

自习社

概率论课后习题 详解

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自习社研究部编

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第一章

习题 1.1

(由于第一节文字较多且原答案较丰富, 所以直接复制过来了...)

1. 略;
2. 设 ω_i 表示“出现 i 点” ($i = 1, 2, \dots, 6$), 则
 - (1) 样本点为 $\omega_1, \omega_2, \omega_3, \omega_4, \omega_5, \omega_6$; 样本空间为 $\Omega = \{\omega_1, \omega_2, \omega_3, \omega_4, \omega_5, \omega_6\}$,
 - (2) $A = \{\omega_2, \omega_4, \omega_6\}$; $B = \{\omega_3, \omega_6\}$.
 - (3) $\bar{A} = \{\omega_1, \omega_3, \omega_5\}$, 表示“出现奇数点”; $\bar{B} = \{\omega_1, \omega_2, \omega_4, \omega_5\}$, 表示“出现的点数不能被3整除”; $A \cup B = \{\omega_2, \omega_3, \omega_4, \omega_6\}$, 表示“出现的点数能被2或3整除”; $AB = \{\omega_6\}$, 表示“出现的点数能被2整除且能被3整除”; $\overline{A \cup B} = \{\omega_1, \omega_5\}$, 表示“出现的点数既不能被2整除也不能被3整除”.
3. (1) 设 ω_i 表示“点数之和等于 i ” ($i = 3, 4, \dots, 18$), 则
$$\Omega = \{\omega_3, \omega_4, \dots, \omega_{18}\};$$
$$A = \{\omega_{11}, \omega_{12}, \dots, \omega_{18}\}; B = \{\omega_3, \omega_4, \dots, \omega_{14}\}.$$
 - (2) 设 ω_{ijk} 表示“出现号码为 i, j, k ” ($i, j, k = 1, 2, \dots, 5, i \neq j \neq k$), 则
$$\Omega = \{\omega_{123}, \omega_{124}, \omega_{125}, \omega_{134}, \omega_{145}, \omega_{135}, \omega_{234}, \omega_{235}, \omega_{245}, \omega_{345}\}$$
$$C = \{\omega_{123}, \omega_{124}, \omega_{125}, \omega_{134}, \omega_{145}\}.$$
4. (1) $A_1 A_2 \cdots A_n$;
 - (2) $\overline{A_1 A_2 \cdots A_n}$ 或 $\bar{A}_1 \cup \bar{A}_2 \cup \cdots \cup \bar{A}_n$;
 - (3) $\bar{A}_1 A_2 \cdots A_n \cup A_1 \bar{A}_2 \cdots A_n \cup \cdots \cup A_1 A_2 \cdots \bar{A}_n$
 - (4) $A_1 \cup A_2 \cup \cdots \cup A_n$ 或 $\overline{\bar{A}_1 \bar{A}_2 \cdots \bar{A}_n}$.

习题 1.2

1.
 - (1) 自己看书, 关系是 $\lim_{n \rightarrow \infty} f_n(A) = P(A)$
 - (2) 自己看书
2. $\because A \cap B = \emptyset \therefore A \cap \bar{B} = A, \therefore P(A\bar{B}) = 0.4$
3. $\because P(A \cup B) = P(A) + P(B) + P(AB) = 0.3$
 $P(A) = 0.1, P(AB) = 0, \therefore P(B) = 0.2$
4. $P(A \cup B \cup C) = P(A) + P(B) + P(C) - p(AB) - p(BC) - p(AC) + P(ABC) = \frac{1}{3} + \frac{1}{3} + \frac{1}{3} - 0 - 0 - \frac{1}{4} + 0 = 0.75$
5. $P(A \cup B) = P(A) + P(B) - P(AB) = 0.8 + 0.6 - 0.48 = 0.92$

习题 1.3

1. 自己看书

$$2. P = \frac{C_{13}^5 C_{13}^3 C_{13}^3 C_{13}^2}{C_{52}^{13}} = 0.01293$$

$$3. P = \frac{C_9^1 \cdot A_8^6 + C_9^1 \cdot C_6^1 \cdot A_7^5}{C_9^1 \cdot 10^6} = 0.0605$$

$$4. P = \frac{C_{95}^{50} + C_{95}^{49} \cdot C_5^1}{C_{100}^{50}} = 0.181$$

5. 直白的说, 夫妇中任何一个人身边一定坐着两个人, 而只要他妻子出现在两侧即可, 就是 $P = \frac{2}{9}$

还可以说, 我们可以给座位号标上 1-10, 那么总共有 A_{10}^{10} 中排列, 而夫妇在一起的情况有两种, 一个 1 一个 10 或者是其他, 1 个 1 一个 10 共有 $2 \times A_8^8$ 可能; 其他情况是 $2 \times A_9^9$,

$$\text{故 } P = \frac{2 \times (A_8^8 + A_9^9)}{A_{10}^{10}} = \frac{2}{9}$$

6. 由几何概型可得



通过计算面积, 可得 $p = \frac{3}{4}$

习题 1.4

1. 自己看书

$$2. \text{解: } p(B|A) = \frac{P(AB)}{P(A)} = 0.95, p(B|\bar{A}) = \frac{P(\bar{A}B)}{P(\bar{A})} = 0.8, p(\bar{B}|A) = \frac{P(A\bar{B})}{P(A)} = 0.05, p(\bar{B}|\bar{A}) =$$

$$\frac{P(\bar{A}\bar{B})}{P(\bar{A})} = 0.2$$

$$3. \text{解: } P(A) = 0.8, P(B) = 0.2, p(B|A) = \frac{P(AB)}{P(A)} = 0.15, p(A|B) = \frac{P(AB)}{P(B)} = 0.6,$$

$$P(AB) = 0.12, P(C) = 0.4, p(C|A) = \frac{P(AC)}{P(A)} = 0.4, p(\bar{A}|\bar{B}) = \frac{P(\bar{A}\bar{B})}{P(\bar{B})} = 0.15, P(AC) = 0.32$$

$$4. \text{解: } P(A) = \frac{C_4^1}{C_{10}^1} = 0.4, P(B) = \frac{4}{10} \cdot \frac{3}{9} = \frac{2}{15}, p(C) = \frac{6}{10} \cdot \frac{4}{9} = \frac{4}{15}, P(D) = P(B) \cdot \frac{2}{8} = \frac{1}{30}$$

$$5. \text{解: } p = \frac{2}{5} \cdot \frac{3}{6} \cdot \frac{3}{7} \cdot \frac{4}{8} = \frac{3}{70}$$

习题 1.5

1. (1) — (6) 略; (7) 先验概率就是这件事没发生之前你去计算他的概率; 后验概率就是这件事发生了你去校正他的概率。

$$2. \text{解: } p = (X = 1, 2, 3, 4 \text{ 概率同}) \cdot \frac{1}{4} \cdot \left(0 + \frac{1}{2} + \frac{1}{3} + \frac{1}{4}\right) = \frac{13}{48}$$

3. 解: 设 A 为硬币是正品, T 为扔 r 次每次均出现国徽, 则 $p(A|T) = \frac{P(AT)}{P(T)} =$

$$\frac{P(T|A)P(A)}{P(T|A)P(A) + P(T|\bar{A})P(\bar{A})} = \frac{\frac{1}{2^r} \frac{m}{m+n}}{\frac{1}{2^r} \frac{m}{m+n} + \frac{n}{m+n}} = \frac{m}{m+n \cdot 2^r}$$

习题 1.6

- 略
- 解: $p = 0.9 * 0.8 * 0.7 + 0.1 * 0.8 * 0.7 + 0.9 * 0.2 * 0.7 + 0.9 * 0.8 * 0.3 = 0.902$
- 解: $p = 0.3 + 0.7 * 0.2 * 0.2 = 0.328$
- 解: $p = \frac{1}{5} + \frac{4}{5} \cdot \frac{1}{3} + \frac{4}{5} \cdot \frac{2}{3} \cdot \frac{1}{4} = \frac{3}{5}$
- 解: 设 A 为飞机被击落, X 为飞机被击中次数。

$$P\{X = 0\} = 0.6 \times 0.5 \times 0.3 = 0.09$$

$$P\{X = 1\} = 0.4 \times 0.5 \times 0.3 + 0.6 \times 0.5 \times 0.3 + 0.6 \times 0.5 \times 0.7 = 0.36$$

$$P\{X = 2\} = 0.4 \times 0.5 \times 0.3 + 0.4 \times 0.5 \times 0.7 + 0.6 \times 0.5 \times 0.7 = 0.41 \quad P\{X = 3\} = 0.14$$

$$\text{则 } P(A) = 0.2P\{X = 1\} + 0.6P\{X = 2\} + P\{X = 3\} = 0.458$$

6. 设做出正确决策为事件 A, 做出正确决策的人数为 X 则 $P\{X = k\} = C_9^k 0.7^k 0.3^{9-k}$,

$$P(A) = P\{X \geq 5\} = 0.901$$

7. 即 $1 - (1 - p)^n \geq p, 1 - p \geq (1 - p)^n, n \geq 1$, 至少一次

习题一

一. 填空题

1. 解: $\because A \cap B = \bar{A} \cap \bar{B} \quad \therefore A$ 与 B 为两个对立事件

$$\therefore A \cup B = \Omega, AB = \phi$$

2. 解: $(\bar{A} \cup B) \cap (A \cup B) = B \quad (A \cup \bar{B}) \cap (\bar{A} \cup \bar{B}) = \bar{B}$

$$B \cap \bar{B} = \phi$$

$$\therefore \text{原式} = 0$$

3. 解: $P(A \cup B \cup C) = P(A) + P(B) + P(C) - P(AB) - P(BC) - P(AC) + P(ABC) = \frac{1}{4} +$

$$\frac{1}{4} + \frac{1}{4} - \frac{1}{16} - \frac{1}{16} = \frac{5}{8}$$

$$P = 1 - \frac{5}{8} = \frac{3}{8}$$

二. 选择题

- 解: $\because P(A) = P(A|B) \therefore A$ 与 B 之间没有关系, 即事件 A 与事件 B 相互独立
- $\because A, B$ 为两个互逆的事件, $\therefore P(A|B) = P(B|A) = 0$
- 若事件 A 与事件 B 相互独立, 则 $P(A|B) + P(\bar{A}|\bar{B}) = P(A) + P(\bar{A}) = 1$, 符合条件

三. 计算题

$$1. P = \frac{C_5^2 + C_5^1 C_2^1 C_4^1 C_3^1}{C_{10}^4} = \frac{13}{21}$$

- (1)解: 有两个实根的情况下, 则 $p^2 - 4q \geq 0$, 由几何概型可得 $p = \frac{13}{24}$ (2) 在 (1)

的条件下, 多出 $q \geq 0$, 由几何概型解出 $p = \frac{1}{48}$

- 解: 设第一段长为 x , 第二段长为 y , 第三段为 $L-x-y$
由三角形三边关系可得

$$x+y \geq 1-x-y, x+L-x-y > y, y+L-x-y > x$$

$$\text{即} \begin{cases} x+y > \frac{l}{2} \\ x < \frac{l}{2} \\ y < \frac{l}{2} \end{cases}$$

由几何概型得 $p=0.25$

$$4. \text{解: } p = \frac{C_{18}^1 \cdot C_{17}^1 \cdot C_2^1 + C_{18}^1 \cdot C_2^1 \cdot C_1^1}{C_{20}^3} = 0.089$$

$$5. \text{解: } p = \frac{C_{96}^3}{C_{100}^3} \cdot 0.01^3 + \frac{C_{96}^2 \cdot C_4^1}{C_{100}^3} \cdot 0.01^2 \cdot 0.95^1 + \frac{C_{96}^1 \cdot C_4^2}{C_{100}^3} \cdot 0.01^1 \cdot 0.95^2 + \frac{C_4^3}{C_{100}^3} \cdot 0.95^3 = 0.8629$$

- 设事件 $A(X=k)$ 表示甲取 k 时大于乙数, 事件 T 为甲取得的数大于乙取得的数, 那么

$$P(T) = \frac{1}{5} [P(A(X=2)) + P(A(X=4)) + P(A(X=6)) + P(A(X=8)) + P(A(X=10))] \\ = \frac{1}{5} \left[\frac{1}{5} + \frac{2}{5} + \frac{3}{5} + \frac{4}{5} + \frac{5}{5} \right] = \frac{15}{25} = 0.6$$

- 设击沉潜水艇为事件 A , 击中导弹数目为 x , 则

$$P\{X=0\} = \frac{1}{6^4} = \frac{1}{1296}, P\{X=1\} = C_4^1 \frac{5}{6} \frac{1}{6^3} = \frac{20}{6^4} = \frac{5}{324}, \text{ 则}$$

$$P(A) = 1 - P(\bar{A}) = 1 - [P\{X=0\} + \frac{3}{5} P\{X=1\}] = \frac{1283}{1296}, \text{ (这里为什么是 } \frac{3}{5} \text{ 而不是 } \frac{1}{2} \text{ 是在已经击中的情况下, 击伤的概率为 } \frac{3}{5} \text{)}$$

8.

(1)

$$\text{已知 } P(A|B) > P(A|\bar{B}), \text{ 则 } \frac{P(AB)}{P(B)} > \frac{P(A\bar{B})}{P(\bar{B})} = \frac{P(A) - P(AB)}{1 - P(B)} \Rightarrow P(AB) > P(A)P(B)$$

要证

$$P(B|A) > P(B|\bar{A}) \Rightarrow \frac{P(AB)}{P(A)} > \frac{P(B\bar{A})}{P(\bar{A})} = \frac{P(B) - P(AB)}{1 - P(A)} \Rightarrow \text{即证 } P(AB) > P(A)P(B)$$

证毕。

$$(2) \quad P(A|B) = P(A|\bar{B}) \Leftrightarrow \frac{P(AB)}{P(B)} = \frac{P(A\bar{B})}{P(\bar{B})} = \frac{P(A) - P(AB)}{1 - P(B)}$$

$$\Leftrightarrow P(AB) = P(A)P(B) \Leftrightarrow A, B \text{ 独立}$$

第二章

习题 2.1

1.

(1) 引入随机变量，使我们可以研究一个随机试验中的所有可能结果，特别是随机事件数有可列个数非常多或连续取值以至于无限时。

(2) 若随机变量 x 所取的可能值是有限个或无限个可列个，则 x 称为离散型随机变量；若随机变量 x 所取的可能值不能逐个列举出来，则 x 为非离散型随机变量（连续型离散变量仅为其中一部分）。

$$2. \quad P\{X = k\} = q^{k-1}p, k = 1, 2, \dots, \text{其中 } q = 1 - p$$

3.

(1)

X	0	1	2	3	4
P	0.4	0.24	0.144	0.0864	0.1296

$$(2) \quad P(X \leq 2) = 0.4 + 0.24 + 0.144 = 0.784$$

4. $a = 0.1$

X	1	2	3	4
P	0.1	0.2	0.3	0.4

习题 2.2

1.

(1) 为了更好地掌握某区间内的概率分布情况。

(2)

定义： x 是一个随机变量， x 是任意实数，函数

$$F(x) = P\{X \leq x\}, (-\infty < x < \infty)$$

称为随机变量 x 的分布函数。

此时要将 x 放在整个数轴上进行讨论。(特点自己感受吧...)

(3) 性质:

[1]有界性: $0 \leq F(x) \leq 1$

[2] $\lim_{x \rightarrow -\infty} F(x) = F(-\infty) = 0, \lim_{x \rightarrow +\infty} F(x) = F(+\infty) = 1$

[3]单调性: $F(x)$ 是单调不减函数。

[4] $F(x)$ 是右连续的。

[5]对每个 a , $P\{X=a\}=F(a)-F(a-0)$

$$(4) F(x) = \sum P\{X \leq x\}$$

2. 是(2)区间的分布函数(由题目1中(3)中分布函数的四个性质判断)

3.

$$(1) F(x) = \begin{cases} 0, & x < 0 \\ \frac{1}{3}, & 0 \leq x < 1 \\ \frac{1}{2}, & 1 \leq x < 2 \\ 1, & x \geq 2 \end{cases}$$

$$(2) P\{X \leq \frac{1}{2}\} = \frac{1}{3}$$

$$(3) P\{\frac{1}{2} < X \leq \frac{3}{2}\} = \frac{1}{6}$$

$$(4) P\{1 \leq X \leq 2\} = \frac{2}{3}$$

习题 2.3

1. (1) 略

(2) n 是随机试验的总次数, p 是单次发生某事件的概率

2.

设至少需要 n 次, 记每次抽出 0 的概率为 p , 则

$$P = 0.1$$

记事件 A 为 0 至少出现一次

$$A \sim B(n, p), P(A) = C_n^1 p(1-p)^{n-1} \geq 0.9$$

$$\text{解得 } n \geq 22$$

$$3. P(\text{至少两次击中目标}) = 1 - C_{5000}^1 (0.001)(1 - 0.001)^{5000-1} - C_{5000}^0 (1 - 0.001)^{5000} = 0.95964$$

4.

$$(1) (\text{属于超几何分布, 通式 } P\{X = k\} = \frac{C_4^k C_{16}^{6-k}}{C_{20}^6}, k = 0, 1, 2, 3, 4)$$

X	0	1	2	3	4
P	0.2066	0.4508	0.2817	0.0578	0.0031

(2) (属于二项式分布, 通式为 $P\{X = k\} = C_{20}^k 0.2^k \times 0.8^{20-k}, k = 0, 1, 2, 3, 4, 5, 6$)

X	0	1	2	3	4	5	6
P	0.2621	0.3932	0.2458	0.089	0.0154	0.0015	0.0001

5. 用二项分布计算

$$P = C_{300}^4 (0.01)^4 (1 - 0.01)^{300-4} = 0.168031355$$

用 • 泊松分布计算

$$\lambda = np = 3$$

$$P = \frac{3^4 e^{-3}}{4!} = 0.168031355$$

相对误差: 5%

$$6. P = C_5^3 (0.3)^3 (1 - 0.3)^2 + C_5^4 (0.3)^4 (1 - 0.3)^1 + C_5^5 (0.3)^5 = 0.16308$$

习题 2.4

1.

(1) 定义:

对于随机变量 X 的分布函数 $F(x)$, 存在非负函数 $f(x)$, 使对于任意实数 x , 有

$$F(x) = \int_{-\infty}^x f(t) dt$$

则称 X 为连续型随机变量, 其中 $f(x)$ 为 X 的概率密度函数。

几何意义: 分布函数的值表示概率密度函数图像中小于等于该值的左侧包裹的面积图像 (说白了就是小于等于该值的概率, 非说成几何就是这样), 密度函数相当于概率密度函数图像中等于该值长度为 1 矩形的面积

(2) 性质: (P63 下—P64 上)

$$[1] f(x) \geq 0$$

$$[2] \int_{-\infty}^{+\infty} f(x) dx = 1$$

$$[3] \text{ 对于任意实数 } x_1, x_2 (x_1 \leq x_2), P\{x_1 < X \leq x_2\} = F(x_2) - F(x_1) = \int_{x_1}^{x_2} f(x) dx$$

$$[4] \text{ 若 } f(x) \text{ 在点 } x \text{ 处连续, 则有 } F'(x) = f(x)$$

[5] 对于连续型随机变量 X 来说, 因为其分布函数 $F(x)$ 是连续的, 所以

$$P\{X=a\} = F(a) - F(a-0) = 0$$

(3) 均匀分布特点: x 落在平均分布的区间内的任意等长度的子区间内的可能性是相同的。

(4) 指数分布特点: 无记忆性

应用背景: 作为某些等待时间的概率分布。

无记忆性：如果一个随机变量呈指数分布

当 $s, t \geq 0$ 时有 $P(T > s + t | T > t) = P(T > s)$

即，如果 T 是某一元件的寿命，已知元件使用了 t 小时，它总共使用至少 $s + t$ 小时的条件概率，与从开始使用时算起它使用至少 s 小时的概率相等。

(5) 正态分布：可见 P68 页下。

标准正态分布： $\mu = 0$ ， $\sigma = 1$ ，特点不详叙

标准正态分布是正态分布的一个特殊情况。

(6) 3σ 原则为

数值分布在 $(\mu - \sigma, \mu + \sigma)$ 中的概率为 0.6827

数值分布在 $(\mu - 2\sigma, \mu + 2\sigma)$ 中的概率为 0.9545

数值分布在 $(\mu - 3\sigma, \mu + 3\sigma)$ 中的概率为 0.9973

可以认为， Y 的取值几乎全部集中在 $(\mu - 3\sigma, \mu + 3\sigma)$ 区间内，超出这个范围的可能性仅占不到 0.3%。

(7) 标准正态分布函数分位点：

设随机变量 $X \sim N(0, 1)$ ，若

$$P(X > u_a) = a$$

则称 u_a 为此正态分布的上 a 分位数

即

$$1 - \Phi(u_a) = a$$

$$\Phi(u_a) = 1 - a$$

2.

$$(1) \int_0^{+\infty} A e^{-3x} dx = 1, \text{ 解得 } A=3$$

$$(2) P\{X > 0.1\} = \int_{0.1}^{+\infty} 3e^{-3x} dx = 0.7480$$

$$(3) F(x) = \begin{cases} 0 & (x < 0) \\ 1 - e^{-3x} & (x \geq 0) \end{cases}$$

3.

(1) 由: $\lim_{x \rightarrow +\infty} F(x) = 1$ 和 $\lim_{x \rightarrow +0} F(x) = 0$ 解得 **A=1, B=-1**

$$(2) f(x) = \begin{cases} 0 & (x \leq 0) \\ xe^{-\frac{x^2}{2}} & (x > 0) \end{cases}$$

$$(3) P=F(2)-F(1)=0.4712$$

4. 设汽车刚开走是在 t_0 时刻, 下一辆车到达时刻为 t_0+5 , 乘客到达汽车站的时刻假定为 x , x 服从 (t_0, t_0+5) 的均匀分布, 则有:

$$F(x) = \begin{cases} \frac{1}{5}(t_0 < x < t_0 + 5) \\ 0, \text{其他} \end{cases}$$

$$P\{t_0+2 < X \leq t_0 + 5\} = \int_{t_0+2}^{t_0+5} \frac{1}{5} dt = 0.6$$

5. 设三个都不能使用 1000h 以上为事件 A, 至少有一个能使用 1000h 以上为事件 B

$$P(B) = 1 - P(A) = 1 - \left(\int_0^{1000} f(x) dx\right)^3 = 0.638$$

6. 先写出概率密度函数,

$$f(x) = \frac{1}{2\sqrt{2\pi}} e^{-\frac{(x-1)^2}{2 \times 4}}, -\infty < x < \infty$$

$$F(5) = \int_{-\infty}^5 f(x) dx = 0.9772$$

$$P\{0 < X \leq 1.6\} = \int_0^{1.6} f(x) dx = 0.3049$$

$$P\{|X - 1| \leq 2\} = P\{-1 \leq X \leq 3\} = \int_{-1}^3 f(x) dx = 0.3049$$

7. 设 x 为获奖分数线, 列出下列公式,

$$P\{X \geq x\} = P\left\{\frac{X-65}{10} \geq \frac{x-65}{10}\right\} = 1 - \Phi\left(\frac{x-65}{10}\right) = 0.1$$

查表解得

$$x = 78$$

8.

$$(1) P\{X < 89\} = P\left\{\frac{X-90}{0.5} < -2\right\} = \Phi(-2) = 0.02275$$

(2) 要求

$$P\{X \geq 80\} \geq 0.99 \Rightarrow P\{X < 80\} < 0.01$$

则

$$P\left\{\frac{X-d}{0.5} < 2(80-d)\right\} < 0.01$$

查表可以解得, $d \geq 81.1635$

习题 2.5

1.

(1) $Y = Y_i$ 时, 得到所有满足 $Y_i = g(X)$ 的 X 集合 (不妨设为 $X_1, X_2 \dots X_n$), 那么

$$P\{Y = Y_i\} = \sum_{i=1}^n P\{X = X_i\}$$

(2)

法一: 利用下面的公式先求 Y 的分布函数

$$F_Y(Y \leq y) = F_X(g(x) \leq y),$$

然后求导得密度函数 $f_Y(y)$

法二: 若 $g(x)$ 是严格单调函数, X 具有概率密度 $f_X(x)$, $-\infty < x < \infty$ 又设 $g(x)$ 处处可导且有 $g'(x) > 0$ (或恒有 $g'(x) < 0$), 则 $Y = g(X)$ 是连续型随机变量, 其概率密度为:

$$f_Y(y) = \begin{cases} f_X(h(y)) |h'(y)|, & \alpha < y < \beta \\ 0, & \text{其他} \end{cases}$$

其中 $\alpha = \min\{g(-\infty), g(\infty)\}$, $\beta = \max\{g(-\infty), g(\infty)\}$, $h(y)$ 是 $g(x)$ 的反函数。

(3) 服从, 详见 P79 页下例 2.5.6

2.

Y	0	1	4	9
P	1/5	7/30	1/5	11/30

3. 用上面 T1 (2) 中方法二得:

$$\text{当 } c > 0 \text{ 时, } f_Y(y) = \begin{cases} \frac{1}{c(b-a)}, & ca + d \leq y \leq cb + d, \\ 0, & \text{其他} \end{cases}$$

$$\text{当 } c < 0 \text{ 时, } f_Y(y) = \begin{cases} -\frac{1}{c(b-a)}, & cb + d \leq y \leq ca + d, \\ 0, & \text{其他} \end{cases}$$

4. 先写出,

$$f_X(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$$

(1) 取了绝对值, 实际上就是相当于正的部分乘了 2 倍, 即

$$f_Y(y) = \begin{cases} \sqrt{\frac{2}{\pi}} e^{-\frac{y^2}{2}}, & y > 0 \\ 0, & \text{其他} \end{cases}$$

(2) 利用公式,

$$f_Y(y) = \begin{cases} f_X(h(y))|h'(y)| = \frac{1}{\sqrt{2\pi}} e^{-\frac{(\ln y)^2}{2}} \cdot \frac{1}{y}, & 0 < y < +\infty \\ 0, & \text{其他} \end{cases}$$

$$5. F_Y(y) = P\{Y < y\} = P\{1-2|X| < y\} = P\{X < -\frac{1-y}{2} \text{ 或 } X > \frac{1-y}{2}\} =$$

$$\frac{1}{\sqrt{2\pi}} \left[\int_{-\infty}^{\frac{y-1}{2}} e^{-\frac{x^2}{2}} dx + \int_{\frac{1-y}{2}}^{\infty} e^{-\frac{x^2}{2}} dx \right], \text{ 求导, 得}$$

$$f_Y(y) = \frac{1}{\sqrt{2\pi}} \left[\frac{1}{2} e^{-\frac{(\frac{y-1}{2})^2}{2}} + \frac{1}{2} e^{-\frac{(\frac{y-1}{2})^2}{2}} \right]$$

$$f_Y(y) = \begin{cases} \frac{1}{\sqrt{2\pi}} e^{-\frac{(y-1)^2}{8}}, & y \leq 1 \\ 0, & y > 1 \end{cases}$$

$$6. X' = -\frac{1}{Y^2}, \text{ 在 } (-\infty, 0), (0, \infty) \text{ 上单调, 所以可以套用公式:}$$

$$f_Y(y) = \begin{cases} \frac{1}{\pi(1+\frac{1}{y^2})} \frac{1}{y^2} = \frac{1}{\pi(1+y^2)}, & y < 0 \\ \text{当 } x \rightarrow \infty, f_Y(y) = 0, \text{ 不妨填 } \frac{1}{\pi(1+0^2)}, & y = 0 \\ \frac{1}{\pi(1+\frac{1}{y^2})} \frac{1}{y^2} = \frac{1}{\pi(1+y^2)}, & y > 0 \end{cases}$$

(以下解释请慎信, 完全为了按照答案解释, 有更好说法欢迎私戳) $y=0$ 时, 这个地方本应该是 0, 但是概率密度函数趴, 少一两个点, 换一两个点也没影响, 所以不如让整个表

$$\text{达式看起来连续, 所以 } f_Y(y) = \frac{1}{\pi(1+y^2)}, -\infty < y < \infty$$

习题二

一、填空题

$$1. \text{ 直接套吧, } F(x_0) - F(x_0 - 0)$$

$$2. F(1) - F(0) = \frac{1}{4}$$

$$3. \text{ 先列出式子,}$$

$$P\{\frac{1}{2} < x < \frac{5}{2}\} = P\{x=1, 2\} = \frac{3}{4}A$$

又因为

$$A(\frac{1}{2} + \frac{1}{4} + \frac{1}{8} + \frac{1}{16}) = 1$$

解得

$$A = \frac{16}{15}$$

$$\text{所以 } P\{\frac{1}{2} < x < \frac{5}{2}\} = P\{x=1, 2\} = 0.8$$

4. F(x)单调不减，函数 F(x)右连续，且 F(-∞)=0, F(+∞)=1

5. 不变。 $g(\sigma) = P\{|X - a| < \sigma\} = P\{|\frac{X-a}{\sigma}| < 1\} = \Phi(1) - \Phi(-1)$, 与 σ 无关

6. 0.5。(因为整个图像关于 x=1 对称)

二、计算题

1.

X	0	1	2
P	$\frac{C_{13}^3}{C_{15}^3} = \frac{22}{35}$	$\frac{C_2^1 C_{13}^2}{C_{15}^3} = \frac{12}{35}$	$\frac{C_2^2 C_{13}^1}{C_{15}^3} = \frac{1}{35}$

2.

X	1	2	3
P	0.6	$0.4 \times 0.6 = 0.24$	$1 - 0.6 - 0.24 = 0.16$

3.

$$(1) C_5^2 (0.1)^2 (0.9)^3 = 0.0729$$

$$(2) C_5^3 (0.1)^3 (0.9)^2 + C_5^4 (0.1)^4 (0.9)^1 + (0.1)^5 = 0.00856$$

$$(3) 1 - C_5^4 (0.1)^4 (0.9)^1 - (0.1)^5 = 0.99954$$

$$(4) 1 - (0.9)^5 = 0.40951$$

$$4. P\{X=k\} = (\frac{1}{4})^{k-1} \times \frac{3}{4}, (k = 1, 2, 3, \dots)$$

$$5. \text{用泊松分布计算, 则 } \lambda = np=3, P = \frac{3^4 e^{-3}}{4!} = 0.168031355$$

6. 无实根则

$$16X^2 - 16(X+2) < 0 \Rightarrow \frac{1}{2} < X < \frac{7}{2}$$

又由题得:

$$f(x) = \begin{cases} e^{-x}, & x > 0 \\ 0, & \text{其他} \end{cases}$$

$$P = \int_{1/2}^{7/2} f(x) dx = 1 - e^{-2}$$

7. 有题目可以列出方程式,

$$\int_2^3 f(x) dx = P = 2 \int_1^2 f(x) dx$$

$$\text{且 } P = \int_1^3 f(x) dx = 1$$

解得

$$\begin{cases} A = \frac{1}{3} \\ B = \frac{1}{6} \end{cases}$$

8.

(1) 证明:

$$F(-a) = \int_{-\infty}^{-a} f(x) dx = \int_{\infty}^a f(-x) d(-x) = \int_a^{+\infty} f(x) dx = 1 - F(a)$$

$$F(-a) = \int_a^{+\infty} f(x) dx = \int_0^{+\infty} f(x) dx - \int_0^a f(x) dx = \frac{1}{2} - \int_0^a f(x) dx$$

$$(2) P\{|X| > a\} = 2P\{X > a\} = 2[1 - F(a)]$$

$$(3) P\{|X| < a\} = 1 - 2[1 - F(a)] = 2F(a) - 1$$

9. 先求出 $a = \frac{1}{9}$, 然后对应于 $Y = 2(X - 2)^2$ 列表,

X	0	2	8	18
P	$\frac{1}{3}$	$\frac{1}{4}$	$\frac{11}{36}$	$\frac{1}{9}$

10. 因为 $I = \sqrt{\frac{W}{2}}, I' = \frac{1}{2\sqrt{2W}}$ 单调, 且

$$f_I(i) = \begin{cases} \frac{1}{2}, & 9 < i < 11 \\ 0, & \text{其他} \end{cases}$$

$$F_W(w) = \begin{cases} \frac{1}{2} \times \frac{1}{2\sqrt{2w}} = \frac{1}{4\sqrt{2w}}, & 162 < w < 242 \\ 0, & \text{其他} \end{cases}$$

$$11. y = x^3, f_Y(y) = \begin{cases} f_X(\sqrt[3]{y}) \frac{1}{3y^{\frac{2}{3}}} = \frac{1}{3ay^{\frac{2}{3}}}, & 0 < y < a^3 \\ 0, & \text{其他} \end{cases},$$

12. 合格率表达式为

$$P\{|X - \mu| \leq m\} = P\left\{\left|\frac{X - \mu}{\sigma}\right| \leq \frac{m}{\sigma}\right\} = \Phi\left(\frac{m}{\sigma}\right) - \Phi\left(-\frac{m}{\sigma}\right) = 2\Phi\left(\frac{m}{\sigma}\right) - 1 = 0.95$$

$$\Phi\left(\frac{m}{\sigma}\right) = 0.975 = \Phi(1.96) \Rightarrow m = 1.96\sigma$$

第三章

习题 3.1

1. 不是, 我们来看, 过了分界线 $x+2y=1$, 分布函数 $F(x, y)$ 就从 0 变为 1, 那么可以说,

所有的 x, y 都满足 $x+2y=1$, 那么这就不叫二元随机变量了, 实质上只有一个随机变量, 所以不是二维随机变量的联合分布函数。

2. 证明: 他是不是二维随机变量的概率密度函数, 首先我们确认他是不是概率密度函数, 证明条件需要 (参考 P90 的 4 个性质): (1) 非负性; (2) 归一性; (3) 对概率的计算; (4) 连续性

(1) 首先非负性很好确认, 不加详述。

$$(2) \text{ 归一性: } \iint_D f(x, y) dx dy = \iint_D \frac{2g(\sqrt{x^2+y^2})}{\pi\sqrt{x^2+y^2}} dx dy = \frac{2}{\pi} \int_0^{\frac{\pi}{2}} d\theta \int_0^{\infty} \frac{g(r)}{r} r dr = 1$$

(3) 这点, 如果我们认为他是, 那他就会有这个性质, 无法证伪。

(4) 我们看 $g(x)$ 可以从 0 积到无穷远, 那么根据可积的判断条件 (连续或者有限个

第一类间断点), 判断 $g(x)$ 连续, 那么 $f(x, y)$ 也连续

所以可以是二维随机变量的概率密度函数。

(整个证明并不严谨, 从性质出发并不好, 欢迎有好办法私戳)

3. 如下

$Y \setminus X$	0	1	2	3
0	0	0	$\frac{C_3^2}{C_7^4} = \frac{3}{35}$	$\frac{C_2^1}{C_7^4} = \frac{2}{35}$
1	0	$\frac{C_3^1 C_2^1}{C_7^4} = \frac{6}{35}$	$\frac{C_3^2 C_2^1 C_2^1}{C_7^4} = \frac{12}{35}$	$\frac{C_2^1}{C_7^4} = \frac{2}{35}$
2	$1/C_7^4 = \frac{1}{35}$	$\frac{C_3^1 C_2^1}{C_7^4} = \frac{6}{35}$	$\frac{C_3^2}{C_7^4} = \frac{3}{35}$	0

4. $Y = |X - (3 - X)| = |2X - 3|$

$Y \setminus X$	0	1	2	3	P_Y
1	0	$\frac{3}{8}$	$\frac{3}{8}$	0	$\frac{3}{4}$

3	$\frac{1}{8}$	0	0	$\frac{1}{8}$	$\frac{1}{4}$
P_X	$\frac{1}{8}$	$\frac{3}{8}$	$\frac{3}{8}$	$\frac{1}{8}$	

$$5. \quad f_{Z=|X-Y|}(Z) = \begin{cases} \int_0^{1-z} f_{X,Y}(Z+Y, Y) dy = \int_0^{1-z} dy = 1-z, & X > Y \\ \int_0^{1-z} f_{X,Y}(X, Z+Y) dx = \int_0^{1-z} dx = 1-z, & X \leq Y \end{cases} = \begin{cases} 1-z, & 0 \leq z \leq 1 \\ 0, & \text{其他} \end{cases}$$

6. 首先计算这个正方形域的面积：边长为 a ，面积为 a^2 ，那

$$(1) \quad f(X, Y) = \frac{1}{a^2}, (|x| + |y| \leq \frac{a}{\sqrt{2}})$$

(2)

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy = \begin{cases} \int_{-\frac{a}{\sqrt{2}}-x}^{\frac{a}{\sqrt{2}}+x} f(x, y) dy = \frac{\sqrt{2}a+2x}{a^2}, & -\frac{a}{\sqrt{2}} \leq x < 0 \\ \int_{-\frac{a}{\sqrt{2}}+x}^{\frac{a}{\sqrt{2}}-x} f(x, y) dy = \frac{\sqrt{2}a-2x}{a^2}, & 0 \leq x \leq \frac{a}{\sqrt{2}} \\ 0, & \text{其他} \end{cases} = \begin{cases} \frac{\sqrt{2}a-2|x|}{a^2}, & |x| \leq \frac{a}{\sqrt{2}} \\ 0, & |x| > \frac{a}{\sqrt{2}} \end{cases}$$

$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx = \begin{cases} \frac{\sqrt{2}a-2|y|}{a^2}, & |y| \leq \frac{a}{\sqrt{2}} \\ 0, & |y| > \frac{a}{\sqrt{2}} \end{cases}$$

习题 3.2

1. 如下

(1) 横着顺着累加就可以了

	0	1	2	3
P_Y	$\frac{8}{27}$	$\frac{12}{27} = \frac{4}{9}$	$\frac{6}{27} = \frac{2}{9}$	$\frac{1}{27}$
	1	2	3	
P_X	$\frac{3}{27} = \frac{1}{9}$	$\frac{18}{27} = \frac{2}{3}$	$\frac{6}{27} = \frac{2}{9}$	

(2)

$X=1, Y$

$X=1, Y$	0	1	2	3
$P_{Y X=1}$	$\frac{2}{3}$	0	0	$\frac{1}{3}$

$Y=0, X$

$Y=0, X$	1	2	3
$P_{X Y=0}$	$\frac{1}{4}$	$\frac{3}{4}$	0

$$(3) \quad P\{X=3|Y=2\}=0, P\{Y=2|X=3\}=0$$

$$2. \quad f_x = \int_{-x}^x f(x, y) dy = 2x, 0 < x < 1, \quad f_y = \int_{|y|}^1 f(x, y) dx = 1 - |y|, -1 < y < 1. \quad \text{那么}$$

$$f_{Y|X}(y|x) = \frac{f(x, y)}{f_x(x)} = \begin{cases} \frac{1}{2x}, & -x < y < x, \quad 0 < x < 1 \\ 0, & \text{其他} \end{cases}$$

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_y(y)} = \begin{cases} \frac{1}{1-|y|}, & -x < y < x, \quad 0 < x < 1 \\ 0, & \text{其他} \end{cases}$$

$$3. \quad f_x(x) = \int_x^1 f(x, y) dy = \int_x^1 f_{X|Y}(x|y) f_y(y) dy = \int_x^1 15x^2 y dy = \frac{15}{2}(x^2 - x^4)$$

$$P(X > \frac{1}{2}) = \int_{\frac{1}{2}}^1 \frac{15}{2}(x^2 - x^4) dx = \frac{47}{64}$$

习题 3.3

1.

$Y \setminus X$	2	5	8	P_Y
0.4	0.15	0.3	0.35	0.8
0.8	0.05	0.12	0.03	0.2
P_X	0.2	0.42	0.38	

不独立，例子 $P(y=0.4, x=2) = 0.15 \neq P(y=0.4)P(x=2)$

2.

$$(1) \quad f(x, y) = \frac{\partial^2 F}{\partial x \partial y} = 0.0001e^{-0.01(x+y)}, f_x(x) = 0.01e^{-0.01x}, f_y(y) = 0.01e^{-0.01y}$$

$$f_X(x)f_Y(y)=f(x,y) \quad , \quad \text{相互独立}$$

$$(2) \quad P\{X > 120, Y > 120\} = 1 - F(120, 120) = 2e^{-1.2} - e^{-2.4} \quad (\text{答案错了})$$

3.

(1) 均匀分布，这又是一块面积，所以我们先来算面积， $S = 2 \int_0^1 \sqrt{x} dx = \frac{4}{3}$ ，那么

$$f(x, y) = \frac{1}{\frac{4}{3}} = \begin{cases} \frac{3}{4}, & 0 \leq x \leq 1, y^2 \leq x \\ 0, & \text{其他} \end{cases}$$

$$(2) \quad f_X(x) = \int_{-\sqrt{x}}^{\sqrt{x}} f(x, y) dy = \frac{3}{2} \sqrt{x}, 0 \leq x \leq 1$$

$$f_Y(y) = \int_{y^2}^1 f(x, y) dx = \frac{3}{4} (1 - y^2), -1 \leq y \leq 1$$

$$f(x, y) \neq f_X(x)f_Y(y), \quad \text{所以不独立}$$

$$(3) \quad P(X < \frac{1}{2}) = \int_0^{\frac{1}{2}} \frac{3}{2} \sqrt{x} dx = \frac{1}{\frac{3}{2}}, \quad P(Y < \frac{1}{2}) = \int_{-1}^{\frac{1}{2}} \frac{3}{4} (1 - y^2) dy = \frac{27}{32}$$

$$P(X < \frac{1}{2}, Y < \frac{1}{2}) = \frac{3}{4} S' = \frac{3}{4} \times \frac{1}{6} = \frac{1}{8}$$

(S' 指的是满足条件的面积，为 $2 \int_0^{\frac{1}{4}} \sqrt{x} dx$)

习题 3.4

1.

如下

$Z = X + Y$	2	3	4
$P_{Z=X+Y}$	$\frac{1}{4}$	$\frac{1}{2}$	$\frac{1}{4}$

2. 如下

$Z = \max\{X, Y\}$	0	1
$P_{Z=\max\{X+Y\}}$	$\frac{1}{4}$	$\frac{3}{4}$

$$3. \quad f_{Z=X+Y}(z) = \begin{cases} \int_0^z f_X(z-y)f_Y(y)dy = e^{-\frac{1}{3}z} - e^{-\frac{1}{2}z}, & z > 0 \\ 0, & z \leq 0 \end{cases}$$

4. $Z = \min\{X_1, X_2, X_3, X_4\}$ ，其中 X_1, X_2, X_3, X_4 相互独立，那么

$$P(Z \geq 180) = P^4(X_i \geq 180) = \left(\frac{1}{\sqrt{2\pi}} \int_{\frac{180-160}{\sqrt{202}}}^{\infty} e^{-\frac{x^2}{2}} dx \right)^4 = 0.00063$$

习题三

1.

$$(1) \iint f(x, y) dx dy = a \int_0^1 dx \int_0^2 (3x^2 + xy) dy = 3a = 1, a = \frac{1}{3}$$

(2) 当 $x < 0$ 或 $y < 0$ 时, $F(x, y) = 0$;

$$\text{当 } 0 \leq x \leq 1, 0 \leq y \leq 2 \quad F(X, Y) = \frac{1}{3} \int_0^x dx \int_0^y f(x, y) dx dy = \frac{x^3 y}{3} + \frac{1}{12} x^2 y^2$$

$$\text{当 } 0 \leq x \leq 1, y > 2 \text{ 时, } F(x, y) = \int_0^x dx \int_0^2 \frac{1}{3} (3x^2 + xy) dy = \frac{2}{3} x^3 + \frac{1}{3} x^2$$

$$\text{当 } x > 1, 0 \leq y \leq 2 \text{ 时, } F(x, y) = \int_0^y dy \int_0^1 \frac{1}{3} (3x^2 + xy) dx = \frac{y}{3} + \frac{1}{12} y^2$$

$$\text{当 } x > 1, y > 2 \text{ 时, } F(x, y) = \int_0^2 dy \int_0^1 \frac{1}{3} (3x^2 + xy) dx = 1$$

$$(3) \text{ 通过画图我们可以确定积分区域嘛 } P(X + Y \leq 1) = \int_0^1 dx \int_0^{1-x} \frac{1}{3} (3x^2 + xy) dy = \frac{7}{72}$$

$$P(X + Y \leq 2.3) = \int_0^{0.3} dx \int_0^2 f(x, y) dy + \int_{0.3}^1 dx \int_0^{2.3-x} f(x, y) dy = \frac{6}{125} + 0.6918 = 0.7398$$

2.

(1)

$$f(x, y) = \begin{cases} \frac{1}{(b-a)(d-c)}, & a < x < b, c < y < d \\ 0, & \text{其他} \end{cases}, f_X(x) = \begin{cases} \frac{1}{b-a}, & a < x < b \\ 0, & \text{其他} \end{cases}, f_Y(y) = \begin{cases} \frac{1}{d-c}, & c < y < d \\ 0, & \text{其他} \end{cases}$$

(2) 相互独立

3.

$$(1) f(x, y) = \frac{\partial^2 F(x, y)}{\partial x \partial y} = (\ln 3)^2 3^{-x-y}, x \geq 0, y \geq 0$$

$$f_X(x) = \begin{cases} \int_0^{\infty} f(x, y) dy = \ln 3 \times 3^{-x}, & x \geq 0 \\ 0, & x < 0 \end{cases} \quad \text{同理,}$$

$$f_Y(y) = \begin{cases} \int_0^{\infty} f(x, y) dx = \ln 3 \times 3^{-y}, & y \geq 0 \\ 0, & y < 0 \end{cases},$$

$$(2) 2F_X(x) = \int_{-\infty}^x f_X(x) dx = 1 - 3^{-x}, F_Y(y) = 1 - 3^{-y}, F(x, y) = F_X(x) \cdot F_Y(y), \text{ 独立}$$

$$4. \quad (1) \quad f_X = \begin{cases} \frac{1}{2}, 0 \leq x \leq 2 \\ 0, \text{其他} \end{cases}, f_Y = \begin{cases} \frac{1}{2}, 0 \leq y \leq 2 \\ 0, \text{其他} \end{cases}, f(x, y) = \begin{cases} \frac{1}{4}, 0 \leq x \leq 2, 0 \leq y \leq 2 \\ 0, \text{其他} \end{cases}$$

(2) 因为本来就是均匀分布，就是一块面积，所以我们只要算符合要求的面积即可

$$P\{X+Y \leq \frac{3}{2}\} = \frac{\frac{1}{2} \times \frac{3}{2} \times \frac{3}{2}}{4} = \frac{9}{32}$$

5. 我们可以列表格

(1) (先把已知为0的填进去，再把剩下的通过计算填入)

$Y \backslash X$	-1	0	1	P_Y
0	$\frac{1}{4}$	0	$\frac{1}{4}$	$\frac{1}{2}$
1	0	$\frac{1}{2}$	0	$\frac{1}{2}$
P_X	$\frac{1}{4}$	$\frac{1}{2}$	$\frac{1}{4}$	

(2) 不独立，举个例子， $P(X=-1, Y=0) = \frac{1}{4} \neq P(X=-1)P(Y=0)$

6. 如右图

(1) 首先， $f(x, y) = \begin{cases} 1, y < x < 2-y, y > 0 \\ 0, \text{其他} \end{cases}$ 那么，

$$f_X(x) = \int f(x, y) dy = \begin{cases} x, 0 < x < 1 \\ 2-x, 1 \leq x < 2 \\ 0, \text{其他} \end{cases}$$

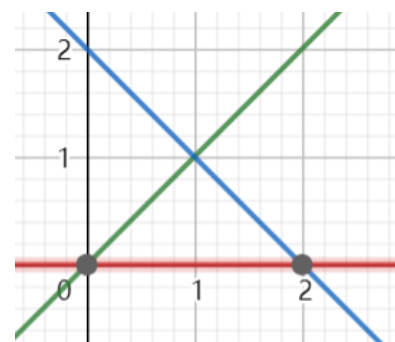
$$f_Y(y) = \int f(x, y) dx = \begin{cases} 2-y-y = 2-2y, 0 \leq y \leq 1 \\ 0, \text{其他} \end{cases}$$

$$(2) \quad f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)} = \begin{cases} \frac{1}{2-2y}, y < x < 2-y, 0 \leq y \leq 1 \\ 0, \text{其他} \end{cases}$$

$$f_{Y|X}(y|x=1.5) = \begin{cases} \frac{f(x=1.5, y)}{f_X(x=1.5)} = 2, 0 < y < \frac{1}{2} \\ 0, \text{其他} \end{cases}$$

(3) 不独立，因为 $f(x, y) \neq f_X(x)f_Y(y)$

$$(4) \quad P(0.1 < Y \leq 0.4 | x=1.5) = \int_{0.1}^{0.4} f_{Y|X}(y|x=1.5) dy = 0.6$$



$$(5) \quad F_{X|Y}(x|y) = \int_{-\infty}^x f_{X|Y}(x|y)dx = \begin{cases} 0, & x < y \\ \frac{x-y}{2-2y}, & y \leq x < 2-y \\ 1, & x \geq 2-y \end{cases}$$

7. 首先我们要确定 A 、 B 、 C 的值, $F(\infty, \infty) = A(B + \frac{\pi}{2})(C + \frac{\pi}{2}) = 1$

$$F(-\infty, y) = A(B - \frac{\pi}{2})(C + \arctan \frac{y}{3}) = 0 \quad F(x, -\infty) = A(B + \arctan \frac{x}{2})(C - \frac{\pi}{2}) = 0,$$

$$\text{可以解出来 } B = \frac{\pi}{2}, C = \frac{\pi}{2}, A = \frac{1}{\pi^2}$$

$$(1) \quad f(x, y) = \frac{\partial^2 F(x, y)}{\partial x \partial y} = \frac{1}{\pi^2} \frac{\frac{1}{2}}{1 + \frac{x^2}{4}} \frac{\frac{1}{3}}{1 + \frac{y^2}{9}} = \frac{6}{\pi^2(4+x^2)(9+y^2)}$$

$$(2) \quad f_X(x) = \int_{-\infty}^{\infty} f(x, y)dy = \frac{2}{\pi(4+x^2)}, \text{同理, } f_Y(y) = \frac{3}{\pi(9+y^2)}$$

8. (1) $\iint f(x, y)dxdy = 1, \frac{A}{2} = 1, A = 2$

$$(2) \quad F(X, Y) = \int_{-\infty}^y dy \int_{-\infty}^x f(x, y)dx = \begin{cases} (1 - e^{-2x})(1 - e^{-y}), & x > 0, y > 0 \\ 0, & \text{其他} \end{cases}$$

(3) 令 $Z = X - Y$,

$$f_{Z=X-Y}(z) = \int_{-\infty}^{\infty} f(y+z, y)dy = \int_{\max\{-z, 0\}}^{\infty} 2e^{-2z-3y}dy = \frac{2}{3}e^{-2z-3\max\{0, -z\}} = \begin{cases} \frac{2}{3}e^z, & z < 0 \\ \frac{2}{3}e^{-2z}, & z \geq 0 \end{cases}$$

$$P(X \leq Y) = P(Z \leq 0) = \int_{-\infty}^0 f(z)dz = \frac{2}{3} \quad (\text{也可以直接})$$

$$\int_0^{\infty} dy \int_0^y 2e^{-2x-y}dx = \frac{2}{3})$$

9.

$$(1) \quad \iint_D f(x, y)dxdy = 1, 8k = 1, k = \frac{1}{8}$$

$$(2) \quad P(X < 1, Y < 3) = \int_0^1 dx \int_2^3 \frac{1}{8}(6-x-y)dy = \frac{3}{8}$$

$$(3) \quad P\{X < 1.5\} = \int_0^{1.5} dx \int_2^4 f(x, y)dy = \frac{27}{32}$$

(4) (这里没有通过求 $Z = X + Y$ 的方法是因为实测有点麻烦, 所以直接算吧)

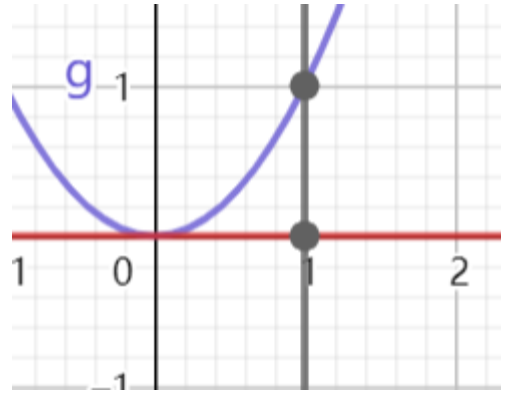
$$P\{X + Y \leq 4\} = \int_0^2 dx \int_2^{4-x} \frac{1}{8}(6-x-y)dy = \frac{2}{3}$$

10.

$$(1) \quad f(x, y) = \begin{cases} \frac{1}{2}(e^{-\frac{y}{2}}), & 0 < x < 1, y > 0 \\ 0, & \text{其他} \end{cases}$$

(2) 即 $P(4X^2 - 4Y \geq 0) = P(X^2 \geq Y)$, 如图, 那

$$\text{么 } P\{X^2 \geq Y\} = \int_0^1 dx \int_0^{\sqrt{x}} \frac{1}{2} e^{-\frac{y}{2}} dy = 12e^{-\frac{1}{2}} - 7 \approx 0.278$$



11.

$$(1) \quad \iint_D C(R - \sqrt{x^2 + y^2}) dx dy = 1, \frac{R^3 \pi}{3} C = 1, C = \frac{3}{\pi R^3}$$

$$(2) \quad P\{X^2 + Y^2 \leq r^2\} = \iint_{D'} f(x, y) dx dy = 3 \frac{r^2}{R^2} - 2 \frac{r^3}{R^3}$$

12. (题目印刷有问题, $f_Y(y) \rightarrow f(x, y)$) $f_X(x) = \begin{cases} \int_{-1}^1 \frac{1}{4}(1+xy) dy = \frac{1}{2}, & |x| < 1 \\ 0, & |x| \geq 1 \end{cases}$ 同

理 $f_Y(y) = \begin{cases} \int_{-1}^1 \frac{1}{4}(1+xy) dx = \frac{1}{2}, & |y| < 1 \\ 0, & |y| \geq 1 \end{cases}$ $f(x, y) \neq f_X(x)f_Y(y)$, 所以 X 与 Y 不独立,

$$F_{U=X^2}(U \leq u) = F_X(-\sqrt{u} \leq X \leq \sqrt{u}) = \begin{cases} 0, & u < 0 \\ \frac{1+\sqrt{u}}{2} - \frac{1-\sqrt{u}}{2} = \sqrt{u}, & 0 \leq u \leq 1 \text{ 同理,} \\ 1, & u > 1 \end{cases}$$

$$F_{V=Y^2}(V \leq v) = F_Y(-\sqrt{v} \leq Y \leq \sqrt{v}) = \begin{cases} 0, & v < 0 \\ \frac{1+\sqrt{v}}{2} - \frac{1-\sqrt{v}}{2} = \sqrt{v}, & 0 \leq v \leq 1 \text{ 那么} \\ 1, & v > 1 \end{cases}$$

$$F_{U=X^2, V=Y^2}(U \leq u, V \leq v) = F_{X,Y}(-\sqrt{u} \leq X \leq \sqrt{u}, -\sqrt{v} \leq Y \leq \sqrt{v})$$

$$= \begin{cases} 0, u \leq 0 \text{ 或 } v \leq 0 \\ \int_{-\sqrt{u}}^{\sqrt{u}} dx \int_{-\sqrt{v}}^{\sqrt{v}} f(x, y) dy = \sqrt{uv}, 0 < u < 1, 0 < v < 1 \\ \int_{-1}^1 dx \int_{-\sqrt{v}}^{\sqrt{v}} f(x, y) dy = \sqrt{v}, u \geq 1, 0 < v < 1 \\ \sqrt{u}, v \geq 1, 0 < u < 1 \\ 1, v \geq 1, u \geq 1 \end{cases}, \text{ 因为 } F_{U,V}(u, v) = F_U(u)F_V(v),$$

所以 X^2, Y^2 独立。

$$13. f_{Z=X+2Y}(z) = \int_{-\infty}^{\infty} f(z-2y, y) dy = \begin{cases} \int_0^{\frac{z}{2}} 2e^{-z} dy = ze^{-z}, & z > 0 \\ 0, & z \leq 0 \end{cases}, \text{ 那么}$$

$$F(Z) = \begin{cases} 1 - e^{-z} - ze^{-z}, & z > 0 \\ 0, & z \leq 0 \end{cases}$$

$$14. \text{ 咱们先写出联合分布, } f(x, y) = \begin{cases} 2, & 0 < x < 1, 0 < y < x \\ 0, & \text{其他} \end{cases}$$

$$(1) f_{Z=XY}(z) = \int_{-\infty}^{\infty} \frac{1}{|y|} f\left(\frac{z}{y}, y\right) dy = \begin{cases} \int_z^{\sqrt{z}} 2 \frac{1}{y} dy = -\ln z, & 0 < z < 1 \\ 0, & \text{其他} \end{cases}$$

$$(2) f_{Z=\frac{Y}{X}}(z) = \int_{-\infty}^{\infty} |x| f(x, xz) dx = \begin{cases} \int_z^1 2x dx = 1, & 0 < z < 1 \\ 0, & \text{其他} \end{cases}$$

15. 这里我们首先要注意到, X, Y 并不独立 ($P\{X \geq 0, Y \geq 0\} \neq P\{X \geq 0\}P\{Y \geq 0\}$)

$P\{\max\{(X, Y) \geq 0\} = 1 - P\{X < 0, Y < 0\}$ 我们可以列表格

($\frac{a}{b}$ 是原来有的数据, a/b 是计算出来的)

$X \setminus Y$	$Y < 0$	$Y \geq 0$	P_X
$X < 0$	$\frac{2}{7}$	$\frac{1}{7}$	$\frac{3}{7}$
$X \geq 0$	$\frac{1}{7}$	$\frac{3}{7}$	$\frac{4}{7}$
P_Y	$\frac{3}{7}$	$\frac{4}{7}$	

$$\text{所以 } P\{\max(X, Y) \geq 0\} = 1 - \frac{2}{7} = \frac{5}{7}$$

16. 直接选用答案给的表

$X \setminus Y$	1	2	3	P_X
1	$\frac{1}{9}$	0	0	$\frac{1}{9}$
2	$2 \times \frac{1}{3} \times \frac{1}{3} = \frac{2}{9}$	$\frac{1}{9}$	0	$\frac{3}{9}$
3	$2 \times \frac{1}{3} \times \frac{1}{3} = \frac{2}{9}$	$2 \times \frac{1}{3} \times \frac{1}{3} = \frac{2}{9}$	$\frac{1}{9}$	$\frac{5}{9}$
P_Y	$\frac{5}{9}$	$\frac{3}{9}$	$\frac{1}{9}$	

$$P\{\xi = \eta\} = \frac{1}{9} + \frac{1}{9} + \frac{1}{9} = \frac{1}{3}$$

17.

(1) 这里主要需要画图, $P\{Y > 0 | Y > X\} = \frac{3}{4}$

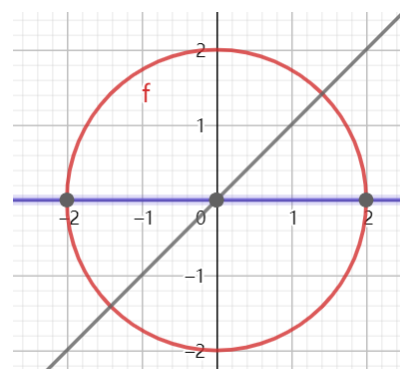
(2) 从图中我们可以看出, 满足条件的就是第一、二、四象限, 则 $P\{M > 0\} = \frac{1}{4}$

18.

$$(1) C_n^m p^m (1-p)^{n-m}$$

$$(2) P\{x=n\} = \frac{\lambda^n e^{-\lambda}}{n!}, P\{y=m | x=n\} = C_n^m p^m (1-p)^{n-m}, 0 \leq m \leq n, n=0, 1, 2. \text{ 则}$$

$$P\{x=n, y=m\} = P\{y=m | x=n\} P\{x=n\} = C_n^m p^m (1-p)^{n-m} \frac{\lambda^n e^{-\lambda}}{n!}, 0 \leq m \leq n, n=0, 1, 2, \dots$$



$$19. f_{Z=\frac{X}{Y}}(z) = |y| \int_{\max\{1000, \frac{1000}{z}\}}^{\infty} f_X(zy) f_Y(y) dy = \frac{10^6}{2z^2} \frac{1}{(\max\{1000, \frac{1000}{z}\})^2} = \begin{cases} \frac{1}{2z^2}, & z \geq 1 \\ \frac{1}{2}, & 0 < z < 1 \\ 0, & \text{其他} \end{cases}$$

20.

$$(1) P\{X=2 | Y=2\} = \frac{0.05}{P\{Y=2\}} = \frac{1}{5} \quad P\{Y=3 | X=0\} = \frac{1}{3}$$

(2) 这个就一个一个列出来趴

$$P\{V = 0\} = 0$$

$$P\{V = 1\} = 0.01 + 0.01 + 0.02 = 0.04$$

$$P\{V = 2\} = 0.01 + 0.03 + 0.05 + 0.04 + 0.03 = 0.16$$

$$P\{V = 3\} = 0.01 + 0.02 + 0.04 + \dots + 0.05 = 0.28$$

$$p\{V = 4\} = 0.24$$

$$P\{V = 5\} = 0.28$$

(3) 同上,

$$P\{U = 0\} = (\text{第一列和第一行总和}) = 0.28$$

$$P\{U = 1\} = (\text{除去第一行第一列的第二行和第二列的总和}) = 0.3$$

$$P\{U = 2\} = (\text{除去上面的第三行和第三列的总和}) = 0.25$$

$$P\{U = 3\} = 0.17$$

(4) 还是直接找就行啦

$$P\{W = 0\} = 0$$

$$P\{W = 1\} = 0.02$$

$$P\{W = 2\} = 0.06$$

$$P\{W = 3\} = 0.13$$

$$P\{W = 4\} = 0.19$$

$$P\{W = 5\} = 0.24$$

$$P\{W = 6\} = 0.19$$

$$P\{W = 7\} = 0.12$$

$$P\{W = 8\} = 0.05$$

第四章

习题 4.1

1. $\frac{1}{2} \quad \frac{5}{4} \quad 4$

$$E(X) = (-1) \times \frac{1}{8} + 0 \times \frac{1}{2} + 1 \times \frac{1}{8} + 2 \times \frac{1}{4} = \frac{1}{2}$$

$$E(X^2) = (-1)^2 \times \frac{1}{8} + 0 \times \frac{1}{2} + 1^2 \times \frac{1}{8} + 2^2 \times \frac{1}{4} = \frac{5}{4}$$

$$E(2X + 3) = 2E(X) + 3 = 4$$

2. 4.125

易知, $3 \leq X \leq 5$,

X	3	4	5
P	$\left(\frac{1}{2}\right)^3 \times 2 = \frac{1}{4}$	$C_3^2 \left(\frac{1}{2}\right)^2 \times \frac{1}{2} \times \frac{1}{2} \times 2 = \frac{3}{8}$	$C_4^2 \left(\frac{1}{2}\right)^4 = \frac{3}{8}$
对 P 的注释	A、B 都可能连胜三场	前三场有一个人赢了两场 $C_3^2 \left(\frac{1}{2}\right)^2 \times \frac{1}{2}$ ，第四场这个人也要赢，再乘 $\frac{1}{2}$ ，而 A、B 都可能是这个人，所以再乘 2	前四场每个人各赢了 2 场，而第五场不能结果都会结束

$$E(X) = 3 \times \frac{1}{4} + 4 \times \frac{3}{8} + 5 \times \frac{3}{8} = 4.125$$

3. 0.4、0.1、0.5

$$p_1 + p_2 + p_3 = 1$$

可以列出三个方程， $E(X) = -p_1 + p_3 = 0.1$

$$E(X^2) = p_1 + p_3 = 0.9$$

可以解得， $p_1 = 0.4, p_2 = 0.1, p_3 = 0.5$

4. 答案

首先设 x 为故障的天数

X	0	1	2	≥ 3
P(X)	$0.8^5 = 0.32768$	$C_5^1 0.2 \times 0.8^4 = 0.4096$	$C_5^2 0.2^2 \times 0.8^3 = 0.2048$	$1 - 0.32768 - 0.4096 - 0.2048 = 0.05792$

$$E(X) = P(X=0) \times 10 + P(X=1) \times 5 - P(X \geq 3) \times 2 = 5.20896$$

5. 答案

(1)

X	0	1	2	3 (不遇到红灯)
P	$\frac{1}{2}$	$\frac{1}{2} \times \frac{1}{2} = \frac{1}{4}$	$\frac{1}{2} \times \frac{1}{2} \times \frac{1}{2} = \frac{1}{8}$	$\frac{1}{2} \times \frac{1}{2} \times \frac{1}{2} = \frac{1}{8}$

$$(2) \quad E\left(\frac{1}{1+X}\right) = \frac{1}{2} \times 1 + \frac{1}{4} \times \frac{1}{2} + \frac{1}{8} \times \frac{1}{3} + \frac{1}{8} \times \frac{1}{4} = \frac{67}{96}$$

6.

$$(1) \quad E(U) = E(2X + 3Y + 1) = 2E(X) + 3E(Y) + 1 = 10 + 33 + 1 = 44$$

$$(2) E(V) = E(YZ - 4X) = E(YZ) - 4E(X) = E(Y)E(Z) - 4E(X) = 88 - 20 = 68$$

7. 独立

$$\Rightarrow E(XY) = E(X)E(Y)$$

$$E(X) = \int_0^1 x \cdot 2x dx = \frac{2}{3}$$

$$E(Y) = \int_5^\infty y \cdot e^{-(y-5)} dy = 6$$

$$E(XY) = 4$$

8. 证明:

$$E(X) = \sum_{k=0}^{\infty} kP(X=k)$$

$$P(X \geq k) = \sum_{j=k}^{\infty} P(X=j)$$

$$\sum_{k=1}^{\infty} P(X \geq k) = \sum_{k=1}^{\infty} \sum_{j=k}^{\infty} P(X=j) = \sum_{k=0}^{\infty} kP(X=k) = E(X)$$

(因为这里在 $P(X=k)$ 时一共出现了 k 次, 所以上述等式成立)

习题 4.2

1. 扫码给了 0.301、0.322, 不准确

X	0	1	2	3
P	$\frac{9}{12} = \frac{3}{4}$	$\frac{3}{12} \times \frac{9}{11} = \frac{9}{44}$	$\frac{3}{12} \times \frac{2}{11} \times \frac{9}{10} = \frac{9}{220}$	$\frac{3}{12} \times \frac{2}{11} \times \frac{1}{10} = \frac{1}{220}$

$$E(X) = \frac{9}{44} \times 1 + \frac{9}{220} \times 2 + \frac{1}{220} \times 3 = \frac{3}{10}$$

$$E(X^2) = \frac{9}{44} \times 1 + \frac{9}{220} \times 4 + \frac{1}{220} \times 9 = \frac{9}{22}$$

$$D(X) = E(X^2) - E(X)^2 = \frac{9}{22} - \frac{9}{100} = \frac{351}{1100}$$

2. 如下,

$$D(XY) = E(X^2Y^2) - E(XY)^2 = E(X^2)E(Y^2) - E(X)^2E(Y)^2 = [D(X) + E(X)^2][D(Y) + E(Y)^2]$$

$$- E(X)^2E(Y)^2 = D(X)D(Y) + D(X)E(Y)^2 + D(Y)E(X)^2$$

3. 如下, $E(X)=0, D(X)=4, E(Y)=2, D(Y)=\frac{4}{3}$

$$E(XY) = E(X)E(Y) = 0$$

$$D(X+Y) = D(X) + D(Y) = \frac{16}{3}$$

$$D(2X-3Y) = 4D(X) + 9D(Y) = 28$$

4. $f(x)$ 的概率密度为 $f(x) = 1, -\frac{1}{2} < x < \frac{1}{2}$

$$E(Y) = \int_0^{\frac{1}{2}} g(x) f(x) dx = \int_0^{\frac{1}{2}} \ln x dx = -\frac{1}{2} \ln 2 - \frac{1}{2}$$

$$E(Y^2) = \int_0^{\frac{1}{2}} \ln^2 x dx = \frac{1}{2} \ln^2 2 + \ln 2 + 1$$

$$D(Y) = E(Y^2) - E(Y)^2 = \frac{1}{4} \ln^2 2 + \frac{1}{2} \ln 2 + \frac{3}{4}$$

5. 如下, $E(X)$ 不存在, $D(X)$ 不存在

若存在, $E(X) = \int_{-\infty}^{\infty} \frac{x}{\pi(1+x^2)} dx$, 不绝对收敛, 所以不存在; 同理 $D(X)$ 不存在

(这里我们复习下连续随机变量的数学期望定义:)

设连续型随机变量 X 的概率密度为 $f(x)$, 若积分

$$\int_{-\infty}^{\infty} xf(x) dx$$

绝对收敛, 则称积分 $\int_{-\infty}^{\infty} xf(x) dx$ 的值为随机变量 X 的数学期望, 记为 $E(X)$.

6. (扫码答案不对) T_i 满足 $f_{T_i}(t) = \frac{1}{5} e^{-\frac{t}{5}}, t > 0$

$$f_{T=T_1+T_2}(t) = \int_{-\infty}^{\infty} f_{T_1}(t-t_2) f_{T_2}(t_2) dt_2 = \int_0^t \frac{1}{25} e^{-\frac{x+t-x}{5}} dx = \frac{1}{25} t e^{-\frac{t}{5}}, t \geq 0$$

$$E(T) = \int_0^{\infty} \frac{1}{25} t^2 e^{-\frac{1}{5}t} dt = 10, D(T) = E(T^2) - E(T)^2 = 150 - 100 = 50$$

假设发生事件 A 的概率为 p , 则

X	0	1
P	1-P	P

$$E(X) = P, D(X) = E(X^2) - E(X)^2 = P(1-P) \leq \frac{1}{4}$$

习题 4.3

$$\begin{aligned}
 1. \quad & Cov(3X - 2Y + 1, X + 4Y - 3) = Cov(3X - 2Y, X + 4Y - 3) + Cov(1, X + 4Y - 3) \\
 & = Cov(3X - 2Y, X + 4Y) + Cov(3X - 2Y, -3) + Cov(1, X + 4Y) + Cov(1, -3) \\
 & = Cov(3X, X) + Cov(3X, 4Y) + Cov(-2Y, X) + Cov(-2Y, 4Y) + Cov(3X, -3) + Cov(-2Y, -3) \\
 & \quad + Cov(1, X) + Cov(1, 4Y) + Cov(1, -3) \\
 & = 3D(X) + 10Cov(X, Y) - 8D(Y) - 8Cov(X, 1) + 10Cov(Y, 1) - 3Cov(1, 1)
 \end{aligned}$$

已知: $Cov(X, Y) = E(XY) - E(X)E(Y) = -1,$

$D(X) = 2, D(Y) = 3, Cov(X, 1) = E(X) - E(X)E(1) = 0 = Cov(Y, 1), Cov(1, 1) = 0$

所以上述表达式 $\Rightarrow 3 \times 2 + 10 \times (-1) - 8 \times 3 = -28$

2.

$P(X = k) = C_n^k \left(\frac{1}{2}\right)^n, \quad E(X) = E(Y) = \frac{n}{2} \quad (\text{概率相同})$

$E(XY) = E(X(n - X)) = nE(X) - E(X^2),$

其中,

$$\begin{aligned}
 E(X^2) &= \sum_{i=0}^n i^2 C_n^i \left(\frac{1}{2}\right)^n = \sum_{i=1}^n i^2 C_n^i \left(\frac{1}{2}\right)^n \\
 &= \sum_{i=1}^n i^2 \frac{n}{i} C_{n-1}^{i-1} \left(\frac{1}{2}\right)^n = \sum_{i=1}^n i n C_{n-1}^{i-1} \left(\frac{1}{2}\right)^n = \sum_{i=2}^n (i-1) n C_{n-1}^{i-1} \left(\frac{1}{2}\right)^n + \sum_{i=1}^n n C_{n-1}^{i-1} \left(\frac{1}{2}\right)^n \\
 &= \sum_{i=2}^n n(n-1) C_{n-2}^{i-2} \left(\frac{1}{2}\right)^n + \sum_{i=1}^n n C_{n-1}^{i-1} \left(\frac{1}{2}\right)^n = n(n-1) 2^{n-2} \left(\frac{1}{2}\right)^n + n 2^{n-1} \left(\frac{1}{2}\right)^n \\
 &= \frac{n(n-1)}{4} + \frac{n}{2} = \frac{n(n+1)}{4}
 \end{aligned}$$

所以, $E(XY) = \frac{n^2}{2} - \frac{n(n+1)}{4} = \frac{n^2 - n}{4}, \quad D(X) = E(X^2) - E(X)^2 = \frac{n(n+1)}{4} - \frac{n^2}{4} = \frac{n}{4} = D(Y)$

$$\rho_{XY} = \frac{Cov(X, Y)}{\sqrt{D(X)}\sqrt{D(Y)}} = \frac{E(XY) - E(X)E(Y)}{D(X)} = \frac{\frac{n^2 - n}{4} - \frac{n^2}{4}}{\frac{n}{4}} = -1$$

3. 如下

(1) 事件 A、B 独立 $\Leftrightarrow P(AB) = P(A)P(B) \Leftrightarrow \rho = 0$

(2) 即证明

$$P(AB) - P(A)P(B) \leq \sqrt{P(A)[1 - P(A)]P(B)[1 - P(B)]}$$

$$P(A)^2 P(B)^2 - 2P(A)P(B)P(AB) + P(AB)^2 \leq P(A)^2 P(B)^2 - P(A)P(B)^2 - P(A)^2 P(B) + P(A)P(B)$$

$$P(AB)^2 \leq [1 - P(B) - P(A) + 2P(AB)]P(A)P(B)$$

$$4. f(x) = \begin{cases} 0, & x \leq 0 \\ \frac{1}{2\pi}, & 0 < x < 2\pi, \\ 0, & x \geq 2\pi \end{cases}$$

$$Cov(Y, Z) = Cov(\sin X, \sin X \cos a + \cos X \sin a)$$

$$= \cos a D(\sin X) + \sin a Cov(\sin X, \cos X)$$

$$\text{其中, } D(Y) = D(\sin x) = E(\sin^2 x) - E(\sin x)^2 = \int_0^{2\pi} \frac{\sin^2 x}{2\pi} dx - \left(\int_0^{2\pi} \frac{\sin x}{2\pi} dx \right)^2 = \frac{1}{2}$$

$$D(Z) = D(\sin(x+a)) = E(\sin^2(x+a)) - E(\sin(x+a))^2 = \int_0^{2\pi} \frac{\sin^2(x+a)}{2\pi} dx - \left(\int_0^{2\pi} \frac{\sin(x+a)}{2\pi} dx \right)^2 = \frac{1}{2}$$

$$Cov(\sin X, \cos X) = E(\sin X \cos X) - E(\sin X)E(\cos X)$$

$$= \int_0^{2\pi} \frac{\sin x \cos x}{2\pi} dx - \int_0^{2\pi} \frac{\sin x}{2\pi} dx \int_0^{2\pi} \frac{\cos x}{2\pi} dx = 0$$

$$\text{所以, } Cov(Y, Z) = \frac{1}{2} \cos a, \rho_{YZ} = \frac{\frac{1}{2} \cos a}{\frac{1}{2}} = \cos a$$

当 $a = \frac{\pi}{2}, \frac{3\pi}{2}$ 时, Y、Z 不相关, 但也不独立; $a \neq \frac{\pi}{2}, \frac{3\pi}{2}$ 时, Y、Z 相关

$$5. \rho_{XY} = 0 \Rightarrow Cov(X, Y) = 0 \Rightarrow E(XY) = E(X)E(Y)$$

$$\text{而 } E(X) = P(A), E(Y) = P(B), E(AB) = P(AB)$$

$$\Rightarrow P(XY) = P(X)P(Y) \Rightarrow \text{相互独立}$$

6. 设这同一分布的数学期望为 $E(x)$, 方差为 $D(x)$, 则 $D(X) = nD(x), D(Y) = nD(x)$,

$$Cov(X, Y) = E(XY) - E(X)E(Y) = n^2 E(x^2) - n^2 E(x)^2 = n^2 D(x)$$

$$\rho_{XY} = \frac{\text{Cov}(X, Y)}{\sqrt{D(X)}\sqrt{D(Y)}} = n$$

习题 4.4

$$1. P(|X - \mu| \geq 3\sigma) \leq \frac{\sigma^2}{9\sigma^2} = \frac{1}{9}$$

$$2. \text{令 } U = X - Y, \text{ 则 } E(U) = E(X) - E(Y) = 0,$$

$$D(U) = D(X) + D(Y) - 2\text{Cov}(X, Y) = D(X) + D(Y) - 2\rho_{XY}\sqrt{D(X)}\sqrt{D(Y)} = 4 + 1 - 2 = 3$$

$$P(|U - \mu| \geq 6) \leq \frac{\sigma^2}{36} = \frac{1}{12}$$

$$3. \text{条件可以转换为 } \lim_{n \rightarrow \infty} D\left(\frac{1}{n} \sum_{i=1}^n X_i^2\right) = \lim_{n \rightarrow \infty} D(\bar{X}_n^2) = 0, \text{ 要证 } \{X_n\} \text{ 服从大数定理, 即}$$

$$\lim_{n \rightarrow \infty} P\{|\bar{X}_n - a_n| < \varepsilon\} = 1, \text{ 即 } \lim_{n \rightarrow \infty} \bar{X}_n - a_n = 0$$

$$4. \text{设总发生故障的设备台数为 } x, \text{ 则}$$

$$P(X \geq 2) = P\left(\frac{X - 400 \times 0.02}{\sqrt{400 \times 0.02 \times 0.98}} \geq \frac{2 - 8}{\sqrt{8 \times 0.16 \times 6}} = -\frac{5}{4}\sqrt{3}\right) = \varphi\left(x \geq -\frac{5}{4}\sqrt{3}\right) = \varphi\left(x \leq \frac{5}{4}\sqrt{3}\right)$$

$$= \Phi\left(\frac{5}{4}\sqrt{3}\right) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\frac{5}{4}\sqrt{3}} e^{-\frac{t^2}{2}} dt = 0.9845$$

$$5. \text{设应供 } A \text{ 瓦电力, 其中工作的机床数目为 } x, \text{ 那么要求}$$

$$P(X \leq A) = P\left(\frac{X - 200 \times 0.6}{\sqrt{200 \times 0.6 \times 0.4}} \leq \frac{A - 120}{4\sqrt{3}}\right) = \Phi\left(\frac{A - 120}{4\sqrt{3}}\right) = 99.9\%$$

$$\Phi\left(\frac{A - 120}{4\sqrt{3}}\right) = 99.9\% = \Phi(3.175), A = 142 \text{ (这里我是根据答案套的)}$$

$$6. \text{利用 Chebyshev 不等式, 设 } A \text{ 是抛掷的次数, } x \text{ 是正面朝上的次数, 则}$$

$$P(|X - 0.5A| \leq 0.1A) = 1 - \frac{\sigma^2}{(0.1A)^2} \geq 0.9, \quad \sigma^2 = A \times \frac{1}{2} \times \frac{1}{2} \leq 10^{-3} A^2, A \geq 250$$

利用中心极限定理, 那么

$$P(0.4A \leq X \leq 0.6A) = P\left(-\frac{1}{5}\sqrt{A} \leq \frac{X - 0.5A}{0.5\sqrt{A}} \leq \frac{1}{5}\sqrt{A}\right) = \Phi\left(\frac{1}{5}\sqrt{A}\right) - \Phi\left(-\frac{1}{5}\sqrt{A}\right) = 2\Phi\left(\frac{1}{5}\sqrt{A}\right) - 1 \geq 0.9$$

$$\Phi\left(\frac{1}{5}\sqrt{A}\right) \geq 0.95 = \Phi(1.65), A \geq 68 \quad (\text{由答案靠近})$$

习题四

一. 填空题

$$1. D(2X - Y) = 4D(X) + D(Y) = 36$$

$$2. \text{标准差为 } \sqrt{np(1-p)}, \quad p = \frac{1}{2} \text{ 时, 标准差最大, 为 } 5$$

$$3. E[(X-1)(X-2)] = E(X^2 - 3X + 2) = E(X^2) - 3E(X) + 2 = D(X) + E(X)^2 - 3\lambda + 2$$

$$\lambda + \lambda^2 - 3\lambda + 2 = 1 \quad \lambda = 1$$

二. 计算题

1.

$$(1) \quad P = \iint_D f(x, y) dx dy = \iint_D f(x) f(y) dx dy = \iint_D \frac{1}{8\pi} e^{-\frac{x^2+y^2}{8}} dx dy = e^{-\frac{1}{8}} - e^{-\frac{1}{2}}$$

$$(2) \quad E(Z) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \sqrt{x^2 + y^2} \frac{e^{-\frac{x^2+y^2}{8}}}{8\pi} dx dy = \sqrt{2\pi}$$

2.

$$(1) \quad E(|X - Y|) = \int_0^1 \int_0^1 |x - y| f(x) f(y) dx dy = \int_0^1 \int_0^1 |x - y| dx dy = \frac{1}{3}$$

$$(2) \quad E(|X - Y|) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} |x - y| f(x) f(y) dx dy$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} |x - y| \frac{1}{2\pi} e^{-\frac{x^2+y^2}{2}} dx dy = 2 \iint_{x \geq y} (x - y) \frac{1}{2\pi} e^{-\frac{x^2+y^2}{2}} dx dy$$

$$\frac{1}{\pi} \int_0^{\infty} r e^{-\frac{r^2}{2}} r dr \int_{\frac{3}{4}\pi}^{\frac{1}{4}\pi} (\cos \theta - \sin \theta) d\theta = \frac{2}{\sqrt{\pi}}$$

3.

$$(1) \quad E(X) = -(a + 0.1) + 0 + (0.4 + c) = 0.2, a - c = 0.1$$

$$P(Y \leq 0 | X \leq 0) = \frac{P(Y = 0, X = 0) + P(Y = -1, X = 0) + P(Y = 0, X = -1) + P(Y = -1, X = -1)}{P(X = 0) + P(X = -1)}$$

$$= \frac{a + b + 0.1}{a + b + 0.2} = 0.5, a + b = 0, \quad (\text{不理解, 这里和答案有出入}) \text{ 又因为}$$

$$a+b+c+0.6=1, a+b+c=0.4, \text{ 由上述三式得, } a=0.2, b=0.1, c=0.1$$

(2)

Z	-2	-1	0	1	2
P	0.2	0.1	0.3	0.3	0.1

$$(3) \quad P(X=Z)=P(Y=0)=0.4$$

$$4. \quad p=P(X>\frac{\pi}{3})=\int_{\frac{\pi}{3}}^{\pi}\frac{1}{2}\cos\frac{x}{2}dx=\frac{1}{2},$$

Y	0	1	2	3	4
P	$(1-p)^4$	$C_4^1 p(1-p)^3$	$C_4^2 p^2(1-p)^2$	$C_4^3 p^3(1-p)$	p^4

$$E(Y^2)=C_4^1 p(1-p)^3+4\times C_4^2 p^2(1-p)^2+9\times C_4^3 p^3(1-p)+16\times p^4=5$$

5.

X	1	2	...	n
P	p	$(1-p)p$		$(1-p)^{n-1}p$

$$E(X)=\sum_{k=1}^{\infty}k(1-p)^{k-1}p=-p[\sum_{k=1}^{\infty}(1-p)^k]'=-p(\frac{1}{p}-1)'=\frac{1}{p},$$

$$D(X)=E(X^2)-E(X)^2=\sum_{k=1}^{\infty}k^2(1-p)^{k-1}p-\frac{1}{p^2}=-p[\sum_{k=1}^{\infty}k(1-p)^k]'-\frac{1}{p^2}$$

$$-p[\sum_{k=1}^{\infty}(k+1)(1-p)^k-(1-p)^k]'-\frac{1}{p^2}=-p\{-[\sum_{k=1}^{\infty}(1-p)^{k+1}]-\sum_{k=1}^{\infty}(1-p)^k\}'-\frac{1}{p^2}=\frac{1-p}{p^2}$$

6. 我们可以假设, 工人修完某一台机, 这台机器是第 k 个 (在因为发生故障概率相同, 所以在哪一台机器是均匀分布的), 那么走向下一台机器的距离为

$$E(Z_k)=\begin{cases} \sum_{i=2}^n(i-1)\frac{1}{n}=\frac{n-1}{2}, k=1 \\ \sum_{i=1}^{k-1}(k-i)\frac{1}{n}+\sum_{i=k+1}^n(i-k)\frac{1}{n}=\frac{k^2}{n}-\frac{n+1}{n}k+\frac{n+1}{2}, 1<k<n=\frac{k^2}{n}-\frac{n+1}{n}k+\frac{n+1}{2}, 1\leq k\leq n \\ \sum_{i=1}^{n-1}(n-i)\frac{1}{n}=\frac{n-1}{2}, k=n \end{cases}$$

$$E(Z) = E(E(Z_k))a = \frac{1}{n} \left[\sum_{k=1}^n \frac{k^2}{n} - \frac{n+1}{n}k + \frac{n+1}{2} \right] a = \frac{1}{3} \left(n - \frac{1}{n} \right) a$$

7.

(1)

$$P(X \geq 96) = P\left(\frac{X-72}{\sigma} \geq \frac{24}{\sigma}\right) = 1 - \Phi\left(\frac{24}{\sigma}\right) = 0.023, \Phi\left(\frac{24}{\sigma}\right) = 0.977 = \Phi(2.00), \sigma = 12$$

$$P(60 \leq X \leq 84) = P\left(-1 \leq \frac{X-72}{12} \leq 1\right) = \Phi(1) - \Phi(-1) = \frac{1}{\sqrt{2\pi}} \int_{-1}^1 e^{-\frac{t^2}{2}} dt = 0.6826$$

$$P(Y = K) = C_{100}^k 0.6826^k \times 0.3174^{100-k}, k = 0, 1, \dots, 100$$

$$(2) \quad E(Y) = np = 68.26, D(X) = np(1-p) = 21.6657$$

8.

$$(1) \quad E(X) = \int_{-\infty}^0 \frac{1}{2} x e^x dx + \int_0^{+\infty} \frac{1}{2} x e^{-x} dx = 0 \quad (\text{也因为 } f(x) \text{ 关于 } y \text{ 轴对称})$$

$$D(X) = E(X^2) - 0 = \int_{-\infty}^0 \frac{1}{2} x^2 e^x dx + \int_0^{+\infty} \frac{1}{2} x^2 e^{-x} dx = 2$$

$$(2) \quad \text{Cov}(X, |X|) = E(X|X|) - E(X)E(|X|) = E(X|X|)$$

令 $Z = X|X|$, 那么 $Z = \begin{cases} -x^2, & x < 0 \\ x^2, & x \geq 0 \end{cases}$, Z 与 x 是单调的, 可以套用 P78 的公式

$$f_Z(z) = \begin{cases} f_X(-\sqrt{-z}) \frac{1}{2\sqrt{-z}} = \frac{1}{4\sqrt{-z}} e^{-\sqrt{-z}} & z < 0 \\ f_X(\sqrt{z}) \frac{1}{2\sqrt{z}} = \frac{1}{4\sqrt{z}} e^{-\sqrt{z}} & z \geq 0 \end{cases} \quad (\text{对称分布}), \text{ 那么}$$

$$\text{Cov}(X, |X|) = E(Z) = 0, \text{ 不相关。}$$

$$(3) \quad \text{首先我们得知道 } f_{Y=|X|}(y), \text{ 易得 } f_{Y=|X|}(y) = \begin{cases} e^{-y}, & y \geq 0 \\ 0, & y < 0 \end{cases} \text{ 所以}$$

$$f_X(x)f_{|X|}(|x|) = \frac{1}{2} e^{-2|x|} \neq f_{X|X|}(x|x|), \text{ 不独立}$$

(写在最后: 在本题里, X 、 $|X|$ 显然是有关联的, 所以 $Z = X|X|$ 用的是一个 x , 但是 X 、 $|X|$ 的分布却不同, 所以造成上述结果)

$$9. \quad \text{Cov}(X, Y) = E(XY) - E(X)E(Y) = \int_{-\pi}^{\pi} \sin \theta \cos \theta d\theta - \int_{-\pi}^{\pi} \sin \theta d\theta \times \int_{-\pi}^{\pi} \cos \theta d\theta = 0, \text{ 不}$$

相关

$$E(XY) = 0 = E(X)E(Y), \text{ 不独立}$$

10.

(1) 如下

X+Y	0	1	2	3
P	0.10	0.40	0.35	0.15

$$(2) E(Z) = 0 + \sin \frac{\pi}{2} \times 0.4 + \sin \pi \times 0.35 + \sin \frac{3}{2} \pi \times 0.15 = 0.25$$

(3)

X	0	1	2
P	0.25	0.45	0.3
Y	0	1	
P	0.5	0.5	

$$E(X) = 1.05, D(X) = E(X^2) - E(X)^2 = 1.65 - 1.05^2 = 0.5475$$

$$E(Y) = 0.5, D(Y) = E(Y^2) - E(Y)^2 = 0.5 - 0.5^2 = 0.25$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y) = 0.2 \times 1 + 2 \times 0.15 - (1 \times 0.45 + 2 \times 0.3) \times (1 \times 0.5) = -0.025$$

$$\rho_{XY} = \frac{\text{Cov}(X, Y)}{\sqrt{D(X)D(Y)}} = -0.0676$$

11.

$$(1) E(Z) = \frac{1}{3}E(X) + \frac{1}{2}E(Y) = \frac{5}{6}, \text{ (答案给 } \frac{1}{3} \text{)}$$

$$D(Z) = \frac{1}{9}D(X) + \frac{1}{4}D(Y) + 2 \times \frac{1}{3} \times \frac{1}{2} \text{Cov}(X, Y) = 1 + 4 + \frac{1}{3} \rho_{XY} \sqrt{D(X)D(Y)} = 5 + \frac{1}{3} \times (-\frac{1}{2}) \times 3 \times 4 = 3$$

(2)

$$\text{Cov}(X, Z) = E(XZ) - E(X)E(Z) = E[X(\frac{X}{3} + \frac{Y}{2})] - \frac{5}{6} = \frac{E(X^2)}{3} + \frac{E(XY)}{2} - \frac{5}{6}$$

$$= \frac{D(X) + E(X)^2}{3} + \frac{\text{Cov}(X, Y) + E(X)E(Y)}{2} - \frac{5}{6} = 0, \rho_{XZ} = 0$$

$$(3) E(XZ) = \frac{5}{6} = E(X)E(Z), \text{ 相互独立}$$

12.

$$(1) E(X) = \int_{-\infty}^{\infty} xf(x)dx = \frac{8}{3}a + \frac{56}{3}b + 6c = 2$$

$$\int_{-\infty}^{\infty} f(x)dx = 2a + 6b + 2c = 1$$

$$P(1 < X < 3) = \int_1^3 f(x)dx = \frac{3}{2}a + \frac{5}{2}b + c = \frac{3}{4}$$

解得, $a = \frac{1}{4}, b = -\frac{1}{4}, c = 1$

$$(2) \quad D(X) = E(X^2) - E(X)^2 = \int_0^4 x^2 f(x)dx - 4 = \frac{2}{3}$$

$$(3) \quad E(Y) = \int_0^4 e^x f(x)dx = \frac{(e^2 - 1)^2}{4},$$

$$D(Y) = E(Y^2) - E(Y)^2 = \int_0^4 e^{2x} f(x)dx - E(Y)^2 = \frac{(e^4 + 1)^2}{16} - \frac{(e^2 - 1)^4}{16} = \frac{1}{4}e^2(e^2 - 1)^2$$

13.

(1)

(X_1, X_2)	(0,0)	(0,1)	(1,0)
P	0.1	0.1	0.8

$$(2) \quad D(X_1) = p(1-p) = 0.8 \times 0.2 = 0.16, D(X_2) = p(1-p) = 0.09$$

$$\text{Cov}(X_1, X_2) = E(X_1 X_2) - E(X_1)E(X_2) = -0.8 \times 0.1 = -0.08$$

$$\rho_{X_1 X_2} = \frac{-0.08}{0.4 \times 0.3} = -\frac{2}{3}$$

14. (完全不会, 请教)

$$15. \quad f_X(x) = \int_0^2 f(x, y)dy = \frac{x+1}{4}, \quad E(X) = \int_0^2 x f_X(x)dx = \frac{7}{6},$$

$$D(X) = E(X^2) - E(X)^2 = \int_0^2 x^2 f_X(x)dx - \frac{49}{36} = \frac{11}{36}, \text{同理,}$$

$$f_Y(y) = \frac{y+1}{4}, E(Y) = \frac{7}{6}, D(Y) = \frac{11}{36}$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y) = \int_0^2 dx \int_0^2 \frac{1}{8} xy(x+y)dy - E(X)^2 = -\frac{1}{36},$$

$$\rho_{XY} = \frac{\text{Cov}(X, Y)}{\sqrt{D(X)}\sqrt{D(Y)}} = \frac{-\frac{1}{36}}{\frac{11}{36}} = -\frac{1}{11}$$

$$D(X+Y) = D(X) + D(Y) + 2\text{Cov}(X, Y) = \frac{20}{36} = \frac{5}{9}$$

16.

(1) 都服从 $f(x) = e^{-x}$ 分布

$$F_X(x) = \int_0^x e^{-x} dx = 1 - e^{-x}, \quad F_V(v) = 1 - [1 - F_X(v)][1 - F_Y(v)] = 1 - e^{-2v},$$

$$f_V(v) = (F_V(v))' = 2e^{-2v} \quad v > 0, \quad f_V(v) = 0 \quad v \leq 0$$

$$(2) \quad F_U(u) = F_X(u)F_Y(u) = (1 - e^{-u})^2, \quad f_U(u) = 2e^{-u} - 2e^{-2u}, u > 0$$

$$E(U + V) = E(U) + E(V) = \int_{-\infty}^{\infty} uf_U(u)du + \int_{-\infty}^{\infty} vf_V(v)dv = 2$$

17. 设治愈人数为 X , 则

$$(1) \quad P(X > 75) = P\left(\frac{X - 80}{\sqrt{100 \times 0.8 \times 0.2}} > \frac{75 - 80}{4}\right) = \Phi(1.25) = 0.8944$$

$$(2) \quad P(X > 75) = P\left(\frac{X - 70}{\sqrt{100 \times 0.7 \times 0.3}} > \frac{75 - 70}{\sqrt{21}}\right) = 1 - \Phi\left(\frac{5}{\sqrt{21}}\right) = 0.1399$$

18.

$$(1) \quad P(X = K) = C_{100}^k 0.2^k \times 0.8^{100-k}, k = 0, 1, 2, \dots, 100$$

$$(2) \quad P(14 \leq X \leq 30) = P\left(-\frac{6}{4} \leq \frac{X - 0.2 \times 100}{\sqrt{0.2 \times 100 \times 0.8}} \leq \frac{10}{4}\right) = \Phi(2.5) - \Phi(-1.5) = 0.927$$

19. 设要有 A 条外线才能满足, x 为需要使用外线通话的电话, 则

$$P(X \leq A) = P\left(\frac{X - 200 \times 0.05}{\sqrt{200 \times 0.05 \times 0.95}} \leq \frac{A - 10}{\sqrt{9.5}}\right) = \Phi\left(\frac{A - 10}{\sqrt{9.5}}\right) = 0.9 = \Phi(1.3), A = 14$$

$$20. \quad P(T > 350) = P\left(\frac{T - 30 \times 10}{\sqrt{30} \times \sqrt{10^2}} > \frac{50}{10\sqrt{30}}\right) = 1 - \Phi\left(\frac{5}{\sqrt{30}}\right) = 0.18 \quad (\text{答案 } 0.1814, \text{ 计算机})$$

不强, 谅解谅解)

21. 如下

(1) 设 x 为相加的误差综合,

$$P(-15 \leq X \leq 15) = P\left(-\frac{3}{\sqrt{5}} \leq \frac{X - 0}{\sqrt{1500} \times \sqrt{\frac{1}{12}}} \leq \frac{3}{\sqrt{5}}\right) = \Phi\left(\frac{3}{\sqrt{5}}\right) - \Phi\left(-\frac{3}{\sqrt{5}}\right) = \frac{1}{\sqrt{2\pi}} \int_{-\frac{3}{\sqrt{5}}}^{\frac{3}{\sqrt{5}}} e^{-\frac{t^2}{2}} dt = 0.82$$

$$\text{那么, } P(|X| > 15) = 1 - 0.82 = 0.18$$

(2) 设最多可有 A 个数相加,

$$P(-10 < X < 10) = P\left(-10\sqrt{\frac{12}{A}} < \frac{X}{\sqrt{\frac{A}{12}}} < 10\sqrt{\frac{12}{A}}\right) = 2\Phi\left(10\sqrt{\frac{12}{A}}\right) - 1 = 0.9$$

$$\Phi\left(10\sqrt{\frac{12}{A}}\right) = 0.95 = \Phi(1.65), A = 443$$

第五章

习题 5.1

$$1. \quad \bar{X} = \frac{1}{n} \sum_{i=1}^n X_i = 3.59, S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2 = 2.881$$

2. 因为 μ 未知, 所以 (2) (5) 并不能通过统计数据可得, 所以不是统计量, 其他四个都是

$$3. \quad f(x_1, x_2, x_3, x_4) = \prod_{i=1}^4 \frac{1}{\theta^4} e^{-\frac{1}{\theta} \sum_{i=1}^4 x_i}, x_i > 0; f(x_1, x_2, x_3, x_4) = 0, x_i \leq 0$$

$$4. \quad E(\bar{X}) = \mu = \frac{a+b}{2}, D(\bar{X}) = \frac{\sigma^2}{n} = \frac{(b-a)^2}{12n}$$

$$5. \quad F_{(x)}^* = \begin{cases} 0, & x \leq 1 \\ \frac{1}{8}, & 1 < x \leq 2 \\ \frac{3}{8}, & 2 < x \leq 3 \\ \frac{6}{8}, & 3 < x \leq 4 \\ \frac{7}{8}, & 4 < x \leq 5 \\ 1, & x > 5 \end{cases}$$

6. 如下

	[335,340)	[340,345)	[345,350)	[350,355)	[355,360)
计数	3	9	22	14	2

图就不画了

习题 5.2

1.

(1) 查表

(2) 查表

2.

$$(1) \quad \frac{X_1 - X_2}{\sqrt{X_3^2 + X_4^2}} \sim \frac{N(0,2)}{\sqrt{\chi^2(2)}} = \frac{N(0,1)}{\sqrt{\frac{\chi^2(2)}{2}}} = t(2)$$

$$(2) \frac{\sqrt{n-1}X_1}{\sqrt{\sum_{i=2}^n X_i^2}} \sim \frac{N(0,1)}{\sqrt{\frac{\chi^2(n-1)}{n-1}}} = t(n-1)$$

$$(3) \frac{(n-3)\sum_{i=1}^3 X_i^2}{3\sum_{i=4}^n X_i^2} \sim \frac{(n-3)\chi^2(3)}{3\chi^2(n-3)} = F(3, n-3)$$

$$3. \text{ 查表, } E(\chi^2(n)) = n, D(\chi^2(n)) = 2n$$

$$4. P(\sum_{i=1}^{10} X_i^2 > 1.44) = P(\chi^2(10) > \frac{1.44}{0.09}) = 16) = 0.1$$

$$5. 2X_1 - X_2 \sim N(0, 10), 3X_3 + 4X_4 \sim N(0, 50), Y = 10a\chi_a^2(1) + 50b\chi_b^2(1) = \chi^2(2), a = \frac{1}{10}, b = \frac{1}{50}$$

公式里 $\chi_a^2(1)$, $\chi_b^2(1)$ 不同, 故两者系数都必须为 1, 和才是 χ^2 分布

$$6. T = \frac{4(X-2)}{\sqrt{\sum_{i=1}^4 Y_i^2}} \sim \frac{4N(0,1)}{\sqrt{4\chi^2(4)}} = \frac{N(0,1)}{\sqrt{\frac{\chi^2(4)}{4}}} = t(4), P(|T| > t_0 = t_{\frac{\alpha}{2}}(4)) = \alpha = 0.01, t_0 = 4.6041$$

习题 5.3

$$1. E(\bar{X}) = \lambda, D(\bar{X}) = \frac{\lambda}{n}, E(S^2) = \lambda \quad (\text{详细的证明见书 P162 下例 5.1.5})$$

2.

$$(1) P(\bar{X} > 13) = P\left(\frac{\bar{X} - 12}{2/\sqrt{5}} > \frac{\sqrt{5}}{2}\right) = 1 - \Phi\left(\frac{\sqrt{5}}{2}\right) \approx 0.1314$$

$$(2) P\{X > 10\} = P\left\{\frac{X - 12}{2} > -1\right\} = 1 - \Phi(-1) = 0.8413,$$

$$P\{X_{\min} > 10\} = P^5\{X > 10\} = 0.4215$$

3.

$$(1) \frac{nS_n^2}{\sigma^2} = \frac{(n-1)S^2}{\sigma^2} \sim \chi^2(n-1)$$

$$(2) \frac{\bar{X} - \mu}{S_n/\sqrt{n-1}} = \frac{\bar{X} - \mu}{\sqrt{S^2/n}} = \frac{\bar{X} - \mu}{S/\sqrt{n}} \sim t(n-1)$$

$$(3) \quad \frac{X-\mu}{\sigma} \sim N(0,1), \left(\frac{X-\mu}{\sigma}\right)^2 \sim \chi^2(1) \Rightarrow \sum_{i=1}^n \left(\frac{X-\mu}{\sigma}\right)^2 \sim \chi^2(n)$$

$$4. \quad P(|\bar{X} - \mu| > 3) = 1 - P(-3 \leq \bar{X} - \mu \leq 3) = 1 - P\left(-\frac{3}{2} \leq \frac{\bar{X} - \mu}{20/\sqrt{100}} \leq \frac{3}{2}\right) = 1 - \frac{1}{\sqrt{2\pi}} \int_{-\frac{3}{2}}^{\frac{3}{2}} e^{-\frac{x^2}{2}} dx = 0.1336$$

$$5. \quad P\left(\sum_{i=1}^{15} X_i^2 > 3.999\right) = P\left(\sum_{i=1}^{15} \frac{X_i^2}{0.4^2} > 24.99375\right) = P(Y > 24.99375), Y \sim \chi^2(15) \Rightarrow P = 0.05$$

$$6. \quad P\left(\frac{(n-1)S^2}{\sigma^2} > \frac{39 \times 5.6}{4}\right) = P(Y > 54.6) = 0.05, \text{ 这里 } Y \sim \chi^2(39)$$

习题五

1.

$$(1) \quad P\{X_1 = x_1 \dots X_n = x_n\} = \frac{\lambda^{\sum_{i=1}^n x_i}}{\prod_{i=1}^n x_i!} e^{-n\lambda}$$

$$(2) \quad P\{X_1 = x_1 \dots X_n = x_n\} = p^{\sum_{i=1}^n x_i} (1-p)^{n^2 - \sum_{i=1}^n x_i} \prod_{i=1}^n C_n^{x_i}$$

$$2. \quad t(n) = \frac{N(0,1)}{\sqrt{\frac{\chi^2(n)}{n}}}, t^2(n) = \frac{\chi^2(1)}{\frac{\chi^2(n)}{n}} = F(1,n)$$

$$3. \quad \bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right), X_{n+1} - \bar{X} \sim N\left(0, \frac{n+1}{n} \sigma^2\right), (X_{n+1} - \bar{X}) \sqrt{\frac{n}{n+1}} \sim N(0, \sigma^2), \quad \text{所 以}$$

$$T = \frac{X_{n+1} - \bar{X}}{S} \sqrt{\frac{n}{n+1}} \sim \frac{N(0, \frac{\sigma^2}{n}) \sqrt{n}}{S} = \frac{\bar{X} - \mu}{S/\sqrt{n}} = t(n-1)$$

$$4. \quad \alpha(\bar{X} - a) + \beta(\bar{Y} - b) \sim \alpha N\left(0, \frac{\sigma^2}{m}\right) + \beta N\left(0, \frac{\sigma^2}{n}\right) \sim N\left(0, \left(\frac{\alpha^2}{m} + \frac{\beta^2}{n}\right) \sigma^2\right)$$

$$\frac{\alpha(\bar{X}-a)+\beta(\bar{Y}-b)}{\sigma\sqrt{\frac{\alpha^2}{m}+\frac{\beta^2}{n}}}\sim N(0,1)$$

$$\frac{mS_1^2}{\sigma^2}=\frac{(m-1)S^2}{\sigma^2}\sim\chi^2(m-1),\frac{nS_2^2}{\sigma^2}\sim\chi^2(n-1)\Rightarrow\frac{mS_1^2+nS_2^2}{\sigma^2}\sim\chi^2(m+n-2) \quad \text{所以}$$

$$\frac{\alpha(\bar{X}-a)+\beta(\bar{Y}-b)}{\sigma\sqrt{\frac{\alpha^2}{m}+\frac{\beta^2}{n}}}\bigg/\sqrt{\frac{mS_1^2+nS_2^2}{\sigma^2}}\bigg/(m+n-2)\sim\frac{N(0,1)}{\sqrt{\frac{\chi^2(m+n-2)}{m+n-2}}}=t(m+n-2)$$

5. 不会

6. 不会

7. (题目表述有一定问题, 没有解释 $X_i, 5 < i < n$ 的情况, 但可以做题)

$$Y=\left(\frac{n-5}{5}\right)\frac{\sum_{i=1}^5 X_i^2}{\sum_{i=6}^n X_i^2}\sim\frac{\frac{\chi^2(5)}{5}}{\frac{\chi^2(n-5)}{n-5}}\sim F(5, n-5)$$

$$8. \quad P(1.4 < \bar{X} < 5.4) = P\left(-\frac{\sqrt{n}}{3} < \frac{\bar{X}-3.4}{6/\sqrt{n}} < \frac{\sqrt{n}}{3}\right) \geq 0.95 \quad \frac{\sqrt{n}}{3} \geq z_{0.025}, n \geq 34.57, n \geq 35$$

9.

$$(1) \quad \bar{X}-\mu \sim N\left(0, \frac{\sigma^2}{n}\right)=\frac{\sigma}{\sqrt{n}} N(0,1), Y=|\bar{X}-\mu|^2 \sim \frac{\sigma^2}{n} \chi^2(1)=\chi^2\left(\frac{\sigma^2}{n}\right)=\chi^2\left(\frac{4}{n}\right)$$

$$E\left(\chi^2\left(\frac{4}{n}\right)\right)=\frac{4}{n} \leq 0.1, n \geq 40$$

$$(2) \quad P(|\bar{X}-\mu| \leq 0.1)=P\left(-\frac{\sqrt{n}}{20} \leq \frac{\bar{X}-\mu}{2/\sqrt{n}} \leq \frac{\sqrt{n}}{20}\right) \geq 0.95, \text{ 可得}$$

$$\frac{\sqrt{n}}{20} \geq 1.96, n \geq 1536.64, n \geq 1537$$

10.

$$(1) \quad \text{已知 } \frac{1}{\sigma^2} \sum_{i=1}^{20} (X_i - \mu)^2 \sim \chi^2(20), \text{ 则 } P(10.9 \leq \chi^2(20) \leq 37.6) = 0.94$$

$$(2) \quad \text{已知 } \frac{1}{\sigma^2} \sum_{i=1}^{20} (X_i - \bar{X})^2 = \frac{(n-1)}{\sigma^2} S^2 \sim \chi^2(19),$$

$$P(11.7 \leq \chi^2(19) \leq 38.6) = 0.895 \quad (\text{计算结果直接摘录于答案})$$

$$11. P(50.8 \leq \bar{X} \leq 53.8) = P(-\frac{8}{7} \leq \frac{\bar{X}-52}{6.3/\sqrt{36}} \leq \frac{12}{7}) = 0.83$$

$$12. P(|\bar{X} - 80| > 3) = P(|\frac{\bar{X}-80}{20/10}| > \frac{3}{2}) = 1 - P(-\frac{3}{2} \leq \frac{\bar{X}-80}{20/10} \leq \frac{3}{2}) = 0.134$$

$$13. P(|\bar{X} - \bar{Y}| > 0.3) = P(|\frac{\bar{X}-\bar{Y}-(\mu_1-\mu_2)}{\sqrt{\frac{9}{10}+\frac{9}{15}}}| > \frac{\sqrt{6}}{10}) = 1 - \frac{1}{\sqrt{2\pi}} \int_{-\frac{\sqrt{6}}{10}}^{\frac{\sqrt{6}}{10}} e^{-\frac{x^2}{2}} dx = 0.806$$

$$14. \text{同上, } P(|\bar{X} - \bar{Y}| > 0.4) = P(|\frac{\bar{X}-\bar{Y}-(\mu_1-\mu_2)}{\sqrt{\frac{9}{20}+\frac{9}{25}}}| > \frac{4}{9}) = 1 - \frac{1}{\sqrt{2\pi}} \int_{-\frac{4}{9}}^{\frac{4}{9}} e^{-\frac{x^2}{2}} dx = 0.657$$

15.

$$(1) P(|\bar{X} - \mu| > 1) = P(|\frac{\bar{X}-\mu}{2/\sqrt{5}}| > \frac{\sqrt{5}}{2}) = 0.264$$

(2) (这题目我为什么不算最大值、最小值的概率密度函数？因为正态分布的分布函数不好算)

$$P\{X \leq 15\} = P\{\frac{X-12}{2} \leq \frac{3}{2}\} = 0.9332, P\{\max > 15\} = 1 - P^5\{X \leq 15\} = 0.2923$$

$$P\{X < 10\} = P\{\frac{X-12}{2} < -1\} = 0.8413, P\{\min < 10\} = P^5\{X < 10\} = 0.4216 \quad (\text{与答案不同})$$

16. 我们首先应该将前 n 项和后 n 项合并 (可以合并的理由是本身也是独立的), 令

$$X_i + X_{i+n} = Y_i \sim N(2\mu, 2\sigma^2), i=1..n, \text{ 则原来的 } \bar{X} = \frac{1}{2n} \sum_{i=1}^{2n} X_i, \text{ 则 } 2\bar{X} = \frac{\sum_{i=1}^n Y_i}{n} = \bar{Y}, \text{ 则}$$

$$E\{\sum_{i=1}^n (X_i + X_{n+i} - 2\bar{X})^2\} = E\{(n-1) \frac{\sum_{i=1}^n (Y_i - \bar{Y})^2}{n-1}\} = (n-1)2\sigma^2 = 2(n-1)\sigma^2$$

17.

$$(1) P(\frac{S^2}{\sigma^2} \leq 2.041) = P(\frac{(16-1) \times S^2}{\sigma^2} \leq 30.615) = P(Y \leq 30.615), \text{ 其中 } Y \sim \chi^2(15)$$

则 $P(Y \leq 30.615) = 0.01$ (直接摘录于答案)

$$(2) D(S^2) = D(\frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2) = \frac{1}{(n-1)^2} D[\sigma^2 \sum_{i=1}^n (\frac{X_i - \mu - (\bar{X} - \mu)}{\sigma})^2] = \frac{\sigma^4}{(n-1)^2} D[\sum_{i=1}^n (Y_i - \bar{Y})^2]$$

其中, $Y_i \sim N(0,1)$, 则原式

$$= \frac{\sigma^4}{(n-1)^2} D[\sum_{i=1}^n Y_i^2 - n\bar{Y}^2] = \frac{\sigma^4}{(n-1)^2} [D(\sum_{i=1}^n Y_i^2) + n^2 D(\bar{Y}^2) - n \text{Cov}(\sum_{i=1}^n Y_i^2, \bar{Y}^2)]$$

其中

$$\begin{aligned} D(\sum_{i=1}^n Y_i^2) &= 2n, n^2 D(\bar{Y}^2) = \frac{n^2}{n} = n, n \text{Cov}(\sum_{i=1}^n Y_i^2, \bar{Y}^2) = n[E(\bar{Y}^2 \sum_{i=1}^n Y_i^2) - E(\sum_{i=1}^n Y_i^2)E(\bar{Y}^2)] \\ &= n[nE(\bar{Y}^2 \bar{Y}^2)] \end{aligned}$$

18.

(1) $Y = 2\lambda X$ 单调, 可以套用公式计算概率密度函数,

$$f_Y(y) = \begin{cases} f_X(h(y)) |h'(y)| = \frac{1}{2} e^{-\frac{y}{2}}, y > 0 \\ 0, y \leq 0 \end{cases}$$

(2) 其实也就是 $2n\lambda\bar{X} = \sum_{i=1}^n Y_i \sim \chi^2(2n)$, 由于 $\chi^2(n)$ 具有可加性, 所以猜测证明

$Y_i \sim \chi^2(2)$, 我们翻到 P270 页找到

$$\chi^2(n): f(x) = \begin{cases} \frac{1}{2^{\frac{n}{2}} \Gamma(n/2)} x^{\frac{n}{2}-1} e^{-\frac{x}{2}}, x > 0 \\ 0, \text{其他} \end{cases}, n=2, f(x) = \begin{cases} \frac{1}{2} e^{-\frac{x}{2}}, x > 0 \\ 0, \text{其他} \end{cases}, \text{正是 } Y_i \text{ 的分布, 所}$$

以 $Y_i \sim \chi^2(2)$, $2n\lambda\bar{X} = \sum_{i=1}^n Y_i \sim \chi^2(2n)$ ($\Gamma(n) = (n-1)!, n$ 是整数)

19. 书上 P173 定理 5.3.2 提到过 \bar{X} 、 S^2 相互独立 $D(T) = D(\bar{X}^2) + \frac{1}{n^2} D(S^2)$, 其中

$$\bar{X} \sim N(0, \frac{1}{n}) = \frac{1}{\sqrt{n}} N(0, 1), S^2 \sim \frac{1}{n-1} t(n-1) \text{ 则}$$

$$D(T) = \frac{1}{n^2} D(\chi^2(1)) + \frac{1}{n^2(n-1)^2} D(t(n-1)) = \frac{2}{n^2} + \frac{\frac{n-1}{n-3}}{n^2(n-1)^2} = \frac{1}{n^2} (2 + \frac{1}{(n-1)(n-3)})$$

第六章

习题 6.1

1. $\hat{\mu} = \frac{\sum_{i=1}^5 x_i}{5} = 52.84$ (答案有误), $\hat{\sigma}^2 = \frac{1}{5} \sum_{i=1}^5 (x_i - \bar{x})^2 = 0.1304$

2. 服从正态分布的话, 其实矩估计值我们已经知道了, 同上, $\hat{\mu} = \frac{\sum_{i=1}^5 x_i}{5} = 2809.4$,

$$\hat{\sigma}^2 = \frac{1}{5} \sum_{i=1}^5 (x_i - \bar{x})^2 = 1170.64$$

3. 服从均匀分布:

(1) 未知参数只有两个, 所以需要两阶:

$$\mu_1 = E(X) = \int_{-\infty}^{\infty} xf(x)dx = \int_{\theta_1}^{\theta_1+\theta_2} \frac{xdx}{\theta_2} = \frac{2\theta_1 + \theta_2}{2}, A_1 = \frac{\sum_{i=1}^n X_i}{n} = \bar{X}$$

$$\mu_2 = E(X^2) = \int_{-\infty}^{\infty} x^2 f(x)dx = \int_{\theta_1}^{\theta_1+\theta_2} \frac{x^2 dx}{\theta_2} = \frac{3\theta_1^2 + 3\theta_1\theta_2 + \theta_2^2}{3},$$

$$A_2 = \frac{\sum_{i=1}^n X_i^2}{n} = E(X^2) = D(X) + E(X)^2 = S_n^2 + \bar{X}^2$$

$$\text{连立方程, } \begin{cases} \mu_1 = A_1 \\ \mu_2 = A_2 \end{cases}$$

$$\text{解得, } \hat{\theta}_1 = \bar{X} - \sqrt{3}S_n, \hat{\theta}_2 = 2\sqrt{3}S_n$$

$$(2) \quad f(x) = \begin{cases} \frac{1}{\theta_2}, & \theta_1 < x < \theta_1 + \theta_2, \\ 0, & \text{其他} \end{cases}, \quad L(x_i; \theta_1, \theta_2) = \frac{1}{\theta_2^n},$$

$$\ln L(\theta_2) = -n \ln \theta_2, \frac{d \ln L(\theta_2)}{d \theta_2} = -\frac{n}{\theta_2}, \text{ 不存在导数为 } 0$$

但是, 我们从极大似然法的定义来看, 要使 $L(\theta_2)$ 最大, 则 θ_2 尽量小, 又因为

$$\theta_2 + \theta_1 \geq \max\{x_1, x_2, \dots, x_n\}, \quad \theta_1 \leq \min\{x_1, x_2, \dots, x_n\}, \text{ 所以}$$

$$\hat{\theta}_2 + \hat{\theta}_1 = \max\{x_1, x_2, \dots, x_n\}, \quad \hat{\theta}_1 = \min\{x_1, x_2, \dots, x_n\}, \quad \text{即 } \hat{\theta}_1 = X_{(1)}, \hat{\theta}_2 = X_{(n)} - X_{(1)}$$

4. 仅有一个参数

$$(1) \quad \mu_1 = \int_{-\infty}^{\infty} xf(x)dx = \frac{\theta+1}{\theta+2}, A_1 = \bar{X} \Rightarrow \hat{\theta} = \frac{1-2\bar{X}}{\bar{X}-1}$$

$$(2) \quad L(\theta) = \prod_{i=1}^n f(x_i; \theta) = (\theta+1)^n \left(\prod_{i=1}^n x_i\right)^{\theta}, \ln L(\theta) = n \ln(\theta+1) + \theta \sum_{i=1}^n \ln x_i$$

$$\frac{d \ln L(\theta)}{d\theta} = \frac{n}{\theta+1} + \sum_{i=1}^n \ln x_i = 0, \theta = -1 - \frac{n}{\sum_{i=1}^n \ln x_i} \quad (\text{与答案不同})$$

5. (摘抄书本 187 页)

$$f(x) = \begin{cases} \frac{1}{2\theta}, & -\theta < x < \theta \\ 0, & \text{其他} \end{cases}$$

$$L(x_i; \theta) = \frac{1}{(2\theta)^n}$$

$$\ln L(\theta) = -n \ln 2\theta, \frac{d \ln L(\theta)}{d\theta} = -\frac{n}{\theta}, \quad \text{不存在导数为 0}$$

但是，我们从极大似然法的定义来看，要使 $L(\theta)$ 最大，则 θ 尽量小，又因为

$$\theta \geq \max\{|x_1|, |x_2|, \dots, |x_n|\}, \quad \text{所以 } \hat{\theta} = \max\{|x_1|, |x_2|, \dots, |x_n|\}$$

习题 6.2

1. 仅有一个参数

$$(1) \quad \mu_1 = \int_0^1 xf(x; \theta)dx = \frac{1+2\theta}{4} = \bar{X}, \hat{\theta} = \frac{4\bar{X}-1}{2} = 2\bar{X} - \frac{1}{2}$$

$$(2) \quad E(4\bar{X}^2) = E[(2\bar{X})^2] = D(2\bar{X}) + E(2\bar{X})^2 > (\theta + \frac{1}{2})^2 > \theta^2, \quad \text{不是无偏估计}$$

2. $\mu_1 = \int_0^{\infty} xf(x)dx = \theta, \therefore \hat{\theta} = \bar{X}$, (第二题摘自浙大版 159 页)

$$Z = \min\{X_1, X_2, \dots, X_n\}$$

$$F(Z) = F_{\min}(x; \theta) = 1 - \prod_{i=1}^n (1 - \int_0^x \frac{1}{\theta} e^{-\frac{x}{\theta}} dx) = 1 - e^{-\frac{nx}{\theta}}$$

$$f(Z) = \frac{dF_{\min}(x)}{dx} = \frac{n}{\theta} e^{-\frac{nx}{\theta}}, \quad x > 0$$

$$E(nZ) = nE(Z) = n \cdot \frac{\theta}{n} = \theta$$

3. 如下,

$$\mu_1 \sim N\left(\frac{1}{5}\mu + \frac{3}{10}\mu + \frac{1}{2}\mu, \frac{1}{25}\sigma^2 + \frac{9}{100}\sigma^2 + \frac{1}{4}\sigma^2\right) = N\left(\mu, \frac{38}{100}\sigma^2\right)$$

$$\mu_2 \sim N\left(\mu, \frac{25}{72}\sigma^2\right)$$

$$\mu_2 \sim N\left(\mu, \frac{7}{18}\sigma^2\right)$$

由上面可知, μ_2 最有效

$$4. \quad F_{\max}(X) = \frac{x^3}{\theta^3}, f_{\max}(X) = \frac{3x^2}{\theta^3}; F_{\min}(X) = 1 - (1 - \frac{x}{\theta})^3 = \frac{x^3}{\theta^3} - 3\frac{x^2}{\theta^2} + 3\frac{x}{\theta}, f_{\min}(x) = \frac{3x^2}{\theta^3} - 6\frac{x}{\theta^2} + 3\frac{1}{\theta}$$

$$Y = \frac{4}{3} \max\{X_1, X_2, X_3\}, E(Y) = \frac{4}{3} \int_0^\theta x \cdot \frac{3x^2}{\theta^3} dx = \theta$$

, 都是无偏估计

$$Z = 4 \min\{X_1, X_2, X_3\}, E(Z) = 4 \int_0^\theta x \left(\frac{3x^2}{\theta^3} - 6\frac{x}{\theta^2} + 3\frac{1}{\theta}\right) dx = \theta$$

$$D(Y) = \frac{16}{9} D(\max\{X_1, X_2, X_3\}) = \frac{16}{9} [E(\max^2\{X_1, X_2, X_3\}) - E(\max\{X_1, X_2, X_3\})^2] = \frac{1}{15} \theta^2$$

$$D(Z) = 16 D(\min\{X_1, X_2, X_3\}) = 16 [E(\min^2\{X_1, X_2, X_3\}) - E(\min\{X_1, X_2, X_3\})^2] = \frac{3}{5} \theta^2$$

所以 $\frac{4}{3} \max\{X_1, X_2, X_3\}$ 更有效

$$5. \quad (\text{法一, 用别的已证明的}) Y = \frac{12}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2 = 12S^2 \text{ 是 } 12\sigma^2 \text{ 的无偏估计, 而}$$

$\sigma^2 = \frac{1}{12} \theta^2$, 所以 Y 是 σ^2 的无偏估计, 那相合也是肯定的, 因为用了已经证明了的。

(法二: 自己写)

$$\begin{aligned} Y &= \frac{12}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2 = \frac{12}{n-1} (\sum_{i=1}^n X_i^2 - n\bar{X}^2), E(Y) = \frac{12}{n-1} [E(\sum_{i=1}^n X_i^2) - nE(\bar{X}^2)] \\ &= \frac{12}{n-1} [\sum_{i=1}^n (\sigma^2 + \mu^2) - n(\frac{\sigma^2}{n} + \mu^2)] = 12\sigma^2 = \theta^2 \end{aligned}$$

要证相合性，就是 $\lim_{n \rightarrow \infty} P\{|\frac{12}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2 - \theta^2| < \varepsilon\} = 1$ ，就要往辛钦大数定律凑，

构造 $Z_i = \frac{12n}{n-1} (X_i - \bar{X})^2$ ，目标是证明 $E(Z_i) = \theta^2$ 则大功告成，但是因为 \bar{X} 的存在工作不好开展，所以我不会……

习题 6.3

（本篇里面答案的值直接摘录于标准答案，如果有错实在抱歉，还请联系我）

1. 首先我们确认含锡量是不是越少越好，所以是单侧置信区间，置信区间是

$$(0, \bar{X} + \frac{\sigma}{\sqrt{n}} z_a), \text{ 绝对误差不超过 } 2.5, \text{ 就是 } \frac{\sigma}{\sqrt{n}} z_a \leq 2.5, n \geq (\frac{4}{2.5} z_a)^2 = 6.88, \text{ 最小值}$$

n 为 7（答案给 10 是按双侧置信区间来的，这个仁者见之）

2. μ 的置信区间是 $(\bar{X} - \frac{S}{\sqrt{n}} t_{\frac{\alpha}{2}}(16-1), \bar{X} + \frac{S}{\sqrt{n}} t_{\frac{\alpha}{2}}(16-1))$, 即 (2.6895, 2.7205)

σ^2 的置信区间是 $(\frac{(n-1)S^2}{\chi_{\frac{\alpha}{2}}^2(n-1)}, \frac{(n-1)S^2}{\chi_{1-\frac{\alpha}{2}}^2(n-1)})$, 即 (4.59e-4, 0.00201)（与答案严重不符）

3. 已知方差， $\bar{X} = 14.95$ ，方差已知，置信区间 $(\bar{X} - \frac{\sigma}{\sqrt{n}} Z_{\frac{\alpha}{2}}, \bar{X} + \frac{\sigma}{\sqrt{n}} Z_{\frac{\alpha}{2}})$ ，即 (14.8, 15.2)

4. 如下

$$(1) \quad (\bar{X} - \frac{\sigma}{\sqrt{n}} Z_{\frac{\alpha}{2}}, \bar{X} + \frac{\sigma}{\sqrt{n}} Z_{\frac{\alpha}{2}}) \Rightarrow (19.87, 20.15)$$

$$(2) \quad (\bar{X} - \frac{S}{\sqrt{n}} t_{\frac{\alpha}{2}}, \bar{X} + \frac{S}{\sqrt{n}} t_{\frac{\alpha}{2}}) \Rightarrow (19.87, 20.17)$$

5. 已知方差，套用左边的公式，即 $(\bar{X} - \bar{Y} - z_{\frac{\alpha}{2}} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}, \bar{X} - \bar{Y} + z_{\frac{\alpha}{2}} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}})$ ，

$$(-0.899, 0.019)$$

6. 比较两个样本的，预计使用 F 模型，套用左边最后一个模型，

$$(\frac{S_X^2}{S_Y^2} \frac{1}{F_{\frac{\alpha}{2}}(n_1-1, n_2-1)}, \frac{S_X^2}{S_Y^2} \frac{1}{F_{1-\frac{\alpha}{2}}(n_1-1, n_2-1)}), (0.2217, 3.6008)$$

7. 含锡量越小越好，又是单侧置信上限，所以是 $\bar{X} + \frac{\sigma}{\sqrt{n}} Z_{\alpha}$ ，即 183.3515

习题 6.4

1. 就同上（题目的上面那道例题）， $\hat{p}_1 = \frac{1}{2a}(-b - \sqrt{b^2 - 4ac})$, $\hat{p}_2 = \frac{1}{2a}(-b + \sqrt{b^2 - 4ac})$

其中， $a = n + (\frac{z_{\frac{\alpha}{2}})^2$, $b = -[2n\bar{x} + (\frac{z_{\alpha}}{2})^2]$, $c = n\bar{x}^2$ ，带入即可。

2. 同上
3. 同上

习题六

- 1.

(1) 一个参数，就只用算一阶矩， $\mu_1 = \sum_{i=0}^3 iP(i) = 3 - 4\theta$, $A_1 = \bar{X} = 2, \therefore \hat{\theta} = \frac{1}{4}$

(2) $L(\theta) = (1 - 2\theta)^4 \cdot [2\theta(1 - \theta)]^2 \cdot \theta^2 \cdot \theta^2 = 4\theta^6(1 - \theta)^2(1 - 2\theta)^4$,

$$\ln L(\theta) = \ln 4 + 6 \ln \theta + 2 \ln(1 - \theta) + 4 \ln(1 - 2\theta),$$

$$\frac{d \ln L(\theta)}{d\theta} = \frac{6}{\theta} - \frac{2}{1 - \theta} - \frac{8}{1 - 2\theta} = 0 \Rightarrow \theta = \frac{7 \pm \sqrt{13}}{12}, \text{要求 } \frac{d^2 \ln L(\theta)}{d\theta^2} = -\frac{6}{\theta^2} - \frac{2}{(1 - \theta)^2} - \frac{16}{(1 - 2\theta)^2} < 0$$

则， $\theta = \frac{7 - \sqrt{13}}{12}$

2. 这 话 的 意 思 就 是 极 大 似 然 法 ， 平 均 数 为 λ ，

$$L(\lambda) = \prod_{i=1}^{50} P(x=i) = \prod_{i=1}^{50} \frac{\lambda^{x_i}}{x_i!} e^{-\lambda} = \frac{\lambda^{\sum_{i=1}^{50} x_i}}{2^{10} \times (3!)^2 \times (4!)} e^{-50\lambda}, \ln L(\lambda) = \sum_{i=1}^{50} x_i \ln \lambda - 50\lambda - \ln C$$

$$\frac{d \ln L(\lambda)}{d\lambda} = \frac{\sum_{i=1}^{50} x_i}{\lambda} - 50 = 0, \hat{\lambda} = \bar{x} = 1$$

- 3.

(1) $L(\theta) = \theta^{2N_1} \cdot [2\theta(1 - \theta)]^{N_2} \cdot (1 - \theta)^{2N_3} = 2^{N_2} \theta^{2N_1 + N_2} \cdot (1 - \theta)^{N_2 + 2N_3}$,

$$\ln L(\theta) = N_2 \ln 2 + (2N_1 + N_2) \ln \theta + (N_2 + 2N_3) \ln(1 - \theta)$$

$$\frac{d \ln L(\theta)}{d\theta} = \frac{2N_1 + N_2}{\theta} - \frac{N_2 + 2N_3}{1-\theta} = 0, \hat{\theta} = \frac{2N_1 + N_2}{2(N_1 + N_2 + N_3)}$$

(2) 带入, 得 $\lambda=0.335$

4. 两个参数, 所以需要列两个方程

$$\mu_1 = np, A_1 = \bar{X}$$

$$\mu_2 = E(X^2) = D(X) + E(X)^2 = np(1-p) + n^2 p^2, A_2 = \frac{\sum_{i=1}^n X_i^2}{n} = \frac{\sum_{i=1}^n (X_i - \bar{X})^2 + n\bar{X}^2}{n} = S_n^2 + \bar{X}^2$$

$$\text{连立方程, 得, } \hat{p} = 1 - \frac{S_n^2}{\bar{X}}, \hat{N} = \frac{\bar{X}}{\hat{p}}$$

5. 若 $p=0$ or 1 , 那么 $P(X=x)=0$, 不满足归一化条件, 所以不可能

$$(1) \text{ 矩估计: 仅有一个参数, } \mu_1 = \sum_{k=1}^{\infty} kp(1-p)^{k-1} = \frac{1}{p} = \bar{X}, \hat{p} = \frac{1}{\bar{X}}$$

$$(2) \quad L(p) = p^n (1-p)^{\sum_{i=1}^n x_i - n}, \ln L(p) = n \ln p + (\sum_{i=1}^n x_i - n) \ln(1-p)$$

$$\frac{d \ln L(p)}{dp} = \frac{n}{p} - \frac{\sum_{i=1}^n x_i - n}{1-p} = 0, \hat{p} = \frac{n}{\sum_{i=1}^n x_i}$$

$$6. \quad L(\theta) = \theta^N \cdot (1-\theta)^{n-N}, \ln L(\theta) = N \ln \theta + (n-N) \ln(1-\theta)$$

$$\frac{d \ln L(\theta)}{d\theta} = \frac{N}{\theta} - \frac{n-N}{1-\theta} = 0, \hat{\theta} = \frac{N}{n}$$

$$7. \quad E\left(\frac{1}{n} \sqrt{\frac{\pi}{2}} \sum_{i=1}^n |X_i - 1|\right) = \frac{1}{n} \sqrt{\frac{\pi}{2}} \cdot 2 \int_0^{\infty} \frac{nx}{\sqrt{2\pi}\sigma} e^{-\frac{x^2}{2\sigma^2}} dx = \frac{1}{n} \sqrt{\frac{\pi}{2}} \cdot \frac{2\sigma n}{\sqrt{2\pi}} = \sigma$$

8. 这里题目我自行补充一下, Y 与 X 同方差, 那么

$$S^2 = \frac{1}{n+m-2} \left[\sum_{i=1}^n (X_i - \bar{X})^2 + \sum_{i=1}^m (Y_i - \bar{Y})^2 \right] = \frac{1}{n+m-2} [(n-1)S_X^2 + (m-1)S_Y^2]$$

其中 S_X^2 、 S_Y^2 都是 σ^2 的无偏估计, 那么 S^2 就是 $\frac{1}{n+m-2} [(n-1)\sigma^2 + (m-1)\sigma^2] = \sigma^2$ 的无偏估计。

$$9. \quad E(\bar{X}) = \lambda, E(\bar{X}^2) = D(\bar{X}) + E(\bar{X})^2 = \frac{n\lambda}{n} + \lambda^2 \Rightarrow \lambda^2 = E(\bar{X}^2) - \frac{E(\bar{X})^2}{n} \quad \text{即} \quad \bar{X}^2 - \frac{\bar{X}}{n}$$

是 λ^2 的无偏估计。

10.

(1) 不会

$$(2) \quad \mu = E[(\bar{X})^2 - CS^2] = E(\bar{X}^2) - CE(S^2) = D(\bar{X}) + E(\bar{X})^2 - C(n-1)\sigma^2 = \frac{\sigma^2}{n} + \mu^2$$

$$11. \quad E(k\bar{X} + (1-k)S^2) = kE(\bar{X}) + (1-k)E(S^2) = k\lambda + (1-k)\lambda = \lambda$$

$$12. \quad E(Z) = aE(S_1^2) + bE(S_2^2) = a\sigma^2 + b\sigma^2 = \sigma^2, \text{ 又因为 } X, Y \text{ 分布独立, 所以 } S_1^2, S_2^2 \text{ 独}$$

立, 不妨设 $D(S_1^2) = D(S_2^2) = \alpha^2$, 所以

$$D(Z) = a^2 D(S_1^2) + b^2 D(S_2^2) = (a^2 + b^2)\alpha^2, \text{ 在 } a = b = \frac{1}{2} \text{ 时最小}$$

13.

$$(1) \quad T_1 = (m-1)S_X^2 + (n-1)S_Y^2, \frac{T_1}{\sigma^2} = \frac{(m-1)S_X^2}{\sigma^2} + \frac{(n-1)S_Y^2}{\sigma^2} \sim \chi^2(m-1) + \chi^2(n-1) = \chi^2(m+n-2)$$

$$T_2 = \sum_{i=1}^m (X_i - \bar{X} + \bar{X} - \mu)^2 + \sum_{i=1}^n (Y_i - \bar{Y} + \bar{Y} - \mu)^2$$

$$= \sum_{i=1}^m [(X_i - \bar{X})^2 + (\bar{X} - \mu)^2] + \sum_{i=1}^n [(Y_i - \bar{Y})^2 + (\bar{Y} - \mu)^2] = T_1 + m(\bar{X} - \mu)^2 + n(\bar{Y} - \mu)^2$$

, 其中, 展开的乘积项 $\sum_{i=1}^m 2(X_i - \bar{X}) \cdot (\bar{X} - \mu) = 2(\bar{X} - \mu)(\sum_{i=1}^m X_i - m\bar{X}) = 0$, Y 同理

$$\text{下面处理非 } T_1 \text{ 项, 即 } \frac{m(\bar{X} - \mu)^2 + n(\bar{Y} - \mu)^2}{\sigma^2} = \left(\frac{\bar{X} - \mu}{\sigma/\sqrt{m}}\right)^2 + \left(\frac{\bar{Y} - \mu}{\sigma/\sqrt{n}}\right)^2 = \chi^2(2),$$

$$\text{所以 } \frac{T_2}{\sigma^2} \sim \chi^2(m+n-2) + \chi^2(2) = \chi^2(m+n)$$

$$(2) \quad \text{由 (1) 中信息以及 } E[\chi^2(n)] = n \text{ 可得 } C_1 = \frac{1}{n+m-2}, C_2 = \frac{1}{n+m},$$

$$(3) \quad D(T_1^*) = \frac{1}{(m+n-2)^2} \cdot 2(m+n-2) = \frac{2}{m+n-2} > \frac{2}{m+n}, \text{ 所以 } T_2^* \text{ 更优.}$$

14.

$$(1) \quad \frac{T}{\sigma^2} = \frac{(n-1)S^2}{\sigma^2} + \sum_{i=n+1}^{2n} \frac{(X_i - \mu)^2}{\sigma^2} \sim \chi^2(n-1) + \chi^2(n) = \chi^2(2n-1)$$

$$(2) \quad P[\chi^2_{1-\frac{a}{2}}(2n-1) < \frac{T}{\sigma^2} < \chi^2_{\frac{a}{2}}(2n-1)] \geq 1-a, \therefore \sigma^2 \in (\frac{T}{\chi^2_{\frac{a}{2}}(2n-1)}, \frac{T}{\chi^2_{1-\frac{a}{2}}(2n-1)})$$

15. σ 已知, 直接套公式, 置信区间长度为 $\frac{2\sigma}{\sqrt{n}} Z_{\frac{a}{2}} \leq L, n \geq (\frac{2\sigma}{L} Z_{\frac{a}{2}})^2$

16. 同上题, $\frac{2\sigma}{\sqrt{n}} Z_{\frac{a}{2}} = 0.658, Z_{\frac{a}{2}} = \frac{0.658 \times 10}{4} = 1.645, a = 0.1, 1-a = 0.9$

17.

$$(1) \quad X = e^Y, Y \sim N(\mu, 1), E(X) = \int_{-\infty}^{\infty} e^y \frac{1}{\sqrt{2\pi}} e^{-\frac{(y-\mu)^2}{2}} dy = \frac{e^{\mu}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{y-\mu} e^{-\frac{(y-\mu)^2}{2}} d(y-\mu)$$

$$= \frac{e^{\mu}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{y-\frac{y^2}{2}} dy = \frac{e^{\mu}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{(y-1)^2+1}{2}} dy = \frac{e^{\mu+\frac{1}{2}}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{(y-1)^2}{2}} d(y-1) = e^{\mu+\frac{1}{2}}$$

$$(2) \quad \bar{Y} = \frac{\sum_{i=1}^4 \ln X_i}{4} = 0, \therefore \mu \text{ 的置信区间 } (-\frac{\sigma}{\sqrt{n}} Z_{\frac{a}{2}}, \frac{\sigma}{\sqrt{n}} Z_{\frac{a}{2}}), \text{ 即 } (-0.98, 0.98) \quad (\text{与答案不同})$$

(3) 题目说让我们利用上述结果, 那我们就用呗, $(e^{-0.98+\frac{1}{2}}, e^{0.98+\frac{1}{2}}) = (0.619, 4.393)$

18. $Y = D(\frac{X^2}{\sigma^2}) = D(\frac{(X-\mu)^2}{\sigma^2} - 2\mu \frac{X}{\sigma^2} + \frac{\mu^2}{\sigma^2}) = D(\frac{(X-\mu)^2}{\sigma^2} - 2\mu \frac{X}{\sigma^2})$, 其中,

$$\frac{(X-\mu)^2}{\sigma^2} \sim \chi^2(1), X \sim N(\mu, \sigma^2)$$

19. 不会

20. 如下:

(1) 套公式, 这里 σ^2 未知, 就套 $(\bar{x} - \frac{S}{\sqrt{n}} t_{\frac{a}{2}}(n-1), \bar{x} + \frac{S}{\sqrt{n}} t_{\frac{a}{2}}(n-1))$, 为

(2.196, 2.230) (与答案不同)

(2) 套公式, 这里 μ 未知, $(\frac{(n-1)S^2}{\chi^2_{\frac{a}{2}}(n-1)}, \frac{(n-1)S^2}{\chi^2_{1-\frac{a}{2}}(n-1)})$, 即 $(1.73 \times 10^{-4}, 2.65 \times 10^{-3})$

21. 这里不知道 σ^2 , 套 $(\bar{x} - \frac{S}{\sqrt{n}} t_{\frac{a}{2}}(n-1), \bar{x} + \frac{S}{\sqrt{n}} t_{\frac{a}{2}}(n-1))$, (750.5, 849.5)

22. 同上, 得(170.44, 173.57)

23. 同上, $\bar{X} = 1147$, $S = 87.06$ 得 (1084.72, 1209.28) (与答案偏差在于小数位的选取)

24. 同上, (30.868, 31.252)

25. 一万户指的是大样本的意思, 那么很没有道理的利用书上 P150 页 4.4.3 第一段: “这种随机变量往往近似的服从正态分布” 的一句话, 我们也不妨假设其满足正态分布, 那就好做了。区间估计为

26. 这就是没给 μ 、 σ^2 , 可以算出, $\bar{X} = 69.8, S^2 = 36.43$

$$(1) \quad (\bar{X} - \frac{S}{\sqrt{16}} t_{\frac{\alpha}{2}}(15), \bar{X} + \frac{S}{\sqrt{16}} t_{\frac{\alpha}{2}}(15)), \text{即 } (66.59, 73.03) \quad (\text{小数保存不同})$$

$$(2) \quad (\frac{(n-1)S^2}{\chi_{\frac{\alpha}{2}}^2(n-1)}, \frac{(n-1)S^2}{\chi_{1-\frac{\alpha}{2}}^2(n-1)}), \text{取根号后, 即 } (4.46, 9.34) \quad (\text{显然不同})$$

27. 标准差已知, 那么 $(\bar{X} - \bar{Y} - z_{\frac{\alpha}{2}} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}, \bar{X} - \bar{Y} + z_{\frac{\alpha}{2}} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}})$, 即 (-6.04, -5.96)

28. $\bar{X} = 130, S_X^2 = 58, \bar{Y} = 135, S_Y^2 = 11.2, S_W = 5.36$, 那么 () (-12.98, 2.98)

29. 同上 28 题, 得 (-0.002, 0.006) (直接摘录于答案)

30. σ^2 未知, $\bar{x} - \frac{S}{\sqrt{n}} t_{\alpha}(n-1)$, 即 40527 (直接摘录)

31. σ^2 已知, 所以单侧置信下限为 $\bar{X} - \frac{\sigma}{\sqrt{n}} z_{\alpha}$, 即 409.88, 置信上限 $\bar{X} + \frac{\sigma}{\sqrt{n}} z_{\alpha}$, 即 420.12

32. 我唯一的想法就是, 近似服从正态分布.....这样

$$\frac{\bar{X} - \lambda}{\lambda / \sqrt{100}} \sim N(0, 1), \text{ 则 } -Z_{0.01} < \frac{\bar{X} - \lambda}{\lambda / 10} < Z_{0.01}, \text{ 计算得 } \frac{\bar{X}}{\frac{Z_{0.01}}{10} + 1} < \lambda < \frac{\bar{X}}{\frac{-Z_{0.01}}{10} + 1}, \text{ 即 } 3.24 < \lambda < 5.22$$

标准答案给的是 (3.56, 4.49) (欢迎会的同学私戳)

第七章

习题 7.2

1. 假设: $H_0: \mu = 4.55$ (无显著变化); $H_1: \mu \neq 4.55$ (显著变化)。则拒绝域为

$$P(|\frac{\bar{x}-4.55}{0.108/\sqrt{5}}| \geq z_{\frac{\alpha}{2}}) = \alpha, \text{ 其中 } z_{\frac{\alpha}{2}} = 1.96, |\frac{\bar{x}-4.55}{0.108/\sqrt{5}}| = 3.93 > z_{\frac{\alpha}{2}}, \text{ 拒绝 } H_0: \mu = 4.55$$

所以有显著变化。

2. 同上, 假设: $H_0: \mu = 1277$ (无明显偏差); $H_1: \mu \neq 1277$ (有明显偏差)。则拒绝域

$$\text{为 } P(|\frac{\bar{x}-1277}{S/\sqrt{5}}| \geq t_{\frac{\alpha}{2}}(4)) = \alpha, \text{ 其中 } t_{\frac{\alpha}{2}}(4) = 2.7764, |\frac{\bar{x}-1277}{S/\sqrt{5}}| = 3.37 > t_{\frac{\alpha}{2}}(4), \text{ 拒绝}$$

$H_0: \mu = 1277$, 所以有明显偏差。(和标准答案有明显差别)

3. 假设: $H_0: \sigma^2 = \sigma_0^2$ (无显著变化); $H_1: \sigma^2 \neq \sigma_0^2$ (有显著变化) 拒绝域为

$$\frac{(n-1)S^2}{\sigma_0^2} > \chi_{\frac{\alpha}{2}}^2(n-1) \cup \frac{(n-1)S^2}{\sigma_0^2} < \chi_{1-\frac{\alpha}{2}}^2(n-1), \text{ 其中 } \chi_{\frac{\alpha}{2}}^2(n-1) = 54.437,$$

$$\chi_{1-\frac{\alpha}{2}}^2(n-1) = 21.336, \frac{(n-1)S^2}{\sigma_0^2} = 51.84, \text{ 不在拒绝域内, 所以接受 } H_0: \sigma^2 = \sigma_0^2,$$

无显著变化

4. 同题 2, 假设: $H_0: \mu < 215$ (未达到预期效果); $H_1: \mu \geq 215$ (达到效果)。则拒绝

$$\text{域为 } \frac{\bar{x}-215}{S/\sqrt{16}} \geq t_{\alpha}(15) \text{ 其中 } t_{\alpha}(15) = 1.7531, \frac{\bar{x}-215}{S/\sqrt{16}} = -2.14 < t_{\alpha}(15), \text{ 接受}$$

$H_0: \mu < 215$, 所以未达到效果。

5. 假设: $H_0: \sigma^2 = \sigma_0^2$ (没有变劣); $H_1: \sigma^2 > \sigma_0^2$ (变劣) 拒绝域为 $\frac{(n-1)S^2}{\sigma_0^2} > \chi_{\alpha}^2(n-1)$,

$$\text{其中 } \chi_{\alpha}^2(n-1) = 24.996, \frac{(n-1)S^2}{\sigma_0^2} = 45.9 > \chi_{\alpha}^2(n-1), \text{ 在拒绝域内, 所以拒绝}$$

$H_0: \sigma^2 = \sigma_0^2$, 显著变劣

习题 7.3

1. 假设: $H_0: \mu_1 - \mu_2 = 0$; $H_1: \mu_1 - \mu_2 \neq 0$, 拒绝域

$$|\frac{\bar{X}-\bar{Y}}{\sqrt{\frac{4}{5}+\frac{6}{5}}}| \geq Z_{\frac{\alpha}{2}}, |\frac{\bar{X}-\bar{Y}}{\sqrt{\frac{4}{5}+\frac{6}{5}}}| = 1.84 < Z_{\frac{\alpha}{2}} = 1.96, \text{ 所以接受 } H_0: \mu_1 - \mu_2 = 0, \text{ 没有差异}$$

2. 假设: $H_0: \mu_1 - \mu_2 = 0; H_1: \mu_1 - \mu_2 \neq 0$, 拒绝域

$$\left| \frac{\bar{X} - \bar{Y}}{S_w \sqrt{\frac{1}{6} + \frac{1}{5}}} \right| \geq t_{\frac{\alpha}{2}}(9), \left| \frac{\bar{X} - \bar{Y}}{S_w \sqrt{\frac{1}{6} + \frac{1}{5}}} \right| = 1.237 < t_{\frac{\alpha}{2}}(9) = 3.2498, \text{ 所以接受}$$

$H_0: \mu_1 - \mu_2 = 0$, 没有差异

3. 假设: $H_0: \sigma_1^2 < \sigma_2^2; H_1: \sigma_1^2 \geq \sigma_2^2$, 拒绝域

$$\frac{S_1^2}{S_2^2} > F_a(m-1, n-1), F_a(m-1, n-1) = 3.48, \frac{S_1^2}{S_2^2} = 0.10813 < F_a(m-1, n-1), \text{ 所以接受}$$

受 $H_0: \sigma_1^2 < \sigma_2^2$, 新生女婴体重的方差冬季的比夏季的小。

4.

- (1) 假设: $H_0: \sigma_1^2 = \sigma_2^2; H_1: \sigma_1^2 \neq \sigma_2^2$, 拒绝域

$$\frac{S_1^2}{S_2^2} > F_{\frac{\alpha}{2}}(m-1, n-1) \text{ 或者 } \frac{S_1^2}{S_2^2} < F_{1-\frac{\alpha}{2}}(m-1, n-1), \text{ 其中}$$

$$F_{\frac{\alpha}{2}}(m-1, n-1) = 5.82, F_{1-\frac{\alpha}{2}}(m-1, n-1) = \frac{1}{F_{\frac{\alpha}{2}}(n-1, m-1)} = 0.172, \frac{S_1^2}{S_2^2} = 2.02,$$

不在拒绝域内, 所以接受 $H_0: \sigma_1^2 = \sigma_2^2$, 无显著差异

- (2) 说方差无变化, 就是说 $\sigma_1^2 = \sigma_2^2 = \sigma^2$, 假设: $H_0: \mu_1 - \mu_2 = 0; H_1: \mu_1 - \mu_2 \neq 0$,

$$\text{拒绝域} \left| \frac{\bar{X} - \bar{Y}}{S_w \sqrt{\frac{1}{7} + \frac{1}{7}}} \right| \geq t_{\frac{\alpha}{2}}(12), \left| \frac{\bar{X} - \bar{Y}}{S_w \sqrt{\frac{1}{7} + \frac{1}{7}}} \right| = 2.46 > t_{\frac{\alpha}{2}}(12) = 2.1788, \text{ 所以拒绝}$$

$H_0: \mu_1 - \mu_2 = 0$, 有显著变化

5. 假设: $H_0: \sigma_1^2 \leq \sigma_2^2; H_1: \sigma_1^2 > \sigma_2^2$, 拒绝域

$$\frac{S_1^2}{S_2^2} > F_a(m-1, n-1), F_a(m-1, n-1) = 1.61, \frac{S_1^2}{S_2^2} = 1.6 < F_a(m-1, n-1), \text{ 所以接受}$$

$H_0: \sigma_1^2 \leq \sigma_2^2$

习题 7.4

1. （本题中列表的方法采用了浙大版教材中的方法，有兴趣可以参阅）按题意需检验假设

$$H_0: P(X=i) = \frac{1}{6}, i=1\dots 6, \text{ 列一个表}$$

X	f_i	P_i	nP_i	f_i^2/np_i
1	8	1/6	10	6.4
2	8	1/6	10	6.4
3	12	1/6	10	14.4
4	11	1/6	10	12.1
5	9	1/6	10	8.1
6	12	1/6	10	14.4

$$\chi^2 = \sum_{i=1}^6 \frac{f_i^2}{np_i} - n = 1.8, \quad \chi_a^2(6-1) = 11.07 \geq \chi^2, \text{ 接受 } H_0: P(X=i) = \frac{1}{6}, i=1\dots 6, \text{ 则这颗}$$

骰子均匀、对称。

2. 按题意需检验假设

$$H_0: P(X=k) = \frac{\lambda^k e^{-\lambda}}{k!}, k=1\dots 6$$

首先咱们先给 λ 做极大似然估计, $\hat{\lambda} = \bar{x} = \frac{\sum_{i=0}^7 f(n_i)n_i}{100} = 1$

X	f_i	P_i	nP_i	f_i^2/np_i
0	36	$\frac{\lambda^0 e^{-\lambda}}{0!} = \frac{1}{e}$	36.79	35.23
1	40	$\frac{1}{e}$	36.79	43.49
2	19	$\frac{1}{2e}$	18.39	19.63
3	2	$\frac{1}{3! e}$	8.03	3.11

4	0	$\frac{1}{4! e}$		
5	2	$\frac{1}{5! e}$		
6	1	$\frac{1}{6! e}$		
≥ 7	0	$1 - \frac{1}{e} \sum_{i=0}^6 \frac{1}{i!}$		

$$\chi^2 = \sum_{i=1}^4 \frac{f_i^2}{np_i} - n = 1.46, \quad \chi_a^2(100-1-1) = \frac{1}{2}(z_a + \sqrt{195})^2 \geq \chi^2, \text{ 接受}$$

$$H_0: P(X=k) = \frac{\lambda^k e^{-\lambda}}{k!}, k=1 \dots 6, \text{ 则一页的错误个数服从泊松分布。}$$

习题七

1. 假设 $H_0: \mu \leq 30; H_1: \mu > 30$, (这里在选取 H_0 时候想想:把不符合要求的当作符合要求的问题比把符合要求的舍弃问题更大, 所以应该避免前一种错误, 即避免在 $H_0: \mu \leq 30$ 的时候拒绝, 认为 $\mu > 30$) σ 未知, 则拒绝域为

$$\frac{\bar{X} - \mu_0}{S/\sqrt{n}} > t_a(n-1), \frac{\bar{X} - \mu_0}{S/\sqrt{n}} = -1.73 < t_a(n-1) = 2.015, \text{ 所以拒绝 } H_0: \mu \leq 30, \text{ 有理由接受符合要求。}$$

2. 假设 $H_0: \mu = 66; H_1: \mu \neq 66$, σ 其实是未知的, 因为你不知道生病的影响, 则拒绝域

$$\text{为 } \left| \frac{\bar{X} - \mu_0}{S/\sqrt{n}} \right| > t_{\frac{\alpha}{2}}(n-1), \frac{\bar{X} - \mu_0}{S/\sqrt{n}} = 2.066 < t_{\frac{\alpha}{2}}(n-1) = 2.1315, \text{ 所以接受 } H_0: \mu = 66,$$

可以断定脉搏的速度没有变化。衡量他有没有恢复, 不仅要看平均的频率, 还要看稳定性。

$$\text{假设 } H_0: \sigma = 5; H_1: \sigma \neq 5, \text{ 则接受域为 } \chi_{1-\frac{\alpha}{2}}^2(n-1) \leq \chi^2 = \frac{(n-1)S^2}{\sigma_0^2} \leq \chi_{\frac{\alpha}{2}}^2(n-1),$$

$$5.629 \leq \chi^2 \leq 26.119, \frac{(n-1)S^2}{\sigma_0^2} = 21.0 \quad \text{所以接受 } H_0: \sigma = 5, \text{ 可以断定脉搏的稳定性}$$

没有变化。综上, 身体已经恢复到受伤前状态。

- 3.

(1) σ^2 已知, 则 μ 的双侧置信区间为 $(\bar{X} - \frac{\sigma}{\sqrt{n}} z_{\frac{\alpha}{2}}, \bar{X} + \frac{\sigma}{\sqrt{n}} z_{\frac{\alpha}{2}})$, 即 (2.11, 2.14)

(2) 不在上面的置信区间内, 所以认为平均长度与 $\mu_0=2.15$ 有明显差异。

4.

(1) σ^2 已知, 则单侧上限为 $\bar{X} + \frac{\sigma}{\sqrt{n}} z_{\alpha} = 69.6$

(2) $\mu = 72$ 不在 $\alpha = 0.05$ 的置信区间内, 所以认为明显低于一般健康成年男子。

5. σ_1^2, σ_2^2 已知, 假设: $H_0: \mu_1 - \mu_2 = 0; H_1: \mu_1 - \mu_2 \neq 0$; 则拒绝域为

$$|\frac{\bar{X} - \bar{Y}}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}| \geq z_{\frac{\alpha}{2}}, |\frac{\bar{X} - \bar{Y}}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}| = 2.39 < z_{\frac{\alpha}{2}} = 2.57, \text{ 则接受 } H_0: \mu_1 - \mu_2 = 0, \text{ 含灰率平}$$

均值没有显著差异。

6. σ_1^2, σ_2^2 未知, $\sigma_1^2 = \sigma_2^2$, 假设: $H_0: \mu_1 - \mu_2 > 0; H_1: \mu_1 - \mu_2 \leq 0$; 则拒绝域为

$$\frac{\bar{X} - \bar{Y}}{S_w \sqrt{\frac{1}{m} + \frac{1}{n}}} \leq -t_{\alpha}(m+n-2), \frac{\bar{X} - \bar{Y}}{S_w \sqrt{\frac{1}{m} + \frac{1}{n}}} = 10.8 > -t_{\alpha}(m+n-2) \approx -z_{\alpha} = -1.64, \text{ 则接受}$$

$H_0: \mu_1 - \mu_2 > 0$, 甲方案明显高于乙种方案。

7. 假设: $H_0: \sigma_1^2 = \sigma_2^2; H_1: \sigma_1^2 \neq \sigma_2^2$ 则接受域为

$$F_{1-\frac{\alpha}{2}}(m-1, n-1) \leq \frac{S_1^2}{S_2^2} \leq F_{\frac{\alpha}{2}}(m-1, n-1) \text{ (为什么我在这里会写接受域? 因为这是接受}$$

域比拒绝域好写一点, 看个人意愿, 标准是拒绝域), 即

$$F_{1-\frac{\alpha}{2}}(m-1, n-1) \leq \frac{S_1^2}{S_2^2} \leq F_{\frac{\alpha}{2}}(m-1, n-1), \text{ (这里我没有查到}$$

$F_{0.025}(49, 51), F_{0.975}(49, 51)$ 的值, 遂直接套用答案), 拒绝 $H_0: \sigma_1^2 = \sigma_2^2$, 认为有差异。

8.

(1) 拒绝域: $|\frac{\bar{X} - 5}{4/2}| > z_{\frac{\alpha}{2}}$, 即 $\bar{X} < 1.08$ 或 $\bar{X} > 8.92$

(2) 第二类错误即 $\mu = 6, 1.08 < \bar{X} < 8.92$ 即 $P(-2.46 < \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} < 1.46) = 0.92$

9. 我们先来算假设的接受域（其实就是置信区间） $(\mu_0 - \frac{\sigma}{\sqrt{n}} z_{\frac{\alpha}{2}}, \mu_0 + \frac{\sigma}{\sqrt{n}} z_{\frac{\alpha}{2}})$ ，然后在

$\mu = \mu_1 < \mu_0$ 时，

$$\begin{aligned} P(\mu_0 - \frac{\sigma}{\sqrt{n}} z_{\frac{\alpha}{2}} < \bar{X} < \mu_0 + \frac{\sigma}{\sqrt{n}} z_{\frac{\alpha}{2}}) &= P(\frac{\mu_0 - \mu_1}{\sigma/\sqrt{n}} - z_{\frac{\alpha}{2}} < \frac{\bar{X} - \mu_1}{\sigma/\sqrt{n}} < \frac{\mu_0 - \mu_1}{\sigma/\sqrt{n}} + z_{\frac{\alpha}{2}}) \\ &= \Phi(\frac{\mu_0 - \mu_1}{\sigma/\sqrt{n}} + z_{\frac{\alpha}{2}}) - \Phi(\frac{\mu_0 - \mu_1}{\sigma/\sqrt{n}} - z_{\frac{\alpha}{2}}) \end{aligned}$$

（为什么和答案不一样？）答案对假设检验问题的接受域认为是 $(\mu_0 - \frac{\sigma}{\sqrt{n}} z_{\alpha}, \infty)$ （其实就是

是根据了后面的情况，默认如果 $\mu \neq \mu_0$ ，那就是 $\mu < \mu_0$ ，笔者觉得这样是不合理的）

10. 这里实际上是给了一个暗含的东西（笔者觉得这样不好） $H_0: \mu \geq 20; H_1: \mu < 20$ 内马

尔接受域 $\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \geq -z_{\alpha}$ ， $\bar{X} \geq \mu - \frac{z_{\alpha}}{\sqrt{n}} \sigma$ 犯第二类错误 $\mu \leq 18$ ，在接受域内

$\bar{X} \geq \mu_1 - \frac{z_{\alpha}}{\sqrt{n}} \sigma$ ，（这里加注 μ_1 指的是这个 $\mu_1 \geq 20$ ）

$$P(\frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} \geq \frac{\mu_1 - \mu_0}{\sigma/\sqrt{n}} - z_{\alpha}) \leq 0.025, \frac{\mu_1 - \mu_0}{\sigma/\sqrt{n}} - z_{\alpha} \geq z_{\beta}, n \geq (\frac{z_{\alpha} + z_{\beta}}{\mu_1 - \mu_0} \sigma)^2 \geq 24.01, n \geq 25$$

（这里加注 μ_0 指的是这个 $\mu_0 \leq 18$ ）

11. 犯第一类错误的概率应该是在 $\mu = 1$ 时，却 $\bar{X} > 1, P(\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} > 0) = \frac{1}{2}$ 。而第二类错误

就是在 $\mu = 2$ 时，在接受域 $\bar{X} \leq 1, P(\frac{\bar{X} - 2}{1/3} \leq \frac{1-2}{1/3}) = \Phi(-3) = 0.0013$

12. 按题意需检验假设

H_0 :服从该指数分布，列一个表

A_i	f_i	P_i	nP_i	f_i^2/nP_i
$A_1(0 \leq t \leq 100)$	121	0.393	117.9	124.18
$A_2(100 < t \leq 200)$	78	0.239	71.7	84.85

$A_3(200 < t \leq 300)$	43	0.145	43.5	42.5
$A_4(t > 300)$	58	0.223	66.9	50.28

$$\chi^2 = \sum_{i=1}^4 \frac{f_i^2}{np_i} - n = 1.81, \quad \chi_a^2(4-1) = 7.815 \geq \chi^2, \text{ 接受 } H_0: \text{服从该指数分布}$$

13. 列一个表

A_i	f_i	P_i	nP_i	f_i^2/np_i
$A_1(X=0)$	1	$\frac{1}{56}$	2(和下面合并)	
$A_2(X=1)$	31 (合并前) 32 (合并后)	$\frac{15}{56}$	30 (合并前) 32 (合并后)	32
$A_3(X=2)$	55	$\frac{15}{28}$	60	50.42
$A_4(X=3)$	25	$\frac{5}{28}$	20	31.25

$$\chi^2 = \sum_{i=1}^3 \frac{f_i^2}{np_i} - n = 1.67, \quad \chi_a^2(3-1) = 5.991 \geq \chi^2, \text{ 接受 } H_0: \text{红球个数为 5}$$

第八章

习题 8.1

$$1. \quad \hat{b} = \frac{L_{XY}}{L_{XX}} = 0.058, \hat{a} = \bar{y} - \hat{b}\bar{x} = 24.628, \quad \hat{Y} = 24.628 + 0.058x$$

$$2. \quad \text{同上}, \quad \hat{Y} = 0.1426 + 0.8662x$$

3.

(1) 先求一些数据, $L_{XX} = 438, L_{XY} = -1643, L_{YY} = 6278$, 那么

$$\hat{b} = \frac{L_{XY}}{L_{XX}} = -3.751, \hat{a} = \bar{y} - \hat{b}\bar{x} = 480.1945, \quad \hat{Y} = 480.1945 - 3.751x$$

$$(2) \quad \hat{\sigma}^2 = \frac{1}{n-2} (L_{YY} - \hat{b}L_{XY}) = 19.1845$$

4.

- (1) 假设: $H_0: b=0; H_1: b \neq 0$; 那么利用 T 检验去求拒绝域:

$$|T| \geq t_{\frac{\alpha}{2}}(n-2), |T| = \frac{|\hat{b}|}{\hat{\sigma}^*} \sqrt{L_{xx}} = 17.92 > t_{\frac{\alpha}{2}}(n-2) = 3.7074, \text{ 回归效果显著}$$

- (2) 为了对 b 进行估计, 我们直接选用 P246 页定理 8.1.1 中 (4):

$$t = \frac{\hat{b} - b}{\hat{\sigma}} \sqrt{L_{xx}} \sim t(n-2), \text{ 则通过计算, } b \in (b_0 - \frac{t_{0.025}(8-2)\hat{\sigma}}{\sqrt{L_{xx}}}, b_0 + \frac{t_{0.025}(8-2)\hat{\sigma}}{\sqrt{L_{xx}}})$$

即 $(-4.2631, -3.2389)$

- (3) 这里就直接套用 P248 页下面的结论,

$$Y \in (\hat{Y}_0 - t_{\frac{\alpha}{2}}(n-2) \cdot \hat{\sigma} \sqrt{1 + \frac{1}{n} + \frac{(x_0 - \bar{x})^2}{L_{xx}}}, \hat{Y}_0 + t_{\frac{\alpha}{2}}(n-2) \cdot \hat{\sigma} \sqrt{1 + \frac{1}{n} + \frac{(x_0 - \bar{x})^2}{L_{xx}}})$$

$$\hat{Y}_0 = 198.8695, t_{\frac{\alpha}{2}}(6) = 2.4469, \hat{\sigma} = 4.3800 \text{ 代入得 } Y \in (187.1582, 210.5808)$$

5. 套用 P246 页公式, $\hat{b} \sim N(b, \frac{\sigma^2}{L_{xx}})$, 那么 $\hat{\hat{b}} \sim N(b, \frac{\sigma^2}{L_{xx}n})$, 置信区间

$$(b - \frac{\sigma/\sqrt{L_{xx}}}{\sqrt{n}} z_{\frac{\alpha}{2}}, b + \frac{\sigma/\sqrt{L_{xx}}}{\sqrt{n}} z_{\frac{\alpha}{2}}), \text{ 即}$$

- (1) 同上, 先求一些数据,

$$\bar{x} = 0.543, \bar{y} = 20.771, L_{xx} = 0.532, L_{xy} = 6.678, L_{yy} = 84.034, \text{ 则}$$

$$\hat{b} = \frac{L_{xy}}{L_{xx}} = 12.55, \hat{a} = \bar{y} - \hat{b}\bar{x} = 13.96, \hat{Y} = 13.96 + 12.55x$$

- (2) $\hat{\sigma}^{*2} = \frac{1}{7-2}(L_{yy} - \hat{b}L_{xy}) = 0.045$

- (3) 利用简单的 t 检验 $|T| \geq t_{\frac{\alpha}{2}}(n-2), |T| = \frac{|\hat{b}|}{\hat{\sigma}^*} \sqrt{L_{xx}} = 43.15 > t_{\frac{\alpha}{2}}(n-2) = 2.5706,$

回归效果显著

- (4) $(\hat{a} + \hat{b}x - \delta(x), \hat{a} + \hat{b}x + \delta(x)), \delta(x) = t_{\frac{\alpha}{2}}(n-2) \cdot \hat{\sigma}^* \sqrt{1 + \frac{1}{n} + \frac{(x_0 - \bar{x})^2}{L_{xx}}} = 0.584,$

区间为 $(19.651, 20.819)$ (与答案的误差都来源于对小数位数的选取)

习题 8.2

1. $\ln Y = \ln a - \frac{b}{x}$, 令 $Y' = \ln Y, x' = \frac{1}{x}, a' = \ln a, b' = -b$, 则 $Y' = a' + b'x'$, 对表格数据进行处理, 最后得 $\hat{Y} = 1.73e^{-\frac{0.146}{x}}$
2. $\ln Y = \ln a + bt$, 令 $Y' = \ln Y, Y' = a + bt$, 对数据 Y 进行处理, 然后即可得
 $\hat{y}' = 5.9665 - 0.2179t$
3. $\frac{1}{y} = b + \frac{a}{x}$, 则 $y' = \frac{1}{y}, x' = \frac{1}{x}$ 即可。

习题 8.3

1. 提出假设: $H_0: \mu_1 = \mu_2 = \mu_3$, (这里详细的数据处理我就不做了, 我利用 excel 帮大家检查一下答案好了) $F = 3.188 < F_{0.05}(3, 9) = 3.862$, 认为没有显著差异
2. 同上, $F = 5.004 > F_{0.05}(2, 13) = 3.885$, 有差异
3. $F = 1.17 < F_{0.1}(4, 20) = 2.25$, 没有显著差异

习题 8.4

1. 提出假设: $H_{01}: \alpha_1 = \alpha_2 = \alpha_3 = 0; H_{02}: \beta_1 = \beta_2 = \beta_3 = 0;$
 $F_A = 0.917 < F_{Aa} = 5.14; F_B = 0.43 < F_{Ba} = 4.76$, 所以都没有显著差异。
2. $F_A = 0.405 < F_{Aa} = 3.84; F_B = 0.449 < F_{Ba} = 4.46$, 所以都没有显著差异。
3. $F_A = 70.047 > F_{Aa} = 5.143; F_B = 21.924 > F_{Ba} = 4.757$, 所以都没有显著差异。

习题八

略(同质化过于严重, 笔者实在不想做了)