## Cemuap 3.

2. Рассмотрим оценку вида  $\tilde{\beta} = ((X'X)^{-1} + \gamma I)X'y$  для вектора коэффициентов регрессионного уравнения  $y = X\beta + \varepsilon$ , удовлетворяющего условиям классической регрессионной модели. Найдите  $E(\tilde{\beta})$  и  $Var(\tilde{\beta})$ .

1.

$$\mathbb{E}(\tilde{\beta}) = ((X'X)^{-1} + \gamma I)X'\mathbb{E}(y) = ((X'X)^{-1} + \gamma I)X'X\beta = \beta + \gamma X'X\beta$$

2.

$$\begin{split} \operatorname{Var}(\tilde{\beta}) &= \operatorname{Var}(((X'X)^{-1} + \gamma I)X'y) = \\ & \operatorname{Var}(((X'X)^{-1} + \gamma I)X'\varepsilon) = \\ & (((X'X)^{-1} + \gamma I)X')\operatorname{Var}(\varepsilon)(((X'X)^{-1} + \gamma I)X')' = \\ & (((X'X)^{-1} + \gamma I)X')\sigma_{\varepsilon}^{2}I(((X'X)^{-1} + \gamma I)X')' = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)X'X((X'X)^{-1} + \gamma I) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1} + \gamma I)(I + \gamma X'X) = \\ & \sigma_{\varepsilon}^{2}((X'X)^{-1}$$

3. Пусть регрессионная модель 
$$y_i = \beta_1 + \beta_2 x_{i2} + \beta_3 x_{i3} + \varepsilon_i, \ i = 1, \dots, n$$
, задана в матричном виде при помощи уравнения  $y = X\beta + \varepsilon$ , где  $\beta = \begin{pmatrix} \beta_1 & \beta_2 & \beta_3 \end{pmatrix}'$ . Известно, что  $E(\varepsilon) = 0$  и  $Var(\varepsilon) = \sigma^2 \cdot I$ . Известно также, что:

$$y = \begin{pmatrix} 2 \\ 3 \\ 4 \\ 5 \end{pmatrix}, X = \begin{pmatrix} 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 1 & 0 \\ 1 & 1 & 1 \end{pmatrix}$$
 7  $\mathbf{y} = \mathbf{3}$  Для удобства расчётов ниже приведены матрицы:

$$(5 \ 2 \ 1)$$
  $(1/3 \ -1/3 \ 0)$ 

$$X'X = \begin{pmatrix} 5 & 2 & 1 \\ 2 & 2 & 1 \\ 1 & 1 & 1 \end{pmatrix} \text{ if } (X'X)^{-1} = \begin{pmatrix} 1/3 & -1/3 & 0 \\ -1/3 & 4/3 & -1 \\ 0 & 1 & 2 \end{pmatrix}.$$

$$X'X = \begin{pmatrix} 2 & 2 & 1 \\ 1 & 1 & 1 \end{pmatrix}$$
  $\bowtie (X'X)^{-1} = \begin{pmatrix} -1/3 & 4/3 & -1 \\ 0 & -1 & 2 \end{pmatrix}$ .

$$\begin{pmatrix} 1 & 1 & 1 \end{pmatrix}$$
  $\begin{pmatrix} 0 & -1 & 2 \end{pmatrix}$ 

$$(x'x)^{-1}x'y = \begin{pmatrix} 2\\2\\1 \end{pmatrix}$$

c) 
$$\beta = (x'x)^{-1}x'y = \begin{pmatrix} 2 \\ 1 \end{pmatrix}$$
  
d)  $TSS = \sum_{i=1}^{2} (y_i - \bar{y})^2 = \sum_{i=1}^{2} (y_i - 3)^2 - (1-3)^2 + (2-3)^2 + o_4^2$ 

$$TSS = \frac{2}{2}(y - \overline{y})^{2} = \frac{5}{2}(y - \overline{y})^{2}$$

$$+ (4-3)^{2} + (5-3)^{2} = 10$$

$$ASS = \sum_{i=1}^{5} (y_{i} - \hat{y}_{i})^{2} = (-1)^{2} + 1^{2} = 2$$

$$\hat{y} = k\beta = \begin{pmatrix} 1000 \\ 1000 \\ 1110 \end{pmatrix} \begin{pmatrix} 2 \\ 2 \\ 4 \end{pmatrix} = \begin{pmatrix} 2 \\ 2 \\ 4 \end{pmatrix}$$

e) 
$$e_{y} = y_{4} - \hat{y}_{4} = 4 - 4 = 0$$
  
f)  $R^{2} = \frac{ESS}{TSS} = \frac{8}{10} = 0.8$   
g)  $\hat{T}^{2}_{z} = \frac{RSS}{n-k} = \frac{2}{5-3} = 1$   
h)  $Var(\beta) = \hat{T}^{2}_{z}(x/x)^{-1} = \frac{1}{2}$ 

 $(1) \quad \text{Cov}(\beta_1, \beta_2) = \sqrt{2} \cdot (\chi'\chi)^{-1}_{12} = 1 \cdot (-\frac{1}{3}) = \frac{1}{3}$ 

k) Var (\beta\_1 + \beta\_2) = Var (\beta\_1) + Var (\beta\_2) +

+2 cov (B1, B2)=1. 1 + 1. 4 + 2. (-13)=

$$= 1. \begin{pmatrix} 1/3 & -1/3 & 0 \\ -1/3 & 4/3 & -1 \\ 0 & -1 & 2 \end{pmatrix}$$

$$= 1 \cdot \begin{pmatrix} 1/3 & -1/3 & 0 \\ -1/3 & 4/3 - 1 \\ 0 & -1 & 2 \end{pmatrix}$$

$$= 1 \cdot \begin{pmatrix} 1/3 & -1/3 & 0 \\ -1/3 & 4/3 - 1 \\ 0 & -1 & 2 \end{pmatrix}$$

$$= 1 \cdot 1/3 = \frac{1}{3}$$

 $=\frac{3}{3}$  = 1











$$\begin{array}{l} (C) & (C) &$$

1
$$= \frac{-1/3}{\sqrt{\frac{4}{9}}} = -\frac{1}{2}$$
m)  $\hat{\mathcal{T}}(\hat{\beta}_1) = \sqrt{\hat{\mathcal{T}}_{\hat{\beta}_1}^2} = \sqrt{|\hat{\alpha}_1|\hat{\beta}_1|^2}$ 

 $= \sqrt{\frac{1}{V_{\varepsilon}^{2}} \cdot (x'x)_{11}} = \sqrt{1 \cdot \frac{1}{3}} = \frac{1}{\sqrt{3}}$