



```

mpg | -49.51222  86.15604  -0.57  0.567  -221.3025  122.278
name: <unnamed>
log:  H:\My Drive\Econ 640\Homework 4\Homework 4 Log.smcl
log type: smcl
opened on:  1 Dec 2025, 10:52:41

```

```

1 .
2 . *Part A*
3 .
4 . sysuse auto
   (1978 automobile data)

```

```
5 . reg price weight mpg
```

Source	SS	df	MS	Number of obs	=	74
Model	186321280	2	93160639.9	F(2, 71)	=	14.74
Residual	448744116	71	6320339.67	Prob > F	=	0.0000
				R-squared	=	0.2934
				Adj R-squared	=	0.2735
Total	635065396	73	8699525.97	Root MSE	=	2514

  

price	Coefficient	Std. err.	t	P> t	[95% conf. interval]
weight	1.746559	.6413538	2.72	0.008	.467736 3.025382
mpg	-49.51222	86.15604	-0.57	0.567	-221.3025 122.278
_cons	1946.069	3597.05	0.54	0.590	-5226.245 9118.382

```
6 . estat hettest
```

Breusch-Pagan/Cook-Weisberg test for heteroskedasticity  
Assumption: Normal error terms  
Variable: Fitted values of **price**

H0: Constant variance

```

chi2(1) = 14.78
Prob > chi2 = 0.0001

```

```

7 .
8 . /* Since the p-value is 0.0001, we would reject the null hypothesis that the
   > variance is homoskedastic. This implies that we have are dealing with
   > heteroskedasticity */
9 .
10 . regress price weight mpg, robust

```

```

Linear regression
Number of obs   = 74
F(2, 71)        = 14.84
Prob > F         = 0.0000
R-squared       = 0.2934
Root MSE       = 2514

```

price	Coefficient	Robust std. err.	t	P> t	[95% conf. interval]
weight	1.746559	.777837	2.25	0.028	.1955963 3.297522
mpg	-49.51222	95.8074	-0.52	0.607	-240.5468 141.5223
_cons	1946.069	4213.793	0.46	0.646	-6455.995 10348.13

```

11 .
12 . /* The standard errors of the robust regression are larger than just the regular
    > regression. By using robust, we relax the assumption that variance of the errors
    > is constant across all observations. This will lead to higher uncertainty,
    > making the robsut standard errors larger */
13 .
14 . *Part B*
15 .
16 . sysuse nls88, clear
    (NLSW, 1988 extract)
17 . reg wage age collgrad

```

Source	SS	df	MS	Number of obs	=	2,246
Model	5394.00567	2	2697.00283	F(2, 2243)	=	87.71
Residual	68973.9617	2,243	30.7507631	Prob > F	=	0.0000
				R-squared	=	0.0725
				Adj R-squared	=	0.0717
Total	74367.9674	2,245	33.1260434	Root MSE	=	5.5453

wage	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
age	-.0643324	.0382481	-1.68	0.093	-.1393378	.010673
collgrad	3.612098	.2752221	13.12	0.000	3.072381	4.151815
_cons	9.430184	1.503989	6.27	0.000	6.480828	12.37954

```
18 . reg wage age collgrad, cluster(occupation)
```

Linear regression	Number of obs	=	2,237
	F(2, 12)	=	14.91
	Prob > F	=	0.0006
	R-squared	=	0.0733
	Root MSE	=	5.5487

(Std. err. adjusted for 13 clusters in occupation)

wage	Coefficient	Robust std. err.	t	P> t	[95% conf. interval]	
age	-.0669641	.0431162	-1.55	0.146	-.1609062	.0269781
collgrad	3.632842	.6751054	5.38	0.000	2.161914	5.10377
_cons	9.539055	2.092498	4.56	0.001	4.979894	14.09822

```

19 .
20 . /* The cluster-robust standard errors are a lot larger than the OLS standard
    > errors. This is because they allow for correlation within each cluster which
    > causes the variance to increase as a result of the dependency. */
21 .
22 . *Part C*
23 .
24 . bcuse cps91, clear

```

Contains data from <http://fmwww.bc.edu/ec-p/data/wooldridge/cps91.dta>

Observations: 5,634

Variables: 24 20 May 2002 11:05

Variable name	Storage type	Display format	Value label	Variable label
husage	byte	%8.0g		husband's age
husunion	byte	%8.0g		=1 if hus. in union
husearns	int	%8.0g		hus. weekly earns
huseduc	byte	%8.0g		husband's yrs schooling
husblack	byte	%8.0g		=1 if hus. black

hushisp	byte	%8.0g	=1 if hus. hispanic
hushrs	byte	%8.0g	hus. weekly hours
kidge6	byte	%8.0g	=1 if have child >= 6
earns	float	%8.0g	wife's weekly earnings
age	byte	%8.0g	wife's age
black	byte	%8.0g	=1 if wife black
educ	byte	%8.0g	wife's yrs schooling
hispanic	byte	%8.0g	=1 if wife hispanic
union	byte	%8.0g	=1 if wife in union
faminc	float	%9.0g	annual family income
husexp	byte	%8.0g	huseduc - husage - 6
exper	byte	%8.0g	age - educ - 6
kidlt6	byte	%8.0g	=1 if have child < 6
hours	int	%9.0g	wife's weekly hours
expersq	int	%8.0g	exper^2
nwifeinc	float	%9.0g	non-wife inc, \$1000s
inlf	byte	%8.0g	=1 if wife in labor force
hrwage	float	%9.0g	earns/hours
lwage	float	%9.0g	log(hrwage)

Sorted by:

```
25 . gen lnexper = ln(exper)
    (20 missing values generated)
```

```
26 . reg hrwage lnexper
```

Source	SS	df	MS	Number of obs	=	3,276
Model	2.93774269	1	2.93774269	F(1, 3274)	=	0.06
Residual	162316.271	3,274	49.5773582	Prob > F	=	0.8077
				R-squared	=	0.0000
				Adj R-squared	=	-0.0003
Total	162319.209	3,275	49.5631171	Root MSE	=	7.0411

  

hrwage	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
lnexper	.0459342	.1886995	0.24	0.808	-.3240469	.4159152
_cons	10.25202	.5408182	18.96	0.000	9.191647	11.3124

```
27 .
28 . /* The coefficient on lnexper is 0.0459. This means that a 1% change in
    > experience is associated with a change in hourly wages of 0.000459. */
29 .
30 . reg hrwage exper
```

Source	SS	df	MS	Number of obs	=	3,286
Model	109.6416	1	109.6416	F(1, 3284)	=	2.22
Residual	162457.898	3,284	49.4695184	Prob > F	=	0.1367
				R-squared	=	0.0007
				Adj R-squared	=	0.0004
Total	162567.54	3,285	49.4878356	Root MSE	=	7.0335

  

hrwage	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
exper	-.0184498	.0123929	-1.49	0.137	-.0427484	.0058488
_cons	10.72282	.268537	39.93	0.000	10.19631	11.24934

```

31 .
32 . /* If we were to regress without logging experience then the coefficient would
> be -0.0184. This would mean that a one unit-change in experince would be
> associated with a -0.0184 change in horuly wages. However, this implies that
> increasing experience would diminish wages. By logging experience, the
> diminishing effects of experience is accounted for */
33 .
34 . *Part D*
35 .
36 . bcuse bwght2, clear

```

Contains data from <http://fmwww.bc.edu/ec-p/data/wooldridge/bwght2.dta>

Observations: 1,832

Variables: 23

24 May 2002 08:47

Variable name	Storage type	Display format	Value label	Variable label
mage	byte	%10.0g		mother's age, years
meduc	byte	%10.0g		mother's educ, years
monpre	byte	%10.0g		month prenatal care began
npvis	byte	%10.0g		total number of prenatal visits
fage	byte	%10.0g		father's age, years
feduc	byte	%10.0g		father's educ, years
bwght	int	%10.0g		birth weight, grams
omaps	byte	%10.0g		one minute apgar score
fmaps	byte	%10.0g		five minute apgar score
cigs	byte	%10.0g		avg cigarettes per day
drink	byte	%10.0g		avg drinks per week
lbw	byte	%9.0g		=1 if bwght <= 2000
vlbw	byte	%9.0g		=1 if bwght <= 1500
male	byte	%9.0g		=1 if baby male
mwhte	byte	%9.0g		=1 if mother white
mblick	byte	%9.0g		=1 if mother black
moth	byte	%9.0g		=1 if mother is other
fwhte	byte	%9.0g		=1 if father white
fblick	byte	%9.0g		=1 if father black
foth	byte	%9.0g		=1 if father is other
lbwght	float	%9.0g		log(bwght)
mage <sup>2</sup>	int	%9.0g		mage^2
npvissq	int	%9.0g		npvis^2

Sorted by:

```
37 . logit lbw cigs mage
```

Iteration 0: Log likelihood = -134.8245

Iteration 1: Log likelihood = -133.21762

Iteration 2: Log likelihood = -132.74617

Iteration 3: Log likelihood = -132.74034

Iteration 4: Log likelihood = -132.74034

Logistic regression

Number of obs = 1,722

LR chi2(2) = 4.17

Prob > chi2 = 0.1244

Pseudo R2 = 0.0155

Log likelihood = -132.74034

	lbw	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
	cigs	.041738	.0315485	1.32	0.186	-.020096	.103572
	mage	-.067341	.0422063	-1.60	0.111	-.1500638	.0153818
	_cons	-2.307042	1.20524	-1.91	0.056	-4.669269	.055186

```
38 . margins, dydx(cigs)
```

Average marginal effects  
Model VCE: OIM

Number of obs = 1,722

Expression: `Pr(lbw), predict()`  
dy/dx wrt: `cigs`

	Delta-method				[95% conf. interval]	
	dy/dx	std. err.	z	P> z		
cigs	.0006189	.0004809	1.29	0.198	-.0003236	.0015613

```
39 .
```

```
40 . /* The marginal effect is about 0.0006. This means that a one-unit change in
> average cigarettes smoked during pregnancy can affect the probability that a
> baby will be born with low weight by 0.0006
>
> For an OLS regression, the coefficient would describe the direct change in the
> dependent variable if there was a one-unit change in the predictor variable.
> However, the coefficient of a logit model would describe the log probability of
> an outcome. The reason why logit coefficients are less straightforward is
> because the effect relies on the base probability of an event happening despite
> the coefficients staying the same
> */
```

```
41 .
```

```
42 . log close
      name: <unnamed>
      log:  H:\My Drive\Econ 640\Homework 4\Homework 4 Log.smcl
      log type: smcl
closed on:  1 Dec 2025, 10:52:46
```