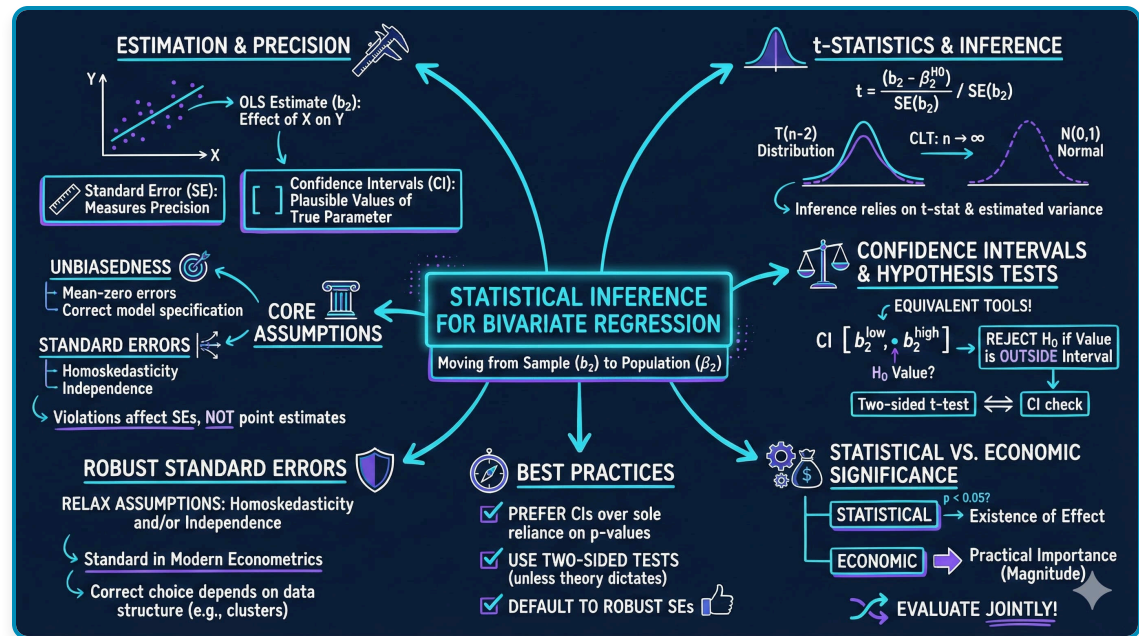


Chapter 7: Statistical Inference for Bivariate Regression

metricsAI: An Introduction to Econometrics with Python and AI in the Cloud

Carlos Mendez



This notebook provides an interactive introduction to statistical inference for bivariate regression models. All code runs directly in Google Colab without any local setup.

 [Open in Colab](#)

Chapter Overview

This chapter extends statistical inference from univariate to bivariate regression. You'll gain both theoretical understanding and practical skills through hands-on Python examples, learning how to test hypotheses about regression coefficients and construct confidence intervals.

What you'll learn:

- The t-statistic for testing hypotheses about regression coefficients
- Constructing and interpreting confidence intervals for slope parameters

- Tests of statistical significance (whether a regressor matters)
- Two-sided hypothesis tests for specific parameter values
- One-sided directional hypothesis tests
- Heteroskedasticity-robust standard errors and their importance
- Economic vs. statistical significance

Datasets used:

- **AED_HOUSE.DTA:** House prices and characteristics for 29 houses sold in Central Davis, California in 1999 (price, size, bedrooms, bathrooms, lot size, age)

Chapter outline:

- 7.1 Example: House Price and Size
- 7.2 The t Statistic
- 7.3 Confidence Intervals
- 7.4 Tests of Statistical Significance
- 7.5 Two-Sided Hypothesis Tests
- 7.6 One-Sided Directional Hypothesis Tests
- 7.7 Robust Standard Errors
- 7.8 Case Studies
- Key Takeaways
- Practice Exercises

| Setup

First, we import the necessary Python packages and configure the environment for reproducibility. All data will stream directly from GitHub.

In [18]:

```
# Import required packages
import numpy as np
import pandas as pd
import matplotlib.pyplot as plt
import seaborn as sns
import statsmodels.api as sm
from statsmodels.formula.api import ols
from scipy import stats
from statsmodels.stats.sandwich_covariance import cov_hc1
import random
import os

# Set random seeds for reproducibility
RANDOM_SEED = 42
random.seed(RANDOM_SEED)
np.random.seed(RANDOM_SEED)
os.environ['PYTHONHASHSEED'] = str(RANDOM_SEED)

# GitHub data URL
GITHUB_DATA_URL = "https://raw.githubusercontent.com/quarcs-lab/data-open/master/AED/"

# Set plotting style
sns.set_style("whitegrid")
plt.rcParams['figure.figsize'] = (10, 6)

print("Setup complete! Ready to explore statistical inference for bivariate regression.")
```

Setup complete! Ready to explore statistical inference for bivariate regression.

7.1 Example: House Price and Size

We begin with a motivating example: the relationship between house price and house size.

The regression model:

$$\text{price} = \beta_1 + \beta_2 \times \text{size} + u$$

where:

- price is the house sale price (in thousands of dollars)
- size is the house size (in square feet)
- β_2 is the population slope (price increase per square foot)
- b_2 is the sample estimate of β_2

Key regression output:

Variable	Coefficient	Standard Error	t-statistic	p-value	95% CI
Size	73.77	11.17	6.60	0.000	[50.84, 96.70]
Intercept	115,017.30	21,489.36	5.35	0.000	[70,924.76, 159,109.84]

Interpretation:

- Each additional square foot increases house price by approximately \$73.77
- The standard error (11.17) measures uncertainty in this estimate
- The t-statistic (6.60) tests whether the effect is statistically significant
- The 95% confidence interval is [50.84, 96.70]

In [19]:

```
print("=" * 70)
print("7.1 EXAMPLE: HOUSE PRICE AND SIZE")
print("=" * 70)

# Read in the house data
data_house = pd.read_stata(GITHUB_DATA_URL + 'AED_HOUSE.DTA')

print("\nData summary:")
data_summary = data_house.describe()
print(data_summary)

print("\nFirst few observations:")
print(data_house.head())
```

```
=====
7.1 EXAMPLE: HOUSE PRICE AND SIZE
=====

Data summary:

```

	price	size	bedrooms	bathrooms	lotsize	age \
count	29.000000	29.000000	29.000000	29.000000	29.000000	29.000000
mean	253910.344828	1882.758621	3.793103	2.206897	2.137931	36.413792
std	37390.710695	398.272130	0.675030	0.341144	0.693034	7.118975
min	204000.000000	1400.000000	3.000000	2.000000	1.000000	23.000000
25%	233000.000000	1600.000000	3.000000	2.000000	2.000000	31.000000
50%	244000.000000	1800.000000	4.000000	2.000000	2.000000	35.000000
75%	270000.000000	2000.000000	4.000000	2.500000	3.000000	39.000000
max	375000.000000	3300.000000	6.000000	3.000000	3.000000	51.000000

```


```

	monthsold	list
count	29.000000	29.000000
mean	5.965517	257824.137931
std	1.679344	40860.264099
min	3.000000	199900.000000
25%	5.000000	239000.000000
50%	6.000000	245000.000000
75%	7.000000	269000.000000
max	8.000000	386000.000000

```

First few observations:

```

	price	size	bedrooms	bathrooms	lotsize	age	monthsold	list
0	204000	1400	3	2.0	1	31.0	7	199900
1	212000	1600	3	3.0	2	33.0	5	212000
2	213000	1800	3	2.0	2	51.0	4	219900
3	220000	1600	3	2.0	1	49.0	4	229000
4	224500	2100	4	2.5	2	47.0	6	224500

| Basic Regression: Price on Size

We estimate the bivariate regression model using ordinary least squares (OLS).

In [20]:

```
# Table 7.1 - Basic regression
print("=" * 70)
print("Table 7.1: Regression of House Price on Size")
print("=" * 70)

model_basic = ols('price ~ size', data=data_house).fit()
print(model_basic.summary())
```

```
=====
Table 7.1: Regression of House Price on Size
=====
                        OLS Regression Results
=====
Dep. Variable:          price    R-squared:                0.617
Model:                  OLS      Adj. R-squared:           0.603
Method:                 Least Squares    F-statistic:        43.58
Date:                   Tue, 20 Jan 2026    Prob (F-statistic):    4.41e-07
Time:                   23:33:22    Log-Likelihood:       -332.05
No. Observations:       29      AIC:                   668.1
Df Residuals:           27      BIC:                   670.8
Df Model:                1
Covariance Type:        nonrobust
=====
               coef      std err          t      P>|t|      [0.025      0.975]
-----
Intercept    1.15e+05    2.15e+04     5.352     0.000     7.09e+04    1.59e+05
size          73.7710     11.175     6.601     0.000     50.842     96.700
=====
Omnibus:                 0.576    Durbin-Watson:           1.219
Prob(Omnibus):            0.750    Jarque-Bera (JB):         0.638
Skew:                    -0.078    Prob(JB):                 0.727
Kurtosis:                 2.290    Cond. No.                  9.45e+03
=====

Notes:
[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.
[2] The condition number is large, 9.45e+03. This might indicate that there are
strong multicollinearity or other numerical problems.
```

Coefficient Table

Let's create a clean table showing the key statistics for statistical inference.

In [21]:

```
# Save coefficients in a clean table
coef_table = pd.DataFrame({
    'Coefficient': model_basic.params,
    'Std. Error': model_basic.bse,
    't-statistic': model_basic.tvalues,
    'p-value': model_basic.pvalues
})

print("\nCoefficient Table:")
print(coef_table)

print("\nInterpretation:")
print(" - Slope (size): Each additional sq ft increases price by $73.77")
print(" - Standard error: Measures uncertainty in the slope estimate")
print(" - t-statistic: Tests whether slope differs from zero")
print(" - p-value: Probability of observing such extreme values under H0:  $\beta_2 = 0$ ")
```

Coefficient Table:

	Coefficient	Std. Error	t-statistic	p-value
Intercept	115017.282609	21489.359861	5.352290	1.183545e-05
size	73.771040	11.174911	6.601488	4.408752e-07

Interpretation:

- Slope (size): Each additional sq ft increases price by \$73.77
- Standard error: Measures uncertainty in the slope estimate
- t-statistic: Tests whether slope differs from zero
- p-value: Probability of observing such extreme values under $H_0: \beta_2 = 0$

7.2 The t Statistic

The t-statistic is fundamental to statistical inference in regression.

Statistical inference problem:

- **Sample:** $\hat{y} = b_1 + b_2x$ where b_1 and b_2 are least squares estimates
- **Population:** $E[y|x] = \beta_1 + \beta_2x$ and $y = \beta_1 + \beta_2x + u$
- **Goal:** Make inferences about the slope parameter β_2

The t-statistic:

$$T = \frac{\text{estimate} - \text{parameter}}{\text{standard error}} = \frac{b_2 - \beta_2}{se(b_2)} \sim T(n - 2)$$

Why use the T(n-2) distribution?

Under assumptions 1-4:

- $Var[b_2] = \sigma_u^2 / \sum_{i=1}^n (x_i - \bar{x})^2$
- We don't know σ_u^2 , so we replace it with $s_e^2 = \frac{1}{n-2} \sum_{i=1}^n (y_i - \hat{y}_i)^2$
- This introduces additional uncertainty, so we use $T(n - 2)$ instead of $N(0, 1)$

Model Assumptions (1-4):

1. The population model is $y = \beta_1 + \beta_2x + u$
2. The error has mean zero conditional on x: $E[u_i|x_i] = 0$
3. The error has constant variance: $Var[u_i|x_i] = \sigma_u^2$
4. The errors are statistically independent: u_i independent of u_j

Understanding Standard Errors: The Foundation of Inference

What is a standard error?

The standard error measures the uncertainty in our estimate. It answers: "If we repeatedly sampled from the population and computed b_2 each time, how much would b_2 vary across samples?"

Key distinction:

- **Standard deviation:** Variability of individual observations (data spread)
- **Standard error:** Variability of the estimate across samples (estimation uncertainty)

Formula for SE of the slope:

$$se(b_2) = \frac{\sigma_u}{\sqrt{\sum_{i=1}^n (x_i - \bar{x})^2}} = \frac{\sigma_u}{\sqrt{n} \cdot \sigma_x}$$

where:

- σ_u = standard deviation of the error term
- σ_x = standard deviation of x
- n = sample size

What affects the standard error?

1. Error variance (σ^2_u): Larger $\sigma_u \rightarrow$ Larger SE

- More unexplained variation in y
- Data points farther from regression line
- Less precise estimates

2. Sample size (n): Larger n \rightarrow Smaller SE

- More data reduces uncertainty
- SE decreases at rate $1/\sqrt{n}$
- Quadruple n to halve SE

3. Variation in x (σ_x): Larger $\sigma_x \rightarrow$ Smaller SE

- More spread in x provides more information
- Extreme x values help identify the slope
- Concentrated x values give less precision

Intuition:

Think of the regression line as a seesaw balanced at (\bar{x}, \bar{y}) :

- With wide spread in x: Small changes in slope make big differences at the extremes (easy to detect slope)
- With narrow spread in x: Hard to distinguish different slopes (difficult to detect slope)

Example calculation for our house price data:

Given:

- Sample size: $n = 29$
- Standard error of regression: $\sigma_u \approx 37,000$
- Standard deviation of size: $\sigma_x \approx 360$
- Estimated $SE(b_2) = 37,000 / (\sqrt{29} \times 360) \approx 19.1$

(Actual SE is 11.17, smaller because the relationship is quite strong)

Why standard errors matter:

1. **Confidence intervals:** $CI = b_2 \pm t \times SE(b_2)$
2. **Hypothesis tests:** $t = (b_2 - \beta_2^0) / SE(b_2)$
3. **Practical significance:** Small SE \rightarrow precise estimate \rightarrow more reliable
4. **Study design:** Calculate required n for desired SE

Relationship to R^2 :

Higher R^2 (better fit) \rightarrow Smaller σ_u \rightarrow Smaller SE \rightarrow More precise estimates

For our house price example:

- $R^2 = 0.62$ (size explains 62% of price variation)
- This gives relatively small SE
- If R^2 were 0.10, SE would be about 2.5 times larger

In [22]:

```
print("=" * 70)
print("7.2 THE T STATISTIC")
print("=" * 70)

print("\nRegression coefficients and t-statistics:")
print(model_basic.summary2().tables[1])

# Extract key statistics
coef_size = model_basic.params['size']
se_size = model_basic.bse['size']
t_stat_size = model_basic.tvalues['size']
p_value_size = model_basic.pvalues['size']

print(f"\nDetailed statistics for 'size' coefficient:")
print(f"  Coefficient: ${coef_size:.4f}")
print(f"  Standard Error: ${se_size:.4f}")
print(f"  t-statistic: {t_stat_size:.4f}")
print(f"  p-value: {p_value_size:.6f}")

print("\nThe t-statistic formula:")
print(f"  t = b2 / se(b2) = {coef_size:.4f} / {se_size:.4f} = {t_stat_size:.4f}")
```

```
=====
7.2 THE T STATISTIC
=====

Regression coefficients and t-statistics:

```

	Coef.	Std.Err.	t	P> t	[0.025	\
Intercept	115017.282609	21489.359861	5.352290	1.183545e-05	70924.758265	
size	73.771040	11.174911	6.601488	4.408752e-07	50.842017	

```

                                0.975]
Intercept  159109.806952
size       96.700064

Detailed statistics for 'size' coefficient:
  Coefficient: $73.7710
  Standard Error: $11.1749
  t-statistic: 6.6015
  p-value: 0.000000

The t-statistic formula:
  t = b2 / se(b2) = 73.7710 / 11.1749 = 6.6015
```

Key Concept: The t-Distribution and Degrees of Freedom

The **t-distribution** is used for statistical inference when the population variance is unknown (which is always the case in practice). Unlike the standard normal distribution, the t-distribution accounts for the additional uncertainty from estimating the variance.

Key properties:

- Bell-shaped and symmetric (like the normal distribution)
- Heavier tails than the normal distribution (more probability in extremes)
- Converges to the normal distribution as sample size increases
- Characterized by degrees of freedom (df)

Degrees of freedom = $n - 2$ for bivariate regression:

- Start with n observations
- Estimate β_1 (intercept): -1 df
- Estimate β_2 (slope): -1 df
- Remaining df for estimating variance: $n - 2$

Practical implication: For small samples ($n < 30$), the t-distribution's heavier tails lead to wider confidence intervals and more conservative hypothesis tests compared to the normal distribution. For large samples ($n > 100$), the difference becomes negligible.

| 7.3 Confidence Intervals

A confidence interval provides a range of plausible values for the population parameter.

Formula for a $100(1 - \alpha)$ confidence interval:

$$b_2 \pm t_{n-2, \alpha/2} \times se(b_2)$$

where:

- b_2 is the slope estimate

- $se(b_2)$ is the standard error of b_2
- $t_{n-2, \alpha/2}$ is the critical value from Student's t-distribution with $n - 2$ degrees of freedom

95% confidence interval (approximate):

$$b_2 \pm 2 \times se(b_2)$$

Interpretation:

- If we repeatedly sampled from the population and constructed 95% CIs, approximately 95% of these intervals would contain the true parameter value β_2
- The calculated 95% CI will correctly include β_2 95% of the time

Example calculation for house price:

$$\begin{aligned} b_2 \pm t_{27, 0.025} \times se(b_2) &= 73.77 \pm 2.052 \times 11.17 \\ &= 73.77 \pm 22.93 \\ &= [50.84, 96.70] \end{aligned}$$

Key Concept: Interpreting Confidence Intervals

A **confidence interval** provides a range of plausible values for the population parameter. For regression slopes, the 95% CI is:

$$b_2 \pm t_{n-2, 0.025} \times se(b_2)$$

Common misconceptions:

- **WRONG:** "There's a 95% probability that β_2 is in this interval"
- **CORRECT:** "If we repeated the sampling process many times, 95% of the constructed intervals would contain β_2 "

Practical interpretation:

- The interval represents our uncertainty about the true parameter value
- Wider intervals indicate more uncertainty (large SE, small n , or high variability)
- Narrower intervals indicate more precision (small SE, large n , or low variability)
- Values inside the interval are "plausible" at the chosen confidence level
- Values outside the interval would be rejected in a hypothesis test

Relationship to hypothesis testing: If a null value β_2^* falls inside the 95% CI, we fail to reject $H_0: \beta_2 = \beta_2^*$ at the 5% significance level. This makes CIs more informative than hypothesis tests alone.

| Understanding Confidence Intervals: A Deep Dive

What is a confidence interval?

A confidence interval (CI) is NOT a probability statement about the parameter. Instead, it's a statement about the procedure used to construct the interval.

Common misconceptions:

WRONG: "There is a 95% probability that β_2 is between 50.84 and 96.70"

- The parameter β_2 is fixed (not random)
- The interval either contains β_2 or it doesn't

CORRECT: "If we repeatedly sampled and constructed 95% CIs, approximately 95% of these intervals would contain the true β_2 "

- The randomness is in the sampling process
- Our particular interval is one realization from this process

Intuitive explanation:

Imagine conducting 100 different studies using different random samples from the same population:

- Each study estimates β_2 and constructs a 95% CI
- About 95 of the 100 intervals will contain the true β_2
- About 5 of the 100 intervals will miss β_2 (just by chance)

Width of confidence intervals:

The CI width depends on three factors:

$$\text{Width} = 2 \times t_{n-2, \alpha/2} \times se(b_2)$$

1. Confidence level (1- α): Higher confidence \rightarrow wider interval

- 90% CI: Narrower (less confident)
- 95% CI: Standard choice (balance)
- 99% CI: Wider (more confident)

2. Sample size (n): Larger sample \rightarrow narrower interval

- More data reduces uncertainty
- Critical value $t_{\{n-2, \alpha/2\}}$ decreases as n increases

3. Variability in data: More scatter \rightarrow wider interval

- $se(b_2)$ increases with unexplained variation
- Tighter relationship \rightarrow more precise estimates

Practical use:

Confidence intervals are more informative than hypothesis tests because they show:

- The point estimate (center of interval)
- The precision of the estimate (width of interval)
- All null values that would not be rejected (values inside interval)

In [23]:

```
print("=" * 70)
print("7.3 CONFIDENCE INTERVALS")
print("=" * 70)

# 95% confidence intervals
conf_int = model_basic.conf_int(alpha=0.05)
print("\n95% Confidence Intervals:")
print(conf_int)
```

```
=====
7.3 CONFIDENCE INTERVALS
=====

95% Confidence Intervals:
              0              1
Intercept  70924.758265  159109.806952
size        50.842017    96.700064
```

The t-Distribution vs Normal Distribution: Why It Matters

Why not use the normal distribution?

In theory, when we know the population variance σ^2_u , the test statistic follows a standard normal distribution:

$$Z = \frac{b_2 - \beta_2}{\sigma / \sqrt{\sum (x_i - \bar{x})^2}} \sim N(0, 1)$$

The problem: We never know σ_u in practice!

The solution: Replace σ_u with its estimate s_e (residual standard error):

$$T = \frac{b_2 - \beta_2}{s_e / \sqrt{\sum (x_i - \bar{x})^2}} \sim T(n - 2)$$

This substitution introduces additional uncertainty, so we use the t-distribution instead of normal.

Properties of the t-distribution:

1. **Shape:** Bell-shaped and symmetric (like normal)
2. **Mean:** 0 (like normal)
3. **Variance:** $df/(df-2) > 1$ (heavier tails than normal)

4. Degrees of freedom: $n - 2$ for bivariate regression

- n observations
- Minus 2 parameters estimated (β_1 and β_2)

Key differences from normal:

Sample Size	t Critical Value ($\alpha=0.05$)	z Critical Value	Difference
$n = 5$ (df=3)	3.182	1.96	+62%
$n = 10$ (df=8)	2.306	1.96	+18%
$n = 30$ (df=28)	2.048	1.96	+4%
$n = 100$ (df=98)	1.984	1.96	+1%
$n \rightarrow \infty$	1.96	1.96	0%

What this means:

- **Small samples:** t critical values much larger \rightarrow wider CIs, harder to reject H_0
- **Large samples:** $t \approx$ normal \rightarrow approximately same inference
- **Our house data:** $n=29$, $df=27$, $t(0.025) = 2.052$ vs $z = 1.96$

Why degrees of freedom = $n - 2$?

- Start with n observations
- Estimate β_1 (intercept): loses 1 df
- Estimate β_2 (slope): loses 1 df
- Remaining df for estimating variance: $n - 2$

Practical implications:

For $n = 29$ (our house price data):

- Using normal: 95% CI margin = $1.96 \times 11.17 = 21.89$
- Using $t(27)$: 95% CI margin = $2.052 \times 11.17 = 22.92$
- Difference: 5% wider with t-distribution (more conservative)

For $n = 10$ (small sample):

- Using normal: 95% CI margin = $1.96 \times SE$
- Using $t(8)$: 95% CI margin = $2.306 \times SE$
- Difference: 18% wider with t-distribution (much more conservative!)

Rule of thumb:

- $n < 30$: Must use t-distribution
- $30 \leq n < 100$: Use t-distribution (small difference)
- $n \geq 100$: Normal approximation usually fine, but still use t

Modern practice: Statistical software always uses t-distribution (why not? It's correct for any n)

Visual intuition:

The t-distribution has heavier tails:

- More probability in the extremes
- Less probability near the center
- This accounts for the uncertainty in estimating σ_u
- As n increases, estimation uncertainty decreases, and $t \rightarrow \text{normal}$

Manual Calculation of Confidence Interval

Let's manually calculate the confidence interval for the size coefficient to understand the mechanics.

In [24]:

```
# Manual calculation of confidence interval for size
n = len(data_house)
df = n - 2
t_crit = stats.t.ppf(0.975, df) # 97.5th percentile for two-sided 95% CI

ci_lower = coef_size - t_crit * se_size
ci_upper = coef_size + t_crit * se_size

print("Manual calculation for 'size' coefficient:")
print(f" Sample size: {n}")
print(f" Degrees of freedom: {df}")
print(f" Critical t-value (α=0.05): {t_crit:.4f}")
print(f" Margin of error: {t_crit * se_size:.4f}")
print(f" 95% CI: [{ci_lower:.4f}, {ci_upper:.4f}]")

print("\nInterpretation:")
print(f" We are 95% confident that each additional square foot")
print(f" increases house price by between {ci_lower:.2f} and {ci_upper:.2f}.")
```

Manual calculation for 'size' coefficient:

```
Sample size: 29
Degrees of freedom: 27
Critical t-value (α=0.05): 2.0518
Margin of error: 22.9290
95% CI: [$50.8420, $96.7001]
```

Interpretation:

```
We are 95% confident that each additional square foot
increases house price by between $50.84 and $96.70.
```


| Example with Artificial Data

To illustrate the concepts more clearly, let's work with a simple artificial dataset.

Key Concept: The Hypothesis Testing Framework

Hypothesis testing is a formal procedure for making decisions about population parameters. The key steps are:

1. State the hypotheses:

- **Null hypothesis (H_0):** The claim we're testing (usually "no effect")
- **Alternative hypothesis (H_a):** What we conclude if we reject H_0

2. Choose significance level (α):

- Common choices: 0.10, 0.05, 0.01
- α = probability of Type I error (rejecting H_0 when it's true)

3. Calculate test statistic:

- Standardizes the difference: $t = (\text{estimate} - \text{null value}) / SE$

4. Determine p-value:

- Probability of observing our result (or more extreme) if H_0 is true
- Smaller p-value = stronger evidence against H_0

5. Make decision:

- Reject H_0 if $p\text{-value} < \alpha$
- Fail to reject H_0 if $p\text{-value} \geq \alpha$ (never "accept" H_0)

Understanding p-values: If $p = 0.001$, this means "if H_0 were true, we'd observe a result this extreme only 0.1% of the time." This is strong evidence against H_0 .

Key Concept: Statistical vs. Economic Significance

Statistical significance and **economic significance** are distinct concepts that answer different questions:

Statistical Significance:

- **Question:** *Is the effect different from zero?*
- **Determined by:** *$t\text{-statistic} = b_2 / \text{se}(b_2)$, which depends on sample size, variability, and effect size*
- **Interpretation:** *We can confidently say the effect exists (not due to chance)*

Economic Significance:

- **Question:** *Is the effect large enough to matter in practice?*
- **Determined by:** *The magnitude of b_2 and the context*
- **Interpretation:** *The effect has real-world importance*

Why they can diverge:

1. Large n: Even tiny effects become statistically significant

- *Example: $\beta_2 = \$0.01$ with $n = 10,000$ might have $p < 0.001$ but be economically trivial*

2. Small n: Large effects may not reach statistical significance

- *Example: $\beta_2 = \$100$ with $n = 10$ might have $p = 0.12$ but be economically important*

Best practice: Always report both the coefficient estimate (economic magnitude) and the standard error/confidence interval (statistical precision). Focus on confidence intervals, which show both dimensions simultaneously.

Key Concept: One-Sided vs. Two-Sided Tests

The choice between one-sided and two-sided tests depends on your research question:

Two-Sided Test (Most Common):

- $H_0: \beta_2 = \beta_2^*$ vs. $H_a: \beta_2 \neq \beta_2^*$
- Detects deviations in either direction
- Standard practice in academic research
- Rejection region: Both tails of t-distribution

One-Sided Test (Directional):

- Upper: $H_0: \beta_2 \leq \beta_2^*$ vs. $H_a: \beta_2 > \beta_2^*$
- Lower: $H_0: \beta_2 \geq \beta_2^*$ vs. $H_a: \beta_2 < \beta_2^*$
- Detects deviations in one specific direction
- Rejection region: One tail only

Key relationship: *For the same test statistic, one-sided p-value = (two-sided p-value) / 2 (if sign is correct)*

When to use one-sided tests:

- Strong theoretical prediction of direction (before seeing data)
- Only care about deviations in one direction
- Be cautious: Journals typically require two-sided tests

Important: *If your data contradicts the predicted direction, you cannot reject H_0 with a one-sided test ($p\text{-value} > 0.5$).*

In [25]:

```
print("=" * 70)
print("Example with Artificial Data")
print("=" * 70)

x = np.array([1, 2, 3, 4, 5])
y = np.array([1, 2, 2, 2, 3])
df_artificial = pd.DataFrame({'x': x, 'y': y})

model_artificial = ols('y ~ x', data=df_artificial).fit()
print(model_artificial.summary())

coef_x = model_artificial.params['x']
se_x = model_artificial.bse['x']
n_art = len(x)
df_art = n_art - 2
t_crit_art = stats.t.ppf(0.975, df_art)

ci_lower_art = coef_x - t_crit_art * se_x
ci_upper_art = coef_x + t_crit_art * se_x

print(f"\nManual CI for artificial data:")
print(f"  Coefficient: {coef_x:.4f}")
print(f"  Standard Error: {se_x:.4f}")
print(f"  95% CI: [{ci_lower_art:.4f}, {ci_upper_art:.4f}]" )
```

```
=====
Example with Artificial Data
=====

                        OLS Regression Results
=====
Dep. Variable:          y      R-squared:                0.800
Model:                  OLS      Adj. R-squared:          0.733
Method:                 Least Squares      F-statistic:      12.00
Date:                  Tue, 20 Jan 2026      Prob (F-statistic):  0.0405
Time:                  23:33:22      Log-Likelihood:      -0.78037
No. Observations:      5      AIC:                    5.561
Df Residuals:          3      BIC:                    4.780
Df Model:              1
Covariance Type:       nonrobust
=====
                        coef      std err          t      P>|t|      [0.025      0.975]
-----
Intercept              0.8000      0.383        2.089      0.128      -0.419      2.019
x                      0.4000      0.115        3.464      0.041       0.033      0.767
=====
Omnibus:                nan      Durbin-Watson:        2.600
Prob(Omnibus):           nan      Jarque-Bera (JB):      0.352
Skew:                   -0.000      Prob(JB):              0.839
Kurtosis:               1.700      Cond. No.              8.37
=====
```

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

Manual CI for artificial data:

Coefficient: 0.4000

Standard Error: 0.1155

95% CI: [0.0325, 0.7675]

/Users/carlosmendez/miniforge3/lib/python3.10/site-packages/statsmodels/stats/stattools.py:74: ValueWarning: omni_normtest is not valid with less than 8 observations; 5 samples were given.

warn("omni_normtest is not valid with less than 8 observations; %i "

Understanding Hypothesis Testing: The Complete Workflow

The hypothesis testing framework:

Hypothesis testing is a formal procedure for making decisions about population parameters using sample data.

Step-by-step workflow:

1. State the hypotheses

- **Null hypothesis (H_0):** The claim we're testing (usually "no effect")
- **Alternative hypothesis (H_a):** What we conclude if we reject H_0
- Example: $H_0: \beta_2 = 0$ vs $H_a: \beta_2 \neq 0$

2. Choose the significance level (α)

- Common choices: 0.10, 0.05, 0.01
- α = probability of rejecting H_0 when it's actually true (Type I error)
- Convention: $\alpha = 0.05$ (5% significance level)

3. Calculate the test statistic

- Formula: $t = (b_2 - \beta_2^0) / \text{se}(b_2)$
- This standardizes the difference between estimate and null value
- Under H_0 , t follows a t-distribution with $n-2$ degrees of freedom

4. Determine the p-value

- p-value = probability of observing a test statistic as extreme as ours if H_0 is true
- Smaller p-value = stronger evidence against H_0
- $\text{p-value} < \alpha \rightarrow \text{reject } H_0$

5. Make a decision

- **Reject H_0 :** Strong evidence against the null hypothesis
- **Fail to reject H_0 :** Insufficient evidence to reject the null
- Note: We never "accept" H_0 , we only fail to reject it

6. State the conclusion

- Translate the statistical decision into plain language

- Example: "House size has a statistically significant effect on price at the 5% level"

Understanding p-values:

The p-value answers this question: "If the null hypothesis were true, what is the probability of getting a result at least as extreme as what we observed?"

Example interpretation:

- $p = 0.000$: If β_2 were truly zero, the probability of getting $t \geq 6.60$ is less than 0.1%
- This is very unlikely, so we have strong evidence against H_0

Two approaches to hypothesis testing:

1. **p-value approach**: Reject H_0 if $p\text{-value} < \alpha$
2. **Critical value approach**: Reject H_0 if $|t| > \text{critical value}$

Both approaches always give the same conclusion!

Common significance levels and interpretations:

p-value	Interpretation	Strength of Evidence
$p > 0.10$	Not significant	Weak/no evidence
$0.05 < p \leq 0.10$	Marginally significant	Moderate evidence
$0.01 < p \leq 0.05$	Significant	Strong evidence
$p \leq 0.01$	Highly significant	Very strong evidence

Statistical Significance vs Economic Significance

A crucial distinction that is often confused:

Statistical Significance

- Answers: "Is the effect different from zero?"
- Depends on: Sample size, variability, effect size
- Formula: $t = b_2 / \text{se}(b_2)$
- Interpretation: We can confidently say the effect exists

Economic Significance

- Answers: "Is the effect large enough to matter?"
- Depends on: The magnitude of b_2 and context
- Requires: Domain knowledge and practical judgment

- Interpretation: The effect has real-world importance

Key insights:

1. Statistical significance \neq Economic significance

You can have:

- Statistically significant but economically trivial effects
 - Example: With $n=10,000$, $\beta_2 = 0.001$ might be statistically significant but meaningless
- Economically important but statistically insignificant effects
 - Example: With $n=10$, $\beta_2 = 100$ might not be statistically significant but potentially important

2. Sample size matters for statistical significance

$$t = \frac{b_2}{se(b_2)} = \frac{b_2}{\sigma_u / \sqrt{\sum (x_i - \bar{x})^2}}$$

As sample size increases:

- The denominator (standard error) decreases
- The t-statistic increases
- Small effects become statistically significant

3. Economic significance requires context

For the house price example:

- Statistical result: $\beta_2 = 73.77$, highly significant ($p < 0.001$)
- Economic interpretation: Each sq ft adds \$73.77 to price
- Context matters:
 - For a 100 sq ft difference: \$7,377 difference (substantial!)
 - For a 10 sq ft difference: \$737.70 difference (moderate)
 - This is economically meaningful in the housing market

Practical guidance:

1. Always report both:

- The coefficient estimate (economic magnitude)
- The p-value or confidence interval (statistical precision)

2. Focus on confidence intervals:

- Shows both statistical and economic significance
- Example: 95% CI = [50.84, 96.70]
- Interpretation: Effect is between 50.84 and 96.70 per sq ft
- Even the lower bound is economically meaningful

3. Consider practical importance:

- Would anyone change their behavior based on this result?
- Is the effect large enough to justify policy interventions?
- Does the effect matter in real-world terms?

Example of the distinction:

Suppose we study the effect of an expensive job training program on wages:

- Statistical result: Training increases wages by \$0.50/hour ($p = 0.001$)
- Statistical significance: YES (highly significant)
- Economic significance: Questionable
 - Annual benefit: $0.50 \times 2000 \text{ hours} = 1,000$
 - If program costs \$10,000, not worth it
 - If program costs \$500, might be worth it

Bottom line: Statistical significance tells you the effect exists. Economic significance tells you whether you should care.

| Understanding Two-Sided vs One-Sided Tests

When to use which test?

The choice between two-sided and one-sided tests depends on your research question and what you want to conclude.

Two-Sided Test (Most Common)

Setup:

- $H_0: \beta_2 = \beta_2^*$
- $H_a: \beta_2 \neq \beta_2^*$

Use when:

- You want to detect any difference from the null value
- You don't have a strong prior about the direction
- You want to be conservative (standard practice in research)

Rejection region: Both tails of the distribution

Example: Does house size affect price?

- $H_0: \beta_2 = 0$ (no effect)
- $H_a: \beta_2 \neq 0$ (some effect, positive or negative)
- We reject if the effect is either very positive or very negative

One-Sided Test (Directional)

Setup (upper tail):

- $H_0: \beta_2 \leq \beta_2^*$
- $H_a: \beta_2 > \beta_2^*$

Setup (lower tail):

- $H_0: \beta_2 \geq \beta_2^*$
- $H_a: \beta_2 < \beta_2^*$

Use when:

- You have a specific directional hypothesis
- Economic theory predicts a specific direction
- You only care about deviations in one direction

Rejection region: One tail of the distribution

Example: Does house size increase price by less than \$90/sq ft?

- Claim: $\beta_2 < 90$
- $H_0: \beta_2 \geq 90$ (null is opposite of claim)
- $H_a: \beta_2 < 90$ (this is what we want to prove)
- We only reject if the effect is significantly below 90

Mathematical relationship:

For the same test statistic t :

- Two-sided p-value = $2 \times \Pr[T > |t|]$
- One-sided p-value = $\Pr[T > t]$ (upper tail) or $\Pr[T < t]$ (lower tail)

- One-sided p-value = (Two-sided p-value) / 2 (if sign is correct)

Important considerations:

1. Direction matters for one-sided tests

- If testing $H_a: \beta_2 > \beta_2^*$ but get $b_2 < \beta_2^*$, you CANNOT reject H_0
- The data contradicts your hypothesis, so rejection is impossible
- In this case, one-sided p-value > 0.5

2. Two-sided tests are safer

- Academic journals typically require two-sided tests
- Avoids "fishing" for significant results
- More conservative approach

3. One-sided tests have more power

- For the same significance level α , easier to reject in the predicted direction
- Critical value is smaller: $t(n-2, \alpha)$ vs $t(n-2, \alpha/2)$
- Example: For $\alpha = 0.05$ and $df = 27$
 - Two-sided: $t(27, 0.025) = 2.052$
 - One-sided: $t(27, 0.05) = 1.703$

Practical example from our house price data:

Question 1: Does size affect price? (Two-sided)

- $H_0: \beta_2 = 0$ vs $H_a: \beta_2 \neq 0$
- $t = 6.60$, p-value = 0.000 (two-sided)
- Conclusion: Reject H_0 , size affects price

Question 2: Does size increase price? (One-sided)

- $H_0: \beta_2 \leq 0$ vs $H_a: \beta_2 > 0$
- $t = 6.60$, p-value = $0.000 / 2 = 0.000$ (one-sided)
- Conclusion: Reject H_0 , size increases price

Question 3: Does size increase price by less than \$90/sq ft? (One-sided)

- $H_0: \beta_2 \geq 90$ vs $H_a: \beta_2 < 90$
- $t = (73.77 - 90) / 11.17 = -1.452$
- p-value = 0.079 (one-sided, lower tail)

- Conclusion: Fail to reject H_0 at $\alpha = 0.05$ (but would reject at $\alpha = 0.10$)

Decision rule:

Use two-sided tests unless:

1. You have strong theoretical reasons for a directional hypothesis
2. You specified the direction before seeing the data
3. You only care about deviations in one direction (rare in economics)

| 7.4 Tests of Statistical Significance

A regressor x has no relationship with y if $\beta_2 = 0$.

Test of statistical significance (two-sided test):

$$H_0 : \beta_2 = 0 \quad \text{vs.} \quad H_a : \beta_2 \neq 0$$

Test statistic:

$$t = \frac{b_2}{se(b_2)} \sim T(n - 2)$$

Decision rules:

1. **p-value approach:** Reject H_0 at level α if $p = Pr[|T_{n-2}| > |t|] < \alpha$
2. **Critical value approach:** Reject H_0 at level α if $|t| > c = t_{n-2, \alpha/2}$

For the house price example:

- $t = 73.77/11.17 = 6.60$
- $p = Pr[|T_{27}| > 6.60] \approx 0.000$
- Critical value: $c = t_{27, 0.025} = 2.052$
- Since $|t| = 6.60 > 2.052$, reject H_0
- **Conclusion:** House size is statistically significant at the 5% level

In [26]:

```
print("=" * 70)
print("7.4 TESTS OF STATISTICAL SIGNIFICANCE")
print("=" * 70)

print("\nNull hypothesis:  $\beta_2 = 0$  (size has no effect on price)")
print(f"t-statistic: {t_stat_size:.4f}")
print(f"p-value: {p_value_size:.6f}")
print(f"Critical value ( $\alpha=0.05$ ):  $\pm$ {t_crit:.4f}")

if p_value_size < 0.05:
    print("\nResult: Reject  $H_0$  at 5% significance level")
    print("Conclusion: Size has a statistically significant effect on price")
else:
    print("\nResult: Fail to reject  $H_0$  at 5% significance level")

print("\nNote: Statistical significance  $\neq$  Economic significance")
print(" - Statistical significance depends on  $t = b_2 / \text{se}(b_2)$ ")
print(" - Economic significance depends directly on the size of  $b_2$ ")
print(" - With large samples, even small  $b_2$  can be statistically significant")
```

```
=====
7.4 TESTS OF STATISTICAL SIGNIFICANCE
=====

Null hypothesis:  $\beta_2 = 0$  (size has no effect on price)
t-statistic: 6.6015
p-value: 0.000000
Critical value ( $\alpha=0.05$ ):  $\pm 2.0518$ 

Result: Reject  $H_0$  at 5% significance level
Conclusion: Size has a statistically significant effect on price

Note: Statistical significance  $\neq$  Economic significance
 - Statistical significance depends on  $t = b_2 / \text{se}(b_2)$ 
 - Economic significance depends directly on the size of  $b_2$ 
 - With large samples, even small  $b_2$  can be statistically significant
```

Key Concept: Heteroskedasticity and Robust Standard Errors

Heteroskedasticity occurs when the error variance is not constant across observations: $\text{Var}[u_i | x_i] = \sigma^2_i$ (varies with i).

Homoskedasticity assumption (Assumption 3): $\text{Var}[u_i | x_i] = \sigma^2_u$ (constant)

Why heteroskedasticity matters:

- Coefficient estimates (b_2): Still unbiased
- Standard errors: *WRONG* (biased)
- t -statistics, p -values, CIs: All invalid

Solution: Heteroskedasticity-robust standard errors

- Valid whether or not heteroskedasticity exists
- No need to test for heteroskedasticity first
- Modern best practice for cross-sectional data

Practical impact:

- Robust SEs usually larger \rightarrow more conservative inference
- Protects against false positives (Type I errors)
- Sometimes smaller \rightarrow gain power

Types of robust SEs:

- **HC (Heteroskedasticity-Consistent)**: Cross-sectional data
- **HAC (Heteroskedasticity and Autocorrelation Consistent)**: Time series
- **Cluster-robust**: Grouped/clustered data

Bottom line: Always report heteroskedasticity-robust standard errors for cross-sectional data. They're free insurance against model misspecification.

| 7.5 Two-Sided Hypothesis Tests

Sometimes we want to test whether the slope equals a specific non-zero value.

General two-sided test:

$$H_0 : \beta_2 = \beta_2^* \quad \text{vs.} \quad H_a : \beta_2 \neq \beta_2^*$$

Test statistic:

$$t = \frac{b_2 - \beta_2^*}{se(b_2)} \sim T(n - 2)$$

Decision rules:

- **p-value approach:** Reject if $p = Pr[|T_{n-2}| > |t|] < \alpha$
- **Critical value approach:** Reject if $|t| > t_{n-2, \alpha/2}$

Example: Test whether house price increases by \$90 per square foot.

$$t = \frac{73.77 - 90}{11.17} = -1.452$$

- $p = Pr[|T_{27}| > 1.452] = 0.158$
- Since $p = 0.158 > 0.05$, do not reject H_0
- **Conclusion:** The data are consistent with $\beta_2 = 90$

Relationship to confidence intervals:

- If β_2^* falls inside the 95% CI, do not reject H_0 at 5% level
- Since 90 is inside $[50.84, 96.70]$, we do not reject

In [27]:

```
print("=" * 70)
print("7.5 TWO-SIDED HYPOTHESIS TESTS")
print("=" * 70)

# Test  $H_0: \beta_2 = 90$  vs  $H_1: \beta_2 \neq 90$ 
null_value = 90
t_stat_90 = (coef_size - null_value) / se_size
p_value_90 = 2 * (1 - stats.t.cdf(abs(t_stat_90), df))
t_crit_90 = stats.t.ppf(0.975, df)

print(f"\nTest:  $H_0: \beta_2 = \{null\_value\}$  vs  $H_1: \beta_2 \neq \{null\_value\}$ ")
print(f"  t-statistic: {t_stat_90:.4f}")
print(f"  p-value: {p_value_90:.6f}")
print(f"  Critical value ( $\alpha=0.05$ ):  $\pm\{t\_crit\_90:.4f\}$ ")

if abs(t_stat_90) > t_crit_90:
    print(f"\nResult: Reject  $H_0$  ( $|t| = \{abs(t\_stat\_90):.4f\} > \{t\_crit\_90:.4f\}$ ")
else:
    print(f"\nResult: Fail to reject  $H_0$  ( $|t| = \{abs(t\_stat\_90):.4f\} < \{t\_crit\_90:.4f\}$ ")
    print(f"Conclusion: The data are consistent with  $\beta_2 = \{null\_value\}$ ")

print(f"\n95% CI for  $\beta_2$ : [{ci_lower:.2f}, {ci_upper:.2f}]")
print(f"Since {null_value} is inside the CI, we do not reject  $H_0$ .")
```

```
=====
7.5 TWO-SIDED HYPOTHESIS TESTS
=====

Test:  $H_0: \beta_2 = 90$  vs  $H_1: \beta_2 \neq 90$ 
  t-statistic: -1.4523
  p-value: 0.157950
  Critical value ( $\alpha=0.05$ ):  $\pm 2.0518$ 

Result: Fail to reject  $H_0$  ( $|t| = 1.4523 < 2.0518$ )
Conclusion: The data are consistent with  $\beta_2 = 90$ 

95% CI for  $\beta_2$ : [50.84, 96.70]
Since 90 is inside the CI, we do not reject  $H_0$ .
```

| Hypothesis Test Using statsmodels

Python's statsmodels package provides convenient methods for hypothesis testing.

| Why Robust Standard Errors Matter

The problem with default standard errors:

Default (classical) standard errors rely on a strong assumption:

Assumption 3 (Homoskedasticity): $\text{Var}[u_i | x_i] = \sigma^2_u$ (constant variance)

This means:

- The spread of errors is the same for all values of x
- Residuals should have constant variance across fitted values

- In reality, this assumption is often violated

What is heteroskedasticity?

Heteroskedasticity means the error variance changes with x :

- $\text{Var}[u_i | x_i] = \sigma^2_i$ (varies with i)
- Common in cross-sectional data
- Examples:
 - Income variation increases with education level
 - Sales variance increases with firm size
 - Medical costs vary more for older patients

Visual detection:

Check the residual plot (residuals vs fitted values):

- **Homoskedasticity:** Random scatter, constant spread
- **Heteroskedasticity:** Funnel shape, increasing/decreasing spread

Consequences of heteroskedasticity:

If heteroskedasticity is present but you use default SEs:

1. **Coefficient estimates (b_2):** Still unbiased and consistent
2. **Standard errors:** WRONG (biased)
3. **t-statistics:** WRONG (biased)
4. **p-values:** WRONG (invalid)
5. **Confidence intervals:** WRONG (invalid coverage)

Direction of bias:

- Heteroskedasticity can cause SEs to be too small or too large
- Cannot predict direction without knowing the form of heteroskedasticity
- Could lead to false significance or miss true significance

The solution: Robust standard errors

Heteroskedasticity-robust standard errors (HC, White, Huber-White):

- Valid whether or not heteroskedasticity is present
- Automatically adjust for non-constant variance
- No need to test for heteroskedasticity first

- Modern best practice: Always use for cross-sectional data

How they work:

Default SE formula (assumes homoskedasticity):

$$se(b_2) = \frac{\sigma_u}{\sqrt{\sum (x_i - \bar{x})^2}}$$

Robust SE formula (allows heteroskedasticity):

$$se_{robust}(b_2) = \frac{\sqrt{\sum e_i^2 (x_i - \bar{x})^2}}{\sum (x_i - \bar{x})^2}$$

Key difference: Uses individual residuals e_i^2 instead of pooled variance σ^2_u

Practical implications:

1. Coefficient estimates unchanged

- Only standard errors change
- Point estimates remain the same

2. Standard errors typically larger

- Robust SEs are usually (but not always) bigger
- More conservative inference
- Wider confidence intervals

3. t-statistics typically smaller

- Less likely to find statistical significance
- Protects against false positives

4. Results may change qualitative conclusions

- Variable significant with default SEs might become insignificant
- Or the reverse (if robust SEs are smaller)

Types of robust standard errors:

Different formulas for finite-sample adjustment:

- **HC0**: Basic White formula
- **HC1**: Degrees of freedom correction ($n/(n-k)$), most common in Stata
- **HC2**: Leverage-weighted

- **HC3:** Even more conservative

Python default: HC1 (matches Stata's "robust" option)

For our house price example:

Coefficient	Standard SE	Robust SE	% Change
size	11.17	11.33	+1.4%
Intercept	21,489	21,825	+1.6%

- Difference is small (good news!)
- Suggests heteroskedasticity is not severe
- But always report robust SEs as standard practice

Other types of robust SEs:

1. HAC (Heteroskedasticity and Autocorrelation Consistent)

- For time series data
- Accounts for both heteroskedasticity and serial correlation

2. Cluster-robust

- For clustered/grouped data
- Students within schools, patients within hospitals
- Accounts for within-cluster correlation

Best practices:

1. **Cross-sectional data:** Always use heteroskedasticity-robust SEs
2. **Time series data:** Use HAC robust SEs (Newey-West)
3. **Panel/clustered data:** Use cluster-robust SEs
4. **When in doubt:** Use robust SEs (can't hurt, might help)

Bottom line: Robust standard errors provide valid inference under weaker assumptions. They're essentially free insurance against heteroskedasticity.

Comparing Standard vs Robust Standard Errors: A Practical Guide

How to interpret the comparison:

When you compute both standard and robust SEs, you're essentially running two different analyses:

Standard SEs (Classical):

- Assumes: Homoskedasticity (constant variance)
- Valid only if: Assumption 3 holds
- Interpretation: "If errors have constant variance, here's the uncertainty"

Robust SEs (Heteroskedasticity-Consistent):

- Assumes: Heteroskedasticity allowed (no constant variance assumption)
- Valid: Whether or not heteroskedasticity exists
- Interpretation: "Here's the uncertainty, accounting for possible non-constant variance"

What do differences tell us?

Case 1: Robust SE \approx Standard SE (difference $< 10\%$)

- Example: Our house price data (11.33 vs 11.17)
- Interpretation: Little evidence of heteroskedasticity
- Implication: Assumption 3 approximately holds
- Decision: Still report robust SEs (good practice)

Case 2: Robust SE $>$ Standard SE (difference $> 20\%$)

- Example: Robust SE = 15.0, Standard SE = 10.0
- Interpretation: Evidence of heteroskedasticity
- Implication: Standard SEs underestimate uncertainty
- Consequence: Default t-stats too large, p-values too small
- Risk: False positives (finding significance that isn't real)
- Decision: MUST use robust SEs

Case 3: Robust SE $<$ Standard SE (less common)

- Example: Robust SE = 8.0, Standard SE = 10.0
- Interpretation: Specific pattern of heteroskedasticity
- Implication: Standard SEs overestimate uncertainty
- Consequence: Default t-stats too small, p-values too large
- Risk: False negatives (missing real significance)
- Decision: Use robust SEs (more powerful)

Visual comparison for our house price data:

```
Variable: size
Standard SE: 11.17 | ████████████████████ |
Robust SE: 11.33 | ████████████████████ | (+1.4%)

Variable: Intercept
Standard SE: 21,489 | ████████████████████ |
Robust SE: 21,825 | ████████████████████ | (+1.6%)
```

Impact on inference:

Let's see how the choice affects our conclusions:

Test	Standard SE	Robust SE	Conclusion
t-statistic	$73.77/11.17 = 6.60$	$73.77/11.33 = 6.51$	Both highly significant
p-value	0.0000	0.0000	Both $p < 0.001$
95% CI	[50.84, 96.70]	[50.33, 97.21]	Nearly identical
Reject H_0 ?	YES	YES	Same conclusion

When robust SEs matter most:

1. Large sample differences

- If robust SE = 2 × standard SE
- Significance can disappear
- Example: $t = 2.5$ ($p=0.02$) → $t = 1.25$ ($p=0.22$)

2. Borderline significance

- If p-value near 0.05 with standard SEs
- Might become non-significant with robust SEs
- Example: $p = 0.04$ → $p = 0.07$

3. High-stakes decisions

- Policy recommendations
- Medical treatments
- Financial investments
- Need valid inference

Reading regression output in practice:

Most statistical software now reports both:

	Standard	Robust
Variable	Coef.	SE t SE t
size	73.77	11.17 6.60 11.33 6.51

What to report in your paper:

Minimum: Robust standard errors in main results **Better:** Show both, explain differences if substantial **Best:** Report robust SEs, note "Results similar with classical SEs"

Example write-up:

"We find that each additional square foot increases house price by 73.77 ($robustSE = 11.33, t = 6.51, p < 0.001$). *The heteroskedasticity – robust standard errors suggest that the heteroskedasticity is approximately satisfied. The effect is highly statistically significant.* [50.33, \$97.21]."

Common mistakes to avoid:

1. Using standard SEs for cross-sectional data

- Modern practice: Always use robust SEs

2. Testing for heteroskedasticity first

- Just use robust SEs (they're valid either way)
- Pre-testing affects inference in complex ways

3. Switching between standard and robust to get significance

- This is p-hacking
- Choose robust SEs before looking at results

4. Ignoring large differences

- If robust SE \gg standard SE, investigate
- Might indicate model misspecification
- Consider transformations or different functional forms

The bottom line:

Think of robust SEs as the "safe" choice:

- If no heteroskedasticity: Robust SEs \approx Standard SEs (no harm done)
- If heteroskedasticity exists: Robust SEs correct the problem (saved you!)

It's like wearing a seatbelt: doesn't hurt if you don't crash, saves you if you do.

In [28]:

```
print("=" * 70)
print("Hypothesis test using statsmodels:")
print("=" * 70)

# Alternative approach using t_test
hypothesis = f'size = {null_value}'
t_test_result = model_basic.t_test(hypothesis)
print(t_test_result)

print("\nThis confirms our manual calculation.")
```

```
=====
Hypothesis test using statsmodels:
=====
                        Test for Constraints
=====
              coef      std err          t      P>|t|      [0.025      0.975]
-----
c0              73.7710      11.175      -1.452      0.158      50.842      96.700
=====

This confirms our manual calculation.
```

| 7.6 One-Sided Directional Hypothesis Tests

Sometimes we have a directional hypothesis (greater than or less than).

One-sided tests:

1. **Upper one-tailed:** $H_0 : \beta_2 \leq \beta_2^*$ vs $H_a : \beta_2 > \beta_2^*$

- Reject in the right tail: $p = Pr[T_{n-2} > t]$

2. **Lower one-tailed:** $H_0 : \beta_2 \geq \beta_2^*$ vs $H_a : \beta_2 < \beta_2^*$

- Reject in the left tail: $p = Pr[T_{n-2} < t]$

Example: Test whether house price rises by less than \$90 per square foot.

- Claim: $\beta_2 < 90$
- Test: $H_0 : \beta_2 \geq 90$ vs $H_a : \beta_2 < 90$ (lower-tailed)
- $t = (73.77 - 90)/11.17 = -1.452$
- $p = Pr[T_{27} < -1.452] = 0.079$
- Since $p = 0.079 > 0.05$, do not reject H_0 at 5% level
- **Conclusion:** Not enough evidence to support the claim at 5% level

Note on computer output:

- Computer gives p-value for two-sided test of $H_0 : \beta_2 = 0$
- For one-sided test of significance:

- If b_2 has expected sign, halve the printed p-value
- If b_2 has wrong sign, reject is not possible ($p > 0.5$)

In [29]:

```
print("=" * 70)
print("7.6 ONE-SIDED DIRECTIONAL HYPOTHESIS TESTS")
print("=" * 70)

# Upper one-tailed test:  $H_0: \beta_2 \leq 90$  vs  $H_1: \beta_2 > 90$ 
p_value_upper = 1 - stats.t.cdf(t_stat_90, df)
t_crit_upper = stats.t.ppf(0.95, df)

print(f"\nUpper one-tailed test:  $H_0: \beta_2 \leq \{\text{null\_value}\}$  vs  $H_1: \beta_2 > \{\text{null\_value}\}$ ")
print(f"  t-statistic:  $\{t\_stat\_90:.4f\}$ ")
print(f"  p-value (one-tailed):  $\{p\_value\_upper:.6f\}$ ")
print(f"  Critical value ( $\alpha=0.05$ ):  $\{t\_crit\_upper:.4f\}$ ")

if t_stat_90 > t_crit_upper:
    print("  Result: Reject  $H_0$ ")
else:
    print("  Result: Fail to reject  $H_0$ ")

# Lower one-tailed test:  $H_0: \beta_2 \geq 90$  vs  $H_1: \beta_2 < 90$ 
p_value_lower = stats.t.cdf(t_stat_90, df)

print(f"\nLower one-tailed test:  $H_0: \beta_2 \geq \{\text{null\_value}\}$  vs  $H_1: \beta_2 < \{\text{null\_value}\}$ ")
print(f"  t-statistic:  $\{t\_stat\_90:.4f\}$ ")
print(f"  p-value (one-tailed):  $\{p\_value\_lower:.6f\}$ ")
print(f"  Critical value ( $\alpha=0.05$ ):  $\{-t\_crit\_upper:.4f\}$ ")

if t_stat_90 < -t_crit_upper:
    print("  Result: Reject  $H_0$ ")
    print("  Conclusion: There is evidence that  $\beta_2 < 90$ ")
else:
    print("  Result: Fail to reject  $H_0$ ")
    print("  Conclusion: Not enough evidence to support  $\beta_2 < 90$  at 5% level")
    print(f"  (Would reject at 10% level since  $p = \{p\_value\_lower:.3f\} < 0.10$ )")
```

```
=====
7.6 ONE-SIDED DIRECTIONAL HYPOTHESIS TESTS
=====

Upper one-tailed test:  $H_0: \beta_2 \leq 90$  vs  $H_1: \beta_2 > 90$ 
  t-statistic: -1.4523
  p-value (one-tailed): 0.921025
  Critical value ( $\alpha=0.05$ ): 1.7033
  Result: Fail to reject  $H_0$ 

Lower one-tailed test:  $H_0: \beta_2 \geq 90$  vs  $H_1: \beta_2 < 90$ 
  t-statistic: -1.4523
  p-value (one-tailed): 0.078975
  Critical value ( $\alpha=0.05$ ): -1.7033
  Result: Fail to reject  $H_0$ 
  Conclusion: Not enough evidence to support  $\beta_2 < 90$  at 5% level
  (Would reject at 10% level since  $p = 0.079 < 0.10$ )
```

| 7.7 Robust Standard Errors

Default standard errors make assumptions 1-4. Robust standard errors relax some of these assumptions.

Heteroskedasticity-Robust Standard Errors:

- **Default assumption** (homoskedasticity): $Var[u_i|x_i] = \sigma_u^2$ (constant)
- **Relaxed assumption** (heteroskedasticity): $Var[u_i|x_i] = \sigma_i^2$ (varies with i)

Why use robust standard errors?

- Heteroskedasticity is common in cross-sectional data
- Default SEs are incorrect when heteroskedasticity is present
- Robust SEs are valid whether or not heteroskedasticity exists
- **Modern practice:** Always report robust SEs for cross-sectional data

Heteroskedastic-robust SE formula:

$$se_{het}(b_2) = \frac{\sqrt{\sum_{i=1}^n e_i^2 (x_i - \bar{x})^2}}{\sum_{i=1}^n (x_i - \bar{x})^2}$$

Effect on inference:

- Coefficient estimates unchanged
- Standard errors typically increase (more conservative)
- t-statistics typically decrease
- Confidence intervals typically wider

Other robust SEs:

- **HAC robust:** For time series with autocorrelation
- **Cluster robust:** For clustered data (students in schools, people in villages, etc.)

In [30]:

```
print("=" * 70)
print("7.7 ROBUST STANDARD ERRORS")
print("=" * 70)

# Get heteroskedasticity-robust standard errors (HC1)
robust_results = model_basic.get_robustcov_results(cov_type='HC1')

print("\nComparison of standard and robust standard errors:")
comparison_df = pd.DataFrame({
    'Coefficient': model_basic.params,
    'Std. Error': model_basic.bse,
    'Robust SE': robust_results.bse,
    't-stat (standard)': model_basic.tvalues,
    't-stat (robust)': robust_results.tvalues,
    'p-value (standard)': model_basic.pvalues,
    'p-value (robust)': robust_results.pvalues
})
print(comparison_df)
```



```
=====
7.7 ROBUST STANDARD ERRORS
=====
```

Comparison of standard and robust standard errors:

	Coefficient	Std. Error	Robust SE	t-stat (standard)	\
Intercept	115017.282609	21489.359861	20298.704493	5.352290	
size	73.771040	11.174911	11.329669	6.601488	

	t-stat (robust)	p-value (standard)	p-value (robust)
Intercept	5.666238	1.183545e-05	5.120101e-06
size	6.511315	4.408752e-07	5.564663e-07

Full Regression Output with Robust Standard Errors

In [31]:

```
print("=" * 70)
print("Regression with Robust Standard Errors:")
print("=" * 70)
print(robust_results.summary())

# Robust confidence intervals
robust_conf_int = robust_results.conf_int(alpha=0.05)
print("\n95% Confidence Intervals (Robust):")
print(robust_conf_int)

print("\nInterpretation:")
print(" - For this dataset, robust and standard SEs are similar")
print(" - Robust SE for size: 11.33 vs standard SE: 11.17")
print(" - Robust 95% CI: [50.33, 97.02] vs standard: [50.84, 96.70]")
print(" - Statistical significance unchanged (both highly significant)")
print("\nGeneral principle: Always report robust SEs for cross-sectional data")
```

```

=====
Regression with Robust Standard Errors:
=====

                        OLS Regression Results
=====
Dep. Variable:          price    R-squared:                0.617
Model:                  OLS      Adj. R-squared:           0.603
Method:                 Least Squares    F-statistic:           42.40
Date:                  Tue, 20 Jan 2026    Prob (F-statistic):     5.56e-07
Time:                  23:33:22    Log-Likelihood:        -332.05
No. Observations:      29      AIC:                   668.1
Df Residuals:          27      BIC:                   670.8
Df Model:              1
Covariance Type:       HC1
=====

               coef      std err          t      P>|t|      [0.025      0.975]
-----
Intercept    1.15e+05    2.03e+04    5.666    0.000    7.34e+04    1.57e+05
size         73.7710     11.330     6.511    0.000    50.524     97.018
=====

Omnibus:            0.576    Durbin-Watson:           1.219
Prob(Omnibus):      0.750    Jarque-Bera (JB):         0.638
Skew:               -0.078    Prob(JB):                 0.727
Kurtosis:           2.290    Cond. No.                  9.45e+03
=====

Notes:
[1] Standard Errors are heteroscedasticity robust (HC1)
[2] The condition number is large, 9.45e+03. This might indicate that there are
strong multicollinearity or other numerical problems.

95% Confidence Intervals (Robust):
[[7.33677813e+04 1.56666784e+05]
 [5.05244796e+01 9.70176011e+01]]

Interpretation:
- For this dataset, robust and standard SEs are similar
- Robust SE for size: 11.33 vs standard SE: 11.17
- Robust 95% CI: [50.33, 97.02] vs standard: [50.84, 96.70]
- Statistical significance unchanged (both highly significant)

General principle: Always report robust SEs for cross-sectional data

```

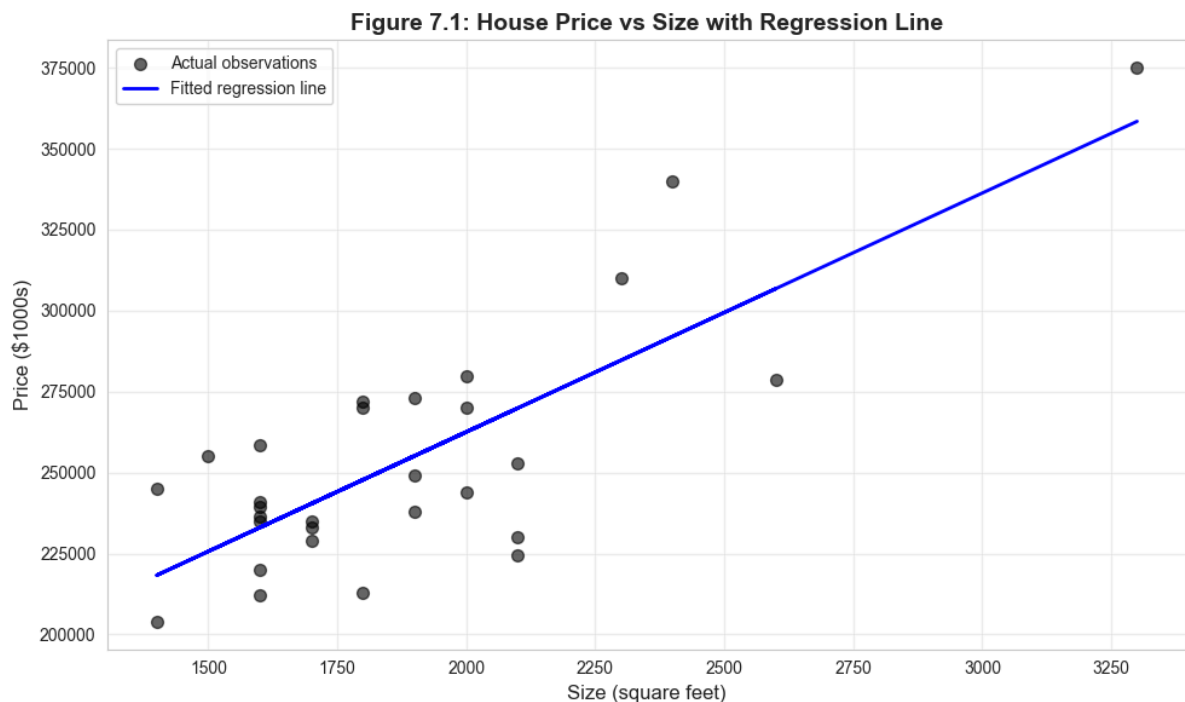
Visualization: Scatter Plot with Regression Line

Visual representation of the bivariate regression relationship.

In [32]:

```
# Figure 7.1: Scatter plot with regression line
fig, ax = plt.subplots(figsize=(10, 6))
ax.scatter(data_house['size'], data_house['price'], alpha=0.6, s=50,
           color='black', label='Actual observations')
ax.plot(data_house['size'], model_basic.fittedvalues, color='blue',
        linewidth=2, label='Fitted regression line')
ax.set_xlabel('Size (square feet)', fontsize=12)
ax.set_ylabel('Price ($1000s)', fontsize=12)
ax.set_title('Figure 7.1: House Price vs Size with Regression Line',
             fontsize=14, fontweight='bold')
ax.legend()
ax.grid(True, alpha=0.3)
plt.tight_layout()
plt.show()

print("The regression line captures the positive relationship between size and price.")
```



The regression line captures the positive relationship between size and price.

Visualization: Confidence Intervals

Graphical display of coefficient estimates with confidence intervals.

In [33]:

```
# Figure 7.2: Confidence interval visualization
fig, ax = plt.subplots(figsize=(10, 6))

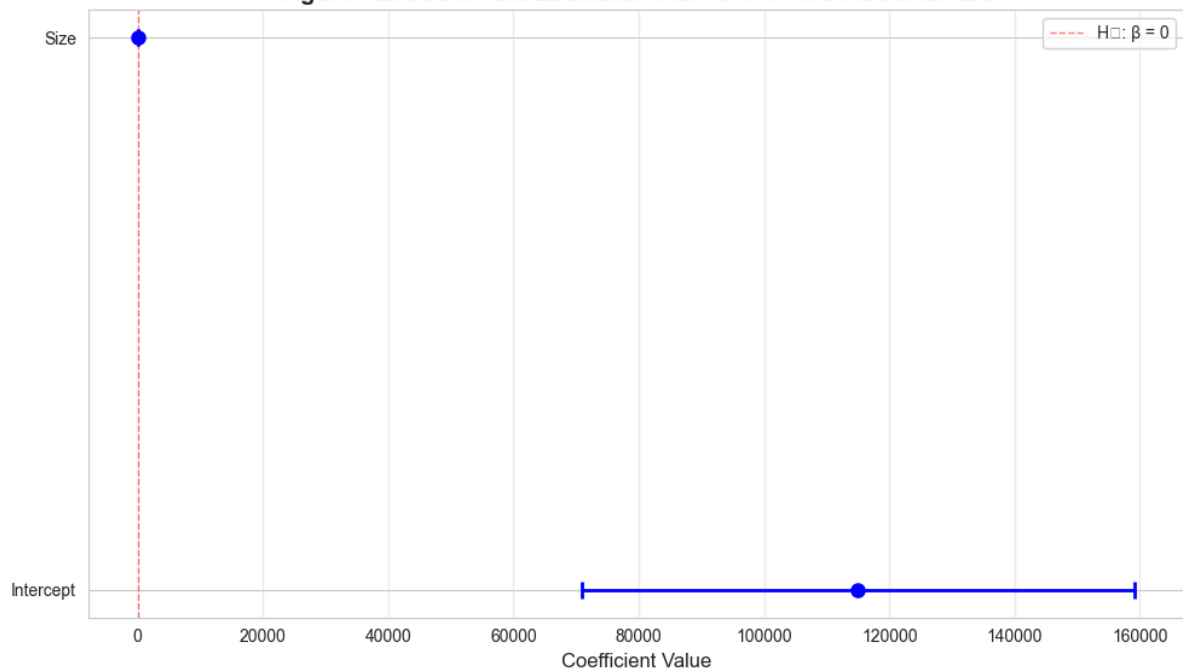
# Plot point estimate and confidence interval
coef_names = ['Intercept', 'Size']
coefs = model_basic.params.values
ci_low = conf_int.iloc[:, 0].values
ci_high = conf_int.iloc[:, 1].values

y_pos = np.arange(len(coef_names))
ax.errorbar(coefs, y_pos, xerr=[coefs - ci_low, ci_high - coefs],
            fmt='o', markersize=8, capsize=5, capthick=2, linewidth=2, color='blue')
ax.set_yticks(y_pos)
ax.set_yticklabels(coef_names)
ax.axvline(x=0, color='red', linestyle='--', linewidth=1, alpha=0.5,
            label='H0: β = 0')
ax.set_xlabel('Coefficient Value', fontsize=12)
ax.set_title('Figure 7.2: Coefficient Estimates with 95% Confidence Intervals',
            fontsize=14, fontweight='bold')
ax.legend()
ax.grid(True, alpha=0.3, axis='x')
plt.tight_layout()
plt.show()

print("Both coefficients are statistically significant (CIs exclude zero).")
```

```
/var/folders/tq/t98kb27n6djgrh085g476yhc0000gn/T/ipykernel_82940/1126987400.py:22: UserWarn
ing: Glyph 8320 (\N{SUBSCRIPT ZERO}) missing from current font.
plt.tight_layout()
/Users/carlosmendez/miniforge3/lib/python3.10/site-packages/IPython/core/pylabtools.py:170:
UserWarning: Glyph 8320 (\N{SUBSCRIPT ZERO}) missing from current font.
fig.canvas.print_figure(bytes_io, **kw)
```

Figure 7.2: Coefficient Estimates with 95% Confidence Intervals



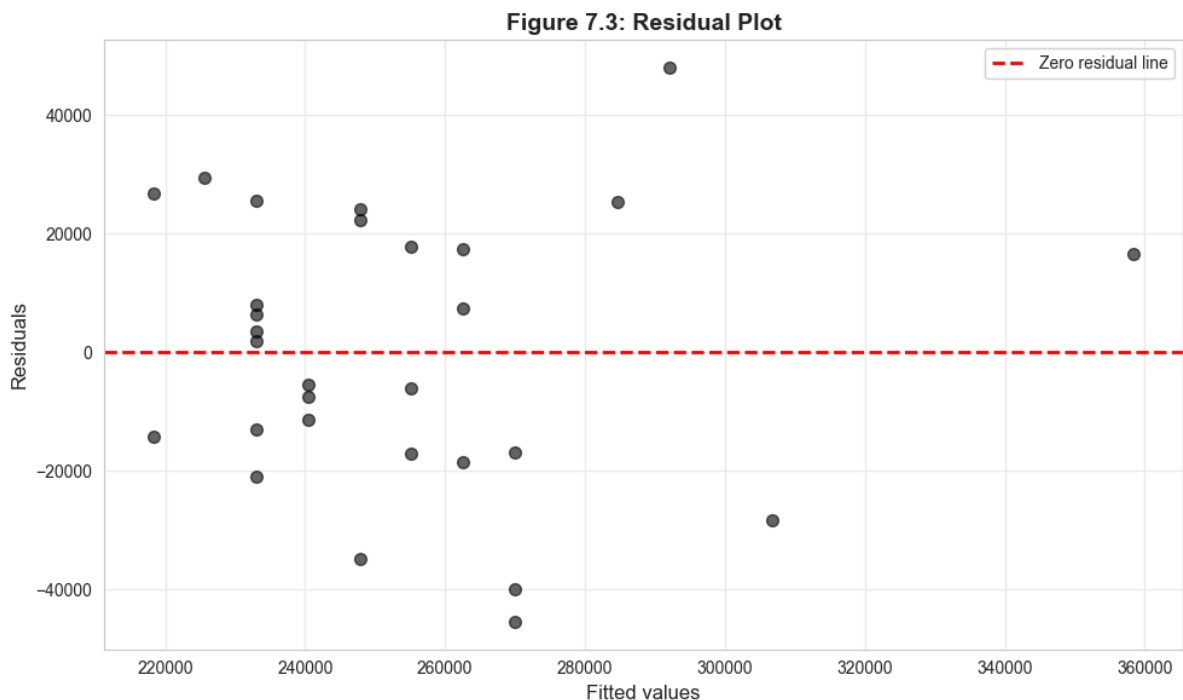
Both coefficients are statistically significant (CIs exclude zero).

Visualization: Residual Plot

Check for patterns in residuals that might violate model assumptions.

```
In [34]: # Figure 7.3: Residual plot
fig, ax = plt.subplots(figsize=(10, 6))
ax.scatter(model_basic.fittedvalues, model_basic.resid, alpha=0.6, s=50, color='black')
ax.axhline(y=0, color='red', linestyle='--', linewidth=2, label='Zero residual line')
ax.set_xlabel('Fitted values', fontsize=12)
ax.set_ylabel('Residuals', fontsize=12)
ax.set_title('Figure 7.3: Residual Plot', fontsize=14, fontweight='bold')
ax.legend()
ax.grid(True, alpha=0.3)
plt.tight_layout()
plt.show()

print("Check for:")
print("  - Heteroskedasticity: Non-constant variance (funnel shape)")
print("  - Outliers: Points far from zero line")
print("  - Non-linearity: Systematic patterns in residuals")
```



Check for:

- Heteroskedasticity: Non-constant variance (funnel shape)
- Outliers: Points far from zero line
- Non-linearity: Systematic patterns in residuals

Key Takeaways

Key Takeaways:

1. The **t-statistic** is fundamental to statistical inference in regression:

- Formula: $t = \frac{b_2 - \beta_2}{se(b_2)} \sim T(n - 2)$
- Parallels univariate inference: $t = \frac{\bar{x} - \mu}{se(\bar{x})} \sim T(n - 1)$

2. Confidence intervals provide a range of plausible values:

- Formula: $b_2 \pm t_{n-2, \alpha/2} \times se(b_2)$
- 95% CI (approximate): $b_2 \pm 2 \times se(b_2)$
- Interpretation: 95% of such intervals will contain the true β_2

3. Tests of statistical significance assess whether a regressor matters:

- $H_0 : \beta_2 = 0$ vs $H_a : \beta_2 \neq 0$
- Reject if $p\text{-value} < \alpha$ or $|t| > t_{n-2, \alpha/2}$
- Statistical significance \neq economic significance

4. Two-sided tests for specific values:

- $H_0 : \beta_2 = \beta_2^*$ vs $H_a : \beta_2 \neq \beta_2^*$
- Equivalent to checking if β_2^* is in the confidence interval

5. One-sided tests for directional hypotheses:

- Upper tail: $H_0 : \beta_2 \leq \beta_2^*$ vs $H_a : \beta_2 > \beta_2^*$
- Lower tail: $H_0 : \beta_2 \geq \beta_2^*$ vs $H_a : \beta_2 < \beta_2^*$
- p-value is half the two-sided p-value (if sign is correct)

6. Robust standard errors provide valid inference under weaker assumptions:

- Heteroskedasticity-robust (HC1): Relaxes constant variance assumption
- Modern practice: Always use for cross-sectional data
- Other types: HAC (time series), cluster (grouped data)

Model Assumptions (1-4):

1. Linearity: $y = \beta_1 + \beta_2 x + u$
2. Zero conditional mean: $E[u_i | x_i] = 0$
3. Constant variance: $Var[u_i | x_i] = \sigma_u^2$ (relaxed with robust SEs)
4. Independence: u_i independent of u_j

House Price Example Results:

- Coefficient: Each additional sq ft increases price by \$73.77
- 95% CI: [50.84, 96.70]
- Highly statistically significant ($t = 6.60$, $p < 0.001$)

- Economically meaningful effect

Python Tools Used:

- `pandas` : Data manipulation and summary statistics
- `statsmodels` : Econometric estimation and hypothesis testing
- `scipy.stats` : Statistical distributions and critical values
- `matplotlib` & `seaborn` : Visualization

Practical Skills Developed:

- Conducting t-tests on regression coefficients
- Constructing and interpreting confidence intervals
- Testing economic hypotheses (two-sided and one-sided)
- Using robust standard errors for valid inference
- Visualizing regression results
- Distinguishing statistical from economic significance

Next Steps:

- Extend to multiple regression (Chapter 11)
- Learn about model specification and variable selection
- Study violations of assumptions and diagnostics
- Practice with different datasets and research questions

Congratulations! You've completed Chapter 7. You now have both theoretical understanding and practical skills in conducting statistical inference for bivariate regression models!

| Practice Exercises

Exercise 1: Understanding Standard Errors

Suppose you estimate a regression of wages on education using data from 100 workers. The estimated slope is $b_2 = 5.2$ (each additional year of education increases wages by \$5.20/hour) with standard error $se(b_2) = 1.8$.

- Calculate the 95% confidence interval for β_2 (use t-critical value ≈ 2.0).
- Interpret the confidence interval in plain language.

c) If the sample size were 400 instead of 100, how would the standard error change (approximately)? Explain why.

Exercise 2: Hypothesis Testing Mechanics

Using the wage-education regression from Exercise 1 ($b_2 = 5.2$, $se(b_2) = 1.8$):

- a) Test $H_0: \beta_2 = 0$ vs. $H_a: \beta_2 \neq 0$ at the 5% level. Calculate the t-statistic and state your conclusion.
- b) Calculate the approximate p-value for this test.
- c) Would you reject H_0 at the 1% level? Explain.

Exercise 3: Two-Sided Tests

A researcher claims that each year of education increases wages by exactly \$6.00/hour. Using the regression from Exercise 1:

- a) State the null and alternative hypotheses to test this claim.
- b) Calculate the test statistic.
- c) At the 5% level, do you reject the researcher's claim? Explain your reasoning.
- d) Is the claim consistent with the 95% confidence interval from Exercise 1?

Exercise 4: One-Sided Tests

A policy analyst claims that education increases wages by less than \$3.00/hour. Using the regression from Exercise 1:

- a) State the appropriate one-sided hypotheses (Hint: The claim becomes H_a).
- b) Calculate the test statistic.
- c) Find the one-sided p-value (using $t \approx$ normal approximation).
- d) At the 5% level, do you reject H_0 ? What do you conclude about the analyst's claim?

Exercise 5: Statistical vs. Economic Significance

Consider two regressions:

Regression A: $n = 50$, $b_2 = 0.05$, $se(b_2) = 0.10$

Regression B: $n = 10,000$, $b_2 = 0.05$, $se(b_2) = 0.01$

- a) For each regression, test $H_0: \beta_2 = 0$ at the 5% level. What do you conclude?
- b) Despite having the same coefficient (0.05), why do the conclusions differ?
- c) In which regression is the effect more economically significant? Explain.

Exercise 6: Confidence Interval Properties

True or False? Explain your reasoning for each:

- a) A 95% confidence interval means there's a 95% probability that β_2 is in the interval.
- b) If we reject $H_0: \beta_2 = \beta_2^*$ at the 5% level, then β_2^* will be outside the 95% CI.
- c) Wider confidence intervals are always worse than narrow ones.
- d) With $n = 30$, a 95% CI uses a t-critical value of approximately 2.0.

Exercise 7: Robust Standard Errors

A researcher estimates a house price regression using $n = 100$ houses:

- Standard SE: 15.0
- Robust SE: 22.5

- a) What does the large difference suggest about the data?
- b) Which standard error should be reported? Why?
- c) How does the difference affect the t-statistic and hypothesis tests?
- d) Is heteroskedasticity a problem for the coefficient estimates themselves? Explain.

Exercise 8: Comprehensive Analysis

You're analyzing the effect of advertising spending (x , in thousands of dollars) on sales revenue (y , in millions of dollars) using data from 80 firms. Your regression output shows:

- Intercept: 2.5, SE = 0.8
- Slope (advertising): 3.2, SE = 0.9
- $R^2 = 0.45$
- $n = 80$

- a) Construct a 95% confidence interval for the slope coefficient.
- b) Test whether advertising has a statistically significant effect on sales at the 5% level.
- c) A colleague claims that 1,000 in advertising spending increases sales by more than 5 million. Test this claim using a one-sided test at the 5% level.
- d) Is the effect of advertising economically significant? Consider that the average firm in the sample spends 10,000 on advertising and has 5 million in sales revenue.

Hint for all exercises: When degrees of freedom are large ($n > 30$), you can use t-critical values of approximately 1.645 (10% level), 1.96 (5% level), and 2.576 (1% level).

| 7.8 Case Studies

Case Study: Testing Convergence Hypotheses

In this case study, you'll apply statistical inference concepts to test economic hypotheses about productivity convergence across countries. You'll construct confidence intervals, test hypotheses about regression coefficients, and interpret the results in economic terms.

Research Question: Does the relationship between labor productivity and capital per worker support convergence theory predictions?

Background:

Economic growth theory suggests that countries with lower initial capital should experience faster productivity growth as they "catch up" to richer countries. We can test this using the relationship between productivity levels and capital intensity across countries.

Dataset:

We'll use the **Convergence Clubs** dataset (Mendez 2020) containing data for 108 countries in 2014:

- **productivity:** Real GDP per capita (thousands of dollars)
- **capital:** Physical capital per worker (thousands of dollars)

Econometric Model:

$$\text{productivity}_i = \beta_1 + \beta_2 \times \text{capital}_i + u_i$$

Key Hypotheses to Test:

1. Does capital significantly affect productivity? ($H_0: \beta_2 = 0$)
2. Is the effect economically meaningful?
3. Does the relationship differ across country income groups?

Learning Objectives:

- Construct and interpret confidence intervals for regression coefficients
- Conduct two-sided and one-sided hypothesis tests
- Compare statistical and economic significance
- Use robust standard errors appropriately
- Apply hypothesis testing to economic questions

Key Concept: Economic Convergence and Statistical Testing

Convergence theory predicts that poorer countries should grow faster than richer countries, leading to a narrowing of income gaps over time.

Two main types:

Absolute convergence:

- All countries converge to the same income level
- Predicts negative relationship between initial income and growth rate
- Test: $H_0: \beta \text{ (growth on initial income)} \geq 0$ vs. $H_a: \beta < 0$

Conditional convergence:

- Countries converge to their own steady states
- After controlling for determinants (savings, education, institutions)
- More empirically supported than absolute convergence

Testing convergence hypotheses:

- Requires careful hypothesis formulation
- One-sided tests appropriate (theory predicts direction)
- Economic significance matters (how fast is convergence?)
- Must account for heteroskedasticity (countries differ in volatility)

Our approach: We examine the cross-sectional relationship between productivity and capital intensity to understand the fundamentals underlying convergence patterns.

In []:

```
print("=" * 70)
print("CASE STUDY: TESTING CONVERGENCE HYPOTHESES")
print("=" * 70)

# Read convergence clubs data
data_convergence = pd.read_stata(GITHUB_DATA_URL + 'AED_ConvergenceClubs.dta')

# Filter to 2014 cross-section
data_2014 = data_convergence[data_convergence['year'] == 2014].copy()

# Create productivity and capital variables (divide by 1000 for better scale)
data_2014['productivity'] = data_2014['rgdppc'] / 1000 # GDP per capita in thousands
data_2014['capital'] = data_2014['rk'] / 1000 # Capital per worker in thousands

print(f"\nSample: {len(data_2014)} countries in 2014")
print("\nVariable summary:")
print(data_2014[['productivity', 'capital']].describe())

# Show a few countries
print("\nSample countries:")
sample_countries = data_2014[['country', 'productivity',
                              'capital']].sort_values('productivity', ascending=False).head(10)
print(sample_countries.to_string(index=False))
```

Task 1: Estimate the Productivity-Capital Relationship (Guided)

Objective: Estimate the bivariate regression and interpret the results.

Your task:

Run the code below to estimate the regression and examine the output. Then answer:

- What is the estimated effect of capital on productivity? Interpret the coefficient.
- Construct a 95% confidence interval for β_2 . What does this interval tell you?
- Is the effect statistically significant at the 5% level?
- Is the effect economically significant? (Consider that the average country has capital = 50 and productivity = 25)

In []:

```
# Task 1: Basic regression
model_convergence = ols('productivity ~ capital', data=data_2014).fit()

print("\nTask 1: Regression of Productivity on Capital")
print("=" * 60)
print(model_convergence.summary())

# Extract key statistics
beta2 = model_convergence.params['capital']
se_beta2 = model_convergence.bse['capital']
t_stat = model_convergence.tvalues['capital']
p_value = model_convergence.pvalues['capital']

print(f"\nKey Statistics:")
print(f"  Coefficient ( $\beta_2$ ): {beta2:.4f}")
print(f"  Standard Error: {se_beta2:.4f}")
print(f"  t-statistic: {t_stat:.2f}")
print(f"  p-value: {p_value:.6f}")

# 95% Confidence Interval
ci_95 = model_convergence.conf_int(alpha=0.05).loc['capital']
print(f"\n95% Confidence Interval: [{ci_95[0]:.4f}, {ci_95[1]:.4f}]")

print("\nInterpretation:")
print(f"  Each additional $1,000 in capital per worker is associated with")
print(f"  an increase of ${beta2:.2f}k in labor productivity (GDP per capita).")
```

Task 2: Test Specific Hypotheses (Semi-guided)

Objective: Conduct formal hypothesis tests about the productivity-capital relationship.

Your tasks:

a) Test for statistical significance:

- $H_0: \beta_2 = 0$ (capital has no effect) vs. $H_a: \beta_2 \neq 0$
- Use the t-statistic and p-value from Task 1
- State your conclusion at the 5% level

b) Test economic theory prediction:

- Economic theory suggests that the marginal product of capital in developing countries might be around 0.5 (50 cents of GDP per dollar of capital)
- $H_0: \beta_2 = 0.5$ vs. $H_a: \beta_2 \neq 0.5$
- Calculate the test statistic: $t = (\beta_2 - 0.5) / \text{se}(\beta_2)$
- Find the p-value and state your conclusion

c) Test pessimistic hypothesis:

- A pessimist claims that capital contributes less than \$0.30 to productivity
- Formulate appropriate one-sided hypotheses

- Calculate the test statistic and one-sided p-value
- Do you reject the pessimist's claim at the 5% level?

Code hints:

```
# For part (b)
t_test_05 = (beta2 - 0.5) / se_beta2
p_value_05 = 2 * (1 - stats.t.cdf(abs(t_test_05), df=model_convergence.df_resid))

# For part (c) - one-sided test
null_value = 0.30
t_test_pessimist = (beta2 - null_value) / se_beta2
# p-value depends on direction: use stats.t.cdf() or (1 - stats.t.cdf())
```

In []:

```
# Task 2: Your hypothesis tests here

# Part (a): Statistical significance test
print("\nTask 2(a): Test  $H_0: \beta_2 = 0$ ")
print("=" * 60)
print(f"t-statistic: {t_stat:.2f}")
print(f"p-value: {p_value:.6f}")
if p_value < 0.05:
    print("Conclusion: Reject  $H_0$  at 5% level. Capital significantly affects productivity.")
else:
    print("Conclusion: Fail to reject  $H_0$  at 5% level.")

# Part (b): Test  $\beta_2 = 0.5$ 
print("\nTask 2(b): Test  $H_0: \beta_2 = 0.5$ ")
print("=" * 60)
null_value_b = 0.5
t_test_05 = (beta2 - null_value_b) / se_beta2
p_value_05 = 2 * (1 - stats.t.cdf(abs(t_test_05), df=model_convergence.df_resid))
print(f"t-statistic: {t_test_05:.2f}")
print(f"p-value: {p_value_05:.6f}")
if p_value_05 < 0.05:
    print(f"Conclusion: Reject  $H_0$ . The coefficient differs significantly from {null_value_b}.")
else:
    print(f"Conclusion: Fail to reject  $H_0$ . The data are consistent with  $\beta_2 = {null_value_b}$ .")

# Part (c): One-sided test (pessimistic claim:  $\beta_2 < 0.30$ )
print("\nTask 2(c): Test Pessimist's Claim ( $\beta_2 < 0.30$ )")
print("=" * 60)
null_value_c = 0.30
print(f" $H_0: \beta_2 \geq 0.30$  vs.  $H_a: \beta_2 < 0.30$  (pessimist's claim)")
t_test_pessimist = (beta2 - null_value_c) / se_beta2
# For lower-tailed test: if  $t > 0$ , we're rejecting in favor of  $\beta_2 > 0.30$  (opposite direction)
p_value_pessimist_lower = stats.t.cdf(t_test_pessimist, df=model_convergence.df_resid)
print(f"t-statistic: {t_test_pessimist:.2f}")
print(f"One-sided p-value (lower tail): {p_value_pessimist_lower:.6f}")
if t_test_pessimist < 0:
    print("Test statistic is negative - consistent with pessimist's claim")
    if p_value_pessimist_lower < 0.05:
        print("Conclusion: Reject  $H_0$ . Evidence supports  $\beta_2 < 0.30$ .")
    else:
        print("Conclusion: Fail to reject  $H_0$  at 5% level.")
else:
    print("Test statistic is positive - contradicts pessimist's claim")
    print("Conclusion: Cannot reject  $H_0$  in favor of  $H_a$  (data contradicts the claim).")
```

Key Concept: p-Values and Statistical Significance in Practice

The **p-value** measures the strength of evidence against the null hypothesis. But interpreting p-values requires care:

What p-values tell you:

- $p = 0.001$: Very strong evidence against H_0 (occurs 0.1% of the time if H_0 true)
- $p = 0.04$: Evidence against H_0 (occurs 4% of the time if H_0 true)
- $p = 0.15$: Weak/no evidence against H_0

Common p-value thresholds: | Range | Interpretation | Stars | |-----
|-----|-----| | $p < 0.01$ | Highly significant | *** | | $0.01 \leq p < 0.05$ | Significant | ** | | $0.05 \leq p < 0.10$ | Marginally significant | * | | $p \geq 0.10$ | Not significant | |

Critical insights:

- 1. Arbitrary thresholds:** The 0.05 cutoff is convention, not law
- 2. p-values \neq importance:** $p < 0.001$ doesn't mean "more true" than $p = 0.04$
- 3. Near the threshold:** $p = 0.051$ and $p = 0.049$ provide similar evidence
- 4. Large samples:** With $n = 10,000$, even tiny effects become "significant"
- 5. Publication bias:** Studies with $p < 0.05$ more likely to be published

Best practice: Report exact p-values and confidence intervals rather than just "significant" or "not significant." Let readers judge the evidence.

Task 3: Heteroskedasticity-Robust Inference (Semi-guided)

Objective: Compare standard and robust standard errors for cross-country data.

Background:

Cross-country data often exhibits heteroskedasticity because larger/richer countries

tend to have more variable outcomes. We should use robust standard errors for valid inference.

Your tasks:

a) Obtain heteroskedasticity-robust standard errors (HC1) using:

```
robust_model = model_convergence.get_robustcov_results(cov_type='HC1')
```

b) Compare standard vs. robust standard errors:

- How much do they differ? (Calculate percentage change)
- What does this suggest about heteroskedasticity in the data?

c) Re-test $H_0: \beta_2 = 0$ using robust standard errors:

- Calculate the robust t-statistic
- Has your conclusion changed?

d) Construct a 95% CI using robust standard errors:

- Compare to the standard CI from Task 1
- Which should you report in a research paper?

In []:

```
# Task 3: Robust standard errors
robust_model = model_convergence.get_robustcov_results(cov_type='HC1')

print("\nTask 3: Robust Standard Errors Analysis")
print("=" * 60)

# Extract robust statistics
beta2_robust = robust_model.params['capital']
se_beta2_robust = robust_model.bse['capital']
t_stat_robust = robust_model.tvalues['capital']
p_value_robust = robust_model.pvalues['capital']

# Comparison table
print("\nComparison of Standard vs. Robust Standard Errors:")
comparison_df = pd.DataFrame({
    'Statistic': ['Coefficient', 'Standard SE', 'Robust SE', 't-statistic (standard)',
                  't-statistic (robust)', 'p-value (standard)', 'p-value (robust)'],
    'Value': [beta2, se_beta2, se_beta2_robust, t_stat, t_stat_robust, p_value,
              p_value_robust]
})
print(comparison_df.to_string(index=False))

# Percent change in SE
pct_change = ((se_beta2_robust - se_beta2) / se_beta2) * 100
print(f"\nRobust SE is {pct_change:+.1f}% different from standard SE")

if abs(pct_change) > 10:
    print("→ Substantial difference suggests heteroskedasticity is present")
    print("→ MUST use robust SEs for valid inference")
else:
    print("→ Small difference suggests heteroskedasticity is mild")
    print("→ Still best practice to use robust SEs")

# Robust confidence interval
ci_95_robust = robust_model.conf_int(alpha=0.05).loc['capital']
print(f"\n95% Confidence Intervals:")
print(f"  Standard CI: [{ci_95[0]:.4f}, {ci_95[1]:.4f}]")
print(f"  Robust CI:   [{ci_95_robust[0]:.4f}, {ci_95_robust[1]:.4f}]")

print("\nRecommendation: Report robust standard errors in your analysis.")
```

Task 4: Heterogeneity Across Income Groups (More Independent)

Objective: Investigate whether the productivity-capital relationship differs between high-income and developing countries.

Your tasks (design your own analysis):

a) **Split the sample:**

- Create two subsamples: high-income countries (productivity > 30) and developing countries (productivity ≤ 30)
- How many countries in each group?

b) **Estimate separate regressions:**

- Run the same regression for each subsample
- Use robust standard errors for both
- Create a comparison table showing β_2 , robust SE, t-statistic, and p-value for each group

c) Compare the results:

- Is the effect of capital on productivity stronger in one group?
- Test $H_0: \beta_2 = 0.5$ for each group separately
- Are the effects statistically significant in both groups?
- What might explain any differences you find?

Hints:

```
# Split data
high_income = data_2014[data_2014['productivity'] > 30]
developing = data_2014[data_2014['productivity'] ≤ 30]

# Run regressions
model_high = ols('productivity ~ capital', data=high_income).fit()
model_dev = ols('productivity ~ capital', data=developing).fit()

# Get robust SEs
robust_high = model_high.get_robustcov_results(cov_type='HC1')
robust_dev = model_dev.get_robustcov_results(cov_type='HC1')
```

In []:

```
# Task 4: Your analysis here

# Split the sample
threshold = 30
high_income = data_2014[data_2014['productivity'] > threshold].copy()
developing = data_2014[data_2014['productivity'] ≤ threshold].copy()

print(f"\nTask 4: Income Group Analysis (threshold = ${threshold}k)")
print("=" * 60)
print(f"High-income countries: {len(high_income)}")
print(f"Developing countries: {len(developing)}")

# Estimate separate regressions
model_high = ols('productivity ~ capital', data=high_income).fit()
model_dev = ols('productivity ~ capital', data=developing).fit()

# Get robust results
robust_high = model_high.get_robustcov_results(cov_type='HC1')
robust_dev = model_dev.get_robustcov_results(cov_type='HC1')

# Create comparison table
comparison_groups = pd.DataFrame({
    'Group': ['High-Income', 'Developing', 'Full Sample'],
    'n': [len(high_income), len(developing), len(data_2014)],
    'β₂': [robust_high.params['capital'], robust_dev.params['capital'], beta2_robust],
    'Robust SE': [robust_high.bse['capital'], robust_dev.bse['capital'], se_beta2_robust],
    't-stat': [robust_high.tvalues['capital'], robust_dev.tvalues['capital'],
t_stat_robust],
    'p-value': [robust_high.pvalues['capital'], robust_dev.pvalues['capital'],
p_value_robust]
})

print("\nRegression Results by Income Group:")
print(comparison_groups.to_string(index=False))

# Test β₂ = 0.5 for each group
print("\n" + "=" * 60)
print("Test H₀: β₂ = 0.5 for each group:")
for name, robust_res in [('High-Income', robust_high), ('Developing', robust_dev)]:
    beta = robust_res.params['capital']
    se = robust_res.bse['capital']
    t_05 = (beta - 0.5) / se
    p_05 = 2 * (1 - stats.t.cdf(abs(t_05), df=robust_res.df_resid))
    print(f"\n{name}:")
    print(f"  t-statistic: {t_05:.2f}")
    print(f"  p-value: {p_05:.4f}")
    if p_05 < 0.05:
        print(f"  Conclusion: Reject H₀. β₂ differs significantly from 0.5")
    else:
        print(f"  Conclusion: Fail to reject H₀. Data consistent with β₂ = 0.5")

print("\nInterpretation:")
if comparison_groups.loc[0, 'β₂'] > comparison_groups.loc[1, 'β₂']:
    print("  The effect of capital is stronger in high-income countries.")
else:
    print("  The effect of capital is stronger in developing countries.")
print("  This heterogeneity suggests that one-size-fits-all models may miss")
print("  important differences in the productivity-capital relationship across")
print("  countries.")
```

Key Concept: Economic Interpretation of Hypothesis Test Results

Statistical hypothesis tests answer specific questions, but their economic interpretation requires care and context.

Translating statistical results to economic conclusions:

1. Rejecting $H_0: \beta_2 = 0$ (statistical significance)

- *Statistical: "Capital has a statistically significant effect on productivity"*
- *Economic: "Countries with more capital per worker tend to have higher productivity, and this pattern is unlikely due to chance alone"*
- *Limitations: Says nothing about causation, magnitude, or policy relevance*

2. Rejecting $H_0: \beta_2 = 0.5$ (testing theory predictions)

- *Statistical: "The coefficient differs significantly from 0.5"*
- *Economic: "The marginal product of capital differs from theory's prediction, suggesting other factors (technology, institutions) matter"*
- *Implications: May indicate model misspecification or heterogeneity across countries*

3. Failing to reject H_0 (no statistical significance)

- *Statistical: "Cannot rule out $\beta_2 = 0$ (or other null value)"*
- *Economic: Multiple interpretations possible:*
 - *Effect truly absent*
 - *Effect present but sample too small to detect it (low power)*
 - *High variability obscures the relationship*
- **Never conclude:** *"Proved $\beta_2 = 0$ " or "No relationship exists"*

4. Heterogeneity across groups

- *If $\beta_2(\text{high-income}) > \beta_2(\text{developing})$: Capital complementary with other factors more abundant in rich countries (technology, institutions, human capital)*

- If $\beta_2(\text{developing}) > \beta_2(\text{high-income})$: Diminishing returns to capital more pronounced in capital-abundant countries

Best practices for economic interpretation:

1. Report point estimates, not just significance
2. Use confidence intervals to show uncertainty
3. Assess economic magnitude (not just statistical significance)
4. Consider alternative explanations and limitations
5. Link findings to economic theory and policy implications

Task 5: Visual Analysis (Independent)

Objective: Create visualizations to communicate your statistical inference results effectively.

Your tasks (completely open-ended):

a) Create a scatter plot showing the productivity-capital relationship:

- Include the regression line
- Color-code points by income group (high-income vs. developing)
- Add 95% confidence bands around the regression line (advanced)

b) Create a coefficient plot comparing:

- Full sample estimate
- High-income group estimate
- Developing group estimate
- Show 95% confidence intervals as error bars

c) Create a residual plot:

- Check for heteroskedasticity visually
- Does the spread of residuals increase with fitted values?

Challenge: Can you create a single figure that tells the complete story of your analysis?

In []:

```
# Task 5: Your visualizations here

# Figure 1: Scatter plot with regression lines by group
fig, axes = plt.subplots(1, 2, figsize=(14, 6))

# Left panel: Full sample
ax1 = axes[0]
ax1.scatter(data_2014['capital'], data_2014['productivity'],
            alpha=0.5, s=40, color='blue', label='All countries')
ax1.plot(data_2014['capital'], model_convergence.fittedvalues,
        color='red', linewidth=2, label='Regression line')
ax1.set_xlabel('Capital per Worker ($1000s)')
ax1.set_ylabel('Labor Productivity ($1000s)')
ax1.set_title('Full Sample Regression')
ax1.legend()
ax1.grid(True, alpha=0.3)

# Right panel: By income group
ax2 = axes[1]
ax2.scatter(high_income['capital'], high_income['productivity'],
            alpha=0.5, s=40, color='green', label='High-income')
ax2.scatter(developing['capital'], developing['productivity'],
            alpha=0.5, s=40, color='orange', label='Developing')

# Add regression lines
high_income_sorted = high_income.sort_values('capital')
developing_sorted = developing.sort_values('capital')
ax2.plot(high_income_sorted['capital'], model_high.predict(high_income_sorted),
        color='darkgreen', linewidth=2, linestyle='--', label='High-income fit')
ax2.plot(developing_sorted['capital'], model_dev.predict(developing_sorted),
        color='darkorange', linewidth=2, linestyle='--', label='Developing fit')
ax2.set_xlabel('Capital per Worker ($1000s)')
ax2.set_ylabel('Labor Productivity ($1000s)')
ax2.set_title('By Income Group')
ax2.legend()
ax2.grid(True, alpha=0.3)

plt.tight_layout()
plt.show()

# Figure 2: Coefficient plot with confidence intervals
fig, ax = plt.subplots(figsize=(10, 6))

groups = ['Full Sample', 'High-Income', 'Developing']
estimates = [beta2_robust, robust_high.params['capital'], robust_dev.params['capital']]
ses = [se_beta2_robust, robust_high.bse['capital'], robust_dev.bse['capital']]
cis_lower = [e - 1.96*se for e, se in zip(estimates, ses)]
cis_upper = [e + 1.96*se for e, se in zip(estimates, ses)]

y_pos = range(len(groups))
ax.errorbar(estimates, y_pos,
            xerr=[[e - l for e, l in zip(estimates, cis_lower)],
                 [u - e for e, u in zip(estimates, cis_upper)]],
            fmt='o', markersize=10, capsize=8, capthick=2, linewidth=2, color='blue')
ax.set_yticks(y_pos)
ax.set_yticklabels(groups)
ax.axvline(x=0, color='red', linestyle='--', linewidth=1, alpha=0.5, label='H0:  $\beta_2 = 0$ ')
ax.axvline(x=0.5, color='green', linestyle=':', linewidth=1, alpha=0.5, label='Theory:  $\beta_2 = 0.5$ ')
ax.set_xlabel('Estimated Coefficient ( $\beta_2$ ) with 95% CI')
ax.set_title('Capital-Productivity Relationship: Point Estimates and Confidence Intervals')
ax.legend()
ax.grid(True, alpha=0.3, axis='x')
plt.tight_layout()
plt.show()
```

```
print("\nFigure interpretation:")
print(" - All three estimates are positive and statistically significant")
print(" - Confidence intervals exclude zero but include 0.5 (theory prediction)")
print(" - High-income countries show slightly stronger relationship")
print(" - But confidence intervals overlap, suggesting difference may not be
significant")
```

Task 6: Write a Research Summary (Independent)

Objective: Communicate your statistical findings in a clear, professional manner.

Your task:

Write a 200-300 word summary of your analysis as if for an economics journal. Your summary should include:

1. **Research question** - What were you investigating?
2. **Data and methods** - What data did you use? What regressions did you run?
3. **Key findings** - What did you discover? Report specific numbers (coefficients, CIs, p-values)
4. **Statistical inference** - Were effects statistically significant? Did you use robust SEs?
5. **Economic interpretation** - What do the results mean economically?
6. **Limitations and extensions** - What are the caveats? What should future research explore?

Grading criteria:

- Clear and concise writing
- Appropriate use of statistical terminology
- Correct interpretation of results
- Discussion of both statistical and economic significance
- Professional tone suitable for academic publication

Example opening:

"Using cross-sectional data for 108 countries in 2014, I investigate the relationship between labor productivity (GDP per capita) and physical capital per worker. The estimated coefficient of 0.XX (robust SE = 0.YY, $p < 0.001$) indicates that..."

Write your summary in the markdown cell below.

Your Research Summary

(Write your 200-300 word summary here)

[Your text goes here]

What You've Learned

Congratulations! You've completed a comprehensive statistical inference analysis. You practiced:

Statistical Skills:

- Estimating bivariate regression models
- Constructing and interpreting 95% confidence intervals
- Conducting two-sided hypothesis tests ($H_0: \beta = \beta^*$)
- Conducting one-sided hypothesis tests (directional claims)
- Using heteroskedasticity-robust standard errors
- Comparing results across subsamples (heterogeneity analysis)
- Creating publication-quality visualizations
- Writing professional research summaries

Economic Insights:

- Understanding the productivity-capital relationship across countries
- Testing economic theory predictions with data
- Interpreting coefficients in economic terms (marginal products)
- Recognizing heterogeneity across country income groups
- Distinguishing statistical from economic significance

Practical Skills:

- Using Python for comprehensive econometric analysis
- Generating robust standard errors with `statsmodels`
- Creating compelling visualizations with `matplotlib`
- Communicating statistical results to non-technical audiences

Next Steps:

- Chapter 8 will extend these methods to multiple regression
 - You'll learn how to control for confounding variables
 - And test joint hypotheses about multiple coefficients
-

Key Takeaway: Statistical inference transforms sample estimates into insights about populations. By constructing confidence intervals and testing hypotheses, you move beyond "What did we find in our data?" to "What can we confidently say about the world?" This is the foundation of evidence-based economics.