University of Chicago, Spring 2025

due Thursday, April 10 at 11:59pm

- 1. (25 points) In this problem, we illustrate the Expectation-Maximization (EM) algorithm using concrete examples. Suppose that the complete dataset consists of $\mathcal{Z} = (\mathcal{X}, \mathcal{Y})$, where \mathcal{X} is observed but \mathcal{Y} is unobserved. The log likelihood for \mathcal{Z} is then denoted by $l(\theta; \mathcal{X}, \mathcal{Y})$, where θ determines the unknown parameter vector. We repeat the E-Step and M-Step below until the sequence of θ_{new} 's converges (which can be guaranteed for some cases for local maximum).
 - **E-Step** (Expectation Step): We compute the expected value of $l(\theta; \mathcal{X}, \mathcal{Y})$, using (a) the information gained from the observed data \mathcal{X} , and (b) the current parameter estimate, θ_{old} . More precisely, let

$$Q(\theta; \theta_{\text{old}}) := \mathbb{E}\left[l(\theta; \mathcal{X}, \mathcal{Y}) \mid \mathcal{X}, \theta_{\text{old}}\right] = \int l(\theta; \mathcal{X}, y) p(y \mid \mathcal{X}, \theta_{\text{old}}) \, \mathrm{d}y. \tag{1}$$

where $p(\cdot | \mathcal{X}, \theta_{\text{old}})$ is the conditional density of \mathcal{Y} given the observed data \mathcal{X} .

M-Step (Maximization Step): We maximize θ over the conditional expectation (1). We simply set $\theta_{\text{new}} := \max_{\theta} Q(\theta; \theta_{\text{old}})$, and afterwards, let $\theta_{\text{old}} = \theta_{\text{new}}$.

(a) We now derive the algorithm above. Let $p(\cdot | \cdot)$ denote an arbitrary conditional probability density function. Show that

$$l(\theta; \mathcal{X}) = \ln p(\mathcal{X} \mid \theta) = \ln \int p(\mathcal{X}, y \mid \theta) \, dy \ge Q(\theta; \theta_{\text{old}}) - \mathbb{E} \left[\ln p(\mathcal{Y} \mid \mathcal{X}, \theta_{\text{old}}) \mid \mathcal{X}, \theta_{\text{old}} \right]. \tag{2}$$

- (b) Denote the rightmost side of (2) by $g(\theta \mid \theta_{\text{old}})$. It is clear that $l(\theta; \mathcal{X}) \geq g(\theta \mid \theta_{\text{old}})$. Prove that we have equality when $\theta = \theta_{\text{old}}$. Why does this imply that the EM algorithm is reasonable for maximizing likelihood?
- (c) Now, consider the multinomial distribution with four classes Mult (n, π_{θ}) where

$$\pi_{\theta} = \left(\frac{1}{2} + \frac{1}{4}\theta, \frac{1}{4}(1-\theta), \frac{1}{4}(1-\theta), \frac{1}{4}\theta\right).$$

let $\mathbf{x} := (x_1, x_2, x_3, x_4)$ be a sample from this distribution. Write down the likelihood $L(\theta; \mathbf{x})$, and log-likelihood $l(\theta; \mathbf{x})$, for sample \mathbf{x} .

(d) We will maximize $l(\theta; \mathbf{x})$ over θ using the EM algorithm (other algorithms will receive no marks), as a toy example. To use EM, we assume that the complete data \mathcal{Z} is given by $\mathbf{y} := (y_1, y_2, y_3, y_4, y_5)$ and that \mathbf{y} has a 5-class Mult (n, π_{θ}^*) distribution where

$$\pi_{\theta}^* = \left(\frac{1}{2}, \frac{1}{4}\theta, \frac{1}{4}(1-\theta), \frac{1}{4}(1-\theta), \frac{1}{4}\theta\right).$$

However, instead of observing \mathbf{y} directly, we are only able to observe $\mathbf{x} = (y_1 + y_2, y_3, y_4, y_5)$. Therefore, we let $\mathcal{X} = (y_1 + y_2, y_3, y_4, y_5)$ and $\mathcal{Y} = y_2$, where \mathcal{Y} remains unobserved. Write down the E-Step and M-Step update equations, with derivations.

2. (15 points) As we saw in class, k-means clustering minimizes the average square distance distortion

$$J_{\text{avg}^2} = \sum_{j=1}^k \sum_{\mathbf{x} \in C_j} d(\mathbf{x}, \mathbf{m}_j)^2,$$
(3)

where $d(\mathbf{x}, \mathbf{x}') = ||\mathbf{x} - \mathbf{x}'||$, and C_j is the set of points belonging to cluster j. Another distortion function that we mentioned is the intra-cluster sum of squared distances,

$$J_{\text{IC}} = \sum_{j=1}^{k} \frac{1}{|C_j|} \sum_{\mathbf{x} \in C_j} \sum_{\mathbf{x}' \in C_j} d(\mathbf{x}, \mathbf{x}')^2.$$

- (a) Given that in k-means, $m_j = \frac{1}{|C_i|} \sum_{\mathbf{x} \in C_i} \mathbf{x}$, show that $J_{\text{IC}} = 2 J_{\text{avg}^2}$.
- (b) Let $\gamma_i \in \{1, ..., k\}$ be the cluster assignment of the *i*'th data point \mathbf{x}_i , and let *n* be the total number of data points. Then

$$J_{\mathrm{avg}^2}(\gamma_1,\ldots,\gamma_n, \boldsymbol{m}_1,\ldots,\boldsymbol{m}_k) = \sum_{i=1}^n d(\mathbf{x}_i, \boldsymbol{m}_{\gamma_i})^2.$$

Recall that k-means clustering alternates the following two steps:

1. Update the cluster assignments:

$$\gamma_i \leftarrow \arg\min_{j \in \{1,\dots,k\}} d(\mathbf{x}_i, \mathbf{m}_j) \quad \forall i = 1,\dots,n.$$

2. Update the centroids:

$$m_j \leftarrow \frac{1}{|C_j|} \sum_{i: \gamma_i = j} \mathbf{x}_i \quad j = 1, \dots, k.$$

Show that step 1 minimizes J_{avg^2} w.r.t. the assignments (holding $\{m_j\}$ fixed), and step 2 minimizes J_{avg^2} w.r.t. the centroids (holding the assignments fixed).

- 3. (10 points) Implement the k-means algorithm in a language of your choice, initializing the cluster centers randomly. The algorithm terminates when no further change in cluster assignments or centroids occurs.
 - (a) Use the toy dataset toydata.txt (500 points in \mathbb{R}^2 , from 3 well-separated clusters). Plot the final clustering assignments (by color or symbol) and also, on a separate figure, plot the distortion value vs. iteration for 20 separate runs. Comment on whether you get the "correct" clusters each time, and on the variability of results across runs.
 - (b) Implement k-means++ initialization and repeat part (a). Compare convergence (speed and final distortion) to the original random initialization.
 - (c) Run both the original and k-means++ algorithms on the MNIST dataset (images are 28×28 pixels, i.e. 784-dimensional vectors). Compare how they converge and how results differ for k = 10 vs. k = 16. You can download MNIST via

Explain any differences you observe in speed, distortion, or cluster quality.

4. (50 points) Recall the Gaussian mixture model for clustering

$$p(\mathbf{x}, z) = \pi_z \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}_z, \boldsymbol{\Sigma}_z),$$

with parameters $\theta = \{\{\pi_j\}, \{\boldsymbol{\mu}_j\}, \{\boldsymbol{\Sigma}_j\}\}_{j=1}^k$.

- (a) Given an i.i.d. sample $\{(\mathbf{x}_i, z_i)\}_{i=1}^n$ from the model, write down the complete-data log-likelihood $\ell(\theta)$, ignoring additive constants that do not affect optimization.
- (b) Let $p_{i,j} = P(z_i = j \mid \mathbf{x}_i)$. Give an expression for $p_{i,j}$ in terms of the mixture parameters.
- (c) Derive the expected complete-data log-likelihood $\bar{\ell}_{\theta_{\text{old}}}(\theta)$ with respect to these posterior probabilities $p_{i,j}$.
- (d) Show that maximizing $\overline{\ell}_{\theta_{\rm old}}(\theta)$ under the constraint $\sum_j \pi_j = 1$ gives

$$\pi_j \leftarrow \frac{1}{n} \sum_{i=1}^n p_{i,j}.$$

- (e) Similarly, derive the updates for μ_j and Σ_j .
- (f) Compare these updates to the k-means updates from Question 2.
- (g) Apply the mixture-of-Gaussians EM algorithm to the toy data and comment on how it clusters the points vs. k-means (both accuracy and convergence speed).

5. (Extra Credit up to 20 points)

Create a dataset for which k-means++ leads to solutions whose final distortion is at least 10 times better (on average) than random initialization. Provide the code used to generate the data.