

## Resolución

1.1.

```
> library(rstan)
> set.seed(2)
> options(mc.cores = parallel::detectCores())
> rstan_options(auto_write = TRUE)
> x1 <- c(4.8, 6.2, 6.1, 5.1, 4.8, 3, 6)
> x2 <- c(4, 4.7, 4.9, 5.7, 4.2, 2.9, 5.6)
> datos <- list(n = length(x1), x1 = x1, x2 = x2)
> codigo <- "
+   data {
+     int<lower=0> n;
+     real<lower=0> x1[n];
+     real<lower=0> x2[n];
+   }
+
+   parameters {
+     real mu1;
+     real mu2;
+   }
+   model{
+     mu1 ~ normal(0,5);
+     mu2 ~ normal(0,5);
+     for(i in 1:n){
+       x1[i] ~ normal(mu1,1);
+     }
+     for(i in 1:n){
+       x2[i] ~ normal(mu2,1);
+     }
+   }
+   generated quantities{
+     real D;
+     D = 0;
+     for (i in 1:n){
+       D += normal_lpdf(x1[i] | mu1,1);
+       D += normal_lpdf(x2[i] | mu2,1);
+     }
+     D*=-2;
+   }
+   "
> fit <- stan(model_code = codigo, data = datos, iter = 1000)
> print(fit)
```

Inference for Stan model: anon\_model.

4 chains, each with iter=1000; warmup=500; thin=1;  
post-warmup draws per chain=500, total post-warmup draws=2000.

	mean	se_mean	sd	2.5%	25%	50%	75%	97.5%
mu1	5.12	0.01	0.39	4.37	4.84	5.12	5.39	5.87
mu2	4.56	0.01	0.38	3.83	4.30	4.56	4.83	5.30
D	41.13	0.07	2.01	39.11	39.68	40.47	41.93	46.57
lp_	-8.64	0.04	1.01	-11.33	-9.03	-8.31	-7.91	-7.63

	n_eff	Rhat
mu1	1776	1
mu2	1709	1
D	734	1
lp_	736	1

Samples were drawn using NUTS(diag\_e) at Wed Apr 19 14:10:36 2023.  
For each parameter, n\_eff is a crude measure of effective sample size,  
and Rhat is the potential scale reduction factor on split chains (at  
convergence, Rhat=1).

```
> D1<- unlist(extract(fit, pars='D'))
> DIC1 <- mean(D1) + 0.5 *var(D1)
> print(DIC1)
```

```
[1] 43.15062
```

1.2.

```
> x1 <- c(4.8, 6.2, 6.1, 5.1, 4.8, 3, 6)
> x2 <- c(4, 4.7, 4.9, 5.7, 4.2, 2.9, 5.6)
> datos <- list(n = length(x1), x1 = x1, x2 = x2)
> codigo <- "
+ data {
+   int<lower=0> n;
+   real<lower=0> x1[n];
+   real<lower=0> x2[n];
+ }
+
+ parameters {
+   real mu;
+ }
+ model{
+   mu ~ normal(0,5);
+   for(i in 1:n){
+     x1[i] ~ normal(mu,1);
+   }
+   for(i in 1:n){
```

```

+           x2[i] ~ normal(mu,1);
+       }
+   }
+   generated quantities{
+     real D;
+     D = 0;
+     for (i in 1:n){
+       D += normal_lpdf(x1[i] | mu,1);
+       D += normal_lpdf(x2[i] | mu,1);
+     }
+     D*=-2;
+   }
+   "
> fit <- stan(model_code = codigo, data = datos, iter=1000)
> print(fit)

```

Inference for Stan model: anon\_model.

4 chains, each with iter=1000; warmup=500; thin=1;

post-warmup draws per chain=500, total post-warmup draws=2000.

	mean	se_mean	sd	2.5%	25%	50%	75%	97.5%
mu	4.84	0.01	0.27	4.32	4.67	4.84	5.03	5.36
D	41.18	0.05	1.36	40.19	40.29	40.65	41.52	44.98
lp_	-8.19	0.02	0.68	-10.11	-8.39	-7.93	-7.75	-7.70

	n_eff	Rhat
mu	706	1
D	834	1
lp_	830	1

Samples were drawn using NUTS(diag\_e) at Wed Apr 19 14:11:33 2023.

For each parameter, n\_eff is a crude measure of effective sample size, and Rhat is the potential scale reduction factor on split chains (at convergence, Rhat=1).

```

> D2<- unlist(extract(fit, pars='D'))
> DIC2 <- mean(D2) + 0.5 *var(D2)
> DIC2 <- mean(D2) + 0.5 *var(D1)
> print(DIC2)

```

```
[1] 43.20344
```

1.3.

```

> cat(sprintf("\nModelo 1. mean D = %8.2f. pD = %8.2f. DIC = %8.2f\n",
+ Modelo 2. mean D = %8.2f. pD = %8.2f. DIC = %8.2f\n",
+ mean(D1), 0.5*var(D1), DIC1, mean(D2), 0.5*var(D2), DIC2))

```

Modelo 1. mean D = 41.13. pD = 2.02. DIC = 43.15

Modelo 2. mean D = 41.18. pD = 0.93. DIC = 43.20

El modelo 2 presenta peor ajuste, pero es más sencillo. Sin embargo, su IC posterior para  $\mu$  no incluye el 0 por lo que con un 95% de confianza no podemos asumir que la diferencia entre ambos grupos sea nula. Elijo por lo tanto el modelo 1.

Conclusión: El rendimiento en orientación espacial es mayor en hombres que en mujeres.

2.1.

```
> y <- c(93, 66, 42, 45, 38, 32, 36, 43, 40, 53)
> x <- c(1, 2, 3, 4, 5, 6, 7, 8, 9, 10)
> datos <- list(n = length(x), x = x, y = y)
> codigo <- "
+   data {
+     int<lower=0> n;
+     real<lower=0> x[n];
+     real<lower=0> y[n];
+   }
+
+   parameters {
+     real phi;
+     real tau;
+     real omega;
+   }
+   model{
+     real mu;
+     real alpha;
+     real beta;
+
+     tau ~ normal(0,5);
+     omega ~ normal(0,5);
+     for(i in 1:n){
+       mu = exp(tau+omega*x[i]);
+       alpha = mu*phi;
+       beta = phi;
+       y[i] ~ gamma(alpha,beta);
+     }
+   }
+   generated quantities{
+     real D;
+     D = 0;
```

```

+   for (i in 1:n){
+     D += gamma_lpdf(y[i] | exp(tau+omega*x[i])*phi,phi);
+   }
+   D*=-2;
+ }
+ "
> fit <- stan(model_code = codigo, data = datos, iter = 2000)
> print(fit)

```

Inference for Stan model: anon\_model.

4 chains, each with iter=2000; warmup=1000; thin=1;

post-warmup draws per chain=1000, total post-warmup draws=4000.

	mean	se_mean	sd	2.5%	25%	50%	75%
phi	0.31	0.00	0.14	0.10	0.21	0.29	0.40
tau	4.24	0.01	0.19	3.86	4.12	4.24	4.36
omega	-0.07	0.00	0.03	-0.14	-0.09	-0.07	-0.05
D	81.53	0.07	2.40	78.72	79.78	80.92	82.66
lp__	-41.12	0.03	1.20	-44.21	-41.69	-40.83	-40.25

	97.5%	n_eff	Rhat
phi	0.62	925	1
tau	4.61	1105	1
omega	0.00	1069	1
D	87.69	1347	1
lp__	-39.72	1342	1

Samples were drawn using NUTS(diag\_e) at Wed Apr 19 14:12:32 2023.

For each parameter, n\_eff is a crude measure of effective sample size, and Rhat is the potential scale reduction factor on split chains (at convergence, Rhat=1).

```

> D1<- unlist(extract(fit, pars='D'))
> DIC1 <- mean(D1) + 0.5 *var(D1)
> print(DIC1)

```

[1] 84.39532

2.2.

```

> y <- c(93, 66, 42, 45, 38, 32, 36, 43, 40, 53)
> x <- c(1, 2, 3, 4, 5, 6, 7, 8, 9, 10)
> datos <- list(n = length(x), x = x, y = y)
> codigo <- "
+   data {
+     int<lower=0> n;
+     real<lower=0> x[n];

```

```

+   real<lower=0> y[n];
+ }
+
+ parameters {
+   real phi;
+   real tau;
+   real omega1;
+   real omega2;
+ }
+ model{
+   real mu;
+   real alpha;
+   real beta;
+
+   tau ~ normal(0,5);
+   omega1 ~ normal(0,5);
+   omega2 ~ normal(0,5);
+   for(i in 1:n){
+     mu = exp(tau+omega1*x[i]+omega2*(x[i])^2);
+     alpha = mu*phi;
+     beta = phi;
+     y[i] ~ gamma(alpha,beta);
+   }
+ }
+ generated quantities{
+   real D;
+   D = 0;
+   for (i in 1:n){
+     D += gamma_lpdf(y[i] | exp(tau+omega1*x[i]+omega2*(x[i])^2)*phi,phi);
+   }
+   D*= -2;
+ }
+ "
> fit <- stan(model_code = codigo, data = datos, iter = 2000)
> print(fit)

```

Inference for Stan model: anon\_model.

4 chains, each with iter=2000; warmup=1000; thin=1;

post-warmup draws per chain=1000, total post-warmup draws=4000.

	mean	se_mean	sd	2.5%	25%	50%	75%
phi	2.69	0.04	1.25	0.82	1.77	2.46	3.45
tau	4.89	0.00	0.09	4.71	4.84	4.89	4.95
omega1	-0.43	0.00	0.04	-0.51	-0.45	-0.43	-0.40
omega2	0.03	0.00	0.00	0.03	0.03	0.03	0.04

D	59.49	0.09	2.91	55.86	57.36	58.84	60.90
lp_	-30.23	0.04	1.45	-34.09	-30.93	-29.91	-29.16
	97.5%	n_eff	Rhat				
phi	5.57	840	1.00				
tau	5.07	995	1.00				
omega1	-0.34	909	1.00				
omega2	0.04	976	1.00				
D	67.22	1071	1.01				
lp_	-28.42	1071	1.01				

Samples were drawn using NUTS(diag\_e) at Wed Apr 19 14:13:34 2023.  
 For each parameter, n\_eff is a crude measure of effective sample size,  
 and Rhat is the potential scale reduction factor on split chains (at  
 convergence, Rhat=1).

```
> D2<- unlist(extract(fit, pars='D'))
> DIC2 <- mean(D1) + 0.5 *var(D2)
> print(DIC2)
```

```
[1] 85.75326
```

2.3.

```
> y <- c(93, 66, 42, 45, 38, 32, 36, 43, 40, 53)
> x <- c(1, 2, 3, 4, 5, 6, 7, 8, 9, 10)
> dados1 <- list(n = length(x), x = x, y = y, modelo=1)
> dados2 <- list(n = length(x), x = x, y = y, modelo=2)
> codigo <- "
+   data {
+     int<lower=0> n;
+     real<lower=0> x[n];
+     real<lower=0> y[n];
+     int modelo;
+   }
+
+   parameters {
+     real phi;
+     real tau;
+     real omega1;
+     real omega2;
+   }
+   model{
+     real mu;
+     real alpha;
+     real beta;
+   }
```

```

+
+   tau ~ normal(0,5);
+   omega1 ~ normal(0,5);
+   omega2 ~ normal(0,5);
+
+   for(i in 1:n){
+       if(modelo==1){
+           mu = exp(tau+omega1*x[i]);
+       }else{
+           mu = exp(tau+omega1*x[i]+omega2*(x[i])^2);
+       }
+       alpha = mu*phi;
+       beta = phi;
+       y[i] ~ gamma(alpha,beta);
+   }
+ }
+ generated quantities{
+   real D;
+   real alpha_D;
+   real beta_D;
+   beta_D = phi;
+   real yPred[n];
+   D = 0;
+   for (i in 1:n){
+       if(modelo == 1){
+           alpha_D = exp(tau+omega1*x[i])*phi;
+       } else {
+           alpha_D = exp(tau+omega1*x[i]+omega2*(x[i])^2)*phi;;
+       }
+       yPred[i] = gamma_rng(alpha_D, beta_D);
+       D += gamma_lpdf(y[i] | alpha_D,beta_D);
+   }
+   D*= -2;
+ }
+ "
+ fit1 <- stan(model_code = codigo, data = datos1, iter = 2000)
+ fit2 <- stan(model_code = codigo, data = datos2, iter = 2000)
+ D1 <- unlist(extract(fit1, pars='D'))
+ D2 <- unlist(extract(fit2, pars='D'))
+ DIC1 <- mean(D1) + 0.5 * var(D1)
+ DIC2 <- mean(D2) + 0.5 * var(D2)
+ cat(sprintf("\nModelo 1. mean D = %8.2f. pD = %8.2f. DIC = %8.2f\n",
+   Modelo 2. mean D = %8.2f. pD = %8.2f. DIC = %8.2f\n",
+   mean(D1), 0.5*var(D1), DIC1, mean(D2), 0.5*var(D2), DIC2))
+
Modelo 1. mean D =      81.40. pD =       2.81. DIC =      84.21

```



Modelo 2. mean D = 59.52. pD = 4.35. DIC = 63.87

```
> print(fit1)
```

Inference for Stan model: anon\_model.

4 chains, each with iter=2000; warmup=1000; thin=1;

post-warmup draws per chain=1000, total post-warmup draws=4000.

	mean	se_mean	sd	2.5%	25%	50%	75%
phi	0.31	0.00	0.13	0.10	0.21	0.29	0.38
tau	4.23	0.00	0.19	3.84	4.12	4.24	4.35
omega1	-0.07	0.00	0.03	-0.13	-0.09	-0.07	-0.05
omega2	-0.03	0.10	5.07	-9.88	-3.36	-0.05	3.36
D	81.40	0.06	2.37	78.70	79.63	80.77	82.49
alpha_D	10.83	0.13	4.85	3.49	7.30	10.15	13.66
beta_D	0.31	0.00	0.13	0.10	0.21	0.29	0.38
yPred[1]	65.61	0.35	19.45	32.23	52.71	63.86	76.21
yPred[2]	60.91	0.30	17.91	30.90	48.86	59.09	70.80
yPred[3]	56.32	0.26	15.85	27.99	45.86	55.02	65.14
yPred[4]	52.92	0.25	15.89	26.08	42.27	51.56	61.50
yPred[5]	49.48	0.24	15.21	24.53	39.37	48.12	57.77
yPred[6]	46.08	0.22	14.07	22.26	36.45	44.71	54.14
yPred[7]	43.19	0.23	14.26	19.50	33.61	41.69	51.00
yPred[8]	40.40	0.23	13.71	17.56	31.07	38.81	48.09
yPred[9]	38.05	0.24	14.27	15.35	28.37	36.40	45.82
yPred[10]	35.50	0.26	13.99	12.93	26.24	34.00	43.14
lp__	-41.57	0.03	1.39	-45.14	-42.30	-41.24	-40.54

	97.5%	n_eff	Rhat
phi	0.63	1354	1
tau	4.59	1848	1
omega1	0.00	1872	1
omega2	10.25	2712	1
D	87.80	1689	1
alpha_D	22.44	1497	1
beta_D	0.63	1354	1
yPred[1]	108.00	3032	1
yPred[2]	100.03	3499	1
yPred[3]	91.51	3683	1
yPred[4]	88.89	3959	1
yPred[5]	82.09	3964	1
yPred[6]	77.20	4059	1
yPred[7]	76.08	3970	1
yPred[8]	71.71	3540	1
yPred[9]	69.79	3579	1
yPred[10]	67.83	2924	1

```
lp__      -39.86  1672    1
```

Samples were drawn using NUTS(diag\_e) at Wed Apr 19 14:14:56 2023.  
 For each parameter, n\_eff is a crude measure of effective sample size,  
 and Rhat is the potential scale reduction factor on split chains (at  
 convergence, Rhat=1).

```
> print(fit2)
```

Inference for Stan model: anon\_model.  
 4 chains, each with iter=2000; warmup=1000; thin=1;  
 post-warmup draws per chain=1000, total post-warmup draws=4000.

	mean	se_mean	sd	2.5%	25%	50%	75%
phi	2.88	0.08	1.37	0.89	1.86	2.68	3.68
tau	4.90	0.00	0.10	4.70	4.84	4.90	4.96
omega1	-0.43	0.00	0.05	-0.52	-0.46	-0.43	-0.40
omega2	0.03	0.00	0.00	0.02	0.03	0.03	0.04
D	59.52	0.09	2.95	55.80	57.32	58.88	60.99
alpha_D	154.75	4.42	74.20	46.34	99.31	142.88	197.93
beta_D	2.88	0.08	1.37	0.89	1.86	2.68	3.68
yPred[1]	90.71	0.20	8.45	74.27	85.28	90.50	96.00
yPred[2]	65.15	0.10	5.84	53.89	61.46	64.90	68.63
yPred[3]	50.16	0.09	5.23	40.44	46.68	49.97	53.39
yPred[4]	41.57	0.08	4.70	32.56	38.44	41.47	44.39
yPred[5]	36.58	0.07	4.40	28.44	33.69	36.42	39.18
yPred[6]	34.58	0.08	4.22	26.97	31.83	34.44	37.16
yPred[7]	34.90	0.08	4.32	26.88	32.03	34.75	37.44
yPred[8]	37.73	0.07	4.45	29.51	34.87	37.61	40.48
yPred[9]	43.37	0.08	4.88	33.86	40.18	43.29	46.38
yPred[10]	53.59	0.11	6.11	41.98	49.61	53.38	57.39
lp__	-30.24	0.05	1.47	-33.92	-30.98	-29.93	-29.14

	97.5%	n_eff	Rhat
phi	6.22	281	1.01
tau	5.09	1004	1.00
omega1	-0.33	972	1.00
omega2	0.04	1034	1.00
D	66.79	1044	1.00
alpha_D	329.07	282	1.01
beta_D	6.22	281	1.01
yPred[1]	107.84	1701	1.00
yPred[2]	77.10	3526	1.00
yPred[3]	60.82	3706	1.00
yPred[4]	51.51	3139	1.00
yPred[5]	45.61	3510	1.00
yPred[6]	42.92	3148	1.00

```

yPred[7]    44.13   3126  1.00
yPred[8]    46.97   3842  1.00
yPred[9]    53.16   4101  1.00
yPred[10]   66.11   3171  1.00
lp__        -28.38   1044  1.00

```

Samples were drawn using NUTS(diag\_e) at Wed Apr 19 14:15:16 2023.  
 For each parameter, n\_eff is a crude measure of effective sample size,  
 and Rhat is the potential scale reduction factor on split chains (at  
 convergence, Rhat=1).

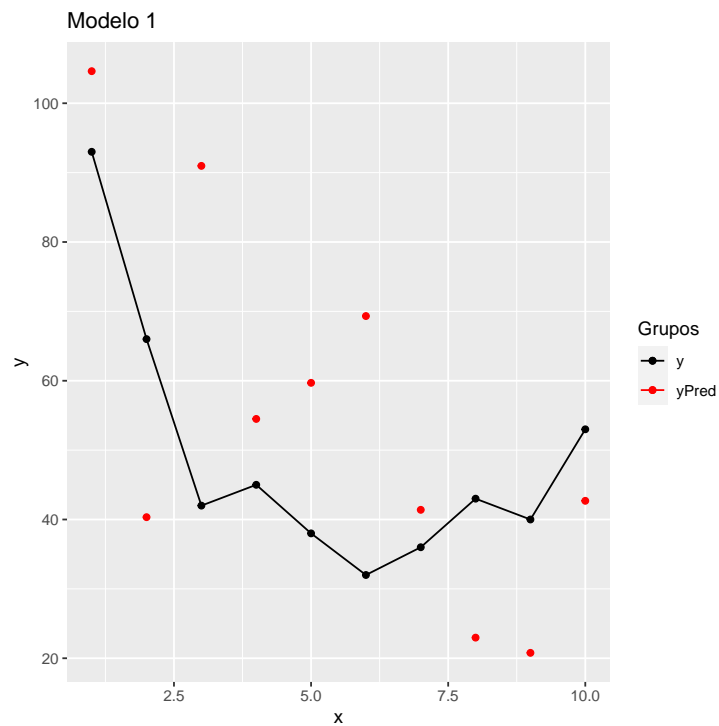
2.4.

```

> library(ggplot2)
> yPred_1 <- extract(fit1,"yPred")
> ## Empleo los últimos 10 valores de la última iteración
> yPred_1 <- yPred_1[[1]]
> yPred_1 <- yPred_1[4000,1:10]
> yPred_2 <- extract(fit2,"yPred")
> yPred_2 <- yPred_2[[1]]
> yPred_2 <- yPred_2[4000,1:10]
> data_1 <- as.data.frame( cbind(x,y,yPred_1))
> data_2 <- as.data.frame(cbind(x,y,yPred_2))
> ggplot(data_1, aes(x, y)) +
+   geom_point(aes(color = "y")) +
+   geom_line(aes(color = "y")) +
+   geom_point(aes(x, yPred_1, color = "yPred")) +
+   scale_color_manual(name = "Grupos",
+
+   labs(color = "Leyenda") +
+   ggtitle("Modelo 1")

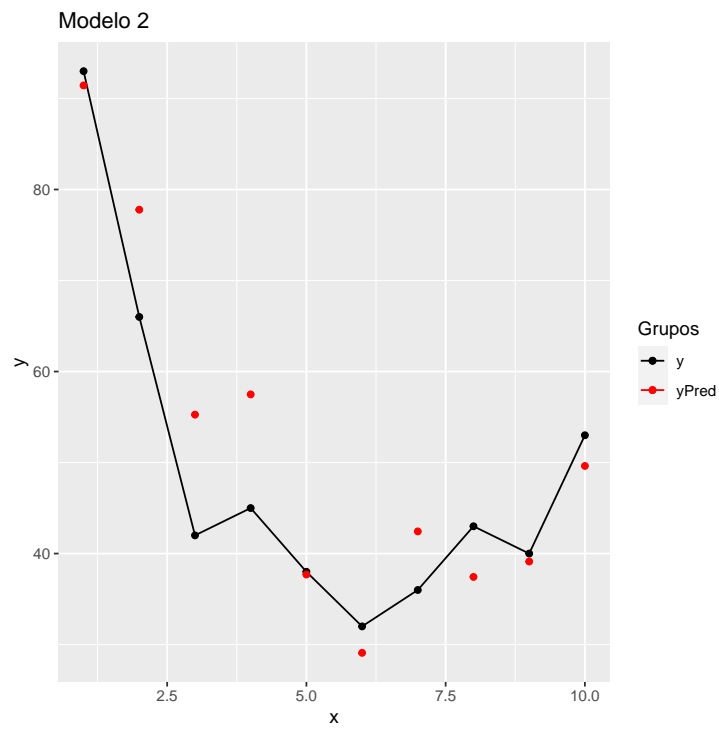
```

values =



```
> ggplot(data_2, aes(x, y)) +
+   geom_point(aes(color = "y")) +
+   geom_line(aes(color = "y")) +
+   geom_point(aes(x, yPred_2, color = "yPred")) +
+   scale_color_manual(name = "Grupos",
+
+   labs(color = "Leyenda") +
+   ggtitle("Modelo 2")
>
```

values =



2.5.

El modelo 1 es más sencillo, pero presenta un ajuste considerablemente menor.  
Elijo el modelo 2 pues.