

EVT and Angular Distributions

Peter Trubey
UCSC Department of Statistics

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Overview

Introduction

Methodology & Results

- Polar Coordinate Transformation

- Projected Gamma Distribution

- Finite Mixture of Projected Gammas

- DP Mixture of Projected Gammas, Independent Prior

- DP Mixture of Multivariate Normals

- DP Mixture of Projected Gammas, LogNormal Prior

Introduction

- ▶ Our goal is a framework for anomaly detection using extreme value theory
- ▶ Currently at least, We're still focusing on the EVT part

EVT - A brief introduction

- ▶ Why EVT?
- ▶ Block Maxima - GEV
- ▶ Thresholding - GP

Block Maxima - GEV

- ▶ GEV (Generalized Extreme Value Distribution) is a unification of three distributions: Frechet, Gumbel, and Weibull.
- ▶ For i.i.d. x_1, \dots, x_n $x_i \sim F$, $M_n = \max_i x_i$, if there exists some sequence $a_n > 0$, b_n such that:

$$\Pr \left[\frac{M_n - b_n}{a_n} \leq z \right] \rightarrow G(z) \quad \text{as } n \rightarrow \infty$$

Then $G(z)$ takes the form:

$$G(z) = \exp \left\{ - \left[1 + \xi \left(\frac{z - \mu}{\sigma} \right) \right]^{-\xi^{-1}} \right\}$$

Then F falls into the *domain of attraction* of said GEV.

Thresholding - GP

- We look now at the distribution of excesses y over a threshold u .

$$\Pr[X > u + y \mid X > u] = \frac{1 - F(u + y)}{1 - F(u)}, \quad y > 0$$

If F is in the domain of attraction of a GEV, then $\Pr[X > u + y \mid X > u]$ converges to

$$H(y) = 1 - \left(1 + \frac{\xi y}{\sigma}\right)^{-\xi^{-1}},$$

the Generalized Pareto distribution.

- We set a high threshold by some criteria, then model excesses over that threshold using GP.

Multivariate EVT

- ▶ Useful to standardize each X_i according to its marginal distribution
- ▶ Standardization occurs as:

$$z_j = \left(1 + \xi \frac{x_j - b_{t,j}}{a_{t,j}} \right)_+^{1/\xi} \quad (1)$$

where $b_{t,j} = \hat{F}^{-1}(1 - 1/t)$.

- ▶ Note that $Z_j > 1 \implies X_j > b_{t,j}$
- ▶ $\max_j Z_j \sim \text{Pareto}$
- ▶ a, ξ are currently found by MLE. Not ideal, I know.

Multivariate EVT

- ▶ Assume the existence of a limit measure μ on Z such that:

$$n\Pr\left(\frac{V_1}{n} \geq v_1 \text{ or } \dots \text{ or } \frac{V_d}{n} \geq v_d\right) \rightarrow \mu\left([0, v]^C\right)$$

- ▶ μ is the asymptotic distribution of Z in extreme regions.
- ▶ μ features the homogeneity property, $\mu(t\cdot) = t^{-1}\mu(\cdot)$.

Pseudo-Polar Representation

Given the homogeneity property, we can decompose Z into two components, radial and angular:

$$R(Z) = \|Z\|_{\infty} = \max_i v_i$$

$$V = \frac{Z}{R} \in S_{\infty}^{d-1}$$

where S_{∞}^{d-1} is the positive orthant of the unit hypercube in R^d .

Spectral Measure

For $B \subset S_{\infty}^{d-1}$, define the *Spectral Measure*:

$$\Omega(B) = \mu[z : R(z) > 1, V \in B].$$

Then,

$$\mu[z : R(z) > t, V \in B] = t^{-1} \Omega(B).$$

Thus t is independent of Ω , the spectral measure. This is completed as a probability measure as:

$$\Pr(V \in B \mid r > 1) = \frac{\Omega(B)}{\Omega(S_{\infty}^{d-1})}.$$

So conditional on at least one dimension exceeding its threshold, we can hold the angular measure independent of the magnitude.

Polar Coordinate Transformation

► One way

$$y_1 = r \cos \theta_1,$$

$$y_2 = r \sin \theta_1 \cos \theta_2$$

$$\vdots$$

$$y_{d-1} = r \sin \theta_1 \dots \sin \theta_{d-2}$$

$$y_d = r \sin \theta_1 \dots \sin \theta_{d-1}$$

► The other way

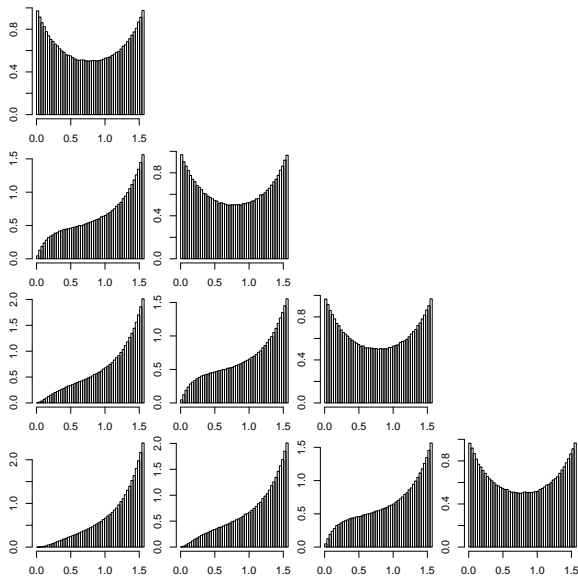
$$\theta_1 = \cos^{-1} \left[\frac{y_1}{\|y_{1:d}\|_2} \right]$$

$$\theta_2 = \cos^{-1} \left[\frac{y_2}{\|y_{2:d}\|_2} \right]$$

$$\vdots$$

$$\theta_{d-1} = \cos^{-1} \left[\frac{y_{d-1}}{\|y_{(d-1):d}\|_2} \right].$$

Polar Coordinate Transformation Effects



Projected Gamma Distribution

Consider d independent Gammas

$$f(\mathbf{y} \mid \alpha, \beta) = \prod_{j=1}^d \text{Ga}(y_j \mid \alpha_j, \beta_j), \quad (2)$$

Use polar coordinate transformation, then integrate out r .

$$\text{PG}(\theta \mid \alpha, \beta) = \frac{\Gamma(A) \beta_d^{\alpha_d}}{B^A \Gamma(a_d)} \left(\prod_{j=1}^{d-1} \frac{\beta_j^{\alpha_j}}{\Gamma(\alpha_j)} (\cos \theta_j)^{\alpha_j-1} (\sin \theta_j)^{(\sum_{h=j+1}^d \alpha_h)-1} \right) \mathcal{I}_{(0, \pi/2)^{d-1}}(\theta) \quad (3)$$

where

$$A = \sum_{j=1}^d \alpha_j \quad \text{and} \quad B = \beta_1 \cos \theta_1 + \sum_{j=2}^{d-1} \left(\beta_j \cos \theta_j \prod_{i=1}^{j-1} \sin \theta_i \right) + \beta_d \prod_{j=1}^{d-1} \sin \theta_j. \quad (4)$$

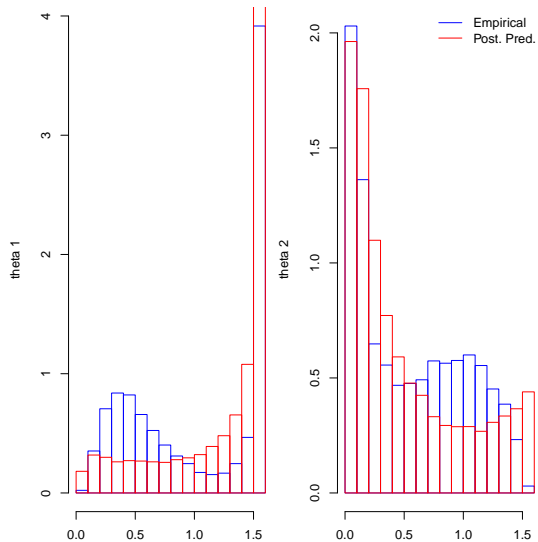
Projected Gamma, Cont.

- ▶ Let $y' = \frac{y}{r}$, $y = ry'$.
- ▶ Gibbs sampling updates to α , β possible with latent r .

$$r_i \mid \alpha, \beta \sim \text{Ga} \left(r_i \mid \sum_j \alpha_j, \sum_j \beta_j y'_{ij} \right)$$

- ▶ For Projected Gamma, information between dimensions is effectively communicated through the latent r .

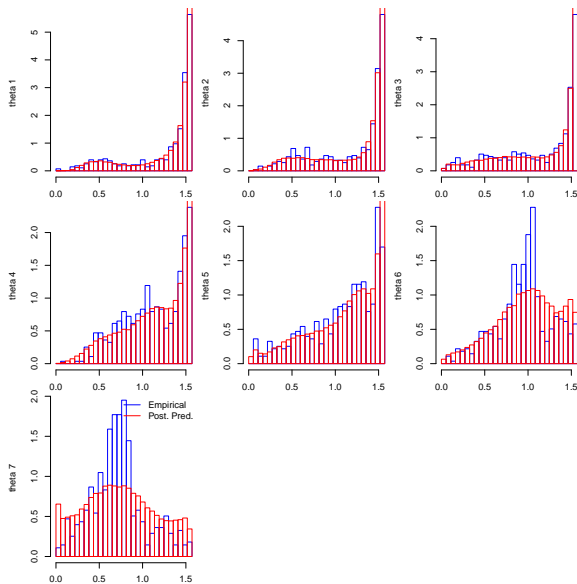
Projected Gamma, Cont.



Finite Mixture of Projected Gammas

$$\begin{aligned}\text{MPG}(\theta \mid \lambda, \alpha, \beta) &= \sum_{j=1}^J \lambda_j \text{PG}(\alpha_j, \beta_j) \\ &= \int \text{MPG}(\theta, \gamma \mid \alpha, \beta) d\gamma \\ &= \int \prod_{j=1}^J [\lambda_j \text{PG}(\theta \mid \alpha_j, \beta_j)]^{\gamma_j} d\gamma \\ &= \int_{\gamma} \int_r \lambda_j^{\gamma_j} \prod_{k=1}^K \text{Ga}(ry' \mid \alpha_j, \beta_j)^{\gamma_j} |\text{Jac}| dr d\gamma\end{aligned}$$

Finite Mixture of Projected Gammas - cont.



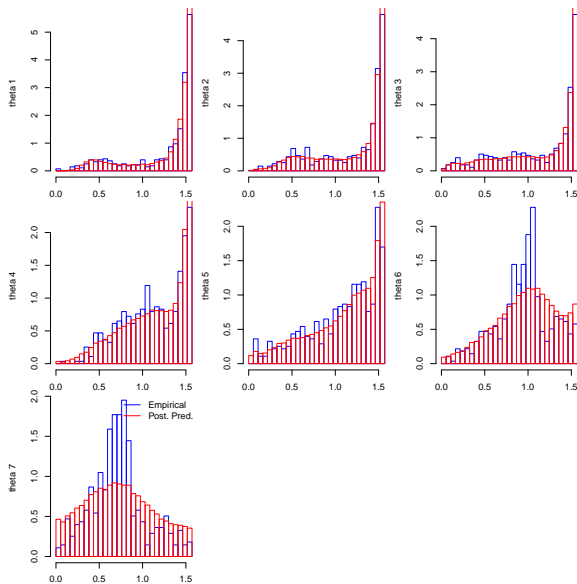
DP Mixture of Projected Gammas, Independent Gamma Prior

$$\begin{aligned}\theta_i &\sim \text{PG}(\theta_i \mid (\alpha_i, \beta_i)) \\ (\alpha_i, \beta_i) &\sim G_i \\ G_i &\sim \text{DP}(\eta, G_0((\alpha_i, \beta_i) \mid (\mathbf{a}_\alpha, \mathbf{b}_\alpha, \mathbf{a}_\beta, \mathbf{b}_\beta))) \\ (\mathbf{a}_\alpha, \mathbf{b}_\alpha, \mathbf{a}_\beta, \mathbf{b}_\beta) &\sim P((\mathbf{a}_\alpha, \mathbf{b}_\alpha, \mathbf{a}_\beta, \mathbf{b}_\beta)) \\ \eta &\sim \text{Ga}(\mathbf{a}_\eta, \mathbf{b}_\eta)\end{aligned}$$

with

$$\begin{aligned}G_0((\alpha_i, \beta_i)) &= \text{Ga}(\alpha_1 \mid \mathbf{a}_{\alpha_1}, \mathbf{b}_{\alpha_1}) \prod_{j=2}^d \text{Ga}(\alpha_j \mid \mathbf{a}_{\alpha_j}, \mathbf{b}_{\alpha_j}) \text{Ga}(\beta_j \mid \mathbf{a}_{\beta_j}, \mathbf{b}_{\beta_j}) \\ P((\mathbf{a}_\alpha, \mathbf{b}_\alpha, \mathbf{a}_\beta, \mathbf{b}_\beta)) &= \text{Ga}(\mathbf{a}_{\alpha_1}) \text{Ga}(\mathbf{b}_{\alpha_1}) \prod_{j=2}^d \text{Ga}(\mathbf{a}_{\alpha_j}) \text{Ga}(\mathbf{b}_{\alpha_j}) \text{Ga}(\mathbf{a}_{\beta_j}) \text{Ga}(\mathbf{b}_{\beta_j})\end{aligned}$$

DP Mixture of Projected Gammas, Independent Gamma Prior - Cont.



DP Mixture of Multivariate Normals on Probit Space

- ▶ $\theta \in [0, \pi/2]$
 - ▶ Jitter and rescale to $\theta' \in [\epsilon, 1 - \epsilon]$
 - ▶ Then $W = \Phi^{-1}(\theta')$
- ▶ Then develop a hierarchical multivariate normal, with length of response $d - 1$.

$$W_i \sim \mathcal{N}_d(\mu_i, \Sigma_i)$$

$$\mu_i, \sigma_i \sim G_i$$

$$G_i \sim \text{DP}(\eta, G_0(\mu_i, \Sigma_i \mid \mu_0, \Sigma_0))$$

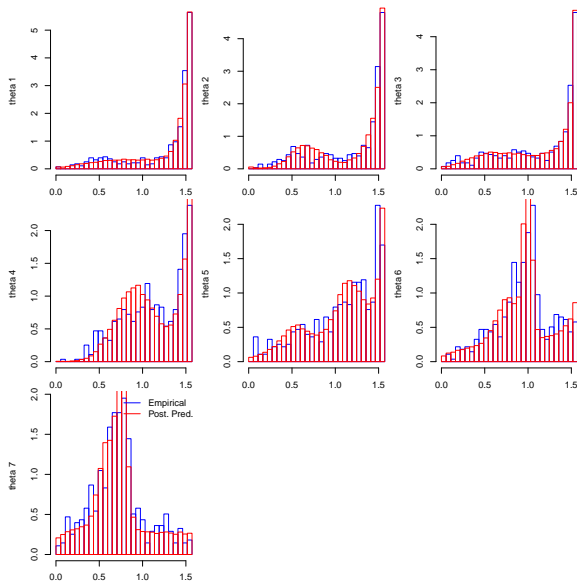
$$G_0(\mu_i, \Sigma_i \mid \mu_0, \Sigma_0) = \mathcal{N}_d(\mu_i \mid \mu_0, \Sigma_0) \text{IW}(\Sigma \mid \nu, \psi)$$

$$\mu_0 \sim \mathcal{N}_d(\mathbf{u}, \mathbf{S})$$

$$\Sigma_0 \sim \text{IW}(\nu_0, \psi_0)$$

$$\eta \sim \text{Ga}(\alpha, \beta)$$

DP Mixture of Multivariate Normals on Probit Space - cont.



DP Mixture of Projected Gammas, LogNormal Prior

- ▶ The Normal Normal model operated in $d - 1$ -dimensional space without considering the transformation that brought it there.
- ▶ If we want to operate in d -dimensional space with $d - 1$ degrees of freedom, then Projected Gamma is a good base distribution.
- ▶ Placing a multivariate log-normal prior on α , we can *potentially* communicate information across dimensions more effectively than with r alone.

$$\theta_i \sim \text{PG}(\theta_i \mid \alpha_i, \beta_i)$$

$$r_i \sim \text{Ga}(r_i \mid \alpha_i, \beta_i)$$

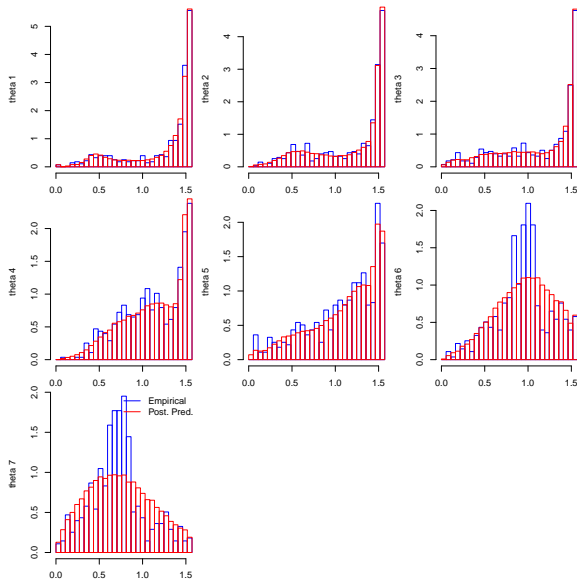
$$(\alpha_i, \beta_i) \sim \text{DP}((\alpha_i, \beta_i) \mid \eta, G_0)$$

$$G_0 = \text{Log}\mathcal{N}(\alpha_i \mid \mu, \Sigma) \prod_{j=2}^d \text{Ga}(\beta_{ij} \mid a, b)$$

$$\mu \sim \mathcal{N}(\mu \mid \mu_0, \Sigma_0)$$

$$\Sigma \sim \text{IG}(\Sigma \mid \nu, \psi)$$

DP Mixture of Projected Gammas, LogNormal Prior - cont.



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