

T-1

Случайная величина распределена равномерно на  $[0, \theta]$ . По выборке объема  $n$  найдены оценки параметра  $\theta$ :

$$\tilde{\theta}_1 = 2\bar{x}; \quad \tilde{\theta}_2 = x_{\min}; \quad \tilde{\theta}_3 = x_{\max};$$

$$\tilde{\theta}_4 = \left( x_1 + \frac{\sum_{k=2}^n x_k}{(n-1)} \right)$$

а)  $\varphi \sim R(0, \theta)$   $\theta$  — неизв. парам.

$\theta > 0$  — вероятн. модель

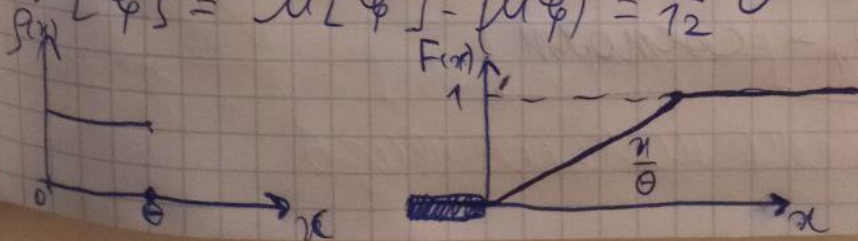
Пусть  $\vec{x}_n$  — выборка объема  $n$

числовые характеристики  $\varphi$ :

$$M[\varphi] = \int_{-\infty}^{+\infty} x p(x) dx = \int_0^{\theta} \frac{x}{\theta} dx = \frac{x^2}{2\theta} \Big|_0^{\theta} = \frac{\theta}{2}$$

$$M\varphi^2 = \int_0^{\theta} x^2 p(x) dx = \int_0^{\theta} \frac{x^2}{\theta} dx = \frac{x^3}{3\theta} \Big|_0^{\theta} = \frac{\theta^2}{3}$$

$$D[\varphi] = M[\varphi^2] - (M\varphi)^2 = \frac{1}{12} \theta^2$$



$$\tilde{\theta}_1 = 2\bar{x} = 2 \cdot \frac{1}{n} \sum_{i=1}^n x_i$$

• Несмещенность:

Определение несмещенности:

$$\forall \theta > 0 \quad M(\tilde{\theta}_1) = \theta$$

$$M\left[\frac{2}{n} \sum_{i=1}^n x_i\right] = \frac{2}{n} \cdot n M[x_i] = 2M[x_i] =$$

$$= \{ \text{т.к. } x_i - \text{одинаково распредел.} \} = 2M[\varphi] =$$

$$= 2 \cdot \frac{\theta}{2} = \theta \rightarrow \tilde{\theta}_1 - \text{несмещенная}$$

• Состоятельность:

Ду:

$\tilde{\theta}$  - несмещенная }  $\Rightarrow \tilde{\theta}$  - состоят.

$$D[\tilde{\theta}] \xrightarrow[n \rightarrow \infty]{\forall \theta \in \Theta} 0$$

$$D[\tilde{\theta}_1] = D\left[\frac{2}{n} \sum_{i=1}^n x_i\right] = \frac{4}{n^2} D\left[\sum_{i=1}^n x_i\right] =$$

$$= \frac{4}{n^2} \cdot n D[x_i] = \frac{4}{n} \cdot D[\varphi] = \frac{4}{n} \cdot \frac{\theta^2}{12} = \frac{\theta^2}{3n} \xrightarrow[n \rightarrow \infty]{} 0$$

$\Rightarrow \tilde{\theta}_1$  - состоят



$$\tilde{\theta}_2 = \min(x_i)$$

• Несмещенность

$$M[\theta_2] = \int_{-\infty}^{+\infty} x p(x) dx = \int_0^{\theta} y \varphi(y) dy$$

$$\Phi(y) = 1 - (1 - F(y))^n$$

$$\Phi'(y) = \varphi(y) = n \underbrace{(1 - F(y))^{n-1}}_{\frac{y'}{\theta}} \underbrace{p(y)}_{\frac{1}{\theta} \mathbb{I}_{(0, \theta)}}$$

$$= n \left(1 - \frac{y}{\theta}\right)^{n-1} \frac{1}{\theta}$$

$$M[\theta_2] = \int_0^{\theta} n \left(1 - \frac{y}{\theta}\right)^{n-1} \frac{1}{\theta} y dy = \int_0^{\frac{t}{\theta}} n \left(1 - \frac{t}{\theta}\right)^{n-1} \frac{t}{\theta} d\frac{t}{\theta} =$$

$$= \int_0^1 n \frac{t}{\theta} \left(1 - \frac{t}{\theta}\right)^{n-1} \frac{1}{\theta} dt = \int_0^1 n \frac{t}{\theta^2} \left(1 - \frac{t}{\theta}\right)^{n-1} dt =$$

$$= \int_0^1 n \frac{t}{\theta^2} \left(1 - \frac{t}{\theta}\right)^{n-1} dt = n \theta \left( \frac{t}{n+1} \Big|_0^1 - \frac{t}{n+1} \Big|_0^1 \right) =$$

$$= n \theta \left( \frac{1}{n} - \frac{1}{n+1} \right) = n \theta \left( \frac{1}{n(n+1)} \right) = \frac{\theta}{n+1} \Rightarrow$$

$\Rightarrow \tilde{\theta}_2$  — смещенная

$\hat{\theta}_2(n+1) \tilde{\theta}_2$  — ~~несмещенная~~ несмещенная

• Следовательно:

$$\begin{aligned}
 M[\tilde{\Theta}_2^2] &= \int_0^{\theta} n \left(1 - \frac{y}{\theta}\right)^{n-1} \frac{1}{\theta} y^2 dy = \\
 &= \left\{ t = 1 - \frac{y}{\theta} \right\} = \int_0^1 n t^{n-1} \theta^2 (1-t)^2 dt = \\
 &= n \theta^2 \left[ \int_0^1 (t^{n-1} - 2t^n + t^{n+1}) dt \right] = \\
 &= n \theta^2 \left[ \frac{1}{n} - 2 \frac{1}{n+1} + \frac{1}{n+2} \right] = n \theta^2 \left[ \frac{(n+1)(n+2) - 2n(n+2) + n(n+1)}{(n+1)(n+2)} \right] = \\
 &= \theta^2 \frac{n^2 + 9n + 2 - 2n^2 - 4n^2 + n^2 + n}{(n+1)(n+2)} = \frac{2\theta^2}{(n+1)(n+2)}
 \end{aligned}$$

$$D[\tilde{\Theta}_2] = \frac{2\theta^2}{(n+1)(n+2)} \neq \frac{\theta^2}{(n+1)^2} = \theta^2 \left[ \frac{n}{(n+1)^2(n+2)} \right] \xrightarrow{n \rightarrow \infty} 0$$

$$\left. \begin{aligned}
 D[\tilde{\Theta}_2] &\xrightarrow{n \rightarrow \infty} 0 \\
 \tilde{\Theta}_2 &\text{ — измерима?}
 \end{aligned} \right\} \Rightarrow \text{т.к. } \tilde{\Theta}_2 \text{ нек. неизб.}$$

$$D[\tilde{\Theta}_2']$$

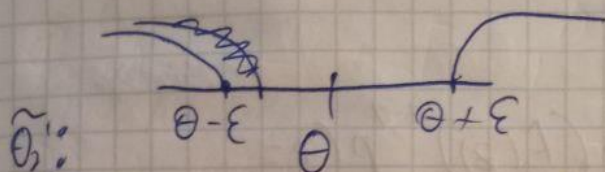
$$D[\tilde{\Theta}_2'] = (n+1)^2 D[\tilde{\Theta}_2] = \frac{\theta^2 n}{n+2} \xrightarrow{n \rightarrow \infty} \theta^2$$

$\Rightarrow$  не измеримо неизб.



.. Неверно по определению

$$\forall \theta > 0 \quad \forall \varepsilon > 0 \quad P(|\tilde{\theta}_2 - \theta| > \varepsilon) \xrightarrow{n \rightarrow \infty} 0$$



$$P(|\tilde{\theta}_2' - \theta| \geq \varepsilon) \geq P(\tilde{\theta}_2' \geq \theta + \varepsilon) =$$

$$= P((n+1) \chi_{\min} \geq \theta + \varepsilon) =$$

$$= P(\chi_{\min} \geq \frac{\theta + \varepsilon}{n+1}) = 1 - P(\chi_{\min} < \frac{\theta + \varepsilon}{n+1}) =$$

$$= 1 - (1 - (1 - F(\frac{\theta + \varepsilon}{n+1}))^n) = \Phi(\frac{\theta + \varepsilon}{n+1})$$

$$= (1 - \frac{\theta + \varepsilon}{\theta(n+1)})^n \xrightarrow{n \rightarrow \infty} e^{-\frac{\theta + \varepsilon}{\theta}} > 0 \Rightarrow \tilde{\theta}_2' \text{ несостоятельна}$$

$$\tilde{\theta}_2: P(|\tilde{\theta}_2 - \theta| \geq \varepsilon) \xrightarrow{n \rightarrow \infty} 0$$

$$\forall \varepsilon > 0$$

$$\forall \theta > 0$$

$$P(\tilde{\theta}_2 < \theta - \varepsilon) + P(\tilde{\theta}_2 \geq \theta + \varepsilon)$$

$$P(\chi_{\min} < \theta - \varepsilon) = \Phi(\theta - \varepsilon) = 1 - (1 - \frac{\theta - \varepsilon}{\theta})^n =$$

$$= 1 - (\frac{\varepsilon}{\theta})^n \xrightarrow{n \rightarrow \infty} 0 \neq 0 \Rightarrow \tilde{\theta}_2 \text{ не состоятельна}$$

$$\tilde{\Theta}_3 = \max_{\text{существование}} (\gamma_{\max})$$

$$M[\tilde{\Theta}_3] = \int_0^{\tilde{\Theta}_3} z \psi(z) dz = \int_0^{\tilde{\Theta}_3} n \frac{z^n}{\tilde{\Theta}^n} dz = \frac{n}{n+1} \tilde{\Theta}$$

$$\Psi = (F(z))^n$$

$$\psi(z) = \Psi'(z) = n(F(z))^{n-1} p(z) =$$

$$= n \left(\frac{z}{\tilde{\Theta}}\right)^{n-1} \frac{1}{\tilde{\Theta}} \{ (0, \tilde{\Theta}) \}$$

$$\tilde{\Theta}_3 = \frac{n+1}{n} \gamma_{\max} - \text{несмещ. оценка}$$

$$M[\tilde{\Theta}_3] = \int_0^{\tilde{\Theta}_3} n \frac{z^n}{\tilde{\Theta}^n} dz = \frac{n}{n+2} \tilde{\Theta}^2$$

• сходимости

$$D[\tilde{\Theta}_3] = \frac{n}{n+2} \tilde{\Theta}^2 - \left(\frac{n}{n+1}\right)^2 \tilde{\Theta}^2 = \tilde{\Theta}^2 \left[ \frac{n(n+1)^2 - n^2(n+2)}{(n+2)(n+1)^2} \right]$$

$$= \tilde{\Theta}^2 \left[ \frac{n^3 + 2n^2 + n - n^3 - 2n^2}{(n+2)(n+1)^2} \right] = \frac{n \tilde{\Theta}^2}{(n+2)(n+1)^2} \xrightarrow{n \rightarrow \infty} 0$$

$$D[\tilde{\Theta}_3] \xrightarrow{n \rightarrow \infty} 0$$

$$\tilde{\Theta}_3 - \text{смещенная}$$

}  $\Rightarrow$  про  $\tilde{\Theta}_3$  ничего сказать нельзя



$\theta_3'$  не ссм:

$$D[\tilde{\theta}_3] = \frac{(n+1)^2}{n^2} D[\tilde{\theta}_3] = \frac{\theta^2}{4(n+2)} \xrightarrow{n \rightarrow \infty} 0 \Rightarrow$$

$\Rightarrow \theta_3'$  - ссм

$\sim$  не ссм.  
 $\theta_3$  не unbiased.

$\forall \varepsilon > 0$

$$\forall \theta > 0: P(|\tilde{\theta}_3 - \theta| \geq \varepsilon) = P(x_{max} < \theta - \varepsilon) +$$

$$+ P(x_{max} > \theta + \varepsilon) = (P(\theta - \varepsilon))^n$$

$$\{ < \theta: \left( \frac{\theta - \varepsilon}{\theta} \right)^n \xrightarrow{n \rightarrow \infty} 0$$

$$\{ > \theta: (0)^n \xrightarrow{n \rightarrow \infty} 0 \Rightarrow \text{ссм.}$$

$$\tilde{\theta}_4 = x_1 + \sum_{k=2}^n x_k / (n-1)$$

• Несмещенность

$$M[\tilde{\theta}_4] = M\left[x_1 + \frac{1}{n-1} \sum_{i=2}^n x_i\right] =$$

$$= M[x_1] + \frac{1}{n-1} \sum_{i=2}^n M[x_i] = \frac{\theta}{2} + \frac{\theta}{2} \hat{=} \theta \Rightarrow$$

$\Rightarrow \tilde{\theta}_4$  - несмещ.

• Соединяем пометы

$$D[\tilde{\theta}_4] = D\left[x_1 + \frac{1}{n-1} \sum_{i=2}^n x_i\right] =$$

$$= D[x_1] + \frac{1}{(n-1)^2} (n-1) D[\varphi] = \frac{\sigma^2}{12} + \frac{\sigma^2}{12(n-1)} =$$

$$= \frac{\sigma^2}{12} \frac{n}{n-1} \xrightarrow{n \rightarrow \infty} 0 \text{ ДУ не падает}$$

$$\tilde{\theta}_4 \xrightarrow{P} \theta$$

$$\text{Св-во: } \varphi_n \xrightarrow{P}, \eta_n \xrightarrow{P} \eta \Rightarrow \varphi_n + \eta_n \xrightarrow{P} \varphi + \eta$$

$$x_1 \xrightarrow{P} x_1$$

$$\frac{1}{n-1} \sum_{i=2}^n x_i \rightarrow \left\{ \begin{array}{l} \text{ЗБЧ хиккика} \\ \varphi_1, \dots, \varphi_n \text{ одинаков. распр.} \\ \text{и независимы} \\ \frac{1}{n} \sum_{i=1}^n \varphi_i \xrightarrow{P} M[\varphi] \end{array} \right\} \rightarrow M[\varphi] =$$

$$= \varphi M \frac{\theta}{2}$$

$$\tilde{\theta}_4 = x_1 + \frac{1}{n-1} \sum_{i=2}^n x_i \rightarrow x_1 + \frac{\theta}{2}$$

$$5) \tilde{\theta}_1 = 2\bar{x}$$

$$\tilde{\theta}_3 = \frac{n+1}{n} x_{\max}$$



$$\left. \begin{aligned} D[\tilde{\theta}_1] &= \frac{\theta^2}{3n} \\ D[\tilde{\theta}_3] &= \frac{\theta^2}{n(n+2)} \end{aligned} \right\} \forall \theta > 0: \frac{\theta^2}{n(n+2)} < \frac{\theta^2}{3n} \Rightarrow$$

$\Rightarrow \tilde{\theta}_3$  — эффективнее  $\tilde{\theta}_1$ ,  $\forall n > 1$   
 $T-3$

Г. В. имеет экспоненц. распред.