

$$1. \int_a^b p(x; a, b) dx = \int_a^b \frac{1}{b-a} dx = \frac{b-a}{b-a} = 1 \quad (a \neq b)$$

So this distribution is normalized.

$$\text{Mean: } \mu_x = \int_a^b x p(x; a, b) dx = \int_a^b \frac{x}{b-a} dx = \frac{1}{2} x \frac{b^2 - a^2}{b-a} = \frac{a+b}{2}$$

$$\begin{aligned} \text{Variance: } \text{Var}(x) &= \mu_{x^2} - (\mu_x)^2 \\ &= \int_a^b x^2 p(x; a, b) dx - \left(\frac{a+b}{2}\right)^2 \\ &= \frac{b^3 - a^3}{3(b-a)} - \frac{a^2 + 2ab + b^2}{4} \\ &= \frac{b^2 + ab + a^2}{3} - \frac{a^2 + 2ab + b^2}{4} \\ &= \frac{(a-b)^2}{12} \end{aligned}$$

2. (a) Let event A = The neighbor has at least one boy.
event B = The neighbor has at least one girl.

$$P(B|A) = \frac{P(AB)}{P(A)} = \frac{\frac{1}{2}}{\frac{1}{2} + \frac{1}{4}} = \frac{2}{3}$$

(b) Let event A = We see one of his children run by, it is a boy.
Let event B = The neighbor has at least one girl.

For event A, since we already seen it is a boy, there could be only three possible event, Boy & Girl, Girl & Boy, Boy & Boy.

$$P(A) = \frac{4}{2 \times 3} = \frac{2}{3} \quad P(AB) = \frac{2}{6} = \frac{1}{3}$$

$$P(B|A) = \frac{P(AB)}{P(A)} = \frac{\frac{1}{3}}{\frac{2}{3}} = \frac{1}{2}$$

$$3. \text{var}[X+Y]$$

$$= \text{cov}[X+Y, X+Y]$$

$$= \text{cov}[X, X] + \text{cov}[X, Y] + \text{cov}[Y, X] + \text{cov}[Y, Y]$$

$$= \text{var}[X] + \text{var}[Y] + 2\text{cov}[X, Y]$$

$$4. \text{cov}[X, Y] = E[(X - E[X])(Y - E[Y])]$$

$$= E[XY - XE[Y] - YE[X] + E[X]E[Y]]$$

$$= E[XY] - E[X]E[Y] - E[Y]E[X] + E[X]E[Y]$$

$$= E[XY] - E[X]E[Y]$$

Since X, Y are independent. $E[XY] = E[X]E[Y]$.

$$\text{So } \text{cov}[X, Y] = E[X]E[Y] - E[X]E[Y] = 0$$

their covariance is zero.

5. (a) proof: since X, Z are independent.

$$E[X+Z] = \int (x+z) dF = \int x dF + \int z dF = E[X] + E[Z]$$

$$(b) \text{ proof: } \text{var}[X+Z] = \text{var}[X] + \text{var}[Z] + 2\text{cov}[X, Z]$$

Since X, Z are independent $\text{cov}[X, Z] = 0$

$$\text{var}[X+Z] = \text{var}[X] + \text{var}[Z]$$

6. apply $x_1=0, x_2=n$ on equation 1,

$$\Rightarrow \tilde{r}(0) + \alpha \tilde{r}(n) = \tilde{r}(\tilde{r}(n)).$$

apply $x_1=1, x_2=n-1$ on equation 1,

$$\Rightarrow \tilde{r}(1) + \alpha \tilde{r}(n-1) = \tilde{r}(\tilde{r}(n-1)).$$

Guess $\tilde{r}(x) = \alpha x$

Let's verify it

$$\tilde{r}(\alpha x_1) + \alpha \tilde{r}(x_2) = \alpha^2 x_1 + \alpha^2 x_2 = \alpha^2 (x_1 + x_2)$$

$$\tilde{r}(\tilde{r}(x_1 + x_2)) = \tilde{r}(\alpha(x_1 + x_2)) = \alpha^2 (x_1 + x_2) = \tilde{r}(\alpha x_1) + \alpha \tilde{r}(x_2).$$

$\Rightarrow \tilde{r}(x) = \alpha x$ satisfies equation 1.

After prediction, G should be $G(x) = \alpha x + \epsilon$

So, the prediction would be $G(\hat{x}) = \alpha \hat{x} + \epsilon$, where α and ϵ could be calculated after testing all functions $f \in C$.

7. Suppose there exists such A . Let's consider two cases:

① each A_i contains only 1 point. In this case, since \mathbb{R} is uncountable, A is uncountable, which is a contradiction. ② some A_i contains at least 2 points, say $A_i = \{x, y\}$. Then $\{x\}$ is a Borel Set which can not be written as a union of elements in A .

So for all 2 cases, Borel σ -algebra $B_{\mathbb{R}}$ is not atomic.

8. (a) $\Omega = \{H, TH, TTH, \dots, T \dots TH\}$.

$$\mathcal{F} = \{\emptyset, \{H\}, \{TH\}, \{TTH\}, \dots, \{T \dots TH\}, \{H, TH\}, \{H, TTH\}, \dots, \Omega\} = 2^\Omega$$

Proof: According to the definition of σ -algebra.

① $\emptyset \in \mathcal{F}$

② $\forall A \in \mathcal{F}, A \subset \Omega$. Then we have $A^c \subset \Omega$.

Since \mathcal{F} contains all subset of Ω , $A^c \in \mathcal{F}$.

③ Given $A_i \in \mathcal{F}$, $A_i \subset \Omega$. $\bigcup_{i=1}^{\infty} A_i \subset \Omega$, so $\bigcup_{i=1}^{\infty} A_i \in \mathcal{F}$

(b) \mathcal{F} is uncountable.

There is a bijection between natural numbers and elements in Ω .

ω	H	TH	TTH	...	$\overbrace{T \dots T}^i$ TH
N	0	1	2	...	i

Suppose \mathcal{F} is countable. If we find a bijection between N and 2^N , there should be a bijection between N and 2^N .

$2^N \backslash N$	0	1	2	3	4	...
2^0	0	1	0	0	...	
2^1	1	0	0	0	...	
2^2	0	0	1	1	...	
2^3	1	0	0	1	...	
\vdots						

Suppose the table on the left is the bijection table, where 2^N_i means i th element in 2^N . In this table, i th row, j th column = 1 means $j \in 2^N_i$, else $j \notin 2^N_i$. Take diagonal elements as a sequence, flip each element i to $1-i$, we'll get a new sequence, whose i th element is different from the i th element in the i th row.

While from the definition of this table, all sequences should be included, where is a contradiction.

So 2^{Ω} is uncountable.

(c) Let $w_i = \underbrace{TT \cdots T}_i H$, $P(w_i) = \frac{1}{2^{i+1}}$

$w_0 = H$, $P(w_0) = \frac{1}{2}$.

For \mathcal{F} , $\forall A_i \in \mathcal{F}$, A_i is the Union of some w_j 's in Ω .

$0 \leq P(A_i) = \sum_j P(w_j) \leq 1$ $P(\emptyset) = 0$, $P(\Omega) = 1$

Justify: ① Since $P(w_i) > 0$, $P(A_i) \geq 0$.

② $P(\emptyset) = 0$

③ $P(\bigcup_{i=0}^{\infty} w_i) = \sum_{i=0}^{\infty} P(w_i)$ (which is the definition of our P)

So P is probability measure.

(d) Let $X(w_i) = \sum_{k=0}^i P(w_k)$

$\lim_{\substack{x \rightarrow +\infty \\ n \rightarrow +\infty}} P(X \leq x) = \lim_{n \rightarrow +\infty} \sum_{i=0}^n P(w_i) = \lim_{n \rightarrow +\infty} \frac{1}{2} + \frac{1}{4} + \cdots + \frac{1}{2^n} = \lim_{n \rightarrow +\infty} (1 - \frac{1}{2^n}) = 1$

$E(x) = \sum_i P(w_i) = \sum_i (P(w_i))^2 = \lim_{n \rightarrow +\infty} (\frac{1}{2})^2 + (\frac{1}{2^2})^2 + (\frac{1}{2^3})^2 + \cdots + (\frac{1}{2^n})^2$
 $= \lim_{n \rightarrow +\infty} \frac{\frac{1}{4} \times (1 - \frac{1}{2^{2n}})}{1 - \frac{1}{4}} = \frac{1}{3}$

9. $\liminf_{n \rightarrow \infty} A_n = \bigcap_{k=1}^{\infty} \bigcup_{n \geq k} A_n.$

Let $B_k = \bigcap_{n \geq k} A_n \Rightarrow \liminf_{n \rightarrow \infty} A_n = \bigcup_{k=1}^{\infty} B_k$

Since B_k is increasing sequence. and $B_k \subset A_n, \forall n \geq k.$

($B_k = B_{k+1} \cap A_{k+1}$)

$P(B_k) \leq \inf_{n \geq k} P(A_n)$

$\Rightarrow P(\liminf_{n \rightarrow \infty} A_n) = P(\bigcup_{k=1}^{\infty} B_k) = \lim_{n \rightarrow \infty} P(B_n) \leq \liminf_{n \rightarrow \infty} \inf_{k \geq n} P(A_k) = \liminf_{n \rightarrow \infty} P(A_n)$