The Bayes Classifier

We have been starting to look at the supervised classification problem: we are given data (\boldsymbol{x}_i, y_i) for i = 1, ..., n, where $\boldsymbol{x}_i \in \mathbb{R}^d$, and $y_i \in \{1, ..., K\}$. In this section, we suppose that we know everything there is to know about the data (in a probabilistic sense): we assume that we know the *joint distribution* of (X, Y). If we have full knowledge of the distribution, then we can design an optimal classifier without seeing any data at all.

We now make the mathematical setup completely concrete. The "feature vector" X is a random vector in \mathbb{R}^d , and the "class label" Y is a discrete random scalar in $\{1,\ldots,K\}$. When we say that we have a joint probability distribution for (X,Y), it means that we have a rule that assigns probabilities to events that obeys the Kolmogorov axioms. Given $\mathcal{X} \subset \mathbb{R}^d$ and $\mathcal{Y} \subset \{1,\ldots,K\}$, the joint distribution gives us a probability

$$P[X \in \mathcal{X}, Y \in \mathcal{Y}] = \text{ probability that } X \text{ is in } \mathcal{X} \text{ and } Y \text{ is in } \mathcal{Y}.$$

We will treat the entries in the feature vector as continuous-valued. Fixing the feature vector X at different points \boldsymbol{x} results in different conditional probability mass functions (pmfs) for the class label Y:

$$\eta_k(\boldsymbol{x}) := p_{Y|X}(k|\boldsymbol{x}) = P[Y = k|X = \boldsymbol{x}]. \tag{1}$$

The pmf $p_{Y|X}(k|\mathbf{x})$, which is also called the **a posteriori distribution**, will play a central role in much of our discussion here and throughout the course. We encounter it so often that it is useful to give it the more compact notation $\eta_k(\mathbf{x})$.

¹We use non-bold capital letters for all random variables in these notes, whether they are scalar-, vector-, matrix-, or whatever-valued.

²https://en.wikipedia.org/wiki/Probability_axioms

It is also useful to note that fixing the class label Y to different values y results in different conditional probability density functions (pdfs) for the feature vector X: $f_{X|Y}(\boldsymbol{x}|y)$, where

$$P[X \in \mathcal{X}|Y = y] = \int_{\mathcal{X}} f_{X|Y}(\boldsymbol{x}|y) d\boldsymbol{x}.$$

 $f_{X|Y}(\boldsymbol{x}|y)$ is the **class conditional distribution** of X, i.e., the distribution of X given that Y belongs to class y.

A classification rule or **classifier** is simply a function $h : \mathbb{R}^d \to \{1, \ldots, K\}$; that is, a function which takes a feature vector and returns a class label. We can specify this classification rule by **paritioning** \mathbb{R}^d into K regions $\Gamma_1(h), \ldots, \Gamma_K(h)$, where $\Gamma_k(h)$ is the set of point that h maps to k:

$$\Gamma_k(h) = \{ \boldsymbol{x} \in \mathbb{R}^d : h(\boldsymbol{x}) = k \}.$$

We will judge the quality of a classifier by the probability that it makes a mistake:

$$R(h) = P[h(X) \neq Y].$$

This is also called the **risk** of h, or the **probability of error**.

We can now ask a very well-defined question which has a clear-cut answer: What is the classifier that minimizes the probability of error? The answer is simple: given $X = \boldsymbol{x}$, choose the class label that maximizes the conditional probability in (1).

Theorem: Define the classifier

$$h^*(\boldsymbol{x}) = \arg\max_{k \in \{1,\dots,K\}} \eta_k(\boldsymbol{x}). \tag{2}$$

Then every other classifier h has

$$R(h) \ge R(h^*).$$

Proof: The optimality of h^* in (2) follows from carefully writing down the risk for an arbitrary classifier h, applying Bayes rule, and then showing that h^* optimizes the resulting expression. We start with an expression for 1 - R(h), which we will show is as *large* as possible when $h = h^*$:

$$1 - R(h) = P[h(X) = Y]$$

$$= \sum_{k=1}^{K} P[Y = k] \cdot P[h(X) = k | Y = k]$$

$$= \sum_{k=1}^{K} P[Y = k] \int_{\Gamma_k(h)} f_{X|Y}(\boldsymbol{x}|k) d\boldsymbol{x}$$

$$= \sum_{k=1}^{K} \int_{\Gamma_k(h)} P[Y = k] f_{X|Y}(\boldsymbol{x}|k) d\boldsymbol{x}.$$

By Bayes rule,

$$\eta_k(\boldsymbol{x}) = rac{\mathrm{P}\left[Y=k
ight]f_{X|Y}(\boldsymbol{x}|k)}{\sum_{\ell=1}^K\mathrm{P}\left[Y=\ell
ight]f_{X|Y}(\boldsymbol{x}|\ell)}.$$

Note that the denominator is a function of \boldsymbol{x} that is independent of k; it is in fact the marginal density $f_X(\boldsymbol{x})$ for X. Using this and the fact that the regions $\Gamma_k(h)$ are disjoint, we can continue the string of equalities:

$$1-R(h) = \int_{\mathbb{R}^d} \left(\sum_{k=1}^K 1_{\Gamma_k(h)}(oldsymbol{x}) f_X(oldsymbol{x}) \eta_k(oldsymbol{x})
ight) \; doldsymbol{x},$$

where $1_{\mathcal{A}}(\boldsymbol{x})$ is the indicator function

$$1_{\mathcal{A}}(\boldsymbol{x}) = \begin{cases} 1, & \boldsymbol{x} \in \mathcal{A}, \\ 0, & \boldsymbol{x} \notin \mathcal{A}. \end{cases}$$

The way we choose h^* in (2) chooses the regions so that the function inside the integral above is as large as possible; it is clear that

$$\sum_{k=1}^K 1_{\Gamma_k(h)}(oldsymbol{x}) f_X(oldsymbol{x}) \eta_k(oldsymbol{x}) \ \le \ \sum_{k=1}^K 1_{\Gamma_k(h^*)}(oldsymbol{x}) f_X(oldsymbol{x}) \eta_k(oldsymbol{x}),$$

for all $\boldsymbol{x} \in \mathbb{R}^d$. Thus

$$1 - R(h) \le \int_{\mathbb{R}^d} \left(\sum_{k=1}^K 1_{\Gamma_k(h)}(\boldsymbol{x}) f_X(\boldsymbol{x}) \eta_k(\boldsymbol{x}) \right) d\boldsymbol{x}$$

= $1 - R(h^*),$

and so $R(h^*) \leq R(h)$.

The nearest neighbor classifier

We have just seen that the Bayes classifier is optimal. Unfortunately, it requires complete knowledge of the conditional probability mass function $\eta_k(\boldsymbol{x})$. In the context of machine learning, this is not a reasonable assumption. The **nearest neighbor classifier** is an extremely simple alternative. For any \boldsymbol{x} , we simply find the closest point \boldsymbol{x}_i in the training set and then assign \boldsymbol{x} the same label as its nearest neighbor.

This is an incredibly simple rule, but perhaps somewhat surprisingly we can show that as $n \to \infty$, i.e., as the size of our training data grows, this simple classifier is near-optimal. To see this, we will consider the risk of the nearest neighbor classifier $h^{\rm NN}$ conditioned on $X = \boldsymbol{x}$ and compare this to the risk of the Bayes classifier h^* .

To make our discussion simpler, we will restrict our attention to the case of binary classification where $y_i \in \{0, 1\}$. We first note that the

risk of the Bayes classifier h^* conditioned on $X = \boldsymbol{x}$ is given by

$$R^*(\boldsymbol{x}) := P[Y \neq h^*(\boldsymbol{x})|X = \boldsymbol{x}].$$

If $h^*(\boldsymbol{x}) = 0$ then we have $R^*(\boldsymbol{x}) = P[Y = 1 | X = \boldsymbol{x}] = \eta_1(\boldsymbol{x})$. Similarly, if $h^*(\boldsymbol{x}) = 1$ we have $R^*(\boldsymbol{x}) = \eta_1(\boldsymbol{x})$. Since by definition $h^*(\boldsymbol{x})$ selects the label that $maximizes\ \eta_k(\boldsymbol{x})$, we thus have that

$$R^*(\boldsymbol{x}) = \min\{\eta_0(\boldsymbol{x}), \eta_1(\boldsymbol{x})\}. \tag{3}$$

For the nearest neighbor classifier, note that

$$R^{\mathrm{NN}}(\boldsymbol{x}) := P\left[h^{\mathrm{NN}}(\boldsymbol{x}) \neq Y | X = \boldsymbol{x}\right].$$

In our analysis, we will treat not only (X, Y) as random, but also the output $h^{\text{NN}}(\boldsymbol{x})$ as random since it depends on the dataset, which is itself drawn at random from the same distribution as (X, Y). This allows us to write

$$R^{\text{NN}}(\boldsymbol{x}) = P[Y = 0|X = \boldsymbol{x}] P[h^{\text{NN}}(\boldsymbol{x}) = 1|X = \boldsymbol{x}] + P[Y = 1|X = \boldsymbol{x}] P[h^{\text{NN}}(\boldsymbol{x}) = 0|X = \boldsymbol{x}].$$
(4)

If $\boldsymbol{x}_{\text{NN}}$ denotes the nearest neighbor to \boldsymbol{x} , then we can write

$$P\left[h^{\mathrm{NN}}(\boldsymbol{x})=k|X=\boldsymbol{x}\right]=P\left[Y=k|X=\boldsymbol{x}_{\mathrm{NN}}
ight]=\eta_k(\boldsymbol{x}_{\mathrm{NN}}).$$

As $n \to \infty$, we have that $\|\boldsymbol{x}_{\text{NN}} - \boldsymbol{x}\| \to 0$, and thus as $n \to \infty$ we have

$$\eta_k(\boldsymbol{x}_{\mathrm{NN}}) \to \eta_k(\boldsymbol{x}).$$

Plugging this back into (4) and simplifying, we obtain

$$R^{\text{NN}}(\boldsymbol{x}) \to \eta_0(\boldsymbol{x})\eta_1(\boldsymbol{x}) + \eta_1(\boldsymbol{x})\eta_0(\boldsymbol{x})$$

$$= 2\eta_0(\boldsymbol{x})\eta_1(\boldsymbol{x})$$

$$\leq 2\min\{\eta_0(\boldsymbol{x}), \eta_1(\boldsymbol{x})\},$$

where the last inequality follows from the fact that both $\eta_1(\boldsymbol{x})$ and $\eta_2(\boldsymbol{x})$ are less than 1. Combining this with (3), this yields

$$\lim_{n\to\infty} R^{\rm NN}(\boldsymbol{x}) \le 2R^*(\boldsymbol{x}),$$

or in words, that asymptotically, the risk of the nearest neighbor classifier is at most twice the Bayes risk.

This can be strengthened by considering the more general k-nearest neighbors classifier. The idea here is to assign a label to \boldsymbol{x} by taking a majority vote over the k training points closest to \boldsymbol{x} . If $R^{kNN}(\boldsymbol{x})$ denotes the risk of the k-nearest neighbor classifier, then one can show via a similar argument that

$$\lim_{n\to\infty} R^{\text{kNN}}(\boldsymbol{x}) \le \left(1 + \sqrt{2/k}\right) R^*(\boldsymbol{x}).$$

Thus, by increasing k it is possible to drive this multiplicative constant arbitrarily close to 1. This results in a property known as **universal consistency**. Specifically, if R^* denotes the Bayes risk and $R_n^{\rm kNN}$ denotes the risk of the k-nearest neighbors classifier based on a dataset of size n, then one can show that as $n \to \infty$, if $k \to \infty$ while $k/n \to 0$, then $R_n^{\rm kNN} \to R^*$.

In words this is simply saying that for any possible distribution on the data, if we are given enough data eventually the risk of the k-nearest neighbor classifier will converge to the Bayes risk (i.e., to the optimal risk). Unfortunately (or fortunately, depending on your perspective), you might have to wait a very long time, so there is still a role for other machine learning algorithms to improve on this situation when we only have a finite amount of data.