reading group summer 24 Location Sorting and Endogenous Amenities: Evidence from Amsterdam (2024)

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#### Outline

#### Introduction

Data

Model - brace yourselves

Empirical Strategy and Results

Counterfactuals

#### Introduction

- build a model of residential choice with hetereogenous households amenities improve in response to location sorting
- but nature of amenities has not been explored
- estimate the model using several data
- run counterfactual exercises

#### Context

- massive expansion in tourism in Amsterdam
- increased supply of private rentals, increased supply of STRs
- New regulation in Amsterdam severely restricting STR supply (hotels and Airbnbs)

#### Related Literature

- spatial equilirium models
- effects of STR entry on the housing market and hotel revenue
- discrete-choice tools from the empirical io literature applied to urban residential markets

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#### Data

- individual-level microdata: Centraal Bureau voor de Statistiek, Netherlands they have complete residential histories, demographics
- housing data: panel on physical characteristics, occupancies, and values. impute rents
- neighbourhood data: Amsterdam City Data consists of demographics, tourist flows, and consumption amenities: restaurants, bars, food stores, non-food stores, nurseries, and "touristic amenities"
- STR data: Inside Airbnb. monthly data on listings, geo-coordinates, prices, reviews. need to identify preoperties that are permanently rented to tourists.

# Fact 1: Tourists and STR listings have grown dramatically and sprawled across Amsterdam

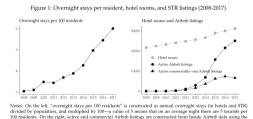


Figure 1:

procedure described in Appendix A.2.7. Hotel, stay and population data are from ACD Tourism and ACD BBGA.

Fact 2: Rents have increased more in neighborhoods with more STR entry using OLS (likely biased) and IV (shift-share)

	Ln (rent/m2)					
	OLS	IV	OLS	IV	OLS	IV
Ln (commercial Airbnb listings)	(0.008)	(0.021)	(0.006)	(0.021)	(0.018)	0.205* (0.093)
Control variables District-year FE			Х	Х	X	X
First stage F-stat		586.89		384.21		69.66
Observations	770	770	763	763	763	763
			Ln (house	sale price)		
	OLS	IV	OLS	IV	OLS	IV
Ln (commercial Airbnb listings)	0.109*** (0.016)	(0.030)	(0.006)	(0.016)	(0.022)	0.326** (0.102)
Control variables District-year FE			Х	Х	X	X
First stage F-stat		572.02		370.87		65.9
Observations	738	738	737	737	737	737

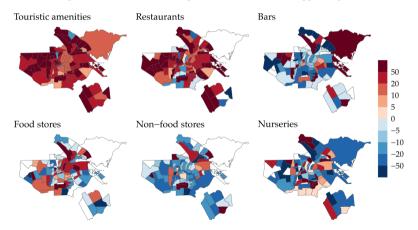
Commercial Airbob listings are constructed from the Inside Airbob data (see Appendix A.2.7 for construction details) Neighborhood-level control variables are: housing stock, average income, high-skill population share, all from ACD BBCA. Standard errors are clustered at the wilk level in parenthesis.

Figure 2:

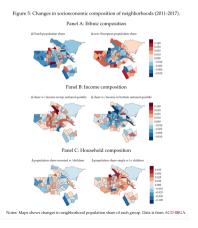
a 1% increase in a neighborhood's commercial STR listings is associated with a rent increase between .06-.11% and a house sale price increase between .04-.11%

#### Fact 3: Amenities have tilted towards tourists and away from locals

Figure 3: Evolution of consumption amenities (2011-2017 pp changes).



Fact 4: The composition of residents has changed heterogeneously across neighborhoods



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#### Notation

- J locations +1 outside option
- ullet  $M^k_{jt}$ : number of type k households in location j
- ullet consumption amenities s in S sectors
- $N_{sjt}$ : the number of varieties in sector s and location j at time t
- amenities  $a_{jt} = [N_{1jt}, N_{2jt}, ..., N_{Sjt}]'$

#### **Amenities Demand**

- ullet Demand: Cobb-Douglas over bundles H and C
- also Cobb-Douglas preferences over amenity sectors, CES over firms/varieties within sector

$$q_{isjt}^k = rac{lpha_s^k \phi^k w_t^k}{p_{isjt}} \left(rac{p_{isjt}}{P_{sjt}}
ight)^{1-\sigma_s}$$
, with  $P_{sjt} \equiv \left(\sum_{i=1}^{N_{sjt}} p_{isjt}^{1-\sigma_s}
ight)^{rac{1}{1-\sigma_s}} q_{isjt} = \sum_k q_{isjt}^k M_{jt}^k$ .

Figure 5:

## **Amenities Supply**

- ullet Firms within s,j in monopolistic competition same marginal cost, free entry
- fixed cost  $F_{sjt}$  assumed to be increasing in  $N_{jt}$  (congestion costs)
- ullet implies that firms choose same price  $p_{sjt}=p_{ijst}$  and quantity  $q_{sjt}=q_{isjt}$

$$(p_{sjt}-c_{sjt})q_{sjt}=F_{sjt}(N_{jt}), \quad \text{where } N_{jt}=\sum_{s}N_{sjt}.$$

Figure 6:

# **Amenities Market Clearing**

- usual market clearing delivers equilibrium number of varieties/firm
- and a mapping from population composition to amenities

# **Housing Supply**

- total housing stock inelastic in the short run
- Landlords face a binary choice: long-term rentals or short-term rentals:

$$\max \{\alpha r_{jt} + \epsilon_{LT}, \quad \alpha p_{jt} - \kappa_{jt} + \epsilon_{ST} \},$$

Figure 7:

# **Housing Supply**

- total housing stock inelastic in the short run
- Landlords face a binary choice: long-term rentals or short-term rentals:

$$\mathcal{H}_{jt}^{LT,S}(r_{jt},p_{jt}) = \frac{\exp(\alpha r_{jt})}{\exp(\alpha r_{jt}) + \exp(\alpha p_{jt} - \kappa_{jt})} \mathcal{H}_{jt},$$
 $\mathcal{H}_{jt}^{ST,S}(r_{jt},p_{jt}) = \mathcal{H}_{jt} - \mathcal{H}_{jt}^{LT,S}(r_{jt},p_{jt}).$ 

Figure 8:

moving costs and location capital

$$MC^{k}(j_{ii}, j_{it-1}) = \begin{cases} 0 & \text{if } j_{ii} = j_{it-1} \\ m_{0}^{k} + m_{1}^{k} \text{dist}(j_{ii}, j_{it-1}) & \text{if } j_{ii} \neq j_{it-1} \text{ and } j_{ii}, j_{it-1} \neq 0 \\ m_{2}^{k} & \text{if } j_{ii} \neq j_{it-1}, \text{ and } j_{it} = 0 \text{ or } j_{it-1} = 0, \end{cases}$$

Figure 9:

$$au_{it} = egin{cases} \min\{ au_{it-1} + 1, ar{ au}\} & \quad ext{if } j_{it} = j_{it-1} \\ 1 & \quad ext{otherwise}. \end{cases}$$

Figure 10:

some macro...

$$V^k_t(x_{it}, \epsilon_{it}) = \max_{j \in \{0, 1, ..., I\}} u^k_t(j, x_{it}) + \epsilon_{ijt} + \beta \mathbb{E}_t \Bigg[ V^k_{t+1}(x_{it+1}, \epsilon_{it+1}) | j, x_{it}, \epsilon_{it} \Bigg].$$

Figure 11:

• choice probabilities

$$\mathbb{P}_{t}^{k}(j|x_{it}) = \frac{\exp\left(u_{t}^{k}(j,x_{it}) + \beta \mathbb{E}_{t} \left[V_{t+1}^{k}(x_{it+1},\epsilon_{it+1})|j,x_{it},\epsilon_{it}\right]\right)}{\sum_{j'} \exp\left(u_{t}^{k}(j',x_{it}) + \beta \mathbb{E}_{t} \left[V_{t+1}^{k}(x_{it+1},\epsilon_{it+1})|j',x_{it},\epsilon_{it}\right]\right)}.$$

Figure 12:

transition matrix

$$\pi_t^k(j,\tau) = \begin{cases} \sum_{\tau'} \sum_{j' \neq j} \mathbb{P}_t^k(j|j',\tau') \pi_{t-1}^k(j',\tau') & \tau = 1 \\ \mathbb{P}_t^k(j|j,\tau-1) \pi_{t-1}^k(j,\tau-1) & \tau \in [2,\bar{\tau}) \\ \mathbb{P}_t^k(j|j,\bar{\tau}-1) \pi_{t-1}^k(j,\bar{\tau}-1) + \mathbb{P}_t^k(j|j,\bar{\tau}) \pi_{t-1}^k(j,\bar{\tau}) & \tau = \bar{\tau}. \end{cases}$$

Figure 13:

• (population) demand for STRs

$$\mathcal{H}_{jt}^{LT,D}(r_t,a_t) = \sum_{k=1}^K M_{jt}^k(r_t,a_t) f_{jt}^k.$$

Figure 14:

$$\mathcal{H}_{jt}^{LT,D}(r_t, a_t) = \sum_{k=1}^{K} M_{jt}^k(r_t, a_t) f_{jt}^k.$$

Figure 15:

### Housing Demand: tourists

• toursists' flow payoff from staying in STR

$$u_{jt}^{ST} = \delta_j^{ST} + \delta_t^{ST} + \delta_p^{ST} \log p_{jt} + \delta_a^{ST} \log a_{jt} + \xi_{jt}^{ST},$$

Figure 16:

# Housing Demand: tourists

• toursists' demand for STR

Figure 17:

# Housing Demand: tourists

- hotels are an outside option
- but no price or bookings data
- so use data on hotel capacity across locations

$$M_{jt}^H(p_t, a_t) = s_{jt}^{beds} \times M_t^H(p_t, a_t)$$

Figure 18:

# Housing Equilibrium

- a stationary equilibrium
- a bunch of market clearing equations
- long-term rental markets clear, short-term rental markets clear, amenities markets clear

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### Amenity supply

estimating equation:

$$\log N_{sjt} = \lambda_j + \lambda_t - \eta \log N_{jt} + \log \left(\sum_k \beta_s^k X_{jt}^k\right) + \omega_{sjt},$$

#### Figure 19:

- $X_{jt}^k$  total expenditure of the type k population in location j on consumption amenities
- ullet  $eta_s^k$  describes how this expenditure is allocated to each amenity sector s
- IV: something that shifts amenities demand

## **Amenity Supply**

- ullet IV:  $Z_{jt}^k=w_t^kS_{jt}^{\gamma(k)}$  , where S is housing stock by tenancy status
- $\bullet$  GMM moment equation to idenfity  $\beta \text{s: } E[\omega_{sjt}Z^k_{sit}] = 0$

### **Amenity Supply**

#### results:

Table 3: Estimates of amenity supply parameters.

	Touristic Amenities	Restaurants	Bars	Food Stores	Non-Food Stores	Nurseries
Older Families	186.3	7.374	0.0	4.469	11.359	980.803***
	[0.0,431.929]	[0.0,30.401]	[0.0,0.0]	[0.0,29.098]	[0.0,50.577]	[368.357,1684.881]
Singles	403.022	90.723	0.0	102.631	11.347	0.0
	[0.0,1816.258]	[0.0,317.227]	[0.0,0.0]	[0.0,327.09]	[0.0,176.969]	[0.0,0.0]
Younger Families	0.0	1.077	11.365	52.846**	194.655***	637.116
	[0.0,0.0]	[0.0, 16.47]	[0.0,38.983]	[0.0,123.07]	[83.304,331.513]	[0.0,1326.254]
Students	984.639*	402.153***	22.562	123.078	2.355	221.204
	[0.0,2014.696]	[198.125,673.928]	[0.0,101.561]	[0.0,319.632]	[0.0,1.365]	[0.0,1759.481]
Immigrant Families	0.122	5.687	25.705**	90.549	127.724**	540.228
	[0.0,0.0]	[0.0,49.542]	[1.126,63.178]	[0.0,210.37]	[3.251,331.895]	[0.0,1740.204]
<b>Dutch Low Income</b>	110.617	9.908	0.0	9.077	0.0	0.0
	[0.0,371.65]	[0.0,56.277]	[0.0,0.0]	[0.0,78.283]	[0.0,0.0]	[0.0,0.0]
Tourists	749.072***	397.274***	211.571***	136.337***	724.223***	0.0
	[522.649,974.412]	[316.498,477.965][	156.406,269.374	][80.554,189.6]	[579.264,892.963]	[0.0,0.0]

Notes: This table reports bootstrap results for coefficients  $\beta_s^k$  from Equation 24 for using a three-way panel of 22 districts in Amsterdam for 2008-2018 over 500 draws. Parameters  $\beta_s^k$  and fixed effects  $\lambda_j$  and  $\lambda_t$  are estimated via GMM, where we restrict parameters to be weakly positive as implied by the microfoundation of the amenity model in Appendix A.3.1. The estimation procedure is outlined in section 5.2 following a Bayesian-bootstrap with random Dirichlet weights. Total expenditure  $X_j^k$  is measured in thousands of Euros. Top rows indicate average estimates of the bootstrap samples. Results inside square brackets indicate 95% confidence intervals. We omit estimates of the location and time fixed effects. \*p < 0.10, \*\*p < 0.05, \*\*\*p < 0.05, \*\*\*p < 0.01, \*\*p < 0.05, \*\*\*p < 0.01, \*\*p <

### Housing Demand: by locals

• "Euler Equations in Conditional Choice Probabilities" (ECCP) estimator: which I didn't understand and didn't have time to study in detail

	Dependent variable: Relative likelihood of renewal paths			
	Older Families	Singles	Younger Families	
High Location Capital	0.187***	0.210***	0.264***	
	(0.017)	(0.013)	(0.014)	
Intra-City Moving Cost	-5.916***	-5.337***	-5.384***	
	(0.015)	(0.011)	(0.012)	
Bilateral Moving Cost	-0.067***	-0.059***	-0.041***	
	(0.000)	(0.000)	(0.000)	
In/Out of City Moving Cost	-4.407***	-4.012***	-4.043***	
	(0.012)	(0.009)	(0.010)	
Log Rent	-10.886***	-2.310**	1.964*	
	(1.205)	(0.999)	(1.027)	
Log Touristic Amerities	-1.319***	-0.496***	0.317*	
	(0.215)	(0.182)	(0.177)	
Log Restaurants	0.288	0.735**	-0.280	
	(0.346)	(0.305)	(0.286)	
Log Bars	-0.757***	-0.526***	-0.104	
	(0.099)	(0.065)	(0.086)	
Log Food Stores	-1.695***	-1.216***	-0.540*	
	(0.327)	(0.281)	(0.282)	
Log Nonfood Stores	0.427	1.533***	1.383***	
	(0.356)	(0.311)	(0.302)	
Log Nurseries	1.631***	0.044	0.246*	
	(0.173)	(0.143)	(0.147)	
N	233772	233772	233772	

Notes: This shell presents regression results of preference parameters for a dynamic knoise their smells for 2 date in the anthonistics for MNN = 1000 Mer similar declaration of a similar shell of MNN = 1000 Mer shell of

Figure 21:

interpret signs

### Housing Demand: by tourists

 no IV here: they use reveiews of Airbnbs to create a score variable for every location

Table 5: Tourist demand across locations.

	Dependent Variable: $\log \mathbb{P}^{ST}_{jt} - \log \mathbb{P}^{H}_{t}$					
	Basel	ine	Controlling for reviews			
Log Price Per Guest	-2.725***	(0.820)	-2.660***	(0.759)		
Log Touristic Amenities	1.009***	(0.376)	0.838**	(0.394)		
Log Restaurants	0.048	(0.259)	0.017	(0.243)		
Log Bars	0.051	(0.155)	0.056	(0.164)		
Log Food Stores	-0.001	(0.300)	0.037	(0.323)		
Log Nonfood Stores	-0.228	(0.417)	-0.185	(0.407)		
Log Nurseries	-0.234*	(0.137)	-0.231*	(0.136)		
Log Review Scores			4.768	(3.699)		
N	371.000		370.000			
$\mathbb{R}^2$	0.529		0.537			

Notes: Table reports estimates of tourists' preference for neighborhood (wijk-level) characteristics for a static model of location choice, using neighborhood-level data for 2015-2018. Construction of Airbnb supply and prices is described in Appendix A.2. Wijk-level clustered standard errors in parenthesis. \*p < 0.10; \*\*\*\*p < 0.05; \*\*\*\*\*\*p < 0.01:

Figure 22:

# Housing supply

estimating equation:

$$\log \mathcal{H}_{jt}^{LT,S} - \log \mathcal{H}_{jt}^{ST,S} = \alpha \left( r_{jt} - p_{jt} \right) + \kappa_j + \kappa_t + \nu_{jt},$$

#### Figure 23:

- OLS is biased use shift-share IV that shifts demand
- shift: worldwide change in STR demand
- share: neighborhood-level exposure to the shift from the historic spatial distribution of touristic attractions
- relevance and exclusion are satisifed. (note: not mention of monotonicty, CALL IVAN)

# Housing supply

 IV results: increase in the gap between STR prices and LTR prices of one standard deviation (29%) would raise the market share of the ST relative to the LT segment by 13.6%

Table 6: Long-term (LT) relative to short-term (ST) housing supply elasticities

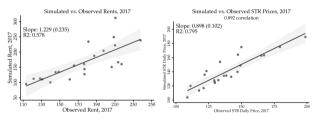
	Dependent variable: ln (LT share) - ln (ST share)				
	OLS	IV	IV	IV	
LT price-ST price	0.242*	0.287**	0.309**	0.385	
	(0.099)	(0.086)	(0.091)	(0.639)	
Year FE			Х	X	
Wijk FE				X	
First stage F-stat		65.68	61.62	3.24	
Observations	275	275	275	275	

Notes: Table reports estimates of landlords' marginal utility of income for a discrete choice model between the shortand long-term rental markets. Data are a panel with 92 locations 2015-2017. Prices are instrumented using a shift-share instrument (Barron et al., 2021) that proxies for demand shocks. Wijk-level clustered standard errors in parenthesis. \*p < 0.10, \*p < 0.00, \*p < 0.00, \*p < 0.00.

#### How does the model do?

pretty well!

Figure 7: Model fit: Rents and STR prices



Notes: The figure presents scatter plots, linear fit, and 95% confidence intervals of simulated rents and STR prices, against observed rents and prices for 22 districts. Rents are in  $Euros/m^2$  per year. STR prices are average daily prices.

Figure 25:

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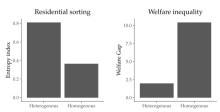
Empirical Strategy and Results

Counterfactuals

### Counterfactual 1: Preference heterogeneity vs homogeneity

- set preference parameters for consumption amenities to the average value across all household types, weighted by the size of groups
- segregation is higher when households have heterogeneous preferences for amenities

Figure 9: Role of preference heterogeneity for spatial sorting and inequality across households.



Notes: The left panel reports the entropy index, a measure of spatial segregation of household types: higher values indicate more segregation (see Appendix A.5.6 for a formal definition). The right panel reports the ratio of the highest consumer surpuls household (in Euros) to that of the lowest household: higher values indicate more inequality.

#### Figure 26:

... but inequality is lower when preferences are heterogeneous

### Counterfactual 1: Preference heterogeneity vs homogeneity

#### • ... but inequality is lower when preferences are heterogeneous

Table 7: Neighborhood differentiation as spatial dispersion of amenities.

	Gini index for each p		
Amenity	Homogenous (HO)	Heterogenous (HE)	НЕ-НО
Touristic amenities	0.34	0.37	0.03
Restaurants	0.43	0.56	0.13
Bars	0.59	0.66	0.07
Food stores	0.32	0.57	0.25
Non-food stores	0.53	0.67	0.14
Nurseries	0.51	0.43	-0.08

Notes: Columns "Homogeneous" and "Heterogeneous" report the Gini index for each amenity sector: how concentrated the number of establishments in each sector is across locations. Higher values indicate most of the sector's establishments are clustered in a few locations. Column HE-HO reports the difference between the "Heterogeneous" and "Homogeneous" columns. Positive values in the HE-HO column indicate the spatial distribution of the amenity becomes more clustered across space when preferences are betweeneous.

#### Figure 27:

 because high income groups do not compete with low income groups for the same locations, allowing low income groups to obtain their preferred amenities without having the high income groups bid up their rents

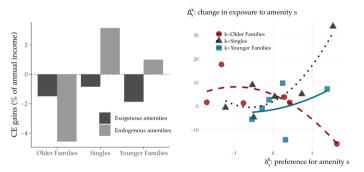
# Counterfactual 2: Decomposing welfare effects of the STR industry

- STR entry reduces rent & change amenity composition
- disentangle: pre-entry baseline, allow STR keeping amenities fixed, then allow amenities to adjust

### Counterfactual 2: Decomposing welfare effects of the STR industry

touristic amenities grow the most in areas populated by old people

Figure 10: Decomposition of welfare effects from STR entry.

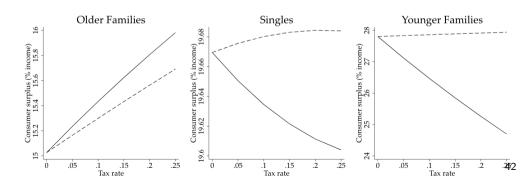


Notes: On the left panel, the consumption equivalent (CE) gains on the vertical axes are computed as how much extra income a household must be given in the baseline equilibrium to obtain the same utility as in the counterfactual equilibrium. Therefore, positive values indicate welfare gains due to STR entry. Details in Appendix for A.5.5. On the right panel, the horizontal axis shows preference parameters for amenity sectors. The vertical axis shows the change in exposure to amenity s after STR entry for a type k household, defined as  $\Delta_s^k \equiv \sum_j \Delta N_{sj} \times \omega_j^k$ , where  $\Delta N_{sj}$  is the change in sector s amenities in location j after STR entry, weighted by  $\omega_j^k = M_j^k/M^k$ , location j's share of the city-wide population of type k before STR entry. Hence,  $\omega_j^k$  is type k's exposure to location j.

# Counterfactual 3: Taxing STRs and/or Touristic amenities

- tax short-term rentals directly reduces rent
- vs tax touristic amenities
- taxing STR has monotonically increasing effects (in tax rates) on welfare

Figure 12: Welfare effects: short-term rental tax vs. touristic amenity tax.



#### Conclusion

- studied the role of preference heterogeneity over a set of endogenous location amenities in shaping within-city sorting and welfare inequality
- there exists hetereogenity preference over amenities
- leads to increased sorting but welfare effects are ambiguous

See ya