Chan Park



# (Efficient) Influence Function

$$\bullet \ O^F=(Y^{(0)},Y^{(1)},A,L) \ \text{and} \ \tau^*=E\big\{\Psi^F(O^F)\big\}\in\mathbb{R}^p$$

• 
$$O=(Y,A,L)$$
 and  $\tau^*=Eig\{\Psi(O)ig\}\in\mathbb{R}^p$  under causal assumptions

• 
$$\mathcal{M} = \{P(O) \mid \text{regular law of } O; \text{ can be parametric, semiparametric, nonparametric}\}, P^* = \text{true law} \in \mathcal{M}$$
  $\eta^*$ : nuisance function at  $P^*$ 

• 
$$\mathcal{M}_t = \{P_t(O) \mid \text{parametric submodel of } \mathcal{M} \text{ indexed by } t\}, \ P_0 = P^*$$

$$\tau_t = E_t \{\Psi_t(O)\}; \quad \eta_t = \text{nuisance function at } P_t; \quad s_O(O;t) = \frac{\partial f(O;t)/\partial t}{f(O;t)} = \text{score function}$$

• Influence function IF $(O; \eta, \tau)$  solves  $\frac{\partial \tau_t}{\partial t}\Big|_{T=0} = E\big\{ \text{IF}(O; \eta^*, \tau^*) s_O(O; t=0) \big\}$ 

• Example: Average treatment effect with 
$$\mathcal{M}=$$
 nonparametric model 
$$\tau_{\mathsf{ATE}}^* = E\big\{Y^{(1)} - Y^{(0)}\big\} = E\big\{E(Y\,\big|\,A=1,X) - E(Y\,\big|\,A=0,X)\big\} \text{ under consistency, ignorability, positivity } \\ \eta^* = (\mu^*,e^*) \text{ where } \mu^*(a,X) = E(Y\,\big|\,A=a,X) \text{ and } e^*(a,X) = \Pr(A=a\,\big|\,X)$$

$$\mathrm{IF}(O; \eta^*, \tau_{\mathsf{ATE}}^*) = \frac{A\{Y - \mu^*(1, X)\}}{e^*(1, X)} - \frac{(1 - A)\{Y - \mu^*(0, X)\}}{e^*(0, X)} + \left\{\mu^*(1, X) - \mu^*(0, X)\right\} - \tau_{\mathsf{ATE}}^*$$

# Decomposition

$$\begin{split} \bullet & \ \widehat{\tau} = N^{-1} \sum_{i=1}^{N} \widetilde{\mathrm{IF}}(O_i; \widehat{\eta}) \\ & \sqrt{N}(\widehat{\tau} - \tau^*) \\ & = \frac{1}{\sqrt{N}} \sum_{i=1}^{N} \left[ \widetilde{\mathrm{IF}}(O_i; \widehat{\eta}) - E \{ \widetilde{\mathrm{IF}}(O_i; \eta^*) \} \right] \\ & = \frac{1}{\sqrt{N}} \sum_{i=1}^{N} \left[ \widetilde{\mathrm{IF}}(O_i; \eta^*) - E \{ \widetilde{\mathrm{IF}}(O_i; \eta^*) \} \right] & \stackrel{D}{\to} N(0, \sigma^2) \\ & + \frac{1}{\sqrt{N}} \sum_{i=1}^{N} \left[ \widetilde{\mathrm{IF}}(O_i; \widehat{\eta}) - E \{ \widetilde{\mathrm{IF}}(O_i; \widehat{\eta}) \} - \widetilde{\mathrm{IF}}(O_i; \eta^*) + E \{ \widetilde{\mathrm{IF}}(O_i; \eta^*) \} \right] & \Rightarrow \text{Empirical process} \overset{Want}{=} o_P(1) \\ & + \frac{1}{\sqrt{N}} \sum_{i=1}^{N} \left[ E \{ \widetilde{\mathrm{IF}}(O_i; \widehat{\eta}) \} - E \{ \widetilde{\mathrm{IF}}(O_i; \eta^*) \} \right] & \Rightarrow \text{Bias} \overset{Want}{=} o_P(1) \end{split}$$

- $\hat{\eta}$  and  $\sum_{i=1}^{N}$  share the same observation  $\Rightarrow$  Unclear to characterize EP and Bias
- Cross-fitting

 $\hat{\eta}^{(-k)}$  (estimate nuisance ft without using  $\mathcal{I}_k$ ) and  $\sum_{i \in \mathcal{I}_k}$  (evaluate nuisance ft over  $\mathcal{I}_k$ )

Clearer characterization of EP and Bias (i.e. show  $o_P(1)$  easily (?)) for any generic ML learners

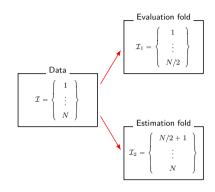
• References: [1, 4] <sup>1</sup>

$$\mathcal{I} = \left\{ \begin{array}{c} 1 \\ \vdots \\ N \end{array} \right\}$$

<sup>&</sup>lt;sup>1</sup> Schick (1986) On asymptotically efficient estimation in semiparametric models. AoS

Chernozhukov, Chetverikov, Demirer, Duflo, Hansen, Newey, Robins (2018) Double/debiased machine learning for treatment and structural parameters. The Econometrics Journal

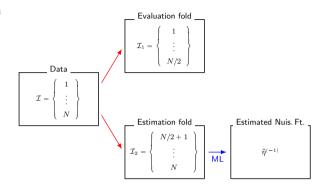
• References: [1, 4] <sup>1</sup>



<sup>&</sup>lt;sup>1</sup> Schick (1986) On asymptotically efficient estimation in semiparametric models. AoS

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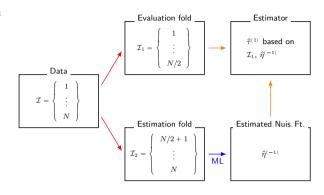
• References: [1, 4] 1



<sup>&</sup>lt;sup>1</sup> Schick (1986) On asymptotically efficient estimation in semiparametric models. AoS

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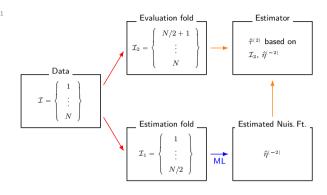
• References: [1, 4] 1



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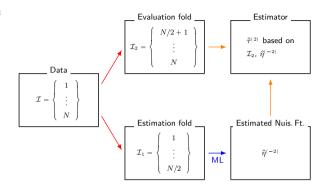
• References: [1, 4] 1



<sup>1</sup> Schick (1986) On asymptotically efficient estimation in semiparametric models. AoS

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• References: [1, 4] <sup>1</sup>

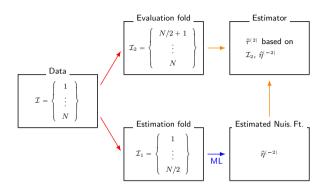


• 
$$\hat{\tau} = \mathsf{aggregate}(\hat{\tau}^{(1)}, \hat{\tau}^{(2)}) \ (\text{e.g.,} \ \frac{\hat{\tau}^{(1)} + \hat{\tau}^{(2)}}{2})$$

<sup>&</sup>lt;sup>1</sup> Schick (1986) On asymptotically efficient estimation in semiparametric models. AoS

Chernozhukov, Chetverikov, Demirer, Duflo, Hansen, Newey, Robins (2018) Double/debiased machine learning for treatment and structural parameters. The Econometrics Journal

• References: [1, 4] <sup>1</sup>



- $\hat{\tau} = \operatorname{aggregate}(\hat{\tau}^{(1)}, \hat{\tau}^{(2)}) \text{ (e.g., } \frac{\hat{\tau}^{(1)} + \hat{\tau}^{(2)}}{2})$
- ML: superlearner; [5]<sup>2</sup>, minimax estimation [2]<sup>3</sup>; Forster-Warmuth Counterfactual Regression [7]<sup>4</sup>

Chernozhukov, Chetverikov, Demirer, Duflo, Hansen, Newey, Robins (2018) Double/debiased machine learning for treatment and structural parameters. The Econometrics Journal

<sup>&</sup>lt;sup>1</sup>Schick (1986) On asymptotically efficient estimation in semiparametric models. AoS

 $<sup>^2</sup>$  van der Laan, Polley, Hubbard (2007). Super learner. Statistical Applications in Genetics and Molecular Biology

<sup>&</sup>lt;sup>3</sup> Ghassami, Ying, Shpitser, Tchetgen Tchetgen (2022). Minimax kernel machine learning for a class of doubly robust functionals with application to proximal causal inference. AISTAT

<sup>&</sup>lt;sup>4</sup>Yang, Kuchibhotla, Tchetgen Tchetgen (2024). Forster-Warmuth counterfactual regression: a unified learning approach. arXiv

## Superlearner by van der Laan et al. [5]

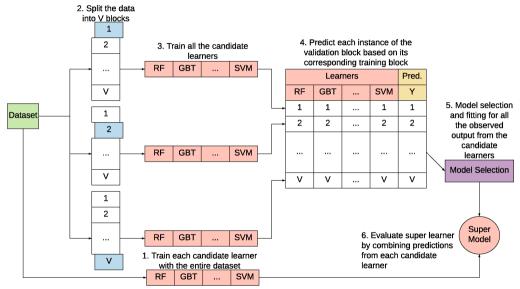


Figure: Reference: [3]

## Superlearner by van der Laan et al. [5]: with an R example

### https://github.com/qkrcks0218/CCI\_Crossfitting

```
SL.Fit ← MvSL(
                    Data = data.
                     locY = response variable columns.
                     locX = explanatory variable columns,
                     Ydist = gaussian() or binomial().
                     SL.list = SL.hpara$SLL.
                     MTRY = SL.hpara$MTRY,
                     MLPL = SL.hpara$MLPL,
                     MLPdecay=SL.hpara$MLPdecay,
                     NMN=SL.hpara$NMN)
                    SL.Fit
SL.Eval ← predict(
                     newdata = newdata[,locX],
                     onlvSL = TRUE)$pred
```

- # ML learners
- # hyperparameters for random forest
- $\ensuremath{\text{\#}}$  hyperparameters for multi-layer perceptron
- # hyperparameters for multi-layer perceptron
- # hyperparameters for GBM

## Minimax Estimation by Ghassami et al. [2]

•  $\tau_1^* = E\{Y^{(1)}\}$ 

• IF
$$(O; \nu^*, \pi^*, \tau_1^*) = A\pi^*(X)\{Y - \nu^*(X)\} + \nu^* - \tau_1^*$$
 where  $\pi^*(X) = 1/\Pr(A = 1 \mid X)$  and  $\nu^*(X) = \mu^*(1, X)$ 

From the DR property:

$$E\{\mathbf{IF}(O; \nu^*, \pi^*, \tau_1^*)\} = E\{\mathbf{IF}(O; \nu^*, \pi^{\dagger}, \tau_1^*)\} = E\{\mathbf{IF}(O; \nu^{\dagger}, \pi^*, \tau_1^*)\} = 0$$

$$\Rightarrow 0 = E\{\mathbf{IF}(O; \nu^*, \pi^{\dagger}, \tau_1^*) - \mathbf{IF}(O; \nu^*, \pi^*, \tau_1^*)\} = E[A\{\pi^{\dagger}(X) - \pi^*(X)\}\{Y - \nu^*(X)\}] = E[p(X)\{AY - A\nu^*(X)\}]$$

One can show

$$\nu^* = \arg\min\max_{p} E[p(X)\{AY - A\nu(X)\} - d^2(X)]$$

• We estimate  $\nu^*$ 

$$\hat{\nu}^{(-k)} = \underset{\nu \in \mathcal{H}_X}{\arg \min} \left[ \max_{p \in \mathcal{H}_X} \left[ \mathbb{P}^{(-k)} [p(X) \{AY - A\nu(X)\} - p^2(X)] - \lambda_p \|p\|_{\mathcal{H}_X}^2 \right] + \lambda_{\nu} \|\nu\|_{\mathcal{H}_X}^2 \right]$$

 $\mathcal{H}_X: \mathsf{RKHS} \ \Rightarrow \ \hat{\nu}^{(-k)}(x) = \ \sum \ \alpha_i K(x,X_i); \ \mathrm{see} \ [2] \ \mathrm{for} \ \mathrm{a} \ \mathrm{closed} - \mathrm{form} \ \mathrm{representation} \ \mathrm{of} \ \alpha_i$ 

Minimax Estimation by Ghassami et al. [2]

· Minimax estimator can be used to solve intergral equations

$$\begin{split} &g(v) = E\{f^*(W) \, \big| \, V = v\} \\ \Rightarrow & 0 = E\big[p(V)\big\{f^*(W) - g(V)\big\}\big] \ , \quad \forall p \\ \Rightarrow & \hat{f}^{(-k)} = \underset{f \in \mathcal{H}_W}{\arg\min} \left[ \underset{p \in \mathcal{H}_V}{\max} \left[\mathbb{P}^{(-k)}\big[p(V)\big\{f(W) - g(V)\big\} - p^2(V)\big] - \lambda_p \|p\|_{\mathcal{H}_V}^2\right] + \lambda_f \|f\|_{\mathcal{H}_W}^2 \right] \end{split}$$

## Minimax Estimation by Ghassami et al. [2]: with an R example

#### https://github.com/qkrcks0218/CCI\_Crossfitting

$$\widehat{\nu}^{(-k)} = \underset{\nu \in \mathcal{H}_X}{\arg\min} \left[ \max_{p \in \mathcal{H}_X} \left[ \mathbb{P}^{(-k)} \left[ p(X) \left\{ AY - A\nu(X) \right\} - p^2(X) \right] - \lambda_p \|p\|_{\mathcal{H}_X}^2 \right] + \lambda_{\nu} \|\nu\|_{\mathcal{H}_X}^2 \right]$$

$$\mathsf{Coef} = -A; \quad \mathsf{Intercept} = AY; \quad \mathsf{Perturb} = p(X); \quad \mathsf{Target} = \nu(X)$$

$$\widehat{\pi}^{(-k)} = \operatorname*{arg\,min}_{\pi \in \mathcal{H}_X} \Big[ \max_{q \in \mathcal{H}_X} \Big[ \mathbb{P}^{(-k)} \big[ q(X) \big\{ A \pi(X) - 1 \big\} - q^2(X) \big] - \lambda_q \|q\|_{\mathcal{H}_X}^2 \Big] + \lambda_\pi \|\pi\|_{\mathcal{H}_X}^2$$

$$\mathsf{Coef} = A; \quad \mathsf{Intercept} = -1; \quad \mathsf{Perturb} = q(X); \quad \mathsf{Target} = \pi(X)$$

## Multiplier Bootstrap

• Reference: [6]1

• 
$$\hat{\tau}$$
 solves

$$\begin{split} 0 &= \frac{1}{N} \sum_{i \in \mathcal{I}_k} \mathrm{IF}(O_i; \hat{\eta}^{(-k)}, \hat{\tau}) \\ \hat{\tau}_{\mathsf{ATE}} &= \frac{1}{N} \sum_{i \in \mathcal{I}_k} \left[ \frac{A_i \{Y_i - \hat{\mu}^{(-k)}(1, X_i)\}}{\hat{e}^{(-k)}(1, X_i)} - \frac{(1 - A_i) \{Y_i - \hat{\mu}^{(-k)}(0, X_i)\}}{\hat{e}^{(-k)}(0, X_i)} + \{\hat{\mu}^{(-k)}(1, X_i) - \hat{\mu}^{(-k)}(0, X_i)\} \right] \end{split}$$

• 
$$\sqrt{N}(\hat{\tau} - \tau^*) \stackrel{D}{\rightarrow} N(0, \sigma^2)$$

• A consistent estimator of  $\sigma^2$  and SE of  $\hat{\tau}$ :

$$\hat{\sigma}^2 = \frac{1}{N} \sum_{i,j} \left[ IF(O_i; \hat{\eta}^{(-k)}, \hat{\tau}) \right]^2 \quad \Rightarrow \quad \mathsf{SE}(\hat{\tau}) = \frac{\hat{\sigma}}{\sqrt{N}}$$

• A multiplier bootstrap standard error of  $\hat{\tau}$ :

$$\mathsf{BSE}(\hat{\tau}) = \mathsf{sd}(\hat{e}^{[1]}, \cdots, \hat{e}^{[B]}) \text{ where } \hat{e}^{[b]} = \frac{1}{N} \sum_{i} \epsilon_i^{[b]} \mathsf{IF}(O_i; \hat{\eta}^{(-k)}, \hat{\tau}) \quad \Leftarrow \quad \epsilon_i^{[b]} \stackrel{iid}{\sim} N(0, 1)$$

### Median Adjustment

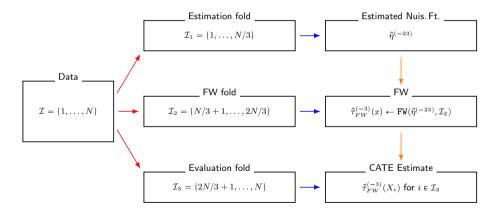
- Cross-fitting estimates depend on particular split samples
- Repeat cross-fitting S times, say S=100
- Obtain estimates  $\hat{\tau}^{[1]},\ldots,\hat{\tau}^{[S]}$  with the corresponding SE  $\hat{\sigma}^{2,[1]},\ldots,\hat{\sigma}^{2,[S]}$
- Median adjustment [1]1

$$\begin{split} \hat{\tau}^{(\text{report})} &= \mathop{\mathrm{median}}_{s=1,...,S} \hat{\tau}^{[s]} \\ \hat{\sigma}^{2,(\text{report})} &= \mathop{\mathrm{median}}_{s=1,...,S} \left[ \hat{\sigma}^{2,[s]} + \left\{ \hat{\tau}^{[s]} - \hat{\tau}^{(\text{report})} \right\}^2 \right] \end{split}$$

<sup>1</sup> Chernozhukov, Chetverikov, Demirer, Duflo, Hansen, Newey, Robins (2018) Double/debiased machine learning for treatment and structural parameters. The Econometrics Journal

# Forster-Warmuth Counterfactual Regression by Yang et al. [7]

- Reference: [7]<sup>1</sup>
- Estimand is now infinite-dimensional function  $\tau^*(x)$



<sup>&</sup>lt;sup>1</sup>Yang, Kuchibhotla, Tchetgen Tchetgen (2024). Forster-Warmuth counterfactual regression: a unified learning approach. arXiv

# Forster-Warmuth Counterfactual Regression by Yang et al. [7]

- $\tau^*(X)$ : target estimand (CATE)
- Find a good f satisfying  $E\{f(O;\eta^*)\,|\,X\}= au^*(X)$  (uncentered EIF in general)

$$\tau^*(X) = \mu^*(1,X) - \mu^*(0,X) \quad \Rightarrow \quad f(O;\eta^*) = \frac{A\{Y - \mu^*(1,X)\}}{e^*(1,X)} - \frac{(1-A)\{Y - \mu^*(0,X)\}}{e^*(0,X)} + \left\{\mu^*(1,X) - \mu^*(0,X)\right\}$$

- Using  $\mathcal{I}_1$ , get  $\hat{\eta}^{(-23)}$
- Using  $\mathcal{I}_2$ , define

$$\hat{\tau}_{FW}^{(-3)}(x) = \frac{\boldsymbol{\phi}^{\intercal}(x) \left[ \boldsymbol{\Phi}_{\mathcal{I}_{2}}^{\intercal} \boldsymbol{\Phi}_{\mathcal{I}_{2}} + \boldsymbol{\phi}(x) \boldsymbol{\phi}^{\intercal}(x) \right]^{-1} \left[ \boldsymbol{\Phi}_{\mathcal{I}_{2}}^{\intercal} \boldsymbol{f}_{\mathcal{I}_{2}}^{(-23)} \right]}{1 - \boldsymbol{\phi}^{\intercal}(x) \left[ \boldsymbol{\Phi}_{\mathcal{I}_{2}}^{\intercal} \boldsymbol{\Phi}_{\mathcal{I}_{2}} + \boldsymbol{\phi}(x) \boldsymbol{\phi}^{\intercal}(x) \right]^{-1} \boldsymbol{\phi}(x)}$$

where

 $\phi(x) = \begin{bmatrix} \phi_1(x) \equiv 1 \\ \phi_2(x) \\ \vdots \\ \phi_J(x) \end{bmatrix} \in \mathbb{R}^J \quad \Phi_{\mathcal{I}_2} = \begin{bmatrix} \phi^\intercal(X_i) \\ \vdots \\ \phi^\intercal(X_{2N/3}) \end{bmatrix} \in \mathbb{R}^{(N/3) \times J} \quad \boldsymbol{f}_{\mathcal{I}_2}^{(-23)} = \begin{bmatrix} f(O_i; \hat{\eta}^{(-23)}) \end{bmatrix}_{i \in \mathcal{I}_2} = \begin{bmatrix} f(O_{1+N/3}; \hat{\eta}^{(-23)}) \\ \vdots \\ f(O_{2N/3}; \hat{\eta}^{(-23)}) \end{bmatrix} \in \mathbb{R}^{N/3}$ 

•  $\phi$ : splines, polynomials,  $\sin/\cos$ , etc; J is chosen from cross-validation

Forster-Warmuth Counterfactual Regression by Yang et al. [7]: with an R example

- https://github.com/qkrcks0218/CCI\_Crossfitting
- Also check Yachong Yang's Github

https://github.com/Elsa-Yang98/Forster\_Warmuth\_counterfactual\_regression

#### References

- [1] Chernozhukov, V., Chetverikov, D., Demirer, M., Duflo, E., Hansen, C., Newey, W., and Robins, J. (2018). Double/debiased machine learning for treatment and structural parameters. *The Econometrics Journal*, 21(1):C1–C68.
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- [3] Neto, H. N. C., Lopez, M. A., Fernandes, N. C., and Mattos, D. M. (2020). Minecap: super incremental learning for detecting and blocking cryptocurrency mining on software-defined networking. Annals of Telecommunications, 75(3):121–131.
- [4] Schick, A. (1986). On asymptotically efficient estimation in semiparametric models. The Annals of Statistics, 14(3):1139-1151.
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- [7] Yang, Y., Kuchibhotla, A. K., and Tchetgen Tchetgen, E. (2024). Forster-warmuth counterfactual regression: A unified learning approach. Preprint arXiv:2307.16798.