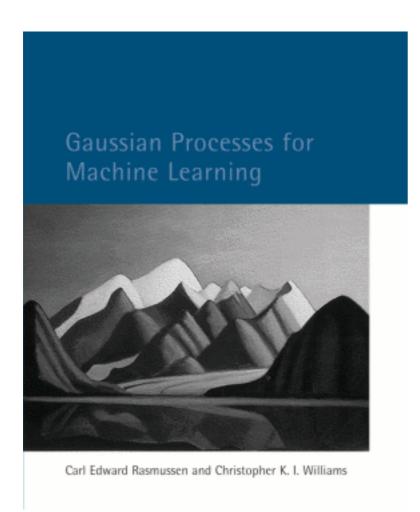


Objectives

- to do regression using a GP
- to find the hyperparameters of the GP by maximizing the (marginal) likelihood
- to use GP regression for uncertainty propagation



The Best Book on the Subject



Gaussian Processes for Machine Learning
Carl Edward Rasmussen and Christopher K. I.
Williams
The MIT Press, 2006. ISBN 0-262-18253-X.

Free online at www.gaussianprocess.org. With Matlab code.



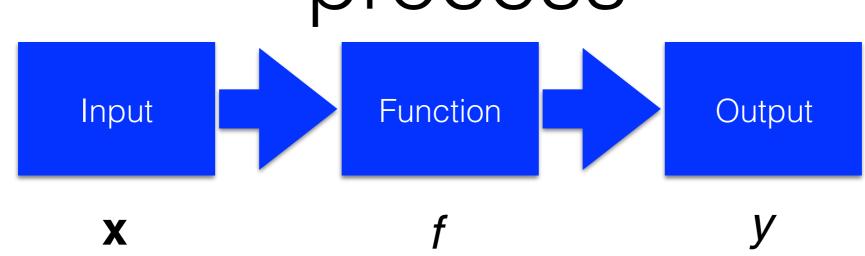
The Best Code on the Subject

GPy (in Python) from the group of N. Lawrence @ University of Sheffield

https://github.com/SheffieldML/GPy



Definition of a Gaussian process



- Treat f as unknown
- Unknown = uncertain = "random", i.e., described with probabilities
- Let us denote our beliefs about f as follows:

$$f(\cdot) \sim p(f(\cdot))$$



Definition of a Gaussian process

A Gaussian process needs two ingredients:

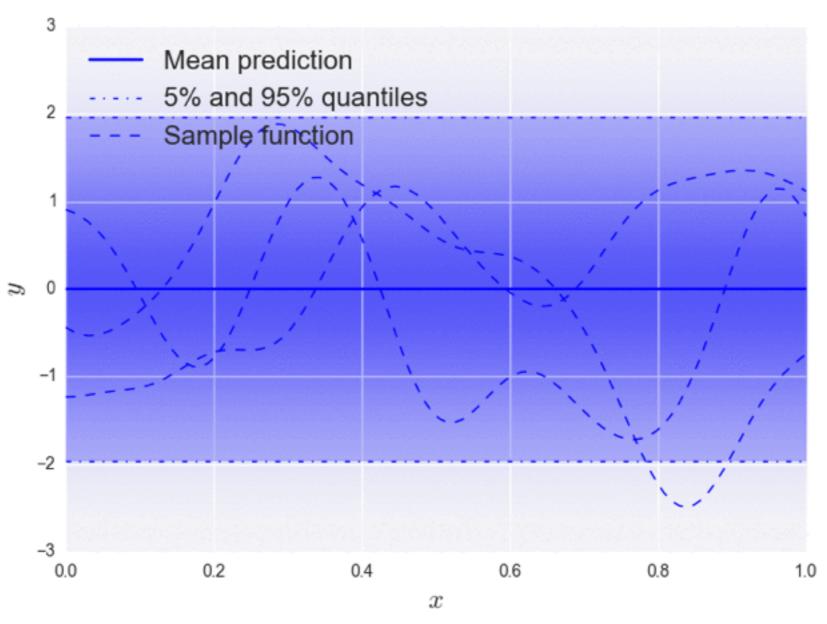
- a mean function
- a covariance function

It uses them to define a probability measure on the space of functions.

We write: $f(\cdot) \sim p(f(\cdot)) = GP(f(\cdot) | m(\cdot), k(\cdot, \cdot))$



Bayesian surrogate





GP Regression 1: No Noise

You have some input/output data:

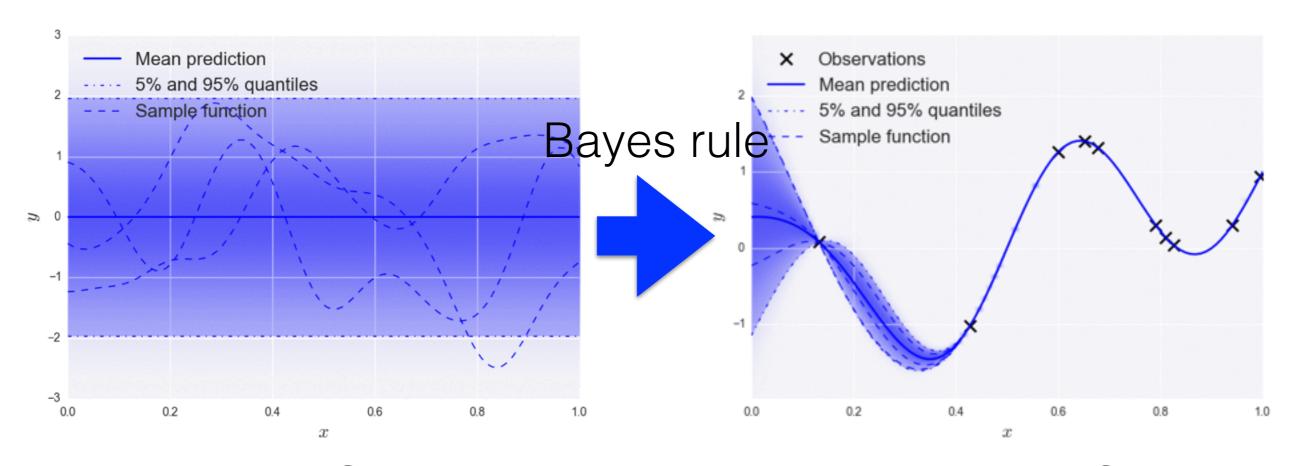
$$\mathbf{X} = \{X_1, \dots, X_N\},\$$

 $\mathbf{f} = \{f(X_1), \dots, f(X_N)\}$

- You wish to learn f.
- Before you start you need to say what you know about f:

$$f(\cdot) \sim \mathsf{GP}(f(\cdot) \mid m(\cdot), k(\cdot, \cdot))$$





Prior GP

Posterior GP



Notebook example Read description



Assume that we have observed:

$$\mathbf{X} = \{X_1, \dots, X_N\},\$$

 $\mathbf{f} = \{f(X_1), \dots, f(X_N)\}$

 and that we want to make predictions at an arbitrary set of *test* inputs:

$$\mathbf{X}^* = \{x_1^*, \dots, x_{N^*}^*\}$$

$$\mathbf{f}^* = \{f(x_1^*), \dots, f(x_{N^*}^*)\}$$



• Since, we have assumed a priori that:

$$p(f(\cdot)) \sim GP(f(\cdot) \mid m(\cdot), k(\cdot, \cdot))$$

then by definition:

$$\rho\begin{pmatrix} \mathbf{f} \\ \mathbf{f}^* \end{pmatrix} = \mathcal{N}\left(\begin{pmatrix} \mathbf{f} \\ \mathbf{f}^* \end{pmatrix} | \begin{pmatrix} \mathbf{m} \\ \mathbf{m}^* \end{pmatrix}, \begin{pmatrix} \mathbf{K}(\mathbf{X}, \mathbf{X}) & \mathbf{K}(\mathbf{X}, \mathbf{X}^*) \\ \mathbf{K}(\mathbf{X}^*, \mathbf{X}) & \mathbf{K}(\mathbf{X}^*, \mathbf{X}^*) \end{pmatrix}\right)$$



Mean on observations

Covariance matrix of observations

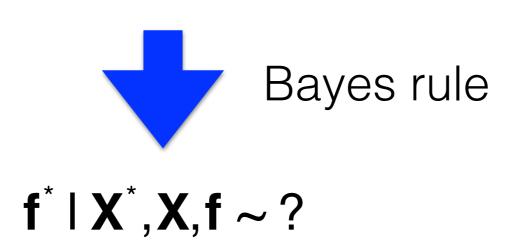
$$\rho\begin{pmatrix} \mathbf{f} \\ \mathbf{f}^* \end{pmatrix} = \mathcal{N}\begin{pmatrix} \begin{pmatrix} \mathbf{f} \\ \mathbf{f}^* \end{pmatrix} | \begin{pmatrix} \mathbf{m} \\ \mathbf{m}^* \end{pmatrix}, \begin{pmatrix} \mathbf{K}(\mathbf{X}, \mathbf{X}) & \mathbf{K}(\mathbf{X}, \mathbf{X}^*) \\ \mathbf{K}(\mathbf{X}^*, \mathbf{X}) & \mathbf{K}(\mathbf{X}^*, \mathbf{X}^*) \end{pmatrix}$$

Mean on test inputs

PREDICTIVE SCIENCE LABORATORY Cross covariance matrix (testobserved)

Covariance matrix of test inputs

$$\rho\begin{pmatrix} \mathbf{f} \\ \mathbf{f}^* \end{pmatrix} = \mathcal{N}\left(\begin{pmatrix} \mathbf{f} \\ \mathbf{f}^* \end{pmatrix} | \begin{pmatrix} \mathbf{m} \\ \mathbf{m}^* \end{pmatrix}, \begin{pmatrix} \mathbf{K}(\mathbf{X}, \mathbf{X}) & \mathbf{K}(\mathbf{X}, \mathbf{X}^*) \\ \mathbf{K}(\mathbf{X}^*, \mathbf{X}) & \mathbf{K}(\mathbf{X}^*, \mathbf{X}^*) \end{pmatrix}\right)$$





$$\rho\begin{pmatrix} \mathbf{f} \\ \mathbf{f}^* \end{pmatrix} = \mathcal{N}\left(\begin{pmatrix} \mathbf{f} \\ \mathbf{f}^* \end{pmatrix} | \begin{pmatrix} \mathbf{m} \\ \mathbf{m}^* \end{pmatrix}, \begin{pmatrix} \mathbf{K}(\mathbf{X}, \mathbf{X}) & \mathbf{K}(\mathbf{X}, \mathbf{X}^*) \\ \mathbf{K}(\mathbf{X}^*, \mathbf{X}) & \mathbf{K}(\mathbf{X}^*, \mathbf{X}^*) \end{pmatrix}\right)$$

Bayes rule

$$\mathbf{f}^* \mid \mathbf{X}^*, \mathbf{X}, \mathbf{f} \sim \mathcal{N}(\mathbf{f}^* \mid \widetilde{\mathbf{m}}, \widetilde{\mathbf{K}}),$$

 $\widetilde{\mathbf{m}} = \mathbf{m}^* + \mathbf{K}(\mathbf{X}^*, \mathbf{X})\mathbf{K}^{-1}(\mathbf{f} - \mathbf{m}),$
 $\widetilde{\mathbf{K}} = \mathbf{K}^* - \mathbf{K}(\mathbf{X}^*, \mathbf{X})\mathbf{K}^{-1}\mathbf{K}(\mathbf{X}, \mathbf{X}^*)$



Proof in Ch. 2.3 Bishop (2006)

The posterior Gaussian process

 Since the choice of test points was arbitrary, the procedure actually defines a posterior Gaussian process:

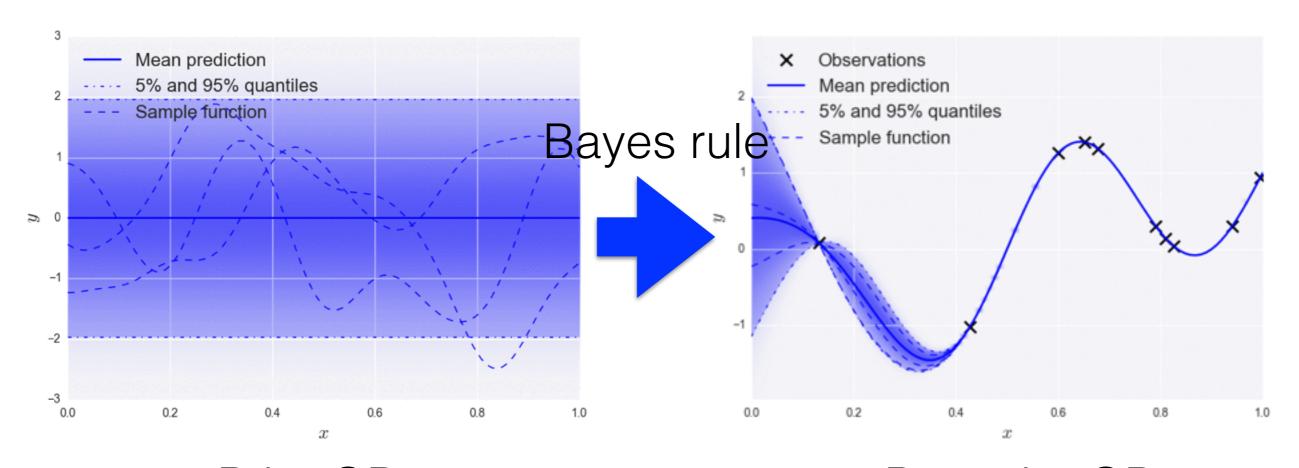
$$p(f(\cdot)|\mathbf{X},\mathbf{f}) = GP(f(\cdot)|\tilde{m}(\cdot),\tilde{k}(\cdot,\cdot)),$$

$$\tilde{m}(\mathbf{x}) = m(\mathbf{x}) + \mathbf{K}(\mathbf{x},\mathbf{X})\mathbf{K}^{-1}(\mathbf{f}-\mathbf{m}),$$

$$\tilde{k}(\mathbf{x},\mathbf{x}') = k(\mathbf{x},\mathbf{x}') - \mathbf{K}(\mathbf{x},\mathbf{X})\mathbf{K}^{-1}\mathbf{K}(\mathbf{X},\mathbf{x}')$$

- This encodes are beliefs about the model output after seeing the data.
- Predictions require a Cholesky decomposition.





Prior GP

Posterior GP



The point predictive distribution

Posterior GP:

$$p(f(\cdot)|\mathbf{X},\mathbf{f}) = GP(f(\cdot)|\tilde{m}(\cdot),\tilde{k}(\cdot,\cdot)),$$

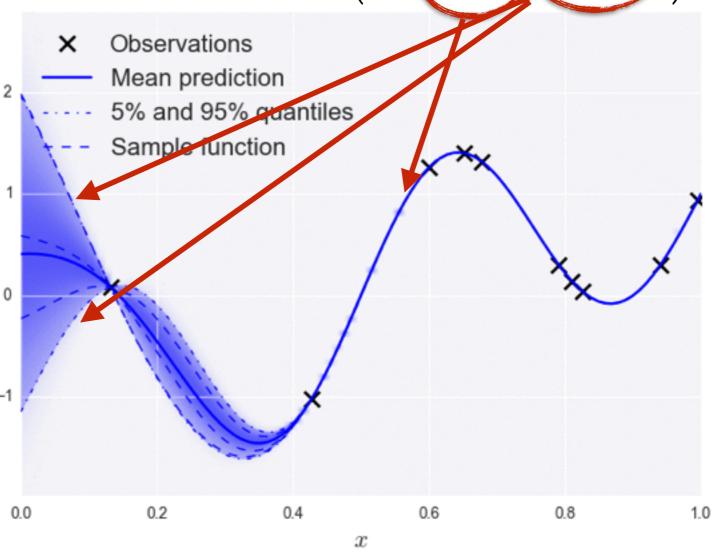
 Looking at just one point, we get the point predictive distribution:

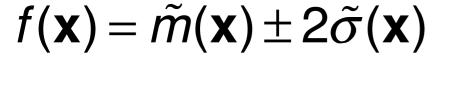
$$p(y \mid \mathbf{x}, \mathbf{X}, \mathbf{f}) = \mathcal{N}(y \mid \tilde{m}(\mathbf{x}), \tilde{\sigma}^{2}(\mathbf{x})),$$
$$\tilde{\sigma}^{2}(\mathbf{x}) = \tilde{k}(\mathbf{x}, \mathbf{x}).$$

You may use the mean as a surrogate.



$$p(y \mid \mathbf{x}, \mathbf{X}, \mathbf{f}) = \mathcal{N}(y \mid \tilde{m}(\mathbf{x}), \tilde{\sigma}^2(\mathbf{x})),$$







GP Regression 2: With Noise

Assume that we have observed:

$$X = \{x_1, ..., x_N\},\$$

 $y = \{y_1, ..., y_N\}$

- where y is a noisy measurement of the ideal f(x) (MD simulation).
- We need to model the measurement process using a likelihood (typically Gaussian):
 Noise (likelihood)

$$p(y_i | f(\mathbf{x}_i)) = \mathcal{N}(y_i | f(\mathbf{x}_i))\sigma^2$$
 variance



Gaussian process regression - Noisy observations

The posterior GP, changes to:

$$p(f(\cdot)|\mathbf{X},\mathbf{f},\sigma^{2}) = GP(f(\cdot)|\tilde{m}(\cdot),\tilde{k}(\cdot,\cdot)),$$

$$\tilde{m}(\mathbf{x}) = m(\mathbf{x}) + \mathbf{K}(\mathbf{x},\mathbf{X})(\mathbf{K} + \sigma^{2}\mathbf{I}_{N})^{-1}(\mathbf{f} - \mathbf{m}),$$

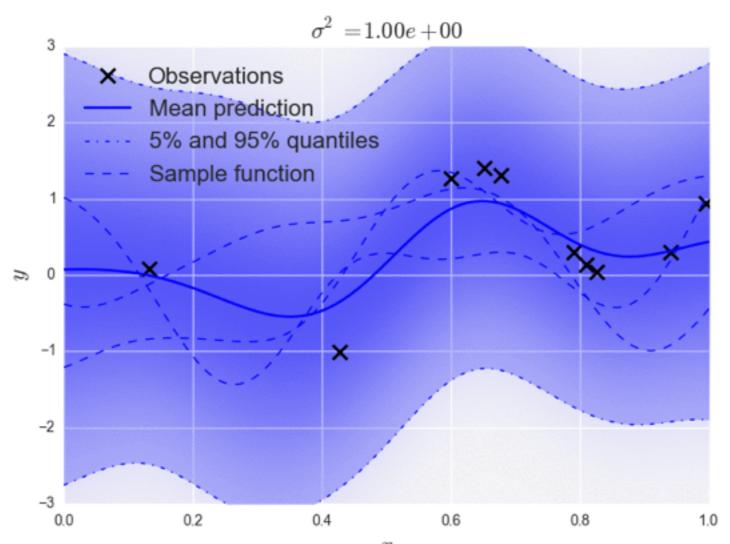
$$\tilde{k}(\mathbf{x},\mathbf{x}') = k(\mathbf{x},\mathbf{x}') - \mathbf{K}(\mathbf{x},\mathbf{X})(\mathbf{K} + \sigma^{2}\mathbf{I}_{N})^{-1}\mathbf{K}(\mathbf{X},\mathbf{x}')$$

and the point predictive distribution to:

$$p(y \mid x, \mathbf{X}, \mathbf{f}) = \mathcal{N}(y \mid \tilde{m}(\mathbf{x}), \tilde{\sigma}^{2}(\mathbf{x})),$$
$$\tilde{\sigma}^{2}(\mathbf{x}) = \tilde{k}(\mathbf{x}, \mathbf{x}) + \sigma^{2}$$



Gaussian process regression - Noisy observations



Each choice of the noise corresponds to a different interpretation of the data.



Noise improves numerical stability

- It is common to use small noise even if there is not any in the data.
- Cholesky fails when covariance is close to being semi-positive definite.
- Adding a small noise improves numerical stability.
- It is known as the "jitter" or as the "nugget" in this case.



Example, Part I, Questions 1-5



 Our prior assumptions were conditional mean and covariance parameters:

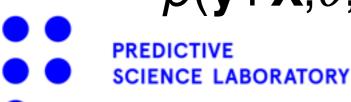
$$p(f(\cdot) \mid \theta) = \mathsf{GP}(f(\cdot) \mid m(\cdot; \theta), k(\cdot, \cdot; \theta))$$

Observations are conditional on the noise level:

$$p(y \mid f(x), \sigma^2) = \mathcal{N}(y \mid f(x), \sigma^2)$$

• Thus, the (marginal) *likelihood* of all the observations is:

$$p(\mathbf{y} \mid \mathbf{X}, \theta, \sigma^2) = p(\mathbf{y} \mid \mathbf{X}, \theta, \sigma^2) = \mathcal{N}(\mathbf{y} \mid \mathbf{m}, \mathbf{K} + \sigma^2 \mathbf{I}_N)$$



• The (marginal) *likelihood* of all the observations is:

$$p(\mathbf{y} \mid \mathbf{X}, \theta, \sigma^2) = \mathcal{N}(\mathbf{y} \mid \mathbf{m}, \mathbf{K} + \sigma^2 \mathbf{I}_N)$$

To complete the prior specification, we must give:

$$p(\theta,\sigma^2) \sim p(\theta,\sigma^2).$$

 Then, after seeing the data, our beliefs about the parameters should change to:

$$p(\theta, \sigma^2 \mid \mathbf{X}, \mathbf{y}) \propto p(\mathbf{y} \mid \mathbf{X}, \sigma^2) p(\theta, \sigma^2)$$

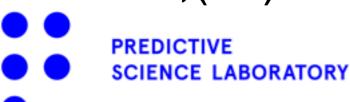


 After seeing the data, our beliefs about the parameters are:

$$p(\theta, \sigma^2 \mid \mathbf{X}, \mathbf{y}) \propto p(\mathbf{y} \mid \mathbf{X}, \sigma^2) p(\theta, \sigma^2)$$

- Ideally, we would sample from this posterior with MCMC.
- Alternatively, we can find the MAP estimate of the parameters:

$$\theta^*, (\sigma^*)^2 = \operatorname{argmax}_{\theta, \sigma} \left\{ \log p(\mathbf{y} \mid \mathbf{X}, \sigma^2) + \log p(\theta, \sigma^2) \right\}$$



MAP estimate of the parameters:

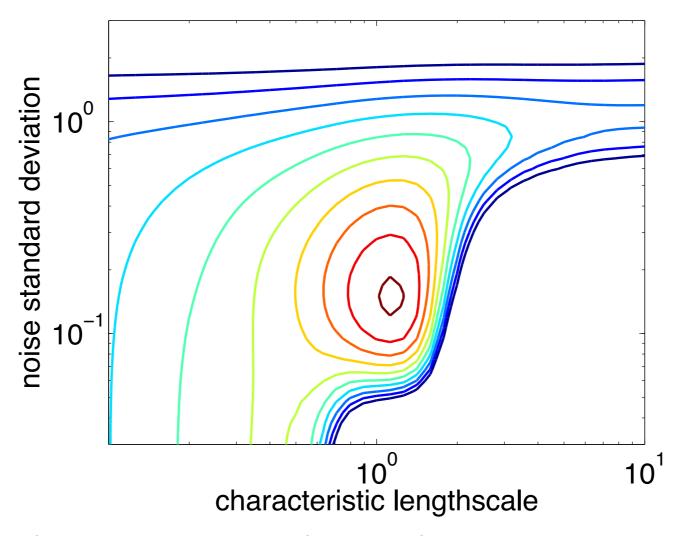
$$\theta^*, (\sigma^*)^2 = \operatorname{argmax}_{\theta, \sigma} \left\{ \log p(\mathbf{y} \mid \mathbf{X}, \sigma^2) + \log p(\theta, \sigma^2) \right\}$$

If our prior assumptions are vague, then

$$\log p(\theta, \sigma^2) = \text{const}$$

and we are effectively just maximizing the likelihood.

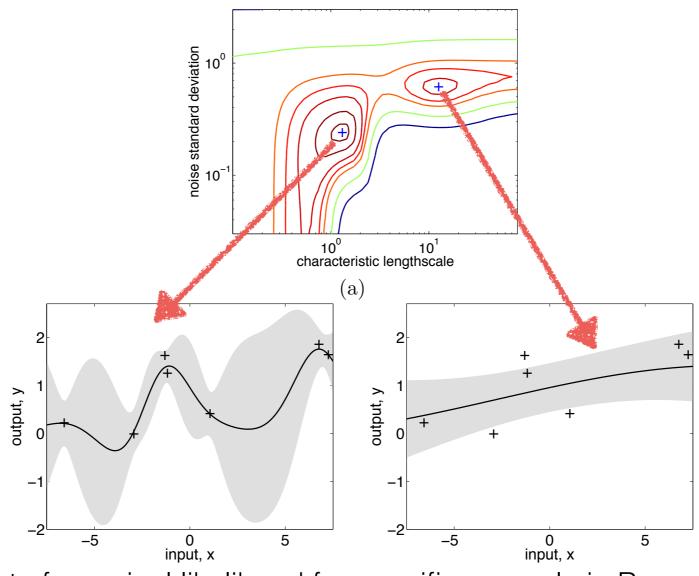








Careful: Different optima correspond to different interpretations







Example, Part I, Questions 6-8 Example, Part II&III

