

DATA130004: Homework 1

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2017.10.10

1. Suppose that X and Y are continuous random variables with density f and g respectively, and a and b are constants. Prove the following arguments:

(a) $E(aX + b) = aE(X) + b$

Proof: Since X is a continuous random variable with density $f(x)$, denote $Z = X + c$ which is also a continuous variable with density $f_z(x)$.

$$E(aX + b) = E(Z) = \int_{-\infty}^{\infty} x f_z(x) dx$$

If $a = 0$, then $E(b) = b$ will be undeniable. So assume $a \neq 0$.

$$\begin{aligned} \therefore F_z(x) &= P\{Z \leq x\} = P\{aX + b \leq x\} = P\{aX \leq x - b\} \\ &= \begin{cases} P\{X \leq \frac{x-b}{a}\} (a > 0) \\ P\{X \geq \frac{x-b}{a}\} (a < 0) \end{cases} \\ &= \begin{cases} F_x\left(\frac{x-b}{a}\right) (a > 0) \\ 1 - F_x\left(\frac{x-b}{a}\right) (a < 0) \end{cases} \end{aligned}$$

$$\begin{aligned} \therefore f_z(x) &= \begin{cases} \frac{1}{a} f_x\left(\frac{x-b}{a}\right) (a > 0) \\ -\frac{1}{a} f_x\left(\frac{x-b}{a}\right) (a < 0) \end{cases} \\ &= \frac{1}{|a|} f_x\left(\frac{x-b}{a}\right) (a \neq 0) \end{aligned}$$

$$E(aX + b) = \int_{-\infty}^{\infty} \frac{x}{|a|} f_x\left(\frac{x-b}{a}\right) dx$$

Denote $t = \frac{x-b}{a}$, which means $x = at + b$

$$\begin{aligned} \Rightarrow E(aX + b) &= \int_{-\infty}^{\infty} \frac{at + b}{|a|} f_x(t) a dx \\ &= a \int_{-\infty}^{\infty} t f_x(t) dx + b \int_{-\infty}^{\infty} f_x(t) dx \\ &= aE(X) + b \end{aligned}$$

- (b) $E(X + Y) = E(X) + E(Y)$

Proof: Assume the joint density of X and Y as $\varphi(x, y)$.

$$\begin{aligned}
 E(X + Y) &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (x + y) \varphi(x, y) dx dy \\
 &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x \varphi(x, y) dx dy + \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} y \varphi(x, y) dx dy \\
 &= \int_{-\infty}^{\infty} x dx \int_{-\infty}^{\infty} \varphi(x, y) dy + \int_{-\infty}^{\infty} y dy \int_{-\infty}^{\infty} \varphi(x, y) dx \\
 &= \int_{-\infty}^{\infty} x f(x) dx + \int_{-\infty}^{\infty} y g(y) dy \\
 &= E(X) + E(Y)
 \end{aligned}$$

- (c) if X and Y are independent, then $E(XY) = E(X)E(Y)$

Proof: Assume the joint density of X and Y as $\varphi(x, y)$. And since X and Y are independent continuous variables, $\varphi(x, y) = f(x)g(y)$.

$$\begin{aligned}
 E(XY) &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy \varphi(x, y) dx dy \\
 &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy f(x)g(y) dx dy \\
 &= \int_{-\infty}^{\infty} x f(x) dx \int_{-\infty}^{\infty} y g(y) dy \\
 &= E(X)E(Y)
 \end{aligned}$$

- (d) $Var(a) = 0$

Proof:

$$Var(a) = E(a - E(a))^2 = E(a - a) = 0$$

- (e) $Var(aX + b) = a^2 Var(X)$

Proof:

$$\begin{aligned}
 Var(aX + b) &= E[aX + b - E(aX + b)]^2 = E[aX + b - aE(X) - b]^2 \\
 &= E[a(X - E(X))]^2 = a^2 E(X - E(X))^2 \\
 &= a^2 Var(X)
 \end{aligned}$$

- (f) $Var(X + Y) = Var(X) + Var(Y) + 2Cov(X, Y)$

Proof:

$$\begin{aligned}
 Var(X + Y) &= E[X + Y - E(X + Y)]^2 = E[X - E(X) + Y - E(Y)]^2 \\
 &= E[(X - E(X))^2 + (Y - E(Y))^2 + 2(X - E(X))(Y - E(Y))] \\
 &= E[X - E(X)]^2 + E[Y - E(Y)]^2 + 2E[(X - E(X))(Y - E(Y))] \\
 &= Var(X) + Var(Y) + 2Cov(X, Y)
 \end{aligned}$$

- (g) if X and Y are independent, then $Var(X + Y) = Var(X) + Var(Y)$

Proof: If X and Y are independent, then according to argument (c), $E(XY) = E(X)E(Y)$.

Therefore,

$$Cov(X, Y) = E[(X - E(X))(Y - E(Y))] = E(XY) - E(X)E(Y) = 0$$

$$\Rightarrow \text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y)$$

2. Rizzo book Exercises 3 (starting from Page 94): 3, 5.

(a) The Pareto(a, b) distribution has cdf

$$F(x) = 1 - \left(\frac{b}{x}\right)^a, \quad x \geq b > 0, a > 0$$

Derive the probability inverse transformation $F^{-1}(U)$ and use the inverse transform method to simulate a random sample from the Pareto(2, 2) distribution. Graph the density histogram of the sample with the Pareto(2, 2) density superimposed for comparison.

$$F^{-1}(U) = \frac{b}{(1-U)^{\frac{1}{a}}}$$

$$\Rightarrow F_{(2,2)}^{-1}(U) = \frac{2}{\sqrt{1-U}}$$

To simulate a random sample from the Pareto(2,2) distribution using inverse transform method, the R script are as follows:

```
n <- 10000
u <- runif(n)
x <- 2/((1-u)^(1/2))
hist(x, prob = TRUE, main = expression(f(x) == (4*x-8)/(x^3)))
y <- seq(0, 120, 0.01)
lines(y, (4*y-8)/(y^3))
```

By running the script, I got Figure 1.

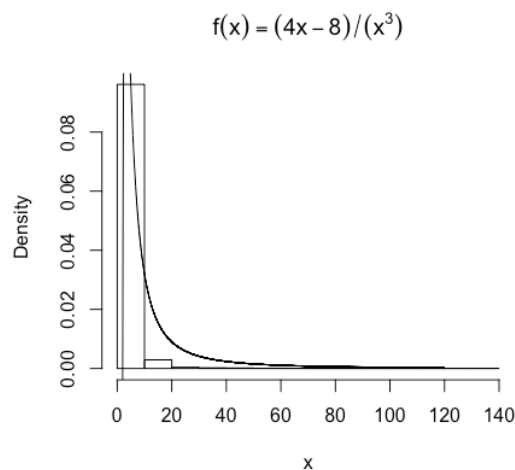


Figure 1: density histogram of sample and density function

(b) A discrete random variable X has probability mass function

x	0	1	2	3	4
$p(x)$	0.1	0.2	0.2	0.2	0.3

Use the inverse transform method to generate a random sample of size 1000 from the distribution of X . Construct a relative frequency table and compare the empirical with the theoretical probabilities. Repeat using the R sample function.

$$x = F_X^{-1}(u) = \begin{cases} 0, & u \leq 0.1 \\ 1, & 0.1 < u \leq 0.3 \\ 2, & 0.3 < u \leq 0.5 \\ 3, & 0.5 < u \leq 0.7 \\ 4, & 0.7 < u \leq 1 \end{cases}$$

Here is the frequency table of the sample (sample size = 1000):

x	0	1	2	3	4
$p(x)$	0.103	0.222	0.198	0.185	0.292

Repeat using the R sample function:

```
x <- sample(c(0,1,2,3,4),1000,replace = TRUE,prob = c(0.1,0.2,0.2,0.2,0.3))
```

Here's the frequency table of the sample:

x	0	1	2	3	4
$p(x)$	0.093	0.208	0.198	0.197	0.304

Comparing the empirical with the theoretical probabilities, the two sample distributions I generated both quite close to the theoretical one, and one using the R sample function seems closer to the theoretical distribution. But more experiments needs to be done to examine this argument.