and votes cannot be cast". The post is sorted by "BEST". A reply from user "manic\_panic" states: "I'm not sure how much detail you are looking for in terms of an answer. I think you are correct, an 'exact'/permutation test is likely to be the better option, all things being equal. A t-test is an approximation of the data assuming the distributional assumptions are met. An exact test will..." To the right of the post, the sidebar for the r/statistics subreddit shows 94.7k members, 158 online statisticians, and a creation date of Mar 13, 2008. There is also a "JOIN" button.

↑ Case\_Control 2 points · 2 years ago  
↓ In practice, its primarily computing time. When appropriate, I can run a t-test which gets me 99.99% of the way there almost instantly or I can wait forever for a permutation test to run across a million records.

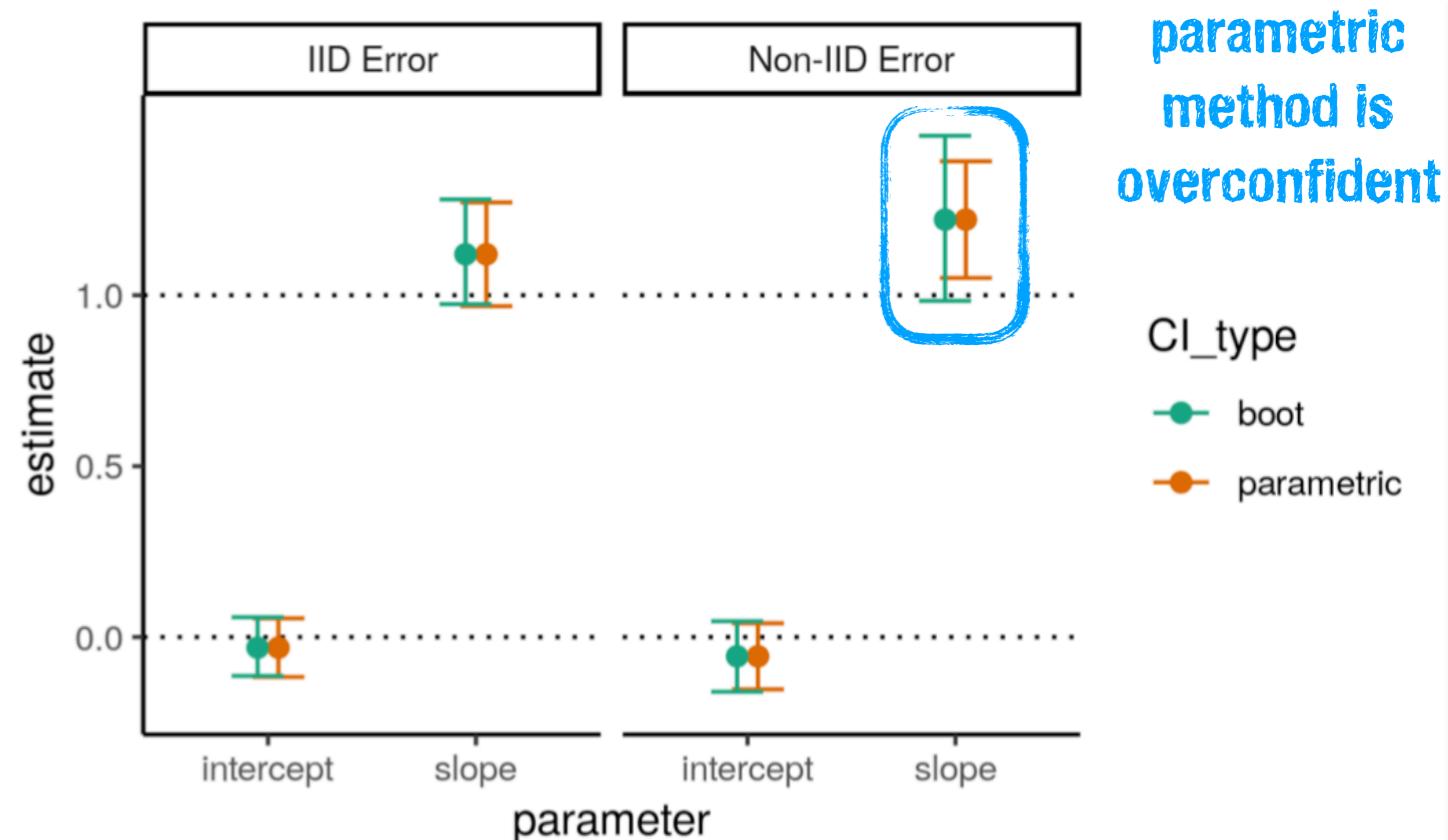
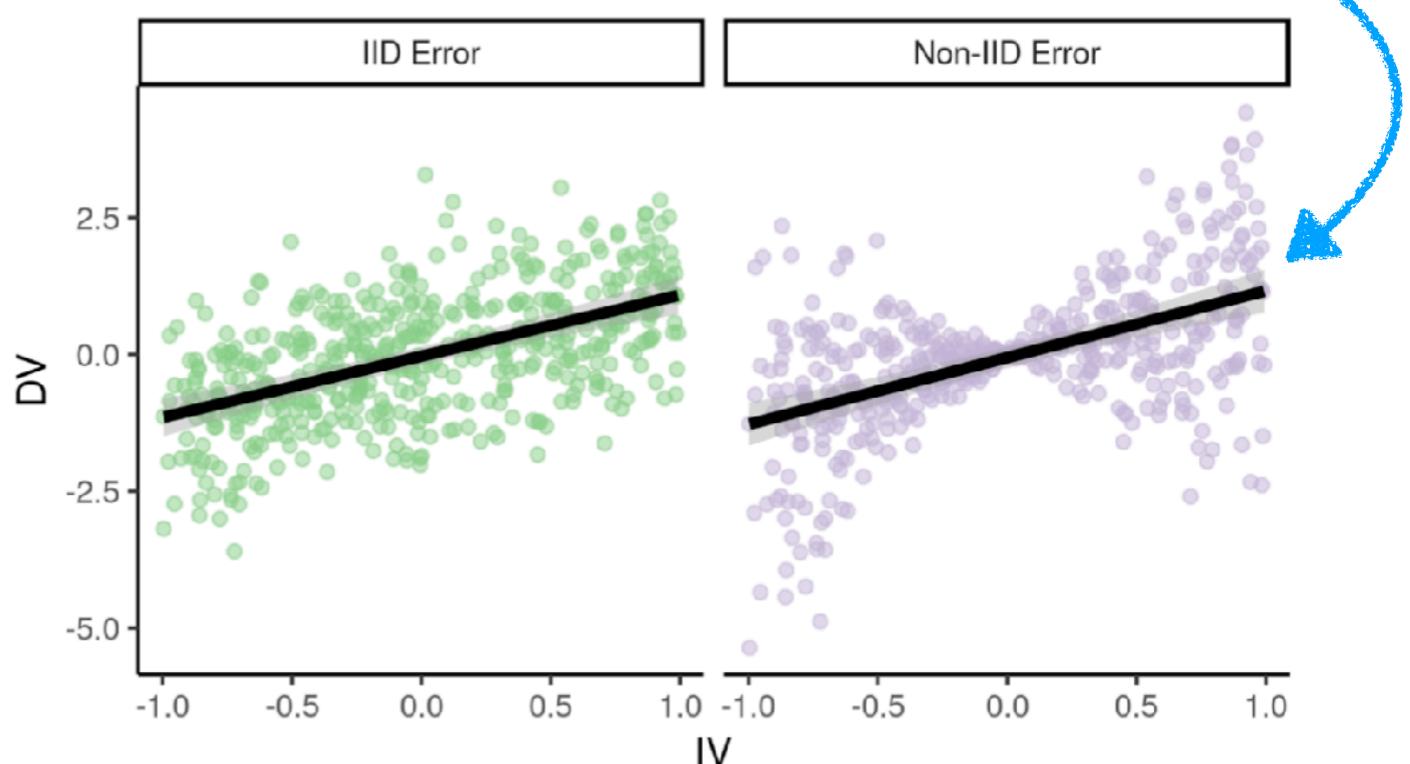
# t-test vs. bootstrap

heteroscedasticity

**t-test assumptions:**

normally distributed residuals

equality of variance

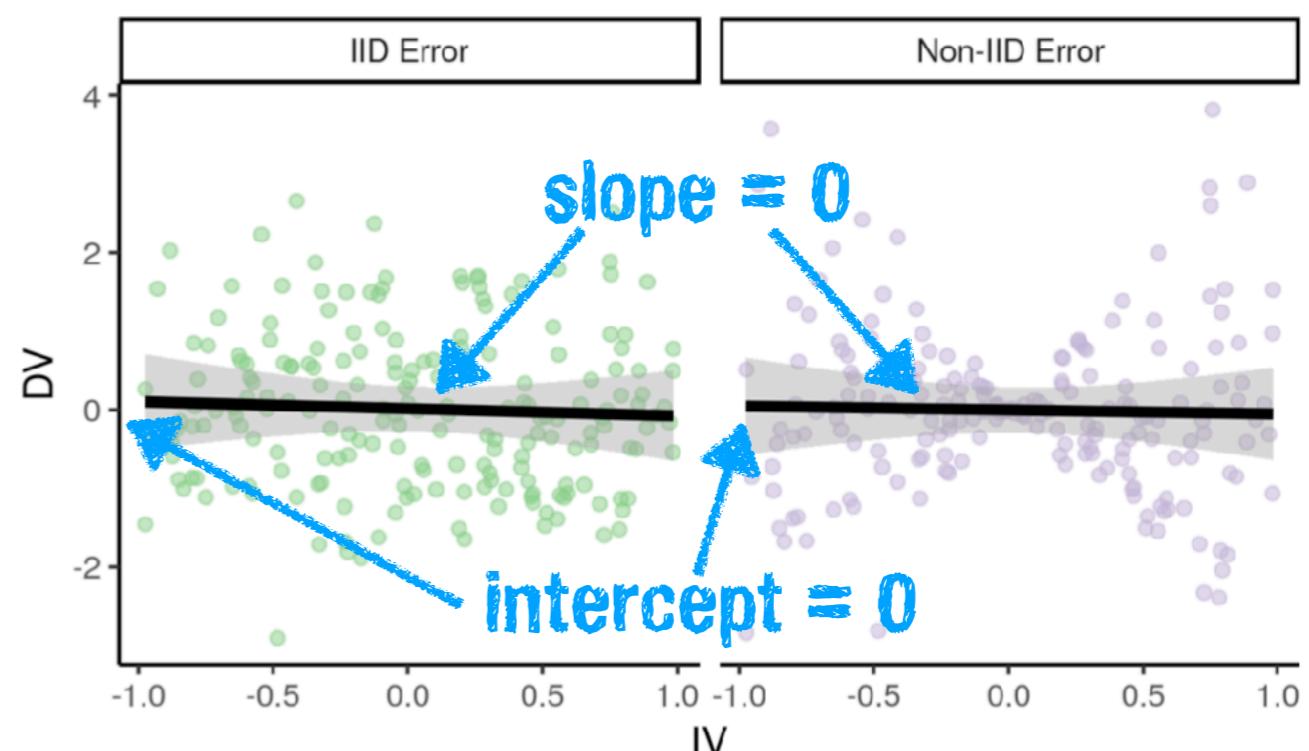


# t-test vs. bootstrap

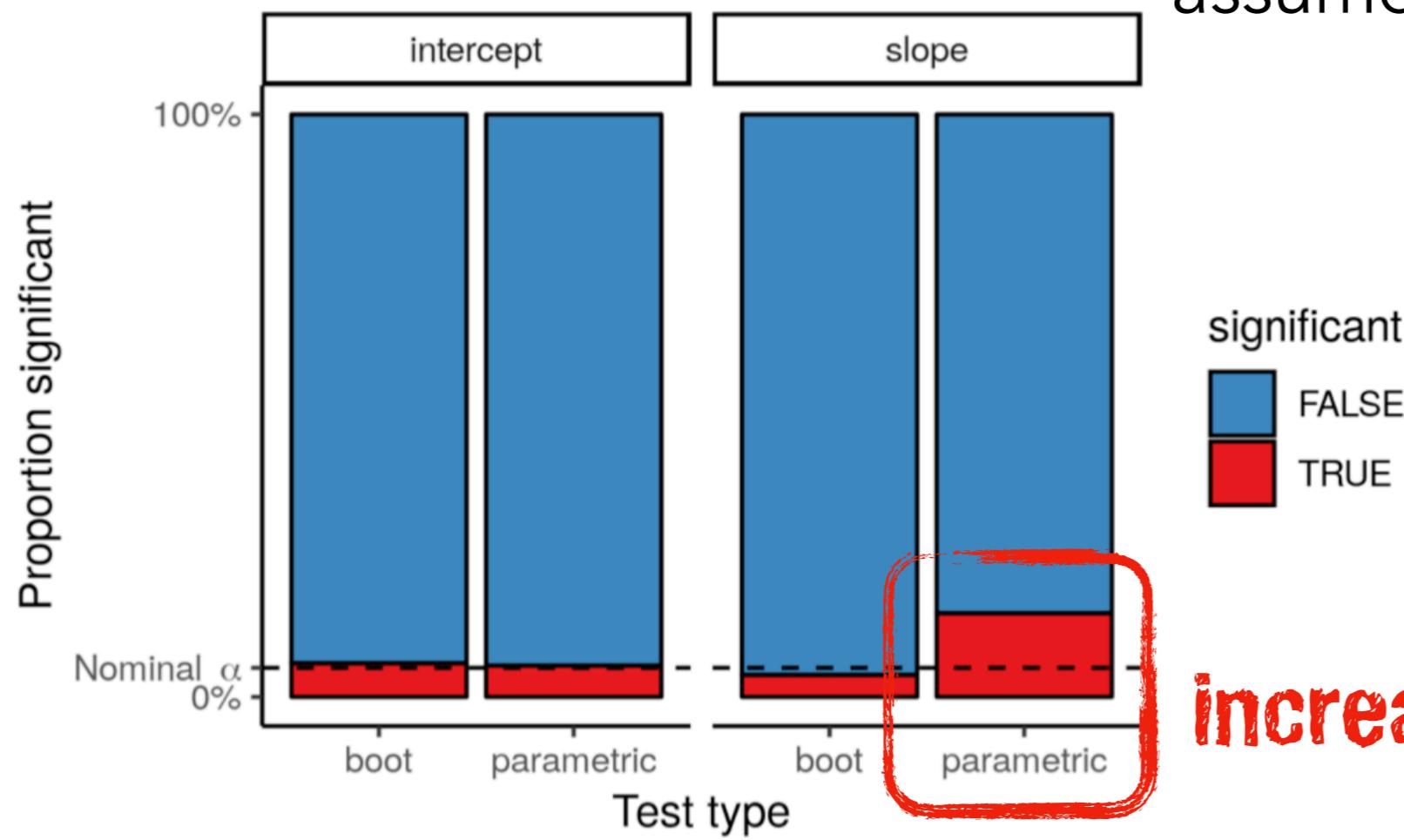
## t-test assumptions:

normally distributed residuals

equality of variance



assume null hypothesis is true



increased type I error!

# t-test vs. bootstrap

**t-test assumptions:**

normally distributed residuals  
equality of variance



Andrew MacDonald 🌈  
@polesasunder

Just saw this in a manuscript. This is a very common practice and it is incorrect. Don't do it!

## Statistical analysis

Using the method described by Shapiro and Wilk (1965), we determined that our response variable (individual infection intensities, Z-standardized based on mean and 2 SDs) did not conform to a normal distribution (at  $P < 0.05$ ; Fig 2). Thus, we

8:52 AM · Feb 5, 2020 · TweetDeck

63 Retweets 210 Likes



Andrew MacDonald 🌈 @polesasunder · Feb 5

Replying to @polesasunder

This is the wrong thing to do because what matters is how your "response variable" is distributed \*AFTER\* you do your model, not before!  
@rlmcelreath called this practice "histomancy", which name I love



2



9



130



Andrew MacDonald 🌈 @polesasunder · Feb 5

If you must do such things, do them to your residuals not your observations



11



1

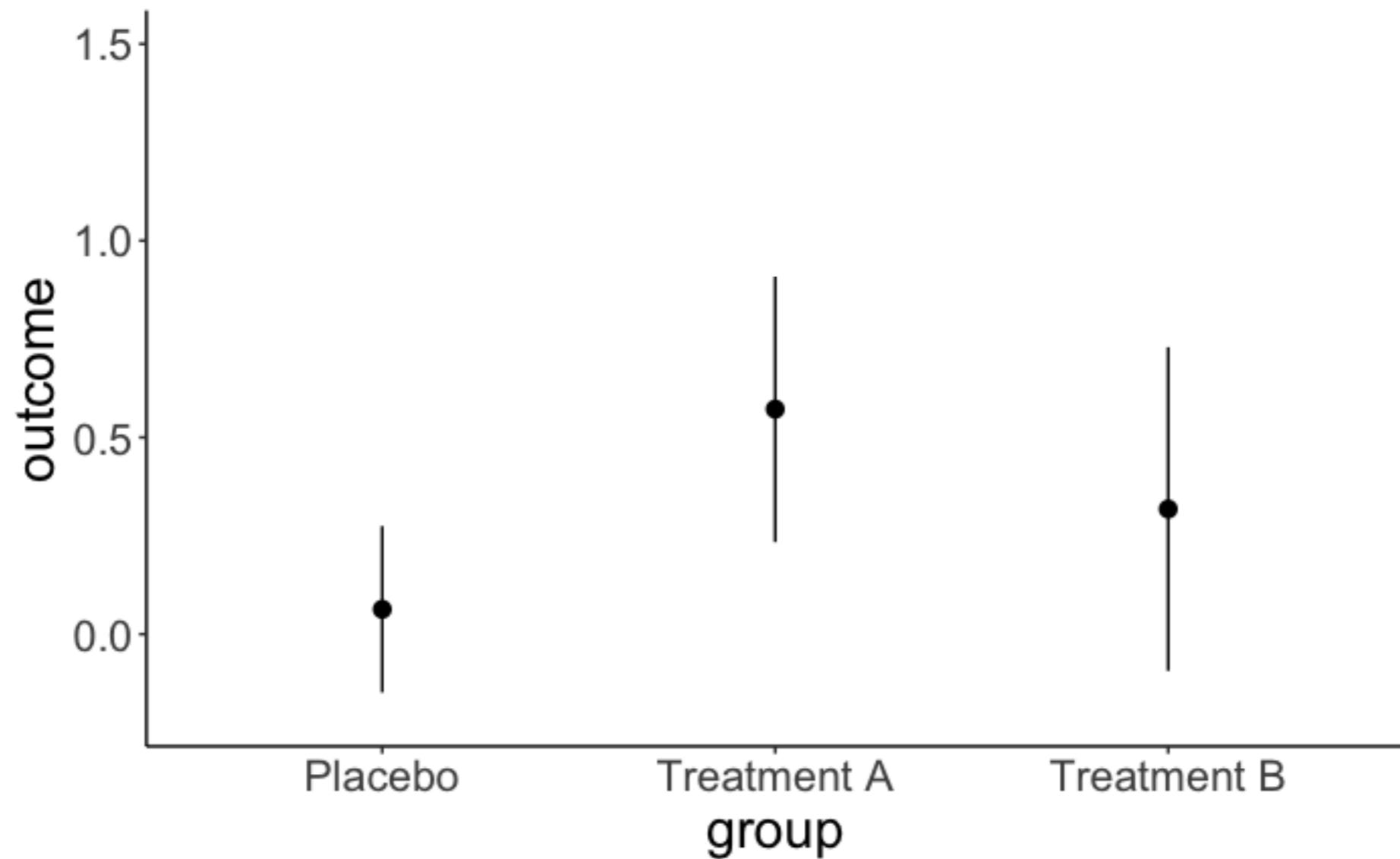


83



**difference in significance  
vs. significant differences**

# significant differences

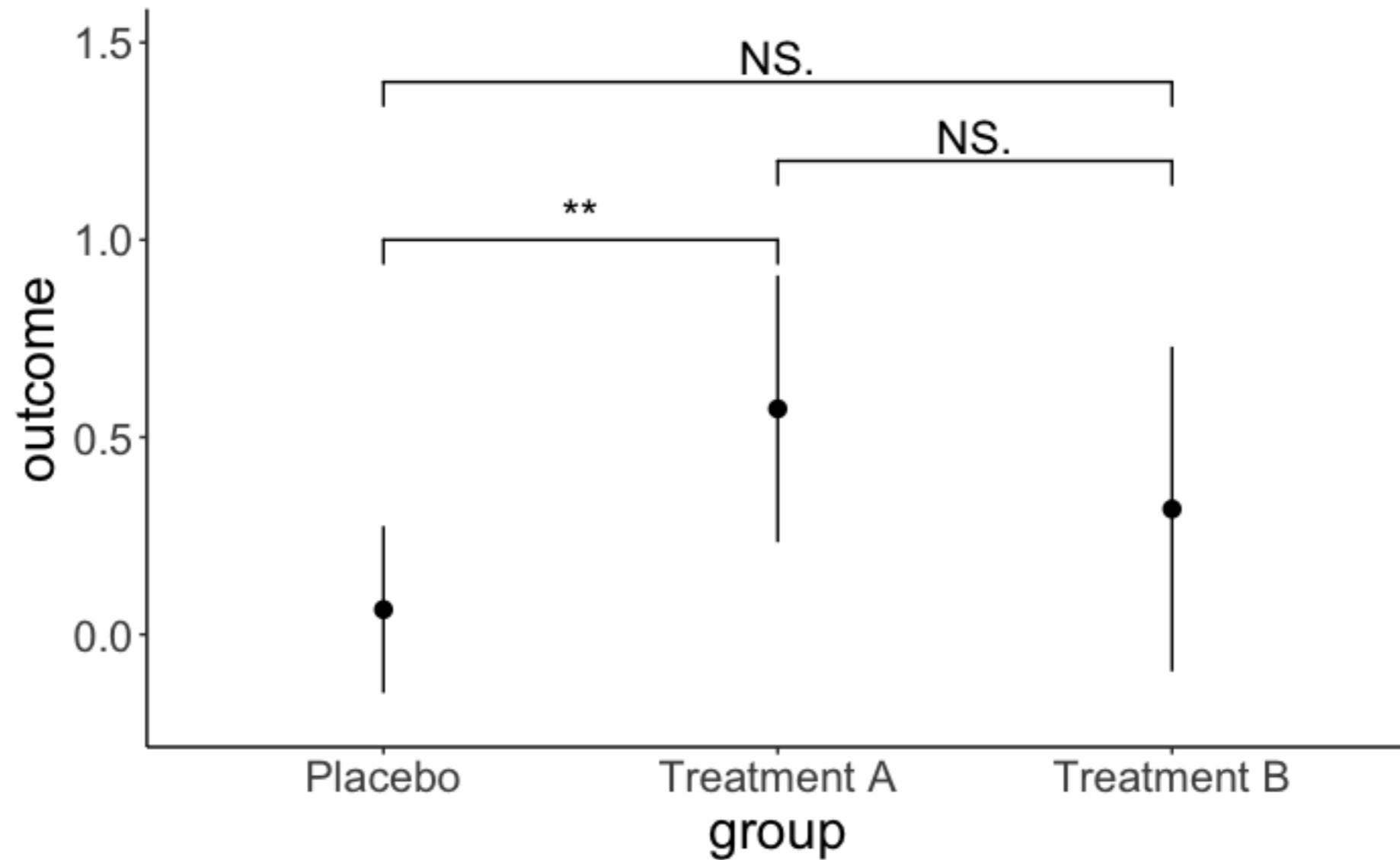


"We compared treatments A and B with a placebo. Treatment A showed a significant benefit over placebo, while treatment B had no statistically significant benefit. Therefore, treatment A is better than treatment B."

<https://www.statisticsonewrong.com/significant-differences.html>

# significant differences

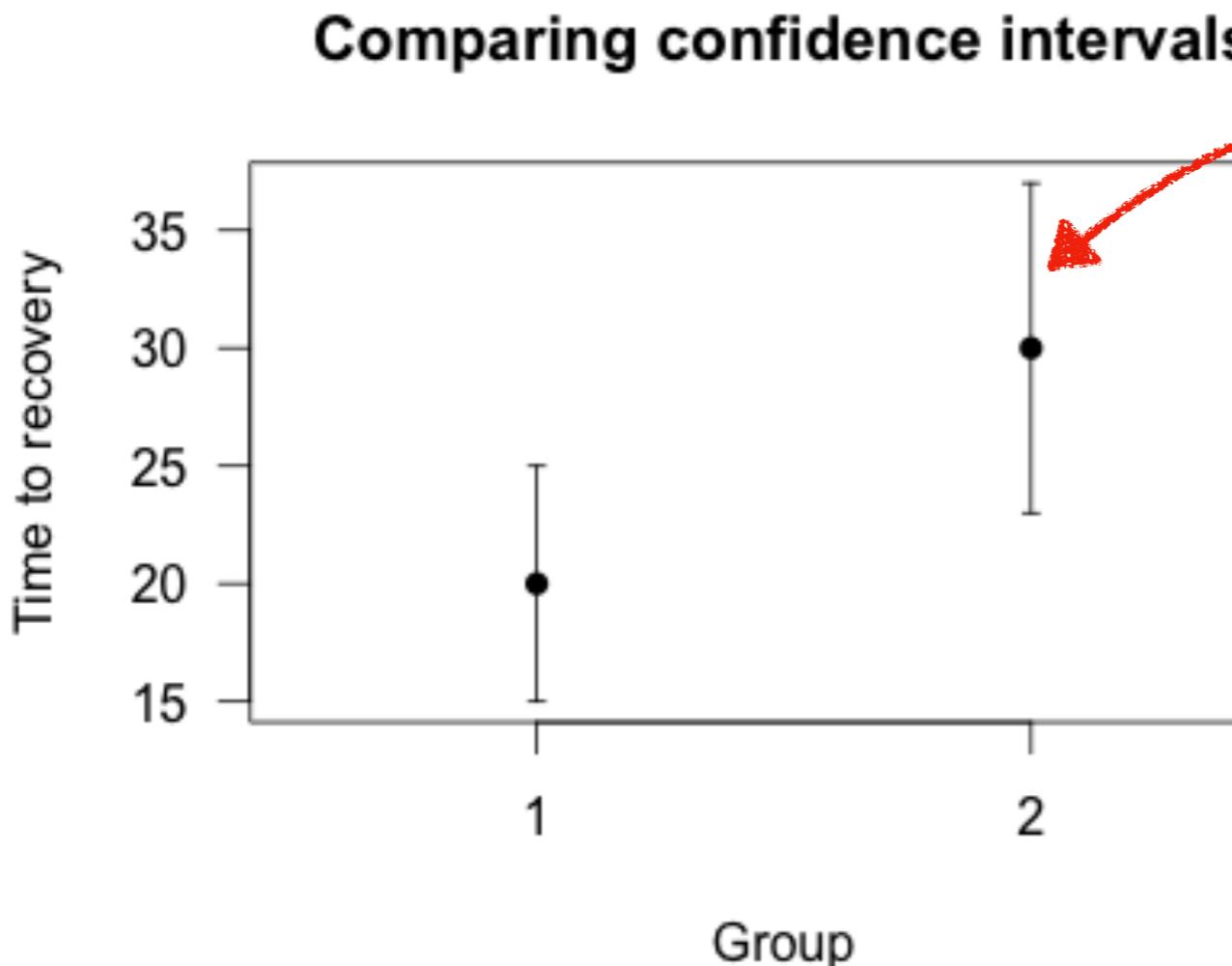
"Our tendency to look for a difference in significance should be replaced by a check for the significance of the difference."



"We compared treatments A and B with a placebo. Treatment A showed a significant benefit over placebo, while treatment B had no statistically significant benefit. Therefore, treatment A is better than treatment B."

# significant differences

## Significant difference between Group 1 and 2?



what do the  
error bars mean?

standard deviation?

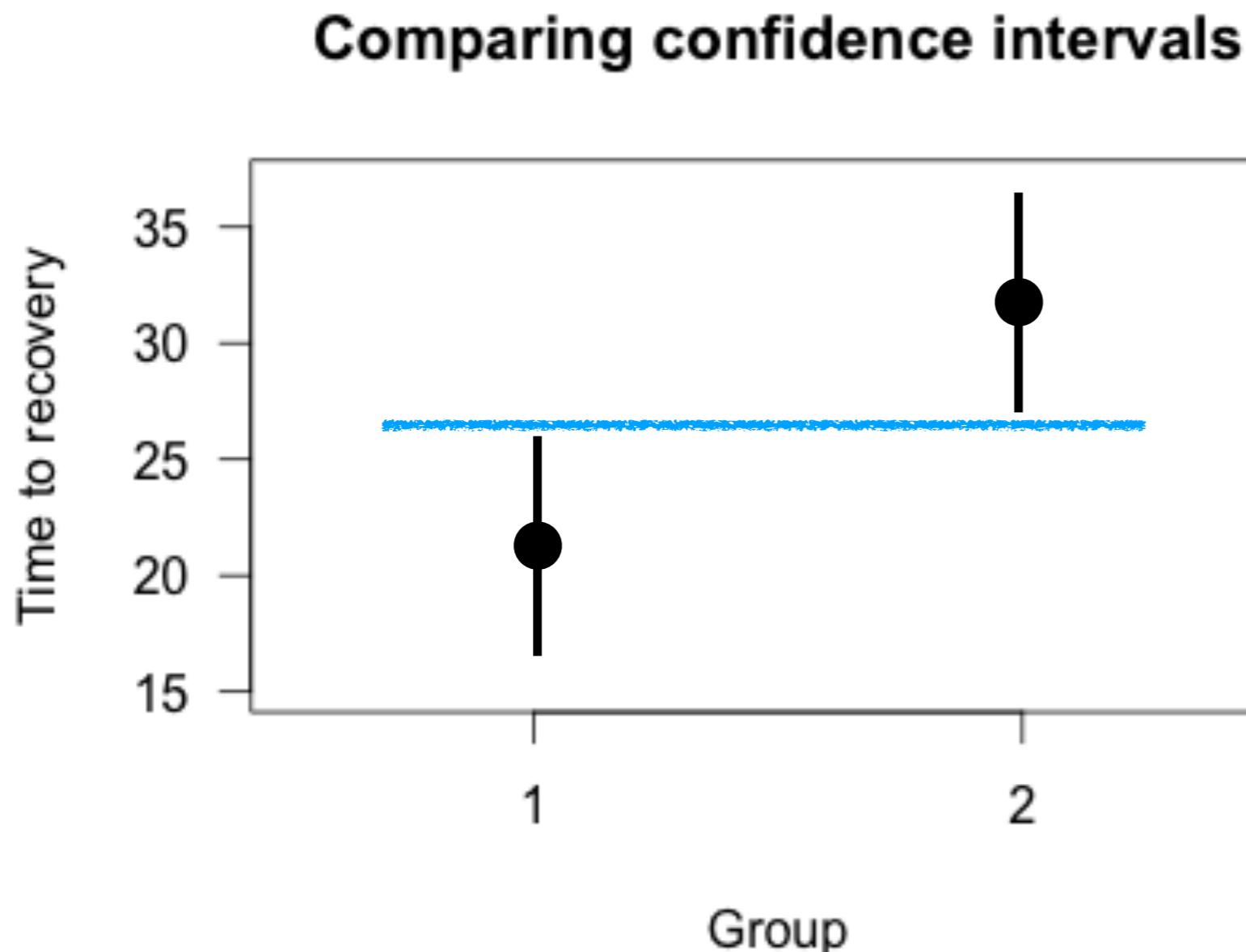
standard error of the mean?

confidence intervals?

What we would like to know, does the confidence interval of the difference between groups exclude 0?

# significant differences

## Significant difference between Group 1 and 2?



95% confidence  
intervals don't overlap

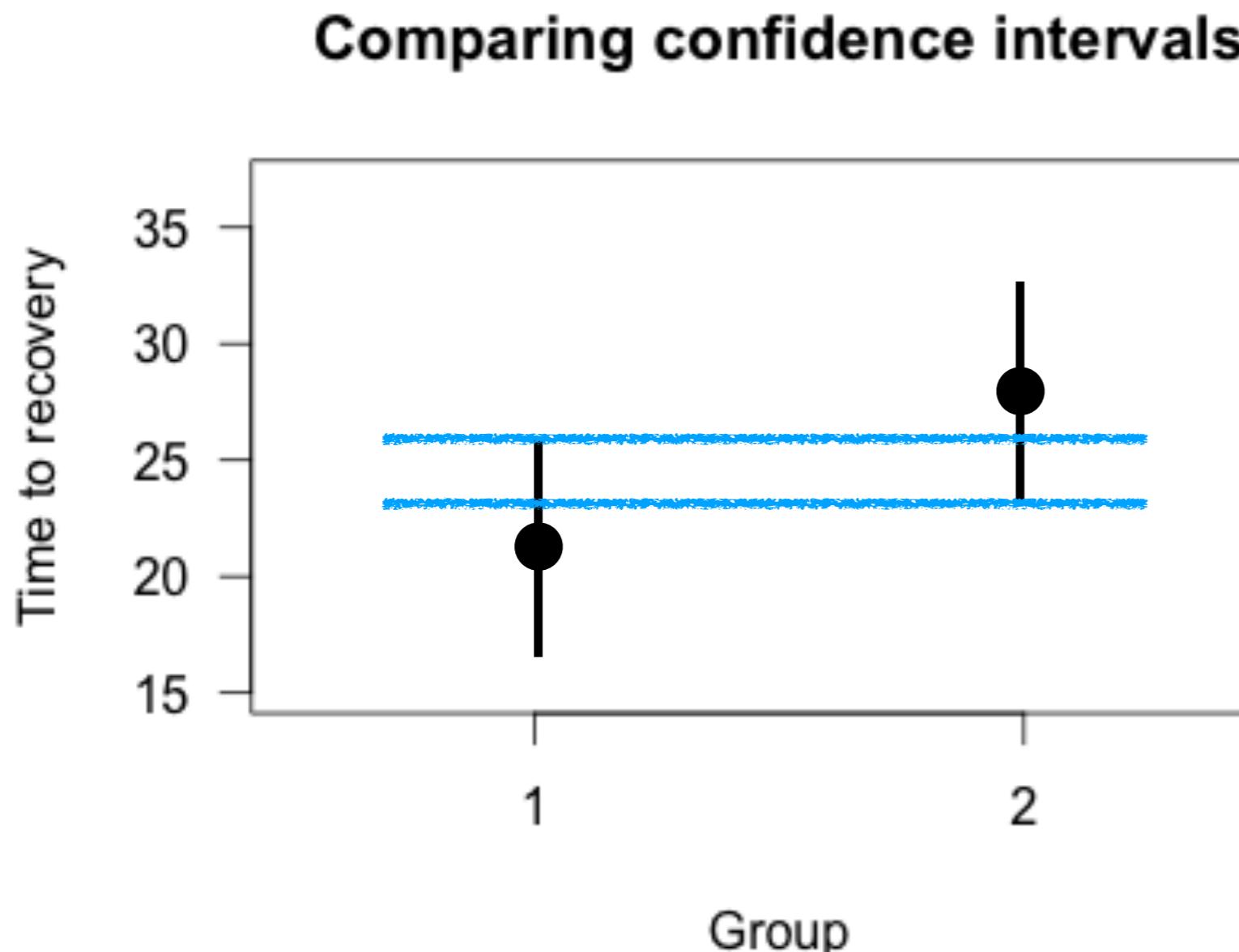
conservative test  
of significance

rejects  $H_0$  less often than the appropriate statistical procedure

Schenker & Gentleman (2001) On judging the significance of differences by examining the overlap between confidence intervals. *The American Statistician*

# significant differences

## Significant difference between Group 1 and 2?



95% confidence  
intervals don't overlap  
with mean

anti-conservative  
test of significance

rejects  $H_0$  more often than the appropriate statistical procedure

Schenker & Gentleman (2001) On judging the significance of differences by examining the overlap between confidence intervals. *The American Statistician*

# **Power analysis (continued)**



Studio<sup>®</sup>

time

# **Controlling for variables**

# Controlling for variables

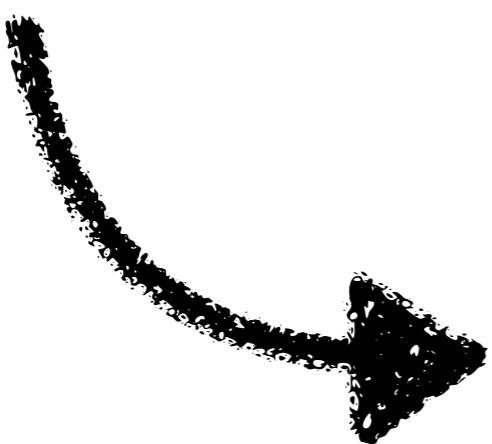
1. What does "controlling" for variables mean?
2. When should I control for variables?

check this out

# What does controlling for variables mean?

we are not actually "**controlling**" the variable

instead, we are taking the variable into consideration when making predictions



**the hope is that we get a better estimate of the parameter that we are interested in by taking into account other factors**

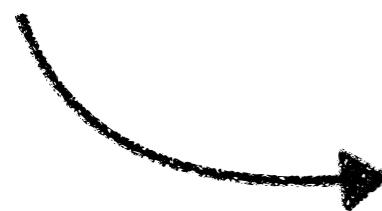
# When should I control for variables?

recent advances in graphical models have produced a simple criterion to distinguish good from bad controls

 **d-separation**  
**directional**

decide from a causal graph whether a set of variables  $X$  is independent of another set  $Y$ , given a third set  $Z$

**Goal:** we want a precise (and unbiased) estimate of the predictive relationship between  $X$  and  $Y$

 **we want to block all other paths from  $X$  to  $Y$**

# When should I control for variables?

## How can I tell whether two variables are independent?

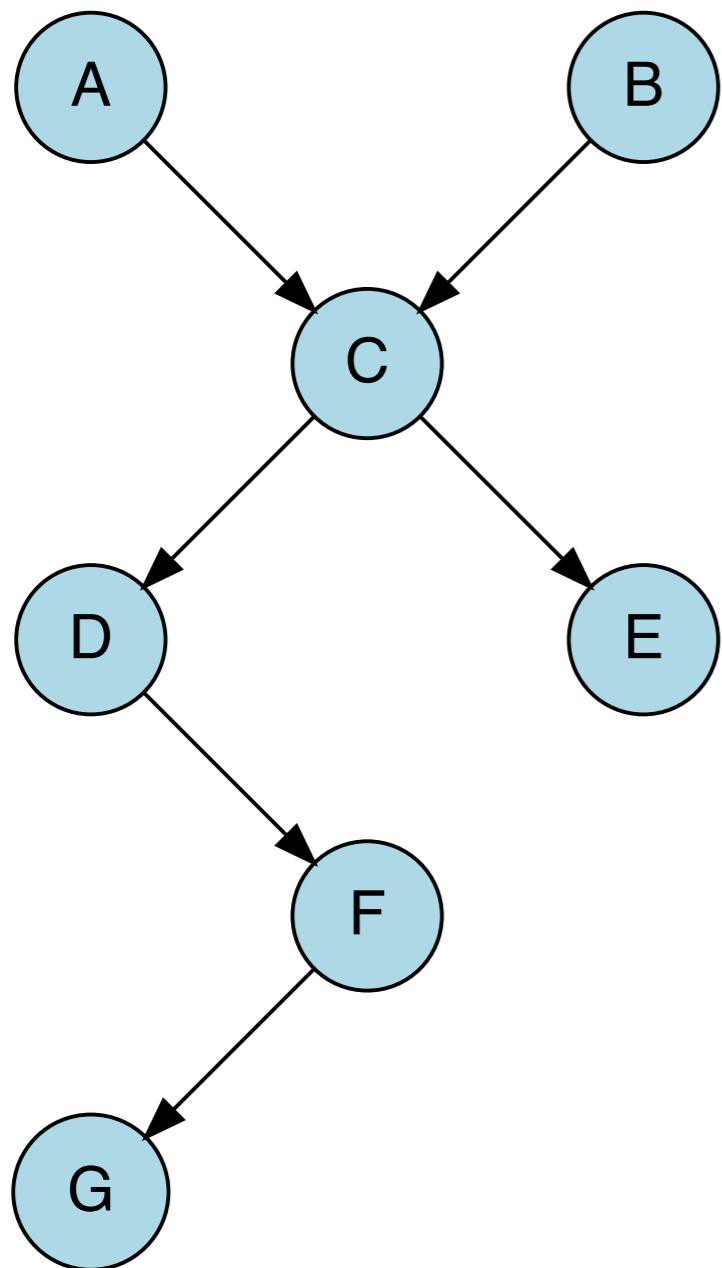
### Recipe for independence

1. Draw the ancestral graph
  2. "Moralize" the graph by "marrying" the parents
  3. "Disorient" the graph by replacing arrows with edges
  4. Delete the givens and their edges
  5. Read the answer off the graph
- if variables are **disconnected** they are independent  
- if variables are connected (have a path between them)  
they are not guaranteed to be independent

# When should I control for variables?

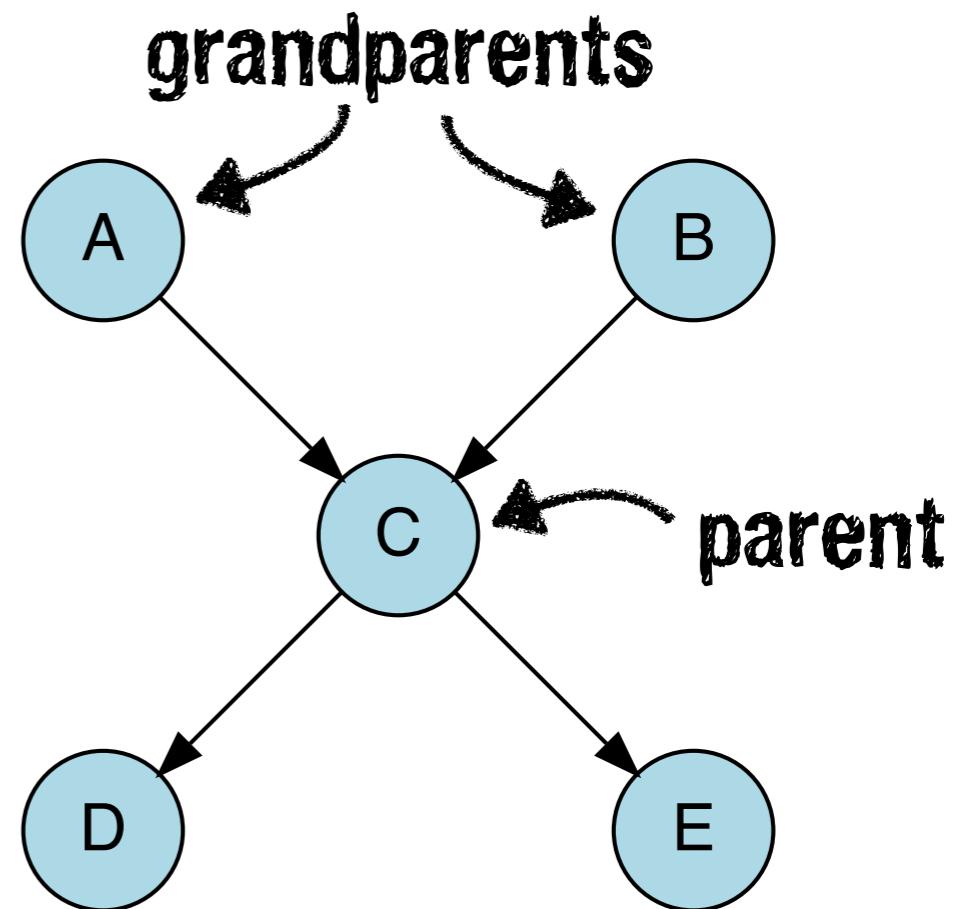
**Are D and E independent?**

$$p(D | E) = p(D) ?$$



## 1. Draw the ancestral graph

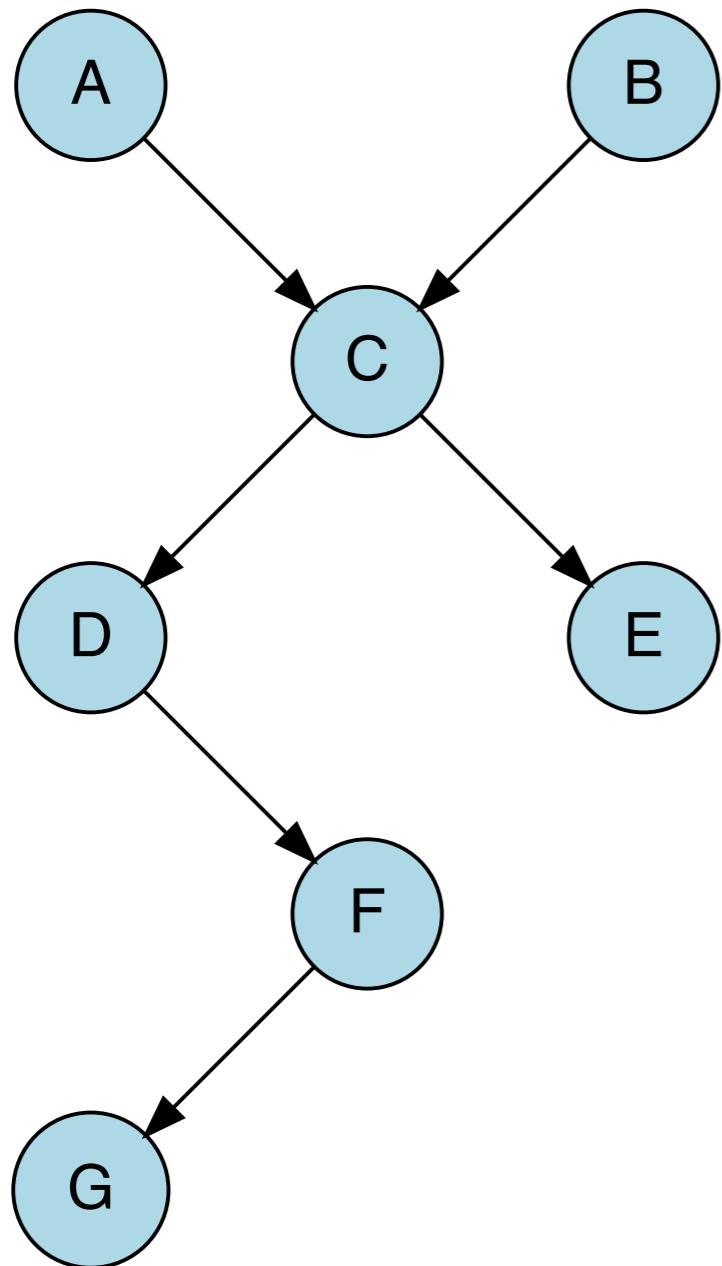
Construct the "ancestral graph" of all variables mentioned in the probability expression. This is a reduced version of the original net, consisting only of the variables mentioned and all of their ancestors (parents, parents' parents, etc.)



# When should I control for variables?

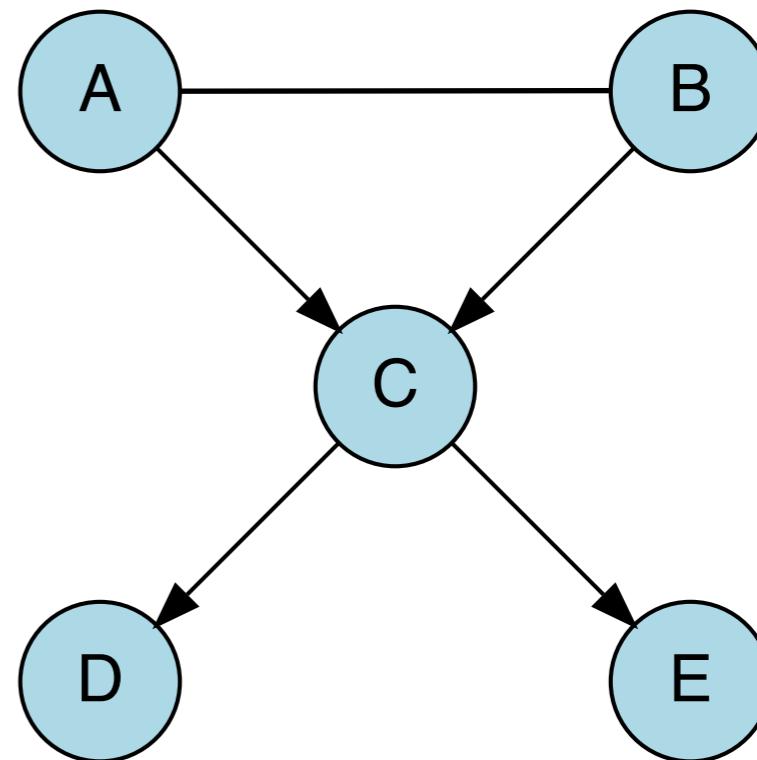
**Are D and E independent?**

$$p(D | E) = p(D) ?$$



**2. "Moralize" the graph**  
**let's get married!**

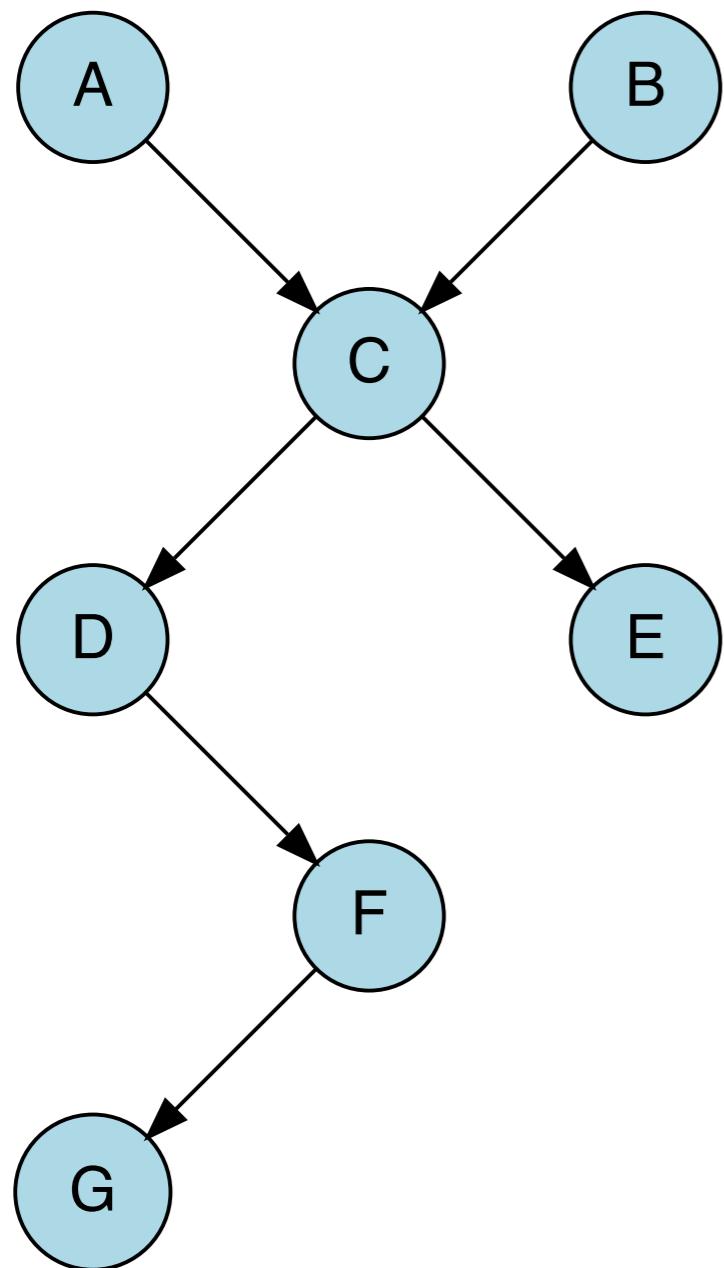
For each pair of variables with a common child, draw an undirected edge (line) between them. (If a variable has more than two parents, draw lines between every pair of parents.)



# When should I control for variables?

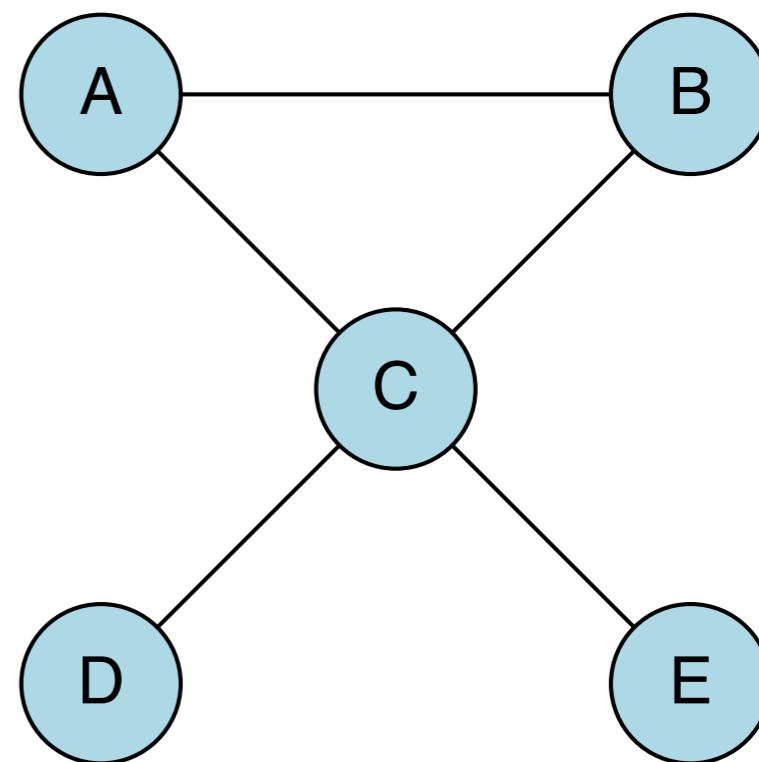
**Are D and E independent?**

$$p(D | E) = p(D) ?$$



**3. "Disorient" the graph**

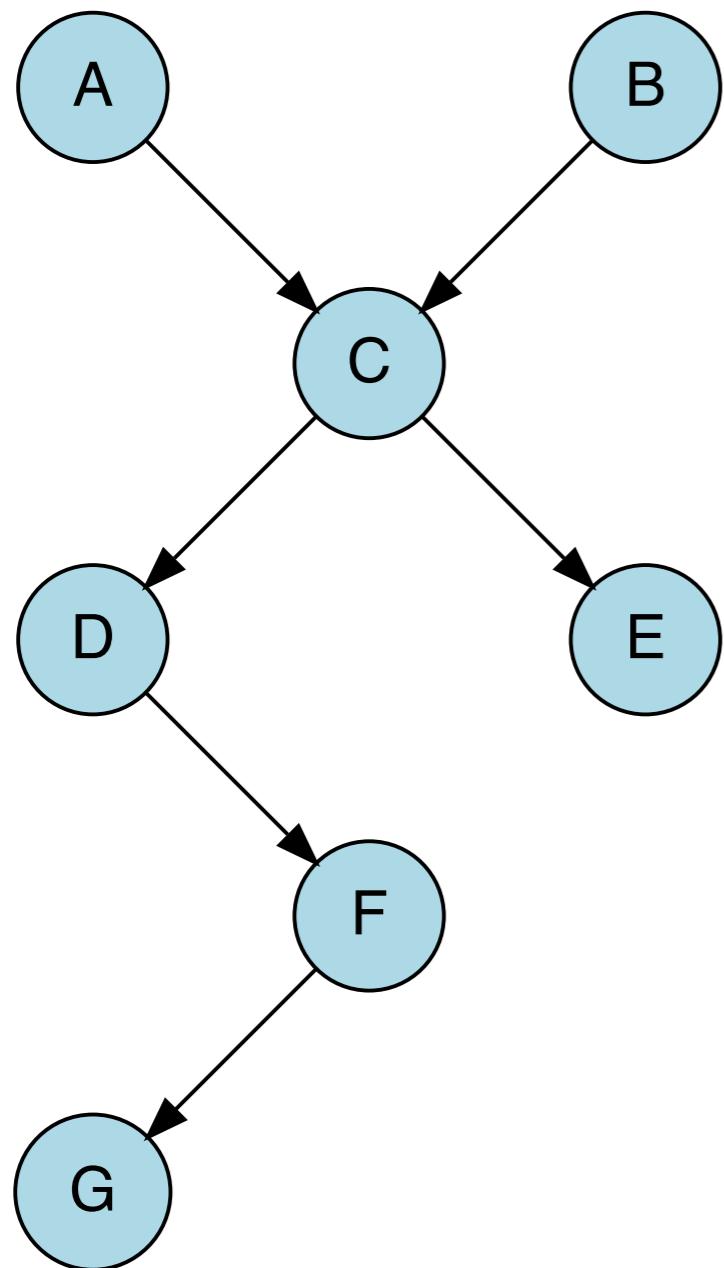
Replace arrows with lines



# When should I control for variables?

**Are D and E independent?**

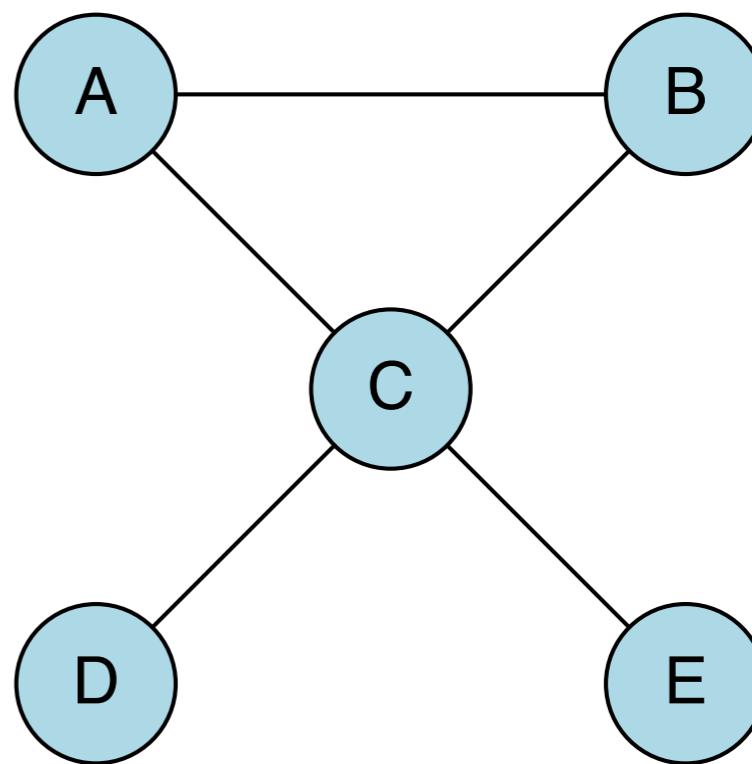
$$p(D | E) = p(D) ?$$



## 4. Delete the givens

Remove the variables that we condition on, as well as their edges

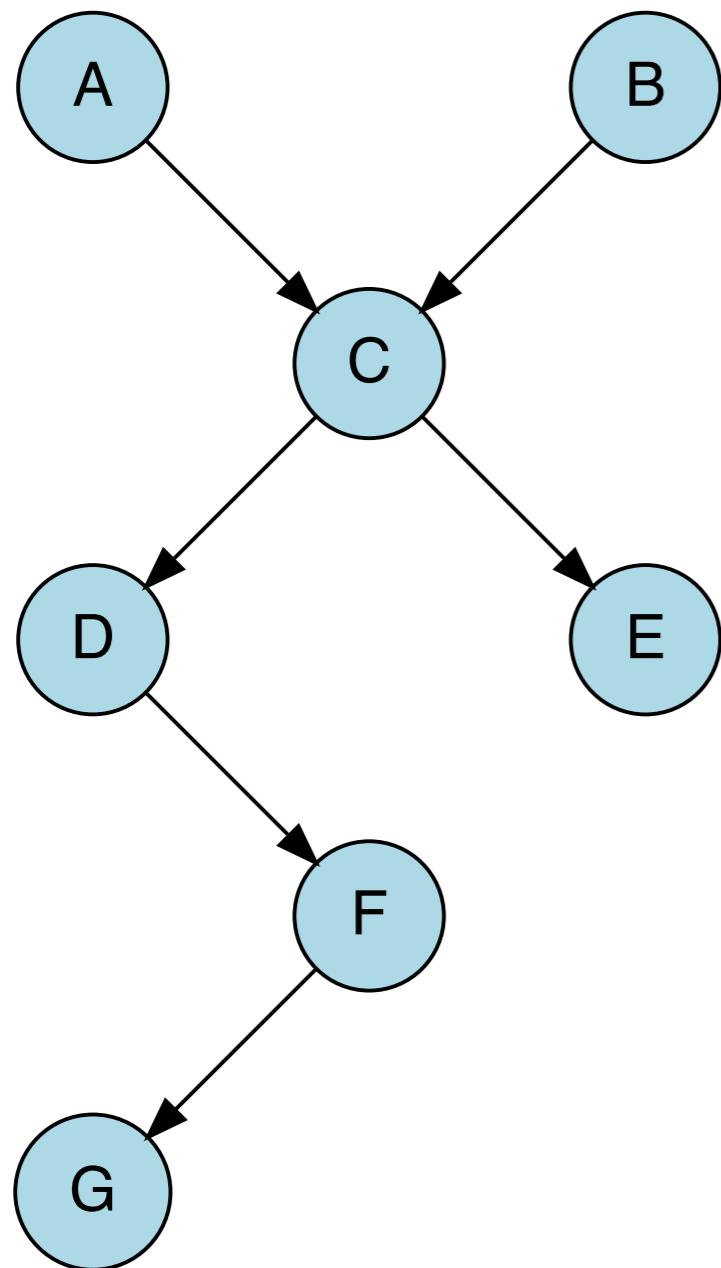
**we didn't condition on anything,  
so there is nothing to delete**



# When should I control for variables?

**Are D and E independent?**

$$p(D | E) = p(D) ?$$



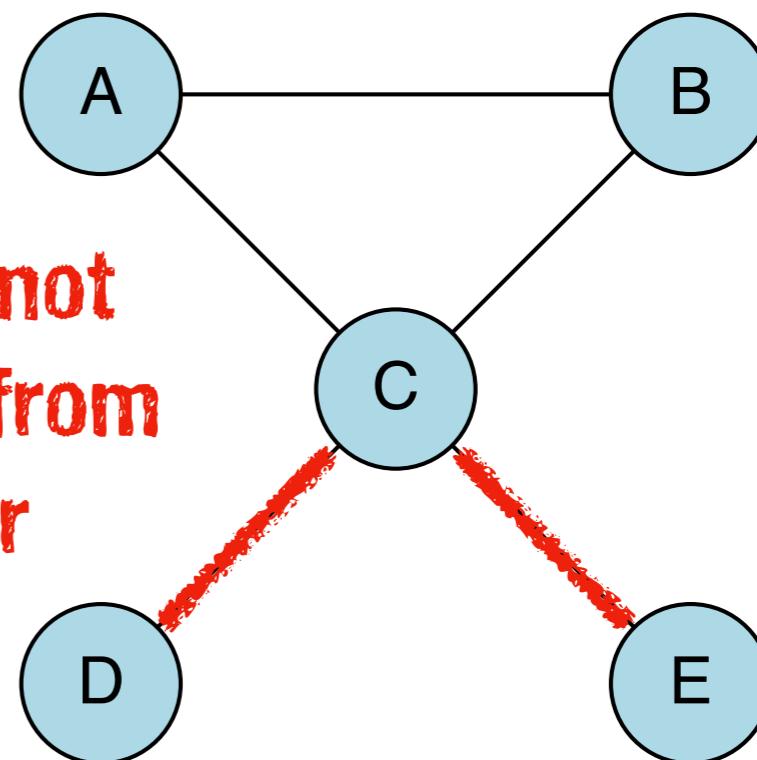
## 5. Read answer off the graph

- if variables are **disconnected** they are independent
- if variables are connected (have a path between them) they are not guaranteed to be independent

D and E are not independent from each other



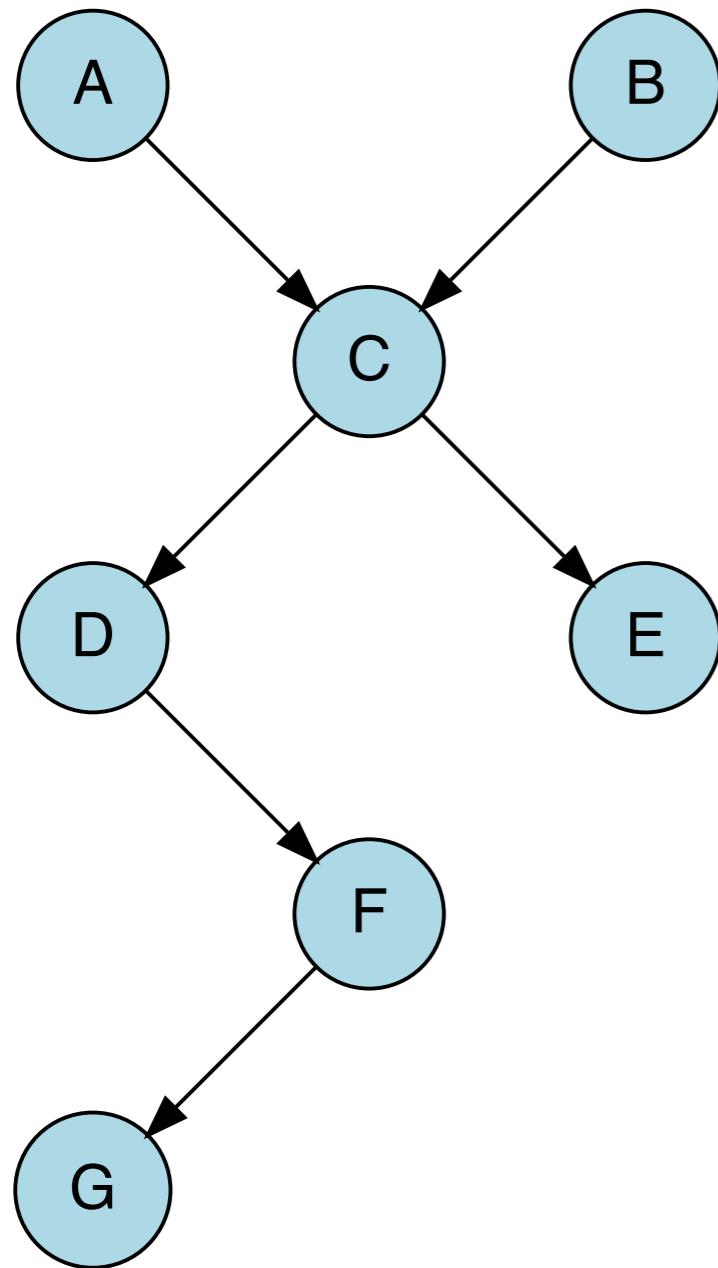
they are connected via at least one path



# When should I control for variables?

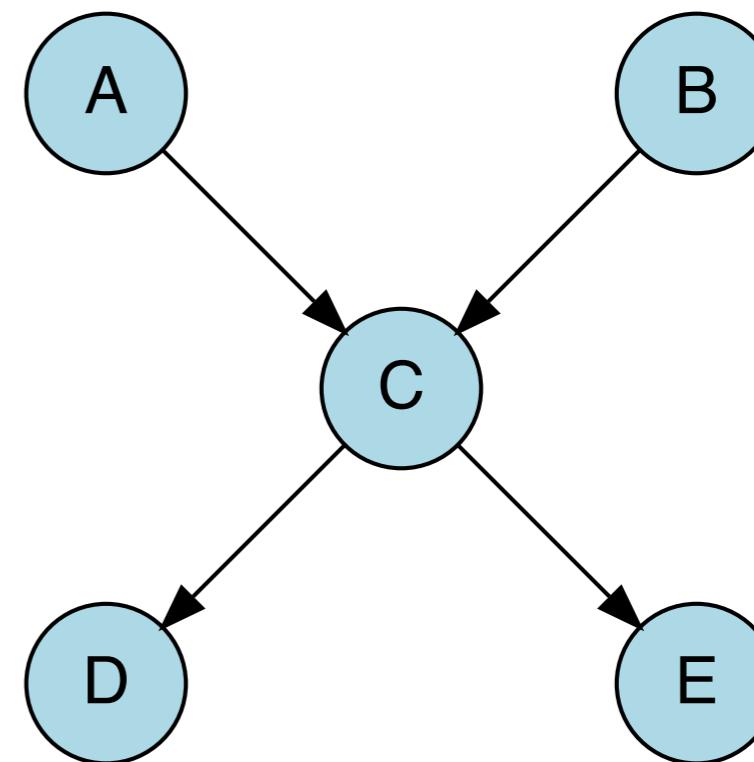
Are D and E independent, given C?

$$p(D | E, C) = p(D | C) ?$$



## 1. Draw the ancestral graph

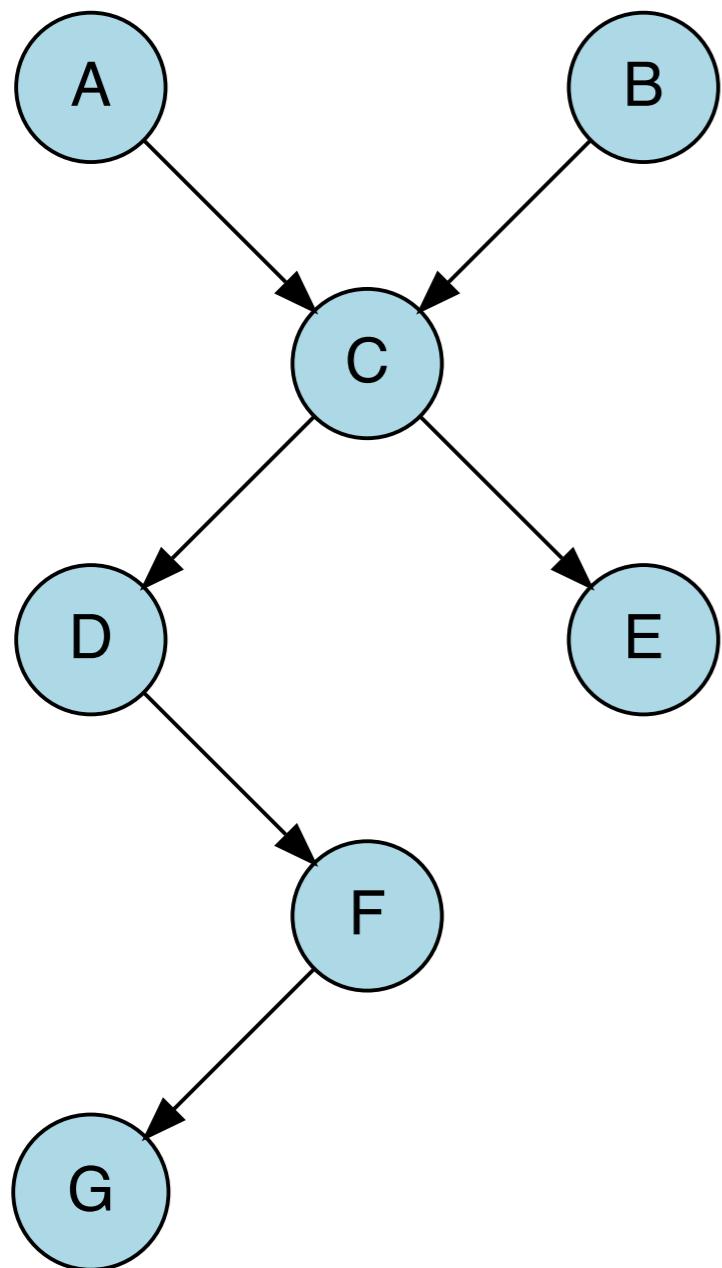
Construct the "ancestral graph" of all variables mentioned in the probability expression. This is a reduced version of the original net, consisting only of the variables mentioned and all of their ancestors (parents, parents' parents, etc.)



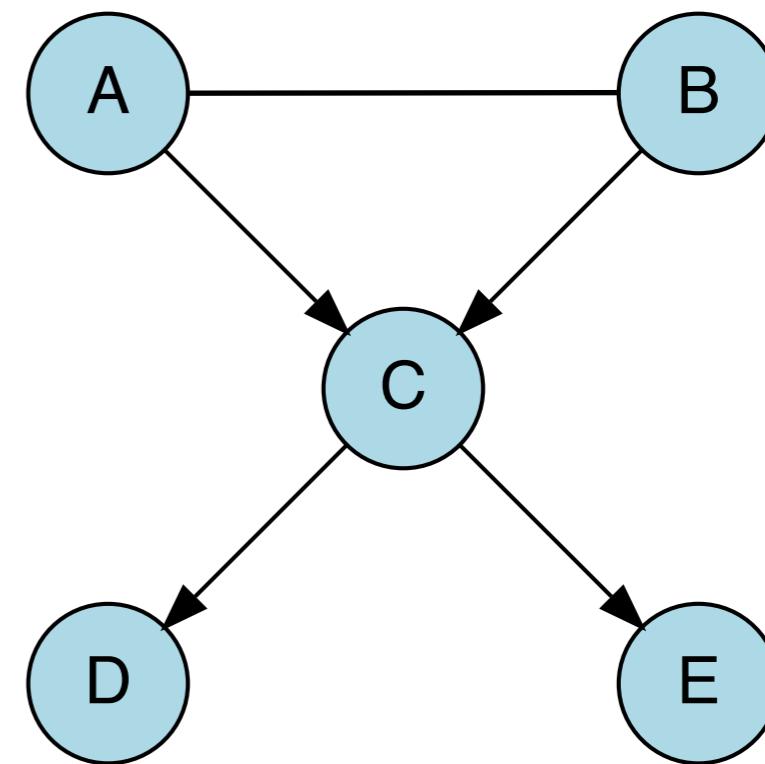
# When should I control for variables?

**Are D and E independent, given C? 2. "Moralize" the graph**

$$p(D | E, C) = p(D | C) ?$$



For each pair of variables with a common child, draw an undirected edge (line) between them. (If a variable has more than two parents, draw lines between every pair of parents.)



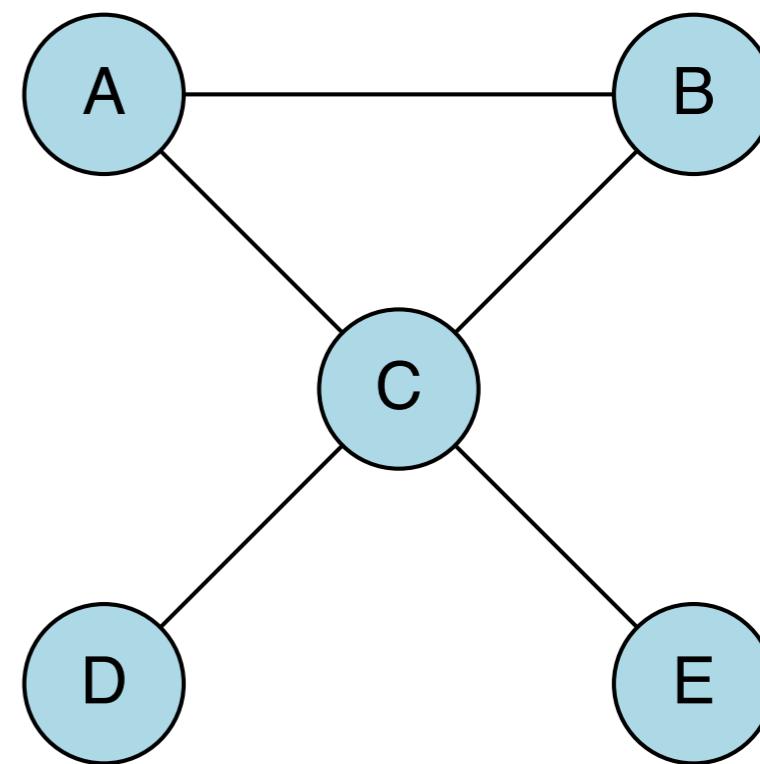
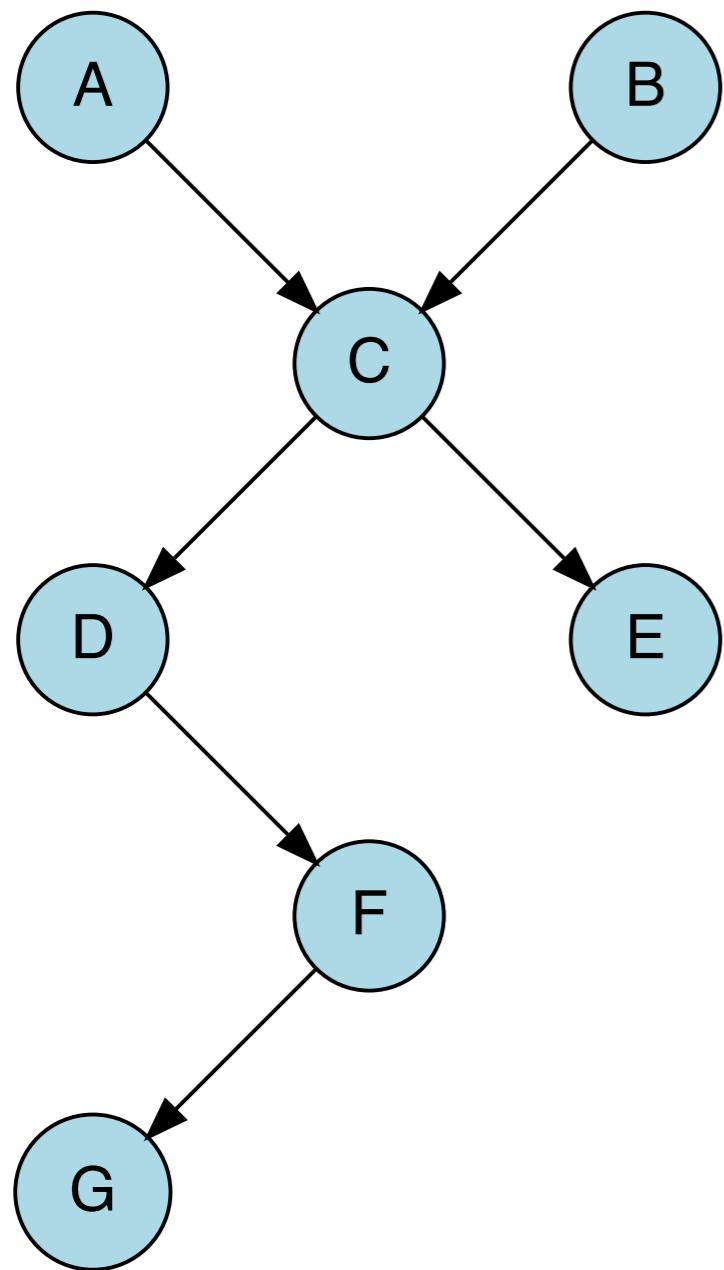
**let's get married!**

# When should I control for variables?

**Are D and E independent, given C? 3. "Disorient" the graph**

$$p(D | E, C) = p(D | C) ?$$

Replace arrows with lines



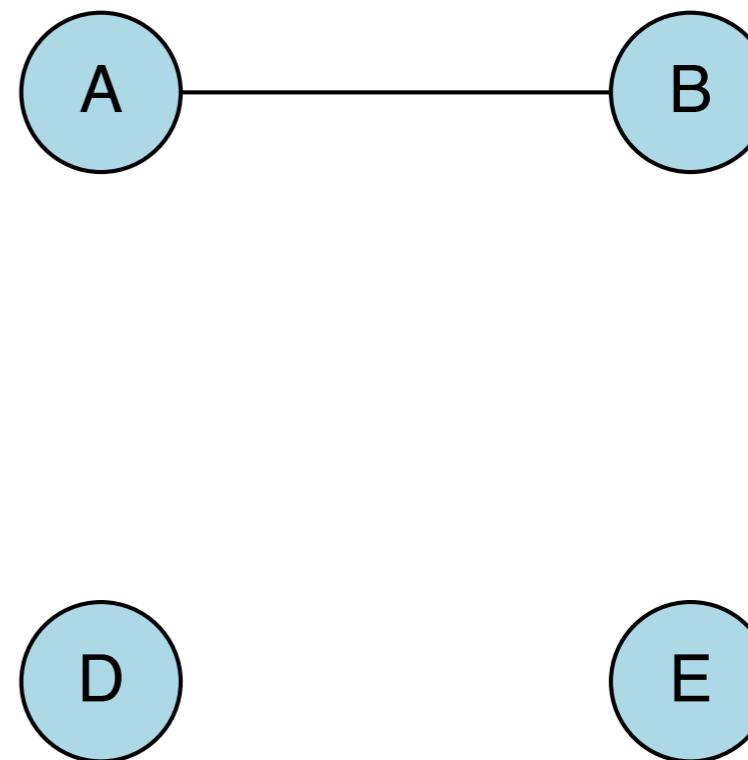
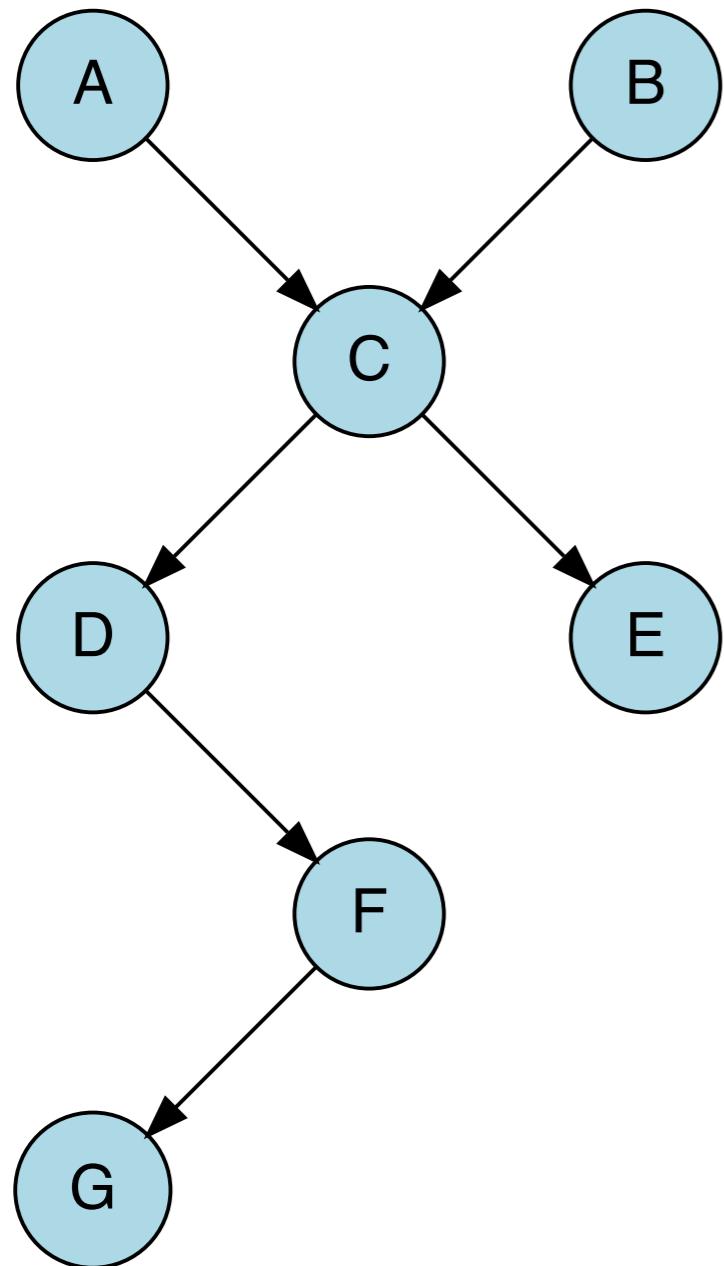
# When should I control for variables?

**Are D and E independent, given C? 4. Delete the givens**

$$p(D | E, C) = p(D | C) ?$$

Remove the variables that we condition on, as well as their edges

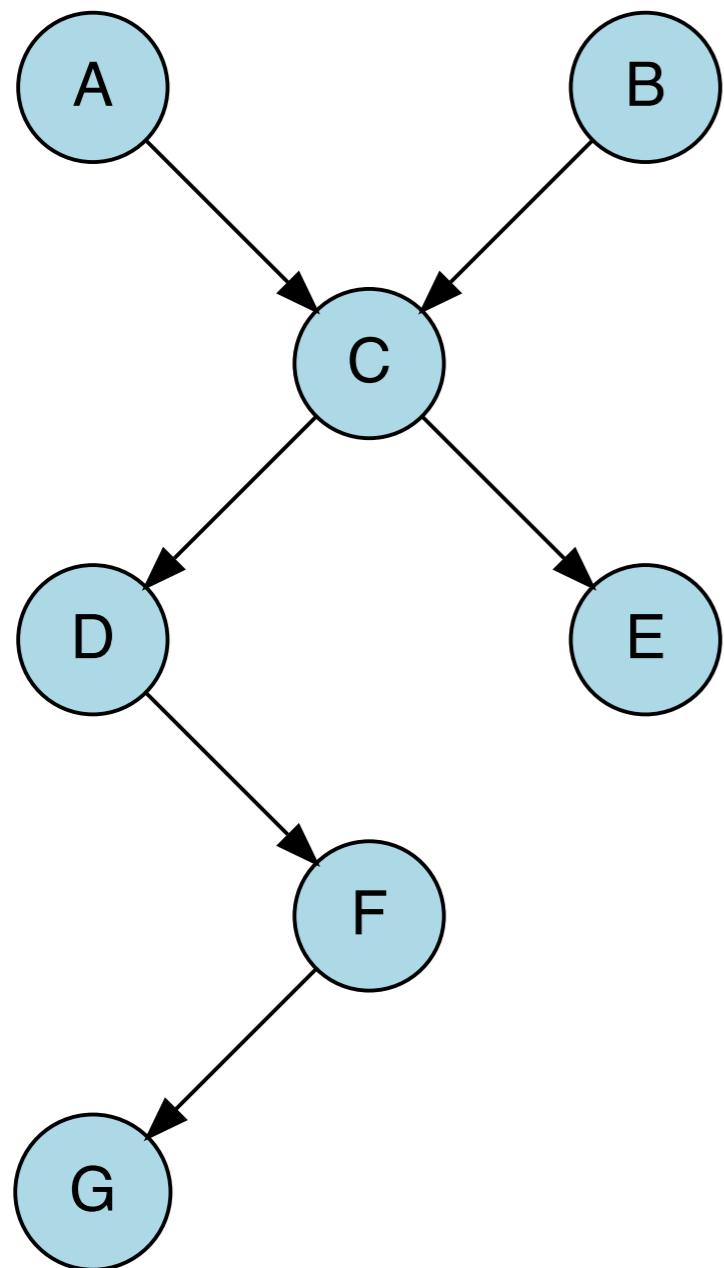
**we conditioned on C!**



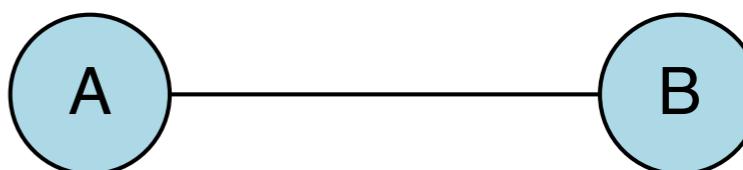
# When should I control for variables?

**Are D and E independent, given C? 5. Read answer off the graph**

$$p(D | E, C) = p(D | C) ?$$



- if variables are **disconnected** they are independent
- if variables are connected (have a path between them) they are not guaranteed to be independent



**D and E are independent from each other conditioned on C**



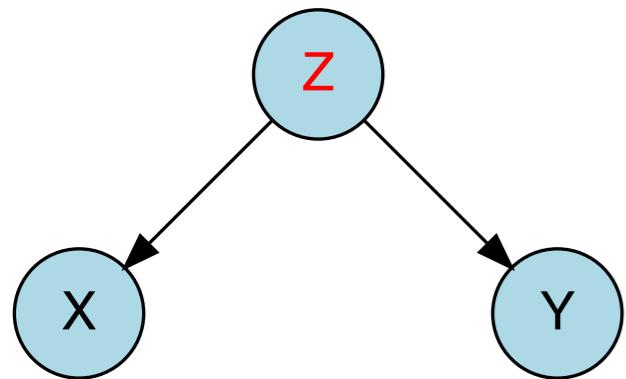
**they aren't connected via a path**

# So what?

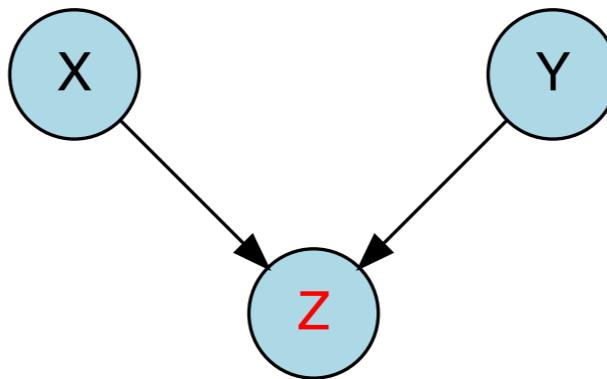


# Patterns of inference

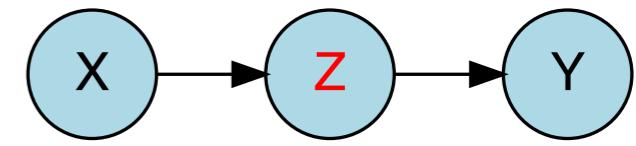
We want to estimate the (causal) relationship between X and Y



common cause



common effect



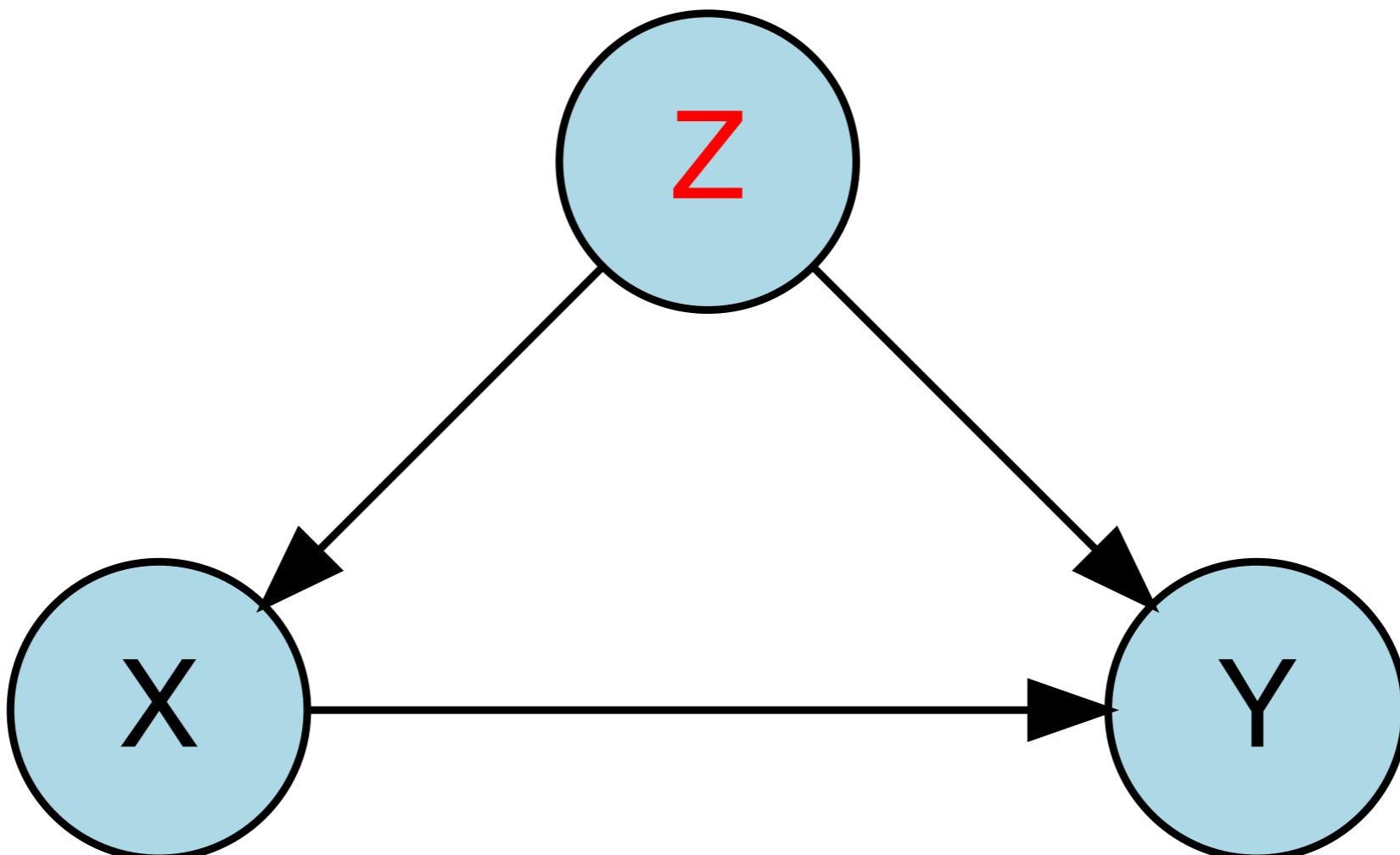
causal chain

by controlling for Z we hope to get a better estimate of the relationship between X and Y

**d-separation** helps us tell apart **good controls** from **bad controls**

# When should I control for variables?

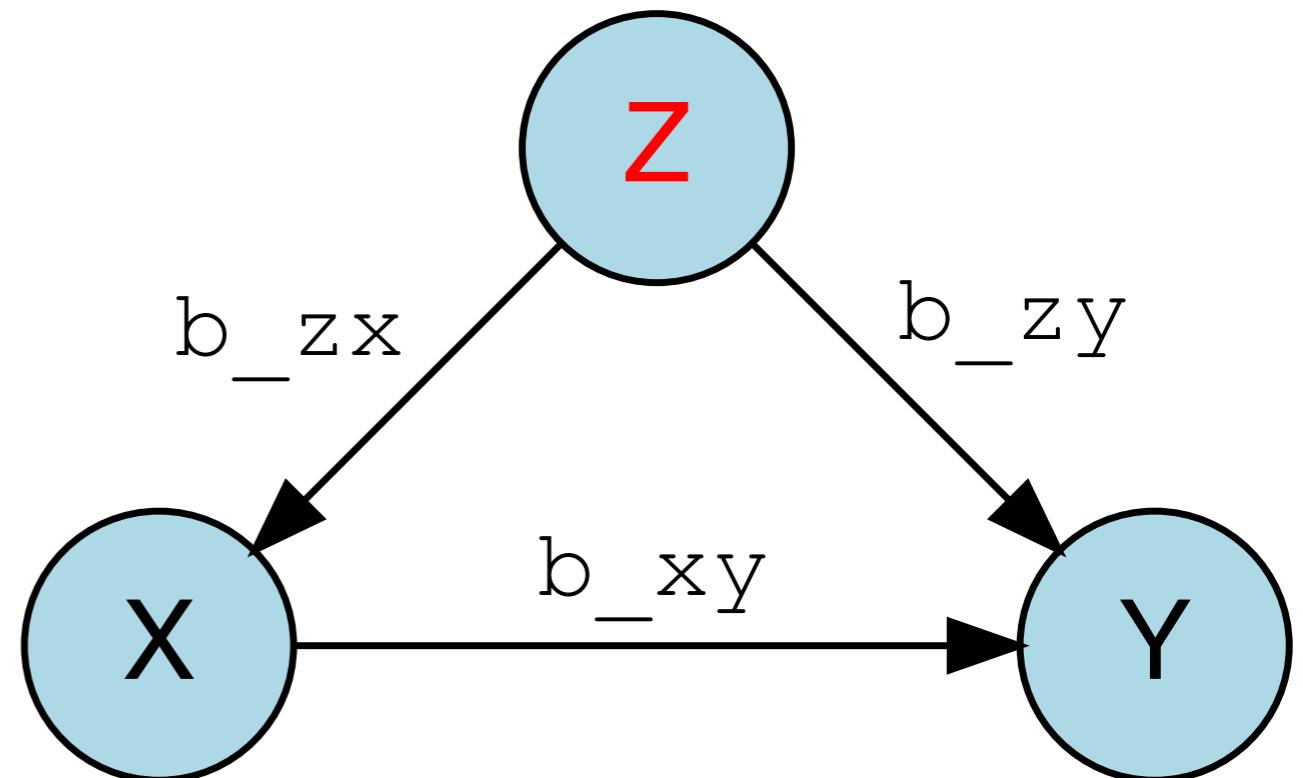
I want to estimate the effect that X has on Y



Is Z a **good** or a **bad** control here?

# When should I control for variables?

```
1 set.seed(1)
2
3 n = 1000
4 b_zx = 2
5 b_xy = 2
6 b_zy = 2
7 sd = 1
8
9 fun_error = function(n, sd) {
10   rnorm(n = n,
11         mean = 0,
12         sd = sd)
13 }
14
15 df = tibble(z = fun_error(n, sd),
16               x = b_zx * z + fun_error(n, sd),
17               y = b_zy * z + b_xy * x + fun_error(n, sd))
```



overestimating  
X's effect on Y

$$Y = b_0 + b_1 \cdot X + e$$

```
1 # without control
2 lm(formula = y ~ x,
3     data = df) %>%
4   summary()
```

```
Call:
lm(formula = y ~ x, data = df)

Residuals:
    Min      1Q  Median      3Q     Max 
-4.6011 -0.9270 -0.0506  0.9711  4.0454 

Coefficients:
            Estimate Std. Error t value Pr(>|t|)    
(Intercept) 0.02449   0.04389   0.558   0.577    
x           2.82092   0.01890 149.225 <2e-16 ***  
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

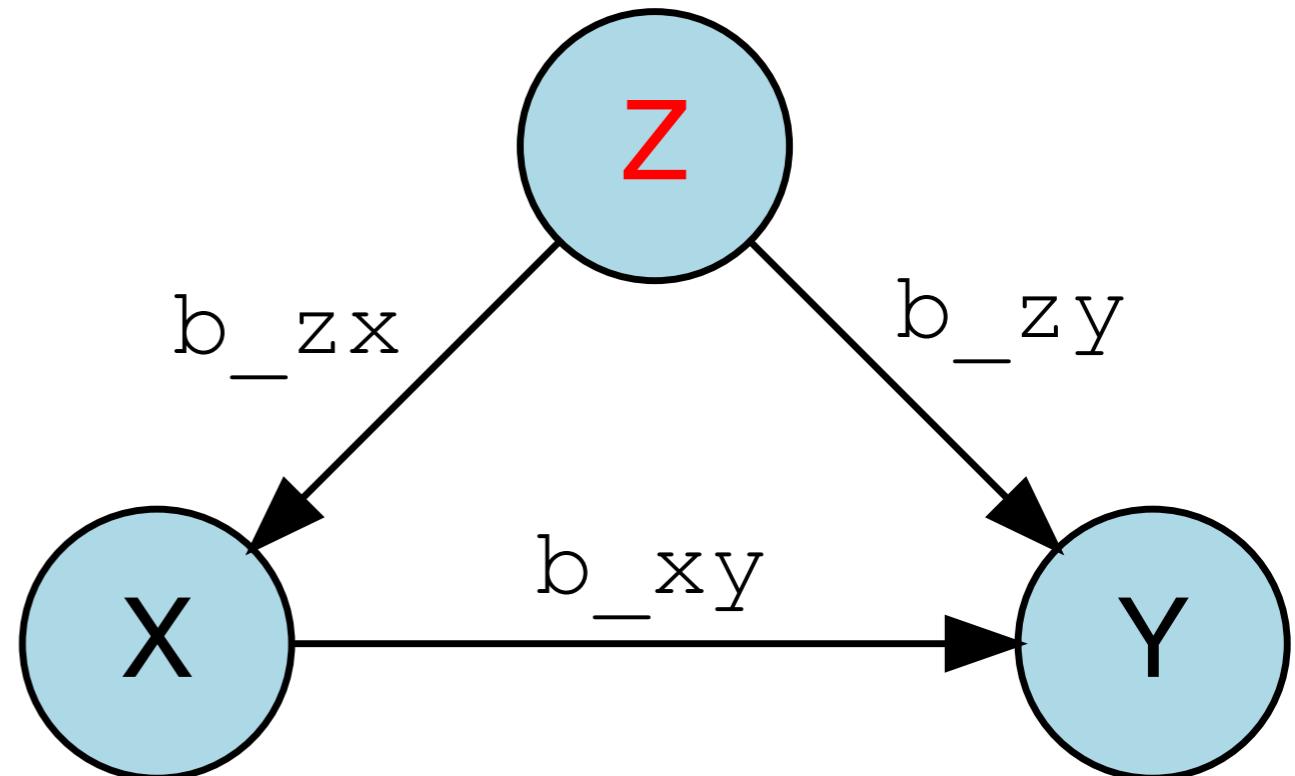
Residual standard error: 1.388 on 998 degrees of freedom
Multiple R-squared:  0.9571,    Adjusted R-squared:  0.9571 
F-statistic: 2.227e+04 on 1 and 998 DF,  p-value: < 2.2e-16
```

# When should I control for variables?

```
1 set.seed(1)
2
3 n = 1000
4 b_zx = 2
5 b_xy = 2
6 b_zy = 2
7 sd = 1
8
9 fun_error = function(n, sd) {
10   rnorm(n = n,
11         mean = 0,
12         sd = sd)
13 }
14
15 df = tibble(z = fun_error(n, sd),
16               x = b_zx * z + fun_error(n, sd),
17               y = b_zy * z + b_xy * x + fun_error(n, sd))
```

$$Y = b_0 + b_1 \cdot X + b_2 \cdot Z + e$$

```
1 # with control
2 lm(formula = y ~ x + z,
3     data = df) %>%
4   summary()
```



Call:  
lm(formula = y ~ x + z, data = df)

Residuals:

Min	1Q	Median	3Q	Max
-3.6151	-0.6564	-0.0223	0.6815	2.8132

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	0.01624	0.03260	0.498	0.618
x	2.02202	0.03135	64.489	<2e-16 ***
z	2.00501	0.07036	28.497	<2e-16 ***

---

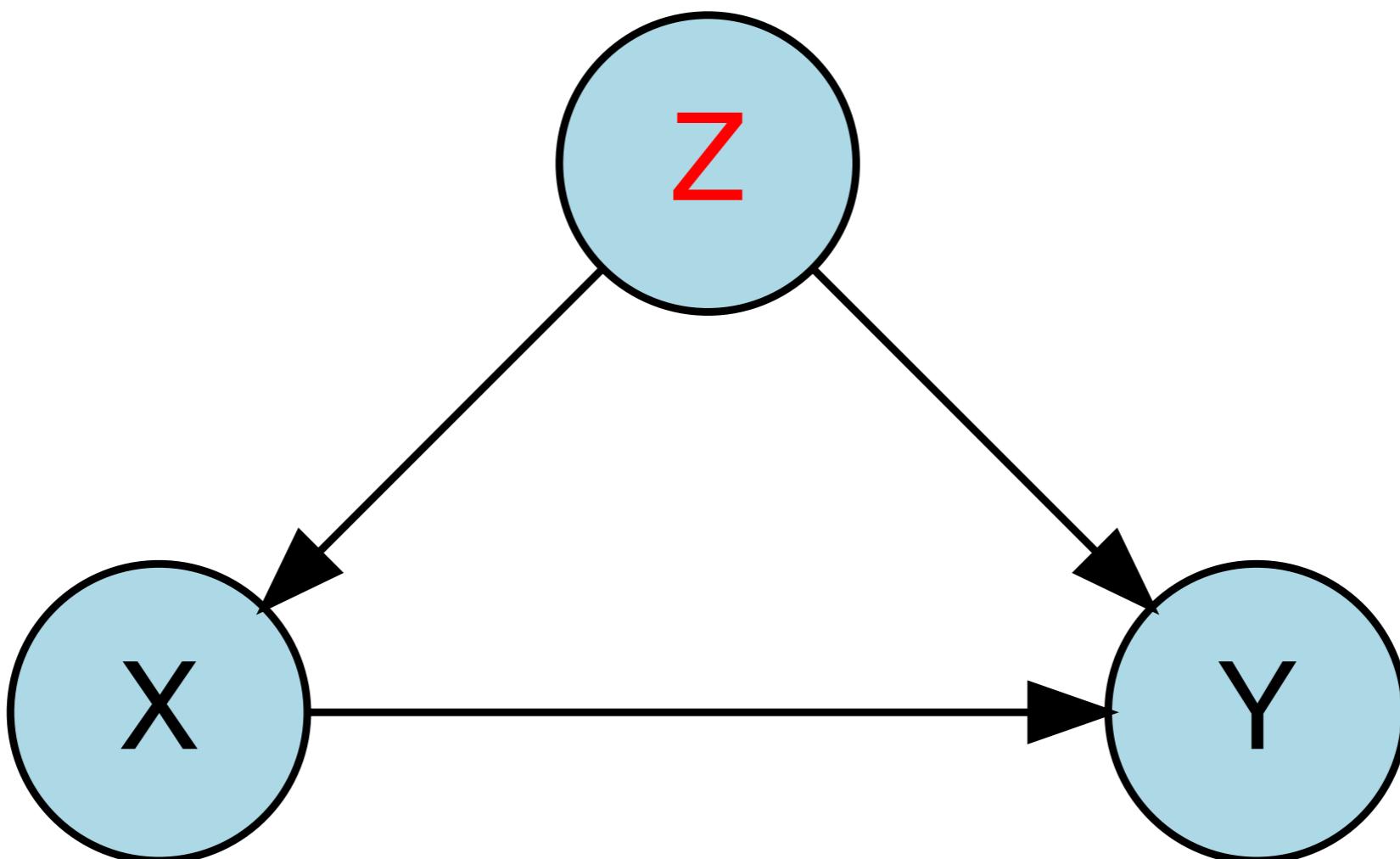
Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 1.031 on 997 degrees of freedom  
Multiple R-squared: 0.9764, Adjusted R-squared: 0.9763  
F-statistic: 2.059e+04 on 2 and 997 DF, p-value: < 2.2e-16

**accurate estimate  
of X's effect on Y**

# When should I control for variables?

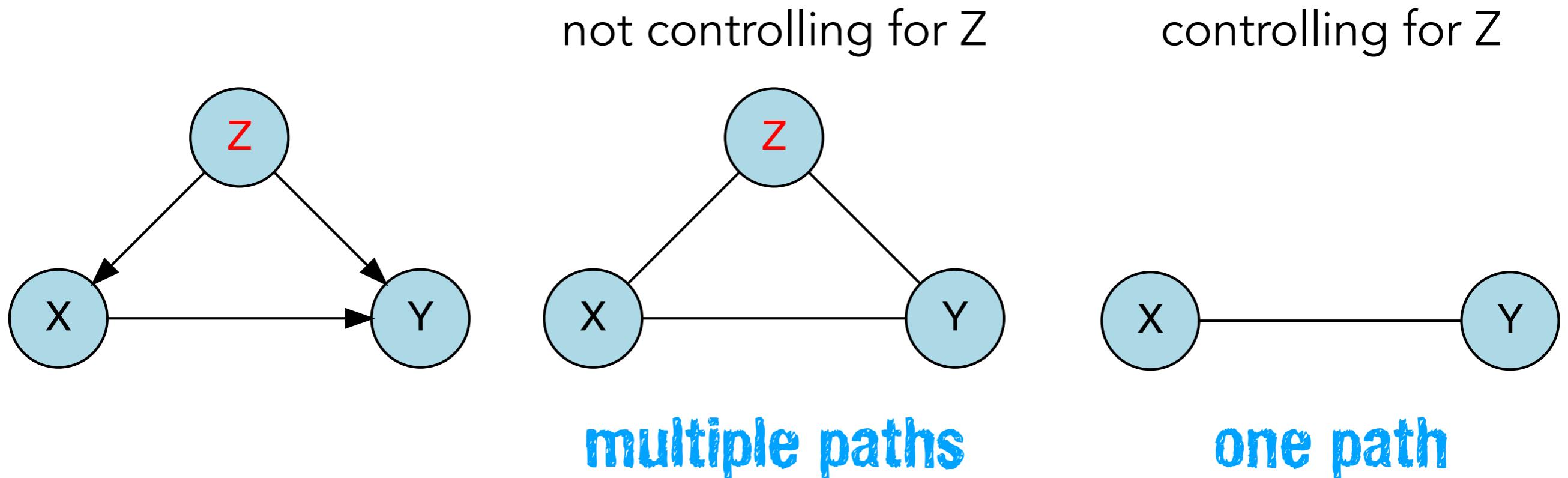
I want to estimate the effect that X has on Y



Z is a **good** control here!

# When should I control for variables?

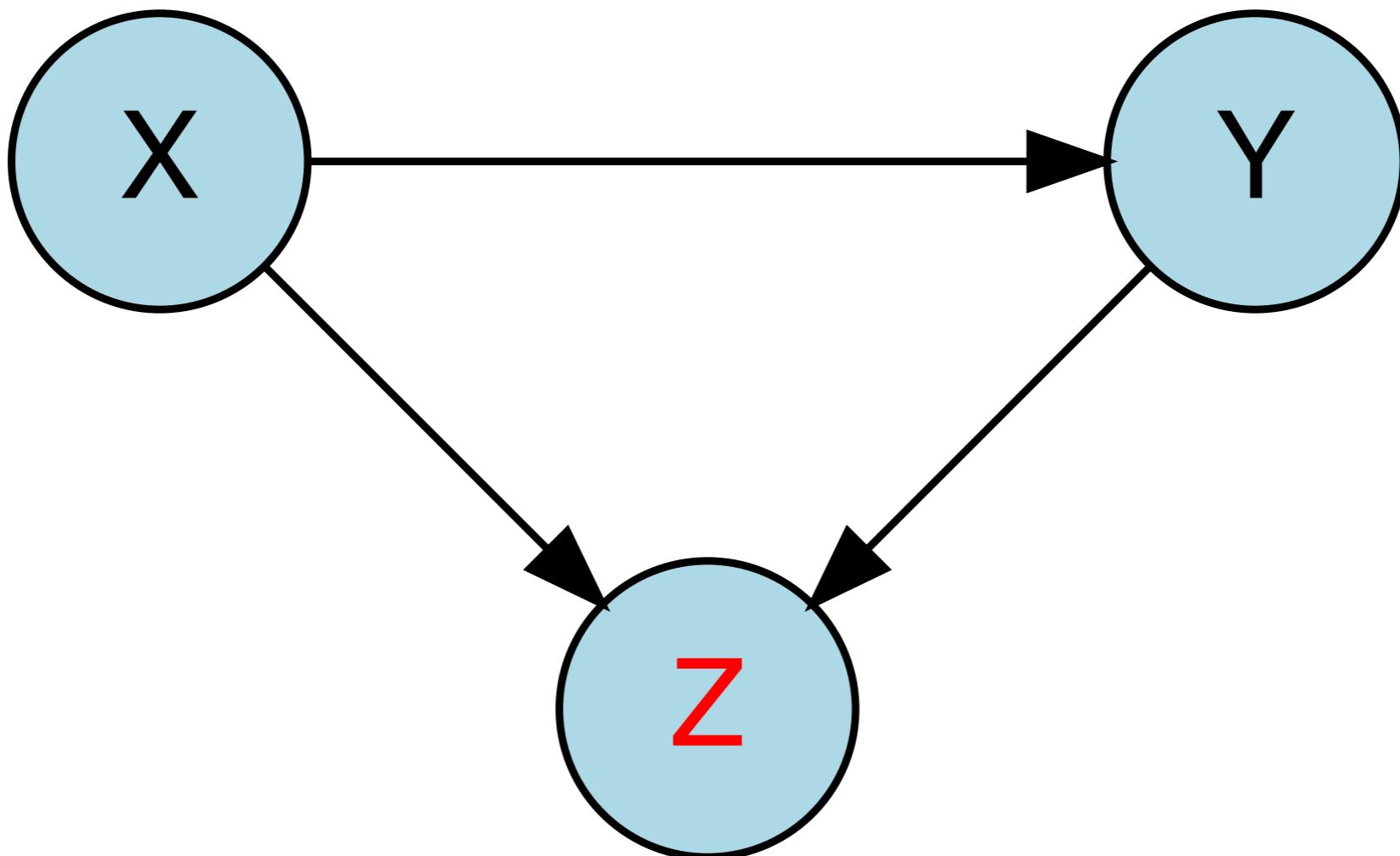
I want to estimate the effect that X has on Y



Z is a **good** control here!

# When should I control for variables?

I want to estimate the effect that X has on Y



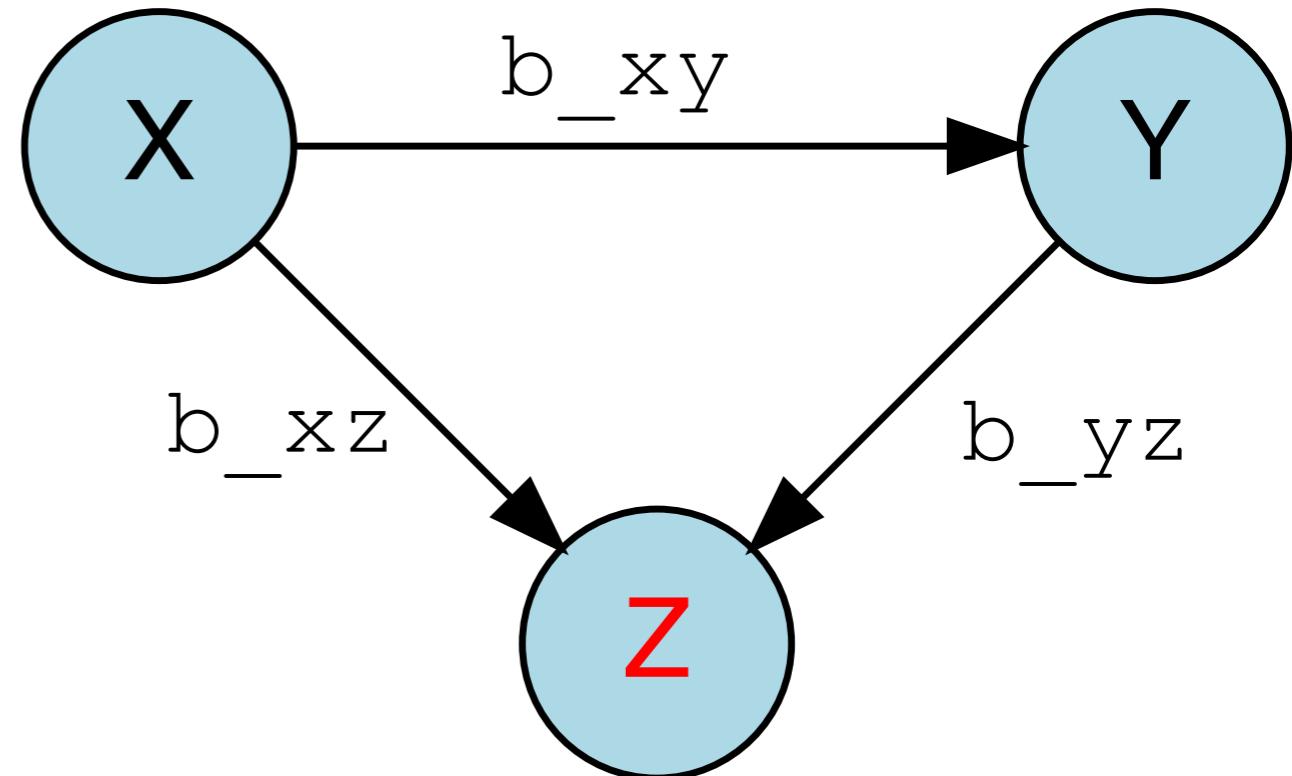
Is Z a **good** or a **bad** control here?

# When should I control for variables?

```
1 set.seed(1)
2 n = 1000
3 b_xy = 2
4 b_xz = 2
5 b_yz = 2
6 sd = 1
7
8 fun_error = function(n, sd) {
9   rnorm(n = n,
10         mean = 0,
11         sd = sd)
12 }
13
14 df = tibble(x = fun_error(n, sd),
15             y = x * b_xy + fun_error(n, sd),
16             z = x * b_xz + y * b_yz + fun_error(n, sd))
```

$$Y = b_0 + b_1 \cdot X + e$$

```
1 # without control
2 lm(formula = y ~ x,
3     data = df) %>%
4   summary()
```



accurate estimate  
of X's effect on Y

```
Call:
lm(formula = y ~ x, data = df)

Residuals:
    Min      1Q  Median      3Q     Max 
-3.2484 -0.6720 -0.0138  0.7554  3.6443 

Coefficients:
            Estimate Std. Error t value Pr(>|t|)    
(Intercept) -0.01619   0.03290  -0.492   0.623    
x            2.00643   0.03181  63.078 <2e-16 ***  
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

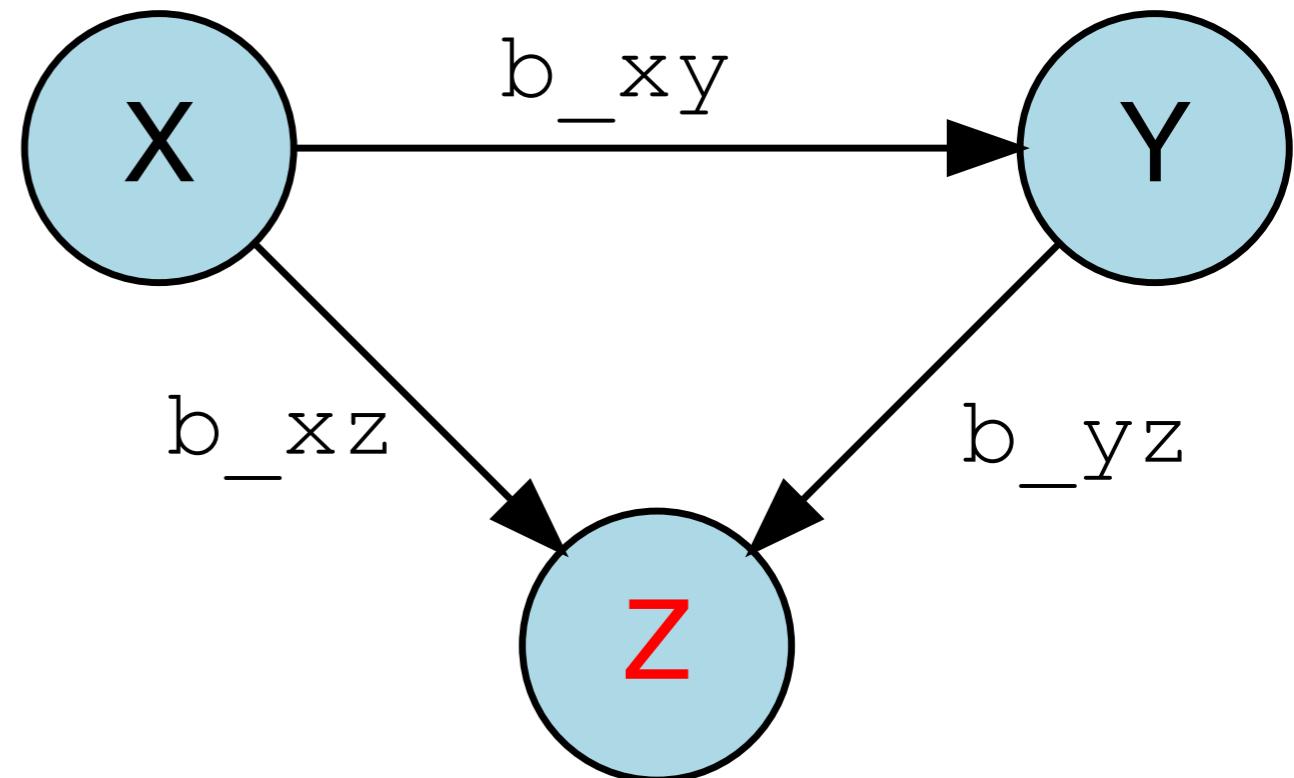
Residual standard error: 1.04 on 998 degrees of freedom
Multiple R-squared:  0.7995,    Adjusted R-squared:  0.7993 
F-statistic: 3979 on 1 and 998 DF,  p-value: < 2.2e-16
```

# When should I control for variables?

```
1 set.seed(1)
2 n = 1000
3 b_xy = 2
4 b_xz = 2
5 b_yz = 2
6 sd = 1
7
8 fun_error = function(n, sd) {
9   rnorm(n = n,
10         mean = 0,
11         sd = sd)
12 }
13
14 df = tibble(x = fun_error(n, sd),
15             y = x * b_xy + fun_error(n, sd),
16             z = x * b_xz + y * b_yz + fun_error(n, sd))
```

$$Y = b_0 + b_1 \cdot X + b_2 \cdot Z + e$$

```
1 # with control
2 lm(formula = y ~ x + z,
3     data = df) %>%
4   summary()
```



in accurate  
estimate of X's  
effect on Y

```
Call:
lm(formula = y ~ x + z, data = df)

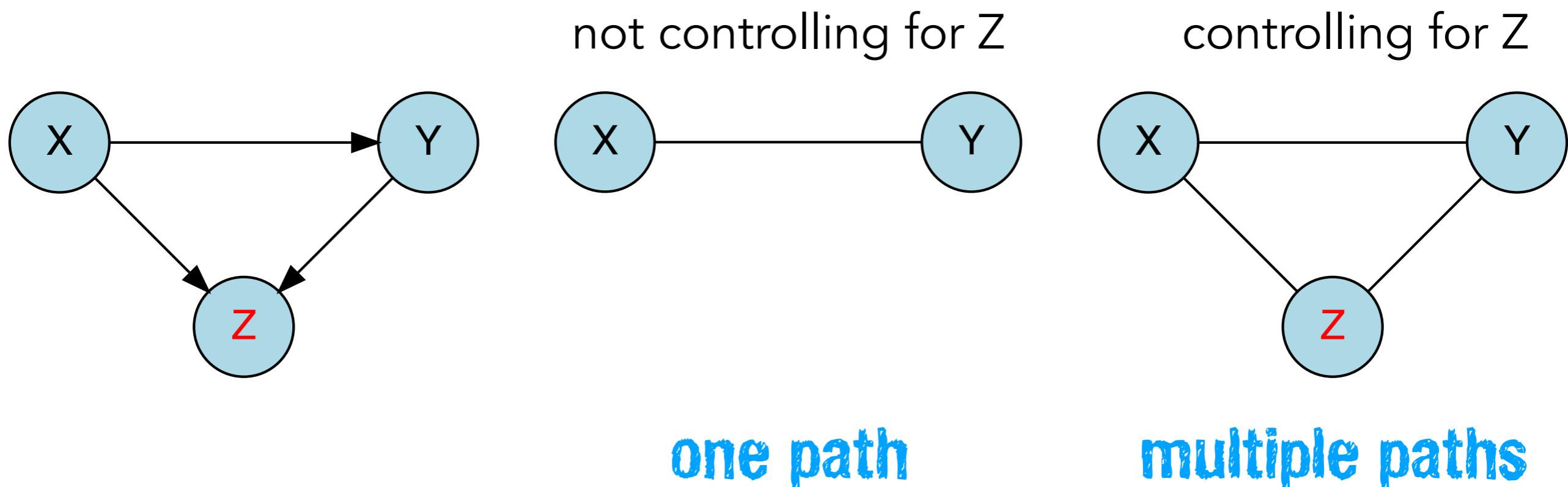
Residuals:
    Min      1Q  Median      3Q     Max 
-1.35547 -0.30016  0.09298  0.31119  1.73408 

Coefficients:
            Estimate Std. Error t value Pr(>|t|)    
(Intercept) -0.009608  0.014477 -0.664   0.507    
x            -0.411847  0.040026 -10.290  <2e-16 ***  
z             0.398921  0.006186  64.489  <2e-16 ***  
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.4578 on 997 degrees of freedom
Multiple R-squared:  0.9612,    Adjusted R-squared:  0.9611 
F-statistic: 1.236e+04 on 2 and 997 DF,  p-value: < 2.2e-16
```

# When should I control for variables?

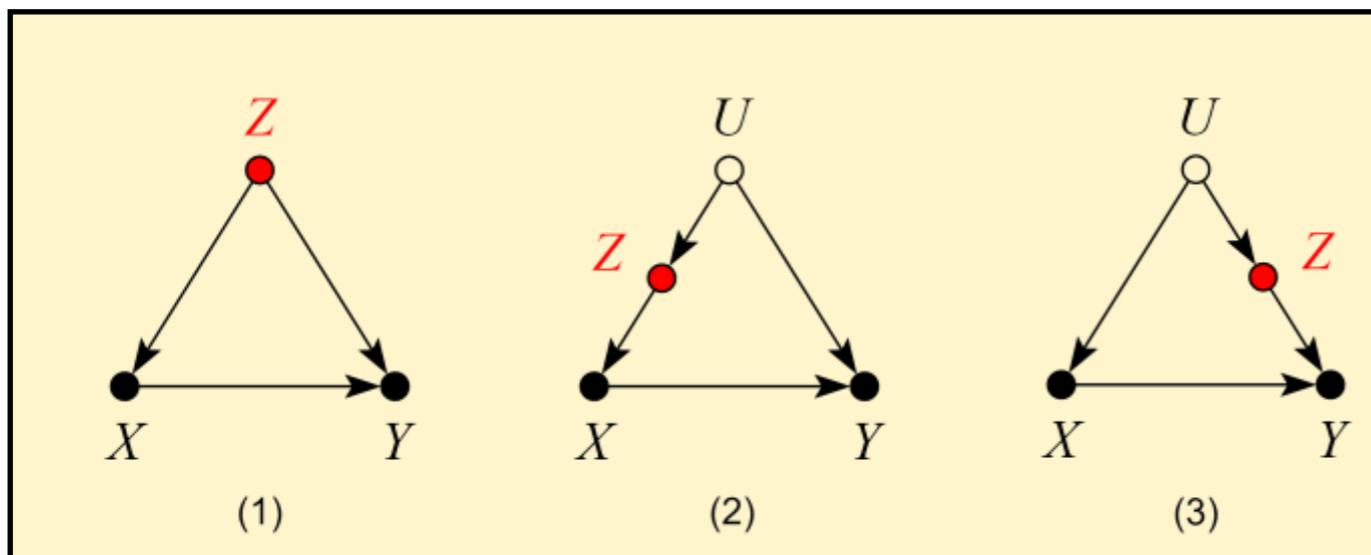
I want to estimate the effect that X has on Y



Z is a **bad** control here!

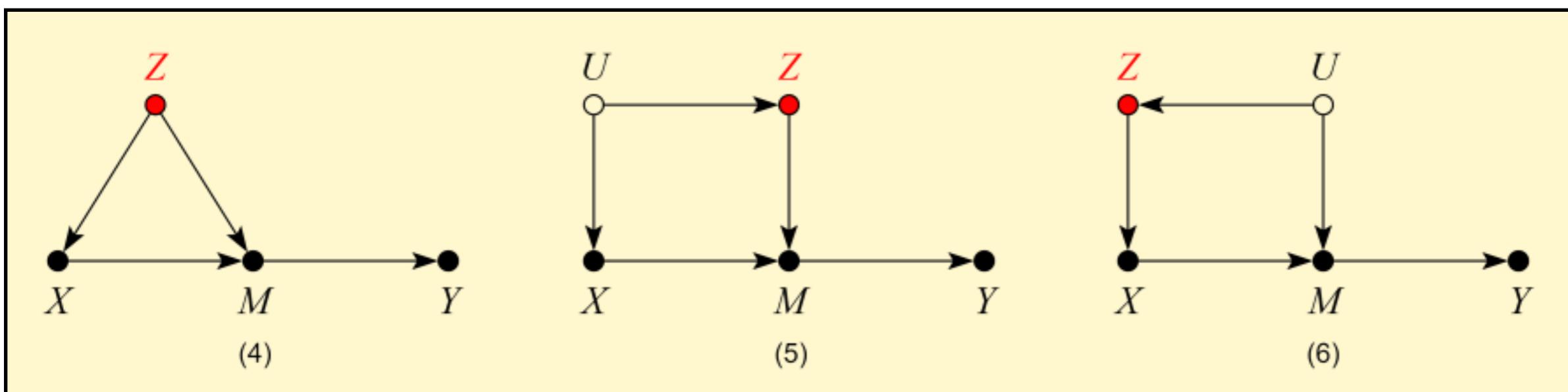
# When should I control for variables?

I want to estimate the effect that X has on Y



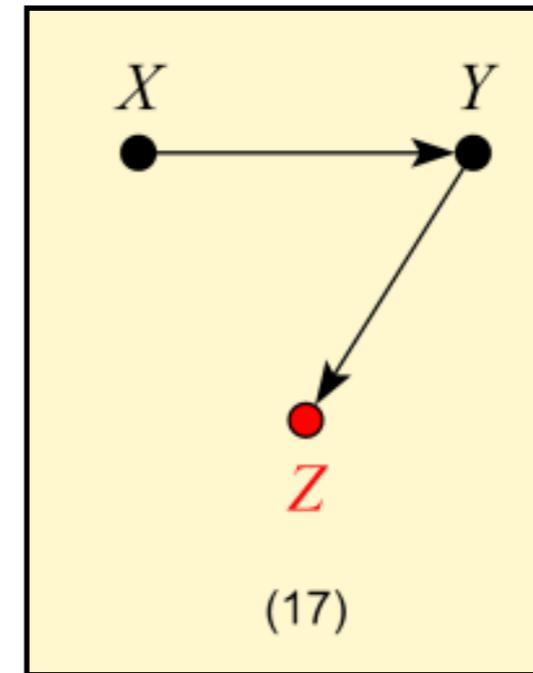
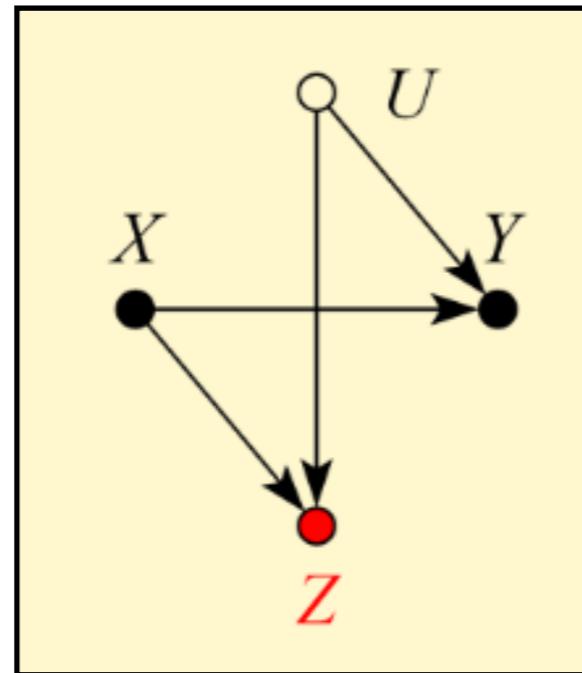
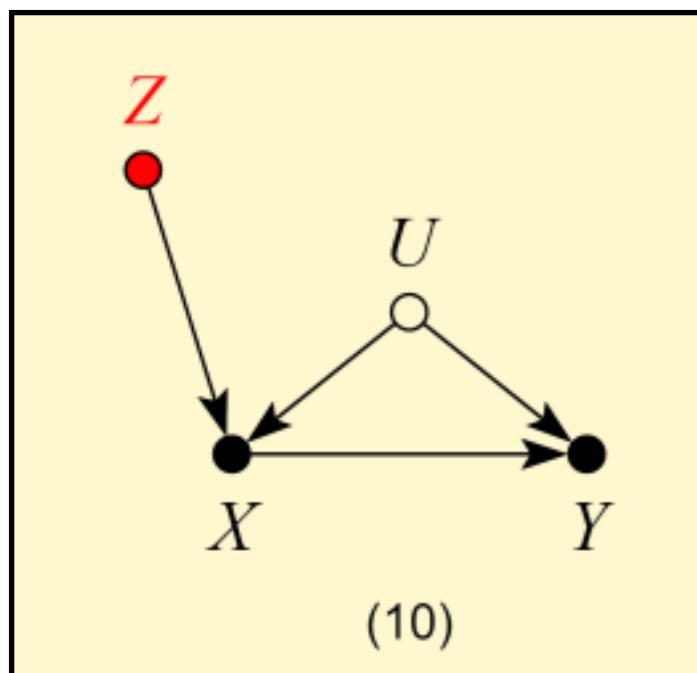
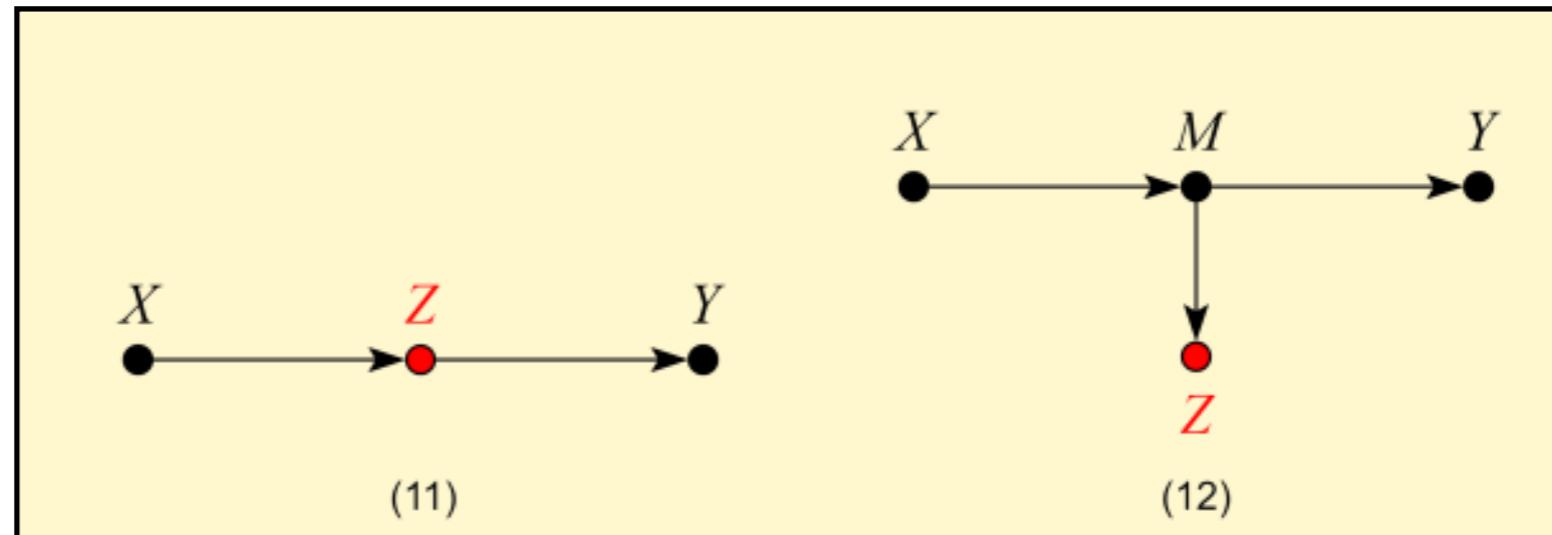
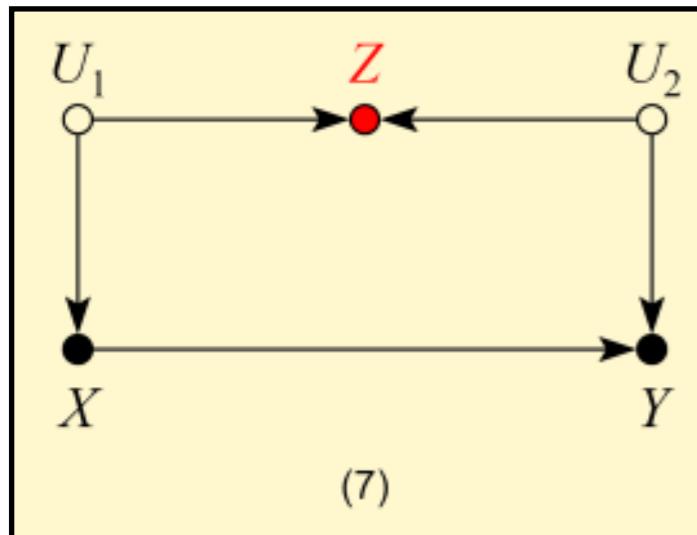
Good controls

U = unobserved variable  
M = mediating variable



# When should I control for variables?

I want to estimate the effect that  $X$  has on  $Y$

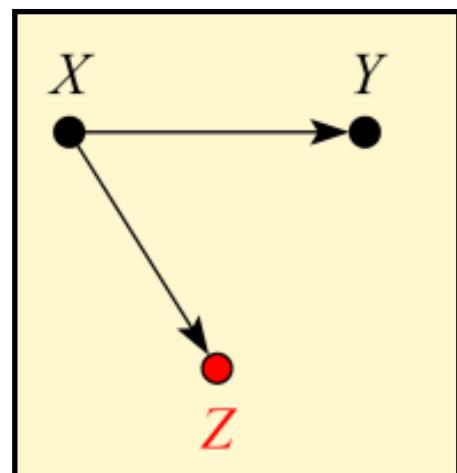


Bad controls

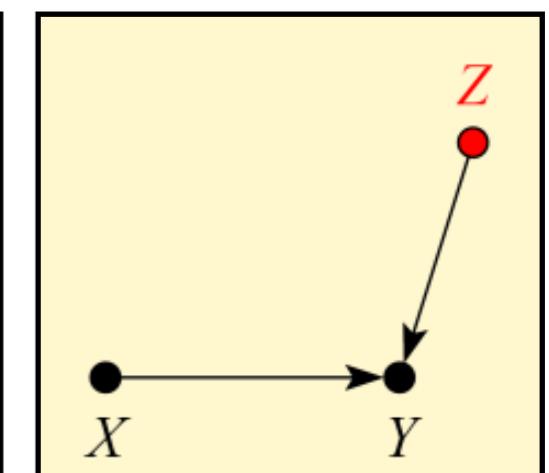
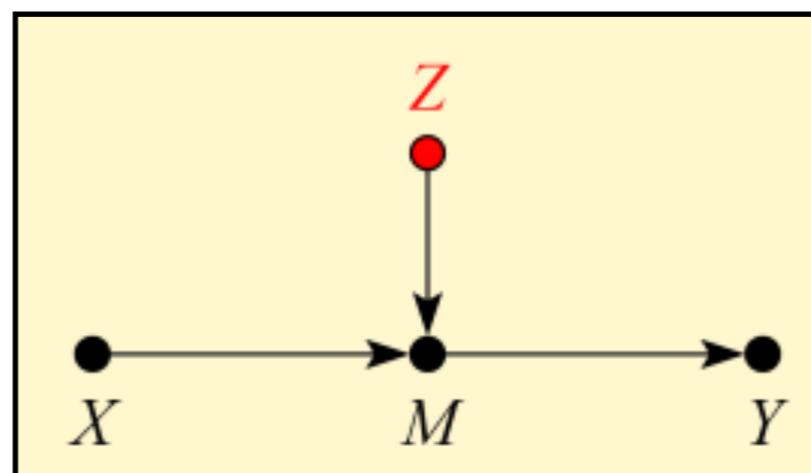
# When should I control for variables?

I want to estimate the effect that X has on Y

possibly bad for precision



possibly good for precision



conditioning on Z reduces  
X's variance

conditioning on Z reduces variance of  
mediator or outcome variable

**Neutral controls**

don't introduce any bias

# When should I control for variables?

## Babies Learning Language

Thoughts on language learning, child development, and fatherhood; experimental methods, reproducibility, and open science; theoretical musings on cognitive science more broadly.

Tuesday, October 8, 2019

### Confounds and covariates

(tl;dr: explanation of confounding and covariate adjustment)

Every year, one of the trickiest concepts for me to teach in my experimental methods course is the difference between experimental confounds and covariates. Although this distinction seems simple, it's pretty deeply related to the definition of what an experiment is and why experiments lead to good causal inferences. It's also caught up in a number of methodological problems that come up again and again in my class. This post is my attempt to explain the distinction and how it relates to different problems and cultural practices in psychology.

#### About



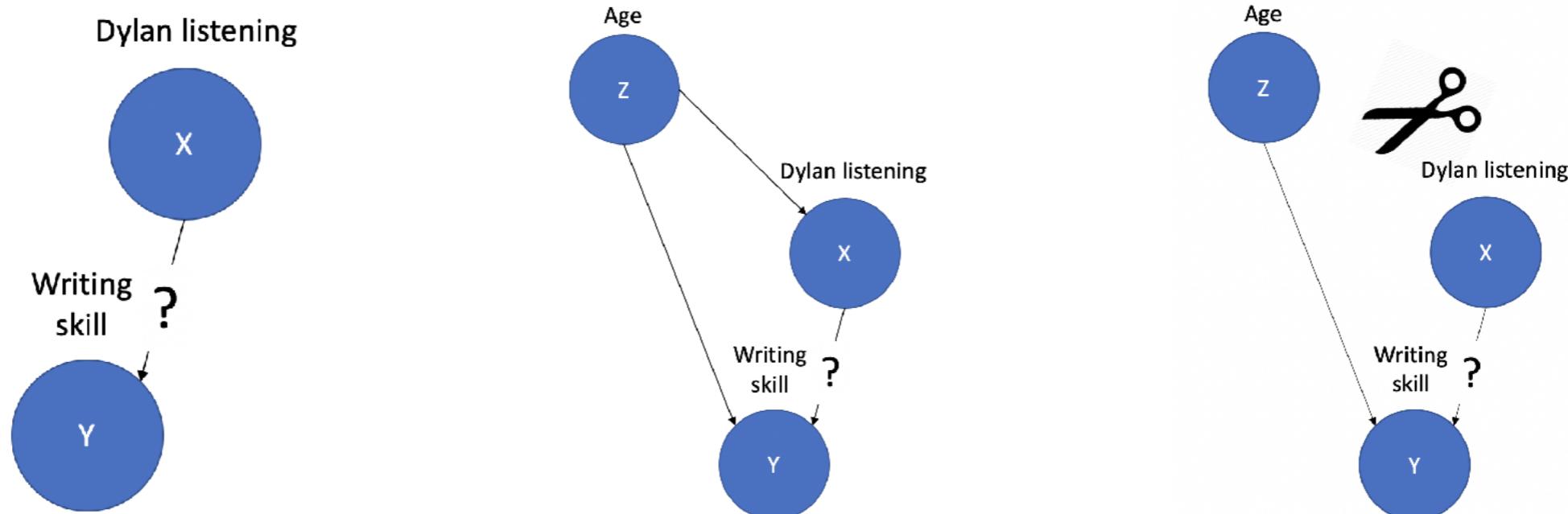
[Michael Frank](#)

Developmental Psychologist at Stanford University. Interested in language acquisition, cognitive development, thoughts, numbers, learning, and teaching. [Visit my lab's website](#).

[View my complete profile](#)

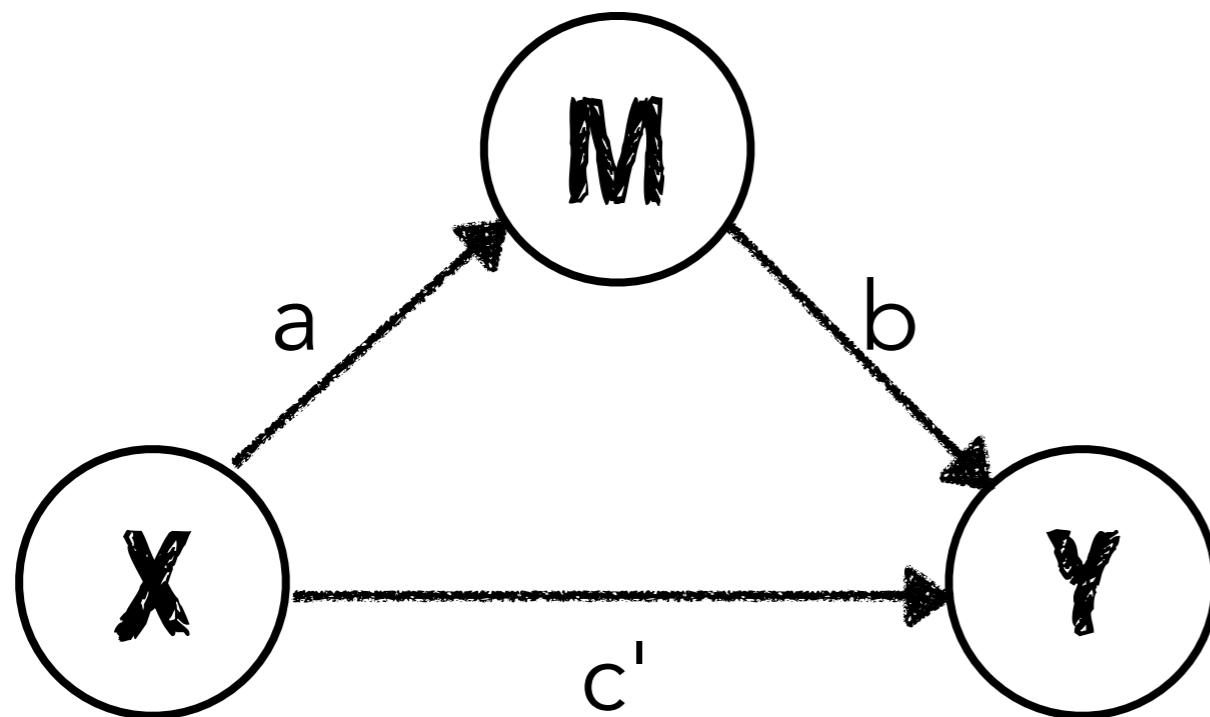
#### Labels

- [Development](#) (40)



# **Mediation**

# Definition

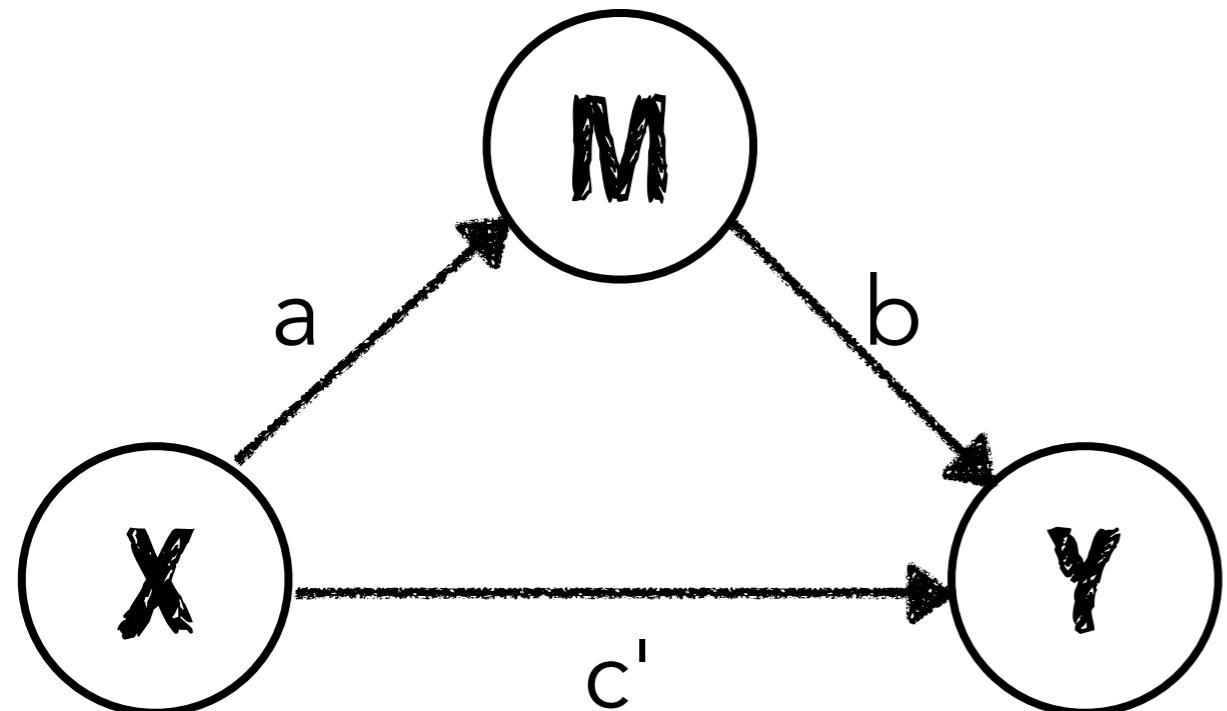


Rather than a direct causal relationship between **X** and **Y**, a mediation model proposes that **X** influences the mediator variable **M**, which in turn influences **Y**. Thus, the mediator variable serves to clarify the nature of the relationship between **X** and **Y**.

**Adapted from Wikipedia**

[https://en.wikipedia.org/wiki/Mediation\\_\(statistics\)](https://en.wikipedia.org/wiki/Mediation_(statistics))

# Example

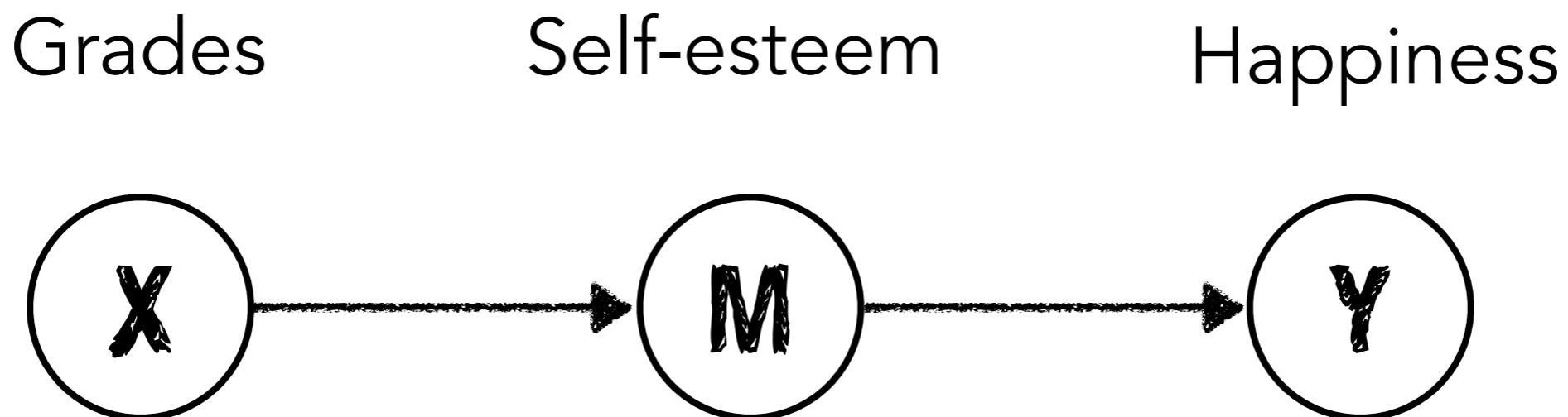


**X** = grades in Psych 252  
**M** = feelings of self-esteem  
**Y** = happiness

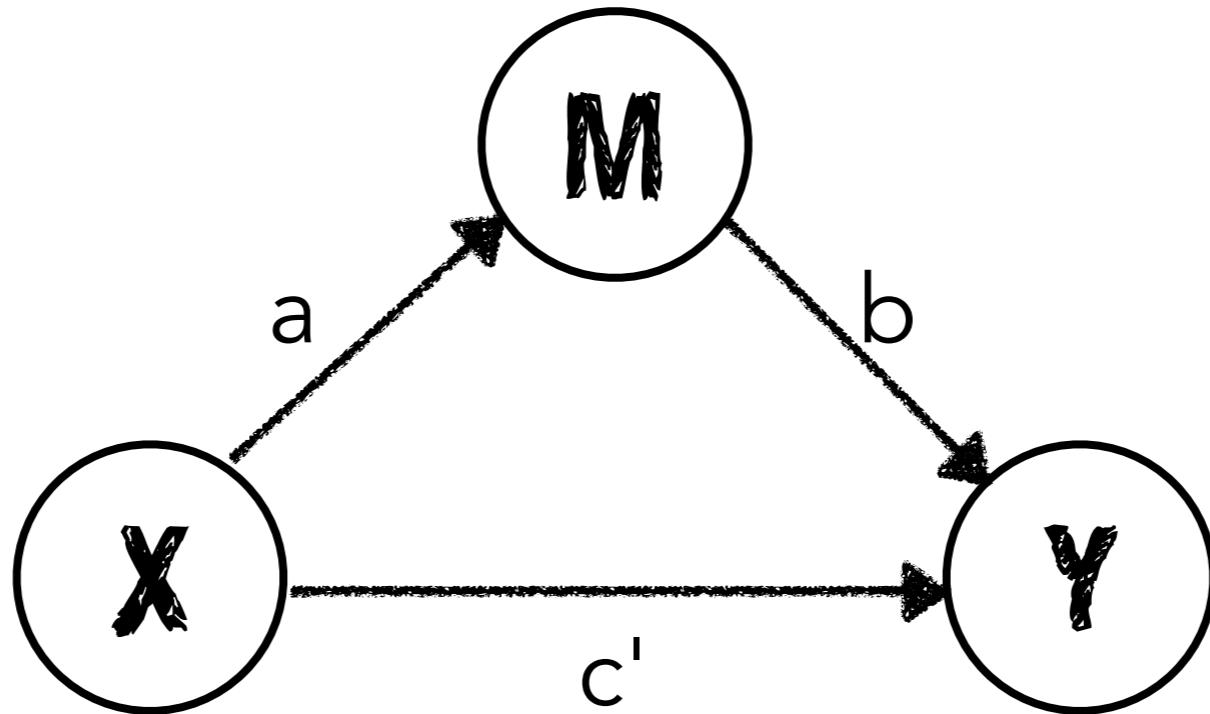
Is the relationship between grades in Psych 252 and happiness mediated by feelings of self-esteem?

# Simulate a mediation analysis

```
1 # number of participants
2 n = 100
3
4 # generate data
5 df.mediation = tibble(
6   x = rnorm(n, 75, 7),           # grades
7   m = 0.7 * x + rnorm(n, 0, 5), # self-esteem
8   y = 0.4 * m + rnorm(n, 0, 5) # happiness
9 )
```



# Baron and Kenny's (1986) steps for mediation



## Sequence of regression models

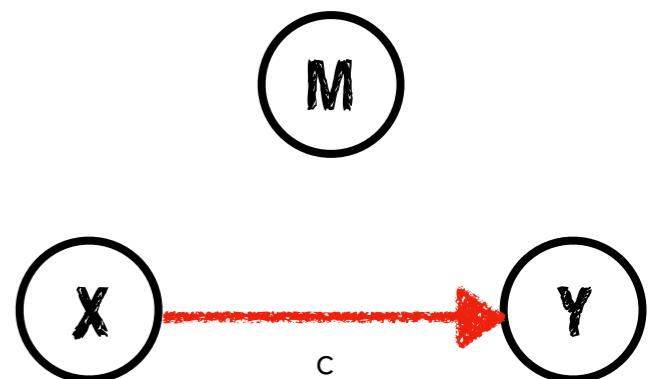
1. Is there a relationship between **X** and **Y**?
2. Is there a relationship between **X** and **M**?
3. Does the relationship between **X** and **Y** change, once we control for **M**?

Baron, R. M. & Kenny, D. A. (1986). The moderator-mediator variable distinction in social psychological research: Conceptual, strategic, and statistical considerations. *Journal of Personality and Social Psychology*, 51(6), 1173-1182.

# Is there a relationship between X and Y?

$$\hat{y} = b_0 + b_1 \cdot x$$

```
1 # fit the model
2 fit.y_x = lm(formula = y ~ 1 + x,
3               data = df.mediation)
4
5 # summarize the results
6 fit.y_x %>% summary()
```



```
Call:
lm(formula = y ~ 1 + x, data = df.mediation)

Residuals:
    Min      1Q  Median      3Q     Max 
-10.917 -3.738 -0.259  2.910 12.540 

Coefficients:
            Estimate Std. Error t value Pr(>|t|)    
(Intercept) 8.78300   6.16002   1.426   0.1571    
x            0.16899   0.08116   2.082   0.0399 *  
                                                 
Signif. codes:  0 '****' 0.001 '***' 0.01 '**' 0.05 '.' 0.1 ' ' 1

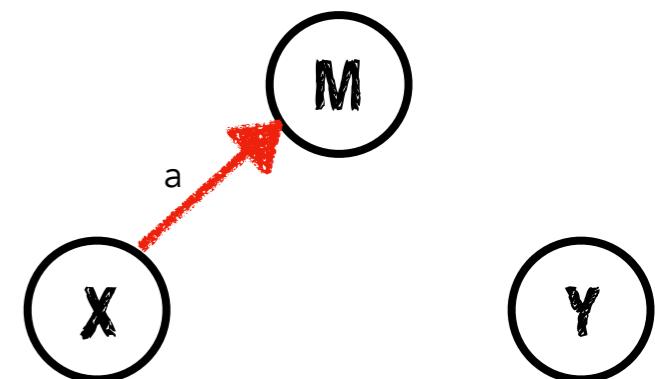
Residual standard error: 5.16 on 98 degrees of freedom
Multiple R-squared:  0.04237, Adjusted R-squared:  0.0326 
F-statistic: 4.336 on 1 and 98 DF,  p-value: 0.03993
```

significant  
relationship

# Is there a relationship between X and M?

$$\hat{m} = b_0 + b_1 \cdot x$$

```
1 # fit the model
2 fit.m_x = lm(formula = m ~ 1 + x,
3               data = df.mediation)
4
5 # summarize the results
6 fit.m_x %>% summary()
```



```
Call:
lm(formula = m ~ 1 + x, data = df.mediation)

Residuals:
    Min      1Q  Median      3Q     Max 
-9.5367 -3.4175 -0.4375  2.9032 16.4520 

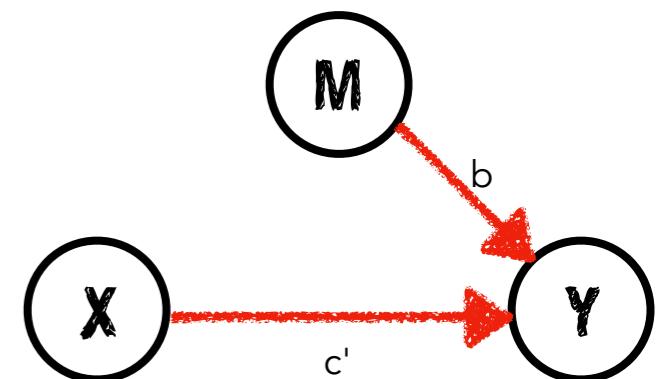
Coefficients:
            Estimate Std. Error t value Pr(>|t|)    
(Intercept) 6.04494   13.41692   0.451   0.653    
x            0.66252    0.07634   8.678 8.87e-14 *** 
---
Signif. codes:  0 '****' 0.001 '***' 0.01 '**' 0.05 '*' 0.1 '.' 1 

Residual standard error: 4.854 on 98 degrees of freedom
Multiple R-squared:  0.4346,    Adjusted R-squared:  0.4288 
F-statistic: 75.31 on 1 and 98 DF,  p-value: 8.872e-14
```

significant  
relationship

# Is there a relationship between X and Y, controlling for M?

$$\hat{y} = b_0 + b_1 \cdot m + b_2 \cdot x$$



```
1 # fit the model
2 fit.y_mx = lm(formula = y ~ 1 + m + x,
3                 data = df.mediation)
4
5 # summarize the results
6 fit.y_mx %>% summary()
```

```
Call:
lm(formula = y ~ 1 + m + x, data = df.mediation)

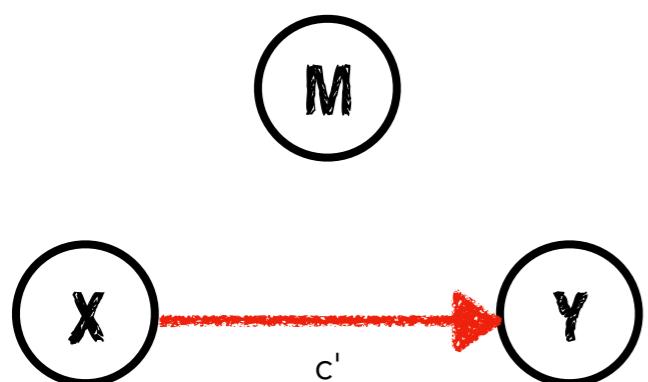
Residuals:
    Min      1Q  Median      3Q     Max 
-9.3651 -3.3037 -0.6222  3.1068 10.3991 

Coefficients:
            Estimate Std. Error t value Pr(>|t|)    
(Intercept) 7.80952   5.68297   1.374   0.173    
m           0.42381   0.09899   4.281 4.37e-05 *** 
x          -0.11179   0.09949  -1.124   0.264    
---
Signif. codes:  0 '****' 0.001 '***' 0.01 '**' 0.05 '*' 0.1 '.' 1 

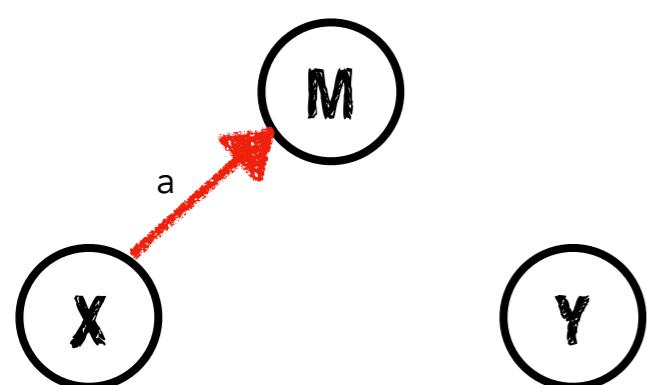
Residual standard error: 4.756 on 97 degrees of freedom
Multiple R-squared:  0.1946, Adjusted R-squared:  0.1779 
F-statistic: 11.72 on 2 and 97 DF,  p-value: 2.771e-05
```

not significant

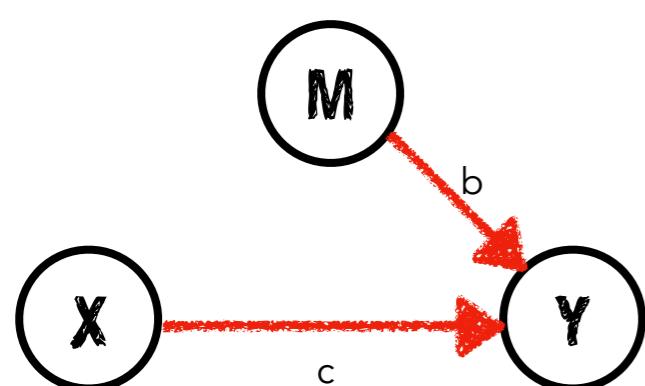
# 3 Step procedure



Relationship between X and Y?  
 $\hat{y} = b_0 + b_1 \cdot x$



Relationship between X and M?  
 $\hat{m} = b_0 + b_1 \cdot x$



Relationship between X and Y,  
controlling for M?

$\hat{y} = b_0 + b_1 \cdot m + b_2 \cdot x$  **significant  
change?**

just because it changes from significant to not  
significant, does not mean the change was significant!

# Is the mediation significant?

## 1. Sobel test

- assumes normally distributed data
- has low power

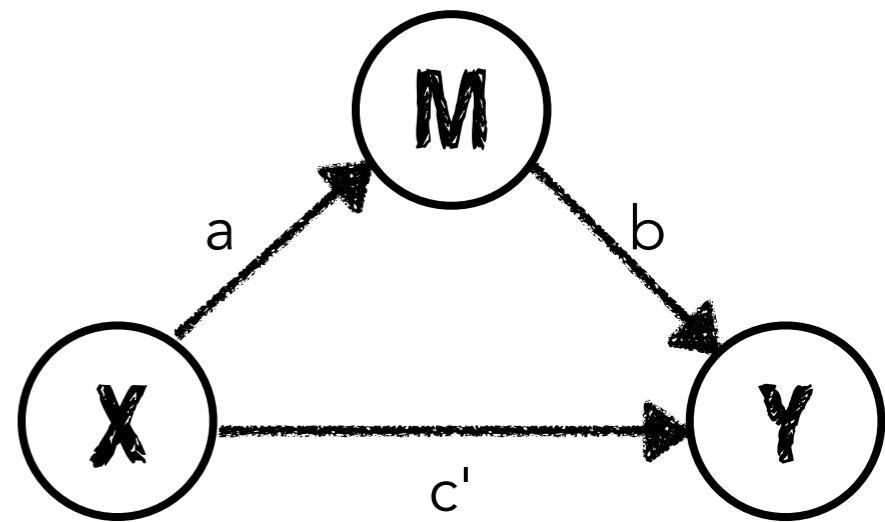
## 2. Bootstrapping

- no assumption about how the data is distributed
- has higher power

## 3. Bayesian mediation analysis

# 1. Sobel test

```
1 library("multilevel")
2
3 # run the sobel test
4 fit.sobel = sobel(pred = df.mediation$x,
5                     med = df.mediation$m,
6                     out = df.mediation$y)
7
8 # calculate the p-value
9 (1 - pnorm(fit.sobel$z.value)) ^ 2
```



$$Z = \frac{ab}{\sqrt{a^2\sigma_b^2 + b^2\sigma_b^2 + \sigma_a^2\sigma_b^2}}$$

*product of the  
coefficients*

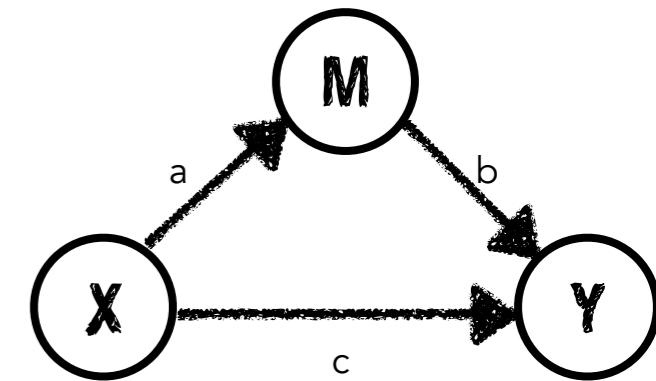
*standard error of  
a x b*

"It is becoming increasingly more difficult to publish tests of mediation based purely on the Baron and Kenny method or tests that make distributional assumptions such as the Sobel test."

A Wikipedia person

## 2. Bootstrapping

```
1 library("mediation")
2
3 # bootstrapped mediation
4 fit.mediation = mediate(model.m = fit.m_x, ←  $\hat{m} = b_0 + b_1 \cdot x$ 
5 model.y = fit.y_mx, ←  $\hat{y} = b_0 + b_1 \cdot m + b_2 \cdot x$ 
6 treat = "x",
7 mediator = "m",
8 boot = T)
9
10 # summarize results
11 fit.mediation %>% summary()
```



```
Causal Mediation Analysis

Nonparametric Bootstrap Confidence Intervals with the Percentile Method

      Estimate 95% CI Lower 95% CI Upper p-value
ACME       0.28078    0.14059        0.42 <2e-16 ***
ADE      -0.11179   -0.29276       0.10     0.272
Total Effect  0.16899   -0.00415       0.34     0.064 .
Prop. Mediated 1.66151   -3.22476      11.46     0.064 .
---
Signif. codes:  0 '****' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Sample Size Used: 100

Simulations: 1000
```

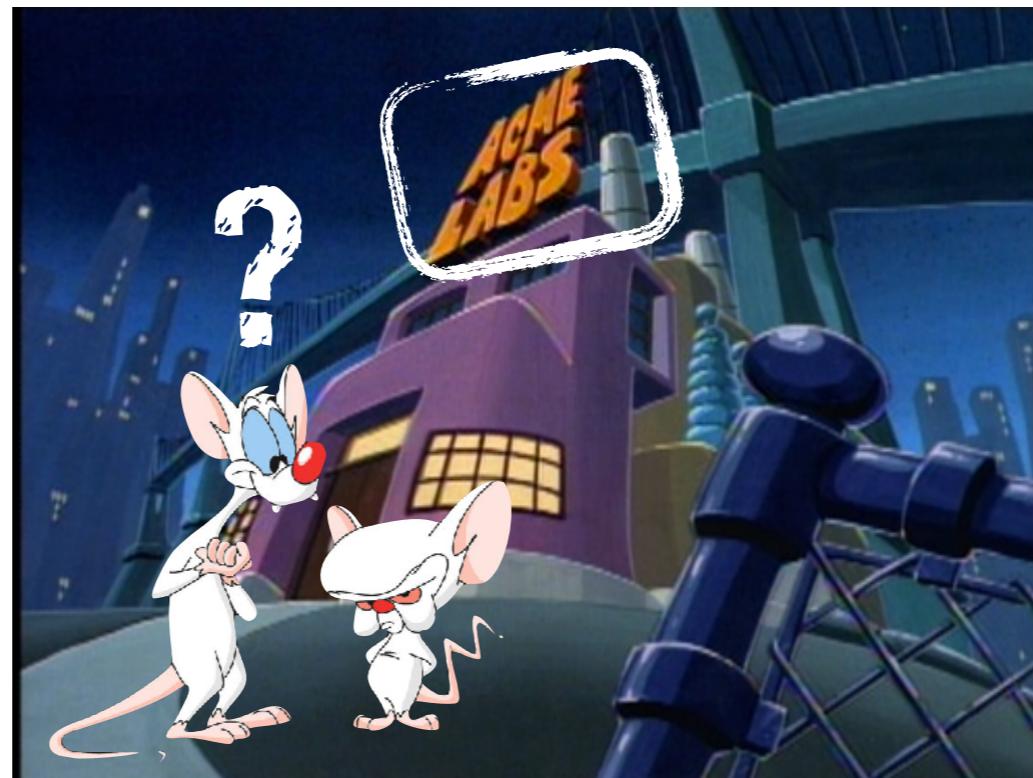
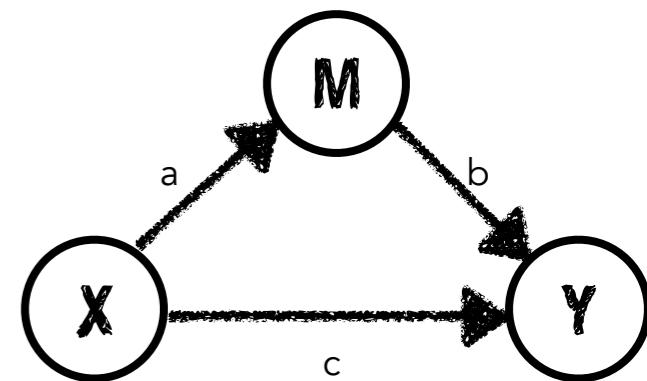
# 2. Bootstrapping

Causal Mediation Analysis

Nonparametric Bootstrap Confidence Intervals with the Percentile Method

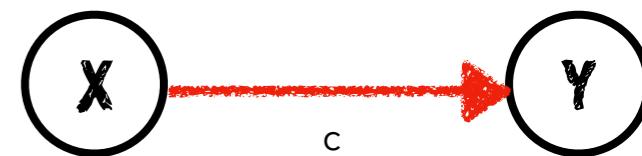
	Estimate	95% CI Lower	95% CI Upper	p-value	
ACME	0.28078	0.14059	0.42	<2e-16	***
ADE	-0.11179	-0.29276	0.10	0.272	
Total Effect	0.16899	-0.00415	0.34	0.064	.
Prop. Mediated	1.66151	-3.22476	11.46	0.064	.
---					
Signif. codes:	0 '***'	0.001 '**'	0.01 '*'	0.05 '.'	0.1 ' '
Sample Size Used:	100				

Simulations: 1000



# 2. Bootstrapping

M



Causal Mediation Analysis

Nonparametric Bootstrap Confidence Intervals with the Percentile Method

	Estimate	95% CI Lower	95% CI Upper	p-value	
ACME	0.28078	0.14059	0.42	<2e-16	***
ADE	-0.11179	-0.29276	0.10	0.272	
Total Effect	0.16899	-0.00415	0.34	0.064	.
Prop. Mediated	1.66151	-3.22476	11.46	0.064	.
---					
Signif. codes:	0 '***'	0.001 '**'	0.01 '*'	0.05 '.'	0.1 ' '
Sample Size Used:	100				

Simulations: 1000

$$\hat{y} = b_0 + b_1 \cdot x$$

Call:

```
lm(formula = y ~ 1 + x, data = df.mediation)
```

Residuals:

Min	1Q	Median	3Q	Max
-10.917	-3.738	-0.259	2.910	12.540

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	8.78300	6.16002	1.426	0.1571
x	0.16899	0.08116	2.082	0.0399 *

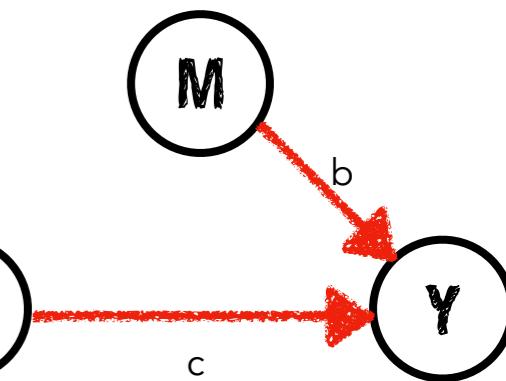
## 2. Bootstrapping

Causal Mediation Analysis

Nonparametric Bootstrap Confidence Intervals with the Percentile Method

	Estimate	95% CI Lower	95% CI Upper	p-value	
ACME	0.28078	0.14059	0.42	<2e-16	***
ADE	-0.11179	-0.29276	0.10	0.272	
Total Effect	0.16899	-0.00415	0.34	0.064	.
Prop. Mediated	1.66151	-3.22476	11.46	0.064	.
---					
Signif. codes:	0 '***'	0.001 '**'	0.01 '*'	0.05 '.'	0.1 ' '
Sample Size Used:	100				

Simulations: 1000



$$\hat{y} = b_0 + b_1 \cdot m + b_2 \cdot x \quad \text{ADE: Average direct effect}$$

Call:

```
lm(formula = y ~ 1 + m + x, data = df.mediation)
```

Residuals:

Min	1Q	Median	3Q	Max
-9.3651	-3.3037	-0.6222	3.1068	10.3991

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	7.80952	5.68297	1.374	0.173
m	0.42381	0.09899	4.281	4.37e-05 ***
x	-0.11179	0.09949	-1.124	0.264

## 2. Bootstrapping

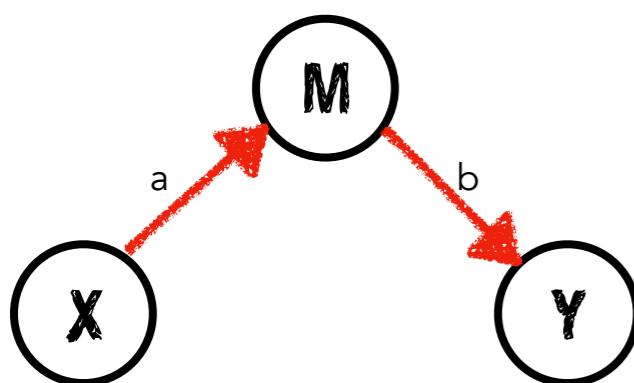
Causal Mediation Analysis

Nonparametric Bootstrap Confidence Intervals with the Percentile Method

	Estimate	95% CI Lower	95% CI Upper	p-value	
ACME	0.28078	0.14059	0.42	<2e-16	***
ADE	-0.11179	-0.29276	0.10	0.272	
Total Effect	0.16899	-0.00415	0.34	0.064	.
Prop. Mediated	1.66151	-3.22476	11.46	0.064	.
---					
Signif. codes:	0 '***'	0.001 '**'	0.01 '*'	0.05 '.'	0.1 ' '
Sample Size Used:	100				

Simulations: 1000

**ACME**



**ADE: Average direct effect**

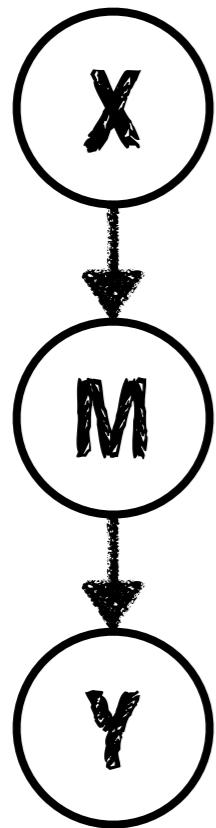
**ACME: Average causal mediation effect**

**ACME = Total effect - ADE**

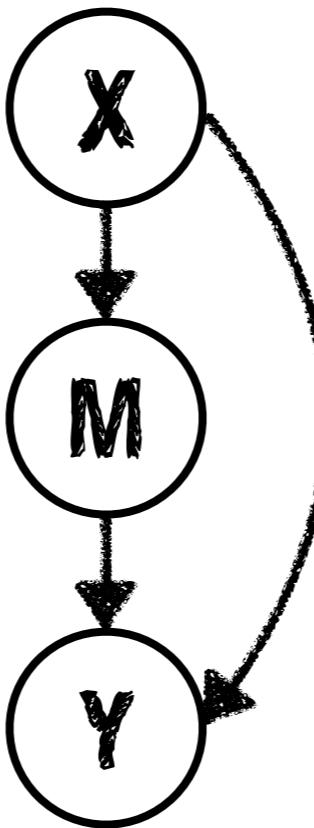
indirect effect:  $a * b$

# Underlying causal model

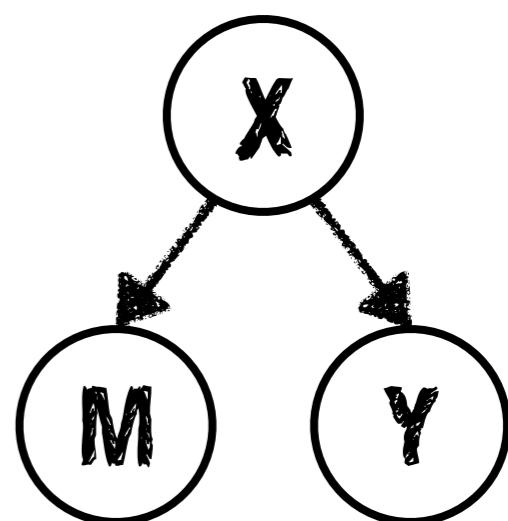
**Full mediation**



**Partial mediation**



**No mediation**

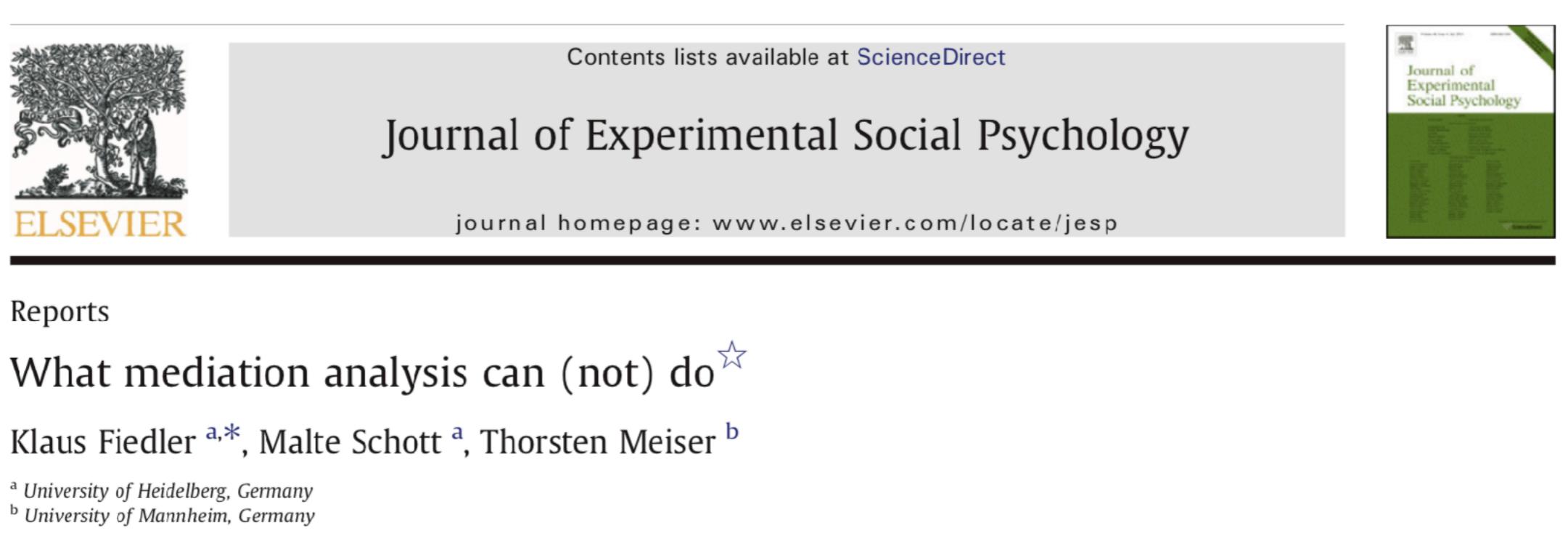


**Full mediation:** When the effect of **X** on **Y** completely disappears, **M** fully mediates between **X** and **Y**.

**Partial mediation:** When the effect of **X** on **Y** still exists, but in a smaller magnitude, **M** partially mediates between **X** and **Y**.

# Limitations

- correlational analysis
  - we need theories / experiments to tease apart causes and effects to properly map our variables onto the diagram



The image shows the cover of a journal article from the Journal of Experimental Social Psychology. The Elsevier logo is on the left, featuring a tree and the word 'ELSEVIER'. The title 'Journal of Experimental Social Psychology' is in the center, with 'Contents lists available at ScienceDirect' above it and 'journal homepage: www.elsevier.com/locate/jesp' below it. To the right is a thumbnail of the journal's cover page.

Reports

What mediation analysis can (not) do<sup>☆</sup>

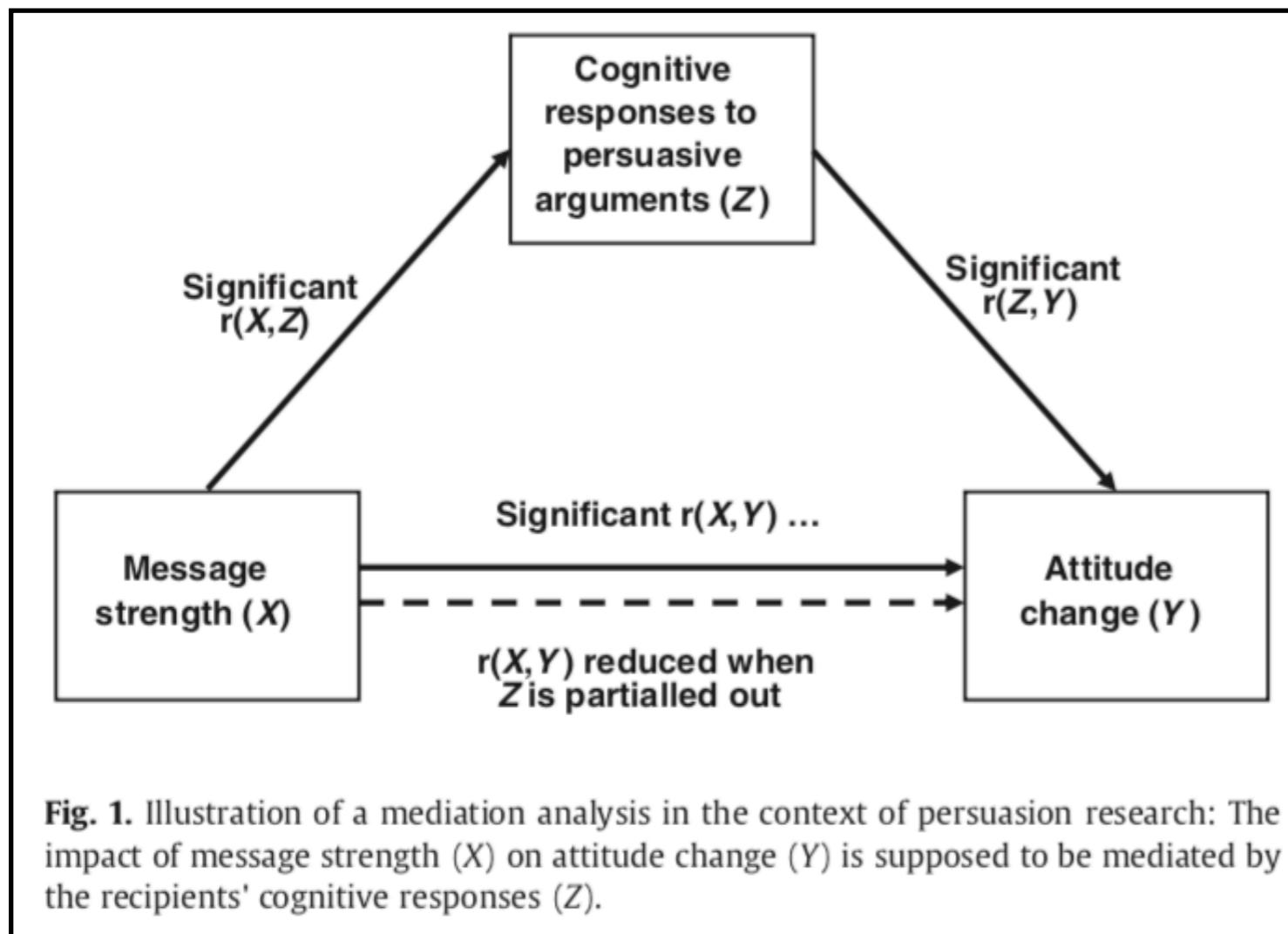
Klaus Fiedler <sup>a,\*</sup>, Malte Schott <sup>a</sup>, Thorsten Meiser <sup>b</sup>

<sup>a</sup> University of Heidelberg, Germany  
<sup>b</sup> University of Mannheim, Germany

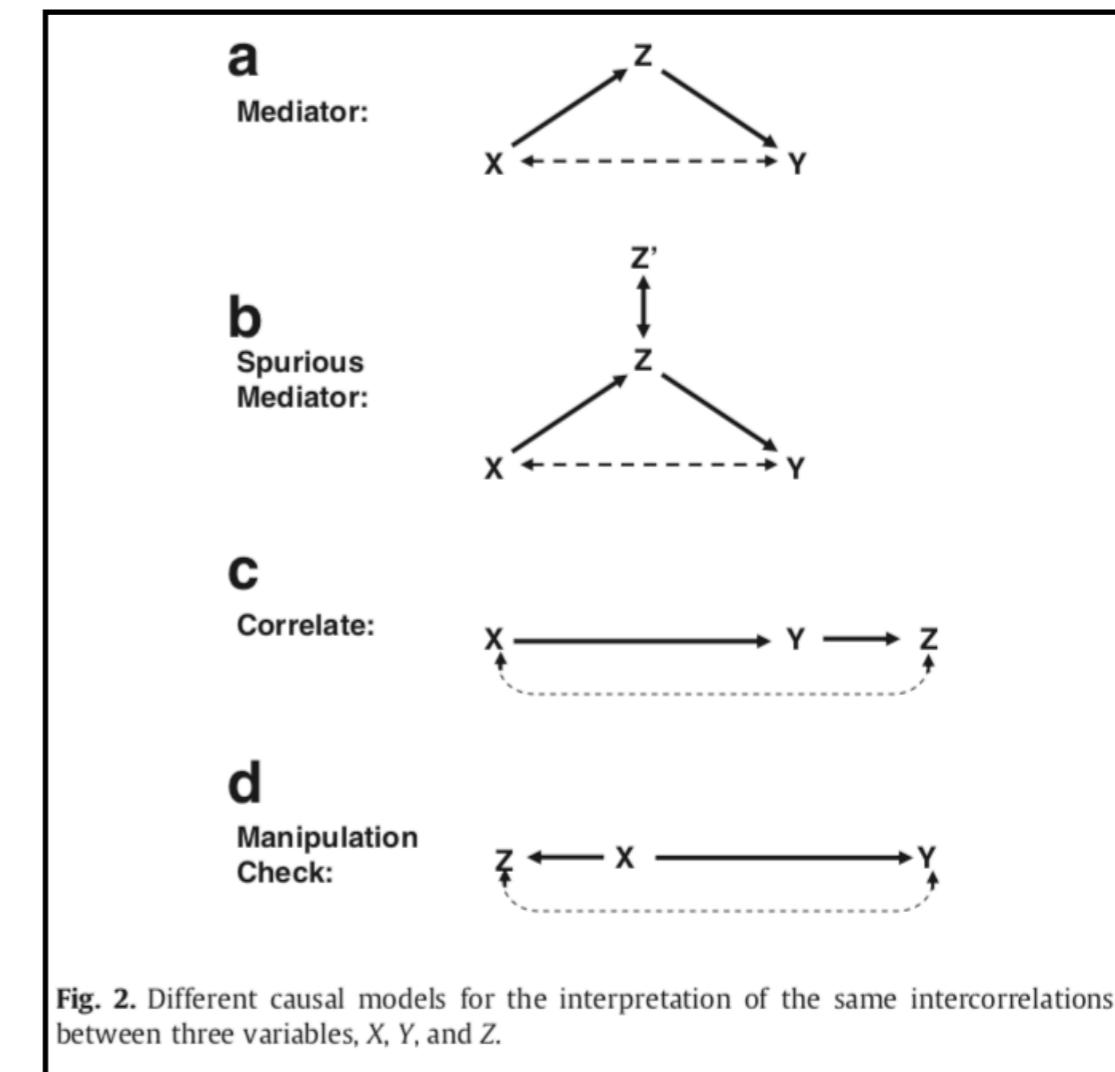
Fiedler, K., Schott, M., & Meiser, T. (2011). What mediation analysis can (not) do. *Journal of Experimental Social Psychology*, 47(6), 1231-1236.

# Limitations

## many-to-one mapping



**Fig. 1.** Illustration of a mediation analysis in the context of persuasion research: The impact of message strength ( $X$ ) on attitude change ( $Y$ ) is supposed to be mediated by the recipients' cognitive responses ( $Z$ ).

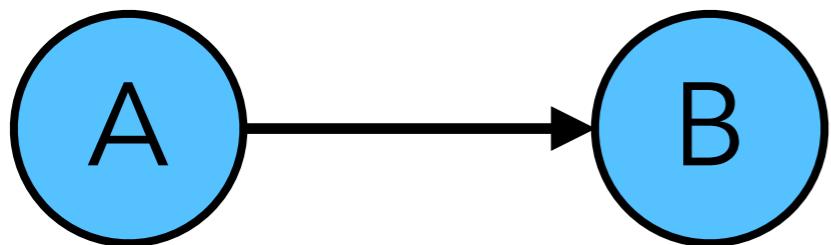
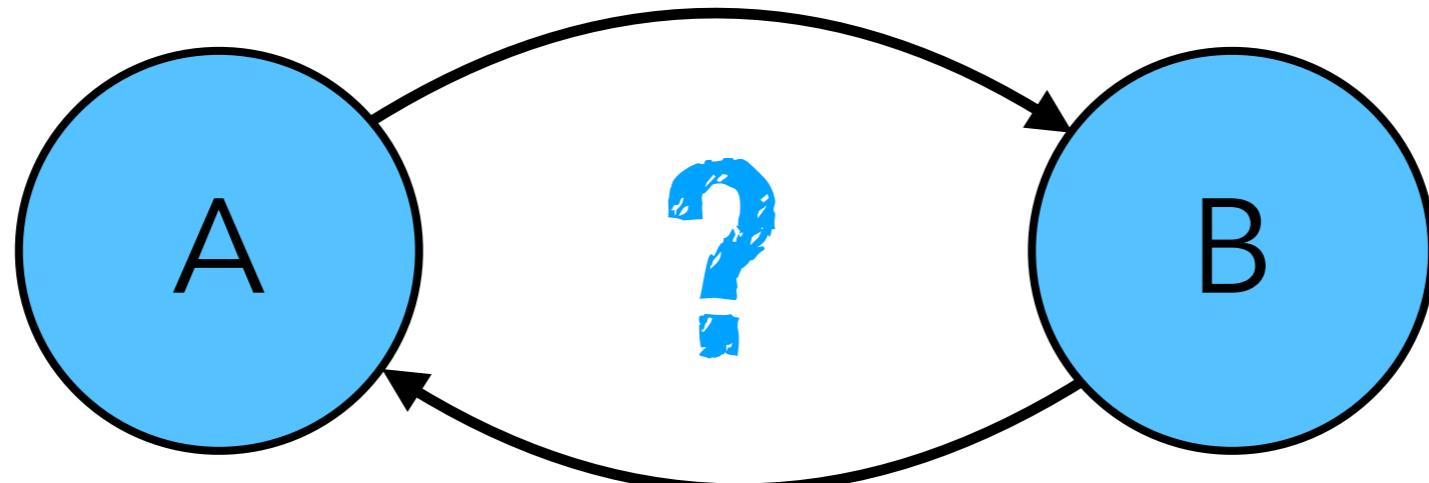


**Fig. 2.** Different causal models for the interpretation of the same intercorrelations between three variables,  $X$ ,  $Y$ , and  $Z$ .

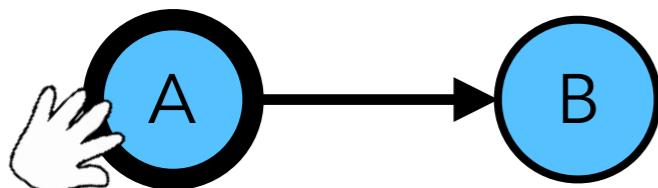
only experiments allow us to tell apart possible causal structures

Fiedler, K., Schott, M., & Meiser, T. (2011). What mediation analysis can (not) do. *Journal of Experimental Social Psychology*, 47(6), 1231-1236.

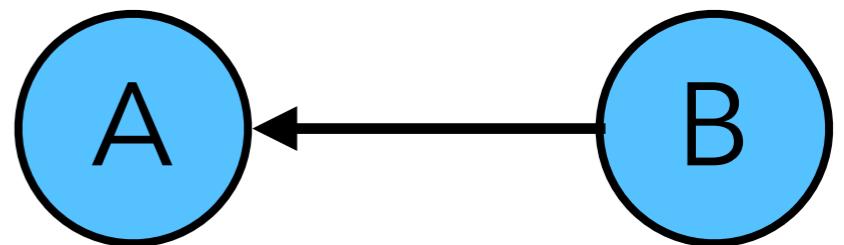
# Inferring causal structure through intervention



$$p(B | \text{do}(A)) = p(B | A)$$



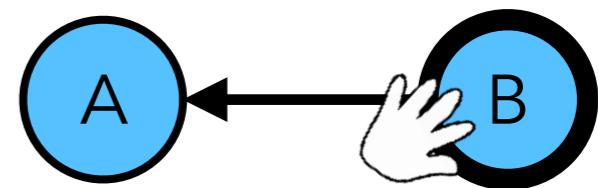
$$p(A | \text{do}(B)) = p(A)$$



$$p(B | \text{do}(A)) = p(B)$$

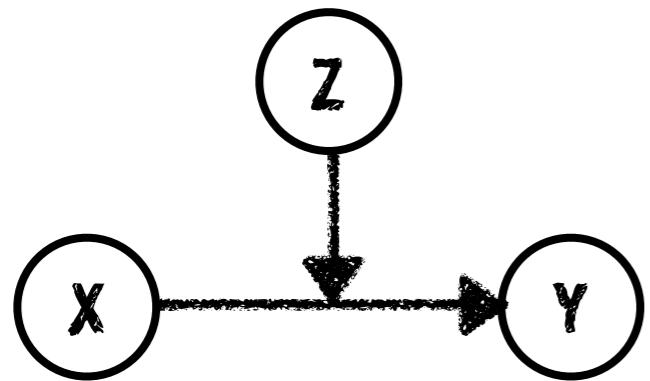


$$p(A | \text{do}(B)) = p(A | B)$$



# Moderation

# Definition

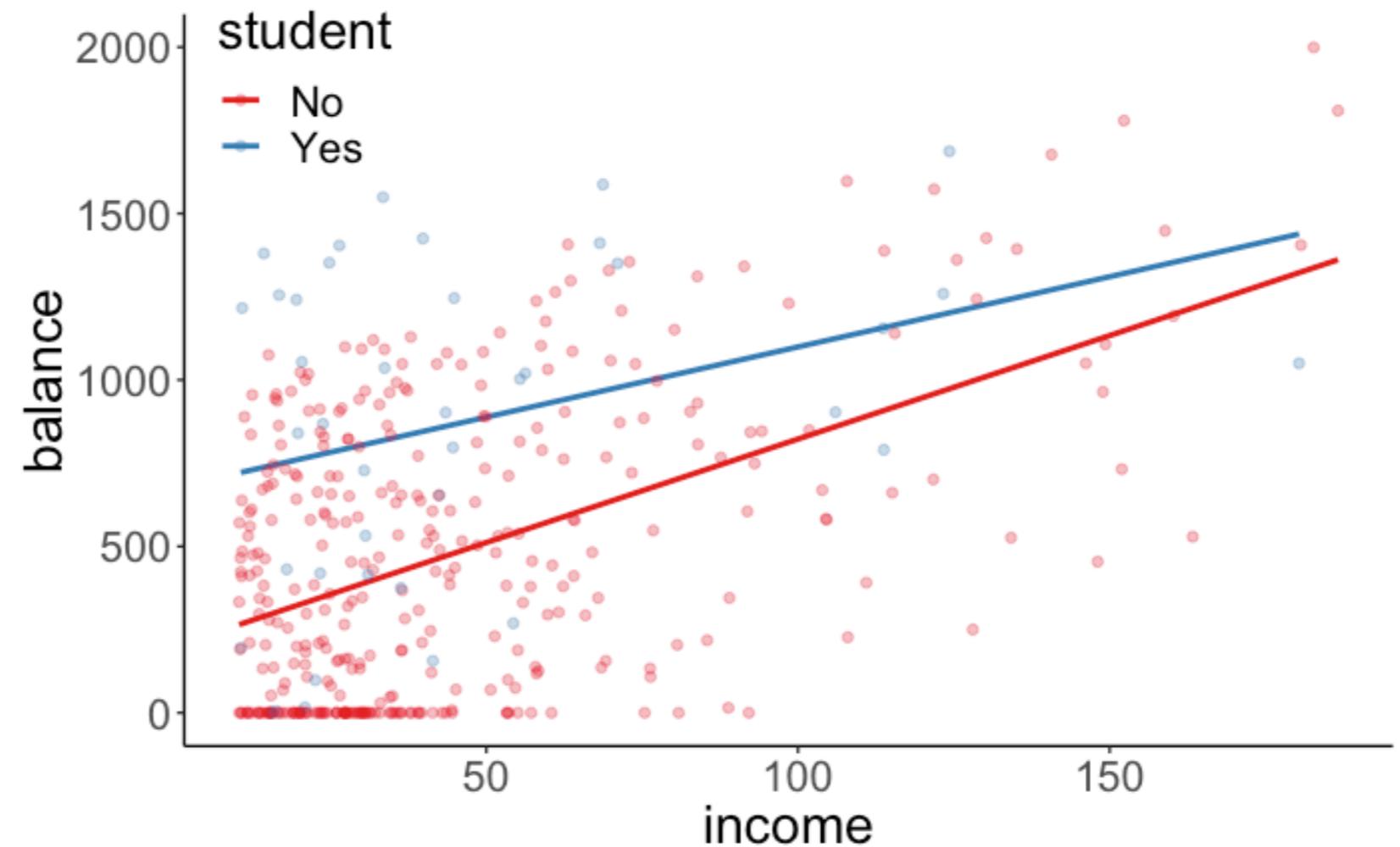


**Moderation** means that the effect of a predictor depends on the value of another.

Here, the nature of the relationship between **X** and **Y** depends on **Z**.

**Have we come across moderation already?**

Relationship  
between credit card  
balance, income,  
and whether the  
person is a student.



$$\widehat{\text{balance}}_i = 200.62 + 6.22 \cdot \text{income}_i + 476.68 \cdot \text{student}_i - 2.00 \cdot (\text{income}_i \times \text{student}_i)$$

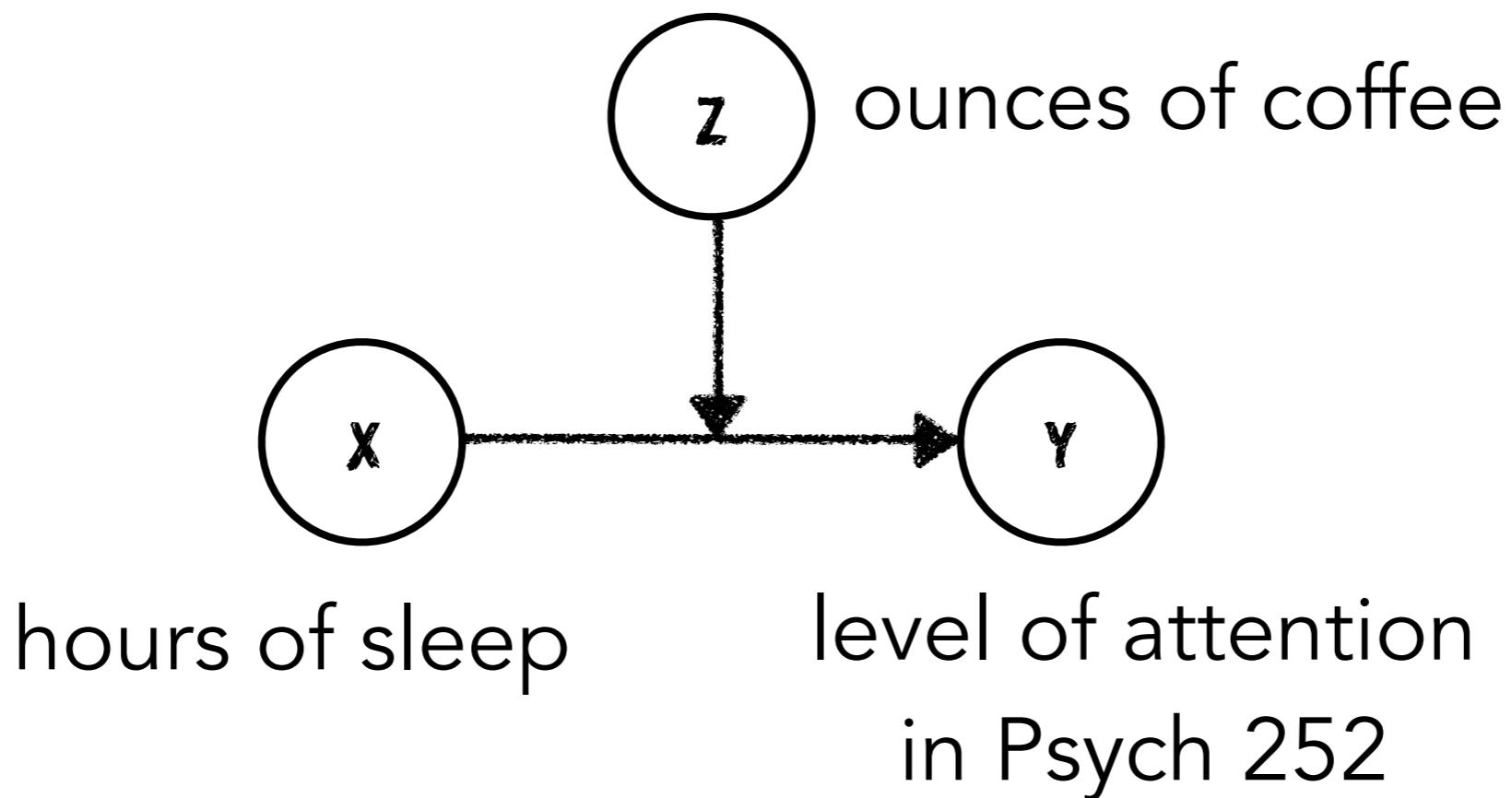
**if student = "No"**  $\widehat{\text{balance}}_i = 200.62 + 6.22 \cdot \text{income}_i$

**if student = "Yes"**

$$\begin{aligned}
 \widehat{\text{balance}}_i &= 200.62 + 6.22 \cdot \text{income}_i + 476.68 \cdot 1 - 2.00 \cdot (\text{income}_i \times 1) \\
 &= 677.3 + 6.22 \cdot \text{income}_i - 2.00 \cdot \text{income}_i \\
 &= 677.3 + 4.22 \cdot \text{income}_i
 \end{aligned}$$

# Simulating a moderation

```
1 # number of participants
2 n = 100
3
4 df.moderation = tibble(
5   x = abs(rnorm(n, 6, 4)), # hours of sleep
6   x1 = abs(rnorm(n, 60, 30)), # adding some systematic variance to our DV
7   z = rnorm(n, 30, 8), # ounces of coffee consumed
8   y = abs((-0.8 * x) * (0.2 * z) - 0.5 * x - 0.4 * x1 + 10 + rnorm(n, 0, 3)) # attention Paid
9 )
```



# Simulating a moderation

```
1 # scale the predictors
2 df.moderation = df.moderation %>%
3   mutate_at(vars(x, z), "scale")
4
5 # run regression model with interaction
6 fit.moderation = lm(formula = y ~ 1 + x * z,
7                      data = df.moderation)
8
9 # summarize result
10 fit.moderation %>%
11   summary()
```

```
Call:
lm(formula = y ~ 1 + x * z, data = df.moderation)

Residuals:
    Min      1Q  Median      3Q     Max 
-21.466 -8.972 -0.233  6.180 38.051 

Coefficients:
            Estimate Std. Error t value Pr(>|t|)    
(Intercept) 48.544     1.173   41.390 < 2e-16 ***
x           17.863     1.196   14.936 < 2e-16 ***
z           8.393     1.181    7.108 2.08e-10 ***
x:z         6.094     1.077    5.656 1.59e-07 ***
---
Signif. codes:  0 '****' 0.001 '***' 0.01 '**' 0.05 '*' 0.1 '.' 1

Residual standard error: 11.65 on 96 degrees of freedom
Multiple R-squared:  0.7661, Adjusted R-squared:  0.7587 
F-statistic: 104.8 on 3 and 96 DF,  p-value: < 2.2e-16
```

significant  
interaction

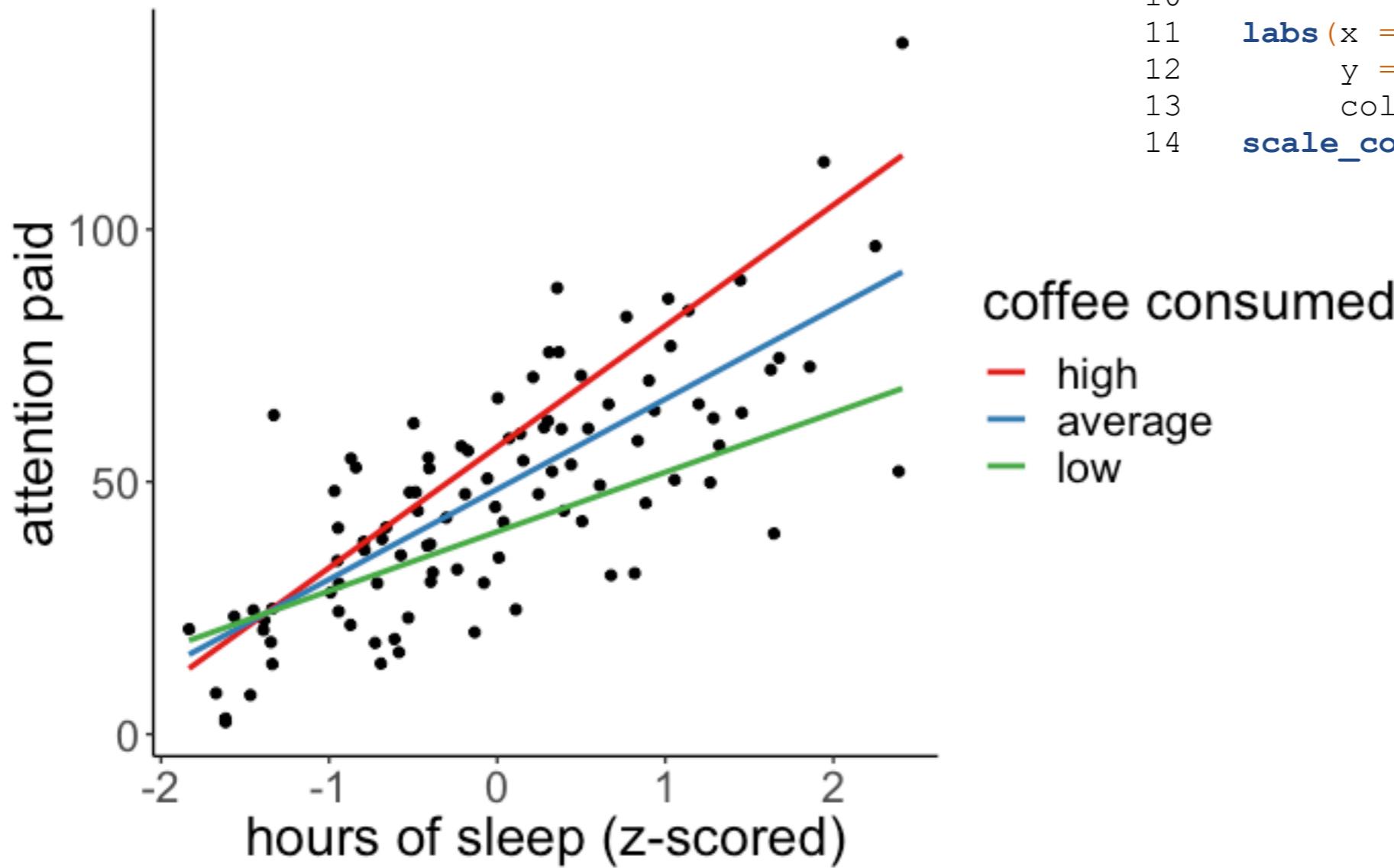
# Simulating a moderation

```
1 # generate data grid with three levels of the moderator
2 df.newdata = df.moderation %>%
3   expand(x = c(min(x),
4                 max(x)),
5         z = c(mean(z) - sd(z),
6                 mean(z),
7                 mean(z) + sd(z))) %>%
8   mutate(moderator = rep(c("low", "average", "high"), nrow(.) / 3))
9
10 # predictions for the three levels of the moderator
11 df.prediction = fit.moderation %>%
12   augment(newdata = df.newdata) %>%
13   mutate(moderator = factor(moderator, levels = c("high", "average", "low")))
```

minimum and maximum → average  $\pm 1SD$

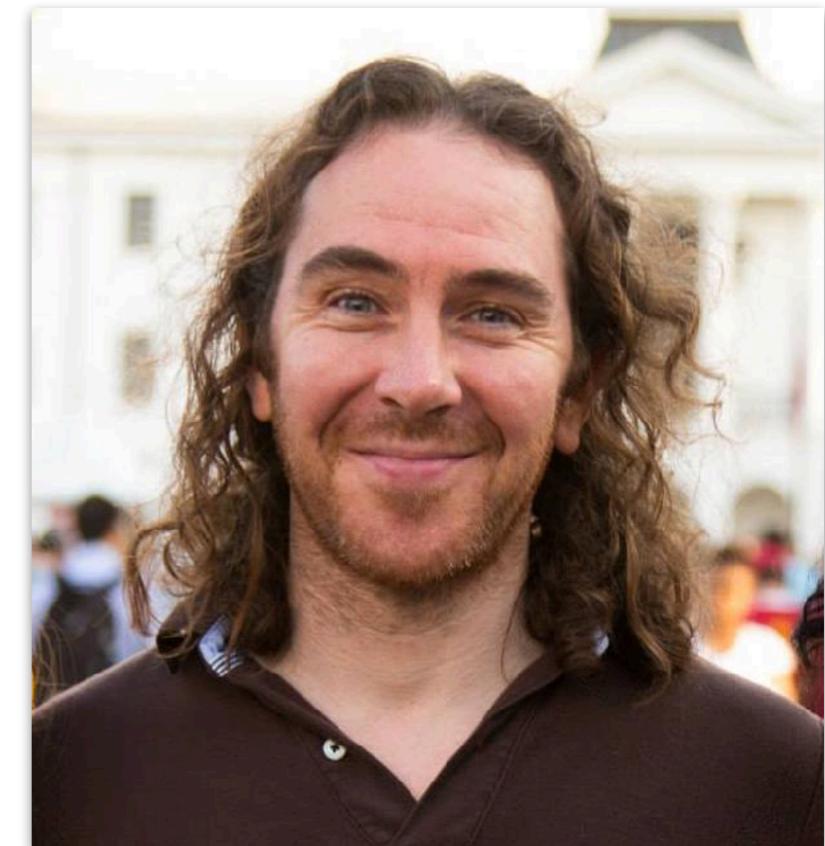
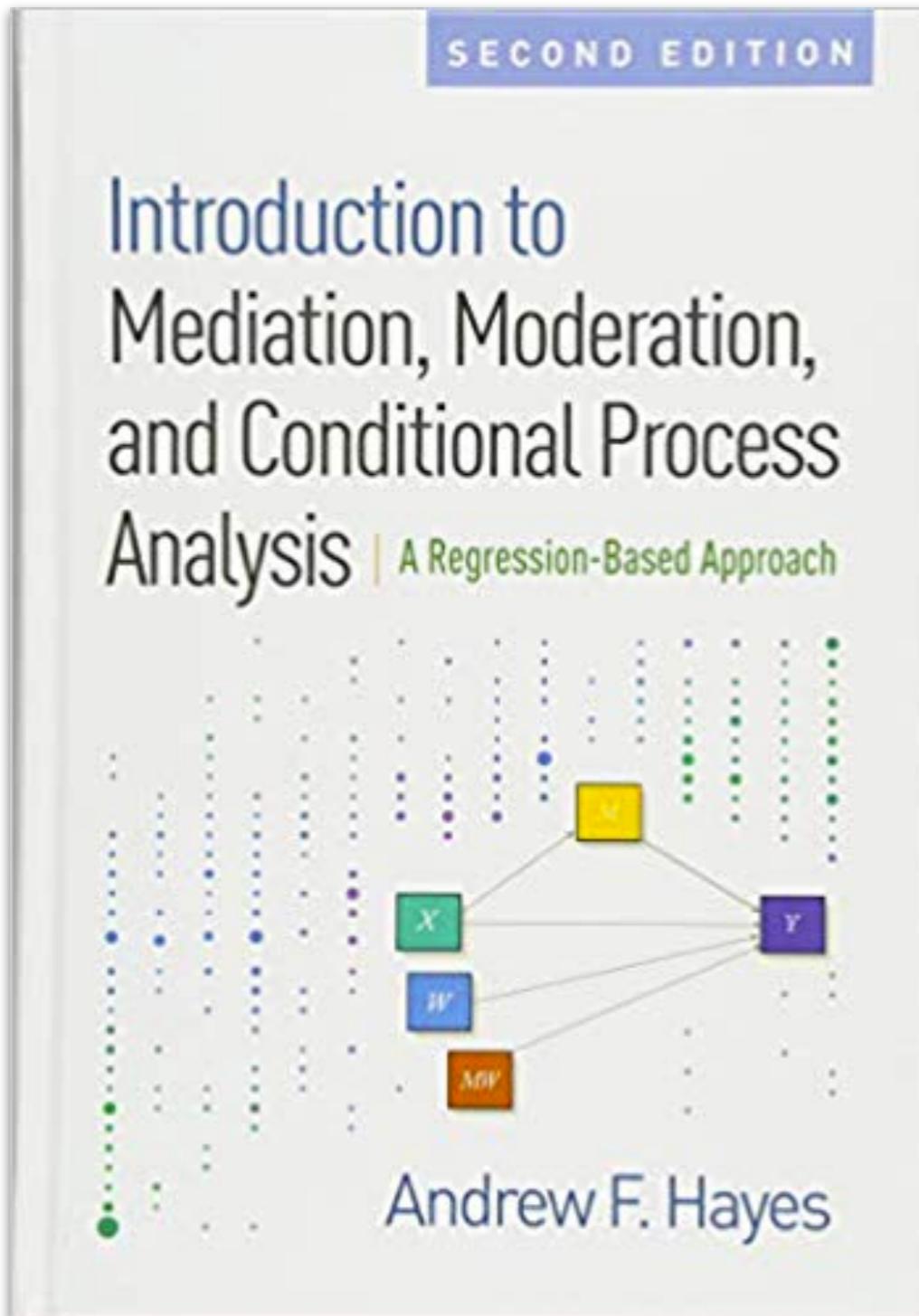
x	z	moderator	.fitted	.se.fit
-1.83	-1	low	18.58	3.75
-1.83	0	average	15.80	2.51
-1.83	1	high	13.02	2.99
2.41	-1	low	68.52	4.32
2.41	0	average	91.60	3.09
2.41	1	high	114.68	4.12

# Simulating a moderation



```
1 # visualize the result
2 df.moderation %>%
3   ggplot(aes(x = x,
4               y = y)) +
5   geom_point() +
6   geom_line(aes(y = .fitted,
7                  group = moderator,
8                  color = moderator),
9             data = df.prediction,
10            size = 1) +
11   labs(x = "hours of sleep (z-scored)",
12         y = "attention paid",
13         color = "coffee consumed") +
14   scale_color_brewer(palette = "Set1")
```

# Learn more about mediation and moderation



Recoded with `brms` by  
Solomon Kurz here:  
[https://bookdown.org/  
connect/#/apps/1523/access](https://bookdown.org/connect/#/apps/1523/access)

# Summary

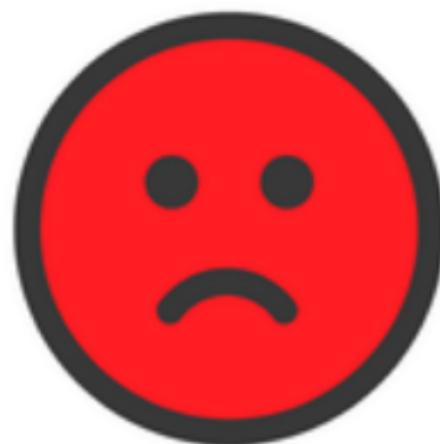
- Questions? Answers!
  - degrees of freedom
  - t-test vs. permutation test
  - overlapping confidence intervals
- Power analysis (continued)
- Controlling for variables
- Mediation
- Moderation

# **Feedback**

# How was the pace of today's class?

much    a little    just    a little    much  
too        too        right      too        too  
slow      slow                                    fast      fast

# How happy were you with today's class overall?



**What did you like about today's class? What could be improved next time?**

Thank you!