

Gaussian processes for Computational Biology

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Computational Biology and Bioinformatics
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October 5, 2017

Introduction

Background

Statistics

A simple example

Bayesian linear
regression

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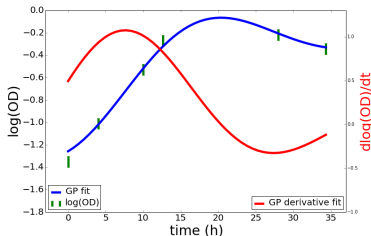
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Conclusions

- ▶ 6th year PhD student
- ▶ Computational Biology and Bioinformatics program
- ▶ Advisors: Amy Schmid and Scott Schmidler
- ▶ Defense: November 14th

Why Gaussian processes?



- ▶ Tonner *et al.*, “Detecting differential growth of microbial populations using Gaussian process regression”, Genome Research 2016
- ▶ Darnell, Tonner *et al.* “Systematic Discovery of Archaeal Transcription Factor Functions in Regulatory Networks through Quantitative Phenotyping Analysis” MSystems, 2017
- ▶ Tonner *et al.* “Random effects modeling of microbial population growth”, *in prep*

Why Gaussian processes?

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- ▶ **Population genomics:** “Flexible Modelling of Genetic Effects on Function-Valued Traits”, Fusi and Listgarten (2016)
- ▶ **Temporal modeling of microbial communities:** “Temporal probabilistic modeling of bacterial compositions derived from 16S rRNA sequencing”, Aijo, Mueller, and Bonneau (2016)
- ▶ **Single cell RNA-seq time-series:** “Order Under Uncertainty: Robust Differential Expression Analysis Using Probabilistic Models for Pseudotime Inference”, Campbell and Yau (2016)

What is your background?

- ▶ statistics: multivariate normal, linear models, Bayesian vs. Frequentist statistics
- ▶ programming: python, jupyter notebooks
- ▶ linear algebra: matrix inversion, positive definiteness

Affiliations

- ▶ CBB
- ▶ Biology
- ▶ CMB
- ▶ MGM
- ▶ UPGG
- ▶ BME

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About this workshop

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Outline of sessions

1. Basic Theory and Applications
 - ▶ GP theory
 - ▶ GP regression & GPy
2. Bayesian Modeling
 - ▶ Simple example
 - ▶ Hierarchical models

Resources

- ▶ github: github.com/ptonner/gp-workshop

Major References

- ▶ *Gaussian Processes for Machine Learning*, Rasmussen and Williams (2006)
- ▶ GPy, a Gaussian process library in python.
sheffieldml.github.io/GPy/
- ▶ Stan: mc-stan.org

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Motivating example

- ▶ **Problem:** choosing $f(x)$ based on observed data, may need to change with new observations. no “best” way to choose f
- ▶ **Question:** what if we could represent $f(x)$ with a distribution?

Answer: Gaussian processes

def: A *Gaussian process* is an infinite collection of random variables, any finite number of which have a joint multivariate normal distribution.

$$f(x) \sim GP(\mu(x), \kappa(x, x'))$$

Most common uses for Gaussian processes

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$$f(x) \sim GP(\mu(x), \kappa(x, x'))$$

regression

$$y(x) = f(x) + \epsilon(x)$$

classification

$$p(y(x) = 1 | f(x)) = \frac{1}{1 + \exp(-f(x))}$$

hierarchical

$$p(y(x) | f(x), \dots)$$

Gaussian processes regression

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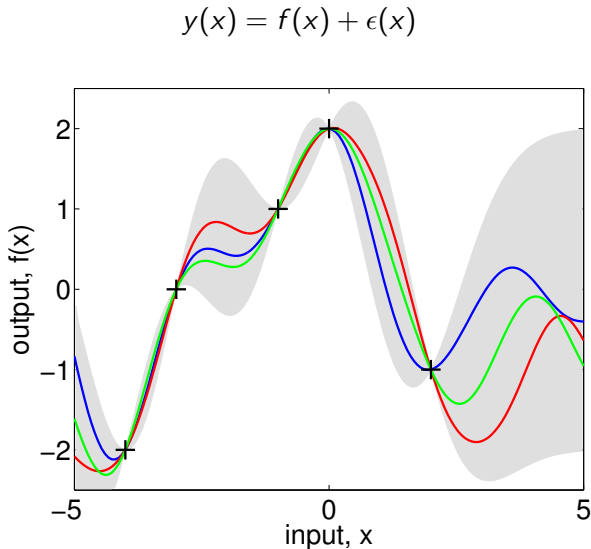
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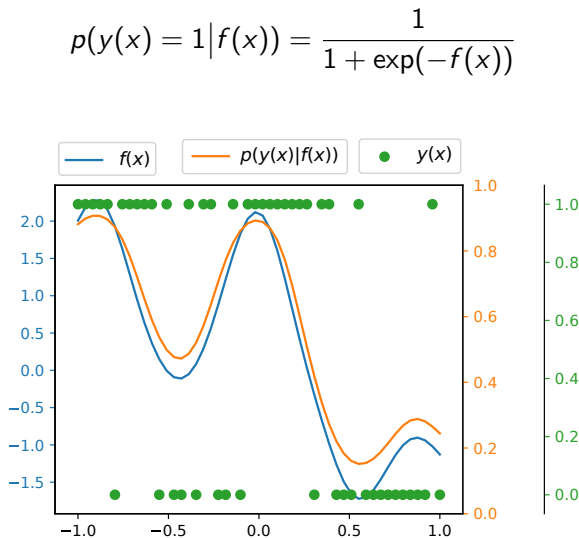
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$$X \sim N(\mu, \Sigma), \quad \mu = \begin{bmatrix} \mu_1 \\ \mu_2 \\ \vdots \\ \mu_p \end{bmatrix}, \quad \Sigma = \begin{pmatrix} \sigma_{1,1}^2 & \cdots \\ \vdots & \ddots \end{pmatrix}$$

$$p(X = x | \mu, \Sigma) = \frac{1}{(2\pi)^{p/2} |\Sigma|^{1/2}} \cdot \exp\left(\frac{1}{2}(x - \mu)^T \Sigma^{-1}(x - \mu)\right)$$

$$X \sim N(\mu, \Sigma), \quad \mu = \begin{bmatrix} \mu_1 \\ \mu_2 \\ \vdots \\ \mu_p \end{bmatrix}, \quad \Sigma = \begin{pmatrix} \sigma_{1,1}^2 & \cdots \\ \vdots & \ddots \end{pmatrix}$$

All subsets of X are multivariate normal:

$$Y = \begin{bmatrix} X_i \\ X_j \\ X_k \end{bmatrix} = N\left(\begin{bmatrix} \mu_i \\ \mu_j \\ \mu_k \end{bmatrix}, \begin{pmatrix} \sigma_{i,i}^2 & \cdots \\ \vdots & \ddots \end{pmatrix}\right)$$

$$X \sim N(\mu, \Sigma), \quad \mu = \begin{bmatrix} \mu_1 \\ \mu_2 \\ \vdots \\ \mu_p \end{bmatrix}, \quad \Sigma = \begin{pmatrix} \sigma_{1,1}^2 & \cdots \\ \vdots & \ddots \end{pmatrix}$$

Any linear combination of X , $AX + b$ is

$$AX + b \sim N(A\mu + b, A\Sigma A^T)$$

$$X \sim N(\mu, \Sigma), \quad \mu = \begin{bmatrix} \mu_1 \\ \mu_2 \\ \vdots \\ \mu_p \end{bmatrix}, \quad \Sigma = \begin{pmatrix} \sigma_{1,1}^2 & \cdots \\ \vdots & \ddots \end{pmatrix}$$

Conditional distribution:

$$X = \begin{bmatrix} X_1 \\ X_2 \end{bmatrix} \sim N\left(\begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix}, \begin{bmatrix} \Sigma_{1,1} & \Sigma_{1,2} \\ \Sigma_{2,1} & \Sigma_{2,2} \end{bmatrix} \right)$$

$$X_1 \mid X_2 = x_2 \sim N(\mu_1 + \Sigma_{1,2}\Sigma_{2,2}^{-1}(x_2 - \mu_2), \\ \Sigma_{1,1} - \Sigma_{1,2}\Sigma_{2,2}^{-1}\Sigma_{2,1})$$

definition

$$\text{cov}(X, Y) = E[(X - \mu_X)(Y - \mu_Y)] \quad (1)$$

$$= \frac{\text{cor}(X, Y)}{\sigma_X \sigma_Y} \quad (2)$$

must be positive (semi)definite

$$\mathbf{x}^T \Sigma \mathbf{x} \geq 0 \quad \forall \mathbf{x}$$

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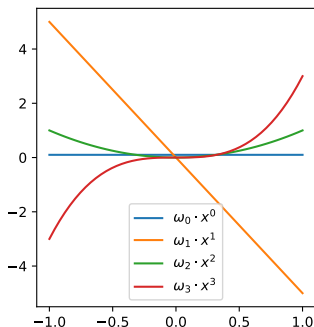
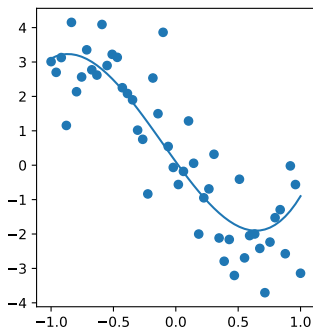
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Bayesian linear regression

$$\omega_0 = 0.1, \omega_1 = -5.0, \omega_2 = 1.0, \omega_3 = 3.0$$



$$y(x) = f(x) + \epsilon(x)$$

$$\phi_p(x) = \{1, x, x^2, \dots, x^p\}$$

$$f(x) = \phi_p(x)^T \omega$$

$$\omega \sim N(0, \Sigma_\omega)$$

$$\epsilon \sim N(0, \sigma_y^2)$$

Distribution of $f(x)$

$$f(x) = \phi_p(x)^T \omega$$

Mean

$$E[f(x)] = \phi(x)^T E[\omega] = 0$$

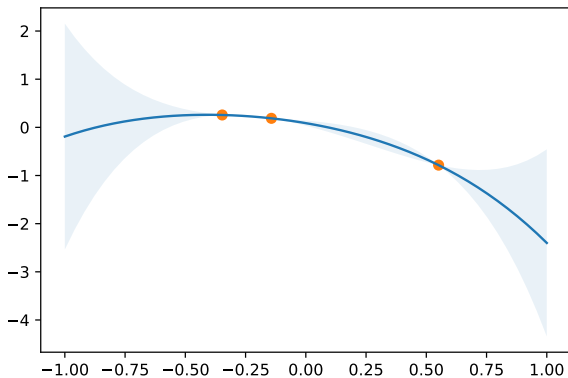
Covariance

$$\begin{aligned} \text{Cov}[f(x), f(x')] &= E[f(x)f(x')] = \phi(x)^T E[\omega\omega^T] \phi(x') \\ &= \phi(x)^T \Sigma_\omega \phi(x') \\ &= \kappa_f(x, x') \end{aligned}$$

Conditional distribution

$$\begin{bmatrix} f(X_1) \\ f(X_2) \end{bmatrix} \sim N\left(\begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} \kappa_1 & \kappa_{1,2} \\ \kappa_{2,1} & \kappa_2 \end{bmatrix} \right)$$

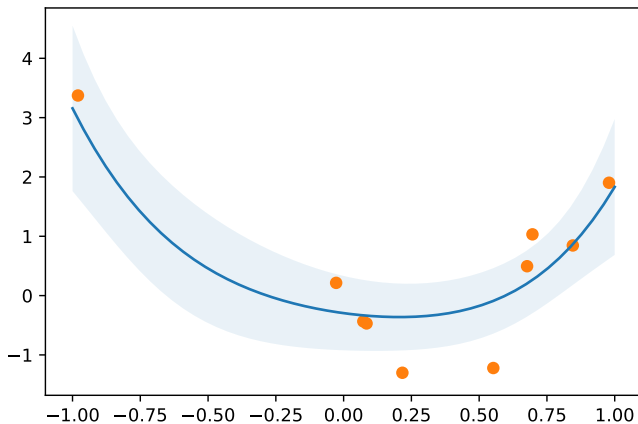
$$f(X_2) \mid f(X_1) \sim N(\kappa_{2,1}\kappa_1^{-1}f(X_1), \kappa_2 - \kappa_{2,1}\kappa_1^{-1}\kappa_{1,2})$$



adding observation noise is easy

$$\begin{bmatrix} y(X_1) \\ f(X_2) \end{bmatrix} \sim N \left(\begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} \kappa_1 + \sigma_y^2 I & \kappa_{1,2} \\ \kappa_{2,1} & \kappa_2 \end{bmatrix} \right)$$

$$f(X_2) \mid y(X_1) \sim N(\kappa_{2,1}(\kappa_1 + \sigma_y^2 I)^{-1} y(X_1), \\ \kappa_2 - \kappa_{2,1}(\kappa_1 + \sigma_y^2 I)^{-1} \kappa_{1,2})$$



What did we learn?

- ▶ A procedure for building a covariance matrix (and mean) can describe a distribution on functions
- ▶ We can compute the conditional distribution of $f(x)$ at any x , given observations
- ▶ The (co)variance of these functions are dependent on the *input* (e.g. x), and not the output (y)

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Distribution of f

- ▶ f has a GP prior: $f(X) \sim GP(\mu(X), \kappa(X))$
- ▶ which is MVN for finite X :

$$p(f(X)) = MVN(\mu(X), \kappa(X))$$

Noisy observation $y(x)$

- ▶ conditional on f : $y(X)|f(X) \sim N(f(x), \sigma_y^2 I)$
- ▶ marginalized:

$$\begin{aligned} p(y) &= \int p(y|f)p(f)\partial f \\ &= N(\mu(X), \kappa(X) + \sigma_y^2 I) \end{aligned}$$

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Idea: what if we changed $k_f(x, x')$ to a different function?

Radial basis function

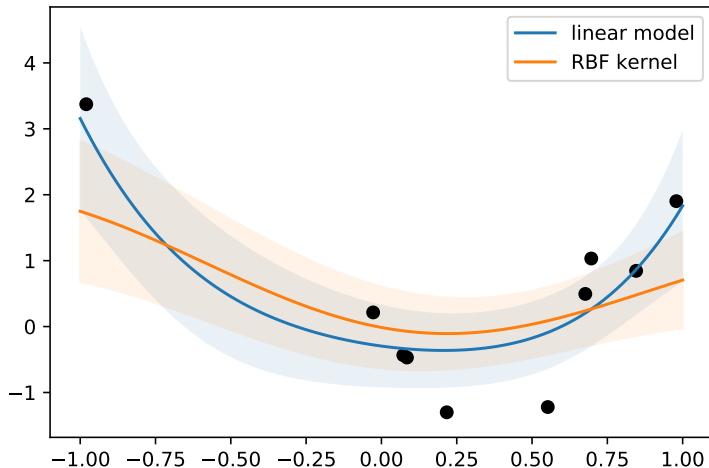
$$\text{cov}(f(x_1), f(x_2)) = \kappa_{\text{RBF}}(x, x') = \sigma^2 \cdot \exp\left(\frac{-|x_1 - x_2|^2}{\ell}\right)$$

Define $f(x)$ as:

$$f(x) \sim N\left(\mathbf{0}, \kappa(x) = \{\kappa_{\text{RBF}}(x_i, x_j)\}\right)$$

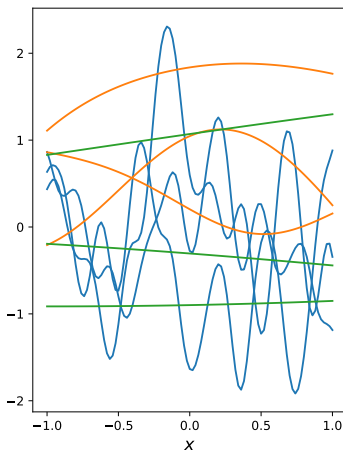
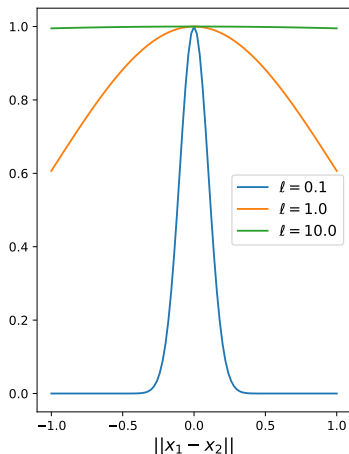
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Replace κ_f with κ_{RBF}



Lengthscale changes the smoothness

$$\text{cov}(f(x_1), f(x_2)) = \kappa_{\text{RBF}}(x, x') = \sigma^2 \cdot \exp\left(\frac{-|x_1 - x_2|^2}{\ell}\right)$$



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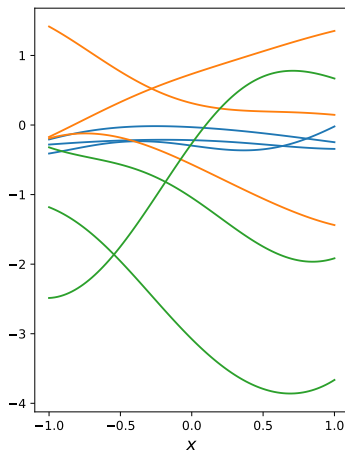
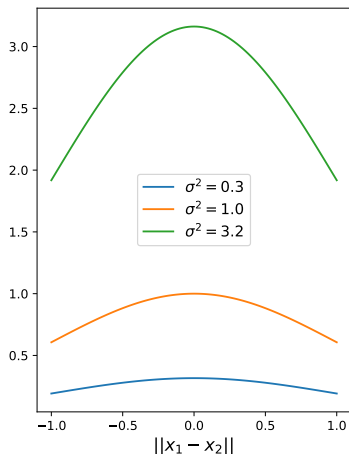
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Variance changes the magnitude

$$\text{cov}(f(x_1), f(x_2)) = \kappa_{\text{RBF}}(x, x') = \sigma^2 \cdot \exp\left(\frac{-|x_1 - x_2|^2}{\ell}\right)$$

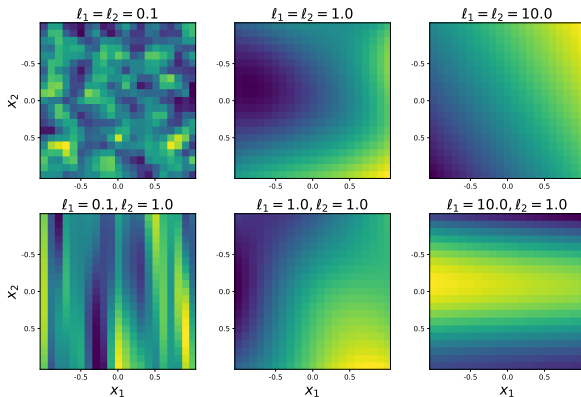


Support for multi-dim x

$$x = \{x_1, x_2, \dots, x_p\},$$

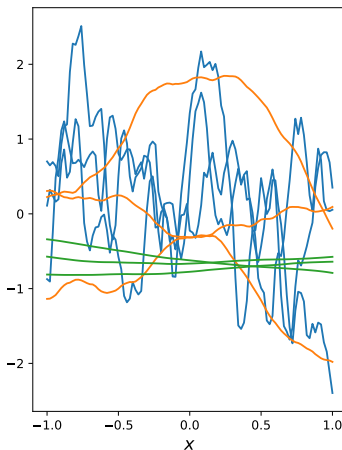
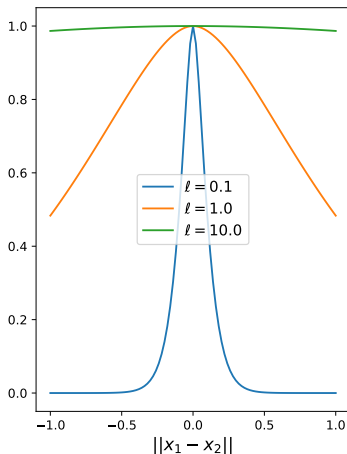
$$\kappa(x_1, x_2) = \sigma^2 \cdot \exp\left(\sum_{i=1}^p \frac{-|x_{1,i} - x_{2,i}|^2}{\ell_p}\right)$$

Referred to as auto-relevance detection (ARD)



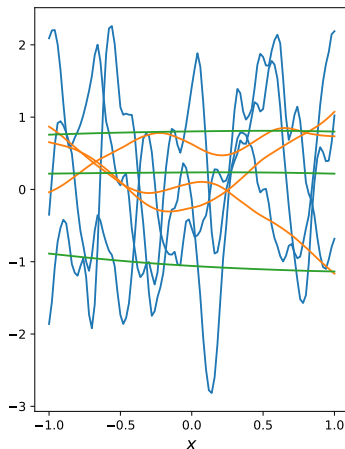
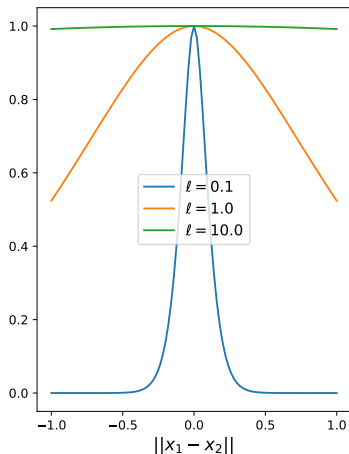
Matern kernel

$$\kappa_{\frac{3}{2}} = \left(1 + \frac{\sqrt{3}r}{\ell}\right) \cdot \exp\left(\frac{-\sqrt{3}r}{\ell}\right)$$

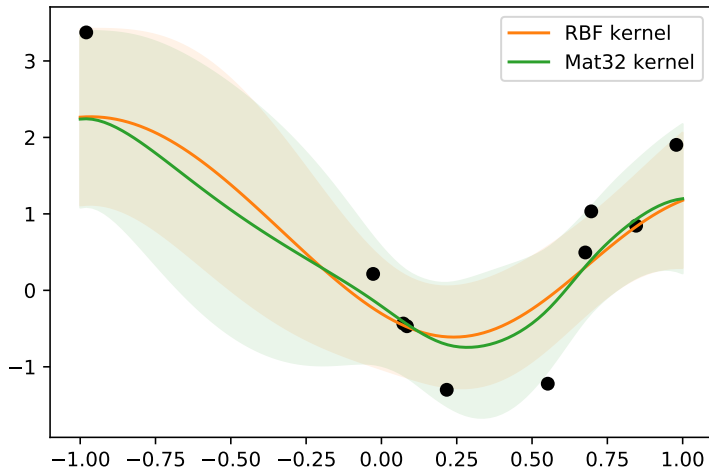


Matern kernel

$$\kappa_{\frac{5}{2}} = \left(1 + \frac{\sqrt{5}r}{\ell} + \frac{5r^2}{3\ell^2}\right) \cdot \exp\left(\frac{-\sqrt{3}r}{\ell}\right)$$



κ_{RBF} VS κ_{Mat32}



if κ_A and κ_B are both valid kernels, then so are:

- ▶ $\kappa_A + \kappa_B$, interpret as

$$y(x) = f_a(x) + f_b(x)$$

- ▶ $\kappa_A * \kappa_B$, interpret as

$$y(x) = f_a(x) \times f_b(x)$$

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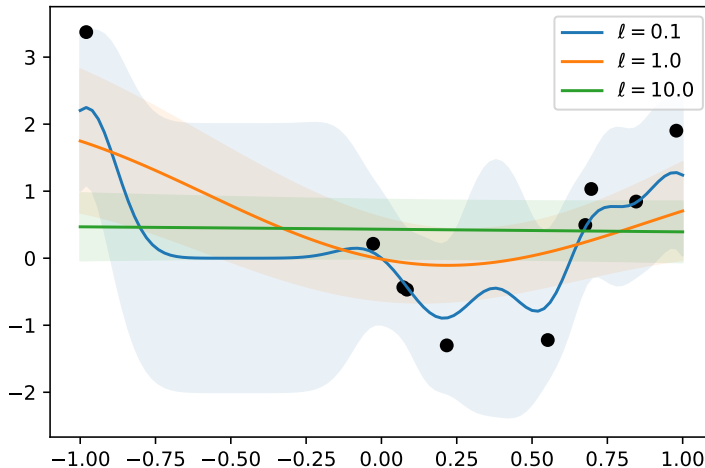
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Hyperparameters change model interpretation



How do we determine hyperparameters?

1. Maximize likelihood

- ▶ gradient optimization, point estimate
- ▶ **pros:** many existing implementations (e.g. *GPY*), often gets answer the fastest
- ▶ **cons:** requires “conjugacy” b/w f and y for analytical sol'n

2. Model hyperparameter uncertainty

- ▶ Posterior sampling (MCMC)
- ▶ **pros:** Doesn't care about form of $p(y)$, can be placed in larger model easily
- ▶ **cons:** sampling time (can be long)

1. Maximal likelihood

$$\mathcal{L}(\theta) = \log p(y|\theta) = \frac{1}{2} \log \det \Sigma(\theta) + \frac{1}{2} y^t \Sigma(\theta)^{-1} y + \frac{N}{2} \log(2\pi)$$

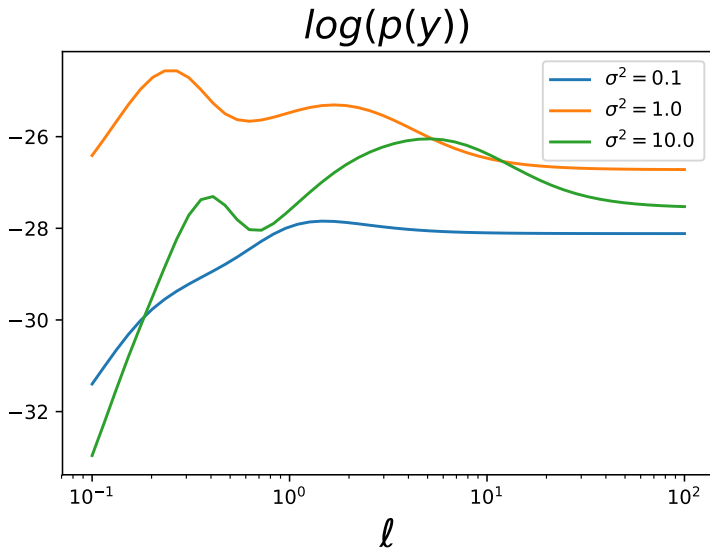
where: $\Sigma(\theta) = \kappa(\theta) + \sigma^2 I$, $\theta = \{\ell, \sigma_f^2, \sigma_y^2\}$

- ▶ $\mathcal{L}(\theta)$ is non-convex
- ▶ Use gradient based optimization:

$$\frac{\partial \mathcal{L}}{\partial \theta_i} = \frac{1}{2} \text{tr} \Sigma(\theta) \frac{\partial \Sigma(\theta)}{\partial \theta_i} + \frac{1}{2} y^t \Sigma(\theta)^{-1} \frac{\partial \Sigma(\theta)}{\partial \theta_i} \Sigma(\theta)^{-1} y$$

- ▶ $\frac{\partial \Sigma(\theta)}{\partial \theta_i}$: need to compute derivative of kernel function wrt each hyperparameter
- ▶ $\Sigma(\theta)^{-1}$: $O(n^3)$ operation

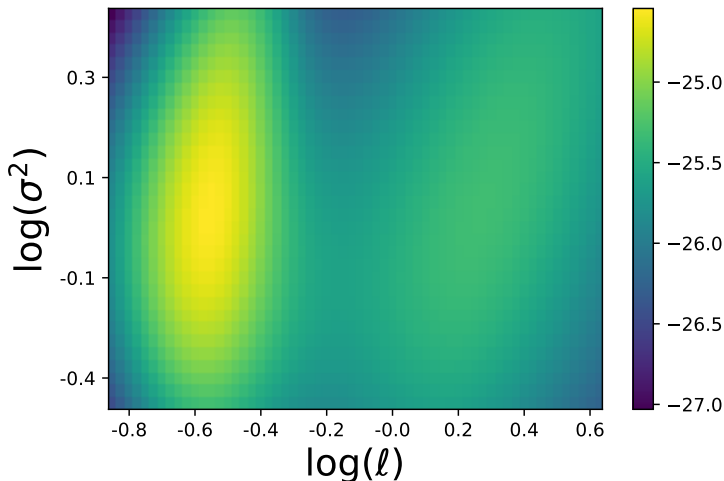
Data likelihood as a function of hyperparameters



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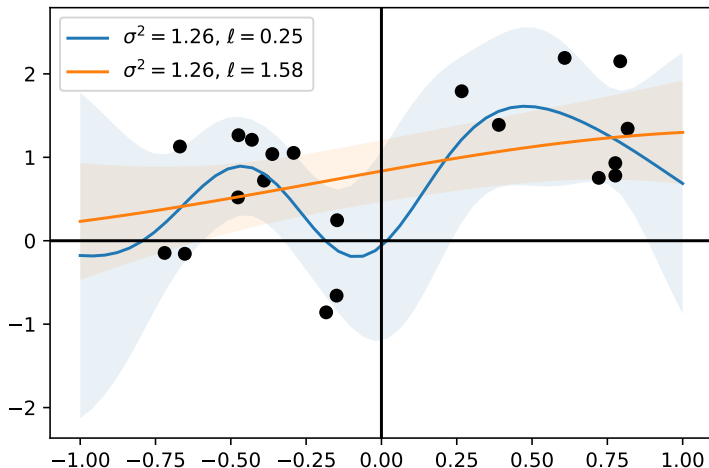
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What next?

Further reading

- ▶ Rasmussen and Williams *Gaussian Processes for Machine Learning*, 2006: Chs. 2,4,5

Remaining sessions

- ▶ Applications (*GPy* and *Stan*)
- ▶ Case studies