Semiparametric Least Squares Inference for Causal Effects with R

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Abstract

This vignette explains how to use the causalSLSE package to estimate causal effects using the semiparametric least squares methods developed by Giurcanu et al. (2023). We describe the classes and methods implemented in the package as well as how they can be used to analyze synthetic and real data.

1 Introduction

This document presents the causalSLSE package describing the functions implemented in the package. It is intended for users interested in the details about the methods presented in Giurcanu et al. (2023) and how they are implemented. We first present the theory and will present the package in the following sections.

The general semiparametric additive regression model is

$$Y = \beta_0(1-Z) + \beta_1 Z + \sum_{l=1}^q f_{l,0}(X_l)(1-Z) + \sum_{l=1}^q f_{l,1}(X_l)Z + \xi$$

$$\equiv \beta_0(1-Z) + \beta_1 Z + f_0(X)(1-Z) + f_1(X)Z + \xi,$$

where $Y \in \mathbb{R}$ is the response variable, Z is the treatment indicator defined as Z = 1 for the treated and Z = 0 for the nontreated, and $X \in \mathbb{R}^q$ is a q-vector of confounders. We approximate this model by the following regression model:

$$Y = \beta_0(1-Z) + \beta_1 Z + \sum_{l=1}^q \psi_{l,0}^T U_{l,0}(1-Z) + \sum_{l=1}^q \psi_{l,1}^T U_{l,1} Z + \zeta$$

$$\equiv \beta_0(1-Z) + \beta_1 Z + \psi_0^T U_0(1-Z) + \psi_1^T U_1 Z + \zeta,$$

where $U_{l,k} = u_{l,k}(X_l) = (u_{j,l,k}(X_l) : 1 \le j \le p_{l,k}) \in \mathbb{R}^{p_{l,k}}$ is a vector of basis functions corresponding to the l^{th} nonparametric component of the k^{th} group $f_{l,k}(X_l)$, $\psi_{l,k} \in \mathbb{R}^{p_{l,k}}$ is an unknown vector of regression coefficients, $U_k = u_k(X) = (u_{l,k}(X_l) : 1 \le l \le q) \in \mathbb{R}^{p_k}$ and $\psi_k = (\psi_{l,k} : 1 \le l \le q) \in \mathbb{R}^{p_k}$, with $p_k = \sum_{l=1}^q p_{l,k}$. In this paper, we propose a data-driven method for selecting the vectors of basis functions $u_0(X)$ and $u_1(X)$. Note that we allow the number of basis functions $(p_{l,k})$ to differ across confounders and groups.

Let the following be the regression model estimated by least squares:

$$Y_i = \beta_0(1 - Z_i) + \beta_1 Z_i + \psi_0^T U_{i,0}(1 - Z_i) + \psi_1^T U_{i,1} Z_i + \zeta_i \text{ for } i = 1, ..., n,$$

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and $\hat{\beta}_0$, $\hat{\beta}_1$, $\hat{\psi}_0$ and $\hat{\psi}_1$ be the least squares estimates of the regression parameters. Then, the semiparametric least squares estimators (SLSE) of the average causal effect (ACE), causal effect on the treated (ACT) and causal effect on the non-treated (ACN) are defined respectively as follows:

ACE =
$$\hat{\beta}_1 - \hat{\beta}_0 + \hat{\psi}_1^T \bar{U}_1 - \hat{\psi}_0^T \bar{U}_0$$

ACT = $\hat{\beta}_1 - \hat{\beta}_0 + \hat{\psi}_1^T \bar{U}_{1,1} - \hat{\psi}_0^T \bar{U}_{0,1}$
ACN = $\hat{\beta}_1 - \hat{\beta}_0 + \hat{\psi}_1^T \bar{U}_{1,0} - \hat{\psi}_0^T \bar{U}_{0,0}$,

where $\bar{U}_k = \frac{1}{n} \sum_{i=1}^n U_{i,k}$, $\bar{U}_{k,1} = \frac{1}{n_1} \sum_{i=1}^n U_{i,k} Z_i$, $\bar{U}_{k,0} = \frac{1}{n_0} \sum_{i=1}^n U_{i,k} (1 - Z_i)$, for k = 0, 1, and n_0 and n_1 are the sample sizes of the nontreated and treated groups respectively. As shown by Giurcanu et al. (2023), under some regularity conditions these estimators are consistent and asymptotically normal.

To derive the variance of these causal effect estimators, note that they can be expressed as a linear combination of the vector of least squares estimates. Let $\hat{\theta} = \{\hat{\beta}_0, \hat{\beta}_1, \hat{\psi}_0^T, \hat{\psi}_1^T\}^T$. Then, the causal effect estimates can be written as $\hat{D}_c^T \hat{\theta}$ for c=ACE, ACT or ACN, with $\hat{D}_{ACE} = \{-1, 1, -\bar{U}_0^T, \bar{U}_1^T\}^T$, $\hat{D}_{ACT} = \{-1, 1, -\bar{U}_{0,1}^T, \bar{U}_{1,1}^T\}^T$ and $\hat{D}_{ACN} = \{-1, 1, -\bar{U}_{0,0}^T, \bar{U}_{1,0}^T\}^T$. Since \hat{D}_c is random, we need a first order Taylor expansion to derive the variance of the estimators. Assuming that the data set is iid and using the asymptotic properties of least squares, we can show that the variance of ACE= $\hat{D}_{ACE}^T \hat{\theta}$ can be consistently estimated as follows (we can derive a similar expression for the ACT and ACN):

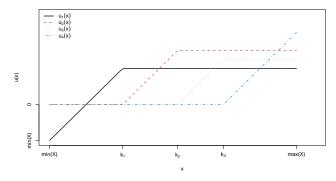
$$\hat{V}_{\text{ACE}} = \begin{pmatrix} -\hat{\beta}_0 & \hat{\beta}_1 & \hat{D}_{\text{ACE}}^T \end{pmatrix} \begin{pmatrix} \hat{\Sigma}_0 & \hat{\Sigma}_{0,1} & \hat{\Sigma}_{0,\hat{\theta}} \\ \hat{\Sigma}_{1,0} & \hat{\Sigma}_1 & \hat{\Sigma}_{1,\hat{\theta}} \\ \hat{\Sigma}_{\hat{\theta},0} & \hat{\Sigma}_{\hat{\theta},1} & \hat{\Sigma}_{\hat{\theta}} \end{pmatrix} \begin{pmatrix} -\hat{\beta}_0 \\ \hat{\beta}_1 \\ \hat{D}_{\text{ACE}} \end{pmatrix},$$

where $\hat{\Sigma}_k = \widehat{\text{var}(\bar{U}_k)}$, $\hat{\Sigma}_{k,l} = \widehat{\text{cov}(\bar{U}_k,\bar{U}_l)}$, $\hat{\Sigma}_{k,\hat{\theta}} = \hat{\Sigma}_{\hat{\theta},k}^T = \widehat{\text{cov}(\bar{U}_k,\hat{\theta})}$, for k,l=0,1, and $\hat{\Sigma}_{\hat{\theta}}$ is a consistent estimator of variance of $\hat{\theta}$. We will discuss the choice of the covariance matrix estimator $\hat{\Sigma}_{\hat{\theta}}$ in the next section.

To understand the package, it is important to know how the $u_{l,k}(X_l)$'s are defined. For clarity, let's write $U_{l,k} = u_{l,k}(X_l)$ as $U = u(X) = (u_j(X) : 1 \le j \le p) \in \mathbb{R}^p$. We just need to keep in mind that it is different for the treated and nontreated groups and also for different confounders. We describe here how to construct the local linear splines for a given confounder X in a given group. To this end, let $\{\kappa_1, \ldots, \kappa_{p-1}\}$ be a set of p-1 knots strictly inside the support of X satisfying $\kappa_1 < \kappa_2 < \ldots < \kappa_{p-1}$. In the case of local linear splines described in the paper, we have:

$$\begin{array}{lcl} u_{1}(x) & = & xI(x \leq \kappa_{1}) + \kappa_{1}I(x > \kappa_{1}) \\ u_{j}(x) & = & (x - \kappa_{j-1})I(\kappa_{j-1} \leq x \leq \kappa_{j}) + (\kappa_{j} - \kappa_{j-1})I(x > \kappa_{j}) \,, \quad 2 \leq j \leq p-1 \\ u_{p}(x) & = & (x - \kappa_{p-1})I(x > \kappa_{p-1}) \end{array}$$

Therefore, if the number of knots is equal to 1, we only have two local linear splines. Since the knots must be strictly inside the support of X, for any categorical variable with two levels, the number of knots must be equal to zero. In this case, u(x) = x. For general ordinal variables, the number of knots cannot exceed the number of levels minus two. The following illustrates local spline functions when the number of knots is equal to 3:



Note that for the sample regression, the knots of X_l for group k, l=1,...,q, must be strictly inside the sample range of $(X_{i,l}: 1 \le i \le n, Z_i = k) \in \mathbb{R}^{n_k}$, where n_k is the sample size in group k, instead of inside the support of X_l .

2 The causalSLSE package

2.1 The starting knots

By default, the knots for each variable are based on the following procedure applied separately for the treated and nontreated. The term sample size refers either to the number of observations in the treated or nontreated group.

- 1. The starting number of knots is a function of the sample size and is determined by the argument nbasis, a function of one argument, the sample size. The floor value of what the function returns is the number of basis functions. The starting number of knots is therefore equal to the floor of what the function returns minus 1 (or 0 if the function returns a value strictly less than 1). The default function is function(n) n^0.3. For example, if the total sample size is 500, with 200 treated and 300 nontreated, the starting number of knots in the treated and nontreated groups are respectively equal to 3=floor(200^0.3)-1 and 4=floor(300^0.3)-1, respectively. It is possible to have a number of knots that does not depend on the sample size. All we need is to set the argument nbasis to a function that returns an integer, e.g., nbasis=function(n) 4 for 4 basis functions or 3 knots.
- 2. Let (p-1) be the number of knots determined in the previous step. The default knots are obtained by computing p+1 quantiles of X for equally spaced probabilities from 0 to 1, and by dropping the first and last quantiles. For example, if the number of knots is 3, then the initial knots are given by quantiles for the probabilities 0.25, 0.5 and 0.75.
- 3. We drop any duplicated knots and any knots equal to either the max or the min of X. If the resulting number of knots is equal to 0, the vector of knots is set to NULL. When the knots is equal to NULL for a variable X, it means that u(x) = x.

The last step implies that the number of knots for all categorical variables with two levels is equal to 0. For nominal variables with a small number of levels, the number of knots, a subset of the levels, may be smaller than the ones defined by nbasis. For example, when the number of levels for a nominal variable is 3, the number of knots cannot exceed 1.

To illustrate how to use the package, we are using the dataset from Lalonde (1986). The dataset, called nsw, contains some continuous and categorical variables, so we can illustrate how knots are selected initially. The dataset is included in the causalSLSE package.

library(causalSLSE)
data(nsw)

The outcome is the real income in 1978 (re78) and the purpose is to estimate the causal effect of a training program (treat) on re78. The dataset includes the continuous variables age (age), education (ed), the 1975 real income (re75), and binary variables (black, hisp, married and nodeg).

The process of selecting knots is the same for the treated and nontreated. In fact, the knots are created by calling the slseKnots function for each group. This function is not restricted to causal models. We could use it to generate knots for any SLSE regression model. But, estimating such models is not yet implemented in the package. The arguments of the function are:

- form: A formula with the right-hand side being the list of confounders. If a left-hand side is provided, the slseKnots function will ignore it.
- data: A data.frame containing all variables included in the formula.
- X: Instead of providing the function with a formula, we can input directly the matrix of confounders.
- nbasis: A function that determines the number of basis functions as explained above.
- **knots**: This argument is used to set the knots manually. We will explain how to use this argument in the next section.

We start by considering the confounders age, re75, ed, and married. To create a set of knots for a specific group using slseKnots, we can restrict the sample using the subset function. For example, we can create the initial knots for the treated and nontreated as follows:

The function returns an object of class slseKnots and its print method produces a nice display separating confounders with and without knots. For example, the following are the starting knots for the treated:

k1

The sample size for the treated is 297. Given the default nbasis, it implies a number of starting knots equal to 4=floor(297^{0.3})-1. This is the number of knots we have for ed and age. However, the number of knots for re75 is 3. The reason is that re75 contains a large fraction of zeros. Since the 20th percentile is equal to 0 and 0 is also the minimum value of ed75, it is dropped. This can be seen as follows (the type argument of the quantile function is the same as it is implemented in the package):

```
quantile(nsw[nsw$treat==1,'re75'], c(.2,.4,.6,.8), type=1)

20%      40%      60%      80%
0.0000      357.9499      1961.8640      5588.6640
```

Note that each object in the package is S3-class, so the elements can be accessed using the operator \$. For example, we can extract the knots for age as follows:

```
k1$age

20% 40% 60% 80%

19 22 25 28
```

Note that the confounders are listed in the "no knots" section when their values are set to NULL. In the above example, it is the case of married because it is a binary variable. As we can see its list of knots it set to NULL:

k1\$married

NULL

Alternatively, we can create a list of slseKnots objects, one for the treated and one for the nontreated, using the cslseKnots function. This approach of generating the knots for each group will allow us to easily extend the package to models with multiple treatments. The function has the same arguments as slseKnots, but form must include a formula linking the outcome and the treatment indication and a formula listing the confounders, separated by the operator |. In the following example, we see the formula linking the outcome re78 and the treatment indicator treat, and the list of confounders, which are the same as in the previous example::

The function returns an object of class cslseKnots. We can use its print method to print the knots, in which case the knots are printed for both groups:

treated Covariates with no knots: married Covariates with knots: 20% 40% 60% 80% Knots 19 22 25 28 40% 60% 80% Knots 357.9 1962 5589 20% 40% 60% 80% 9 10 11 12 Knots nontreated Covariates with no knots: married Covariates with knots: 16.67% 33.33% 50% 66.67% 83.33% Knots 18 20 23 26 re75 : 50% 66.67% 83.33% Knots 823.3 2292 16.67% 33.33% 66.67% 83.33% 9 10 11 12 Knots

Note that cslseKnots objects are lists of slseKnots objects. Therefore, we can print the set of knots for a specific group using \$ followed by its label (treated or nontreated). For example, we can print the set of knots for the treated by typing k\$treated.

2.2 Creating a causal SLSE model

In general, users will not call cslseKnots explicitly. This function is called by the cslseModel function. The function returns an object of class cslseModel (or a causal SLSE model), which contains the information

about the outcome (Y), the treatment indicator (Z), the confounders (X) and their knots $(\kappa_{l,k})$. The arguments of the function are the same as for the cslseKnots function. Therefore, we can create a model with the same set of knots as in the previous section as follows:

```
model1 <- cslseModel(re78 ~ treat | ~ age + re75 + ed + married, data = nsw)
```

Its print method summarizes the characteristics of the model. In particular, it indicates which confounders have a positive number of knots and which ones have no knots.

model1

Number of treated: 297
Number of nontreated: 425
Selection method: Default
Confounders approximated by SLSE:
 treated: age, re75, ed
 nontreated: age, re75, ed
Confounders not approximated by SLSE:
 treated: married
 nontreated: married

Note that the selection method is set to Default. We refer to this method when the knots are automatically selected by the method described in the previous section. Later in the document, we will present methods for optimally selecting a subset of Default.

The element knots of model1 is the object of class cslseKnots created by the cslseKnots function, so we can print the knots as follows:

model1\$knots

ed:

```
treated
Covariates with no knots:
   married
Covariates with knots:
     20% 40% 60% 80%
Knots 19 22 25 28
       40% 60% 80%
Knots 357.9 1962 5589
ed:
     20% 40% 60% 80%
Knots 9 10 11 12
nontreated
Covariates with no knots:
   married
Covariates with knots:
     16.67% 33.33% 50% 66.67% 83.33%
Knots
         18
                20 23
                           26
       50% 66.67% 83.33%
Knots 823.3 2292
```

```
16.67% 33.33% 66.67% 83.33% Knots 9 10 11 12
```

Alternatively, we can print the knots by setting the print method argument which to "selKnots". Also, all print methods offers the possibility of changing the number of digits. For example, the following would print the knots using 4 digits:

```
print(model1, which="selKnots", digits=4)
```

We have also included, in the package, the simulated dataset simDat4, which contains special types of confounders. It helps to further illustrate how the knots are determined. The dataset contains a continuous variable X1 with a large proportion of zeros, the categorical variable X2 with 3 levels, an ordinal variable X3 with 3 levels, and a binary variable X4. The levels for X2 are {"first", "second", "third"} and for X3 the levels are {1,2,3}.

Character-type variables are automatically converted into factors. It is also possible to define a numerical variable like X3 as a factor by using the function as.factor in the formula. We see that the 2 binary variables X2second and X2third are created and X2first is omitted to avoid multicollinearity. For the binary variable X4, the number of knots is set to 0, and for the ordinal variable X3, the number of knots is set to 1 because the min and max values 1 and 3 cannot be selected.

2.2.1 Setting the knots manually

The user has control over the selection of knots through the argument knots. When the argument is missing (the default), all knots for both groups are set automatically as described above. One way to set the number of knots to 0 for all variables in both groups is to set the argument to NULL.

Number of treated: 297

Number of nontreated: 425

Selection method: User Based

Confounders approximated by SLSE:
 treated: None
 nontreated: None

Confounders not approximated by SLSE:
 treated: age, re75, ed, married
 nontreated: age, re75, ed, married

Causal Semiparametric LSE Model

If we want to set the number of knots to 0 for one group, we must specify the group using a named list. In the following, it is set to 0 for the treated only:

Causal Semiparametric LSE Model

Covariates with knots:

```
Number of treated: 297
Number of nontreated: 425
Selection method: User Based
Confounders approximated by SLSE:
    treated: None
    nontreated: age, re75, ed
Confounders not approximated by SLSE:
    treated: age, re75, ed, married
    nontreated: married
```

Notice that the selection method is defined as "User Based" whenever the knots are provided manually by the user. The other option is to provide a list of knots. For each variable, we have three options:

- NA: The knots are set automatically for this variable only.
- NULL: The number of knots is set to 0 for this variable only.
- A numeric vector: The vector cannot contain missing or duplicated values and must be strictly inside the range of the variable for the group.

In the following, we describe all possible formats for the list of knots.

1. An unnamed list of length equal to the number of confounders. In that case, the knots must be defined in the same order of confounders implied by the formula.

Suppose we want to set for the nontreated group an automatic selection for age, no knots for ed, the knots {1000, 5000, 10000} for re75, and the knots to be automatically selected for the treated group. We proceed as follows. Note that setting the value to NA or NULL has the same effect for the binary variable married.

```
selK <- list(nontreated=list(NA, c(1000,5000,10000), NULL, NA))</pre>
model <- cslseModel(re78 ~ treat | ~ age + re75 + ed + married, data = nsw,
                  knots = selK)
print(model, which = "selKnots")
Causal Semiparametric LSE Model: Selected knots
*************
treated
*****
Selection method: Default
Covariates with no knots:
   married
Covariates with knots:
     20% 40% 60% 80%
Knots 19 22 25 28
re75 :
       40% 60% 80%
Knots 357.9 1962 5589
     20% 40% 60% 80%
Knots
      9 10 11 12
nontreated
******
Selection method: User Based
Covariates with no knots:
   ed, married
```

Note that the selection method is set to Default for the treated and User Based for the nontreated. This is because the knots were manually selected only for the nontreated. However, when we print the model, the selection method will be set to User Based whenever at least one of the sets of knots is User Based as in the following:

model

```
************************

Number of treated: 297

Number of nontreated: 425

Selection method: User Based

Confounders approximated by SLSE:
    treated: age, re75, ed
    nontreated: age, re75

Confounders not approximated by SLSE:
    treated: married
    nontreated: ed, married
```

Causal Semiparametric LSE Model

2. A named list of length equal to the number of confounders. In that case, the order of the list of variables does not matter. The cslseModel function will automatically reorder the variables to match the order implied by the formula. The names must match perfectly the confounder names generated by R.

In the following example, we want to add the interaction between ed and age. We want the same set of knots as in the previous example and no knots for the interaction term. The name of the interaction depends on how we enter it in the formula. For example, it is "age:ed" if we enter age*ed in the formula and "ed:age" if we enter ed*age. For factors, the names depend on which binary variable is omitted. Using the above example with the simDat4 model, if we interact X2 and X4 by adding X2*X4 to the formula, the names of the interaction terms are "X2second:X4" and "X2third:X4". When we are uncertain about the names, we can print the knots of a model with the default sets of knots. In the following, we change the order of variables to show that the order does not matter.

```
selK <- list(nontreated=list(married = NA, ed = NULL, 'age:ed' = NULL,</pre>
                              re75 = c(1000,5000,10000), age = NA))
model <- cslseModel(re78 ~ treat | ~ age * ed + re75 + married, data = nsw,
                  knots = selK)
model$knots$nontreated
Covariates with no knots:
    ed, married, age:ed
Covariates with knots:
      16.67% 33.33% 50% 66.67% 83.33%
Knots
          18
                 20 23
                            26
re75 :
        k1
            k2
                   k3
Knots 1000 5000 10000
```

3. A named list of length strictly less than the number of confounders. The names of the selected confounders must match perfectly the names generated by R and the order does not matter. This is particularly useful when the number of confounders is large.

If we consider the previous example, the knots are set manually only for ed, ed:age and re75. By default, all names not included in the list of knots are set to NA. Therefore, we can create the same model from the previous example as follows:

```
selK <- list(nontreated=list(ed = NULL, 'age:ed' = NULL,</pre>
                             re75 = c(1000,5000,10000))
model <- cslseModel(re78 ~ treat | ~ age * ed + re75 + married, data = nsw,
                    knots = selK)
model $knots $nontreated
Covariates with no knots:
    ed, married, age:ed
Covariates with knots:
      16.67% 33.33% 50% 66.67% 83.33%
Knots
         18
                 20 23
                            26
re75 :
        k1 k2
                   k3
Knots 1000 5000 10000
```

Note that the previous case offers an easy way of setting the number of knots to 0 for a subset of confounders. For example, suppose we want to add more interaction terms and set the knots to 0 for all of them. We can proceed as follows.

```
Number of treated: 297
Number of nontreated: 425
Selection method: User Based
Confounders approximated by SLSE:
    treated: age, ed, re75
    nontreated: age, ed, re75
Confounders not approximated by SLSE:
    treated: married, age:ed, ed:re75, ed:married
```

nontreated: married, age:ed, ed:re75, ed:married

Note also that cslseModel deals with interaction terms as any other variable. For example, ed:black is like a continuous variable with a large proportion of zeros. The following shows the default selected knots for ed:black.

```
model <- cslseModel(re78 ~ treat | ~ age + ed * black, data = nsw)
model$knots$nontreated[["ed:black"]]

33.33% 50% 66.67%
9 10 11</pre>
```

2.3 Estimating the model

Given the set of knots from the model object, the estimation is just a least squares method applied to the extended set of confounders defined as the local linear splines corresponding to the set of knots. The regression model is given by:

$$Y = \beta_0(1 - Z) + \beta_1 Z + \psi_0^T U_0(1 - Z) + \psi_1^T U_1 Z + \zeta,$$

where $U_0 = u_0(X)$ and $U_1 = u_1(X)$ are defined above (which depend on the knots of the model). The method that estimates the model is **estSLSE** which has three arguments, but two of them are mainly used internally by other functions. We present them in case they are needed. The arguments are:

• model: A model created by the function cslseModel.

• selKnots: As for the knots argument of cslseModel, it is a list of one or two elements, one for each group. Each element is a list of integers to select knots for the associated group. For example, suppose we have 2 confounders with 5 knots each. If we want to estimate the model with only the first knot for the first confounder and knots 3 and 5 for the second for the treated and all knots for the nontreated, we set selKnots to list(treated=list(1L,c(3L, 5L))). By default it is missing and all the knots from the model are used.

We illustrate the use of estSLSE with a simple model containing 2 confounders and one knot per variable.

```
model <- cslseModel(re78 ~ treat | ~ age + married, data = nsw,</pre>
                  nbasis = function(n) 2)
model$knots
treated
Covariates with no knots:
   married
Covariates with knots:
     50%
Knots 23
nontreated
******
Covariates with no knots:
   married
Covariates with knots:
     50%
Knots 23
fit <- estSLSE(model)</pre>
fit
Semiparametric LSE
*******
Selection method: Default
factor(treat)0 factor(treat)1
                                       Xf0age_1
                                                       XfOage_2
                                                                     Xf0married
       4558.28
                       3754.98
                                         27.80
                                                         -12.51
                                                                        -115.82
      Xf1age_1
                                    Xf1married
                      Xf1age_2
         89.25
                         22.22
                                       1435.28
```

The object of class slseFit returned by estSLSE has its own print method that returns the coefficient estimates. A more detailed presentation of the results can be obtained using the summary method. The following is an example with the previous model.

```
summary(fit)
```

```
Semiparametric LSE
*******
Selection method: Default
              Estimate Std. Error t value Pr(>|t|)
factor(treat)0 4558.28
                         2739.40
                                  1.664
                                          0.0961
factor(treat)1 3754.98
                         3704.37
                                  1.014
                                          0.3107
                                  0.203
Xf0age_1
                27.80
                          136.61
                                          0.8387
Xf0age_2
               -12.51
                           56.06 -0.223
                                          0.8234
Xf0married
               -115.82
                          795.35 -0.146
                                          0.8842
Xf1age_1
                 89.25
                          185.53
                                   0.481
                                          0.6305
Xf1age_2
                 22.22
                          82.52
                                   0.269
                                          0.7877
Xf1married
               1435.28
                         1226.68
                                  1.170
                                          0.2420
```

11

```
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Multiple R-squared: 0.009618, Adjusted R-squared: -9.119e-05
```

For example, the coefficient of Xf0age_1 is the effect of age for the control on re78 when age ≤ 23 and Xf0age_2 is the effect when age ≥ 23 . Note that the R^2 and adjusted R^2 are different from what we obtain using the summary of the 1m object:

```
summary(fit$lm.out)[c("r.squared","adj.r.squared")]
$r.squared
[1] 0.4379272
```

\$adj.r.squared
[1] 0.4316295

This is because our model does not explicitly contain an intercept and the R^2 is computed differently for models without an intercept. But, assuming that our model does not have an intercept is wrong since it includes both Z and (1-Z).

2.4 The predict and plot methods

The predict method is very similar to the predict.lm method. We use the same arguments: object, interval, se.fit, newdata and level. The difference is that it returns the predicted outcome for the treated and nontreated separately, and the argument vcov. provides a way of changing how the least squares covariance matrix is computed. By default, it is computed using vcovHC from the sandwich package (Zeileis (2006)). The function returns a list of 2 elements, treated and nontreated. By default (se.fit=FALSE and interval="none"), each element contains a vector of predictions. Here is an example with the previously fitted model fit:

```
predict(fit,
    newdata = data.frame(treat = c(1,1,0,0),age = 20:23, married = 1))
```

\$treated

[1] 6975.337 7064.591

\$nontreated

[1] 5054.036 5081.834

If interval is set to "confidence", but se.fit remains equal to FALSE, each element contains a matrix containing the prediction, and the lower and upper confidence limits, with the confidence level determined by the argument level (set to 0.95 by default). Here is an example with the same fitted model:

```
predict(fit,
    newdata = data.frame(treat = c(1,1,0,0),age = 20:23, married = 1),
    interval = "confidence")
```

\$treated

```
fit lower upper
1 6975.337 4646.673 9304.001
2 7064.591 4741.653 9387.528
$nontreated
```

fit lower upper 3 5054.036 3574.096 6533.975 4 5081.834 3544.849 6618.820

If se.fit is set to TRUE, each element, treated or nontreated, is a list with the elements pr, containing the predictions, and se.fit, containing the standard errors. In the following, we only show the result for the same fitted model:

\$treated
\$treated\$fit

[1] 6975.337 7064.591

\$treated\$se.fit 1 2 1188.116 1185.194

\$nontreated
\$nontreated\$fit
[1] 5054.036 5081.834

The predict method is called by the plot method to visually assess the predicted outcome for the treated and nontreated with respect to a given confounder, controlling for the other variables in the model. The arguments of the plot method are:

- x: An object of class slseFit.
- y: An alias for which for compatibility with the generic plot function.
- which: confounder to plot against the outcome variable. It could be an integer (the position of the confounder) or a character (the name of the confounder)
- interval: The type of confidence interval to display. The default is "none". The alternative is "confidence".
- level: The confidence level when interval="confidence". The default is 0.95.
- fixedCov: Optional named lists of fixed values for some or all other confounders in each group. The values of the confounders not specified are determined by the argument FUN. To fix some covariates for both groups, fixedCov is just a named list with the names being the variable names. To fix them to different values for the treated and nontreated, fixedCov is a named list of 1 or 2 elements (for the treated, nontreated or both), each element being a named list of values for the covariates. See the examples below.
- vcov.: An optional function to compute the estimated matrix of covariance of the least squares estimators. This argument only affects the confidence intervals. The default is vcovHC with type="HC3".
- add: Should the curves be added to an existing plot? The default is FALSE.
- addToLegend: An optional character string to add to the legend next to "treated" and "nontreated".
- addPoints: Should we include the scatterplot of the outcome and confounder to the graph? The default is FALSE.
- **FUN**: A function to determine how the other confounders are fixed. The default is **mean**. Note that the function is applied to each group separately.
- plot: By default, the method produces a graph. Alternatively, we can set this argument to FALSE and it returns one data.frame per group with the variable selected by which and the prediction.
- graphPar: A list of graphical parameters if not satisfied with the default ones.
- ...: Other arguments are passed to the vcov. function. For example, it is possible to change the type of vcovHC from the default HC3 to any available methods included in the sandwich package (Zeileis (2006)).

The default set of graphical parameters can be obtained by running the function <code>causalSLSE:::.initPar()</code>. The function returns a list of four elements: <code>treated</code>, <code>nontreated</code>, <code>common</code>, <code>legend</code>. The first two are lists of two elements: <code>points</code> for the list of parameters of the scatterplot produced when <code>addPoints=TRUE</code> and <code>lines</code> for the line parameters. For example, we can see that the type of points for the treated is initially set to <code>pch=21</code> and their colour to 2:

```
causalSLSE:::.initPar()$treated$points

$pch
[1] 21
```

\$col [1] 2

The element common are for parameters not specific to a group like the main title or the axis labels and legend are the parameters that control the legend. Note, however, that the colour and line shapes for the legend are automatically determined by the lines and points parameters of the treated and nontreated elements.

The default parameters can be modified by the argument graphPar. This argument must follow the structure of causalSLSE:::.initPar(). For example, if we want a new title, new x-axis label, new type of lines for the treated, new type of points for the nontreated and a different position for the legend, we create the following graphPar:

In the following, we illustrate some examples.

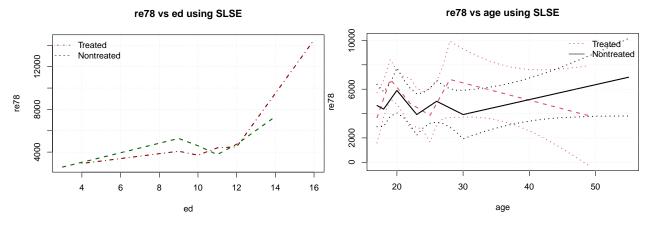
2.4.1 Plot examples

Example 1:

Consider the model:

```
model1 <- cslseModel(re78 ~ treat | ~ age + re75 + ed + married, data = nsw)
fit1 <- estSLSE(model1)</pre>
```

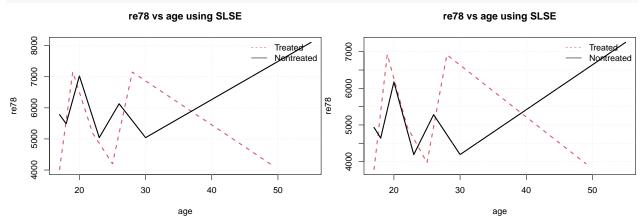
Suppose we want to compare the predicted income between the two treatment groups with respect to age or education, holding the other variables fixed to their group means (the default). The following are two examples with some of the default arguments modified. Note that vcov.lm is used in the first plot function and vcovHC (the default) of type HC1 in the second plot.



Example 2:

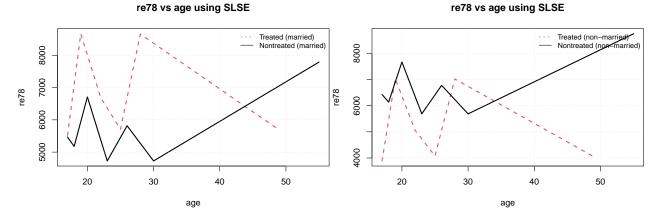
If we want to fix the other confounders using another function, we can change the argument FUN. The new function must be a function of one argument. For example, if we want to fix the other confounders to their group medians, we set FUN to median (no quotes). We proceed the same way for any function that requires only one argument. If the function requires more than one argument, we have to create a new function. For example, if we want to fix them to their 20% group empirical quantiles, we can set the argument to function(x) quantile(x, .20). The following illustrates the two cases:





Example 3:

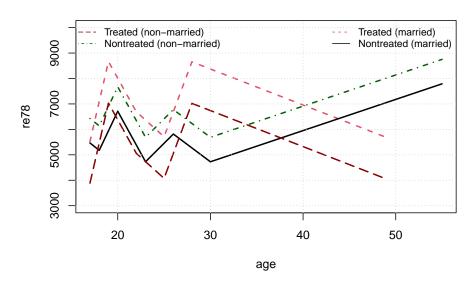
It is also possible to set some of the other confounders to a specific value by changing the argument fixedCov. To fix some variables to the same values for both groups, fixedCov must be a named list with the names corresponding to the variables you want to fix. You can also add a description to the legend with the argument addToLegend.



Example 4:

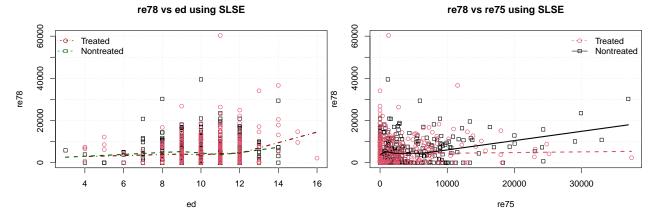
To better compare the two groups, it is also possible to have them plotted on the same graph by setting the argument add. to TRUE. We just need to adjust some of the arguments to better distinguish the different curves. In the following example, we set the colors and line shapes to different values and change the position of the legend for the second set of lines.

re78 vs age using SLSE



Example 5:

Finally, it is also possible to add the observed points to the graph.



${\bf 2.4.2} \quad {\bf Factors, interactions \ and \ functions \ of \ confounders}$

The package allows some of the confounders to be factors, functions of other confounders or interactions. For example, the dataset simDat4 includes one factor, X2, with levels equal to "first", "second" and "third". We can include this confounder directly to the list of confounders. For example,

```
data(simDat4)
mod <- cslseModel(Y ~ Z | ~ X1 + X2 + X4, data = simDat4)
mod</pre>
```

```
Causal Semiparametric LSE Model
*****************************

Number of treated: 246

Number of nontreated: 254
Selection method: Default
Confounders approximated by SLSE:
    treated: X1
    nontreated: X1
Confounders not approximated by SLSE:
    treated: X2second, X2third, X4
    nontreated: X2second, X2third, X4
```

We see that R has created 2 binary variables, one for X2="second" and one for X2="third". These two variables are automatically included in the group of confounders not approximated by SLSE because they are binary variables like X4. If we want to plot Y against X1, the binary variables X2second, X2third and X4 are fixed to their group averages which, in case of binary variables, represent the proportions of ones in each group.

For interaction terms or functions of confounders, FUN is applied to the functions of confounders. This is how we have to proceed to obtain the average prediction in regression models. For example, if we interact X2 and X4, we obtain:

```
data(simDat4)
mod <- cslseModel(Y ~ Z | ~ X1 + X2 * X4, data = simDat4)
mod

Causal Semiparametric LSE Model
********************************

Number of treated: 246
Number of nontreated: 254
Selection method: Default
Confounders approximated by SLSE:
    treated: X1
    nontreated: X1
    Confounders not approximated by SLSE:
    treated: X2second, X2third, X4, X2second:X4, X2third:X4
    nontreated: X2second, X2third, X4, X2second:X4, X2third:X4</pre>
```

In this case, when FUN=mean, X2second: X4 is replaced by the proportion of ones in X2second \times X4 for each group. It is not replaced by the proportion of ones in X2second times the proportion of ones in X4. The same applies to functions of confounders. For functions of confounders, which can be defined in the formula using a built-in function like log or using the identity function I() (e.g. we can interact X1 and X4 by using I(X1*X4)), FUN is applied to the function (e.g. the average log(X) or the average I(X1*X4)).

To fix a factor to a specific level, we just set its value in the fixedCov. In the following example, we fix X2 to "first", so X2second and X2third are set to 0.

```
fit <- estSLSE(mod)
plot(fit, "X1", fixedCov = list(X2 = "first"))</pre>
```

Note that if a function of confounders (or an interaction) involves the confounder we want to plot the outcome against, we factorize the confounder out, apply FUN to the remaining of the function and add the confounder back. For example, if we interact X1 with X4 and FUN=mean, X1:X4 is replaced by X1 times the proportion of ones in X4 for each group.

2.5 Optimal selection of the knots

We have implemented two methods for selecting the knots: the backward semiparametric LSE (BLSE) and the forward semiparametric LSE (FLSE) methods. For each method, we have 3 criteria: the p-value threshold (PVT), the Akaike Information criterion (AIC), and the Bayesian Information criterion (BIC). The two selection methods can be summarized as follows:

BLSE:

- 1. We estimate the model with all knots included in the model.
- 2. For each knot, we test if the slopes of the basis functions adjacent to the knot are the same, and return the p-value.
- 3. The knots are selected using one of the following criteria
 - PVT: We remove all knots with a p-value greater than a specified threshold.
 - AIC or BIC: We order the p-values in descending order. Then, going from the largest to the smallest, we remove the knot associated with the p-value one by one, estimate the model and return the information criterion. We keep the model with the smallest information citerion.

FLSE:

- 1. We estimate the model by including a subset of the knots, one variable at the time. When we test a knot for one confounder, the number of knots is set to 0 for all other variables.
- 2. For each knot, we test if the adjacent slopes to the knot is the same, and return the p-value. The set of knots used for each test depends on the following:
 - Variables with 1 knot: we return the p-value of the test of equality of the slopes adjacent to the knot.
 - Variables with 2 knots: we include the two knots and return the p-values of the test of equality of the slopes adjacent to each knot.
 - Variables with p knots (p > 2): We test the equality of the slopes adjacent to knot i, for i = 1, ..., p, using the sets of knots $\{1, 2\}, \{1, 2, 3\}, \{2, 3, 4\}, ..., \{p 2, p 1, p\}$ and $\{p 1, p\}$ respectively.
- 3. The knots are selected using one of the following criteria
 - PVT: We remove all knots with a p-value greater than a specified threshold.
 - AIC or BIC: We order the p-values in ascending order. Then, starting with a model with no knots and going from the smallest to the highest highest p-value, we add the knot associated with the smallest remaining p-value one by one, estimate the model and return the information criterion. We keep the model with the smallest information citerion.

The knot selection is done using the selSLSE method. The arguments are:

- model: An object of class cslseModel.
- selType: This is the selection method. We have the choice between "FLSE" and "BLSE" (the default).
- **selCrit**: This is the criterion used by the selection method. We have the choice between "AIC" (the default), "BIC" or "PVT".
- pvalT: This is a function that returns the p-value threshold. It is a function of one argument, the average number of basis functions per confounder. The default is function(p) 1/log(p) and it is applied to each group separately. Therefore, the threshold may be different for the treated and non-treated. It is also possible to set it to a fix threshold. For example, function(p) 0.20 sets the threshold to 0.2. This argument affects the result only when selCrit is set to "PVT".
- vcov.: An optional function to compute the least squares standard errors. By default, the p-values are computed using the vcovHC method from the sandwich package with type="HC3" (Zeileis (2006)).
- ...: This is used to pass arguments to the vcov. function.

The function returns a model of class cslseModel with the optimal selection of knots. For example, we can compare the starting knots of model1, with the model selected by the default arguments.

```
print(model1, which = "selKnots")
Causal Semiparametric LSE Model: Selected knots
*************
Selection method: Default
treated
*****
Covariates with no knots:
   married
Covariates with knots:
     20% 40% 60% 80%
Knots 19 22 25 28
re75 :
      40% 60% 80%
Knots 357.9 1962 5589
     20% 40% 60% 80%
Knots 9 10 11 12
nontreated
Covariates with no knots:
  married
Covariates with knots:
    16.67% 33.33% 50% 66.67% 83.33%
Knots 18 20 23 26 30
re75 :
      50% 66.67% 83.33%
Knots 823.3 2292 6567
    16.67% 33.33% 66.67% 83.33%
Knots 9 10 11
model2 <- selSLSE(model1)</pre>
print(model2, which = "selKnots")
Causal Semiparametric LSE Model: Selected knots
**************
Selection method: BLSE
Criterion: AIC
treated
*****
Covariates with no knots:
  re75, married
Covariates with knots:
    20% 60% 80%
Knots 19 25 28
ed:
     80%
Knots 12
```

For example, the BLSE-AIC method has removed all knots from re75 for the treated and kept two knots for the nontreated. The print method indicates which method was used to select the knots. It is possible to recover the p-values of all original knots by setting the argument which to Pvalue.

```
print(model2, which="Pvalues")
```

```
treated
*****
Covariates with no knots:
   married
Covariates with knots:
age :
            20%
                   40%
                           60%
Knots 19.00000 22.0000 25.0000 28.0000
P-Value 0.03472 0.7597 0.3167 0.2257
re75 :
            40%
                     60%
                               80%
Knots 357.9499 1961.8640 5588.6640
P-Value 0.8648
                0.7035
                           0.8177
                 40%
                         60%
          20%
Knots 9.0000 10.0000 11.0000 12.0000
P-Value 0.6366 0.5837 0.8017 0.3286
******
Covariates with no knots:
   {\tt married}
Covariates with knots:
age :
        16.67% 33.33%
                         50% 66.67% 83.33%
Knots 18.0000 20.0000 23.0000 26.0000 30.0000
P-Value 0.4466 0.1067 0.1567 0.3234 0.2884
re75 :
            50%
                 66.67%
                            83.33%
Knots 823.2544 2292.1710 6567.3290
P-Value 0.3261 0.6295
ed:
       16.67% 33.33% 66.67% 83.33%
Knots 9.0000 10.0000 11.0000 12.0000
P-Value 0.2813 0.9158 0.1879 0.8806
```

In the following example, we see BLSE as selection method and BIC as criterion. Note that the BIC selects 0 knots for all confounders.

```
model3 <- selSLSE(model1, selType = "BLSE", selCrit = "BIC")</pre>
model3
Causal Semiparametric LSE Model
Number of treated: 297
Number of nontreated: 425
Selection method: BLSE
Criterion: BIC
Confounders approximated by SLSE:
   treated: None
   nontreated: None
Confounders not approximated by SLSE:
   treated: age, re75, ed, married
   nontreated: age, re75, ed, married
Since the selSLSE method returns a new model, we can apply the estSLSE to it:
estSLSE(selSLSE(model1, selType = "FLSE", selCrit = "BIC"))
Semiparametric LSE
*******
Selection method: FLSE
Criterion: BIC
factor(treat)0 factor(treat)1
                                       Xf0age
                                                      Xf0re75
                                                                        Xf0ed
    4.826e+03
               -3.890e+02
                                   -2.011e+01
                                                    2.982e-01
                                                                    2.500e+00
   Xf0married
                                      Xf1re75
                       Xf1age
                                                       Xf1ed
                                                                   Xf1married
   -1.094e+03
                    4.105e+01
                                    2.676e-02
                                                    4.849e+02
                                                                    1.417e+03
```

2.6 The causalSLSE method for slseFit objects

The method causalSLSE estimates the causal effects from slseFit objects using the knots included in the estimated model. The arguments of the method are:

- object: An object of class slseFit.
- causal: What causality measure should the function compute? We have the choice between "ALL" (the default), "ACE", "ACT" or "ACN".
- vcov.: An alternative function used to compute the covariance matrix of the least squares estimates. This is the $\hat{\Sigma}_{\hat{\theta}}$ defined in the Introduction section. By default, vcovHC is used with type="HC3". Simulations show that using vcovHC with type="HC3" produces the most accurate estimate of the variance of ACE, ACT and ACN in small and large samples.
- ...: This is used to pass arguments to the vcov. function.

In the following example, we estimate the causal effect with the initial knots (without selection).

```
model1 <- cslseModel(re78 ~ treat | ~ age + re75 + ed + married, data=nsw)
fit1 <- estSLSE(model1)
causalSLSE(fit1)</pre>
```

```
Causal Effect using Semiparametric LSE
*******************
Selection method: Default

ACE = 814.3
ACT = 831.9
```

ACN = 802

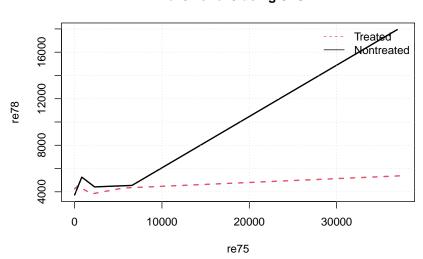
We see that the selection method used to select the knots are set to Default because the knots were first selected by the default method and no additional selection method was used. The method returns an object of class cslse and its print method only prints the causal effect estimates. For more details about the estimation, which includes standard errors and significance tests, we can use the summary method:

```
ce <- causalSLSE(fit1)</pre>
sce <- summary(ce)</pre>
sce
Causal Effect using Semiparametric LSE
Selection method: Default
    Estimate Std. Error t value Pr(>|t|)
ACE
                                      0.108
       814.3
                   506.4
                            1.608
                            1.579
ACT
       831.9
                   526.8
                                      0.114
ACN
       802.0
                   514.8
                            1.558
                                      0.119
```

The summary method returns an object of class summary.cslse and the above output is produced by its print method. If needed, we can extract the above table using \$causal, the least squares coefficients table using \$beta or the list of knots using \$knots.

The cslse object inherits from the class slseFit, so we can apply the plot (or the predict) method directly on this object as shown below:

```
plot(ce, "re75")
```



re78 vs re75 using SLSE

2.6.1 The extract method

The package comes with an extract method for objects of class cslse, which is a required method for creating Latex tables using the texreg package (Leifeld (2013)). For example, we can compare different methods in a single table. In the following example, we compare the SLSE, BLSE-AIC and FLSE-AIC:

```
library(texreg)
c1 <- causalSLSE(fit1)
fit2 <- estSLSE(selSLSE(model1, selType="BLSE"))
fit3 <- estSLSE(selSLSE(model1, selType="FLSE"))
c2 <- causalSLSE(fit2)
c3 <- causalSLSE(fit3)
texreg(list(SLSE=c1, BLSE=c2, FLSE=c3), table=FALSE, digits=4)</pre>
```

	SLSE	BLSE	FLSE
ACE	814.3083	818.1598	824.4901
	(506.3839)	(497.7885)	(496.4032)
ACT	831.8856	837.0768	852.4659
	(526.8465)	(519.2572)	(513.7079)
ACN	802.0249	804.9401	804.9401
	(514.7871)	(503.3558)	(502.7779)
Num. knots (Nontreated)	12	8	6
Num. knots (Treated)	11	4	4
Num. confounders	4	4	4
Num. obs. (Nontreated)	425	425	425
Num. obs. (Treated)	297	297	297
\mathbb{R}^2	0.0869	0.0852	0.0840
Adj. R ²	0.0445	0.0577	0.0592
$***_n < 0.001 \cdot **_n < 0.01 \cdot *_n < 0.05$			

p < 0.001; **p < 0.01; *p < 0.05

Note that we refer to SLSE when no optimal selection method is used, which includes the Default and User Based methods. The option table=FALSE, from the texreg package, is used to remove the Latex floating table environment. With this option, the table appears right after the code instead of being placed somewhere else by Latex. The arguments of the extract methods, which control what is printed and can be modified through the texreg function, are:

- include.nobs: Include the number of observations. The default is TRUE.
- include.nknots: Include the number of knots. The default is TRUE.
- include.rsquared: Include the R^2 . The default is TRUE.
- include.adjrs: Include the adjusted R^2 . The default is TRUE.
- which: Which causal effects should be printed? The options are "ALL" (the default), "ACE", "ACT", "ACN", "ACE-ACT", "ACE-ACN" or "ACT-ACN".

Here is one example on how to change some arguments:

texreg(list(SLSE=c1, BLSE=c2, FLSE=c3), table=FALSE, which="ACE-ACT", include.adjrs=FALSE)

	SLSE	BLSE	FLSE
ACE	814.31	818.16	824.49
	(506.38)	(497.79)	(496.40)
ACT	831.89	837.08	852.47
	(526.85)	(519.26)	(513.71)
Num. knots (Nontreated)	12	8	6
Num. knots (Treated)	11	4	4
Num. confounders	4	4	4
Num. obs. (Nontreated)	425	425	425
Num. obs. (Treated)	297	297	297
\mathbb{R}^2	0.09	0.09	0.08
*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$			

The causalSLSE method for cslseModel objects

When applied directly to cslseModel objects, the causalSLSE method offers the possibility to select the knots and estimate the causal effects all at once. The method also returns an object of class cslse. The arguments are the same as the method for slseFit objects, plus the necessary arguments for the knots selection. The following are the arguments not already defined for objects of class slseFit. The details of these arguments are presented in Section 2.5.

• object: An object of class cslseModel.

2.7

- selType: This is the selection method. We have the choice between "SLSE" (the default), "FLSE" and "BLSE". The SLSE method performs no selection, so all knots from the model are kept.
- selCrit: This is the criterion used by the selection method when selType is set to "FLSE" or "BLSE". The default is "AIC".

• pvalT: This is a function that returns the p-value threshold. We explained this argument when we presented the selSLSE method.

For example, we can generate the previous table as follows.

```
c1 <- causalSLSE(model1, selType="SLSE")
c2 <- causalSLSE(model1, selType="BLSE")
c3 <- causalSLSE(model1, selType="FLSE")
texreg(list(SLSE=c1, BLSE=c2, FLSE=c3), table=FALSE, digits=4)</pre>
```

	SLSE	BLSE	FLSE
ACE	814.3083	818.1598	824.4901
	(506.3839)	(497.7885)	(496.4032)
ACT	831.8856	837.0768	852.4659
	(526.8465)	(519.2572)	(513.7079)
ACN	802.0249	804.9401	804.9401
	(514.7871)	(503.3558)	(502.7779)
Num. knots (Nontreated)	12	8	6
Num. knots (Treated)	11	4	4
Num. confounders	4	4	4
Num. obs. (Nontreated)	425	425	425
Num. obs. (Treated)	297	297	297
\mathbb{R}^2	0.0869	0.0852	0.0840
$Adj. R^2$	0.0445	0.0577	0.0592
$***_{n} < 0.001 \cdot **_{n} < 0.01 \cdot *_{n} < 0.05$			

^{***}p < 0.001; **p < 0.01; *p < 0.05

2.8 The causalSLSE method for formula objects

This last method, offers an alternative way of estimating the causal effects. It allows the estimation in one step without having to first create a model. The arguments are the same as for the cslseModel function and the causalSLSE method for cslseModel objects. It creates the model, selects the knots and estimates the causal effects in one step. For example, we can create the previous table as follows:

	SLSE	BLSE	FLSE
ACE	814.3083	818.1598	824.4901
	(506.3839)	(497.7885)	(496.4032)
ACT	831.8856	837.0768	852.4659
	(526.8465)	(519.2572)	(513.7079)
ACN	802.0249	804.9401	804.9401
	(514.7871)	(503.3558)	(502.7779)
Num. knots (Nontreated)	12	8	6
Num. knots (Treated)	11	4	4
Num. confounders	4	4	4
Num. obs. (Nontreated)	425	425	425
Num. obs. (Treated)	297	297	297
\mathbb{R}^2	0.0869	0.0852	0.0840
$Adj. R^2$	0.0445	0.0577	0.0592
$^{***}p < 0.001; ^{**}p < 0.01; ^{*}p < 0.05$			

Note that this method calls cslseModel, selSLSE, estSLSE and the method causalSLSE for slseFit objects sequentially. It is easier to simply work with this method, but manually going through all steps may be beneficial to better understand the procedure. Also, it is more convenient to work with a model when we want to compare the different selection methods, or if we want to compare estimations with different standard errors.

3 Examples

3.1 A simulated data set from Model 1

In the package, the data set datSim1 is generated using the following data generating process with a sample size of 300.

$$Y(0) = 1 + X + X^{2} + \epsilon(0)$$

 $Y(1) = 1 - 2X + \epsilon(1)$
 $Z = \text{Bernoulli}[\Lambda(1 + X)]$
 $Y = Y(1)Z + Y(0)(1 - Z)$

where X, $\epsilon(0)$ and $\epsilon(1)$ are independent standard normal and $\Lambda(x)$ is the CDF of the standard logistic distribution. The causal effects ACE, ACT and ACN are approximately equal to -1, -1.6903 and 0.5867 (estimated using a sample size of 10^7). We can start by building the starting model:

```
data(simDat1)
mod <- cslseModel(Y ~ Z | ~ X, data = simDat1)</pre>
```

Then we can compare three different methods:

```
c1 <- causalSLSE(mod, selType = "SLSE")
c2 <- causalSLSE(mod, selType = "BLSE", selCrit = "BIC")
c3 <- causalSLSE(mod, selType = "FLSE", selCrit = "BIC")
texreg(list(SLSE = c1, BLSE = c2, FLSE = c3), table = FALSE, digits = 4)</pre>
```

	SLSE	BLSE	FLSE
ACE	-1.4396***	-1.4530***	-1.4533***
	(0.2770)	(0.2742)	(0.2700)
ACT	-1.9316***	-1.9316***	-1.9320***
	(0.3180)	(0.3176)	(0.3131)
ACN	-0.0865	-0.1369	-0.1369
	(0.3338)	(0.3271)	(0.3254)
Num. knots (Nontreated)	2	2	1
Num. knots (Treated)	4	0	0
Num. confounders	1	1	1
Num. obs. (Nontreated)	80	80	80
Num. obs. (Treated)	220	220	220
\mathbb{R}^2	0.7434	0.7386	0.7303
$Adj. R^2$	0.7354	0.7342	0.7266
*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$			

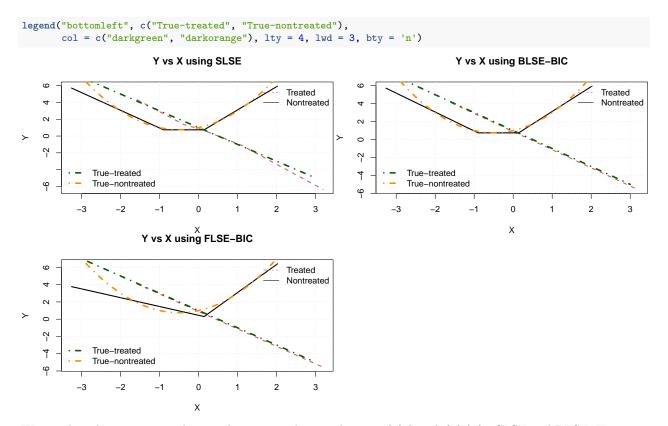
We see that both selection methods choose to assign 0 knots for the treated group, which is not surprising since the true $f_1(x)$ is linear. We can compare the different fits.

```
list(common = list(main = "Y vs X using BLSE-BIC"))
```

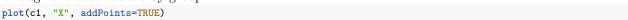
```
$common
```

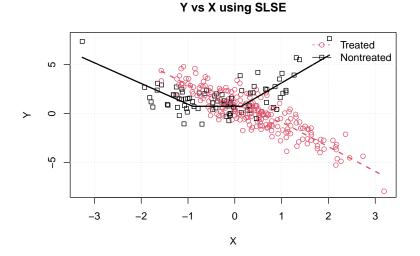
\$common\$main

```
[1] "Y vs X using BLSE-BIC" \,
```



We see that the piecewise polynomials are very close to the true $f_1(x)$ and $f_2(x)$ for SLSE and BLSE. However, the FLSE based on BIC does not do a good job. We can see from the following graph how the lines are fit through the observations by group.





3.2 A simulated data set from Model 2

The dataset datSim2 is a change point regression model (with unknown location of change points) defined as follows:

$$\begin{array}{lcl} Y(0) & = & (1+X)I(X \leq -1) + (-1-X)I(X > -1) + \epsilon(0) \\ Y(1) & = & (1-2X)I(X \leq 0) + (1+2X)I(X > 0) + \epsilon(1) \\ Z & = & \mathrm{Bernoulli}[\Lambda(1+X)] \\ Y & = & Y(1)Z + Y(0)(1-Z) \end{array}$$

where I(A) is the indicator function equal to 1 if A is true, and X, $\epsilon(0)$ and $\epsilon(1)$ are independent standard normal. The causal effects ACE, ACT and ACN are approximately equal to 3.763, 3.858 and 3.545 (estimated with a sample size of 10^7). We can compare the SLSE, BLSE-AIC and BLSE-BIC.

```
data(simDat2)
mod <- cslseModel(Y-Z | -X, data=simDat2)

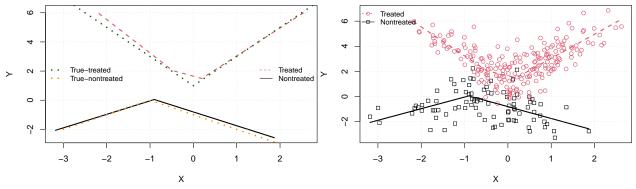
c1 <- causalSLSE(mod, selType = "SLSE")
c2 <- causalSLSE(mod, selType = "BLSE", selCrit = "BIC")
c3 <- causalSLSE(mod, selType = "BLSE", selCrit = "AIC")
texreg(list(SLSE = c1, BLSE.BIC = c2, BLSE.AIC = c3), table = FALSE, digits = 4)</pre>
```

	SLSE	BLSE.BIC	BLSE.AIC	
ACE	3.9290***	3.9201***	3.9201***	
	(0.1772)	(0.1756)	(0.1756)	
ACT	3.9552***	3.9404***	3.9404***	
	(0.1966)	(0.1938)	(0.1938)	
ACN	3.8670***	3.8721***	3.8721***	
	(0.2352)	(0.2337)	(0.2337)	
Num. knots (Nontreated)	2	1	1	
Num. knots (Treated)	3	2	2	
Num. confounders	1	1	1	
Num. obs. (Nontreated)	89	89	89	
Num. obs. (Treated)	211	211	211	
\mathbb{R}^2	0.7833	0.7829	0.7829	
Adj. \mathbb{R}^2	0.7774	0.7784	0.7784	
p = 10001; p = 10001	*** $p < 0.001;$ ** $p < 0.01;$ * $p < 0.05$			

The following shows the fit of BLSE-AIC with the true $f_1(x)$ and $f_0(x)$, and the observations.



Y vs X using BLSE-AIC



3.3 A simulated data set from Model 3

The data set datSim3 is generated from model with multiple confounders defined as follows:

$$\begin{array}{lll} Y(0) & = & [1+X_1+X_1^2] + [(1+X_2)I(X_2 \leq -1) + (-1-X_2)I(X_2 > -1)] + \epsilon(0) \\ Y(1) & = & [1-2X_1] + [(1-2X_2)I(X_2 \leq 0) + (1+2X_2)I(X_2 > 0)] + \epsilon(1) \\ Z & = & \mathrm{Bernoulli}[\Lambda(1+X_1+X_2)] \\ Y & = & Y(1)Z + Y(0)(1-Z) \,, \end{array}$$

where X_1 , X_2 , $\epsilon(0)$ and $\epsilon(1)$ are independent standard normal. The causal effects ACE, ACT and ACN are approximately equal to 2.762, 2.204 and 3.922 (estimated with a sample size of 10^7). We can compare the SLSE, FLSE with AIC and FLSE with BIC.

```
data(simDat3)
mod <- cslseModel(Y ~ Z | ~ X1 + X2, data = simDat3)

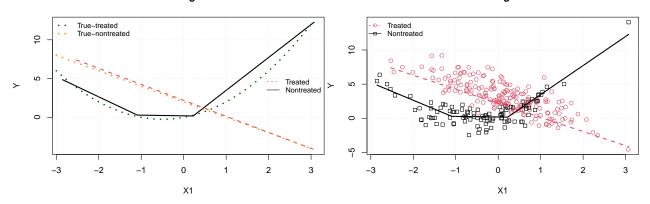
c1 <- causalSLSE(mod, selType = "SLSE")
c2 <- causalSLSE(mod, selType = "FLSE", selCrit = "BIC")
c3 <- causalSLSE(mod, selType = "FLSE", selCrit = "AIC")
texreg(list(SLSE = c1, FLSE.BIC = c2, FLSE.AIC = c3), table = FALSE, digits = 4)</pre>
```

	SLSE	FLSE.BIC	FLSE.AIC
ACE	2.4699***	2.4866***	2.4725***
	(0.3021)	(0.2997)	(0.2998)
ACT	2.0653***	2.0688***	2.0688***
	(0.3772)	(0.3731)	(0.3731)
ACN	3.2323***	3.2739***	3.2334***
	(0.3540)	(0.3523)	(0.3529)
Num. knots (Nontreated)	6	5	5
Num. knots (Treated)	6	3	4
Num. confounders	2	2	2
Num. obs. (Nontreated)	104	104	104
Num. obs. (Treated)	196	196	196
\mathbb{R}^2	0.8630	0.8614	0.8625
$Adj. R^2$	0.8547	0.8551	0.8558
*** $p < 0.001$; ** $p < 0.01$; * p	< 0.05		

To illustrate the method, since we have two confounders, we need to plot the outcome against one confounder holding the other fixed. The default is to fix it to its sample mean. For the true curve, we fix it to its population mean, which is 0. We first look at the outcome against X_1 . By fixing X_2 to 0, the true curve is $X_1 + X_1^2$ for the untreated and $2 - 2X_1$ for the treated. The following graphs show how the FLSE-BIC method fits the curves.



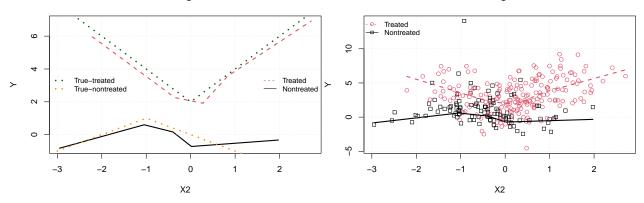
Y vs X1 using FLSE-AIC



If we fix X_1 to 0, the true curve is $1 + [(1 + X_2)I(X_2 \le -1) + (-1 - X_2)I(X_2 > -1)]$ for the nontreated and $1 + [(1 - 2X_2)I(X_2 \le 0) + (1 + 2X_2)I(X_2 > 0)]$ for the treated. The following graphs illustrates how these curves are approximated by FLSE-AIC.

Y vs X2 using FLSE-AIC

Y vs X2 using FLSE-AIC



3.4 A simulated data set with interactions

The data set datSim5 is generated using the following data generating process with a sample size of 300.

$$\begin{array}{lcl} Y(0) & = & \left[1+X_1+X_1^2\right]+\left[(1+X_2)I(X_2\leq -1)+(-1-X_2)I(X_2>-1)\right]\\ & & +\left[1+X_1X_2+(X_1X_2)^2\right]+\epsilon(0)\\ Y(1) & = & \left[1-2X_1\right]+\left[(1-2X_2)I(X_2\leq 0)+(1+2X_2)I(X_2>0)\right]\\ & & +\left[1-2X_1X_2\right]+\epsilon(1)\\ Z & = & \mathrm{Bernoulli}[\Lambda(1+X_1+X_2+X_1X_2)]\\ Y & = & Y(1)Z+Y(0)(1-Z)\,, \end{array}$$

where X_1 , X_2 , e and u are independent standard normal. The causal effects ACE, ACT and ACN are

approximately equal to 1.763, 0.998 and 3.194 (estimated with a sample size of 10^7). We can compare the SLSE, FLSE-AIC and FLSE-BIC.

```
data(simDat5)
mod <- cslseModel(Y ~ Z | ~ X1 * X2, data = simDat5)

c1 <- causalSLSE(mod, selType = "SLSE")
c2 <- causalSLSE(mod, selType = "FLSE", selCrit = "BIC")
c3 <- causalSLSE(mod, selType = "FLSE", selCrit = "AIC")
texreg(list(SLSE = c1, FLSE.BIC = c2, FLSE.AIC = c3), table = FALSE, digits = 4)</pre>
```

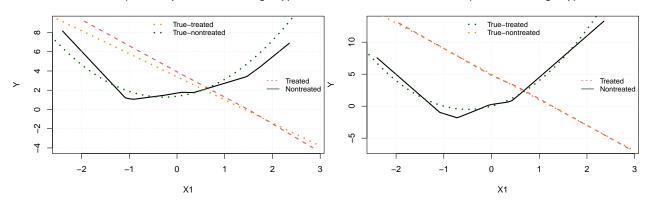
	SLSE	FLSE.BIC	FLSE.AIC
ACE	1.7990***	1.7797***	1.7744***
	(0.3889)	(0.3962)	(0.3940)
ACT	1.2582*	1.2091*	1.2091*
	(0.5029)	(0.5143)	(0.5143)
ACN	2.8183***	2.8550***	2.8399***
	(0.4524)	(0.4535)	(0.4475)
Num. knots (Nontreated)	9	8	8
Num. knots (Treated)	9	5	6
Num. confounders	3	3	3
Num. obs. (Nontreated)	104	104	104
Num. obs. (Treated)	196	196	196
\mathbb{R}^2	0.8909	0.8879	0.8894
$Adj. R^2$	0.8809	0.8799	0.8811
*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$			

In the case of multiple confounders with interactions, the shape of the fitted outcome with respect to one confounder depends on the value of the other confounders. Without interaction, changing the value of the other confounders only shifts the fitted line without changing its shape. The following graphs compare the estimated relationship between Y and X_1 for X_2 equal to the group means (left graph) and 1 (right graph). Using a sample of 10^7 , we obtain that $E(X_2|Z=1)$ and $E(X_2|Z=0)$ are approximately equal to 0.1982 and -0.3698, respectively. Therefore, the true curves are $(1.3698+0.6302x+1.1368x^2)$ for the nontreated and (3.3964-2.3964x) for the treated. If $X_2=1$, the true curves become $2x+2x^2$ for the treated and (5-4x) for the nontreated.

```
x20 <- mean(subset(simDat5, Z == 0)$X2)</pre>
x21 <- mean(subset(simDat5, Z == 1)$X2)
arg <- list(common = list(main = "Y vs X1 (X2 = sample mean for each group)"),</pre>
             legend = list(x = "right", cex = 0.8))
plot(c2, "X1", fixedCov = list(nontreated = list(X2 = x20), treated = list(X2 = x21)),
     graphPar = arg)
curve(1.3698 + 0.6302 * x + 1.1368 * x^2, -3, 3,
     col = "darkgreen", lty = 3, lwd = 3, add = TRUE)
curve(3.3964 - 2.3964 * x, -3, 3, col = "darkorange", lty = 3, lwd = 3, add = TRUE)
legend("top", c("True-treated", "True-nontreated"),
      col=c("darkorange", "darkgreen"), lty = 3, lwd = 3, bty = 'n', cex = .8)
arg <- list(common = list(main = "Y vS X1 (X2 = 1 for each group)"),
            legend = list(x = "right", cex = 0.8))
plot(c2, "X1", fixedCov = list(X2 = 1), graphPar = arg)
curve(2 * x + 2 * x^2, -3, 3, col = "darkgreen", lty = 3, lwd = 3, add = TRUE)
curve(5 - 4 * x, -3, 3, col = "darkorange", lty = 3, lwd = 3, add = TRUE)
legend("top", c("True-treated", "True-nontreated"),
      col = c("darkgreen", "darkorange"), lty = 3, lwd = 3, bty = 'n', cex = .8)
```

Y vs X1 (X2 = sample mean for each group)

Y vS X1 (X2 = 1 for each group)

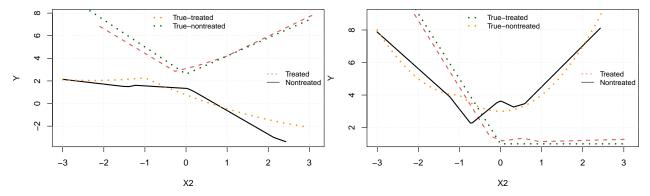


The following graphs illustrate the relationship between Y and X_2 for a given X_1 . When X_1 is equal its population group means (they are equal to the population means of X_2), the true curves are $[1.6036 - 0.3964x)(x \le 0) + (1+2x)(x > 0)]$ for the treated and $[(1.767 - 0.3698x + 0.1368x^2) + (1+x)(x \le -1) + (-1-x)(x > -1)]$ for the nontreated. If $X_1 = 1$, the true curves become $[-2x + (1-2x)(x \le 0) + (1+2x)(x > 0)]$ for the treated and $[(4+x+x^2) + (1+x)(x \le -1) + (-1-x)(x > -1)]$ for the nontreated.

```
x10 <- mean(subset(simDat5, Z == 0)$X1)</pre>
x11 <- mean(subset(simDat5, Z == 1)$X1)</pre>
arg <- list(common = list(main = "Y vs %2 (%1 = sample mean for each group)"),
             legend = list(x = "right", cex = 0.8))
plot(c2, "X2", fixedCov = list(nontreated = list(X1 = x10), treated = list(X1 = x11)),
     graphPar = arg)
curve(1.603900 - .3964 * x + (1 - 2 * x) * (x <= 0) + (1 + 2 * x) * (x > 0), -3, 3,
      col = "darkgreen", lty = 3, lwd = 3, add = TRUE)
curve(1.767 - 0.3698 * x + 0.1368 * x^2 + (1 + x) * (x <= -1) + (-1 - x) * (x > -1),
      -3, 3, col = "darkorange", lty = 3, lwd = 3, add = TRUE)
legend("top", c("True-treated", "True-nontreated"),
       col = c("darkorange", "darkgreen"), lty = 3, lwd = 3, bty = 'n', cex = .8)
arg$common$main <- "Y vS X2 (X1 = 1 for each group)"</pre>
plot(c2, "X2", fixedCov = list(X1 = 1), graphPar = arg)
curve (-2 * x + (1 - 2 * x) * (x \le 0) + (1 + 2 * x) * (x > 0), -3, 3,
      col = "darkgreen", lty = 3, lwd = 3, add = TRUE)
curve(4 + (1 + x) * (x \le -1) + (-1 - x) * (x > -1) + x + x^2,
-3, 3, col = "darkorange", lty = 3, lwd = 3, add = TRUE) legend("top", c("True-treated", "True-nontreated"),
       col = c("darkgreen", "darkorange"), lty = 3, lwd = 3, bty = 'n', cex = .8)
```

Y vs X2 (X1 = sample mean for each group)

Y vS X2 (X1 = 1 for each group)



4 Summary of methods and objects

The following is a list of all objects from the package. For each object, we explain how it is constructed and give a list of the registered methods. For mode details about the arguments of the different methods, see the help files. Note, however, that no help files exist for non-exported methods and the latter must be called using causalSLSE::: before the method names.

- slseKnots: The object is created by the function slseKnots and the only exported registered method is print. The method update, which is used by estSLSE to select knots before estimating the model is not exported.
- cslseKnots: The object is created by the function cslseKnots and it is a list of slseKnots objects. As for slseKnots object, the only exported registered method is print and there is an non-exported method update.
- cslseModel: The object is created by the function cslseModel and the exported registered methods are print, estSLSE (estimate the regression model), selSLSE (optimal selection of knots) and causalSLSE (to compute the causal effects). There are two non-exported methods: pvalSLSE (used to compute the p-values) and model.matrix (to extract the matrix of confounders).
- slseFit: The object is created by the method estSLSE and the exported registered methods are print, causalSLSE (to compute the causal effects), predict (to predict the outcome), plot (to plot the outcome as a function of one confounder) and summary (to give more details about the least squares estimation)
- summary.slseFit: The object is created by the summary method for slseFit objects. The only exported registered method is print.
- cslse: The object is created by any causalSLSE method. It inherits from slseFit object. The methods that are common through this inheritance are plot and predict. The exported registered methods specific to cslse objects are print, summary (to give more details about the causal effect estimation) and extract (a method needed for texreg)

Note that the method causalSLSE is also registered for objects of class formula.

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