

Scalable ML

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Space Complexity of Frequency Moments

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The space complexity of approximating the frequency moments

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Problem Statement

We are given a sequence $A = [a_1, a_2, \dots, a_m]$

$$a_i \in \{1, 2, \dots, n\}, \forall 1 \leq i \leq m$$

The elements a_1, a_2, \dots, a_m are given one by one,
and they cannot be stored in the memory or hard drive.

Example: $A = [1, 2, 1, 1, 3, 2, 1, 1]$,

$$a_i \in \{1, 2, 3\}, m = 8, n = 3$$

Problem Statement

Let m_i denote how many times we observed $i \in \{1, \dots, n\}$ in $A = [a_1, \dots, a_m]$

$$m_i = |\{j : a_j = i, 1 \leq j \leq m\}|$$

Observation: $\sum_{i=1}^n m_i = m$

Type	Number
1	m_1
2	m_2
...	...
i	m_i
...	...
n	m_n
Sum	m

Problem Statement

Goal: Estimate $F_k = \sum_{i=1}^n m_i^k$ frequency moments.

$$F_0 = \sum_{i=1}^n \underbrace{m_i^0}_1 = n$$

$$F_1 = \sum_{i=1}^n m_i = m$$

$$F_2 = \sum_{i=1}^n m_i^2$$

...

$$F_k = \sum_{i=1}^n m_i^k$$

...

$$F_{\infty}^* = \max_{1 \leq i \leq n} m_i$$

The Naïve Method

If we have at least $n \log m$ bit of memory,
then calculating $F_k = \sum_{i=1}^n m_i^k$ is easy:

- * Calculate m_1, m_2, \dots, m_n from $A = [a_1, \dots, a_m]$
- * Each can be stored on $\log m$ bits storage since $m_i \leq m$
- * We have n of these numbers
- * This together requires $n \log m$ bits of memory.

Surprising Fact

Informal (somewhat) surprising statement:

F_k can be estimated randomly with using only

$$O\left(n^{1-1/k}(\log n + \log m)\right) \quad \text{bits}$$

Formal Statement

Theorem 1: For every $k \geq 1$
 $\lambda > 0$
 $\epsilon > 0$

\Rightarrow There exists a randomized algorithm \mathcal{A} such that
given a sequence $A = [a_1, \dots, a_m]$ (where $a_i \in N \doteq \{1, 2, \dots, n\}$)

★ in one pass \mathcal{A} computes random $Y \in \mathbb{R}$ so that

$$\Pr(|Y - F_k| > \lambda F_k) < \epsilon \quad (*1)$$

★ \mathcal{A} is using only

$$O\left(\frac{k \log(1/\epsilon)}{\lambda^2} n^{1-1/k} (\log n + \log m)\right) \quad \text{memory bits.} \quad (*2)$$

Main Idea of the Algorithm

Main idea to estimate F_k :

Define a random variable whose

- expected value is F_k
- variance is relatively small

Preliminaries

Let $\lambda > 0$, $\epsilon > 0$ be fixed

$$\text{Let } S_1 \doteq \frac{12kn^{1-1/k}}{\lambda^2}$$

$$\text{Let } S_2 \doteq 2\log(1/\epsilon)$$

This 12 is only 8 in the paper,
but I can only see the proof with 12

We want to prove that \mathcal{A} will only use

$$O\left(\underbrace{\frac{k\log(1/\epsilon)}{\lambda^2}n^{1-1/k}}_{O(S_1S_2)}(\log n + \log m)\right) \quad \text{memory bits.}$$

Here we assume m is known in advance, i.e. \mathcal{A} can use m .

The paper also discusses the case when m is not known in advance, i.e. \mathcal{A} cannot use m .

The Random Algorithm to Calculate Y

The Algorithm \mathcal{A} will sample $\{X_{ij}\}$ i.i.d. random variables.

$$1 \leq j \leq S_1$$

$$1 \leq i \leq S_2$$

From these X_{ij} random variables we will calculate Y_1, Y_2, \dots, Y_{S_2} random variables

$$\begin{pmatrix} X_{1,1} & \dots & X_{1,S_1} \\ \vdots & & \vdots \\ X_{i,1} & \dots & X_{i,S_1} \\ \vdots & & \vdots \\ X_{S_2,1} & \dots & X_{S_2,S_1} \end{pmatrix} \Rightarrow \begin{aligned} Y_1 &\doteq \frac{1}{S_1} (X_{1,1} + \dots + X_{1,S_1}) \\ Y_i &\doteq \frac{1}{S_1} (X_{i,1} + \dots + X_{i,S_1}) \\ Y_{S_2} &\doteq \frac{1}{S_1} (X_{S_2,1} + \dots + X_{S_2,S_1}) \end{aligned}$$

Let $Y \doteq \text{median}(Y_1, Y_2, \dots, Y_{S_2})$

We will show that Y can be used to estimate F_k , i.e.

$$\Pr(|Y - F_k| > \lambda F_k) < \epsilon$$

Proof of Main Steps

To compute $X_{i,j}$ for a given i and j , we will use $O(\log n + \log m)$ bits

Therefore, to compute $\{X_{i,j}\}_{i=1}^{S_2},_{j=1}^{S_1}$, the algorithm \mathcal{A} will use

$$O(S_1 S_2 (\log n + \log m)) \text{ bits}$$

and this is what we had to prove for (*2) part of Theorem 1.

The remaining parts are (i) to show how to compute X_{ij} and (ii) to show that

$$\Pr\left(|Y - F_k| > \lambda F_k\right) < \epsilon$$

The Random Algorithm to Calculate X_{ij}

Let $p \sim U[1, 2, \dots, m]$


$\Rightarrow a_p$ is a random member of $A = [a_1, a_2, \dots, a_m]$

Suppose $a_p = l \in \{1, 2, \dots, n\}$

Let $r = |\{q : n \geq q \geq p, a_q = l\}|$

= num of occurances of l in A following a_p

Example: $A = [1, 2, 1, 1, 3, 2, 1, 1]$


 $p = 4$

$\Rightarrow r = |\{4, 7, 8\}| = 3$

Let $X \doteq m(r^k - (r - 1)^k)$

The Random Algorithm to Calculate X_{jj}

This is what we know so far:

$$r = |\{q : n \geq q \geq p, a_q = l\}|$$

= num of occurancies of l in A following a_p

$$X = m(r^k - (r - 1)^k)$$

Remark:

To be able to calculate $X = m(r^k - (r - 1)^k)$, we need

$\log n$ bits to store $a_p = l \in \{1, 2, \dots, n\}$.

(a_p is used to calculate r)

$\log m$ bits to store r ($1 \leq r \leq m$)

That is $\log m + \log n$ bits together.

This is what we wanted to prove for (*2) part of Theorem 1

Example

Example: $A = [a_1, a_2, \dots, a_m]$
 $= [1, 2, 1, 1, 3, 2, 1, 1]$

$$m_1 = 5$$



$$m_2 = 2$$



$$F_k = 5^k + 2^k + 1^k$$

$$m_3 = 1$$

$$\text{If } p = 1 \Rightarrow a_1 = 1 \Rightarrow r = |1, 3, 4, 7, 8| = 5$$

$$\text{If } p = 2 \Rightarrow a_2 = 2 \Rightarrow r = |2, 6| = 2$$

$$\text{If } p = 3 \Rightarrow a_3 = 1 \Rightarrow r = |3, 4, 7, 8| = 4$$

$$\text{If } p = 4 \Rightarrow a_4 = 1 \Rightarrow r = |4, 7, 8| = 3$$

$$\text{If } p = 5 \Rightarrow a_5 = 3 \Rightarrow r = |5| = 1$$

$$\text{If } p = 6 \Rightarrow a_6 = 2 \Rightarrow r = |6| = 1$$

$$\text{If } p = 7 \Rightarrow a_7 = 1 \Rightarrow r = |7, 8| = 2$$

$$\text{If } p = 8 \Rightarrow a_8 = 1 \Rightarrow r = |8| = 1$$

Proof of Theorem 1

Example (Continued)

We will calculate $\mathbb{E}[X]$ and $\text{Var}(X)$.

By definition, $X = \underbrace{m(r^k - (r-1)^k)}_{f(r)}$

Therefore,

$$\begin{aligned}\mathbb{E}[X] &= \frac{1}{8} [f(5) + f(2) + f(4) + f(3) + f(1) + f(1) + f(2) + f(1)] \\ &= \frac{1}{8} \begin{pmatrix} f(1) + f(2) + f(3) + f(4) + f(5) \\ + f(1) + f(2) \\ + f(1) \end{pmatrix} \quad \leftarrow \text{n rows}\end{aligned}$$

m elements

$$p \sim U[\{1, 2, \dots, 8\}]$$

Expected Value of X

More generally,

$$\mathbb{E}[X] = \frac{1}{m} \begin{pmatrix} f(1) + f(2) + \dots + f(m_1) \\ + f(1) + f(2) + \dots + f(m_2) \\ \vdots \\ + f(1) + f(2) + \dots + f(m_i) \\ \vdots \\ + f(1) + f(2) + \dots + f(m_n) \end{pmatrix} \quad \leftarrow \text{n rows}$$

$p \sim U[\{1, 2, \dots, m\}]$ m elements

$$= \frac{m}{m} \begin{pmatrix} 1^k + (2^k - 1^k) + \dots + (m_1^k - (m_1 - 1)^k) \\ + 1^k + (2^k - 1^k) + \dots + (m_2^k - (m_2 - 1)^k) \\ \vdots \\ + 1^k + (2^k - 1^k) + \dots + (m_i^k - (m_i - 1)^k) \\ \vdots \\ + 1^k + (2^k - 1^k) + \dots + (m_n^k - (m_n - 1)^k) \end{pmatrix} = \sum_{i=1}^n m_i^k = F_k$$

We have proved that X is unbiased estimator of F_k , i.e. $\mathbb{E}[X] = F_k$

Preliminaries

Lemma 1 [Bound on $a^k - b^k$]

If $0 \leq b \leq a$, then $a^k - b^k \leq (a - b)ka^{k-1}$

Proof of Lemma 1

$$\begin{aligned} a^k - b^k &= (a - b)(a^{k-1} + a^{k-2}b + \dots + ab^{k-2} + b^{k-1}) \\ &\leq (a - b)(ka^{k-1}) \end{aligned}$$



Since $0 \leq b \leq a$

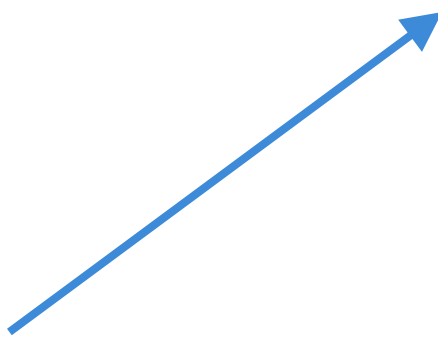
This finishes the proof of Lemma 1. Q.E.D.

Lemma 2

Lemma 2

If $1 \leq a$, then $(a^k - (a - 1)^k)^2 \leq ka^{2k-1} - k(a - 1)^{2k-1}$

Proof

$$\begin{aligned}(a^k - (a - 1)^k)^2 &= (a^k - (a - 1)^k)(a^k - (a - 1)^k) \\ &\leq ka^{k-1}(a^k - (a - 1)^k) \\ &= ka^{2k-1} - ka^{k-1}(a - 1)^k \\ &\leq ka^{2k-1} - k(a - 1)^{2k-1}\end{aligned}$$


From Lemma 1 [If $0 \leq b \leq a$, $\Rightarrow a^k - b^k \leq (a - b)ka^{k-1}$]

This finishes the proof of Lemma 2. Q.E.D

Variance of X

Let us calculate and bound the variance of X .

$$\text{Var}(X) = \mathbb{E}[X^2] - \underbrace{\mathbb{E}^2[X]}_{F_k^2}$$

Lemma 3 [Bound on $\mathbb{E}[X^2]$] $\mathbb{E}[X^2] \leq kmF_{2k-1}$
 $= kF_1F_{2k-1}$

Proof of Lemma 3

By definition,

$$X = \underbrace{m(r^k - (r-1)^k)}_{f(r)}$$

Therefore,

$$X^2 = \underbrace{m^2(r^k - (r-1)^k)^2}_{f^2(r)}$$

Expected Value of X^2

We know

$$X^2 = \underbrace{m^2(r^k - (r-1)^k)^2}_{f^2(r)}$$

Therefore,

$$\begin{aligned}\mathbb{E}[X^2] &= \frac{1}{m} \begin{pmatrix} f^2(1) + f^2(2) + \dots + f^2(m_1) \\ + f^2(1) + f^2(2) + \dots + f^2(m_2) \\ \vdots \\ + f^2(1) + f^2(2) + \dots + f^2(m_i) \\ \vdots \\ + f^2(1) + f^2(2) + \dots + f^2(m_n) \end{pmatrix} \\ &= \frac{m^2}{m} \begin{pmatrix} 1^{2k} + (2^k - 1^k)^2 + \dots + (m_1^k - (m_1 - 1)^k)^2 \\ + 1^{2k} + (2^k - 1^k)^2 + \dots + (m_2^k - (m_2 - 1)^k)^2 \\ \vdots \\ + 1^{2k} + (2^k - 1^k)^2 + \dots + (m_i^k - (m_i - 1)^k)^2 \\ \vdots \\ + 1^{2k} + (2^k - 1^k)^2 + \dots + (m_n^k - (m_n - 1)^k)^2 \end{pmatrix}\end{aligned}$$

Expected Value of X^2

Therefore from Lemma 2

$$[\text{If } 1 \leq a, \text{ then } (a^k - (a - 1)^k)^2 \leq ka^{2k-1} - k(a - 1)^{2k-1}]$$

$$\begin{aligned} \mathbb{E}[X^2] &= m \left(\underbrace{1^{2k}}_{\leq k1^{2k-1}} + \underbrace{(2^k - 1^k)^2}_{\leq k2^{2k-1} - k1^{2k-1}} + \dots + \underbrace{(m_1^k - (m_1 - 1)^k)^2}_{\leq km_1^{2k-1} - k(m_1 - 1)^{2k-1}} \right. \\ &\quad \vdots \\ &\quad \underbrace{1^{2k}}_{\leq k1^{2k-1}} + \underbrace{(2^k - 1^k)^2}_{\leq k2^{2k-1} - k1^{2k-1}} + \dots + \underbrace{(m_i^k - (m_i - 1)^k)^2}_{\leq km_i^{2k-1} - k(m_i - 1)^{2k-1}} \\ &\quad \vdots \\ &\quad \left. \underbrace{1^{2k}}_{\leq k1^{2k-1}} + \underbrace{(2^k - 1^k)^2}_{\leq k2^{2k-1} - k1^{2k-1}} + \dots + \underbrace{(m_n^k - (m_n - 1)^k)^2}_{\leq km_n^{2k-1} - k(m_n - 1)^{2k-1}} \right) \\ &\leq m(km_1^{2k-1} + \dots + km_n^{2k-1}) \\ &= mkF_{2k-1} \\ &= kF_1F_{2k-1} \quad \text{This completes the proof of Lemma 3. Q.E.D} \end{aligned}$$

Power Mean Inequality

Lemma 4

$$\begin{array}{l} \text{If } m_1 \geq 0 \\ m_2 \geq 0 \\ \vdots \\ m_n \geq 0 \\ k \geq 1 \end{array} \quad \Rightarrow \quad \sum_{i=1}^n m_i \leq n \left(\sum_{i=1}^n m_i^k / n \right)^{1/k}$$

Proof of Lemma 4

According to the power-mean inequality,

$$\text{If } k \geq 1, \text{ then } \sum_{i=1}^n \frac{1}{n} m_i \leq \left(\sum_{i=1}^n \frac{1}{n} m_i^k \right)^{1/k}$$

See, e.g. Cauchy-Schwarz master class, Problem 8.3

This finishes the proof of Lemma 4. Q.E.D

Variance of X

By Lemma 3, we have that $\mathbb{E}[X^2] \leq kmF_{2k-1} = kF_1F_{2k-1}$

Let us bound the r.h.s

Lemma 5 (Bound on $\mathbb{E}[X^2]$)

$$\begin{array}{l} \text{If } m_1 \geq 0 \\ m_2 \geq 0 \\ \vdots \\ m_n \geq 0 \\ k \geq 1 \end{array} \quad \Rightarrow \quad F_1F_{2k-1} \leq n^{1-1/k} F_k^2$$

Proof of Lemma 5

We need to prove:

$$\left(\sum_{i=1}^n m_i \right) \left(\sum_{i=1}^n m_i^{2k-1} \right) \leq n^{1-1/k} \left(\sum_{i=1}^n m_i^k \right)^2$$
$$F_1 \qquad F_{2k-1} \qquad F_k^2$$

Proof of Lemma 5 (Continued)

Observation 1

Let $M \doteq \max_{1 \leq i \leq n} m_i$

$$\begin{aligned} \text{Since } M^k &= \max_{1 \leq i \leq n} m_i^k \\ &\leq \sum_{1 \leq i \leq n} m_i^k \end{aligned}$$

therefore, $M \leq \left(\sum_{i=1}^n m_i^k \right)^{1/k}$

Proof of Lemma 5 (Continued)

We need to prove: $\left(\sum_{i=1}^n m_i\right) \left(\sum_{i=1}^n m_i^{2k-1}\right) \leq n^{1-1/k} \left(\sum_{i=1}^n m_i^k\right)^2$

$$\left(\sum_{i=1}^n m_i\right) \left(\sum_{i=1}^n m_i^{2k-1}\right) \leq \left(\sum_{i=1}^n m_i\right) \left(\sum_{i=1}^n M^{k-1} m_i^k\right)$$

Since $M \doteq \max_{1 \leq i \leq n} m_i$

$$\begin{aligned} &= \left(\sum_{i=1}^n m_i\right) \underbrace{M^{k-1}}_{\leq \left(\sum_{i=1}^n m_i^k\right)^{\frac{k-1}{k}}} \left(\sum_{i=1}^n m_i^k\right) \\ &\leq \left(\sum_{i=1}^n m_i\right) \left(\sum_{i=1}^n m_i^k\right)^{\frac{k-1}{k}} \end{aligned}$$

Since $M \leq \left(\sum_{i=1}^n m_i^k\right)^{1/k}$ by Observation 1

$$\leq \left(\sum_{i=1}^n m_i\right) \left(\sum_{i=1}^n m_i^k\right)^{\frac{k-1}{k}} \left(\sum_{i=1}^n m_i^k\right)$$

Proof of Lemma 5 (Continued)

This is what we know so far:

$$\begin{aligned} \left(\sum_{i=1}^n m_i \right) \left(\sum_{i=1}^n m_i^{2k-1} \right) &\leq \left(\sum_{i=1}^n m_i \right) \left(\sum_{i=1}^n m_i^k \right)^{\frac{k-1}{k}} \left(\sum_{i=1}^n m_i^k \right) \\ &= \left(\sum_{i=1}^n m_i \right) \underbrace{\left(\sum_{i=1}^n m_i^k \right)^{\frac{2k-1}{k}}}_{\leq n \left(\sum_{i=1}^n m_i^k / n \right)^{1/k}} \end{aligned}$$

By Lemma 4

$$\begin{aligned} \sum_{i=1}^n m_i &\leq n \left(\sum_{i=1}^n m_i^k / n \right)^{1/k} \longrightarrow \leq n \left(\sum_{i=1}^n m_i^k / n \right)^{1/k} \\ &\leq n \left(\sum_{i=1}^n m_i^k / n \right)^{1/k} \left(\sum_{i=1}^n m_i^k \right)^{\frac{2k-1}{k}} \\ &= n^{1-1/k} \left(\sum_{i=1}^n m_i^k \right)^{1/k} \left(\sum_{i=1}^n m_i^k \right)^{\frac{2k-1}{k}} \\ &= n^{1-1/k} \left(\sum_{i=1}^n m_i^k \right)^2 \end{aligned}$$

This is what we had to prove for Lemma 5. Q.E.D

Summary

So far we know

$$\begin{aligned}\mathbb{E}[X_{ij}] &= F_k & \forall i, j & \quad 1 \leq j \leq S_1 \\ \mathbb{E}[X_{ij}^2] &\leq kF_1F_{2k-1} \leq kn^{1-1/k}F_k^2 & & \quad 1 \leq i \leq S_2\end{aligned}$$

By Definition,

$$\begin{pmatrix} X_{1,1} & \dots & X_{1,S_1} \\ \vdots & & \vdots \\ X_{i,1} & \dots & X_{i,S_1} \\ \vdots & & \vdots \\ X_{S_2,1} & \dots & X_{S_2,S_1} \end{pmatrix} \Rightarrow Y_i \doteq \frac{1}{S_1} (X_{i,1} + \dots + X_{i,S_1})$$

Therefore,

$$\mathbb{E}[Y_i] = \frac{1}{S_1} \sum_{j=1}^{S_1} \mathbb{E}[X_{i,j}] = F_k$$

That is, Y_i is unbiased estimator of F_k .

Variance of Y_i

Let us bound the variance of Y_i

Lemma 6

$$\text{Var}[Y_i] \leq \frac{1}{S_1} \mathbb{E}[X^2]$$

Proof

$$\begin{aligned} \text{Var}[Y_i] &= \text{Var} \left[\frac{1}{S_1} \sum_{j=1}^{S_1} X_{ij} \right] \\ &= \frac{1}{S_1^2} \sum_{j=1}^{S_1} \underbrace{\text{Var}[X_{ij}]}_{\mathbb{E}[X^2] - F_k^2} \quad \text{Since } \mathbb{E}[X] = F_k \\ &= \frac{1}{S_1} (\mathbb{E}[X^2] - F_k^2) \\ &\leq \frac{1}{S_1} \mathbb{E}[X^2] \end{aligned}$$

This is what we had to prove for Lemma 6. Q.E.D

Variance of Y_i

Lemma 7

$$\text{Var } [Y_i] \leq \frac{k}{S_1} n^{1-1/k} F_k^2$$

Proof of Lemma 7

$$\text{Var } [Y_i] \leq \frac{1}{S_1} \mathbb{E}[X^2]$$

$$\leq \frac{1}{S_1} k F_1 F_{2k-1}$$

$$\leq \frac{k}{S_1} n^{1-1/k} F_k^2$$

By Lemma 6

By Lemma 3: $\mathbb{E}[X^2] \leq k F_1 F_{2k-1}$

By Lemma 5: $F_1 F_{2k-1} \leq n^{1-1/k} F_k^2$

This completes the proof of Lemma 7. Q.E.D

Concentration Bound on Y_i

Chebyshev's inequality:

If $a > 0$, then $Pr(|Y - \mathbb{E}[Y]| > a) \leq \frac{\text{Var}[Y]}{a^2}$

Lemma 8

Let i be fixed ($1 \leq i \leq S_2$). We have that $Pr(|Y_i - F_k| > \lambda F_k) \leq \frac{1}{12}$

Proof of Lemma 8 From Chebyshev's inequality, we have that

For all fixed i :

$$\begin{aligned} Pr(|Y_i - \underbrace{\mathbb{E}[Y]}_{F_k}| > \underbrace{\lambda F_k}_a) &\leq \frac{\text{Var}[Y_i]}{\lambda^2 F_k^2} \\ &\leq \frac{\frac{k}{S_1} n^{1-1/k} F_k^2}{\lambda^2 F_k^2} \quad \leftarrow \begin{array}{l} \text{Since } \text{Var}[Y_i] \leq \frac{k}{S_1} n^{1-1/k} F_k^2 \\ \text{by Lemma 7} \end{array} \\ &= \frac{kn^{1-1/k}}{\lambda^2 S_1} \quad \leftarrow \begin{array}{l} \text{Since } S_1 \doteq \frac{12kn^{1-1/k}}{\lambda^2} \\ \text{This completes the} \\ \text{proof of Lemma 8.} \\ \text{Q.E.D} \end{array} \\ &\leq \frac{1}{12} \end{aligned}$$

Concentration Bound on Y

We know from Lemma 8 that

For all $1 \leq i \leq S_2$, we have that $Pr(|Y_i - F_k| > \lambda F_k) \leq \frac{1}{12}$

This is a bound on the probability that (for a fixed i) Y_i deviates from $\mathbb{E}[Y_i] = F_k$ more than λF_k

Now let us bound the probability that at least $S_2/2$ among $\{Y_1, Y_2, \dots, Y_{S_2}\}$ deviates from F_k more than λF_k

We will prove that this probability is $\leq \epsilon$

Since $Y \doteq \text{median}(Y_1, Y_2, \dots, Y_{S_2})$, this will imply that

$$Pr\left(|Y - F_k| > \lambda F_k\right) \leq \epsilon$$

This will complete the proof of Theorem 1

Chernoff Bound

Lemma [Chernoff, 1952, Tail bound for binomial distribution]

Let B be a binomial random variable $B \sim B(n, p)$

Let $0 < p < \delta < 1$

Then

$$Pr(B > n\delta) \leq \exp \left(-n \underbrace{\left[\delta \log \frac{\delta}{p} + (1 - \delta) \log \frac{1 - \delta}{1 - p} \right]}_{\geq 0} \right)$$

Proof: [Gyorfi et al, Distribution Free Theory of Nonparametric Regression, Page 592]

Application of Chernoff Bound

Let $Z_i \doteq 1_{\{|Y_i - F_k| \geq \lambda F_k\}} \in \{0, 1\}$

We already know from Lemma 8 that

$$\mathbb{E}[Z_i] = \mathbb{P}(|Y_i - F_k| > \lambda F_k) \leq \frac{1}{12}$$

Observation:

$$\sum_{i=1}^{S_2} Z_i \sim \text{Binomial}(S_2, p) \quad \text{where } p = \mathbb{E}[Z_i] \leq \frac{1}{12}$$


To see that (*1) part of Theorem 1 holds, we need to prove that

$$\Pr \left(\sum_{i=1}^{S_2} Z_i > \frac{S_2}{2} \right) \leq \epsilon$$


Application of Chernoff Bound

From Chernoff $Pr(B > n\delta) \leq \exp \left(-n \left[\delta \log \frac{\delta}{p} + (1 - \delta) \log \frac{1 - \delta}{1 - p} \right] \right)$


Let $B \doteq \sum_{i=1}^{S_2} Z_i$, $n \doteq S_2$, $\delta \doteq 1/2$ in the Chernoff bound


$$Pr \left(\sum_{i=1}^{S_2} Z_i > \frac{S_2}{2} \right) \leq \exp \left(-S_2 \left[\frac{1}{2} \log \frac{1}{2p} + \frac{1}{2} \log \frac{1}{2(1-p)} \right] \right)$$

Since $S_2 = 2 \log(1/\epsilon)$
 $-S_2 = 2 \log(\epsilon)$


$$\begin{aligned} &= \exp \left((\log \epsilon) \left[\log \frac{1}{2p} + \log \frac{1}{2(1-p)} \right] \right) \\ &= \exp \left((\log \epsilon) \left[\log \frac{1}{4p(1-p)} \right] \right) \end{aligned}$$

If $\log \frac{1}{4p(1-p)} > 1$
i.e $\frac{1}{4p(1-p)} > e$
i.e $1 > e4p(1-p)$


$$\begin{aligned} &\leq \exp(\log \epsilon) \\ &= \epsilon \end{aligned}$$

This holds if $p < \frac{1}{12}$. **This completes the proof of Theorem 1 Q.E.D.**

Thanks for your Attention! 😊

Another Surprising Fact

We know that

F_k can be estimated randomly with using only

$$O\left(n^{1-1/k}(\log n + \log m)\right) \quad \text{bits}$$

Theorem:

F_2 can be estimated randomly with using only

$$O\left((\log n + \log m)\right) \quad \text{bits}$$