Online supporting information for

Studying predator foraging mode and hunting success at the

individual level with an online videogame

Supporting Information: Materials and methods

Principal component analysis

We computed a principal component analysis (PCA) using the packages 'FactoMineR' version 2.4 (Lê et al. 2008) and 'factoextra' version 1.0.7 (Kassambara and Mundt 2020 Apr 1) in the R software version 4.0.4 (R Core Team, 2021), under a Windows 10 Home computer OS (version 21H2, OS build 19044.1466). The 'PCA()' function in 'FactoMineR' uses singular value decomposition. Before running the PCA, we divided each variable (except travel speed) by the match duration (seconds) to account for differences in game lenght among matches. We applied a square-root transformation to all the varialbes. They were then standardized to mean and unit variance (Z scores) before running the PCA. We then ranked the variables based on their relative contribution (in %) to a given principal component axis. To do so, for each variable, we used the ratio between their cos2 multiplied by 100 and the cos2 of the principal component.

Parametrization of the Bayesian multivariate mixed-model

We used a half-normal prior with mean of 0 and standard deviation of 1 for all the variance parameters, and applied a weakly informative LKJ prior with a shape parameter (n) set to 2 for the random effects variance-covariance matrices. We applied a Gaussian prior with mean of 0 and standard deviation of 2 for the fixed effects (i.e. prey travel speed x_1 and rate of space covered x_2) of all the submodels (i.e. for each response variable). We set the model to run 4 chains with 2500 iterations where samples were drawn at every 8 intervals (thinning), and to use the first 500 iterations as warmups. We visually inspected the trace plots, effective sample sizes, and residuals to assess convergence and stability. We also evaluated the model's prediction accuracy using posterior predictive checks. The rhat values of all the model's main parameters (intercept, slopes, interactions, variances, correlations) were lower than 1.05, while the ratio of effective samples sizes to total sample size was greater than 0.5 for all the parameters. The model computed a variance-covariance matrix (Ω_k) for each random effect (k). We extracted the among-environment, among-individual, and within-individual variance components and behavioral correlations using the function 'as_draws_df()' in the 'brms' package (for details, consult the code files on the GitHub repository : https://github.com/quantitative-ecologist/predator-foraging-mode-videogames). We could thus compute the mean of the variance and correlation sample values, and the HDP intervals to obtain their 95% credible intervals. The variance-covariance matrices were parametrized as :

$$\begin{bmatrix} en_{0y1,g} \\ en_{0y2,g} \\ en_{0y3,g} \\ en_{0y4,g} \end{bmatrix} = MVN(0, \Omega_{en}) : \text{where } \Omega_{en} = \begin{bmatrix} V_{en_{0y1}} & & & \\ Cov_{en_{0y1}en_{0y2}} & V_{en_{0y2}} & & \\ Cov_{en_{0y1}en_{0y3}} & Cov_{en_{0y1}en_{0y4}} & V_{en_{0y3}} & \\ Cov_{en_{0y2}en_{0y3}} & Cov_{en_{0y2}en_{0y4}} & Cov_{en_{0y3}en_{0y4}} & V_{en_{0y4}} \end{bmatrix}$$
(S1)

$$\begin{bmatrix} av_{0y1,h} \\ av_{0y2,h} \\ av_{0y3,h} \\ av_{0y4,h} \end{bmatrix} = MVN(0, \Omega_{av}) : \text{where } \Omega_{av} = \begin{bmatrix} V_{av_{0y1}} & & & \\ Cov_{av_{0y1}av_{0y2}} & V_{av_{0y2}} & & \\ Cov_{av_{0y1}av_{0y3}} & Cov_{av_{0y1}av_{0y4}} & V_{av_{0y3}} & \\ Cov_{av_{0y2}av_{0y3}} & Cov_{av_{0y2}av_{0y4}} & Cov_{av_{0y3}av_{0y4}} & V_{av_{0y4}} \end{bmatrix}$$
(S2)

$$\begin{bmatrix} id_{0y1,i} \\ id_{0y2,i} \\ id_{0y3,i} \\ id_{0y4,i} \end{bmatrix} = MVN(0, \Omega_{id}) : \text{where } \Omega_{id} = \begin{bmatrix} V_{id_{0y1}} & & & \\ Cov_{id_{0y1}id_{0y2}} & V_{id_{0y2}} & & \\ Cov_{id_{0y1}id_{0y3}} & Cov_{id_{0y1}id_{0y4}} & V_{id_{0y3}} & \\ Cov_{id_{0y2}id_{0y3}} & Cov_{id_{0y2}id_{0y4}} & Cov_{id_{0y3}id_{0y4}} & V_{id_{0y4}} \end{bmatrix}$$
(S3)

$$\begin{bmatrix} \varepsilon_{0y1,ghij} \\ \varepsilon_{0y2,ghij} \\ \varepsilon_{0y3,ghij} \\ \varepsilon_{0y4,ghij} \end{bmatrix} = MVN(0, \Omega_{\varepsilon}) : \text{where } \Omega_{\varepsilon} = \begin{bmatrix} V_{\varepsilon_{0y1}} & & & \\ Cov_{\varepsilon_{0y1}\varepsilon_{0y2}} & V_{\varepsilon_{0y2}} & & \\ Cov_{\varepsilon_{0y1}\varepsilon_{0y3}} & Cov_{\varepsilon_{0y1}\varepsilon_{0y4}} & V_{\varepsilon_{0y3}} & \\ Cov_{\varepsilon_{0y2}\varepsilon_{0y3}} & Cov_{\varepsilon_{0y2}\varepsilon_{0y4}} & Cov_{\varepsilon_{0y3}\varepsilon_{0y4}} & V_{\varepsilon_{0y4}} \end{bmatrix}$$
(S4)

where the diagonals represent the random effect (k) variance ($V_{k_{0y_n}}$) for each hunting trait (y_n), and the lower off-diagonals the covariance between the random effect intercepts for each combination of hunting traits (travel speed ($y1$), the rate of space covered ($y2$), the time spent guarding ($y3$), and the time before first capture ($y4$)).

We ran three additional multivariate mixed-models:

1. Different structure: without prey behavior as fixed effects
2. Same structure: only on novice players
3. Same structure: only on experienced players

For the first model, we investigated if acounting for prey behavior in the predator behavioral variation (i.e. removing prey as a control) had an effect on the variance parameters and the behavioral correlations. We used this approach to test the assertion that behavioral changes in prey behavior could modulate the tactics used by individual predators, as well as the effect of the game environments on predator behavior. We thus computed the same multivariate model, and excluded prey travel speed and rate of space covered as the linear fixed effects.

For the second and third models, we investigated if the relationships between the predator behaviors (i.e. their correlations) changed with player experience. To do so, we applied the same model procedure on two distinct datasets based on player experience by splitting the total experience of individual players from the main dataset in quartiles. This resulted in 75% of the players having played up to 31 matches, which represented ~5 hours of gameplay. Based on their cumulative number of matches, we then assigned novice players as those who played between 1 and 31 matches, and players with experience as those having above 31 cumulative matches played. Individuals who played above 31 matches could thus be found in both the novice and experienced players datasets. We split our data this way because if we separated players based on their total number of matches, we would have had matches for experienced players where they were still novice.

Parametrization of the Bayesian mixed-models for hunting success

We computed 4 different models to investigate the relationship between predator behavior, prey behavior, and their interactions with hunting success. For all models, we used a half-normal prior for the variance parameters with a mean of 0 and a standard deviation of 1, and a Gaussian prior with a mean of 0 and a standard deviation of 2 for the fixed effects. We setup the 4 models to run 4 chains using 2500 iterations where samples were drawn at every 8 intervals (thinning), and used the first 500 iterations for warmup. We visually inspected trace plots, effective sample sizes, and residuals to assess convergence and stability. We also evaluated the models' prediction accuracies using posterior predictive checks. The rhat values of the models' main parameters (intercept, slopes, interactions, variances) were lower than 1.05, while the ratio of

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effective samples sizes to total sample size was greater than 0.5 for all except the general intercept (0.43). We calculated the ICC_k estimate for each random effect (k) using the function as_draws_df() in the 'brms' package. To do so, we divided each random effect variance sample value by the total variance (following Nakagawa et al. 2017). We used the mean of the posterior variance parameters to compute the ICC values of each random effect, and used the HDP intervals to obtain their corresponding 95% credible intervals

For the first model, we aimed at investigating the linear relationship between hunting success and predator behavior exclusively. The model was parametrized as :

$$
\begin{aligned}
logit(P_{hij}) = (\beta_0 + env_{0,h} + id_{0,i} + \varepsilon_{0,hij}) \\
+ \ \beta_{1,pred}speed_{hi} \\
+ \ \beta_{2,pred}space_{hi} \\
+ \ \beta_{3,pred}guarding\ time_{hi} \\
+ \ \beta_{4,pred}time\ 1^{st}capture_{hi}
\end{aligned}
\tag{S5}
$$

We included prey behavior in the second model and kept the same structure as equation S5. The model was parametrized as :

$$
\begin{aligned}
logit(P_{hij}) = (\beta_0 + env_{0h} + id_{0i} + \varepsilon_{0,hij}) \\
+ \ \beta_{1,pred}speed_{hi} \\
+ \ \beta_{2,pred}space_{hi} \\
+ \ \beta_{3,pred}guarding\ time_{hi} \\
+ \ \beta_{4,pred}time\ 1^{st}capture_{hi} \\
+ \ \beta_{5,prey}speed'_{hi} \\
+ \ \beta_{6,prey}space'_{hi}
\end{aligned}
\tag{S6}
$$

In the third model, we were interested in investigating nonlinear relationships between hunting success and predator behavior exlusively. We thus added quadratic effects for each hunting trait. We also investigated how predator behaviors interact to affect hunting success. This model was computed as :

$$
\begin{aligned}
logit(P_{hij}) = & (\beta_0 + env_{0h} + id_{0i} + \varepsilon_{0,hij}) \\
& + \beta_{1,pred}speed_{hi} + \beta_{2,pred}space_{hi} \\
& + \beta_{3,pred}guarding\ time_{hi} + \beta_{4,pred}time\ 1^{st}capture_{hi} \\
& + \frac{1}{2}\gamma_{1,pred}speed_{hi}^2 + \frac{1}{2}\gamma_{2,pred}space_{hi}^2 \\
& + \frac{1}{2}\gamma_{3,pred}guarding\ time_{hi}^2 + \frac{1}{2}\gamma_{4,pred}time\ 1^{st}capture_{hi}^2 \\
& + \gamma_{1,pred}speed_{hi} \times space_{hi} \\
& + \gamma_{2,pred}speed_{hi} \times guarding\ time_{hi} \\
& + \gamma_{3,pred}speed_{hi} \times time\ 1^{st}capture_{hi} \\
& + \gamma_{4,pred}space_{hi} \times guarding\ time_{hi} \\
& + \gamma_{5,pred}space_{hi} \times time\ 1^{st}capture_{hi} \\
& + \gamma_{6,pred}guarding\ time_{hi} \times time\ 1^{st}capture_{hi}
\end{aligned}
\tag{S7}
$$

The fourth and final model included prey behaviors with their quadratic effects as well as their interactions with predator behavior to investigate how they jointly affect hunting success. The model is described as :

$$
\begin{aligned}
logit(P_{hij}) = \ & (\beta_0 + env_{0h} + id_{0i} + \varepsilon_{0,hij}) \\
& + \beta_{1,pred} speed_{hi} + \beta_{2,pred} space_{hi} \\
& + \beta_{3,pred} guarding\ time_{hi} + \beta_{4,pred} time\ 1^{st} capture_{hi} \\
& + \beta_{5,prey} speed'_{hi} + \beta_{6,prey} space'_{hi} \\
& + \frac{1}{2}\gamma_{1,pred} speed^2_{hi} + \frac{1}{2}\gamma_{2,pred} space^2_{hi} \\
& + \frac{1}{2}\gamma_{3,pred} guarding\ time^2_{hi} + \frac{1}{2}\gamma_{4,pred} time\ 1^{st} capture^2_{hi} \\
& + \frac{1}{2}\gamma_{5,prey} speed'^2_{hi} + \frac{1}{2}\gamma_{6,prey} space'^2_{hi} \\
& + \gamma_{1,pred} speed_{hi} \times space_{hi} \\
& + \gamma_{2,pred} speed_{hi} \times guarding\ time_{hi} \\
& + \gamma_{3,pred} speed_{hi} \times time\ 1^{st} capture_{hi} \\
& + \gamma_{4,pred} space_{hi} \times guarding\ time_{hi} \\
& + \gamma_{5,pred} space_{hi} \times time\ 1^{st} capture_{hi} \\
& + \gamma_{6,pred} guarding\ time_{hi} \times time\ 1^{st} capture_{hi} \\
& + \gamma_{7,pred\ prey} speed_{hi} \times speed'_{hi} \\
& + \gamma_{8,pred\ prey} speed_{hi} \times space'_{hi} \\
& + \gamma_{9,pred\ prey} space_{hi} \times speed'_{hi} \\
& + \gamma_{10,pred\ prey} space_{hi} \times space'_{hi} \\
& + \gamma_{11,pred\ prey} guarding\ time_{hi} \times speed'_{hi} \\
& + \gamma_{12,pred\ prey} guarding\ time_{hi} \times space'_{hi} \\
& + \gamma_{13,pred\ prey} time\ 1^{st} capture_{hi} \times speed'_{hi} \\
& + \gamma_{14,pred\ prey} time\ 1^{st} capture_{hi} \times space'_{hi}
\end{aligned}
\tag{S8}
$$

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Principal component analysis

The first principal component (PC1) explained 20.9% of the total variation, and the second principal component (PC2) explained 19.6% of the total variation in the data (figure S1). We found that the time before the first capture, the rate of space covered, and the time spent guarding had the highest contributions (36.71, 27.04, and 23.38 respectively) to the first principal component (table S1), while the travel speed had the highest contribution (33.29) to the second principal component (table S1). Although the amount of times the predator damaged a generator explained almost an equivalent portion of the variation of the second principal component, we only kept travel speed because its contribution was higher and more ecologically meaningful.
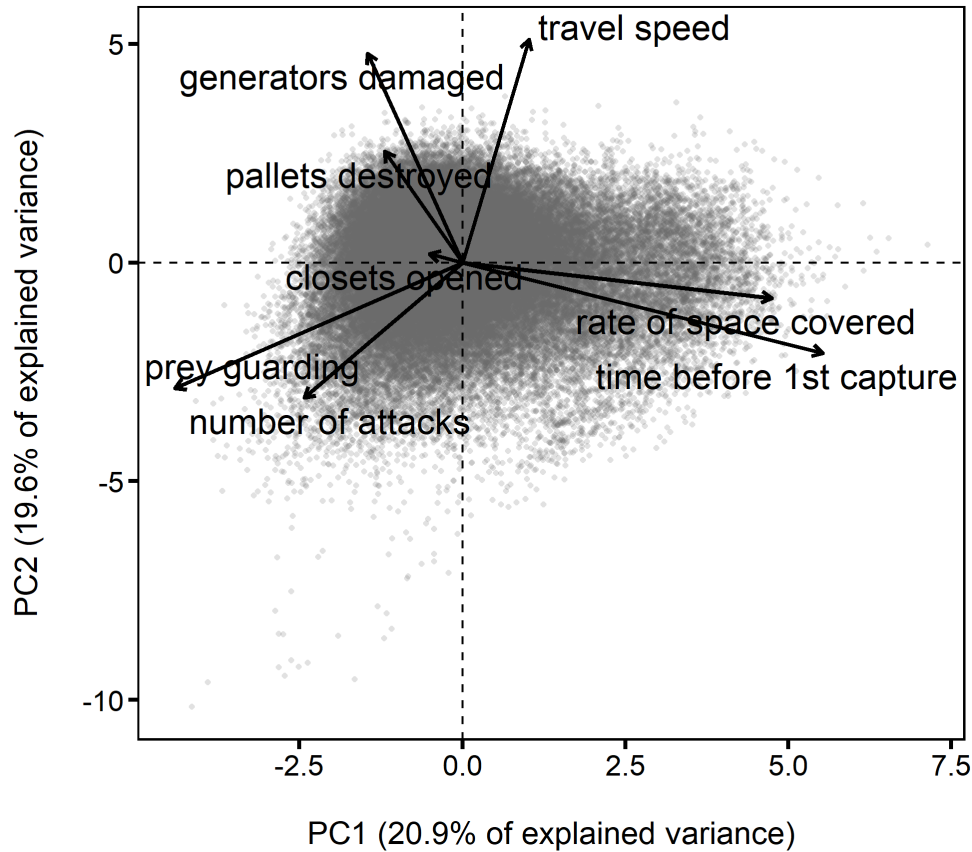

Figure S 1: Biplot of the principal component analysis on predator hunting behaviors. The variables presented in the main text were selected based on their relative contribution (%) to the first and second principal components.

Table S1: Principal component loadings for all the predator behavioral traits.

	PC1 (% variance = 21.0)		PC2 (% variance = 15.3)	
	Correlation	% variance	Correlation	% variance
ambush time	-0.66	23.19	-0.03	0.05
travel speed	0.75	29.38	0.15	1.73
rate of space covered	0.00	0.00	0.82	48.12
closets opened	-0.06	0.21	-0.09	0.56
pallets destroyed	0.21	2.42	-0.08	0.43
generators damaged	0.46	11.17	-0.37	9.79
normal attacks	-0.61	19.89	-0.13	1.31
special attacks	0.48	12.02	0.18	2.47
time before 1st capture	-0.18	1.72	0.70	35.55

Intra-class correlation coefficients of the predator behaviors for the three additional multivariate

mixed-models

After removing prey behavior as linear fixed effects (model 2), the relative contribution (%) of the game environments to the variation in travel speed, the time spent guarding, and the time required to capture the first prey was comparable to the first model (table S2). Thus, the expression of these behaviors was similar among the game environments. However, among environmental differences in the rate of space covered increased (table S2). Results show that novice and experimented players obtained similar ICCs for several of their random effects. However, experimented players were more similar in their average behavior than novice players.

Table S2: Posterior means of the ICC estimates for the three additional multivariate mixed-models.

Behavior	Random effect	ICC (95% CI) Without prey	ICC (95% CI) Novices	ICC (95% CI) Experienced
travel speed	player ID	0.28 (0.26, 0.30)	0.35 (0.32, 0.38)	0.19 (0.16, 0.22)
	game environment	0.00 (0.00, 0.00)	0.00 (0.00, 0.00)	0.00 (0.00, 0.00)
	avatar	0.09 (0.04, 0.14)	0.09 (0.04, 0.15)	0.10 (0.05, 0.17)
	residuals	0.63 (0.59, 0.67)	0.56 (0.51, 0.59)	0.71 (0.65, 0.76)
space covered	player ID	0.08 (0.06, 0.09)	0.06 (0.05, 0.07)	0.06 (0.05, 0.07)
	game environment	0.19 (0.12, 0.29)	0.04 (0.02, 0.06)	0.19 (0.12, 0.28)
	avatar	0.01 (0.01, 0.03)	0.02 (0.01, 0.03)	0.02 (0.01, 0.03)
	residuals	0.72 (0.64, 0.79)	0.88 (0.86, 0.91)	0.73 (0.65, 0.80)
guard time	player ID	0.31 (0.28, 0.33)	0.23 (0.21, 0.25)	0.15 (0.13, 0.18)
	game environment	0.00 (0.00, 0.00)	0.02 (0.01, 0.02)	0.00 (0.00, 0.00)
	avatar	0.03 (0.01, 0.05)	0.02 (0.01, 0.04)	0.04 (0.01, 0.07)
	residuals	0.67 (0.64, 0.69)	0.73 (0.71, 0.75)	0.81 (0.77, 0.84)
time 1st capture	player ID	0.26 (0.24, 0.28)	0.21 (0.20, 0.23)	0.15 (0.13, 0.17)
	game environment	0.01 (0.00, 0.01)	0.03 (0.02, 0.04)	0.01 (0.00, 0.01)
	avatar	0.03 (0.01, 0.05)	0.02 (0.01, 0.04)	0.03 (0.01, 0.06)
	residuals	0.70 (0.68, 0.73)	0.74 (0.71, 0.76)	0.81 (0.78, 0.84)

Table S3: Among-individual and within-individual behavioral correlations for the three additional multivariate mixed-models.

	speed	space	prey guarding	time 1st cap.
Model: without prey				
speed	-	0.21 (0.21, 0.22)	-0.05 (-0.05,-0.06)	-0.07 (-0.06,-0.07)
space	-0.03 (-0.09, 0.04)	-	-0.25 (-0.24,-0.26)	0.09 (0.08, 0.10)
prey guarding	0.06 (0.00, 0.13)	-0.48 (-0.42,-0.54)	-	-0.43 (-0.43,-0.44)
time 1st cap.	-0.41 (-0.35,-0.46)	0.59 (0.53, 0.64)	-0.73 (-0.70,-0.76)	-
Model: novice players				
speed	-	0.22 (0.21, 0.23)	-0.13 (-0.12,-0.14)	-0.02 (-0.01,-0.03)
space	0.50 (0.44, 0.57)	-	-0.19 (-0.18,-0.21)	0.07 (0.05, 0.08)
prey guarding	0.08 (0.02, 0.14)	0.05 (-0.04, 0.14)	-	-0.45 (-0.44,-0.46)
time 1st cap.	-0.49 (-0.45,-0.55)	-0.21 (-0.13,-0.29)	-0.62 (-0.58,-0.66)	-
Model: experienced players				
speed	-	0.22 (0.21, 0.23)	0.00 (-0.01, 0.01)	-0.11 (-0.10,-0.12)
space	-0.03 (-0.15, 0.09)	-	-0.23 (-0.21,-0.24)	0.04 (0.02, 0.06)
prey guarding	-0.21 (-0.10,-0.32)	-0.04 (-0.16, 0.09)	-	-0.39 (-0.38,-0.40)
time 1st cap.	-0.29 (-0.19,-0.39)	0.38 (0.27, 0.48)	-0.41 (-0.32,-0.50)	-

* The among-individual correlations are on the lower off-diagonal and the residual within-individual correlations on the upper off-diagonal

Selection of the best hunting success model

The model with the lowest expected log predictive density was chosen as the best model for our interpretations of the relationship between behavior and hunting success.

Table S4: Summary table comparing the hunting success models based on approximate leave-one-out cross-validation.

model	elpd difference	sd difference	elpd loo value	elpd loo standard error
quadratic model pred-prey	0.00	0.00	-96 276.18	239.14
linear model pred-prey	-2 356.80	94.08	-98 632.99	239.65
quadratic model pred	-7 243.63	138.56	-103 519.82	250.09
linear model pred	-8 905.85	156.73	-105 182.03	248.13

Figure S2 shows the interacting effects of all the combinations of predator behaviors on predator hunting success drawn from the parameters presented on table 1 in the main text.
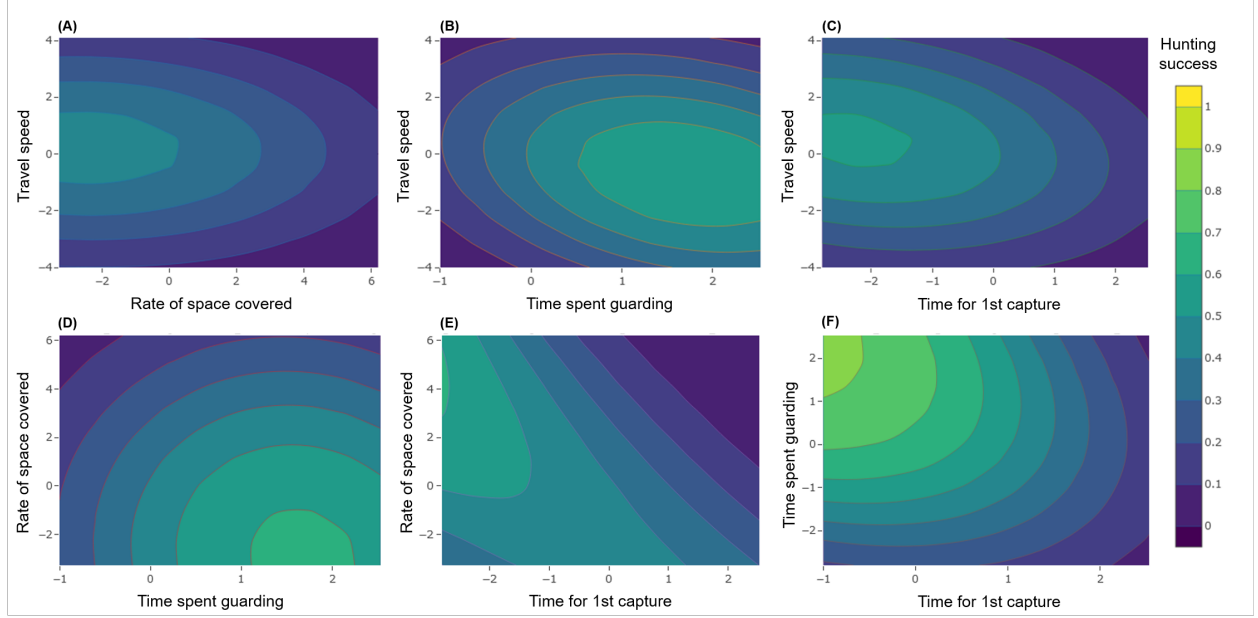

Figure S 2: Interacting effects of the predator behaviors on hunting success. Hunting success is represented by the color gradient. We computed the plots by predicting the mean probability of capturing four prey based on the best quadratic approximation of the predator behavior interaction terms. (A) Travel speed and the rate of space covered. (B) Travel speed and the time spent guarding prey. (C) Time before the first capture and travel speed. (D) The rate of space covered and the time spent guarding. (E) The rate of space covered and the time before the first capture. (F) The time spent guarding prey and the time before the first capture.

Literature Cited

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