1 公式

$$\begin{split} \sum_{k=0}^{n-1} k^3 &= \frac{1}{4} n^2 (n+1)^2 \\ \sum_{k=0}^{n-1} r^k &= \frac{1-r^n}{1-r}, r \neq 1 \\ \sum_{k=1}^{n} k r^k &= r \frac{1-(n+1)r^n + nr^{n+1}}{(1-r)^2}, r \neq 1 \\ \sin x &= \sum_{n=0}^{\infty} \frac{(-1)^n}{(2n+1)!} x^{2n+1} = x - \frac{x^3}{3!} + \frac{x^5}{5!} - \cdots \\ \cos x &= \sum_{n=0}^{\infty} \frac{(-1)^{n-1}}{(2n)!} x^{2n} = 1 - \frac{x^2}{2!} + \frac{x^4}{4!} - \cdots \\ \ln x &= \sum_{n=0}^{\infty} \frac{(-1)^{n-1}}{n} (x-1)^n = (x-1) - \frac{(x-1)^2}{2} - \cdots \\ \frac{1}{1-x} &= \sum_{n=0}^{\infty} x^n = 1 + x + x^2 + x^3 + \cdots \\ \frac{d}{dx} \tan x &= \sec^2 x \\ \frac{d}{dx} \cot x &= -\csc^2 x \\ \frac{d}{dx} \cot x &= -\csc^2 x \\ \frac{d}{dx} \operatorname{arccos} x &= -\frac{1}{\sqrt{1-x^2}} \\ \frac{d}{dx} \operatorname{arccos} x &= -\frac{1}{\sqrt{1-x^2}} \\ \frac{d}{dx} \operatorname{arccos} x &= -\frac{1}{\sqrt{1-x^2}} \\ \frac{d}{dx} \operatorname{arccan} x &= \frac{1}{1+x^2} \\ \int \frac{1}{x^2 + \alpha^2} \mathrm{d} x &= \frac{\ln(\frac{x \mp \alpha}{2x + \alpha})}{2\alpha} \\ \int \frac{1}{4x^2 + \beta} \mathrm{d} x &= \frac{1}{\sqrt{ab}} \arctan \frac{\sqrt{ax}}{\sqrt{b}} \\ \int \frac{1}{\sqrt{a^2 - x^2}} \mathrm{d} x &= \arcsin \frac{x}{a} &= -\arccos \frac{x}{a} \\ \int \sec^2 x \mathrm{d} x &= \tan x \\ \int \csc^2 x \mathrm{d} x &= - \ln|\cos x| &= \ln|\sec x| \\ \int \cot x \mathrm{d} x &= \ln|\sin x| \\ \int \sec x \mathrm{d} x &= \ln|\sec x + \tan x| \\ \int \csc x \mathrm{d} x &= \ln|\sec x + \tan x| \\ \int \csc x \mathrm{d} x &= \ln|\csc x - \cot x| &= \ln|\frac{\tan x - \sin x}{\sin x \tan x}| \\ \int x^n e^{ax} \mathrm{d} x &= \frac{1}{a} x^n e^{ax} - \frac{n}{a} \int x^{n-1} e^{ax} \mathrm{d} x \\ \int_{-\infty}^{\infty} e^{-ax^2 + bx + c} \mathrm{d} x &= \sqrt{\frac{\pi}{a}} e^{\frac{b^2}{4a} + c} \end{split}$$

2 Probability

2.1 Distribution

2.1.1 Poisson Distribution

本质上是一个 $n \to \infty$ 的二项分布, $\lambda = np$ 。 性质: $\mathbb{E}(X) = \lambda$, $\mathrm{Var}(X) = \lambda$ Approximate Bin: n large, p small $(n \ge 50, np \le 5)$

2.1.2 Hypergeometric Distribution

记号: $X \sim \text{Hypergeomet}(n,N,m)$ 概率: $p(k) = \frac{\binom{m}{k}\binom{N-m}{n-k}}{\binom{N}{n}}$ N 个球,m 个红球,不放回取出 n 个,有 k 红球。 $\mathbb{E}(X) = n \cdot \frac{m}{N}, \text{Var}(X) = n \cdot \frac{m}{N} \left(1 - \frac{m}{N}\right) \left(1 - \frac{n-1}{N-1}\right)$

2.1.3 Normal Distribution

Approximate Bin: $np(1-p) \ge 10$ $Z \sim N(0,1), \mathbb{E}(g'(Z)) = \mathbb{E}(Zg(Z)),$ assuming that $\lim_{x\to\infty} \frac{g(x)}{e^{\frac{x^2}{2}}} = 0.$ So $\mathbb{E}(Z^{n+1}) = n\mathbb{E}(Z^{n-1}).$

2.1.4 Exponential Distribution

CDF:
$$F(X) = 1 - e^{-\lambda x}, x \ge 0$$

 $\mathbb{P}r(X > x) = e^{-\lambda x}, x \ge 0$
 $\mathbb{E}(X^n) = \frac{n}{\lambda} \mathbb{E}(X^{n-1}) = \frac{n!}{\lambda^n}$
 $\mathbb{E}(X) = \frac{1}{\lambda}, \text{Var}(X) = \frac{1}{\lambda^2}$

2.1.5 Gamma Distribution

考试都用 $\Gamma(\alpha,\beta)$ 的形式 $\Gamma(x) = \int_0^\infty u^{x-1} e^{-u} du, x > 0$ $\Gamma(x) = (x-1)\Gamma(x-1)$ $\Gamma(n) = (n-1)!$ PDF: $f(x) = \frac{\lambda^{\alpha}}{\Gamma(a)} x^{\alpha-1} e^{-\lambda x}, x > 0$ α : 发生次数 $\mathbb{E}(X) = \frac{\alpha}{\lambda}, \operatorname{Var}(X) = \frac{\alpha^2}{\lambda^2}$ $\mathbb{E}(X^n) = \frac{n+\alpha-1}{\lambda} \mathbb{E}(X^{n-1}) = \frac{\alpha^n}{\lambda^n}$

2.1.6 Chi-Squared Distribution

$$X \sim \chi^{2}(k)$$

$$\mathbb{E}(X) = k$$

$$\mathcal{N}(0, 1)^{2} \sim \chi_{1}^{2}$$

$$\chi_{n}^{2} \sim \Gamma\left(\frac{n}{2}, 2\right)$$

2.1.7 t-Distribution

$$T_k = \frac{C}{\sqrt{D/k}}, C \sim \mathcal{N}(0, 1), D = \chi_k^2$$

$$\frac{\sqrt{n}(\overline{X} - \mu)}{S} \sim t_{n-1} \text{ for } X_i \sim \mathcal{N}(\mu, \sigma^2)$$

$$\sigma^2 = \frac{v}{v-2} \text{ for } v > 2, \infty \text{ for } 1 < v \le 2$$

$$f(t) = \frac{\Gamma[(n+1)/2]}{\sqrt{n\pi}\Gamma(n/2)} \left(1 + \frac{t^2}{n}\right)^{-\frac{n+1}{2}}$$

2.1.8 F-Distribution

$$\begin{split} F(m,n) &= \frac{U/m}{V/n}, U \sim \chi_m^2, V \sim \chi_n^2 \\ f(w) &= \frac{\Gamma\left(\frac{n+m}{2}\right)}{\Gamma\left(\frac{m}{2}\right)\Gamma\left(\frac{n}{2}\right)} \cdot \left(\frac{m}{n}\right)^{\frac{m}{2}} \cdot w^{\frac{m}{2}-1} \cdot \left(1 + \frac{m}{n}w\right)^{-\frac{m+n}{2}} \end{split}$$

2.2 MGF

$$M(X) = \mathbb{E}\left[e^{tX}\right]$$
$$M^{(m)}(X) = \mathbb{E}\left[X^m\right]$$

Distribution	MGF	PMF/PDF
Bernoulli(p)	$pe^t + 1 - p$	p(1) = p
Binomial (n, p)	$(1 - p + pe^t)^n$	$\binom{n}{k} p^k (1-p)^{n-k}$
$Poisson(\lambda)$	$e^{\lambda(e^t-1)}$	$p(k) = \frac{\lambda^k}{k!} e^{-\lambda}$
Geo(p)	$\frac{pe^t}{1 - (1 - p)e^t}$	$(1-p)^{k-1}p$
$\mathcal{N}(\mu, \sigma^2)$	$e^{t\mu + \frac{1}{2}\sigma^2 t^2}$	$\frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{(x-\mu)^2}{2\sigma^2}}$
$\operatorname{Exp}(\lambda)$	$\frac{\lambda}{\lambda - t}$	$\lambda e^{-\lambda x}$
$\Gamma(\alpha, \beta)$	$(1-\beta t)^{-\alpha}$	$\frac{1}{\beta^k \Gamma(k)} x^{k-1} e^{-\frac{x}{\beta}}$
χ_k^2	$(1-2t)^{-\frac{k}{2}}$	$\frac{1}{2^{\frac{k}{2}}\Gamma(\frac{k}{2})}x^{\frac{k}{2}-1}e^{-\frac{x}{2}}$

2.3 Central Limit Theorem

2.3.1 Inequalities

Markov's inequality:
$$\Pr\{X \ge t\} \le \frac{\mathbb{E}(X)}{t}$$
, 要求是 $X \ge 0, t > 0$

Chebyshev's Inequality:
$$\Pr\{|X - \mathbb{E}(X)| \ge t\} \le \frac{\operatorname{Var}(X)}{t^2}$$

$$\Pr\{|X - \mathbb{E}(X)| \ge k\sigma\} \le \frac{1}{k^2}$$

2.3.2 Law of Large Numbers

$$| \overline{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$$
Weak: $\lim_{n \to \infty} \mathbb{P}r \{ | \overline{X}_n - \mu | > \epsilon \} = 0$
Strong: $\mathbb{P}r \{ \lim_{n \to \infty} \overline{X}_n = \mu \} = 1$

2.3.3 CLT

$$\begin{cases} S_n = \sum_{i=1}^n X_i \\ \lim_{n \to \infty} \mathbb{P}r \left\{ \frac{S_n - n\mu}{\sigma \sqrt{n}} \le x \right\} = \Phi(x) \end{cases}$$

Estimation

3.1 Maximum Likelihood (MLE)

$$L(\boldsymbol{x} \mid \boldsymbol{\theta}) = \prod_{i=1}^{n} f(x_i \mid \boldsymbol{\theta})$$

Standard conditions:

- 1. $L(\theta) > 0$ for all $\theta \in (a, b)$
- 2. $\frac{\partial L(\theta)}{\partial \theta}$ exists for all $\theta \in (a, b)$ 3. $\lim_{\theta \to a^+} L(\theta) = \lim_{\theta \to b^-} L(\theta) = 0$

$$Bias(\theta) = \mathbb{E}_{\theta}(\theta) - \theta$$