ST455: Reinforcement Learning

Lecture 11: Off-Policy Evaluation

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Lecture Outline

- 1. Off-Policy Evaluation (OPE) Introduction
- 2. OPE in Contextual Bandits
- 3. OPE in Reinforcement Learning
- 4. State-of-the-Art OPE Algorithms

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What is Off-Policy Evaluation

- Objective: Evaluate the impact of a target policy offline using historical data generated from a different behavior policy
- **Setting**: Offline RL with a precollected data



(a) Health Care



(c) Ridesharing



(b) Robotics



(d) Auto-driving

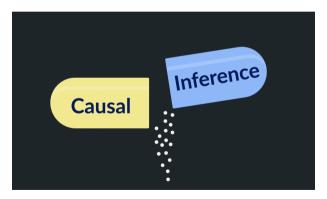
Why Off-Policy Evaluation

In many applications, it can be **dangerous** to evaluate a **target policy** by directly running this policy

- Healthcare: which medical treatment to suggest for a patient
- Ridesharing: which driver to assign for a call order
- Eduction: which curriculum to recommend for a student

Causal Inference

Off-policy evaluation is closely related to **causal inference**, whose objective is to learn the difference between a new treatment and a standard treatment



Causal Inference (Cont'd)

home / insights / agenda / causality and natural experiments the 2021 nobel prize in economic sciences

Causality and natural experiments: the 2021 Nobel Prize in Economic Sciences

26 NOV 2021

Offline Policy Optimisation

- Off-policy evaluation is also related to **offline** policy learning (Lecture 10), whose objective is to learn an optimal policy offline using historical data
- Suppose we are able to evaluate the **value** of any policy, it suffices to pick the policy that maximises the value

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Recap: Contextual Bandits

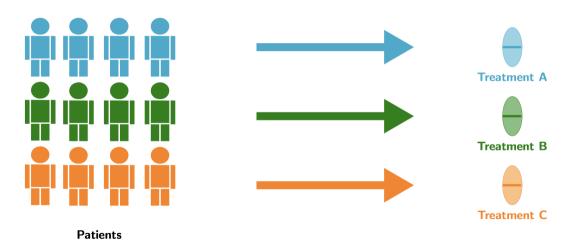
- Extension of MAB with **contextual** information.
- A widely-used model in medicine and technological industries.
- At time **t**, the agent
 - Observe a context S_t;
 - Select an action A_t;
 - Receives a reward R_t (depends on both S_t and A_t).
- **Objective**: Given an i.i.d. offline dataset $\{(S_t, A_t, R_t) : 0 \le t < T\}$ generated by a behavior policy b, i.e.,

$$\Pr(\mathbf{A_t} = \mathbf{a}|\mathbf{S_t} = \mathbf{s}) = \mathbf{b}(\mathbf{a}|\mathbf{s}),$$

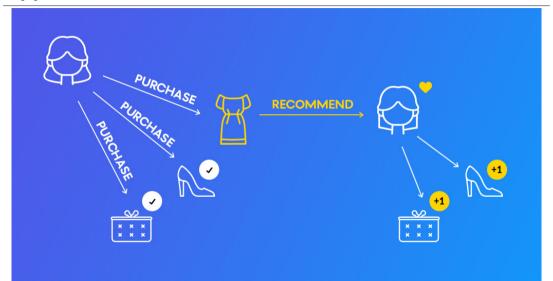
we aim to evaluate the mean outcome under a target policy π , i.e.,

$$\Pr(\mathbf{A_t} = \mathbf{a}|\mathbf{S_t} = \mathbf{s}) = \pi(\mathbf{a}|\mathbf{s}).$$

Application I: Precision Medicine

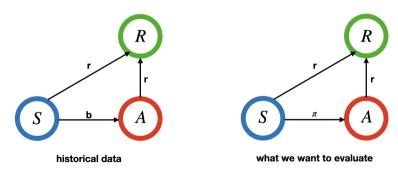


Application II: Personalized Recommendation



Challenge

- Confounding: State serves as confounding variables that confound the action-reward pair
- Distributional shift: The target policy generally differs from the behavior policy



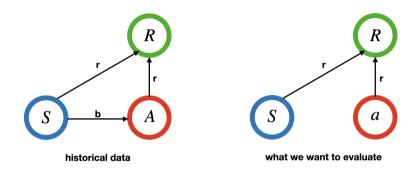
Challenge (Cont'd)

• Suppose π is a nondynamic policy, i.e., there exists some a such that $\pi(a|s) = 1$ for any s. We aim to evaluate the value under a given action a. A naive estimator is

$$\frac{\sum_{t=0}^{T-1} R_t \mathbb{I}(A_t = a)}{\sum_{t=0}^{T-1} \mathbb{I}(A_t = a)} \stackrel{P}{\to} \mathbb{E}(R|A = a)$$

• This estimator is valid only when no confounding variables exist

Challenge (Cont'd)



According to the causal diagram, the target policy's value equals

$$\mathbb{E}[\mathbb{E}(R|A=a,S)] \neq \mathbb{E}(R|A=a)$$

OPE Estimators

• With a general target policy π , the target policy's value equals

$$\sum_{\mathbf{a}} \mathbb{E}[\pi(\mathbf{a}|\mathbf{S})\mathbb{E}(R|\mathbf{A}=\mathbf{a},\mathbf{S})] = \sum_{\mathbf{a}} \mathbb{E}[\pi(\mathbf{a}|\mathbf{S})r(\mathbf{S},\mathbf{a})],$$

where
$$r(s, a) = \mathbb{E}(R|A = a, S = s)$$

- Direct estimator
- Importance sampling estimator
- Doubly robust estimator

Direct Estimator

Given that the target policy's value is given by

$$\sum_{\mathbf{a}} \mathbb{E}[\pi(\mathbf{a}|\mathbf{S})r(\mathbf{S},\mathbf{a})]$$

• The expectation can be approximated by the sample average, i.e.,

$$\frac{1}{T} \sum_{\mathbf{a}} \sum_{t=0}^{T-1} [\pi(\mathbf{a}|\mathbf{S}_t)r(\mathbf{S}_t,\mathbf{a})]$$

• The reward function can be replaced with some estimator \hat{r} . This yields the direct estimator

$$\frac{1}{T} \sum_{\mathbf{a}} \sum_{t=0}^{T-1} [\pi(\mathbf{a}|\mathbf{S}_t) \hat{\mathbf{r}}(\mathbf{S}_t, \mathbf{a})]$$

Direct Estimator (Cont'd)

• \hat{r} estimated using supervised learning

$$S_0, A_0 \rightarrow R_0$$

$$S_1, A_1 \rightarrow R_1$$

$$\vdots$$

$$S_{T-1}, A_{T-1} \rightarrow R_{T-1}$$

- Loss function: least square/Huber loss
- Computer parameter that minimizes empirical loss

Importance Sampling Estimator

• Given that the target policy's value is given by

$$\sum_{\mathbf{a}} \mathbb{E}[\pi(\mathbf{a}|\mathbf{S})r(\mathbf{S},\mathbf{a})]$$

• By the change of measure theory, it equals

$$\sum_{\mathbf{a}} \mathbb{E}\left[\mathbf{b}(\mathbf{a}|\mathbf{S}) \frac{\pi(\mathbf{a}|\mathbf{S})}{\mathbf{b}(\mathbf{a}|\mathbf{S})} \mathbf{r}(\mathbf{S}, \mathbf{a})\right] = \mathbb{E}\left[\frac{\pi(\mathbf{A}|\mathbf{S})}{\mathbf{b}(\mathbf{A}|\mathbf{S})} \mathbf{r}(\mathbf{S}, \mathbf{A})\right] = \mathbb{E}\left[\frac{\pi(\mathbf{A}|\mathbf{S})}{\mathbf{b}(\mathbf{A}|\mathbf{S})} R\right]$$

• This yields the following importance sampling (IS) estimator [Zhang et al., 2012]

$$\frac{1}{T}\sum_{t=0}^{T-1}\frac{\pi(\mathbf{A}_t|\mathbf{S}_t)}{\widehat{b}(\mathbf{A}_t|\mathbf{S}_t)}R_t,$$

for a given estimator $\hat{\boldsymbol{b}}$

Importance Sampling Estimator (Cont'd)

- The ratio $\pi(a|s)/b(a|s)$ is referred to as the **importance sampling ratio**
- It measures the difference between the behavior and target policies
- When $\pi = \boldsymbol{b}$, the ratio equals $\boldsymbol{1}$ for any \boldsymbol{a} and \boldsymbol{s}
- In general, the ratio centres at 1

$$\mathbb{E}\left[rac{\pi(oldsymbol{A}|oldsymbol{S})}{oldsymbol{b}(oldsymbol{A}|oldsymbol{S})}
ight]=1$$

Importance Sampling Estimator (Cont'd)

- In randomized studies, b is known
- In **observational studies**, **b** needs to be estimated from data
- \hat{b} estimated using supervised learning

- Loss function: logistic regression loss
- Computer parameter that minimizes empirical loss

Direct Estimator v.s. IS Estimator

- Bias/Variance Trade-Off
- The direct estimator has **some bias**, since r needs to be estimated from data
- The IS estimator has **zero bias** when **b** is known as in randomized studies
- The IS estimator might have a large variance when π differs significantly from \boldsymbol{b}
- Suppose $R = r(S, A) + \varepsilon$ for some ε independent of (S, A),

$$\operatorname{Var}\left[\frac{\pi(\boldsymbol{A}|\boldsymbol{S})}{\boldsymbol{b}(\boldsymbol{A}|\boldsymbol{S})}R\right] = \mathbb{E}\left[\frac{\pi(\boldsymbol{A}|\boldsymbol{S})}{\boldsymbol{b}(\boldsymbol{A}|\boldsymbol{S})}\{R - \boldsymbol{r}(\boldsymbol{S}, \boldsymbol{A})\}\right]^2 + \text{some term}$$
$$= \sigma^2 \mathbb{E}\left[\frac{\pi^2(\boldsymbol{A}|\boldsymbol{S})}{\boldsymbol{b}^2(\boldsymbol{A}|\boldsymbol{S})}\right] + \text{some term},$$

where
$$\sigma^2 = \operatorname{Var}(\varepsilon)$$

Extensions

- When π differs from **b** significantly, IS estimator suffers from **large variance** and becomes **unstable**
- Solutions sought by using self-normalized and/or truncated IS
- Self-normalized IS

$$\left[\frac{1}{T}\sum_{t=0}^{T-1}\frac{\pi(\boldsymbol{A}_{t}|\boldsymbol{S}_{t})}{\boldsymbol{b}(\boldsymbol{A}_{t}|\boldsymbol{S}_{t})}\right]^{-1}\frac{1}{T}\sum_{t=0}^{T-1}\frac{\pi(\boldsymbol{A}_{t}|\boldsymbol{S}_{t})}{\boldsymbol{b}(\boldsymbol{A}_{t}|\boldsymbol{S}_{t})}R_{t}$$

• Equivalent to IS estimator in large samples, by noting that

$$\frac{1}{T} \sum_{t=0}^{T-1} \frac{\pi(A_t|S_t)}{b(A_t|S_t)} \stackrel{P}{\to} \mathbb{E}\left[\frac{\pi(A|S)}{b(A|S)}\right] = 1$$

• Stabilize the important sampling ratio in finite samples

Extensions (Cont'd)

Truncated IS

$$\frac{1}{T} \sum_{t=0}^{T-1} \frac{\pi(\mathbf{A}_t | \mathbf{S}_t)}{\max(\widehat{\mathbf{b}}(\mathbf{A}_t | \mathbf{S}_t), \varepsilon)} R_t,$$

for some $\varepsilon > 0$

- Truncate the behavior policy whose value is smaller than arepsilon
- Avoid extremely large ratio, thus reducing the variance of the estimator

Doubly Robust Estimator

Direct estimator

$$\frac{1}{T} \sum_{\mathbf{a}} \sum_{t=0}^{T-1} [\pi(\mathbf{a}|\mathbf{S}_t) \hat{\mathbf{r}}(\mathbf{S}_t, \mathbf{a})]$$

requires \hat{r} to be consistent

IS estimator

$$\frac{1}{T}\sum_{t=0}^{T-1}\frac{\pi(\mathbf{A}_t|\mathbf{S}_t)}{\widehat{b}(\mathbf{A}_t|\mathbf{S}_t)}R_t,$$

requires $\widehat{\boldsymbol{b}}$ to be consistent

• Doubly robust (DR) estimator combines both, and requires **either** \hat{r} **or** \hat{b} to be consistent ("doubly-robustness" property)

Doubly Robust Estimator (Cont'd)

Consider the estimating function

$$\phi(\mathbf{S}, \mathbf{A}, R) = \sum_{\mathbf{a}} \pi(\mathbf{a}|\mathbf{S}) \widehat{r}(\mathbf{S}, \mathbf{a}) + \frac{\pi(\mathbf{A}|\mathbf{S})}{\widehat{b}(\mathbf{A}|\mathbf{S})} [R - \widehat{r}(\mathbf{S}, \mathbf{A})]$$

- First term on the RHS is the estimating function of the direct estimator
- Second term corresponds to the augmentation term
 - Zero mean when $\hat{r} = r$
 - Debias the bias of the direct estimator
 - Offering additional robustness against model misspecification of \hat{r}
- DR estimator given by $T^{-1} \sum_{t=0}^{T-1} \phi(S_t, A_t, R_t)$

Fact 1: Double Robustness

The estimating function

$$\phi(\mathbf{S}, \mathbf{A}, R) = \sum_{\mathbf{a}} \pi(\mathbf{a}|\mathbf{S}) \hat{\mathbf{r}}(\mathbf{S}, \mathbf{a}) + \frac{\pi(\mathbf{A}|\mathbf{S})}{\hat{\mathbf{b}}(\mathbf{A}|\mathbf{S})} [R - \hat{\mathbf{r}}(\mathbf{S}, \mathbf{A})]$$

- In large sample size, DR estimator converges to $\mathbb{E}\phi(S, A, R)$
- When $\hat{r} = r$, the augmentation term has zero mean. It follows that

$$\mathbb{E}\phi(S,A,R) = \sum_{a} \mathbb{E}[\pi(a|S)r(S,a)] = \text{target policy's value}$$

• When $\hat{b} = b$, it has the same mean as the IS estimator

$$\mathbb{E}\phi(S, A, R) = \mathbb{E}\left[\frac{\pi(A|S)}{b(A|S)}R\right] + \mathbb{E}\left[\sum_{a}\pi(a|S)\widehat{r}(S, a) - \frac{\pi(A|S)}{b(A|S)}\widehat{r}(S, A)\right]$$
$$= \mathbb{E}\left[\frac{\pi(A|S)}{b(A|S)}R\right] = \text{target policy's value}$$

Fact 2: Efficiency

• When $\hat{\boldsymbol{b}} = \boldsymbol{b}$, the estimating function

$$\phi(\mathbf{S}, \mathbf{A}, R) = \sum_{\mathbf{a}} \pi(\mathbf{a}|\mathbf{S}) \hat{r}(\mathbf{S}, \mathbf{a}) + \frac{\pi(\mathbf{A}|\mathbf{S})}{b(\mathbf{A}|\mathbf{S})} [R - \hat{r}(\mathbf{S}, \mathbf{A})]$$

• The MSE of DR estimator is proportional to the variance of $\phi(S, A, R)$

$$\operatorname{Var}(\phi(\boldsymbol{S},\boldsymbol{A},R)) = \mathbb{E}[\operatorname{Var}(\phi(\boldsymbol{S},\boldsymbol{A},R)|\boldsymbol{S},\boldsymbol{A})] + \operatorname{Var}[\mathbb{E}(\phi(\boldsymbol{S},\boldsymbol{A},R)|\boldsymbol{S},\boldsymbol{A})]$$

- The first term on the RHS is independent of \hat{r}
- The second term is minimized when $\hat{r} = r$
- ullet A good working model for $oldsymbol{r}$ improves the estimator's efficiency
- When $\hat{r} = r$, the estimator achieves the **efficiency bound** [e.g., smallest MSE among a class of regular estimators; see Tsiatis, 2007]

Fact 3: Efficiency

- When $\hat{\boldsymbol{b}}$ is estimated from data and the model is **correctly specified**, the estimator's MSE would be **generally smaller than** the one that uses the oracle behavior policy \boldsymbol{b} [Tsiatis, 2007]
- Estimating $\hat{\boldsymbol{b}}$ yields a more efficient estimator, even if we know the oracle \boldsymbol{b}
- Multi-armed bandit example without context information
 - Objective: evaluate $\mathbb{E}(R|A=a)$ for a given a
 - IS estimator with **known** Pr(A = a)

$$\frac{\sum_{t=0}^{T-1} \mathbb{I}(A_t = a) R_t}{T \Pr(A_t = a)}$$

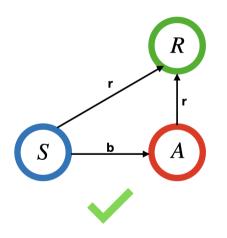
• IS estimator with **estimated** Pr(A = a) has a **smaller** asymptotic variance

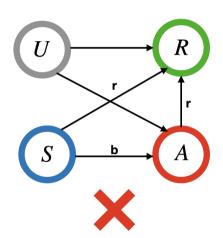
$$\frac{\sum_{t=0}^{T-1} \mathbb{I}(\mathbf{A}_t = \mathbf{a}) R_t}{\sum_{t=0}^{T-1} \mathbb{I}(\mathbf{A}_t = \mathbf{a})}$$

Assumption: No Unmeasured Confounders

- All three estimators (direct estimator, IS, DR) rely on the no unmeasured confounders assumption
- They are biased when this assumption is violated
- It requires all confounders that confound the action-reward relationship are included in the state
- This assumption is cannot be verified in practice
- When violated, we may use some auxiliary variable (e.g., instrumental variables, mediators) for consistent policy evaluation [Angrist et al., 1996, Pearl, 2009]

Assumption: No Unmeasured Confounders (Cont'd)





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General OPE Problem

• Objective: Given an offline dataset $\{(S_{i,t}, A_{i,t}, R_{i,t}) : 1 \le i \le N, 0 \le t \le T\}$ generated by a behavior policy b, where i indexes the ith episode and t indexes the tth time point, we aim to evaluate the mean return under a target policy π

$$\mathbb{E}^{\pi}\left[\sum_{t=0}^{\infty} \gamma^{t} R_{t}\right] = \mathbb{E} \boldsymbol{V}^{\pi}(\boldsymbol{S}_{0})$$

When $\gamma = 1$, the task is assumed to be episodic

- We focus on the case where both π and \boldsymbol{b} are stationary policies
- Challenge: **Distributional shift**
 - In the offline dataset, actions are generated according to **b**
 - ullet The target policy π we wish to evaluate is different from $oldsymbol{b}$
- Existing prediction algorithms (e.g., MC, TD) designed in online settings are not applicable

Recap: MC Prediction

- **Objective**: learns V^{π} from experience under π
- MC Policy Evaluation: $V(s) \leftarrow average[Returns(s)]$
- Incremental update for every-visit MC prediction:

$$V(S_t) \leftarrow V(S_t) + \alpha_t[G_t - V(S_t)]$$

where α_t is $\frac{1}{\#[\mathsf{Returns}(S_t)]}$ at time t

- The update can be performed after return G_t is observed
- i.e. after the episode is completed

Recap: TD Prediction

- Unlike MC methods, TD methods wait only until next time step
- The simplest TD method (so called TD(0)) considers the update

$$V(S_t) \leftarrow V(S_t) + \alpha_t[R_t + \gamma V(S_{t+1}) - V(S_t)]$$

- This update rule has $R_t + \gamma V(S_{t+1})$ as the target
- Considered as a **bootstrap** method: update in part based on an existing estimate

Direct Estimator

• The target policy's value is given by $\mathbb{E} V^{\pi}(S_0)$, or equivalently,

$$\mathbb{E}[\sum_{m{a}}\pi(m{a}|m{S_0})m{Q}^\pi(m{S_0},m{a})]$$

- The expectation can be approximated via the **empirical initial state distribution**
- Q-learning is an **off-policy** algorithm. Can be applied to learn Q^{π} offline
- This yields the direct estimator

$$\frac{1}{N}\sum_{i=1}^{N}\sum_{\boldsymbol{a}}\pi(\boldsymbol{a}|\boldsymbol{S}_{i,0})\widehat{Q}(\boldsymbol{S}_{i,0},\boldsymbol{a})$$

ullet It remains to compute $\widehat{oldsymbol{Q}}$

Recap: Fitted Q-Iteration in Offline Setting

- Offline data: $\{(S_{i,t}, A_{i,t}, R_{i,t}) : 0 \le t \le T, 1 \le i \le N\}$
- Fitted Q-Iteration can be naturally applied by repeating
 - 1. Compute \widehat{Q} as the argmin of

$$rg \min_{m{Q}} \sum_{m{t}} \left[m{R}_{i,t} + \gamma \max_{m{a}} \widetilde{m{Q}}(m{S}_{i,t+1}, m{a}) - m{Q}(m{S}_{i,t}, m{A}_{i,t})
ight]^2$$

- 2. Set $\widetilde{\pmb{Q}} = \widehat{\pmb{Q}}$
- Designed for learning $Q^{\pi^{\text{opt}}}$
- Do not require actions to follow the greedy policy

Fitted Q-Evaluation [Le et al., 2019]

Bellman equation

$$\mathbb{E}\left[R_t + \gamma \pi(\mathbf{a}|\mathbf{S}_{t+1})Q^{\pi}(\mathbf{S}_{t+1},\mathbf{a})|\mathbf{S}_t,\mathbf{A}_t\right] = Q^{\pi}(\mathbf{S}_t,\mathbf{A}_t)$$

- ullet Both LHS and RHS involves $oldsymbol{Q}^{\pi}$
- Repeat the following procedure
 - 1. Compute $\widehat{\boldsymbol{Q}}$ as the argmin of

$$\arg\min_{\boldsymbol{Q}} \sum_{\boldsymbol{t}} \left[R_{i,t} + \gamma \sum_{\boldsymbol{a}} \pi(\boldsymbol{a}|\boldsymbol{S}_{i,t+1}) \widetilde{\boldsymbol{Q}}(\boldsymbol{S}_{i,t+1},\boldsymbol{a}) - \boldsymbol{Q}(\boldsymbol{S}_{i,t},\boldsymbol{A}_{i,t}) \right]^2$$

- 2. Set $\widetilde{\pmb{Q}} = \widehat{\pmb{Q}}$
- Designed for learning Q^{π}
- Do **not** require actions to follow the target policy

Stepwise IS Estimator [Zhang et al., 2013]

- Consider episodic task where T is the termination time
- Standard MC prediction is **not** applicable under **distributional shift**
- Importance sampling ratio needs to be employed

$$\mathbb{E}^{\pi}R_{0} = \mathbb{E}^{b} \left[\frac{\pi(A_{0}|S_{0})}{b(A_{0}|S_{0})} R_{0} \right]$$

$$\mathbb{E}^{\pi}R_{1} = \mathbb{E}^{b} \left[\frac{\pi(A_{0}|S_{0})}{b(A_{0}|S_{0})} \frac{\pi(A_{1}|S_{1})}{b(A_{1}|S_{1})} R_{1} \right]$$

$$\vdots$$

$$\mathbb{E}^{\pi}R_{t} = \mathbb{E}^{b} \left[\frac{\pi(A_{0}|S_{0})}{b(A_{0}|S_{0})} \cdots \frac{\pi(A_{t}|S_{t})}{b(A_{t}|S_{t})} R_{t} \right]$$

Stepwise IS Estimator (Cont'd)

• According to this logic, the target policy's value can be represented by

$$\mathbb{E}\left[\sum_{t=0}^{T} \gamma^{t} \left\{\prod_{j=0}^{t} \frac{\pi(\mathbf{A}_{j}|\mathbf{S}_{j})}{b(\mathbf{A}_{j}|\mathbf{S}_{j})}\right\} R_{t}\right]$$

• This yields the stepwise IS estimator

$$\frac{1}{N} \sum_{i=1}^{N} \left[\sum_{t=0}^{T} \gamma^{t} \left\{ \prod_{j=0}^{t} \frac{\pi(\mathbf{A}_{i,j}|\mathbf{S}_{i,j})}{\widehat{b}(\mathbf{A}_{i,j}|\mathbf{S}_{i,j})} \right\} R_{i,t} \right]$$

for a given estimator $\hat{\boldsymbol{b}}$ computed using supervised learning algorithms

Limitation

- Stepwise IS suffers from a large variance
- In particular, the IS ratio at time t is the product of individual ratios from the **initial** time to time t

$$\prod_{j=0}^t rac{\pi(extsf{A}_j| extsf{S}_j)}{b(extsf{A}_j| extsf{S}_j)}$$

- Variance of the ratio grows exponentially with respect to t, referred to as the curse of horizon [Liu et al., 2018]
- Extension: Doubly-robust estimator by [Jiang and Li, 2016]

Pros & Cons of Direct v.s. Stepwise IS

- Stepwise IS is similar to an offline version of MC
- SIS learns from **complete** sequences
- SIS only works for episodic (terminating) environments

- Direct estimator (DE) is similar to an offline version of TD
- DE can learn from incomplete sequences
- DE works in **continuing** environments

Pros & Cons of Direct v.s. Stepwise IS (Cont'd)

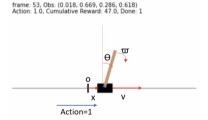
- Bias/Variance Trade-Off
- When **b** is known, stepwise IS is an **unbiased** estimator since

$$\mathbb{E}^{\boldsymbol{\pi}} R_t = \mathbb{E}^{\boldsymbol{b}} \left[\frac{\pi(\boldsymbol{A_0}|\boldsymbol{S_0})}{\boldsymbol{b}(\boldsymbol{A_0}|\boldsymbol{S_0})} \cdots \frac{\pi(\boldsymbol{A_t}|\boldsymbol{S_t})}{\boldsymbol{b}(\boldsymbol{A_t}|\boldsymbol{S_t})} R_t \right]$$

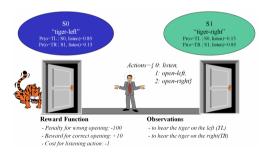
- Direct estimator has **some bias**, since Q^{π} needs to be estimated from data
- Stepwise IS suffers from curse of horizon and a large variance
- Direct estimator has a much lower variance

Pros & Cons of Direct v.s. Stepwise IS (Cont'd)

- Direct estimator exploits Markov & stationary properties
- Relies on the Bellman equation
- More **efficient** in MDP environments



- SIS does **not** exploit these properties
- More flexible in non-MDP environments (e.g., POMDP)



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Recap: RL Models

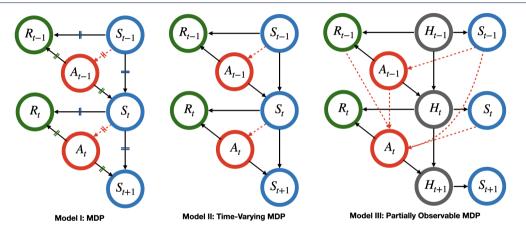


Figure: Causal diagrams for MDPs, TMDPs and POMDPs. Solid lines represent the causal relationships. Dashed lines indicate the information needed to implement the optimal policy. $\{H_t\}_t$ denotes latent variables. The parallel sign \parallel indicates that the conditional probability function given parent nodes is equal.

Marginalized IS Estimator

- As we have discussed, stepwise IS suffers from curse of horizon
- Curse of horizon is unavoidable in general Non-Markov decision processes (e.g., POMDP)
- Under some additional model assumptions (e.g., Markovianity & time-homogeneity), it is possible to break the curse of horizon using **marginalized IS** estimator
- Stepwise IS does **not** exploit these properties

Marginalized IS Estimator (Cont'd)

• Stepwise IS uses the **cumulative** IS ratio

$$\mathbb{E}^{\pi}R_t = \mathbb{E}^{b}\left[\frac{\pi(A_0|S_0)}{b(A_0|S_0)}\cdots\frac{\pi(A_t|S_t)}{b(A_t|S_t)}R_t\right]$$

• Under Markovianity (TMDP), marginalized IS uses the marginalized IS ratio

$$\mathbb{E}^{\pi} R_{t} = \mathbb{E}^{b} \left[\frac{\boldsymbol{p}_{t}^{\pi}(\boldsymbol{S}_{t}, \boldsymbol{A}_{t})}{\boldsymbol{p}_{t}^{b}(\boldsymbol{S}_{t}, \boldsymbol{A}_{t})} R_{t} \right]$$
(1)

where $m{p}_t^{\pi}$ and $m{p}_t^{m{b}}$ are the marginal density functions of $(m{S}_t, m{A}_t)$ under π and $m{b}$

• The resulting marginalized IS estimator can be derived from (1)

Marginalized IS Estimator

• Under Markovianity and time-homogeneity (MDP),

$$\mathbb{E}V^{\pi}(S_0) = \mathbb{E}^{b} \left[\frac{\sum_{t=0}^{\infty} \gamma^{t} p_t^{\pi}(S, A)}{p_{\infty}(S, A)} R \right]$$
 (2)

where p_{∞} denotes the limiting state-action distribution under b and the numerator corresponds to the γ -discounted state-action visitation probability

• The resulting marginalized IS estimator can be derived from (2)

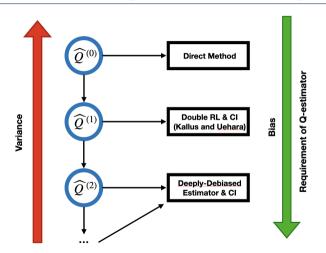
Double RL [Kallus and Uehara, 2019]

- Double RL extends DR in **contextual bandits** to the general RL problem
- Similar to DR, the estimator can be represented as

Direct Estimator + Augmentation Term

- Augmentation term is to debias the bias of direct estimator and offer protection against model misspecification of Q^{π} ; it relies on the marginalized IS ratio
- Similar to DR, the estimator is **doubly-robust**, e.g., consistent when either Q^{π} or the marginalized IS ratio is correct
- Similar to DR, the estimator achieves the efficiency bound in MDPs

Deeply-Debiased OPE [Shi et al., 2021]



- Ensures the bias decays much faster than standard deviation
- Allows to provide valid **uncertainty quantification** (e.g., confidence interval)

Summary

- Off-policy evaluation
- Direct estimator
- Importance sampling estimator
- Doubly robust estimator

- Fitted Q-evaluation
- Stepwise IS/Marginalized IS
- Double reinforcement learning
- Deeply-debiased estimator

References I

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Questions