R-package FME: inverse modelling, sensitivity, monte carlo - applied to a nonlinear model

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Abstract

Rpackage **FME** (Soetaert 2009) contains functions for model calibration, sensitivity, identifiability, and monte carlo analysis of nonlinear models.

This vignette (vignette("FMEother")), applies the FME functions to a simple non-linear model.

A similar vignette (vignette("FMEdyna")), applies the functions to a dynamic similation model, solved with integration routines from package **deSolve**

A third vignette, (vignette("FMEsteady")), applies **FME** to a partial differential equation, solved with a steady-state solver from package **rootSolve**

vignette("FMEmcmc") tests the markov chain monte carlo (MCMC) implementation

Keywords: steady-state models, differential equations, fitting, sensitivity, Monte Carlo, identifiability, R.

1. Fitting a Monod function

1.1. the model

This example is discussed in (Laine 2008) (who quotes Berthoux and Brown, 2002. Statistics for environmental engineers, CRC Press).

The following model:

$$y = \theta_1 \cdot \frac{x}{x + \theta_2} + \epsilon$$
$$\epsilon \sim N(\theta, I\sigma^2)$$

is fitted to data.

1.2. implementation in R

> require(FME)

First we input the observations

```
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```

```
> Obs <- data.frame(x=c(</pre>
                            28, 55,
                                        83, 110, 138, 225, 375),
                                                                          # mg COD/1
                     y=c(0.053,0.06,0.112,0.105,0.099,0.122,0.125))
                                                                          # 1/hour
The Monod model returns a data.frame, with elements x and y :
> Model <- function(p,x) return(data.frame(x=x,y=p[1]*x/(x+p[2])))</pre>
We first fit the model to the data.
Function Residuals estimates the deviances of model versus the data.
> Residuals <- function(p) (Obs$y-Model(p,Obs$x)$y)
This function is input to modFit which fits the model to the observations.
> print(system.time(
+ P
         <- modFit(f=Residuals,p=c(0.1,1))</pre>
+ ))
   user system elapsed
   0.02
           0.00
                    0.01
We can estimate and print the summary of fit
> sP
        <- summary(P)
> sP
Parameters:
     Estimate Std. Error t value Pr(>|t|)
[1,] 0.14542
                 0.01564
                            9.296 0.000242 ***
[2,] 49.05292
                 17.91196
                            2.739 0.040862 *
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.01278 on 5 degrees of freedom
Parameter correlation:
       [,1]
               [,2]
[1,] 1.0000 0.8926
[2,] 0.8926 1.0000
We also plot the residual sum of squares, the residuals and the best-fit model
> x
         <-0:375
> par(mfrow=c(2,2))
> plot(P,mfrow=NULL)
> plot(Obs,pch=16,cex=2,xlim=c(0,400),ylim=c(0,0.15),
```

xlab="mg COD/1",ylab="1/hr",main="best-fit")

> lines(Model(P\$par,x))
> par(mfrow=c(1,1))

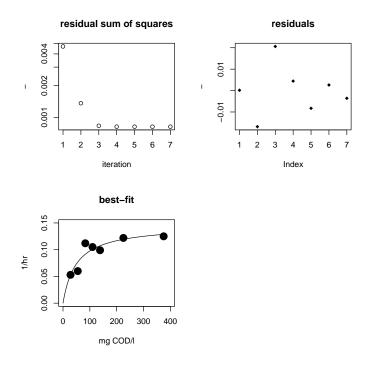


Figure 1: Fit diagnostics of the Monod function - see text for R-code

Finally, we run an MCMC analysis. The -scaled- parameter covariances returned from the summary function are used as estimate of the proposal covariances (jump). Scaling is as in (Gelman, Varlin, Stern, and Rubin 2004).

For the initial model variance (var0) we use the residual mean squares also returned by the summary function. We give equal weight to prior and modeled mean squares (wvar0=1)

The MCMC method adopted here is the metropolis-hastings algorithm; the MCMC is run for 3000 steps; we use the best-fit parameter set (P\$par) to initiate the chain (p). A lower bound (0) is imposed on the parameters (lower).

The plotted results demonstrate (near-) convergence of the chain.

```
> plot(MCMC,Full=TRUE)
```

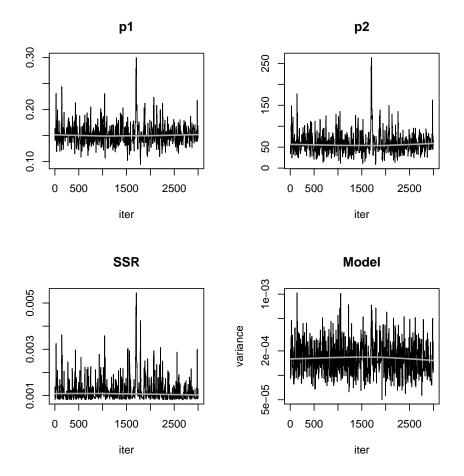


Figure 2: The mcmc - see text for ${\sf R\text{-}code}$

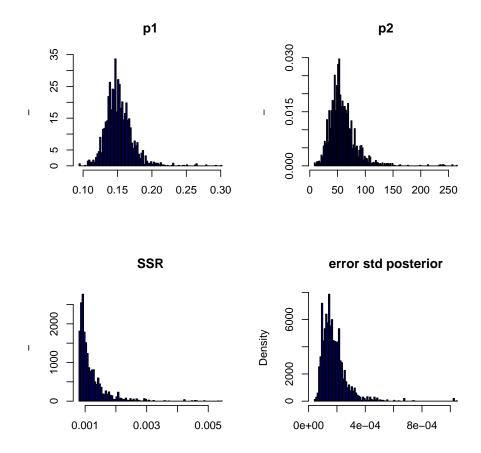


Figure 3: Hist plot - see text for R-code

The posterior distribution of the parameters, the sum of squares and the model's error standard deviation.

> hist(MCMC,Full=TRUE,col="darkblue")

The pairs plot shows the relationship between the two parameters

> pairs(MCMC)

The parameter correlation and covariances from the MCMC results can be calculated and compared with the results obtained by the fitting algorithm.

> cor(MCMC\$pars)

```
p1 p2
p1 1.0000000 0.9242684
p2 0.9242684 1.0000000
```

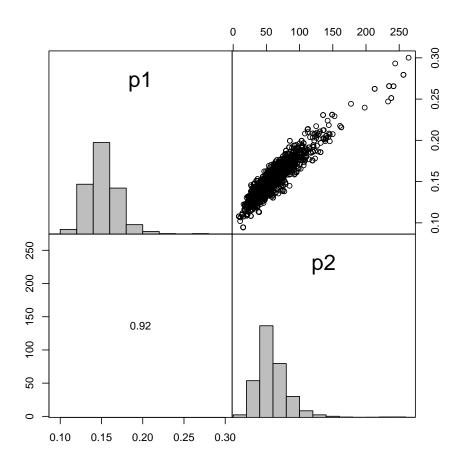


Figure 4: Pairs plot - see text for R-code

```
> cov(MCMC$pars)
```

```
p1 p2
p1 0.0003900295 0.461481
p2 0.4614809662 639.166434
```

> sP\$cov.scaled

```
[,1] [,2]
[1,] 0.0002447075 0.2501157
[2,] 0.2501157147 320.8381590
```

The Raftery and Lewis's diagnostic from package **coda** gives more information on the number of runs that is actually needed. First the MCMC results need to be converted to an object of type mcmc, as used in **coda**.

```
> MC <- as.mcmc(MCMC$pars)
> raftery.diag(MC)

Quantile (q) = 0.025
Accuracy (r) = +/- 0.005
Probability (s) = 0.95
```

You need a sample size of at least 3746 with these values of q, r and s

Also interesting is function cumuplot from coda:

```
> cumuplot(MC)
```

The predictive posterior distribution of the model is easily estimated by running function sensRange, using a randomly selected subset of the parameters in the chain (MCMC\$pars; we use the default of 100 parameter combinations.

```
> sR<-sensRange(parInput=MCMC$pars,func=Model,x=1:375)
```

The distribution is plotted and the data added to the plot:

```
> plot(summary(sR),quant=TRUE)
> points(Obs)
```

By toggling on covariance adaptation (updatecov and delayed rejection (ntrydr), the acceptance rate is increased:

```
> print(system.time(
+ MCMC2 <- modMCMC(f=Residuals,p=P$par,jump=Covar,niter=3000, ntrydr=3,
+ var0=s2prior,wvar0=1,updatecov=100,lower=c(0,0))
+ ))</pre>
```

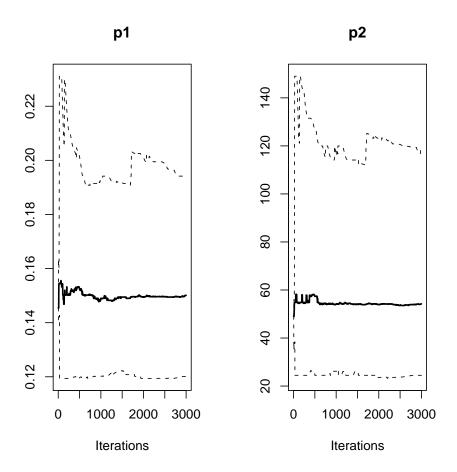


Figure 5: Cumulative quantile plot - see text for $\mathsf{R}\text{-}\mathsf{code}$

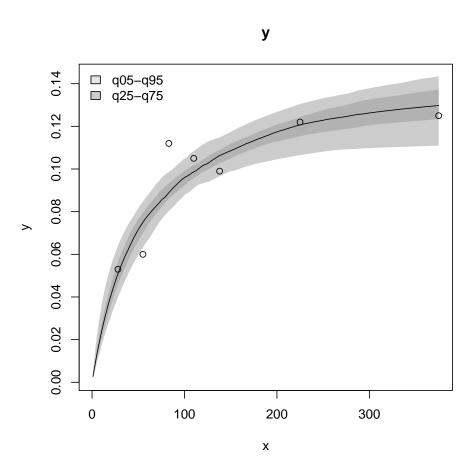


Figure 6: Predictive envelopes of the model - see text for $\mathsf{R}\text{-}\mathsf{code}$

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```
number of accepted runs: 2532 out of 3000 (84.4%) user system elapsed 4.11 0.00 4.18
```

> MCMC2\$count

```
dr_steps Alfasteps num_accepted num_covupdate 2727 11757 2532 30
```

2. finally

This vignette is a Sweave (Leisch 2002) translation of part of the **FME** examples.

References

Gelman A, Varlin JB, Stern HS, Rubin DB (2004). Bayesian Data Analysis, second edition. Chapman and Hall / CRC, Boca Raton.

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