

Consider the model $y_i = \mu + \beta \sin(x_i + \eta) + \epsilon_i$ where $x_i = \frac{2\pi i}{n}$ for $i = 0, 1, \dots, n$, $\epsilon_i \sim \mathcal{N}(0, 1)$ independently for each i , and $\eta \in [0, \pi]$.

1. (25 pts) Assume for now that η is known. What is the maximum likelihood estimator for (μ, β) ?

Solution: The model can be written as $\mathbf{y} = \mathbf{X}\boldsymbol{\theta} + \boldsymbol{\epsilon}$, where

$$\mathbf{y} = \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix} \quad \mathbf{X} = \begin{pmatrix} 1 & \sin(x_0 + \eta) \\ \vdots & \vdots \\ 1 & \sin(x_n + \eta) \end{pmatrix} \quad \boldsymbol{\theta} = \begin{pmatrix} \mu \\ \beta \end{pmatrix} \quad \boldsymbol{\epsilon} = \begin{pmatrix} \epsilon_1 \\ \vdots \\ \epsilon_n \end{pmatrix}$$

The MLE is simply $\boldsymbol{\theta} = \{\mathbf{X}^T \mathbf{X}\}^{-1} \mathbf{X}^T \mathbf{y}$ with

$$\mathbf{X}^T \mathbf{X} = \begin{pmatrix} n & \sum_{i=0}^n \sin(x_i + \eta) \\ \sum_{i=0}^n \sin(x_i + \eta) & \sum_{i=0}^n \sin^2(x_i + \eta) \end{pmatrix}$$

$$\mathbf{X}^T \mathbf{y} = \begin{pmatrix} \sum_{i=0}^n y_i \\ \sum_{i=0}^n y_i \sin(x_i + \eta) \end{pmatrix}$$

so that

$$\hat{\mu} = \frac{\{\sum_{i=0}^n y_i\} \{\sum_{i=0}^n \sin^2(x_i + \eta)\} - \{\sum_{i=0}^n y_i \sin(x_i + \eta)\} \{\sum_{i=0}^n \sin(x_i + \eta)\}}{n \sum_{i=0}^n \sin^2(x_i + \eta) - \{\sum_{i=0}^n \sin(x_i + \eta)\}^2}$$

and

$$\hat{\beta} = \frac{n \{\sum_{i=0}^n y_i \sin(x_i + \eta)\} - \{\sum_{i=0}^n y_i\} \{\sum_{i=0}^n \sin(x_i + \eta)\}}{n \sum_{i=0}^n \sin^2(x_i + \eta) - \{\sum_{i=0}^n \sin(x_i + \eta)\}^2}.$$

2. (35 pts) Assume now that η is unknown. Compute the maximum likelihood estimator for (μ, β, η) (Hint: You can use simple trigonometric identities to write $\beta \sin(x_i + \eta) = \beta \sin x_i \cos \eta + \beta \cos x_i \sin \eta$. Also, note that $\sum_{i=0}^n \sin x_i = 0$ and $\sum_{i=0}^n \sin x_i \cos x_i = 0$.)

Solution: The proposed alternative representation is valid because of the well-known trigonometric identity $\sin(\alpha + \gamma) = \sin \alpha \cos \gamma + \cos \alpha \sin \gamma$. Note that the model can now be written as

$$\mathbf{y} = \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix} \quad \mathbf{X} = \begin{pmatrix} 1 & \sin x_0 & \cos x_0 \\ \vdots & \vdots & \vdots \\ 1 & \sin x_n & \cos x_n \end{pmatrix} \quad \boldsymbol{\theta} = \begin{pmatrix} \mu \\ \phi_1 \\ \phi_2 \end{pmatrix} \quad \boldsymbol{\epsilon} = \begin{pmatrix} \epsilon_1 \\ \vdots \\ \epsilon_n \end{pmatrix}$$

where $\phi_1 = \beta \cos \eta$ and $\phi_2 = \beta \sin \eta$. In this case

$$\mathbf{X}^T \mathbf{X} = \begin{pmatrix} n & \sum_{i=0}^n \sin x_i & \sum_{i=0}^n \cos x_i \\ \sum_{i=0}^n \sin x_i & \sum_{i=0}^n \sin^2 x_i & \sum_{i=0}^n \sin x_i \cos x_i \\ \sum_{i=0}^n \cos x_i & \sum_{i=0}^n \sin x_i \cos x_i & \sum_{i=0}^n \cos^2 x_i \end{pmatrix},$$

$$\mathbf{X}^T \mathbf{y} = \begin{pmatrix} \sum_{i=0}^n y_i \\ \sum_{i=0}^n y_i \sin x_i \\ \sum_{i=0}^n y_i \cos x_i \end{pmatrix}.$$

They can proceed with this general form, but it is easier to note that $\sum_{i=0}^n \sin x_i = 0$ and $\sum_{i=0}^n \sin x_i \cos x_i = 0$ (since both are odd functions around π and the x_i s are equally spaced), so that

$$\mathbf{X}^T \mathbf{X} = \begin{pmatrix} n & 0 & \sum_{i=0}^n \cos x_i \\ 0 & \sum_{i=0}^n \sin^2 x_i & 0 \\ \sum_{i=0}^n \cos x_i & 0 & \sum_{i=0}^n \cos^2 x_i \end{pmatrix}.$$

Now, if we work with (ϕ_1, μ, ϕ_2) instead of (μ, ϕ_1, ϕ_2) we obtain a block diagonal matrix that is easier to invert and

$$\hat{\mu} = \frac{\{\sum_{i=0}^n y_i\} \{\sum_{i=0}^n \cos^2 x_i\} - \{\sum_{i=0}^n y_i \cos x_i\} \{\sum_{i=0}^n \cos x_i\}}{n \sum_{i=0}^n \cos^2 x_i - \{\sum_{i=0}^n \cos x_i\}^2},$$

$$\hat{\phi}_1 = \frac{\sum_{i=0}^n y_i \sin x_i}{\sum_{i=0}^n \sin^2 x_i},$$

and

$$\hat{\phi}_2 = \frac{n \{\sum_{i=0}^n y_i \cos x_i\} - \{\sum_{i=0}^n y_i\} \{\sum_{i=0}^n \cos x_i\}}{n \sum_{i=0}^n \cos^2 x_i - \{\sum_{i=0}^n \cos x_i\}^2}.$$

Because of invariance, the MLE of η is simply $\hat{\eta} = \arctan \left\{ \frac{\hat{\phi}_2}{\hat{\phi}_1} \right\}$.

3. (40 pts) Assume that you are interested in testing the hypotheses $H_0 : \eta = \pi/4$ vs. $H_a : \eta \neq \pi/4$. Derive the likelihood ratio test for this pair of hypotheses. Please keep in mind that we are assuming that the variance is known!

Solution: Under the null, the model is equivalent to setting $\phi_1 = \phi_2 = \beta$. Hence, testing $H_0 : \eta = \pi/4$ vs. $H_a : \eta \neq \pi/4$ is equivalent to testing the general linear hypotheses:

$$H_0 : (1, 0, -1) \begin{pmatrix} \phi_1 \\ \mu \\ \phi_2 \end{pmatrix} = 0 \quad \text{vs.} \quad H_a : (1, 0, -1) \begin{pmatrix} \phi_1 \\ \mu \\ \phi_2 \end{pmatrix} \neq 0,$$

Note that

$$(1, 0, -1) \begin{pmatrix} \hat{\phi}_1 \\ \hat{\mu} \\ \hat{\phi}_2 \end{pmatrix} = \frac{\sum_{i=0}^n y_i \sin x_i}{\sum_{i=0}^n \sin^2 x_i} - \frac{n \{ \sum_{i=0}^n y_i \cos x_i \} - \{ \sum_{i=0}^n y_i \} \{ \sum_{i=0}^n \cos x_i \}}{n \sum_{i=0}^n \cos^2 x_i - \{ \sum_{i=0}^n \cos x_i \}^2},$$

and that

$$\begin{aligned} (1, 0, -1) \begin{pmatrix} \frac{1}{\sum_{i=0}^n \sin^2 x_i} & 0 & 0 \\ 0 & \frac{\sum_{i=0}^n \cos^2 x_i}{n \sum_{i=0}^n \cos^2 x_i - \{ \sum_{i=0}^n \cos x_i \}} & \frac{-\sum_{i=0}^n \cos x_i}{n \sum_{i=0}^n \cos^2 x_i - \{ \sum_{i=0}^n \cos x_i \}} \\ 0 & \frac{-\sum_{i=0}^n \cos x_i}{n \sum_{i=0}^n \cos^2 x_i - \{ \sum_{i=0}^n \cos x_i \}} & \frac{n}{n \sum_{i=0}^n \cos^2 x_i - \{ \sum_{i=0}^n \cos x_i \}} \end{pmatrix} \begin{pmatrix} 1 \\ 0 \\ -1 \end{pmatrix} \\ = \frac{1}{\sum_{i=0}^n \sin^2 x_i} + \frac{n}{n \sum_{i=0}^n \cos^2 x_i - \{ \sum_{i=0}^n \cos x_i \}}. \end{aligned}$$

Therefore

$$U = \frac{\left(\frac{\sum_{i=0}^n y_i \sin x_i}{\sum_{i=0}^n \sin^2 x_i} - \frac{n \{ \sum_{i=0}^n y_i \cos x_i \} - \{ \sum_{i=0}^n y_i \} \{ \sum_{i=0}^n \cos x_i \}}{n \sum_{i=0}^n \cos^2 x_i - \{ \sum_{i=0}^n \cos x_i \}^2} \right)^2}{\frac{1}{\sum_{i=0}^n \sin^2 x_i} + \frac{n}{n \sum_{i=0}^n \cos^2 x_i - \{ \sum_{i=0}^n \cos x_i \}}} \sim \chi_1^2$$

So the test proceeds by computing U_{obs} using the sample and rejecting H_0 if $U_{obs} > \chi_1^2(1 - \alpha)$ where $\chi_1^2(1 - \alpha)$ denotes the $1 - \alpha$ quantile of the chi squared distribution with one degree of freedom.