

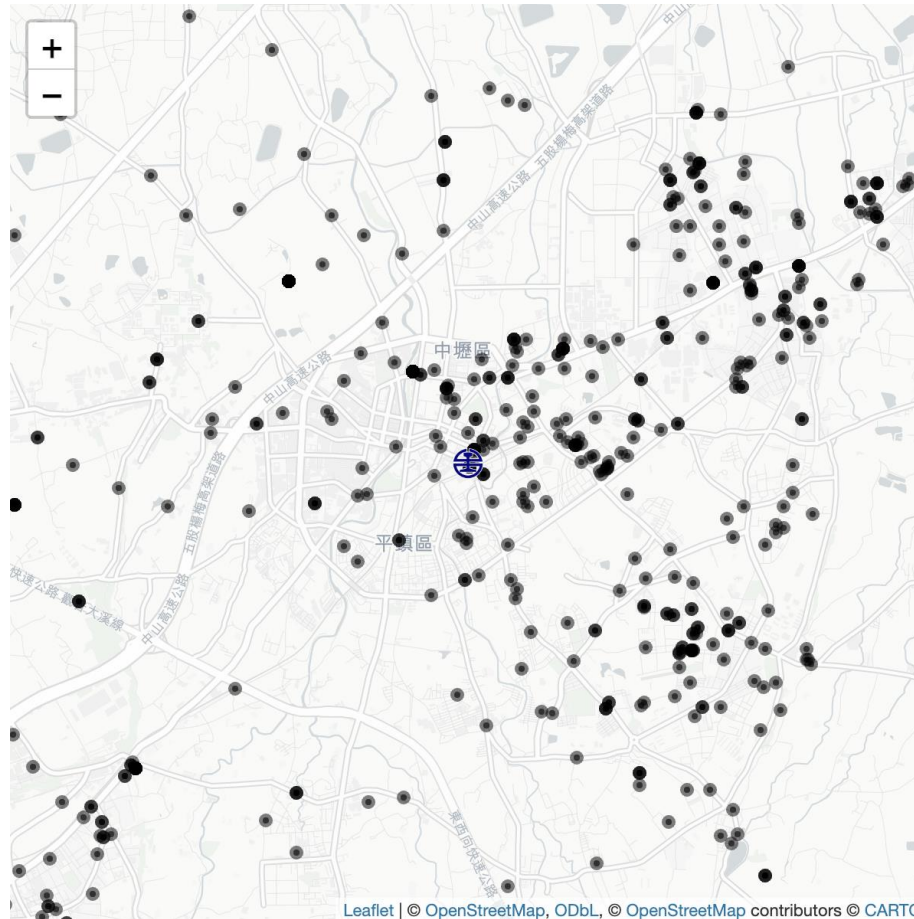
Introduction to spatial point process and its Fourier analysis

Qi-Wen Ding

Fundamentals

Point pattern data

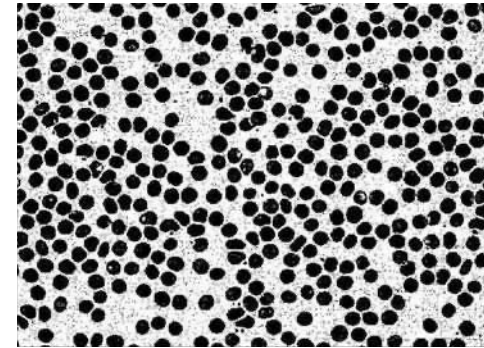
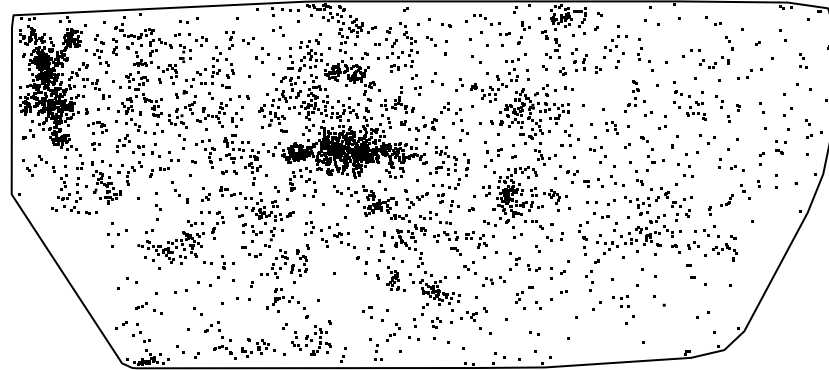
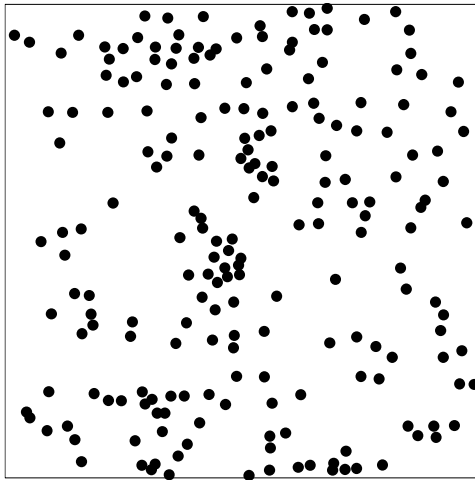
Spatial **locations** of things or events



For a study area, **all locations are observable!**
Locations without a point = observed but no event happened!

Point pattern data

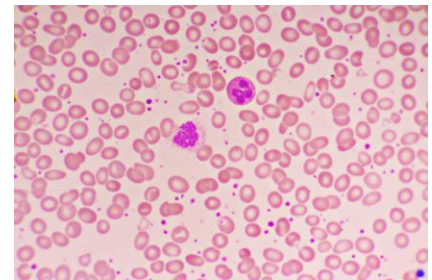
Spatial **locations** of things or events



Japanese black pine seedlings



Sky positions of 4215 galaxies in the Shapley Supercluster



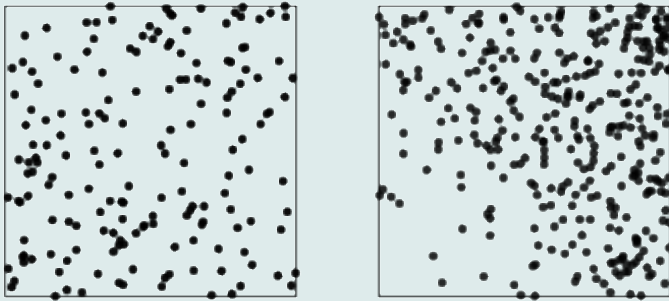
Red blood particles

For a study area, **all locations are observable!**
Locations without a point = observed but no event happened!

Characteristics of point process

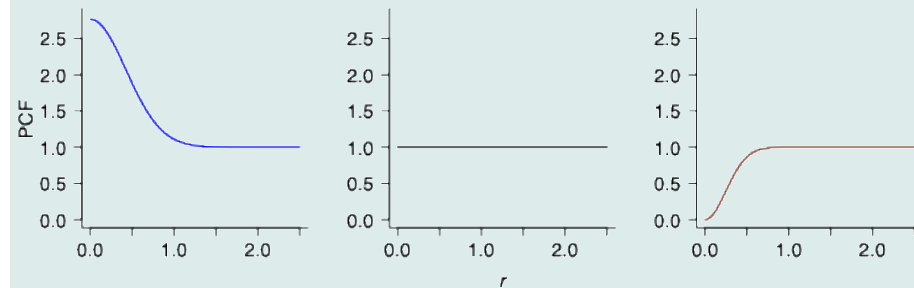
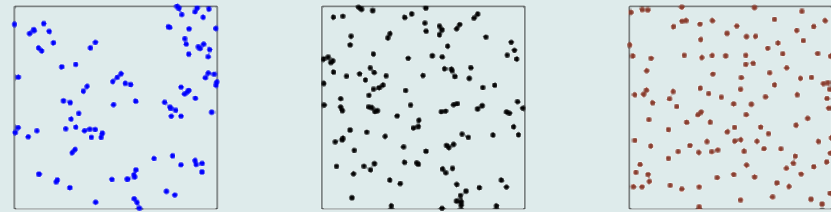
Intensity function: $\lambda(\mathbf{x})$

$\propto \mathbb{P}(\text{Having a point at location } \mathbf{x})$

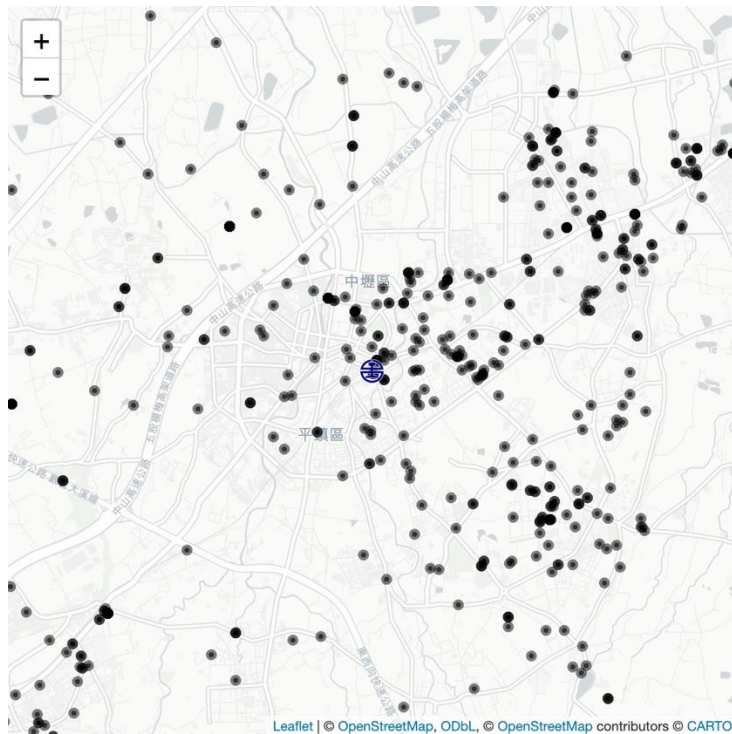


Pair correlation function: $g(r)$

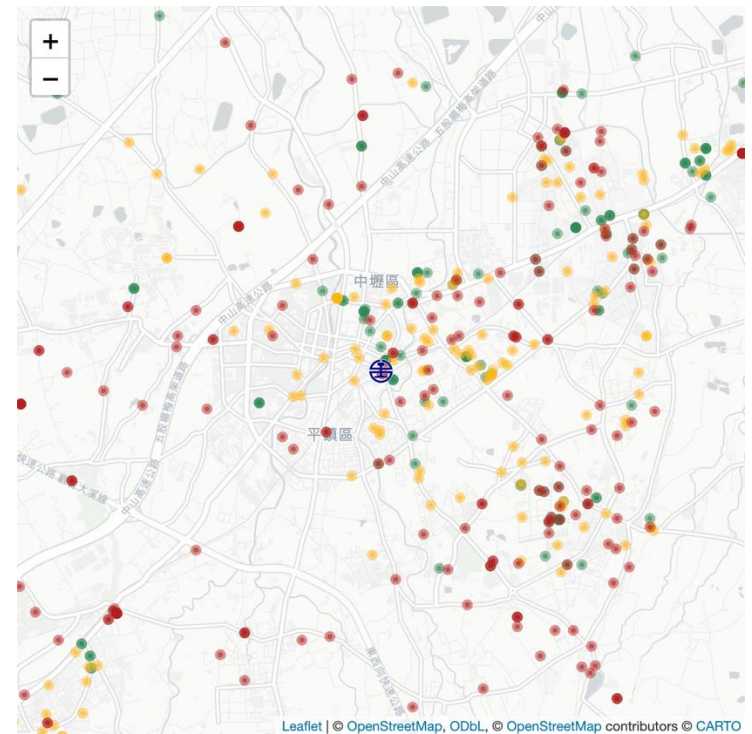
How events interact at radius r



From univariate to multivariate



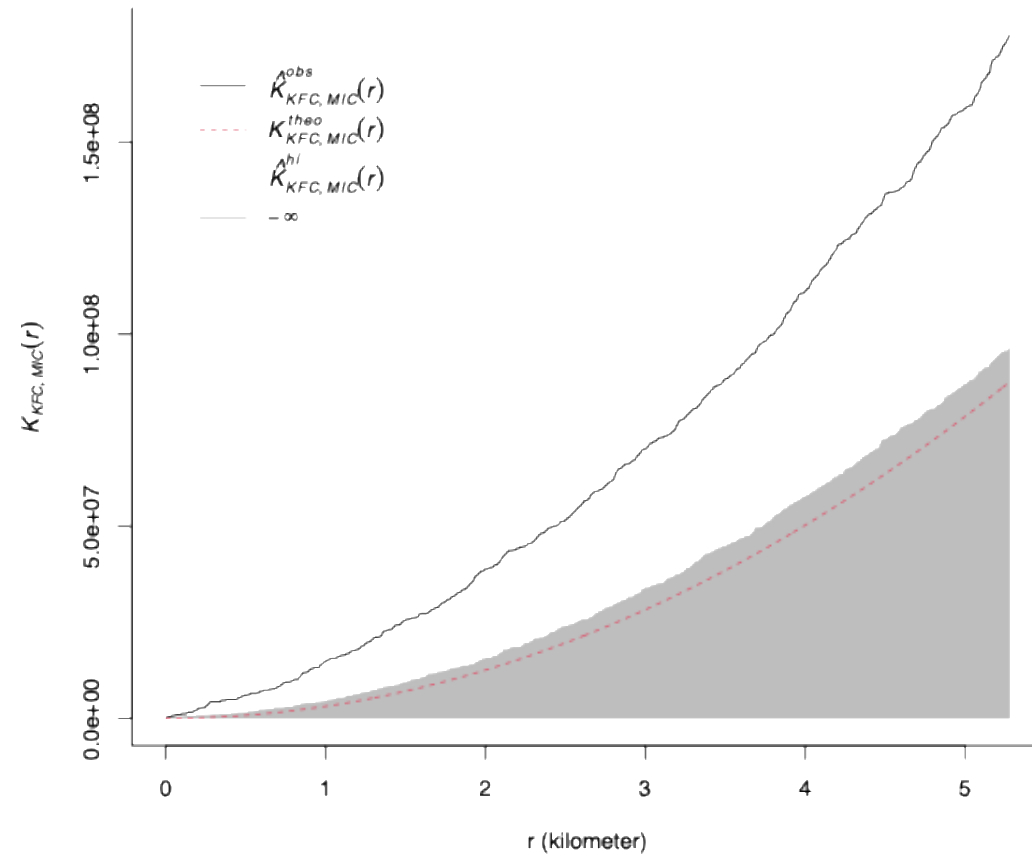
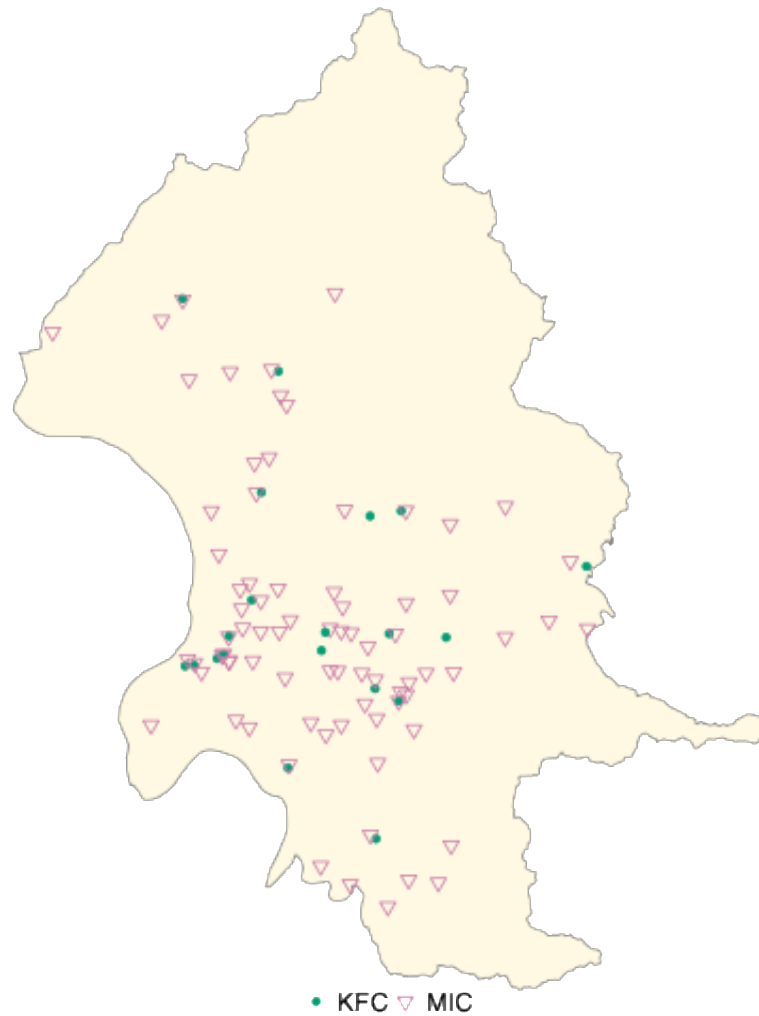
● Burglary



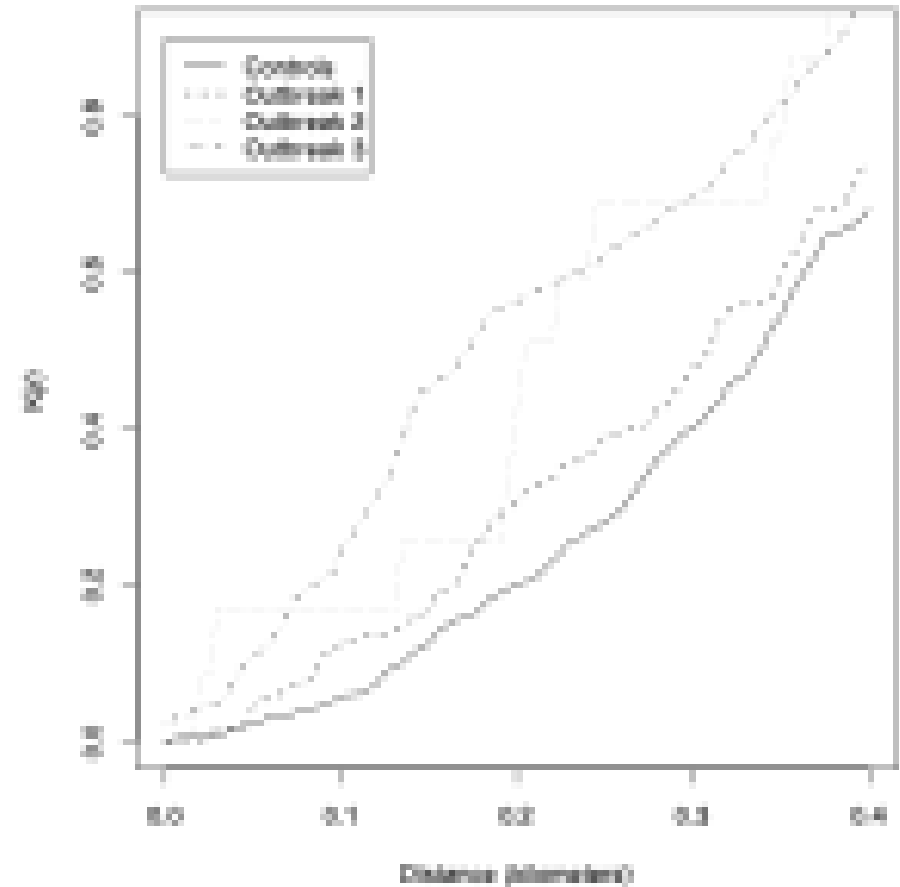
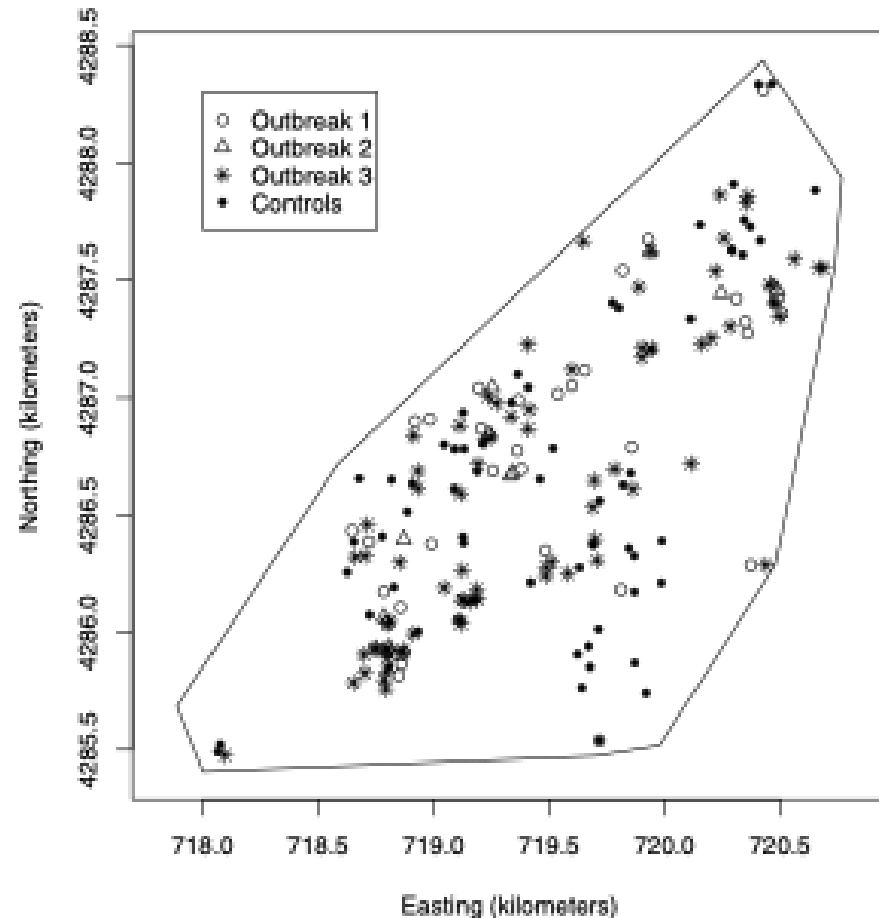
Type of burglary

● Motorcycle ● Car ● Bicycle

Example 1: Type interaction



Example 2: Source detection of infectious disease⁸



Fourier analysis of point processes

Time series vs. spatial point processes

Time series model

- White noise : $e_t \stackrel{\text{iid}}{\sim} \mathcal{N}(0, \sigma^2)$
- MA(1) : $X_t = e_t + \theta e_{t-1}$
- AR(1) : $X_t = \alpha X_{t-1} + e_t$

ACF

- $\gamma(k) = \sigma^2 \mathbf{1}\{k = 0\}$
- $\gamma(k) = \theta \sigma^2, k = \pm 1$
- $\gamma(k) = \alpha^{|k|} \sigma_X^2$

Spectral density

- $f(\omega) = \sigma^2 / (2\pi)$
- $f(\omega) = \sigma^2 (1 + \theta^2 + 2\theta \cos \omega) / (2\pi)$
- $f(\omega) = \frac{\sigma^2}{2\pi(1 - 2\alpha \cos \omega + \alpha^2)}$

Point process model

- Homogeneous Poisson process
- Thomas cluster process
- Determinantal point process

PCF

- $g(r) = 1$
- $g(r) = 1 + \frac{e^{-r^2/4\sigma^2}}{4\pi\kappa\sigma^2}$
- $g(r) = 1 - \left(e^{-r^2/\sigma^2}\right)^2$

Spectral density

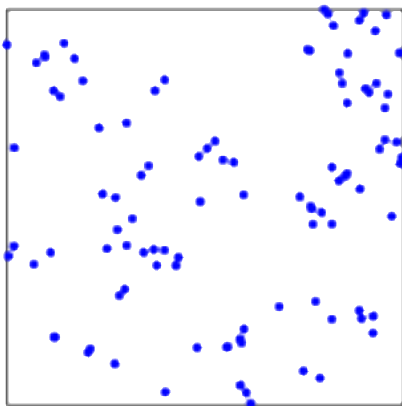
- $f(\omega) = (2\pi)^{-d} \lambda$
- $f(\omega) = (2\pi)^{-d} \kappa \mu \left[1 + \mu e^{-\sigma^2 \|\omega\|^2}\right]$
- $f(\omega) = (2\pi)^{-d} \left[\lambda - \lambda^2 \left(\frac{\pi\sigma^2}{2}\right)^{\frac{d}{2}} e^{\frac{-\sigma^2 \|\omega\|^2}{8}} \right]$

For stationary and isotropic case, $g(\mathbf{x}_i, \mathbf{x}_j) = g(\|\mathbf{x}_i - \mathbf{x}_j\|) = g(r)$.
 Parameter: $\boldsymbol{\theta}_{\text{Thomas}} = (\kappa, \mu, \sigma^2)^\top$, $\boldsymbol{\theta}_{\text{Determinantal}} = (\lambda, \sigma^2)^\top$.

PCF and spectral density

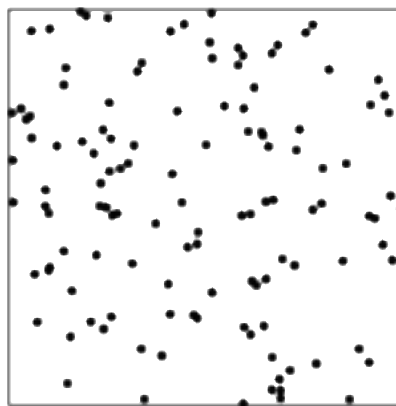
Thomas process

($\kappa = 0.5$, $\sigma = 0.3$, $\mu = 2$)



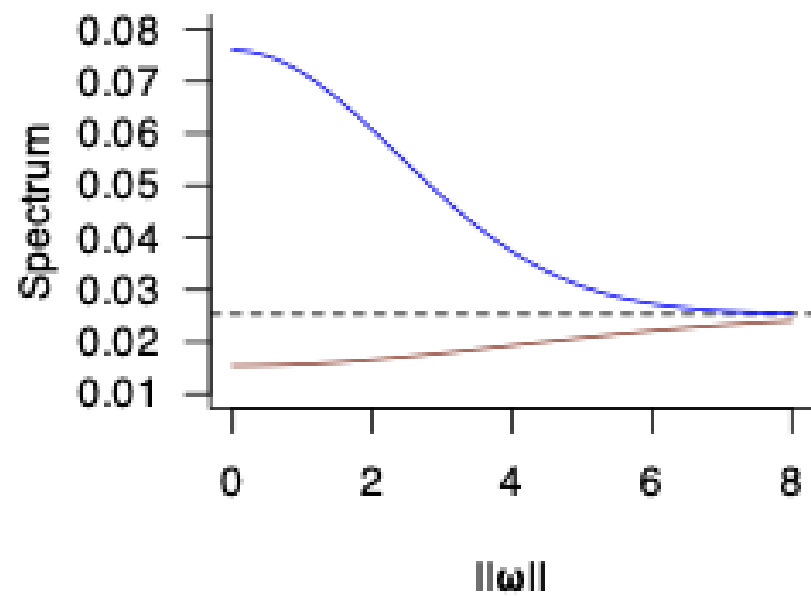
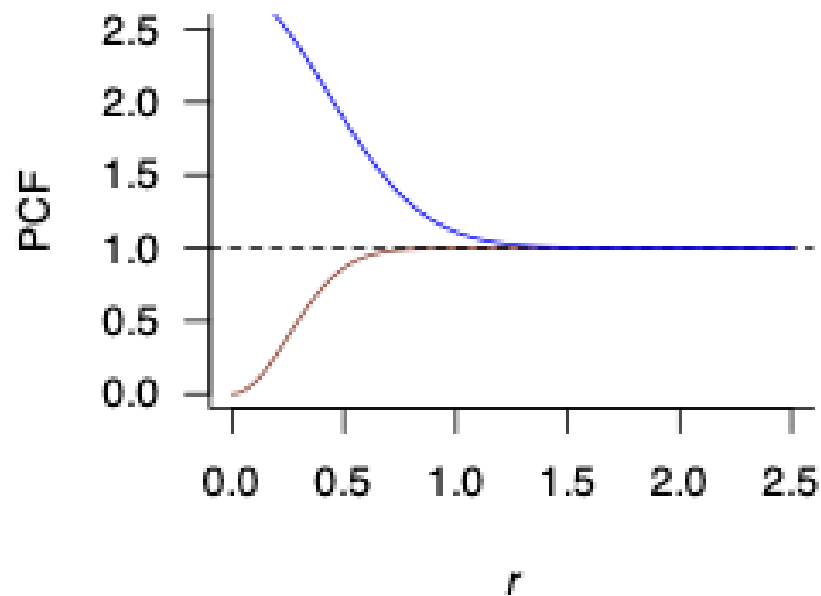
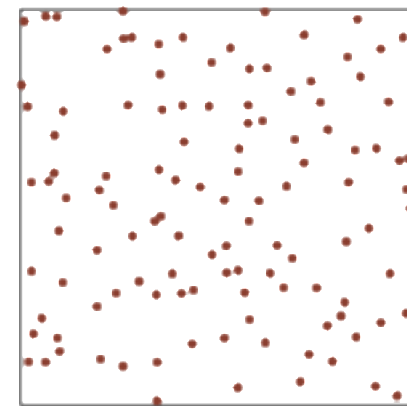
Poisson process

($\lambda = 1$)



DPP with Gaussian kernel

($\lambda = 1$, $\sigma = 0.5$)



Frequency domain parameter estimation

Whittle-type likelihood

Let $\{X_\theta\}$ be a family of 2nd-order stationary point processes with parameter $\theta \in \Theta$. The associated spectral density is denoted as f_θ . Then, we fit the model using the pseudo-likelihood

$$L(\theta) = \sum_{\omega_k \in D} \left\{ \frac{\hat{I}(\omega_k)}{f_\theta(\omega_k)} + \log f_\theta(\omega_k) \right\}.$$

Intuition

- $\hat{I}(\omega)$ can be regarded as the “**truth**” since $\mathbb{E}[\hat{I}_n(\omega)] \rightarrow f(\omega)$.
- $f_\theta(\omega)$ are the parameter family of spectral densities.
- $L(\theta)$ is the **spectral divergence between the truth and our guess. The smaller the better!**

Proposed model parameter estimator

$$\hat{\theta} = \arg \min_{\theta \in \Theta} L(\theta)$$

Frequency domain parameter estimation

| Model | Window | Parameter | Method | |
|-------|------------------------|-------------------|--------------|---------------|
| | | | Ours | ML |
| TCP | [−5, 5] ² | κ | −0.04 (0.11) | −0.02 (0.11) |
| | | μ | 0.72 (3.52) | −0.24 (5.93) |
| | | σ^2 | 0.02 (0.07) | −0.04 (0.22) |
| | | Time (sec) | 0.74 | 0.38 |
| | [−10, 10] ² | κ | −0.02 (0.05) | −0.02 (0.05) |
| | | μ | 0.60 (1.77) | 0.33 (2.79) |
| | | σ^2 | 0.01 (0.02) | −0.01 (0.11) |
| | | Time (sec) | 2.38 | 5.67 |
| | [−20, 20] ² | κ | −0.01 (0.04) | 0.00 (0.03) |
| | | μ | 0.25 (1.03) | 0.15 (1.23) |
| | | σ^2 | 0.01 (0.02) | 0.00 (0.03) |
| | | Time (sec) | 9.15 | 173.66 |

The bias and the standard errors (in parentheses) of the estimated parameters based on two different approaches for the Thomas clustered process (TCP)

Spectral analysis

Time series model

- White noise : $e_t \stackrel{\text{iid}}{\sim} \mathcal{N}(0, \sigma^2)$
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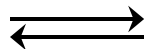
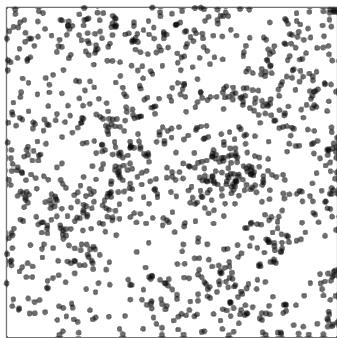
ACF

- $\gamma(k) = \sigma^2 \mathbf{1}\{k = 0\}$
- $\gamma(k) = \theta \sigma^2, k = \pm 1$
- $\gamma(k) = \alpha^{|k|} \sigma_X^2$

Spectrum

- $f(\omega) = \sigma^2 / (2\pi)$
- $f(\omega) = \sigma^2 (1 + \theta^2 + 2\theta \cos \omega) / (2\pi)$
- $f(\omega) = \frac{\sigma^2}{2\pi(1 - 2\alpha \cos \omega + \alpha^2)}$

2nd-order stationary
point process



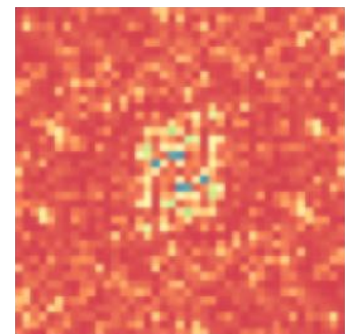
Spectrum
 $f(\omega)$



Its estimator is



Periodogram
 $\hat{l}(\omega)$



Spectral analysis

Time series model

- White noise : $e_t \stackrel{\text{iid}}{\sim} \mathcal{N}(0, \sigma^2)$
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ACF

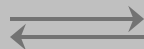
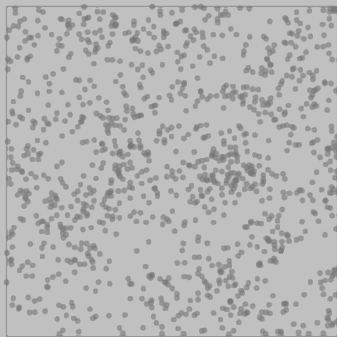
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Spectrum

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- $f(\omega) = \frac{\sigma^2}{2\pi(1 - 2\alpha \cos \omega + \alpha^2)}$

Does NOT work for inhomogeneous point processes!

2nd-order stationary
point process



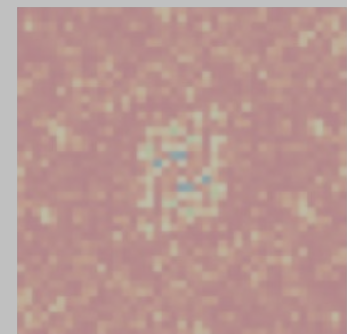
Spectrum
 $f(\omega)$



Its estimator is

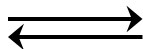
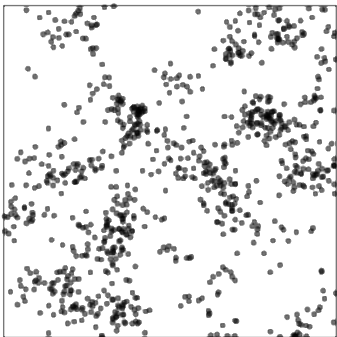


Periodogram
 $\hat{l}(\omega)$

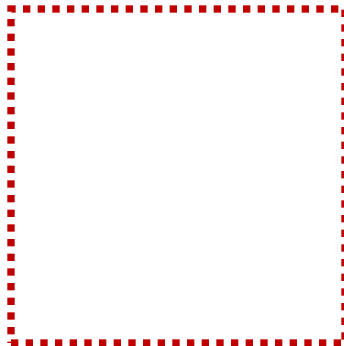


Spectral analysis under inhomogeneity

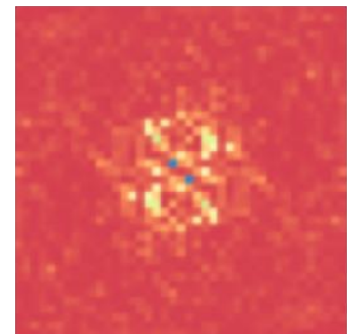
Inhomogeneous
point process



???

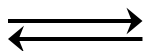
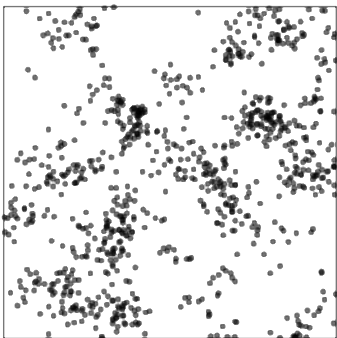


Periodogram
 $\hat{l}(\omega)$

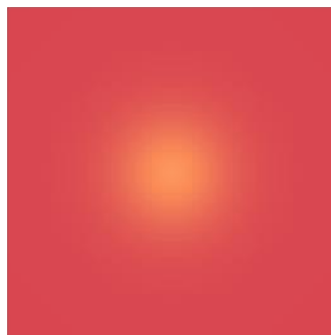


Spectral analysis under inhomogeneity

Inhomogeneous
point process



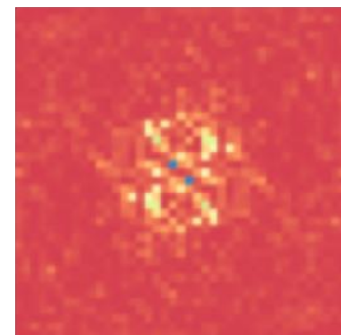
Pseudo-spectrum
 $f_P(\omega)$



Infer



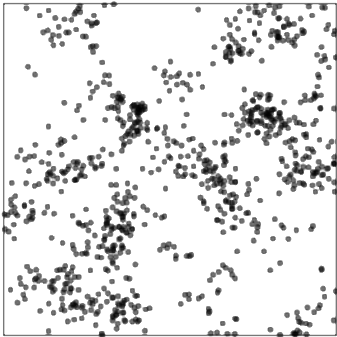
Periodogram
 $\hat{I}(\omega)$



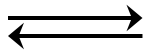
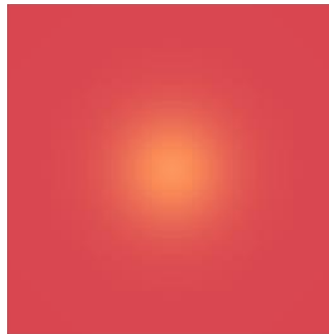
$$f_P(\omega) := \lim_{n \rightarrow \infty} \mathbb{E}[\hat{I}_n(\omega)], \quad \forall \omega \in \mathbb{R}^d.$$

Estimating the pseudo-spectrum

Inhomogeneous
point process

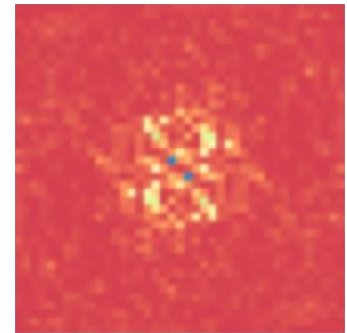


Pseudo-spectrum
 $f_p(\omega)$



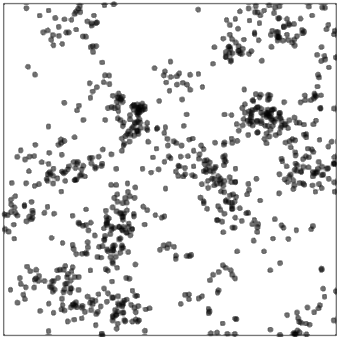
Estimators:

① Periodogram
 $\hat{l}(\omega)$

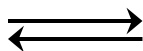
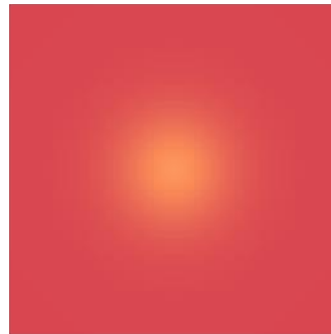


Estimating the pseudo-spectrum

Inhomogeneous
point process

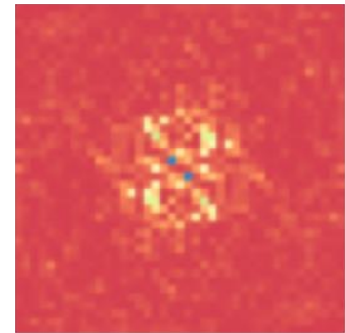


Pseudo-spectrum
 $f_P(\omega)$

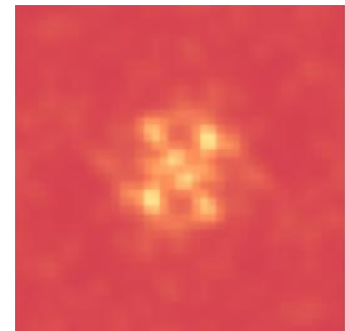


Estimators:

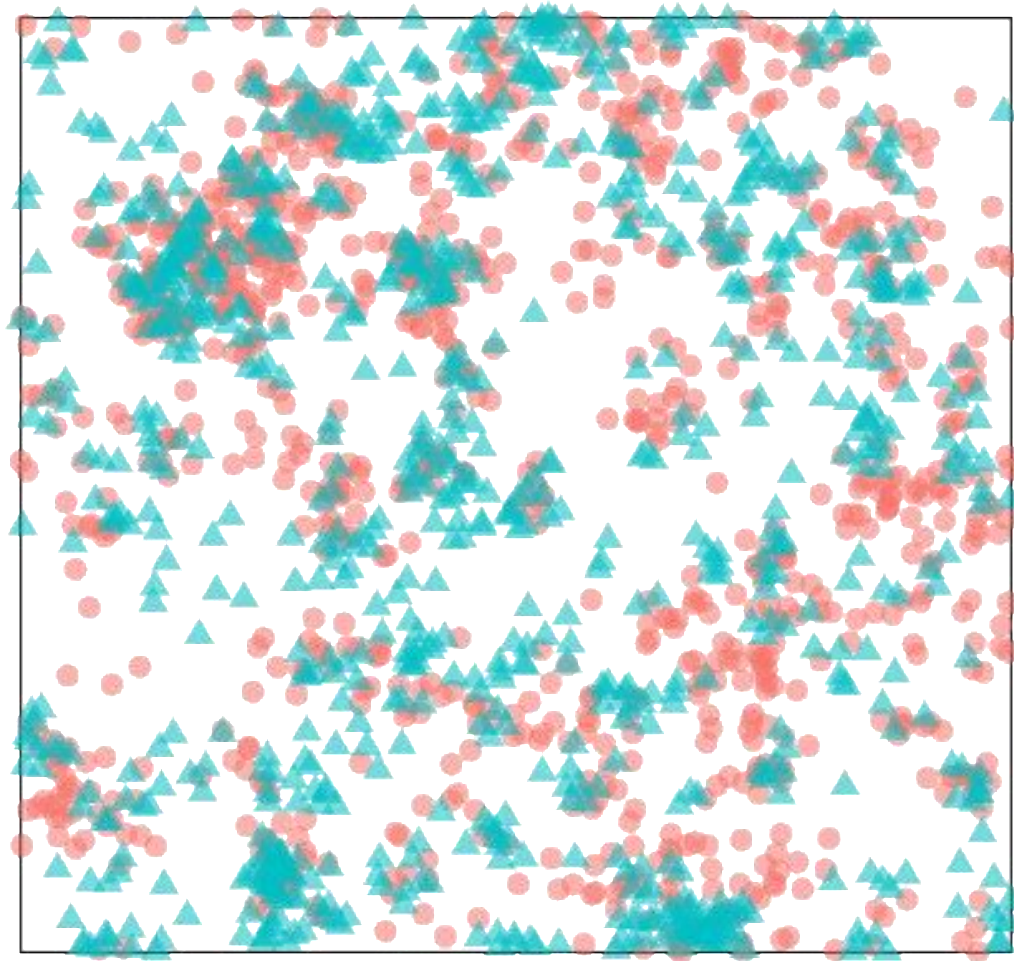
① Periodogram
 $\hat{l}(\omega)$



② Kernel estimator
 $\hat{f}_P(\omega)$



Simulation: Data generating process

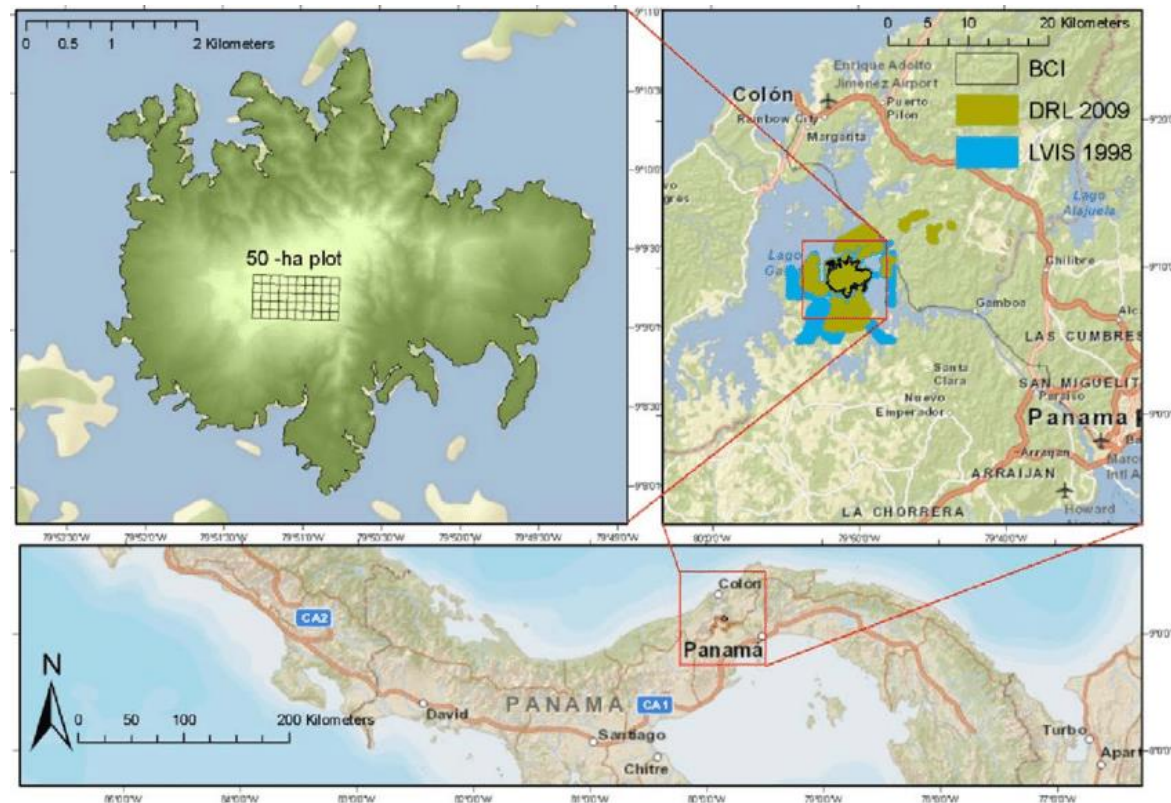


Bivariate point process with marginal and joint clustering interactions

Simulation: Result

| Pseudo-spectrum | Window | $\hat{l}(\omega)$ | | $\hat{f}_P(\omega)$ | |
|-----------------|---------------|-------------------|-------|---------------------|------|
| | | Bias | MSE | Bias | MSE |
| Marginal | $[-5, 5]^2$ | 0.03 | 1.49 | 0.02 | 0.49 |
| | $[-10, 10]^2$ | 0.01 | 1.09 | 0.00 | 0.17 |
| | $[-20, 20]^2$ | 0.00 | 1.02 | 0.00 | 0.13 |
| Cross | $[-5, 5]^2$ | 0.06 | 28.14 | 0.01 | 3.76 |
| | $[-10, 10]^2$ | 0.05 | 22.61 | 0.01 | 2.45 |
| | $[-20, 20]^2$ | 0.05 | 20.18 | 0.01 | 2.11 |

Barro Colorado Island (BCI) data



207,718 alive trees
310 species
7 censuses



Figure 1 in Fricker et al. (2012)

Barro Colorado Island (BCI) data

cappfr



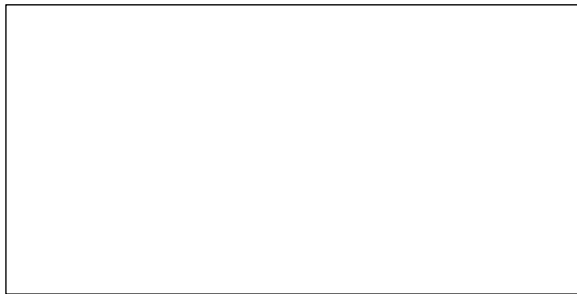
hirttr



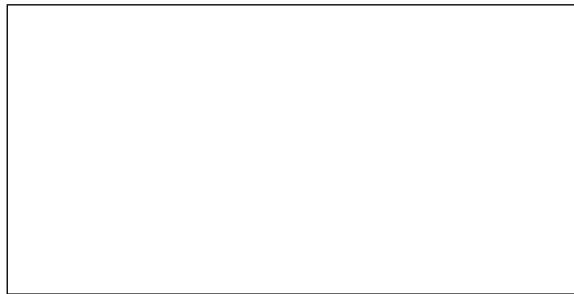
protpa



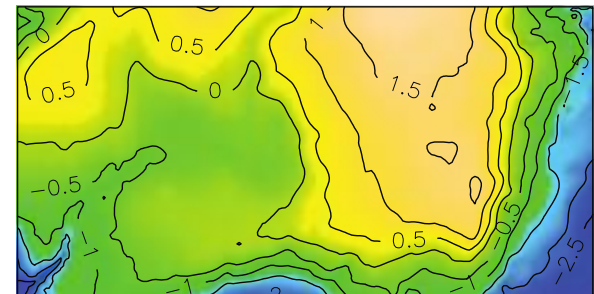
protte



tet2pa



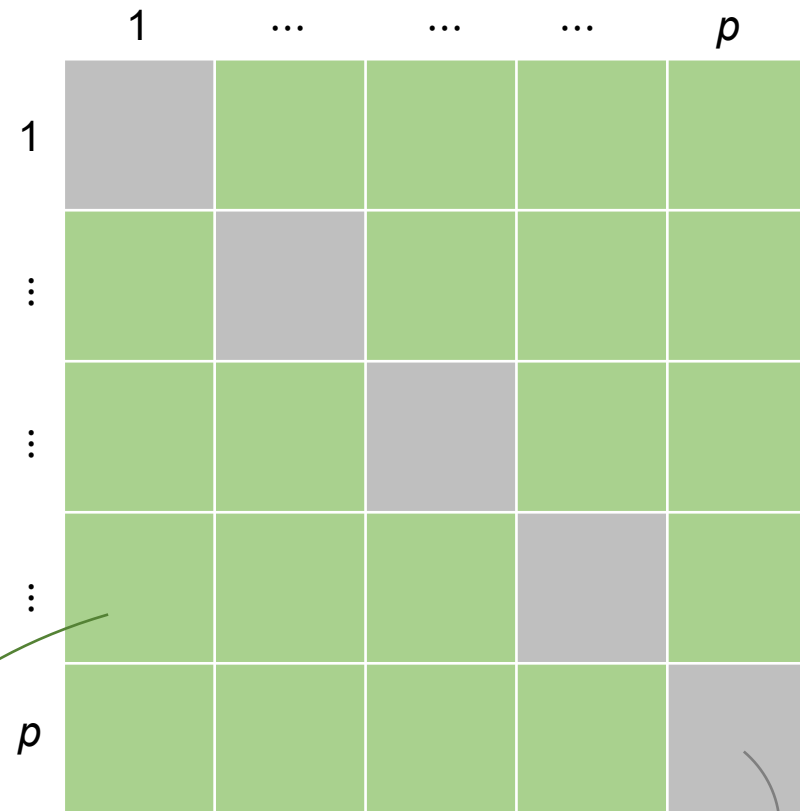
elevation



Point patterns of five species in the BCI dataset and image of elevation in the study region.

Multivariate point pattern

Consider a p -variate process: $i, j \in \mathcal{V} = \{1, 2, \dots, p\}$



Spectral coherence:

$$R_{ij}(\omega) = \frac{f_{ij}(\omega)}{[f_{ii}(\omega)f_{jj}(\omega)]^{\frac{1}{2}}}$$

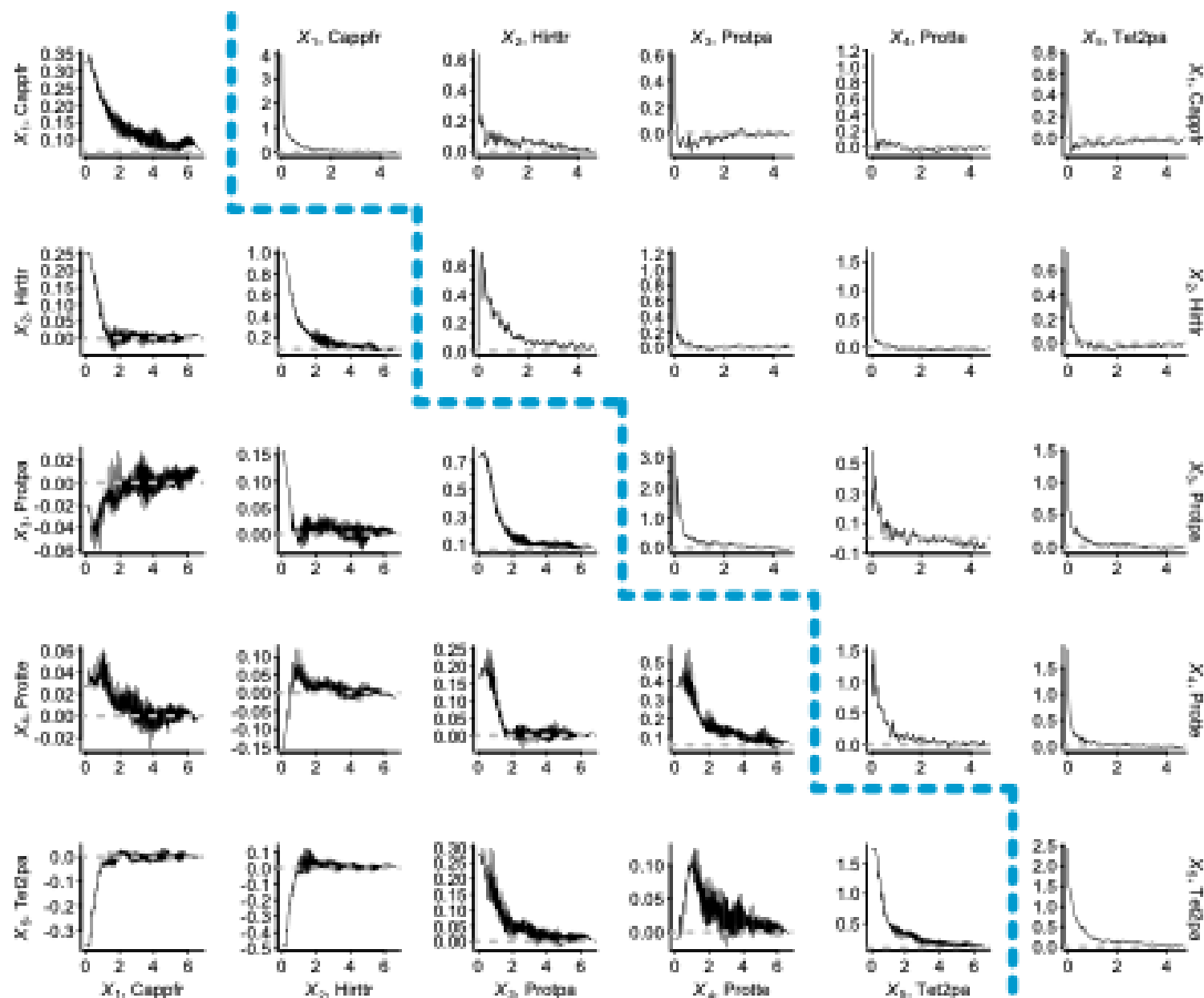
Cross-spectrum:

$$f_{ij}(\omega) = (2\pi)^{-d} \int_{\mathbb{R}^d} C_{ij}(\mathbf{x}_1 - \mathbf{x}_2) e^{-i\mathbf{x}^\top \omega} d\mathbf{x}$$

Marginal spectrum:

$$f_{ii}(\omega) = (2\pi)^{-d} \int_{\mathbb{R}^d} C_{ii}(\mathbf{x}_1 - \mathbf{x}_2) e^{-i\mathbf{x}^\top \omega} d\mathbf{x}$$

Pseudo-spectra



Coherence analysis

